## Gaussian Mixture Models

Julia Kempe & David S. Rosenberg

CDS, NYU

May 7, 2019

### Contents

- Gaussian Mixture Models
- Mixture Models
- 3 Learning in Gaussian Mixture Models
- 4 Issues with MLE for GMM
- 5 The EM Algorithm for GMM



• Let's consider the following generative model (i.e. a way to generate data).

- Let's consider the following generative model (i.e. a way to generate data).
- Suppose
  - **1** There are k clusters (or "**mixture components**").

- Let's consider the following generative model (i.e. a way to generate data).
- Suppose
  - **1** There are *k* clusters (or "**mixture components**").
  - 2 We have a probability density for each cluster.

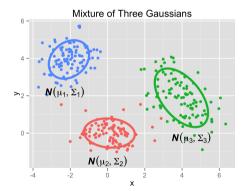
- Let's consider the following generative model (i.e. a way to generate data).
- Suppose
  - **1** There are k clusters (or "mixture components").
  - 2 We have a probability density for each cluster.
- Generate a point as follows
  - ① Choose a random cluster  $z \in \{1, 2, ..., k\}$ .

- Let's consider the following generative model (i.e. a way to generate data).
- Suppose
  - **1** There are k clusters (or "mixture components").
  - 2 We have a probability density for each cluster.
- Generate a point as follows
  - ① Choose a random cluster  $z \in \{1, 2, ..., k\}$ .
  - Choose a point from the distribution for cluster z.

- Let's consider the following generative model (i.e. a way to generate data).
- Suppose
  - **1** There are *k* clusters (or "**mixture components**").
  - 2 We have a probability density for each cluster.
- Generate a point as follows
  - **①** Choose a random cluster  $z \in \{1, 2, ..., k\}$ .
  - 2 Choose a point from the distribution for cluster z.
- Data generated in this way is said to have a mixture distribution.

# Gaussian Mixture Model (k = 3)

- **1** Choose  $z \in \{1, 2, 3\}$  with  $p(1) = p(2) = p(3) = \frac{1}{3}$ .
- **2** Choose  $x \mid z \sim \mathcal{N}(X \mid \mu_z, \Sigma_z)$ .

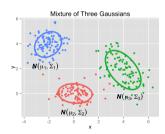


# Gaussian Mixture Model Parameters (k Components)

Cluster probabilities:  $\pi = (\pi_1, \dots, \pi_k)$ 

Cluster means:  $\mu = (\mu_1, ..., \mu_k)$ 

Cluster covariance matrices:  $\Sigma = (\Sigma_1, \dots \Sigma_k)$ 

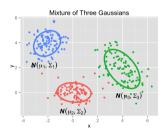


# Gaussian Mixture Model Parameters (k Components)

Cluster probabilities:  $\pi = (\pi_1, \dots, \pi_k)$ 

Cluster means:  $\mu = (\mu_1, ..., \mu_k)$ 

Cluster covariance matrices:  $\Sigma = (\Sigma_1, \dots \Sigma_k)$ 



For now, suppose all these parameters are known.

We'll discuss how to learn or estimate them later.

• Factorize the joint density:

$$p(x,z) = p(z)p(x \mid z)$$

• Factorize the joint density:

$$p(x,z) = p(z)p(x \mid z)$$
  
=  $\pi_z \mathcal{N}(x \mid \mu_z, \Sigma_z)$ 

• Factorize the joint density:

$$p(x,z) = p(z)p(x | z)$$
  
=  $\pi_z \mathcal{N}(x | \mu_z, \Sigma_z)$ 

•  $\pi_z$  is probability of choosing cluster z.

• Factorize the joint density:

$$p(x,z) = p(z)p(x \mid z)$$
  
=  $\pi_z \mathcal{N}(x \mid \mu_z, \Sigma_z)$ 

- $\pi_z$  is probability of choosing cluster z.
- $x \mid z$  has distribution  $\mathcal{N}(\mu_z, \Sigma_z)$ .

• Factorize the joint density:

$$p(x,z) = p(z)p(x \mid z)$$
  
=  $\pi_z \mathcal{N}(x \mid \mu_z, \Sigma_z)$ 

- $\pi_z$  is probability of choosing cluster z.
- $x \mid z$  has distribution  $\mathcal{N}(\mu_z, \Sigma_z)$ .
- z corresponding to x is the true cluster assignment.

Factorize the joint density:

$$p(x,z) = p(z)p(x \mid z)$$
  
=  $\pi_z \mathcal{N}(x \mid \mu_z, \Sigma_z)$ 

- $\pi_z$  is probability of choosing cluster z.
- $x \mid z$  has distribution  $\mathcal{N}(\mu_z, \Sigma_z)$ .
- z corresponding to x is the true cluster assignment.
- Suppose we know the model parameters  $\pi_z$ ,  $\mu_z$ ,  $\Sigma_z$ .

Factorize the joint density:

$$p(x,z) = p(z)p(x \mid z)$$
  
=  $\pi_z \mathcal{N}(x \mid \mu_z, \Sigma_z)$ 

- $\pi_z$  is probability of choosing cluster z.
- $x \mid z$  has distribution  $\mathcal{N}(\mu_z, \Sigma_z)$ .
- z corresponding to x is the true cluster assignment.
- Suppose we know the model parameters  $\pi_z$ ,  $\mu_z$ ,  $\Sigma_z$ .
- Then we can easily evaluate the joint density p(x, z).

