Bootstrap, Bagging, and Random Forests

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Bias and Variance

Parameters

- Suppose we have a probability distribution *P*.
- Often went to estimate some characteristic of P.
 - e.g. expected value, variance, kurtosis, median, etc...
- These things are called **parameters** of *P*.
- A parameter $\mu = \mu(P)$ is any function of the distribution P.

Statistics and Estimators

- Suppose $\mathfrak{D}_n = (x_1, x_2, \dots, x_n)$ is an i.i.d. sample from P.
- A statistic $s = s(\mathcal{D}_n)$ is any function of the data.
- A statistic $\hat{\mu} = \hat{\mu}(\mathcal{D}_n)$ is a **point estimator** of μ if $\hat{\mu} \approx \mu$.

Bias and Variance for Real-Valued Estimators

- Let $\mu: P \mapsto R$ be a real-valued parameter.
- Let $\hat{\mu}: \mathcal{D}_n \mapsto \mathbf{R}$ be an estimator of μ .
- Examples?
- We define the bias of $\hat{\mu}$ to be $Bias(\hat{\mu}) = \mathbb{E}\hat{\mu} \mu$.
- We define the variance of $\hat{\mu}$ to be $Var(\hat{\mu}) = \mathbb{E}\hat{\mu}^2 (\mathbb{E}\hat{\mu})^2$.
- An estimator is **unbiased** if $Bias(\hat{\mu}) = \mathbb{E}\hat{\mu} \mu = 0$.

Neither bias nor variance depend on a specific sample \mathcal{D}_n . We are taking expectation over \mathcal{D}_n .

Estimating Variance of an Estimator

- To estimate $Var(\hat{\mu})$ we need estimates of $\mathbb{E}\hat{\mu}$ and $\mathbb{E}\hat{\mu}^2$.
- Instead of a single sample \mathcal{D}_n of size n, suppose we had
 - B independent samples of size $n: \mathcal{D}_n^1, \mathcal{D}_n^2, \dots, \mathcal{D}_n^B$
- Can then estimate

$$\mathbb{E}\hat{\mu} \approx \frac{1}{B} \sum_{i=1}^{B} \hat{\mu} \left(\mathcal{D}_{n}^{i} \right)$$

$$\mathbb{E}\hat{\mu}^{2} \approx \frac{1}{B} \sum_{i=1}^{B} \left[\hat{\mu} \left(\mathcal{D}_{n}^{i} \right) \right]^{2}$$

and

$$\operatorname{Var}(\hat{\mu}) \approx \frac{1}{B} \sum_{i=1}^{B} \left[\hat{\mu} \left(\mathcal{D}_{n}^{i} \right) \right]^{2} - \left[\frac{1}{B} \sum_{i=1}^{B} \hat{\mu} \left(\mathcal{D}_{n}^{i} \right) \right]^{2}.$$

Putting "Error Vars" on Estimator

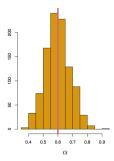
- Why do we even care about estimating variance?
- Would like to report a confidence interval for our point estimate:

$$\hat{\mu} \pm \sqrt{\widehat{Var(\hat{\mu})}}$$

• (This confidence interval assumes $\hat{\mu}$ is unbiased.)

Histogram of Estimator

- Want to estimate $\alpha = \alpha(P)$ for some known P. (some complicated α)
- Point estimator $\hat{\alpha} = \hat{\alpha}(\mathcal{D}_{100})$ for samples of size 100.
- Histogram of $\hat{\alpha}$ for 1000 random datasets of size 100:



Pink line indicates true value of α . This is Figure 5.10 from *An Introduction to Statistical Learning, with applications in R* (Springer, 2013) with permission from the authors: G. James, D. Witten, T. Hastie and R. Tibshirani

Practical Issue

- We typically get only one sample \mathfrak{D}_n .
- We could divide it into B groups.
- Our estimator would be $\hat{\mu} = \hat{\mu} (\mathcal{D}_{n/B})$.
- And we could get a variance estimate for $\hat{\mu}$.
- But the estimator itself would not be as good as if we used all data:

$$\hat{\mu} = \hat{\mu}(\mathcal{D}_n).$$

- Can we get the best of both worlds?
 - A good point estimate AND a variance estimate?

The Bootstrap

The Bootstrap Sample

Definition

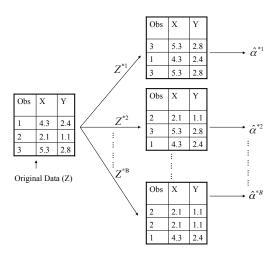
A **bootstrap sample** from $\mathcal{D}_n = \{x_1, \dots, x_n\}$ is a sample of size n drawn with replacement from \mathcal{D}_n .

