## Basic Statistics and a Bit of Bootstrap

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Bias and Variance

#### **Parameters**

- Suppose we have a probability distribution *P*.
- Often we want to estimate some characteristic of P.
  - e.g. expected value, variance, kurtosis, median, etc...
- These things are called **parameters** of *P*.
- A parameter  $\mu = \mu(P)$  is any function of the distribution P.
- Question: Is μ random?
- Answer: Nope. For example if P has density f(x) on R, then mean is

$$\mu = \int_{-\infty}^{\infty} x f(x) \, dx,$$

which is just an integral - nothing random.

#### Statistics and Estimators

- Suppose  $\mathcal{D}_n = (x_1, x_2, \dots, x_n)$  is an i.i.d. sample from P.
- A statistic  $s = s(\mathcal{D}_n)$  is any function of the data.
- A statistic  $\hat{\mu} = \hat{\mu}(\mathcal{D}_n)$  is a **point estimator** of  $\mu$  if  $\hat{\mu} \approx \mu$ .
- Question: Are statistics and/or point estimators random?
- Answer: Yes, since we're considering the data to be random.
  - The function  $s(\cdot)$  isn't random, but we're plugging in random inputs.

### **Examples of Statistics**

- Mean:  $\bar{x}(\mathfrak{D}_n) = \frac{1}{n} \sum_{i=1}^n x_i$ .
- Median:  $m(\mathcal{D}_n) = \text{median}(x_1, \dots, x_n)$
- Sample variance:  $\sigma^2(\mathcal{D}_n) = \frac{1}{n-1} \sum_{i=1}^n (x_i \bar{x}(\mathcal{D}_n))^2$

#### Fancier:

- A data histogram is a statistic.
- Empirical distribution function.
- A confidence interval.

#### Statistics are Random

- Statistics are random, so they have probability distributions.
- The distribution of a statistic is called a **sampling distribution**.
- We often want to know some parameters of the sampling distribution.
  - Most commonly the mean and the standard deviation.
- The standard deviation of the sampling distribution is called the **standard error**.
- Question: Is standard error random?
- Answer: Nope. It's a parameter of a distribution.

### Bias and Variance for Real-Valued Estimators

- Let  $\mu: P \mapsto \mathbf{R}$  be a real-valued parameter.
- Let  $\hat{\mu}: \mathcal{D}_n \mapsto \mathbf{R}$  be an estimator of  $\mu$ .
- For short, write  $\mu = \mu(P)$  and  $\hat{\mu} = \hat{\mu}(\mathcal{D}_n)$ .
- We define the bias of  $\hat{\mu}$  to be Bias( $\hat{\mu}$ ) =  $\mathbb{E}\hat{\mu} \mu$ .
- We define the variance of  $\hat{\mu}$  to be  $Var(\hat{\mu}) = \mathbb{E}\hat{\mu}^2 (\mathbb{E}\hat{\mu})^2$ .
- An estimator is **unbiased** if  $Bias(\hat{\mu}) = \mathbb{E}\hat{\mu} \mu = 0$ .

Neither bias nor variance depend on a specific sample  $\mathcal{D}_n$ . We are taking expectation over  $\mathcal{D}_n$ .

# Estimating Variance of an Estimator

- To estimate  $Var(\hat{\mu})$  we need estimates of  $\mathbb{E}\hat{\mu}$  and  $\mathbb{E}\hat{\mu}^2$ .
- Instead of a single sample  $\mathcal{D}_n$  of size n, suppose we had
  - B independent samples of size n:  $\mathcal{D}_n^1, \mathcal{D}_n^2, \dots, \mathcal{D}_n^B$
- Can then estimate

$$\mathbb{E}\hat{\mu} \approx \frac{1}{B} \sum_{i=1}^{B} \hat{\mu} \left( \mathcal{D}_{n}^{i} \right)$$

$$\mathbb{E}\hat{\mu}^{2} \approx \frac{1}{B} \sum_{i=1}^{B} \left[ \hat{\mu} \left( \mathcal{D}_{n}^{i} \right) \right]^{2}$$

and

$$\operatorname{Var}(\hat{\mu}) \approx \frac{1}{B} \sum_{i=1}^{B} \left[ \hat{\mu} \left( \mathcal{D}_{n}^{i} \right) \right]^{2} - \left[ \frac{1}{B} \sum_{i=1}^{B} \hat{\mu} \left( \mathcal{D}_{n}^{i} \right) \right]^{2}.$$