- We observe x.
- We don't observe z (the cluster assignment).

- We observe x.
- We don't observe z (the cluster assignment).
- Cluster assignment z is called a hidden variable or latent variable.

- We observe x.
- We don't observe z (the cluster assignment).
- Cluster assignment z is called a hidden variable or latent variable.

#### Definition

A latent variable model is a probability model for which certain variables are never observed.

- We observe x.
- We don't observe z (the cluster assignment).
- Cluster assignment z is called a hidden variable or latent variable.

#### **Definition**

A latent variable model is a probability model for which certain variables are never observed.

e.g. The Gaussian mixture model is a latent variable model.

• We observe x. We want to know its cluster assignment z.

- We observe x. We want to know its cluster assignment z.
- The conditional probability for cluster z given x is

$$p(z \mid x) = p(x, z)/p(x)$$

- We observe x. We want to know its cluster assignment z.
- The conditional probability for cluster z given x is

$$p(z \mid x) = p(x, z)/p(x)$$

• The conditional distribution is a **soft assignment** to clusters.

- We observe x. We want to know its cluster assignment z.
- The conditional probability for cluster z given x is

$$p(z \mid x) = p(x, z)/p(x)$$

- The conditional distribution is a **soft assignment** to clusters.
- A hard assignment is

$$z^* = \underset{z \in \{1, \dots, k\}}{\operatorname{arg\,max}} p(z \mid x).$$

- We observe x. We want to know its cluster assignment z.
- The conditional probability for cluster z given x is

$$p(z \mid x) = p(x, z)/p(x)$$

- The conditional distribution is a **soft assignment** to clusters.
- A hard assignment is

$$z^* = \underset{z \in \{1, \dots, k\}}{\operatorname{arg\,max}} p(z \mid x).$$

• So if we know the model parameters, clustering is trival.

# Mixture Models

• Let S be a set of k probability distributions ("mixture components").

- Let S be a set of k probability distributions ("mixture components").
- Let  $\pi = (\pi_1, ..., \pi_k)$  be a distribution on  $\{1, ..., k\}$  ("mixture weights")

- Let S be a set of k probability distributions ("mixture components").
- Let  $\pi = (\pi_1, ..., \pi_k)$  be a distribution on  $\{1, ..., k\}$  ("mixture weights")
- Suppose we generate *x* with the following procedure:

- Let S be a set of k probability distributions ("mixture components").
- Let  $\pi = (\pi_1, ..., \pi_k)$  be a distribution on  $\{1, ..., k\}$  ("mixture weights")
- Suppose we generate *x* with the following procedure:
  - Choose a distribution randomly from S according to  $\pi$ .

- Let S be a set of k probability distributions ("mixture components").
- Let  $\pi = (\pi_1, ..., \pi_k)$  be a distribution on  $\{1, ..., k\}$  ("mixture weights")
- Suppose we generate *x* with the following procedure:
  - Choose a distribution randomly from S according to  $\pi$ .
  - 2 Sample x from the chosen distribution.

- Let S be a set of k probability distributions ("mixture components").
- Let  $\pi = (\pi_1, ..., \pi_k)$  be a distribution on  $\{1, ..., k\}$  ("mixture weights")
- Suppose we generate *x* with the following procedure:
  - Choose a distribution randomly from S according to  $\pi$ .
  - 2 Sample x from the chosen distribution.
- Then we say x has a mixture distribution.

#### Mixture Densities

- Suppose we have a mixture distribution with
  - mixture components represented as densities  $p_1, \ldots, p_k$ , and

- Suppose we have a mixture distribution with
  - mixture components represented as densities  $p_1, \ldots, p_k$ , and
  - mixture weights  $\pi = (\pi_1, \dots, \pi_k)$ , then

- Suppose we have a mixture distribution with
  - mixture components represented as densities  $p_1, \ldots, p_k$ , and
  - mixture weights  $\pi = (\pi_1, \dots, \pi_k)$ , then
- $\bullet$  the corresponding probability density for x is

$$p(x) = \sum_{i=1}^k \pi_i p_i(x).$$

- Suppose we have a mixture distribution with
  - mixture components represented as densities  $p_1, \ldots, p_k$ , and
  - mixture weights  $\pi = (\pi_1, \dots, \pi_k)$ , then
- $\bullet$  the corresponding probability density for x is

$$p(x) = \sum_{i=1}^k \pi_i p_i(x).$$

Note that p is a convex combination of the mixture component densities.

- Suppose we have a mixture distribution with
  - mixture components represented as densities  $p_1, \ldots, p_k$ , and
  - mixture weights  $\pi = (\pi_1, \dots, \pi_k)$ , then
- $\bullet$  the corresponding probability density for x is

$$p(x) = \sum_{i=1}^k \pi_i p_i(x).$$

- Note that p is a convex combination of the mixture component densities.
- p(x) is called a mixture density.

