## Expectation Maximization Algorithm

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November 1, 2015

# Kullback-Leibler Divergence

- Let p(x) and q(x) be PMFs on  $\mathfrak{X}$ .
- How can we measure how "different" p and q are?
- The Kullback-Leibler or "KL" Diverence is defined by

$$KL(p||q) = \sum_{x} p(x) \log \frac{p(x)}{q(x)}.$$

(Assumes q(x) = 0 implies p(x) = 0.)

Can also write this as

$$\mathrm{KL}(p\|q) = \mathbb{E}_p \log \frac{p(X)}{q(X)},$$

where  $X \sim p(x)$ .

# Gibbs Inequality $(KL(p||q) \geqslant 0)$

#### Theorem (Gibbs Inequality)

Let p(x) and q(x) be PMFs on  $\mathfrak{X}$ . Then

$$KL(p||q) \geqslant 0$$
,

with equality iff p(x) = q(x) for all  $x \in \mathcal{X}$ .

- KL divergence measures the "distance" between distributions.
- Note:
  - KL divergence not a metric.
  - KL divergence is **not symmetric**.

## Jensen's Inequality

#### Theorem (Jensen's Inequality)

If  $f: \mathcal{X} \to \mathbf{R}$  is a **conve**x function, and  $X \in \mathcal{X}$  is a random variable, then

$$\mathbb{E}f(X) \geqslant f(\mathbb{E}X).$$

Moreover, if f is **strictly convex**, then equality implies that  $X = \mathbb{E}X$  with probability 1 (i.e. X is a constant).

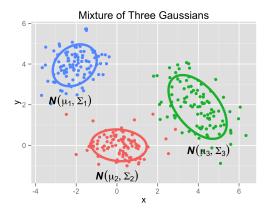
• e.g.  $f(x) = x^2$  is convex. So  $\mathbb{E}X^2 \geqslant (\mathbb{E}X)^2$ . Thus

$$\operatorname{Var} X = \mathbb{E} X^2 - (\mathbb{E} X)^2 \geqslant 0.$$

• Jensen's inequality is used to prove Gibbs inequality  $(\log(x))$  is strictly concave).

# Gaussian Mixture Model (k = 3)

- **1** Choose  $Z \in \{1, 2, 3\} \sim \text{Multi}\left(\frac{1}{3}, \frac{1}{3}, \frac{1}{3}\right)$ .
- 2 Choose  $X \mid Z = z \sim \mathcal{N}(X \mid \mu_z, \Sigma_z)$ .



# Gaussian Mixture Model (k Components)

GMM Parameters

Cluster probabilities: 
$$\pi = (\pi_1, ..., \pi_k)$$

Cluster means: 
$$\mu = (\mu_1, ..., \mu_k)$$

Cluster covariance matrices: 
$$\Sigma = (\Sigma_1, \dots \Sigma_k)$$

- Let  $\theta = (\pi, \mu, \Sigma)$ .
- Marginal log-likelihood

$$\log p(x \mid \theta) = \log \left\{ \sum_{z=1}^{k} \pi_z \mathcal{N}(x \mid \mu_z, \Sigma_z) \right\}$$

#### General Latent Variable Model

- Two sets of random variables: Z and X.
- Z consists of unobserved hidden variables.
- X consists of observed variables.
- Joint probability model parameterized by  $\theta \in \Theta$ :

$$p(x,z \mid \theta)$$

#### Notation abuse

Notation  $p(x, z \mid \theta)$  suggests a Bayesian setting, in which  $\theta$  is a r.v. However we are **not** assuming a Bayesian setting.  $p(x, z \mid \theta)$  is just easier to read than  $p_{\theta}(x, z)$ , once  $\theta$  gets more complicated.

## Complete and Incomplete Data

- An observation of X is called an incomplete data set.
- An observation (X, Z) is called a **complete data set**.
  - We never have a complete data set for latent variable models.
  - But it's a useful construct.
- Suppose we have an incomplete data set  $\mathcal{D} = (x_1, \dots, x_n)$ .
- To simplify notation, take X to represent the entire dataset

$$X = (X_1, \ldots, X_n)$$
,

and Z to represent the corresponding unobserved variables

$$Z = (Z_1, \ldots, Z_n)$$
.