# Putting "Error Bars" on Estimator

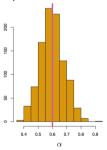
- Why do we even care about estimating variance?
- Would like to report a confidence interval for our point estimate:

$$\hat{\mu} \pm \sqrt{\widehat{Var(\hat{\mu})}}$$

- (This confidence interval assumes  $\hat{\mu}$  is unbiased.)
- $\bullet$  Our estimate of standard error is  $\sqrt{\widehat{Var}(\hat{\mu})}.$

### Histogram of Estimator

- Want to estimate  $\alpha = \alpha(P) \in \mathbf{R}$  for some unknown P, and some complicated  $\alpha$ .
- Point estimator  $\hat{\alpha} = \hat{\alpha}(\mathcal{D}_{100})$  for samples of size 100.
- How to get error bars on  $\hat{\alpha}$ ?
- Histogram of  $\hat{\alpha}$  for 1000 random datasets of size 100 (estimates sampling distribution of  $\hat{\alpha}$ ):



Pink line indicates true value of  $\alpha$ . This is Figure 5.10 from An Introduction to Statistical Learning, with applications in R (Springer, 2013) with permission from the authors: G. James, D. Witten, T. Hastie and R. Tibshirani.

#### Practical Issue

- We typically get only one sample  $\mathcal{D}_n$ .
- We could divide it into *B* groups.
- Our estimator would be  $\hat{\mu} = \hat{\mu} \left( \mathcal{D}_{n/B} \right)$ .
- And we could get a variance estimate for  $\hat{\mu}$ .
- But the estimator itself would not be as good as if we used all data:

$$\hat{\mu} = \hat{\mu}(\mathcal{D}_n).$$

- Can we get the best of both worlds?
  - A good point estimate AND a variance estimate?

The Bootstrap

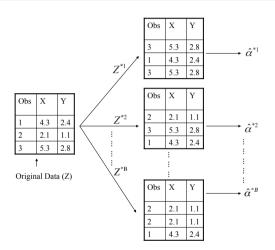
# The Bootstrap Sample

- A **bootstrap sample** from  $\mathcal{D}_n = (x_1, ..., x_n)$  is a sample of size n drawn with replacement from  $\mathcal{D}_n$ .
- In a bootstrap sample, some elements of  $\mathfrak{D}_n$ 
  - will show up multiple times, and
  - some won't show up at all.
- Each  $X_i$  has a probability of  $(1-1/n)^n$  of not being selected.
- Recall from analysis that for large n,

$$\left(1-\frac{1}{n}\right)^n \approx \frac{1}{e} \approx .368.$$

• So we expect  $^{\sim}63.2\%$  of elements of  $\mathcal D$  will show up at least once.

# The Bootstrap Sample



From An Introduction to Statistical Learning, with applications in R (Springer, 2013) with permission from the authors: G. James, D. Witten, T. Hastie and R. Tibshirani.

## The Bootstrap Method

#### Definition

A **bootstrap method** is when you *simulate* having B independent samples from P by taking B bootstrap samples from the sample  $\mathfrak{D}_n$ .

- Given original data  $\mathcal{D}_n$ , compute B bootstrap samples  $D_n^1, \ldots, D_n^B$ .
- For each bootstrap sample, compute some function

$$\phi(D_n^1), \ldots, \phi(D_n^B)$$

- Work with these values as though  $D_n^1, \ldots, D_n^B$  were i.i.d. P.
- Amazing fact: Things often come out very close to what we'd get with independent samples from *P*.

### Independent vs Bootstrap Samples

- Want to estimate  $\alpha = \alpha(P)$  for some unknown P and some complicated  $\alpha$ .
- Point estimator  $\hat{\alpha} = \hat{\alpha}(\mathcal{D}_{100})$  for samples of size 100.
- Histogram of  $\hat{\alpha}$  based on
  - 1000 independent samples of size 100, vs
  - 1000 bootstrap samples of size 100

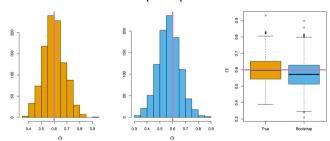


Figure 5.10 from ISLR (Springer, 2013) with permission from the authors: G. James, D. Witten, T. Hastie and R. Tibshirani.

# The Bootstrap in Practice

- Suppose we have an estimator  $\hat{\mu} = \hat{\mu}(\mathcal{D}_n)$ .
- To get error bars, we can compute the "bootstrap variance".
  - Draw B bootstrap samples.
  - Compute sample or empirical variance of  $\hat{\mu}(\mathcal{D}_n^1), \ldots, \hat{\mu}(\mathcal{D}_n^B)$ ...
- Could report

$$\hat{\mu}(\mathcal{D}_n) \pm \sqrt{\mathsf{Bootstrap Variance}}$$