- Suppose we have a mixture distribution with
  - mixture components represented as densities  $p_1, \ldots, p_k$ , and
  - mixture weights  $\pi = (\pi_1, \dots, \pi_k)$ , then
- ullet the corresponding probability density for x is

$$p(x) = \sum_{i=1}^k \pi_i p_i(x).$$

- Note that p is a convex combination of the mixture component densities.
- p(x) is called a mixture density.
- Conversely, if x has a density of this form, then x has a mixture distribution.

# Gaussian Mixture Model (GMM): Marginal Distribution

#### For example:

ullet The marginal distribution for a single observation x in a GMM is

$$p(x) = \sum_{z=1}^{k} p(x, z)$$

# Gaussian Mixture Model (GMM): Marginal Distribution

#### For example:

• The marginal distribution for a single observation x in a GMM is

$$p(x) = \sum_{z=1}^{k} p(x, z)$$
$$= \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x \mid \mu_{z}, \Sigma_{z})$$

Learning in Gaussian Mixture Models

### The GMM "Learning" Problem

- Given data  $x_1, \ldots, x_n$  drawn from a GMM,
- Estimate the parameters:

Cluster probabilities:  $\pi = (\pi_1, \dots, \pi_k)$ 

Cluster means:  $\mu = (\mu_1, ..., \mu_k)$ 

Cluster covariance matrices:  $\Sigma = (\Sigma_1, \dots \Sigma_k)$ 

## The GMM "Learning" Problem

- Given data  $x_1, \ldots, x_n$  drawn from a GMM,
- Estimate the parameters:

Cluster probabilities: 
$$\pi = (\pi_1, \dots, \pi_k)$$

Cluster means: 
$$\mu = (\mu_1, \dots, \mu_k)$$

Cluster covariance matrices: 
$$\Sigma = (\Sigma_1, \dots \Sigma_k)$$

• Once we have the parameters, we're done.

## The GMM "Learning" Problem

- Given data  $x_1, \ldots, x_n$  drawn from a GMM,
- Estimate the parameters:

Cluster probabilities: 
$$\pi = (\pi_1, \dots, \pi_k)$$

Cluster means: 
$$\mu = (\mu_1, \dots, \mu_k)$$

Cluster covariance matrices: 
$$\Sigma = (\Sigma_1, \dots \Sigma_k)$$

- Once we have the parameters, we're done.
- Just do "inference" to get cluster assignments.

- One approach to learning is maximum likelihood
  - find parameter values with highest likelihood for the observed data.

- One approach to learning is maximum likelihood
  - find parameter values with highest likelihood for the observed data.
- The model likelihood for  $\mathcal{D} = (x_1, \dots, x_n)$  sampled iid from a GMM is

$$L(\pi, \mu, \Sigma) = \prod_{i=1}^{n} p(x_i)$$

- One approach to learning is maximum likelihood
  - find parameter values with highest likelihood for the observed data.
- The model likelihood for  $\mathcal{D} = (x_1, \dots, x_n)$  sampled iid from a GMM is

$$L(\pi, \mu, \Sigma) = \prod_{i=1}^{n} p(x_i)$$
$$= \prod_{i=1}^{n} \sum_{z=1}^{k} \pi_z \mathcal{N}(x_i \mid \mu_z, \Sigma_z).$$

- One approach to learning is maximum likelihood
  - find parameter values with highest likelihood for the observed data.
- The model likelihood for  $\mathcal{D} = (x_1, \dots, x_n)$  sampled iid from a GMM is

$$L(\pi, \mu, \Sigma) = \prod_{i=1}^{n} p(x_i)$$
$$= \prod_{i=1}^{n} \sum_{z=1}^{k} \pi_z \mathcal{N}(x_i \mid \mu_z, \Sigma_z).$$

• As usual, we'll take our objective function to be the log of this:

$$J(\pi, \mu, \Sigma) = \sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}$$

• Recall that the density for  $x \sim \mathcal{N}(\mu, \Sigma)$  is

$$p(x \mid \mu, \Sigma) = \frac{1}{\sqrt{|2\pi\Sigma|}} \exp\left(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right)$$

• Recall that the density for  $x \sim \mathcal{N}(\mu, \Sigma)$  is

$$p(x \mid \mu, \Sigma) = \frac{1}{\sqrt{|2\pi\Sigma|}} \exp\left(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right)$$

• And the log-density is

$$\log p(x \mid \mu, \Sigma) = -\frac{1}{2} \log |2\pi\Sigma| - \frac{1}{2} (x - \mu)^{T} \Sigma^{-1} (x - \mu)$$

• Recall that the density for  $x \sim \mathcal{N}(\mu, \Sigma)$  is

$$p(x \mid \mu, \Sigma) = \frac{1}{\sqrt{|2\pi\Sigma|}} \exp\left(-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)\right)$$

• And the log-density is

$$\log p(x \mid \mu, \Sigma) = -\frac{1}{2} \log |2\pi\Sigma| - \frac{1}{2} (x - \mu)^{T} \Sigma^{-1} (x - \mu)$$

• To estimate  $\mu$  and  $\Sigma$  from a sample  $x_1, \ldots, x_n$  i.i.d.  $\mathcal{N}(\mu, \Sigma)$ , we'll maximize the log joint density:

$$\sum_{i=1}^{n} \log p(x_i \mid \mu, \Sigma) = -\frac{n}{2} \log |2\pi\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

