Basic Statistics and a Bit of Bootstrap

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Bias and Variance

Parameters

- Suppose we have a probability distribution *P*.
- Often went to estimate some characteristic of P.
 - e.g. expected value, variance, kurtosis, median, etc...
- These things are called **parameters** of *P*.
- A parameter $\mu = \mu(P)$ is any function of the distribution P.
- Question: Is μ random?
- Answer: Nope. For example if P has density f(x) on R, then mean is

$$\mu = \int_{-\infty}^{\infty} x f(x) \, dx,$$

which just an integral - nothing random.

Statistics and Estimators

- Suppose $\mathcal{D}_n = (x_1, x_2, \dots, x_n)$ is an i.i.d. sample from P.
- A statistic $s = s(\mathcal{D}_n)$ is any function of the data.
- A statistic $\hat{\mu} = \hat{\mu}(\mathcal{D}_n)$ is a **point estimator** of μ if $\hat{\mu} \approx \mu$.
- Question: Are statistics and/or point estimators random?
- Answer: Yes, since we're considering the data to be random.
 - The function $s(\cdot)$ isn't random, but we're plugging in random inputs.

Examples of Statistics

- Mean: $\bar{x}(\mathfrak{D}_n) = \frac{1}{n} \sum_{i=1}^n x_i$.
- Median: $m(\mathcal{D}_n) = \text{median}(x_1, \dots, x_n)$
- Sample variance: $\sigma^2(\mathcal{D}_n) = \frac{1}{n-1} \sum_{i=1}^n (x_i \bar{x}(\mathcal{D}_n))^2$

Fancier:

- A data histogram is a statistic.
- Empirical distribution function.
- A confidence interval.

Statistics are Random

- Statistics are random, so they have probability distributions.
- The distribution of a statistic is called a **sampling distribution**.
- We often want to know some parameters of the sampling distribution.
 - Most commonly the mean and the standard deviation.
- The standard deviation of the sampling distribution is called the **standard error**.
- Question: Is standard error random?
- Answer: Nope. It's a parameter of a distribution.

Bias and Variance for Real-Valued Estimators

- Let $\mu: P \mapsto \mathbf{R}$ be a real-valued parameter.
- Let $\hat{\mu}: \mathcal{D}_n \mapsto \mathbf{R}$ be an estimator of μ .
- We define the bias of $\hat{\mu}$ to be $Bias(\hat{\mu}) = \mathbb{E}\hat{\mu} \mu$.
- We define the variance of $\hat{\mu}$ to be $Var(\hat{\mu}) = \mathbb{E}\hat{\mu}^2 (\mathbb{E}\hat{\mu})^2$.
- An estimator is **unbiased** if $Bias(\hat{\mu}) = \mathbb{E}\hat{\mu} \mu = 0$.

Neither bias nor variance depend on a specific sample \mathcal{D}_n . We are taking expectation over \mathcal{D}_n .

Estimating Variance of an Estimator

- To estimate $Var(\hat{\mu})$ we need estimates of $\mathbb{E}\hat{\mu}$ and $\mathbb{E}\hat{\mu}^2$.
- Instead of a single sample \mathcal{D}_n of size n, suppose we had
 - B independent samples of size $n: \mathcal{D}_n^1, \mathcal{D}_n^2, \dots, \mathcal{D}_n^B$
- Can then estimate

$$\mathbb{E}\hat{\mu} \approx \frac{1}{B} \sum_{i=1}^{B} \hat{\mu} \left(\mathcal{D}_{n}^{i} \right)$$

$$\mathbb{E}\hat{\mu}^{2} \approx \frac{1}{B} \sum_{i=1}^{B} \left[\hat{\mu} \left(\mathcal{D}_{n}^{i} \right) \right]^{2}$$

and

$$\operatorname{Var}(\hat{\mu}) \approx \frac{1}{B} \sum_{i=1}^{B} \left[\hat{\mu} \left(\mathcal{D}_{n}^{i} \right) \right]^{2} - \left[\frac{1}{B} \sum_{i=1}^{B} \hat{\mu} \left(\mathcal{D}_{n}^{i} \right) \right]^{2}.$$