## Log-Likelihood

• The log-likelihood of  $\theta$  for observation X = x is

$$\log p(x \mid \theta) = \log \left\{ \sum_{z} p(x, z \mid \theta) \right\}.$$

- (We write discrete case everything same for continuous case.)
- For exponential families,
  - Without the sum " $\sum_{z}$ ", things simplify.
  - The log and the exp cancel out.
- Assumption for the EM algorithm:
  - Optimization for complete data is relatively easy

$$\underset{\theta \in \Theta}{\operatorname{arg\,max}} \log p(x, z \mid \theta)$$

• (We'll actually need slightly more than this.)

# The EM Algorithm Key Idea

Marginal log likelihood is hard to optimize:

$$\max_{\theta} \log \left\{ \sum_{z} p(x, z \mid \theta) \right\}$$

Full log-likelihood would be easy to optimize:

$$\max_{\theta} \log p(x, z \mid \theta)$$

- What if we had a **distribution** q(z) for the latent variables Z?
  - e.g.  $q(z) = p(z \mid x, \theta)$
- Could maximize the expected complete data log-likelihood:

$$\max_{\theta} \sum_{z} q(z) \log p(x, z \mid \theta)$$

## A Lower Bound for Marginal Likelihood

• Let q(z) be any PMF on  $\mathcal{Z}$ , the support of Z:

$$\log p(x \mid \theta) = \log \sum_{z} p(x, z \mid \theta)$$

$$= \log \sum_{z} q(z) \left[ \frac{p(x, z \mid \theta)}{q(z)} \right]$$

$$\geqslant \sum_{z} q(z) \log \left( \frac{p(x, z \mid \theta)}{q(z)} \right)$$

$$=: \mathcal{L}(q, \theta).$$

• The inequality is by Jensen's, by concavity of the log.

### Lower Bound and Expected Complete Log-Likelihood

• Consider maximizing the lower bound  $\mathcal{L}(q, \theta)$ :

$$\mathcal{L}(q,\theta) = \sum_{z} q(z) \log \left( \frac{p(x,z \mid \theta)}{q(z)} \right)$$

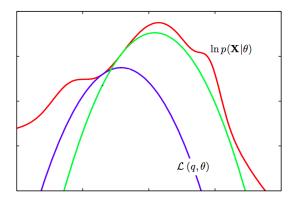
$$= \sum_{z} q(z) \log p(x,z \mid \theta) - \sum_{z} q(z) \log q(z)$$

$$\mathbb{E}[\text{complete log-likelihood}] \quad \text{no } \theta \text{ here}$$

• Maximizing  $\mathcal{L}(q, \theta)$  equivalent to maximizing  $\mathbb{E}[\text{complete data log-likelihood}].$ 

# A Family of Lower Bounds

- Each q gives a different lower bound:  $\log p(x \mid \theta) \geqslant \mathcal{L}(q, \theta)$
- Two lower bounds, as functions of  $\theta$ :



From Bishop's Pattern recognition and machine learning, Figure 9.14.

#### EM: Coordinate Ascent on Lower Bound

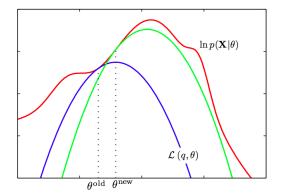
• In EM algorithm, we maximize the lower bound  $\mathcal{L}(q, \theta)$ :

$$\log p(x \mid \theta) \geqslant \mathcal{L}(q, \theta).$$

- EM Algorithm (high level):
  - Choose initial  $\theta^{\text{old}}$ .
  - 2 Let  $q^* = \arg\max_{q} \mathcal{L}(q, \theta^{\text{old}})$
  - 3 Let  $\theta^{\text{new}} = \arg\max_{\theta} \mathcal{L}(q^*, \theta)$ .
  - 4 Go to step 2, until converged.
- Will show:  $p(x \mid \theta^{new}) \ge p(x \mid \theta^{old})$
- Get sequence of  $\theta$ 's with monotonically increasing likelihood.

#### EM: Coordinate Ascent on Lower Bound

- **1** Start at  $\theta^{\text{old}}$ . Find best lower bound at  $\theta^{\text{old}}$ :  $\mathcal{L}(q,\theta)$ .



From Bishop's Pattern recognition and machine learning, Figure 9.14.