• To estimate  $\mu$  and  $\Sigma$  from a sample  $x_1, \ldots, x_n$  i.i.d.  $\mathcal{N}(\mu, \Sigma)$ , we'll maximize the log joint density:

$$J(\mu, \Sigma) = \sum_{i=1}^{n} \log p(x \mid \mu, \Sigma) = -\frac{n}{2} \log |2\pi\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

• To estimate  $\mu$  and  $\Sigma$  from a sample  $x_1, \ldots, x_n$  i.i.d.  $\mathcal{N}(\mu, \Sigma)$ , we'll maximize the log joint density:

$$J(\mu, \Sigma) = \sum_{i=1}^{n} \log p(x \mid \mu, \Sigma) = -\frac{n}{2} \log |2\pi\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

ullet This is a solid exercise in vector and matrix differentiation. Find  $\hat{\mu}$  and  $\hat{\Sigma}$  satisfying

$$abla_{\mu}J(\mu,\Sigma)=0$$
 $abla_{\Sigma}J(\mu,\Sigma)=0$ 

• To estimate  $\mu$  and  $\Sigma$  from a sample  $x_1, \ldots, x_n$  i.i.d.  $\mathcal{N}(\mu, \Sigma)$ , we'll maximize the log joint density:

$$J(\mu, \Sigma) = \sum_{i=1}^{n} \log p(x \mid \mu, \Sigma) = -\frac{n}{2} \log |2\pi\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

ullet This is a solid exercise in vector and matrix differentiation. Find  $\hat{\mu}$  and  $\hat{\Sigma}$  satisfying

$$abla_{\mu}J(\mu,\Sigma)=0$$
 $abla_{\Sigma}J(\mu,\Sigma)=0$ 

• We get a closed form solution:

$$\hat{\mu}_{\mathsf{MLE}} = \frac{1}{n} \sum_{i=1}^{n} x_{i}$$

$$\hat{\Sigma}_{\mathsf{MLE}} = \frac{1}{n} \sum_{i=1}^{n} (x_{i} - \hat{\mu}_{\mathsf{MLE}})^{T} (x_{i} - \hat{\mu}_{\mathsf{MLE}})$$

• GMM log-likelihood:

$$J(\pi, \mu, \Sigma) = \sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \frac{\pi_z}{\sqrt{|2\pi\Sigma_z|}} \exp\left(-\frac{1}{2}(x-\mu_z)^T \Sigma_z^{-1}(x-\mu_z)\right) \right\}$$

• GMM log-likelihood:

$$J(\pi, \mu, \Sigma) = \sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \frac{\pi_z}{\sqrt{|2\pi\Sigma_z|}} \exp\left(-\frac{1}{2}(x-\mu_z)^T \Sigma_z^{-1}(x-\mu_z)\right) \right\}$$

• Let's compare to the log-likelihood for a single Gaussian:

$$J(\mu, \Sigma) = -\frac{n}{2} \log |2\pi\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

• GMM log-likelihood:

$$J(\pi, \mu, \Sigma) = \sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \frac{\pi_z}{\sqrt{|2\pi\Sigma_z|}} \exp\left(-\frac{1}{2}(x-\mu_z)^T \Sigma_z^{-1}(x-\mu_z)\right) \right\}$$

• Let's compare to the log-likelihood for a single Gaussian:

$$J(\mu, \Sigma) = -\frac{n}{2} \log |2\pi\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

- For a single Gaussian, the log cancels the exp in the Gaussian density.
  - $\implies$  Things simplify a lot.

• GMM log-likelihood:

$$J(\pi, \mu, \Sigma) = \sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \frac{\pi_z}{\sqrt{|2\pi\Sigma_z|}} \exp\left(-\frac{1}{2}(x-\mu_z)^T \Sigma_z^{-1}(x-\mu_z)\right) \right\}$$

• Let's compare to the log-likelihood for a single Gaussian:

$$J(\mu, \Sigma) = -\frac{n}{2} \log |2\pi\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)^T \Sigma^{-1} (x_i - \mu)$$

- For a single Gaussian, the log cancels the exp in the Gaussian density.
  - $\implies$  Things simplify a lot.
- For the GMM, the sum inside the log prevents this cancellation.
  - Expression more complicated. No closed form expression for MLE.

Issues with MLE for GMM

Suppose we have found parameters

Cluster probabilities:  $\pi = (\pi_1, \dots, \pi_k)$ 

Cluster means:  $\mu = (\mu_1, \dots, \mu_k)$ 

Cluster covariance matrices:  $\Sigma = (\Sigma_1, \dots \Sigma_k)$ 

• Suppose we have found parameters

Cluster probabilities:  $\pi = (\pi_1, \dots, \pi_k)$ 

Cluster means:  $\mu = (\mu_1, \dots, \mu_k)$ 

Cluster covariance matrices:  $\Sigma = (\Sigma_1, \dots \Sigma_k)$ 

that are at a local minimum.

• What happens if we shuffle the clusters? e.g. Switch the labels for clusters 1 and 2.