Putting "Error Vars" on Estimator

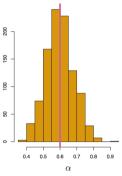
- Why do we even care about estimating variance?
- Would like to report a confidence interval for our point estimate:

$$\hat{\mu} \pm \sqrt{\widehat{Var}(\hat{\mu})}$$

- (This confidence interval assumes $\hat{\mu}$ is unbiased.)
- \bullet Our estimate of standard error is $\sqrt{\widehat{Var}(\hat{\mu})}.$

Histogram of Estimator

- Want to estimate $\alpha = \alpha(P)$ for some known P, and some complicated α .
- Point estimator $\hat{\alpha} = \hat{\alpha}(\mathcal{D}_{100})$ for samples of size 100.
- Histogram of $\hat{\alpha}$ for 1000 random datasets of size 100:



Practical Issue

- We typically get only one sample \mathfrak{D}_n .
- We could divide it into B groups.
- Our estimator would be $\hat{\mu} = \hat{\mu} (\mathcal{D}_{n/B})$.
- And we could get a variance estimate for $\hat{\mu}$.
- But the estimator itself would not be as good as if we used all data:

$$\hat{\mu} = \hat{\mu}(\mathcal{D}_n).$$

- Can we get the best of both worlds?
 - A good point estimate AND a variance estimate?

The Bootstrap

The Bootstrap Sample

Definition

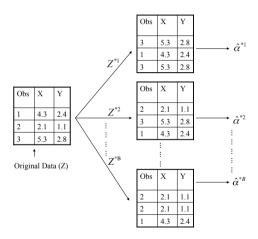
A **bootstrap sample** from $\mathcal{D}_n = \{x_1, \dots, x_n\}$ is a sample of size n drawn with replacement from \mathcal{D}_n .

- In a bootstrap sample, some elements of \mathfrak{D}_n
 - will show up multiple times,
 - some won't show up at all.
- Each X_i has a probability $(1-1/n)^n$ of not being selected.
- Recall from analysis that for large n,

$$\left(1-\frac{1}{n}\right)^n \approx \frac{1}{e} \approx .368.$$

• So we expect $^{\sim}63.2\%$ of elements of $\mathcal D$ will show up at least once.

The Bootstrap Sample



From An Introduction to Statistical Learning, with applications in R (Springer, 2013) with permission from the authors: G. James, D. Witten, T. Hastie and R. Tibshirani.

The Bootstrap Method

Definition

A **bootstrap method** is when you *simulate* having B independent samples from P by taking B bootstrap samples from the sample \mathfrak{D}_n .

- Given original data \mathcal{D}_n , compute B bootstrap samples D_n^1, \ldots, D_n^B .
- For each bootstrap sample, compute some function

$$\phi(D_n^1), \ldots, \phi(D_n^B)$$

- Work with these values as though D_n^1, \ldots, D_n^B were i.i.d. P.
- Amazing fact: Things often come out very close to what we'd get with independent samples from *P*.

Independent vs Bootstrap Samples

- Want to estimate $\alpha = \alpha(P)$ for some known P and some complicated α .
- Point estimator $\hat{\alpha} = \hat{\alpha}(\mathcal{D}_{100})$ for samples of size 100.
- ullet Histogram of \hat{lpha} based on
 - 1000 independent samples of size 100, vs
 - 1000 bootstrap samples of size 100

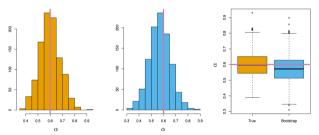


Figure 5.10 from ISLR (Springer, 2013) with permission from the authors: G. James, D. Witten, T. Hastie and R. Tibshirani.

The Bootstrap in Practice

- Suppose we have an estimator $\hat{\mu} = \hat{\mu}(\mathcal{D}_n)$.
- To get error bars, we can compute the "bootstrap variance".
 - Draw B bootstrap samples.
 - Compute empirical variance of $\hat{\mu}(\mathcal{D}_n^1), \ldots, \hat{\mu}(\mathcal{D}_n^B)$..
- Could report

$$\hat{\mu}(\mathcal{D}_n) \pm \sqrt{\mathsf{Bootstrap Variance}}$$