#### The Lower Bound

Let's investigate the lower bound:

$$\begin{split} \mathcal{L}(q,\theta) &= \sum_{z} q(z) \log \left( \frac{p(x,z \mid \theta)}{q(z)} \right) \\ &= \sum_{z} q(z) \log \left( \frac{p(z \mid x,\theta) p(x \mid \theta)}{q(z)} \right) \\ &= \sum_{z} q(z) \log \left( \frac{p(z \mid x,\theta)}{q(z)} \right) + \sum_{z} q(z) \log p(x \mid \theta) \\ &= -\mathrm{KL}[q(z), p(z \mid x,\theta)] + \log p(x \mid \theta) \end{split}$$

• Amazing! We get back an equality for the marginal likelihood:

$$\log p(x \mid \theta) = \mathcal{L}(q, \theta) + \text{KL}[q(z), p(z \mid x, \theta)]$$

#### The Best Lower Bound

Find q maximizing

$$\mathcal{L}(q, \theta^{\text{old}}) = -\text{KL}[q(z), p(z \mid x, \theta^{\text{old}})] + \underbrace{\log p(x \mid \theta^{\text{old}})}_{\text{no } q \text{ here}}?$$

- Recall  $KL(p||q) \ge 0$ , and KL(p||p) = 0.
- Best q is  $q^*(z) = p(z \mid x, \theta^{\text{old}})$ :

$$\mathcal{L}(q^*, \theta^{\text{old}}) = -\underbrace{\text{KL}[p(z \mid x, \theta^{\text{old}}), p(z \mid x, \theta^{\text{old}})]}_{=0} + \log p(x \mid \theta^{\text{old}})$$

Summary:

$$\log p(x \mid \theta^{\text{old}}) = \mathcal{L}(q^*, \theta^{\text{old}}) \text{ (tangent at } \theta^{\text{old}}).$$
$$\log p(x \mid \theta) \geqslant \mathcal{L}(q^*, \theta) \quad \forall \theta$$

# General EM Algorithm

- Choose initial  $\theta^{\text{old}}$ .
- Expectation Step
  - Let  $q^*(z) = p(z \mid x, \theta^{\text{old}}).$
  - Let

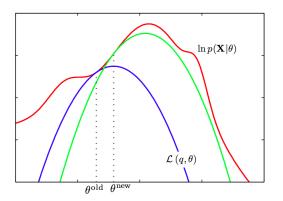
$$J(\theta) = \mathcal{L}(q^*, \theta) = \sum_{z} q^*(z) \log \left( \frac{p(x, z \mid \theta)}{q^*(z)} \right)$$

- Note that  $J(\theta)$  is an **expectation** w.r.t.  $Z \sim q^*(z)$ .
- Maximization Step

$$\theta^{\mathsf{new}} = \underset{\theta}{\mathsf{arg}\,\mathsf{max}}\,J(\theta).$$

Go to step 2, until converged.

# EM Gives Monotonically Increasing Likelihood: By Picture



# EM Gives Monotonically Increasing Likelihood: By Math

- Start at  $\theta^{\text{old}}$ .
- ② Choose  $q^*(z) = \arg\max_{q} \mathcal{L}(q, \theta^{\text{old}})$ . We've shown

$$\log p(x \mid \theta^{\text{old}}) = \mathcal{L}(q^*, \theta^{\text{old}})$$

 ${\color{red} \bullet} \ \, \mathsf{Choose} \,\, \theta^{\mathsf{new}} = \mathsf{arg}\, \mathsf{max}_{\theta}\, \mathcal{L}(q^*,\theta^{\mathsf{old}}). \,\, \mathsf{So}$ 

$$\mathcal{L}(q^*, \theta^{\mathsf{new}}) \geqslant \mathcal{L}(q^*, \theta^{\mathsf{old}}).$$

Putting it together, we get

$$\begin{array}{ll} \log p(x \,|\, \theta^{\mathsf{new}}) & \geqslant & \mathcal{L}(q^*, \theta^{\mathsf{new}}) & \mathcal{L} \text{ is a lower bound} \\ & \geqslant & \mathcal{L}(q^*, \theta^{\mathsf{old}}) & \text{By definition of } \theta^{\mathsf{new}} \\ & = & \log p(x \,|\, \theta^{\mathsf{old}}) & \text{Bound is tight at } \theta^{\mathsf{old}}. \end{array}$$

# EM Gives Monotonically Increasing Likelihood: And so?

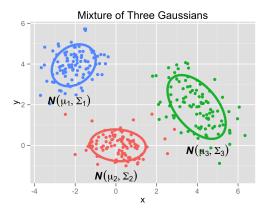
- Let  $\theta_n$  be value of EM algorithm after n steps.
- Are there conditions for which
  - $\theta_n$  converges to the maximum likelihood?
  - $\theta_n$  converges to a local maximum?
  - $\theta_n$  converges to a stationary point of likelihood?
  - $\theta_n$  converges?
- There are conditions for each of these (to happen and not to happen).
- See "On the Convergence Properties of the EM Algorithm" by C. F. Jeff Wu, The Annals of Statistics, Mar. 1983.
  - http://web.stanford.edu/class/ee378b/papers/wu-em.pdf
- In practice, can run EM multiple times with random starts.