• Suppose we have found parameters

Cluster probabilities:  $\pi = (\pi_1, \dots, \pi_k)$ 

Cluster means:  $\mu = (\mu_1, ..., \mu_k)$ 

Cluster covariance matrices:  $\Sigma = (\Sigma_1, \dots \Sigma_k)$ 

- What happens if we shuffle the clusters? e.g. Switch the labels for clusters 1 and 2.
- We'll get the same likelihood. How many such equivalent settings are there?

• Suppose we have found parameters

Cluster probabilities:  $\pi = (\pi_1, \dots, \pi_k)$ 

Cluster means:  $\mu = (\mu_1, \dots, \mu_k)$ 

Cluster covariance matrices:  $\Sigma = (\Sigma_1, \dots \Sigma_k)$ 

- What happens if we shuffle the clusters? e.g. Switch the labels for clusters 1 and 2.
- We'll get the same likelihood. How many such equivalent settings are there?
- Assuming all clusters are distinct, there are k! equivalent solutions.

• Suppose we have found parameters

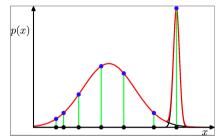
Cluster probabilities:  $\pi = (\pi_1, \dots, \pi_k)$ 

Cluster means:  $\mu = (\mu_1, \dots, \mu_k)$ 

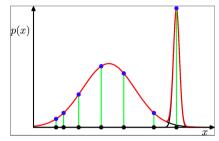
Cluster covariance matrices:  $\Sigma = (\Sigma_1, \dots \Sigma_k)$ 

- What happens if we shuffle the clusters? e.g. Switch the labels for clusters 1 and 2.
- We'll get the same likelihood. How many such equivalent settings are there?
- Assuming all clusters are distinct, there are k! equivalent solutions.
- Not a problem per se, but something to be aware of.

• Consider the following GMM for 7 data points:

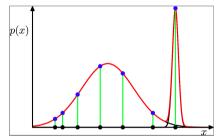


• Consider the following GMM for 7 data points:



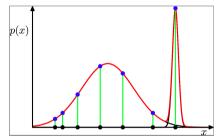
- Let  $\sigma^2$  be the variance of the skinny component.
- What happens to the likelihood as  $\sigma^2 \to 0$ ?

• Consider the following GMM for 7 data points:



- Let  $\sigma^2$  be the variance of the skinny component.
- What happens to the likelihood as  $\sigma^2 \to 0$ ?
- In practice, we end up in local minima that do not have this problem.
  - Or keep restarting optimization until we do.

• Consider the following GMM for 7 data points:



- Let  $\sigma^2$  be the variance of the skinny component.
- What happens to the likelihood as  $\sigma^2 \to 0?$
- In practice, we end up in local minima that do not have this problem.
  - Or keep restarting optimization until we do.
- Bayesian approach or regularization will also solve the problem.

From Bishop's Pattern recognition and machine learning, Figure 9.7.

#### Gradient Descent / SGD for GMM

• What about running gradient descent or SGD on

$$J(\pi, \mu, \Sigma) = -\sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}?$$

<sup>&</sup>lt;sup>1</sup>See Hosseini and Sra's Manifold Optimization for Gaussian Mixture Models for discussion and further references

• What about running gradient descent or SGD on

$$J(\pi, \mu, \Sigma) = -\sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}?$$

Can be done, in principle – but need to be clever about it.

<sup>&</sup>lt;sup>1</sup>See Hosseini and Sra's Manifold Optimization for Gaussian Mixture Models for discussion and further references.

What about running gradient descent or SGD on

$$J(\pi, \mu, \Sigma) = -\sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}?$$

- Can be done, in principle but need to be clever about it.
- Each matrix  $\Sigma_1, \ldots, \Sigma_k$  has to be positive semidefinite.

<sup>&</sup>lt;sup>1</sup>See Hosseini and Sra's Manifold Optimization for Gaussian Mixture Models for discussion and further references.

What about running gradient descent or SGD on

$$J(\pi, \mu, \Sigma) = -\sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}?$$

- Can be done, in principle but need to be clever about it.
- Each matrix  $\Sigma_1, \ldots, \Sigma_k$  has to be positive semidefinite.
- How to maintain that constraint?

<sup>&</sup>lt;sup>1</sup>See Hosseini and Sra's Manifold Optimization for Gaussian Mixture Models for discussion and further references.

• What about running gradient descent or SGD on

$$J(\pi, \mu, \Sigma) = -\sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}?$$

- Can be done, in principle but need to be clever about it.
- Each matrix  $\Sigma_1, \ldots, \Sigma_k$  has to be positive semidefinite.
- How to maintain that constraint?
  - Rewrite  $\Sigma_i = M_i M_i^T$ , where  $M_i$  is an unconstrained matrix.

<sup>&</sup>lt;sup>1</sup>See Hosseini and Sra's Manifold Optimization for Gaussian Mixture Models for discussion and further references.

What about running gradient descent or SGD on

$$J(\pi, \mu, \Sigma) = -\sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}?$$

- Can be done, in principle but need to be clever about it.
- Each matrix  $\Sigma_1, \ldots, \Sigma_k$  has to be positive semidefinite.
- How to maintain that constraint?
  - Rewrite  $\Sigma_i = M_i M_i^T$ , where  $M_i$  is an unconstrained matrix.
  - Then  $\Sigma_i$  is positive semidefinite.

<sup>&</sup>lt;sup>1</sup>See Hosseini and Sra's Manifold Optimization for Gaussian Mixture Models for discussion and further references.

What about running gradient descent or SGD on

$$J(\pi, \mu, \Sigma) = -\sum_{i=1}^{n} \log \left\{ \sum_{z=1}^{k} \pi_{z} \mathcal{N}(x_{i} \mid \mu_{z}, \Sigma_{z}) \right\}?$$

- Can be done, in principle but need to be clever about it.
- Each matrix  $\Sigma_1, \ldots, \Sigma_k$  has to be positive semidefinite.
- How to maintain that constraint?
  - Rewrite  $\Sigma_i = M_i M_i^T$ , where  $M_i$  is an unconstrained matrix.
  - Then  $\Sigma_i$  is positive semidefinite.
- Even then, pure gradient-based methods have trouble.<sup>1</sup>

<sup>&</sup>lt;sup>1</sup>See Hosseini and Sra's Manifold Optimization for Gaussian Mixture Models for discussion and further references.

# The EM Algorithm for GMM

- Let's start by considering the MLE for the Gaussian model.
- For data  $\mathcal{D} = \{x_1, \dots, x_n\}$ , the log likelihood is given by

$$\sum_{i=1}^{n} \log \mathcal{N}(x_i \mid \mu, \Sigma) = -\frac{nd}{2} \log (2\pi) - \frac{n}{2} \log |\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)' \Sigma^{-1}(x_i - \mu).$$

- Let's start by considering the MLE for the Gaussian model.
- For data  $\mathcal{D} = \{x_1, \dots, x_n\}$ , the log likelihood is given by

$$\sum_{i=1}^{n} \log \mathcal{N}(x_i \mid \mu, \Sigma) = -\frac{nd}{2} \log (2\pi) - \frac{n}{2} \log |\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)' \Sigma^{-1}(x_i - \mu).$$

With some calculus, we find that the MLE parameters are

$$\mu_{\mathsf{MLE}} = \frac{1}{n} \sum_{i=1}^{n} x_i$$

- Let's start by considering the MLE for the Gaussian model.
- For data  $\mathcal{D} = \{x_1, \dots, x_n\}$ , the log likelihood is given by

$$\sum_{i=1}^{n} \log \mathcal{N}(x_i \mid \mu, \Sigma) = -\frac{nd}{2} \log (2\pi) - \frac{n}{2} \log |\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)' \Sigma^{-1}(x_i - \mu).$$

With some calculus, we find that the MLE parameters are

$$\mu_{\text{MLE}} = \frac{1}{n} \sum_{i=1}^{n} x_{i}$$

$$\Sigma_{\text{MLE}} = \frac{1}{n} \sum_{i=1}^{n} (x_{i} - \mu_{\text{MLE}}) (x_{i} - \mu_{\text{MLE}})^{T}$$

- Let's start by considering the MLE for the Gaussian model.
- For data  $\mathcal{D} = \{x_1, \dots, x_n\}$ , the log likelihood is given by

$$\sum_{i=1}^{n} \log \mathcal{N}(x_i \mid \mu, \Sigma) = -\frac{nd}{2} \log (2\pi) - \frac{n}{2} \log |\Sigma| - \frac{1}{2} \sum_{i=1}^{n} (x_i - \mu)' \Sigma^{-1}(x_i - \mu).$$

With some calculus, we find that the MLE parameters are

$$\mu_{\text{MLE}} = \frac{1}{n} \sum_{i=1}^{n} x_{i}$$

$$\Sigma_{\text{MLE}} = \frac{1}{n} \sum_{i=1}^{n} (x_{i} - \mu_{\text{MLE}}) (x_{i} - \mu_{\text{MLE}})^{T}$$

- For GMM, If we knew the cluster assignment  $z_i$  for each  $x_i$ ,
  - we could compute the MLEs for each cluster.

• Suppose we observe  $(x_1, z_1), \ldots, (x_n, z_n)$  i.i.d. from GMM p(x, z).

- Suppose we observe  $(x_1, z_1), \ldots, (x_n, z_n)$  i.i.d. from GMM p(x, z).
- Them find MLE is easy:

- Suppose we observe  $(x_1, z_1), \ldots, (x_n, z_n)$  i.i.d. from GMM p(x, z).
- Them find MLE is easy:

$$n_z = \sum_{i=1}^n 1(z_i = z)$$

- Suppose we observe  $(x_1, z_1), \ldots, (x_n, z_n)$  i.i.d. from GMM p(x, z).
- Them find MLE is easy:

$$n_z = \sum_{i=1}^n 1(z_i = z)$$

$$\hat{\pi}(z) = \frac{n_z}{n}$$

- Suppose we observe  $(x_1, z_1), \ldots, (x_n, z_n)$  i.i.d. from GMM p(x, z).
- Them find MLE is easy:

$$n_z = \sum_{i=1}^n 1(z_i = z)$$

$$\hat{\pi}(z) = \frac{n_z}{n}$$

$$\hat{\mu}_z = \frac{1}{n_z} \sum_{i: z_i = z} x_i$$

- Suppose we observe  $(x_1, z_1), \ldots, (x_n, z_n)$  i.i.d. from GMM p(x, z).
- Them find MLE is easy:

$$n_{z} = \sum_{i=1}^{n} 1(z_{i} = z)$$

$$\hat{\pi}(z) = \frac{n_{z}}{n}$$

$$\hat{\mu}_{z} = \frac{1}{n_{z}} \sum_{i:z_{i}=z} x_{i}$$

$$\hat{\Sigma}_{z} = \frac{1}{n_{z}} \sum_{i:z_{i}=z} (x_{i} - \hat{\mu}_{z}) (x_{i} - \hat{\mu}_{z})^{T}.$$