# Homework: Derive EM for GMM from General EM Algorithm

- Subsequent slides may help set things up.
- Key skills:
  - MLE for multivariate Gaussian distributions.
  - Lagrange multipliers

# Gaussian Mixture Model (k = 3)

- **1** Choose  $Z \in \{1, 2, 3\} \sim \text{Multi}\left(\frac{1}{3}, \frac{1}{3}, \frac{1}{3}\right)$ .
- ② Choose  $X \mid Z = z \sim \mathcal{N}(X \mid \mu_z, \Sigma_z)$ .



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Cluster covariance matrices: 
$$\Sigma = (\Sigma_1, \dots \Sigma_k)$$

- Let  $\theta = (\pi, \mu, \Sigma)$ .
- Marginal log-likelihood

$$\log p(x \mid \theta) = \log \left\{ \sum_{z=1}^{k} \pi_z \mathcal{N}(x \mid \mu_z, \Sigma_z) \right\}$$

# $q^*(z) =$ Soft Assignments

At each step, we take

$$q^*(z) = p(z \mid x, \theta^{\text{old}})$$

This corresponds to "soft assignments" we had last time:

$$\gamma_{i}^{j} = \mathbb{P}(Z = j \mid X = x_{i}) \\
= \frac{\pi_{j} \mathcal{N}(x_{i} \mid \mu_{j}, \Sigma_{j})}{\sum_{c=1}^{k} \pi_{c} \mathcal{N}(x_{i} \mid \mu_{c}, \Sigma_{c})}$$

## Expectation Step

• The complete log-likelihood is

$$\log p(x, z \mid \theta) = \sum_{i=1}^{n} \log [\pi_z \mathcal{N}(x_i \mid \mu_z, \Sigma_z)]$$

$$= \sum_{i=1}^{n} \left( \log \pi_z + \underbrace{\log \mathcal{N}(x_i \mid \mu_z, \Sigma_z)}_{\text{simplifies nicely}} \right)$$

• Take the expected complete log-likelihood w.r.t.  $q^*$ :

$$J(\theta) = \sum_{z} q^{*}(z) \log p(x, z \mid \theta)$$
$$= \sum_{i=1}^{n} \sum_{j=1}^{k} \gamma_{i}^{j} [\log \pi_{j} + \log \mathcal{N}(x_{i} \mid \mu_{j}, \Sigma_{j})]$$

# Maximization Step

• Find  $\theta^*$  maximizing  $J(\theta)$ . Result is what we had last time:

$$\begin{array}{lll} \boldsymbol{\mu}_{c}^{\text{new}} & = & \frac{1}{n_{c}} \sum_{i=1}^{n} \boldsymbol{\gamma}_{i}^{c} \boldsymbol{x}_{i} \\ \\ \boldsymbol{\Sigma}_{c}^{\text{new}} & = & \frac{1}{n_{c}} \sum_{i=1}^{n} \boldsymbol{\gamma}_{i}^{c} \left( \boldsymbol{x}_{i} - \boldsymbol{\mu}_{\text{MLE}} \right) \left( \boldsymbol{x}_{i} - \boldsymbol{\mu}_{\text{MLE}} \right)^{T} \\ \\ \boldsymbol{\pi}_{c}^{\text{new}} & = & \frac{n_{c}}{n}, \end{array}$$

for each  $c = 1, \ldots, k$ .

## Machine Learning

- Look at other course notes at this level.
  - Every course covers different subset of topics.
  - Different perspectives. (e.g. Bayesian / Probabilistic)
- Read on some "second semester" topics
  - LDA / Topic Models (DS-GA 1005?)
  - Sequence models: Hidden Markov Models / MEMMs / CRFs (DS-GA 1005)
  - Bayesian methods
  - Collaborative Filtering / Recommendations
  - Ranking
  - Bandit problems (Thompson sampling / UCB methods)
  - Gaussian processes

#### Other Stuff To Learn

- Statistics
- Data Structures & Algorithms (Theoretical)
- Some production programming language (e.g. Java, C++)