- Suppose we observe  $(x_1, z_1), \ldots, (x_n, z_n)$  i.i.d. from GMM p(x, z).
- Them find MLE is easy:

$$n_{z} = \sum_{i=1}^{n} 1(z_{i} = z)$$

$$\hat{\pi}(z) = \frac{n_{z}}{n}$$

$$\hat{\mu}_{z} = \frac{1}{n_{z}} \sum_{i:z_{i}=z} x_{i}$$

$$\hat{\Sigma}_{z} = \frac{1}{n_{z}} \sum_{i:z_{i}=z} (x_{i} - \hat{\mu}_{z}) (x_{i} - \hat{\mu}_{z})^{T}.$$

• In the EM algorithm we will modify the equations to handle our evolving soft assignments, which we will call responsibilities.

$$\gamma_i^j = p(z = j \mid x = x_i).$$

• Denote the probability that observed value  $x_i$  comes from cluster j by

$$\gamma_i^j = p(z = j \mid x = x_i).$$

• The **responsibility** that cluster j takes for observation  $x_i$ .

$$\gamma_i^j = p(z = j \mid x = x_i).$$

- The **responsibility** that cluster j takes for observation  $x_i$ .
- Computationally,

$$\gamma_i^j = p(z=j|x_i).$$

$$\gamma_i^j = p(z = j \mid x = x_i).$$

- The **responsibility** that cluster j takes for observation  $x_i$ .
- Computationally,

$$\gamma_i^j = p(z=j|x_i).$$

$$= p(z=j,x_i)/p(x_i)$$

$$\gamma_i^j = p(z = j \mid x = x_i).$$

- The **responsibility** that cluster j takes for observation  $x_i$ .
- Computationally,

$$\gamma_i^j = p(z=j|x_i).$$

$$= p(z=j,x_i)/p(x_i)$$

$$= \frac{\pi_j \mathcal{N}(x_i | \mu_j, \Sigma_j)}{\sum_{c=1}^k \pi_c \mathcal{N}(x_i | \mu_c, \Sigma_c)}$$

• Denote the probability that observed value  $x_i$  comes from cluster j by

$$\gamma_i^j = p(z = j \mid x = x_i).$$

- The **responsibility** that cluster j takes for observation  $x_i$ .
- Computationally,

$$\gamma_i^j = p(z = j \mid x_i). 
= p(z = j, x_i) / p(x_i) 
= \frac{\pi_j \mathcal{N}(x_i \mid \mu_j, \Sigma_j)}{\sum_{c=1}^k \pi_c \mathcal{N}(x_i \mid \mu_c, \Sigma_c)}$$

• The vector  $(\gamma_i^1, \dots, \gamma_i^k)$  is exactly the **soft assignment** for  $x_i$ .

$$\gamma_i^j = p(z = j \mid x = x_i).$$

- The **responsibility** that cluster j takes for observation  $x_i$ .
- Computationally,

$$\gamma_i^j = p(z=j|x_i). 
= p(z=j,x_i)/p(x_i) 
= \frac{\pi_j \mathcal{N}(x_i|\mu_j, \Sigma_j)}{\sum_{c=1}^k \pi_c \mathcal{N}(x_i|\mu_c, \Sigma_c)}$$

- The vector  $(\gamma_i^1, \dots, \gamma_i^k)$  is exactly the **soft assignment** for  $x_i$ .
- Let  $n_c = \sum_{i=1}^n \gamma_i^c$  be the "number" of points "soft assigned" to cluster c.

• If we know  $\mu_j$ ,  $\Sigma_j$ ,  $\pi_j$  for all clusters j, then easy to find

$$\gamma_i^j = p(z = j \mid x_i)$$

• If we know  $\mu_j$ ,  $\Sigma_j$ ,  $\pi_j$  for all clusters j, then easy to find

$$\gamma_i^j = p(z = j \mid x_i)$$

• If we know the (soft) assignments, we can easily find estimates for  $\pi, \Sigma, \mu$ .

• If we know  $\mu_j$ ,  $\Sigma_j$ ,  $\pi_j$  for all clusters j, then easy to find

$$\gamma_i^j = p(z = j \mid x_i)$$

- If we know the (soft) assignments, we can easily find estimates for  $\pi, \Sigma, \mu$ .
- Repeatedly alternate these two steps.

**1** Initialize parameters  $\mu$ ,  $\Sigma$ ,  $\pi$  (e.g. using k-means).

- **1** Initialize parameters  $\mu$ ,  $\Sigma$ ,  $\pi$  (e.g. using k-means).
- 2 "E step". Evaluate the responsibilities using current parameters:

$$\gamma_i^j = \frac{\pi_j \mathcal{N}(x_i \mid \mu_j, \Sigma_j)}{\sum_{c=1}^k \pi_c \mathcal{N}(x_i \mid \mu_c, \Sigma_c)},$$

for i = 1, ..., n and j = 1, ..., k.

- **1** Initialize parameters  $\mu$ ,  $\Sigma$ ,  $\pi$  (e.g. using k-means).
- (2) "E step". Evaluate the responsibilities using current parameters:

$$\gamma_i^j = \frac{\pi_j \mathcal{N}(x_i \mid \mu_j, \Sigma_j)}{\sum_{c=1}^k \pi_c \mathcal{N}(x_i \mid \mu_c, \Sigma_c)},$$

for i = 1, ..., n and j = 1, ..., k.

(3) "M step". Re-estimate the parameters using responsibilities:

$$\mu_c^{\text{new}} = \frac{1}{n_c} \sum_{i=1}^n \gamma_i^c x_i$$

- **1** Initialize parameters  $\mu$ ,  $\Sigma$ ,  $\pi$  (e.g. using k-means).
- (2) "E step". Evaluate the responsibilities using current parameters:

$$\gamma_i^j = \frac{\pi_j \mathcal{N}(x_i \mid \mu_j, \Sigma_j)}{\sum_{c=1}^k \pi_c \mathcal{N}(x_i \mid \mu_c, \Sigma_c)},$$

for i = 1, ..., n and j = 1, ..., k.

(3) "M step". Re-estimate the parameters using responsibilities:

$$\mu_c^{\text{new}} = \frac{1}{n_c} \sum_{i=1}^n \gamma_i^c x_i$$

$$\Sigma_c^{\text{new}} = \frac{1}{n_c} \sum_{i=1}^n \gamma_i^c (x_i - \mu_c^{\text{new}}) (x_i - \mu_c^{\text{new}})^T$$

- **1** Initialize parameters  $\mu$ ,  $\Sigma$ ,  $\pi$  (e.g. using k-means).
- (2) "E step". Evaluate the responsibilities using current parameters:

$$\gamma_i^j = \frac{\pi_j \mathcal{N}(x_i \mid \mu_j, \Sigma_j)}{\sum_{c=1}^k \pi_c \mathcal{N}(x_i \mid \mu_c, \Sigma_c)},$$

for i = 1, ..., n and j = 1, ..., k.

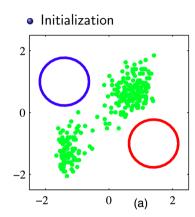
**1** "M step". Re-estimate the parameters using responsibilities:

$$\mu_c^{\text{new}} = \frac{1}{n_c} \sum_{i=1}^n \gamma_i^c x_i$$

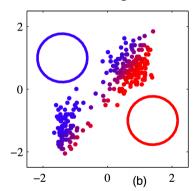
$$\Sigma_c^{\text{new}} = \frac{1}{n_c} \sum_{i=1}^n \gamma_i^c (x_i - \mu_c^{\text{new}}) (x_i - \mu_c^{\text{new}})^T$$

$$\pi_c^{\text{new}} = \frac{n_c}{n_c},$$

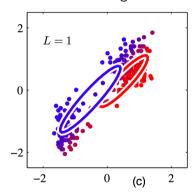
Repeat from Step 2, until log-likelihood converges.



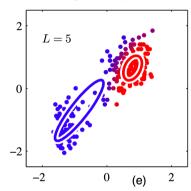
### • First soft assignment:



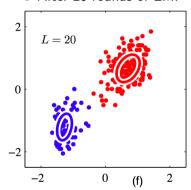
• First soft assignment:



#### • After 5 rounds of EM:



#### • After 20 rounds of EM:



• EM for GMM seems a little like k-means.

- EM for GMM seems a little like k-means.
- In fact, k-means is a limiting case of a **restricted** version of GMM.

- EM for GMM seems a little like k-means.
- In fact, k-means is a limiting case of a restricted version of GMM.
- First, fix each cluster covariance matrix to be  $\sigma^2 I$ .

- EM for GMM seems a little like k-means.
- In fact, k-means is a limiting case of a restricted version of GMM.
- First, fix each cluster covariance matrix to be  $\sigma^2 I$ .
  - (This is the restriction: covariance matrices are fixed, and not iteratively estimated.)

- EM for GMM seems a little like k-means.
- In fact, k-means is a limiting case of a restricted version of GMM.
- First, fix each cluster covariance matrix to be  $\sigma^2 I$ .
  - (This is the restriction: covariance matrices are fixed, and not iteratively estimated.)
- As we take  $\sigma^2 \rightarrow 0$ , the update equations converge to doing k-means.

- EM for GMM seems a little like k-means.
- In fact, k-means is a limiting case of a restricted version of GMM.
- First, fix each cluster covariance matrix to be  $\sigma^2 I$ .
  - (This is the restriction: covariance matrices are fixed, and not iteratively estimated.)
- As we take  $\sigma^2 \to 0$ , the update equations converge to doing k-means.
- If you do a quick experiment yourself, you'll find
  - Soft assignments converge to hard assignments.
  - Has to do with the tail behavior (exponential decay) of Gaussian.