- \bullet In a bootstrap sample, some elements of ${\mathfrak D}$
 - will show up multiple times,
 - some won't show up at all.
- Each X_i has a probability $(1-1/n)^n$ of not being selected.
- Recall from analysis that for large n,

$$\left(1-\frac{1}{n}\right)^n \approx \frac{1}{e} \approx .368.$$

 \bullet So we expect ~63.2% of elements of ${\mathfrak D}$ will show up at least once.

The Bootstrap Sample



From An Introduction to Statistical Learning, with applications in R (Springer, 2013) with permission from the authors: G. James. D. Witten, T. Hastie and R. Tibshirani.

The Bootstrap Method

Definition

A **bootstrap method** is when you *simulate* having B independent samples from P by taking B bootstrap samples from the sample \mathcal{D}_n .

- Given original data \mathcal{D}_n , compute B bootstrap samples D_n^1, \ldots, D_n^B .
- For each bootstrap sample, compute some function

$$\phi(D_n^1), \ldots, \phi(D_n^B)$$

- Work with these values as though D_n^1, \ldots, D_n^B were i.i.d. P.
- Amazing fact: Things often come out very close to what we'd get with independent samples from *P*.

Independent vs Bootstrap Samples

- Want to estimate $\alpha = \alpha(P)$ for some known P. (some complicated α)
- Point estimator $\hat{\alpha} = \hat{\alpha}(\mathcal{D}_{100})$ for samples of size 100.
- Histogram of $\hat{\alpha}$ based on
 - 1000 independent samples of size 100, vs
 - 1000 bootstrap samples of size 100

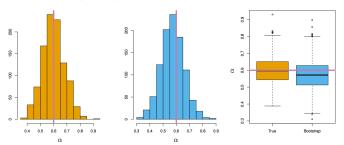


Figure 5.10 from An Introduction to Statistical Learning, with applications in R (Springer, 2013) with permission from the authors: G. James. D. Witten, T. Hastie and R. Tibshirani.

The Bootstrap in Practice

- Suppose we have an estimator $\hat{\mu} = \hat{\mu}(\mathcal{D}_n)$.
- To get error bars, we can compute the "bootstrap variance".
 - Draw B bootstrap samples.
 - Compute empirical variance of $\hat{\mu}(\mathcal{D}_n^1), \ldots, \hat{\mu}(\mathcal{D}_n^B)$...
- Could report

$$\hat{\mu}(\mathfrak{D}_n) \pm \sqrt{\mathsf{Bootstrap Variance}}$$

The Benefits of Averaging

A Lousy Estimator

- Let Z, Z_1, \ldots, Z_n i.i.d. $\mathbb{E}Z = \mu$ and $\text{Var}Z = \sigma^2$.
- We could use any single Z_i to estimate μ .
- Performance?
 - Unbiased: $\mathbb{E}Z_i = \mu$.
 - Variance of estimator would be σ^2 .

Variance of a Mean

- Let Z, Z_1, \ldots, Z_n i.i.d. $\mathbb{E}Z = \mu$ and $\text{Var}Z = \sigma^2$.
- Let's consider the average of the Z_i 's.
 - Average has the same expected value but smaller variance:

$$\mathbb{E}\left[\frac{1}{n}\sum_{i=1}^{n}Z_{i}\right] = \mu \qquad \operatorname{Var}\left[\frac{1}{n}\sum_{i=1}^{n}Z_{i}\right] = \frac{\sigma^{2}}{n}.$$

- Clearly the average is preferred to a single Z_i as estimator.
- Can we apply this to reduce variance of general decision functions?

Averaging Independent Prediction Functions

- Suppose we have *B* independent training sets from same distribution.
- Learning algorithm gives B decision functions: $\hat{f}_1(x), \hat{f}_2(x), \dots, \hat{f}_B(x)$
- Define the average prediction function as:

$$\hat{f}_{\text{avg}} = \frac{1}{B} \sum_{b=1}^{B} \hat{f}_b$$

• What's random here?

Averaging Independent Prediction Functions

- Fix some $x \in \mathfrak{X}$.
- Then average prediction on x is

$$\hat{f}_{avg}(x) = \frac{1}{B} \sum_{b=1}^{B} \hat{f}_b(x).$$

- Consider $\hat{f}_{avg}(x)$ and $\hat{f}_1(x), \ldots, \hat{f}_B(x)$ as random variables. (They are.)
- $\hat{f}_1(x), \ldots, \hat{f}_B(x)$ are i.i.d.
- $\hat{f}_{avg}(x)$ and $\hat{f}_b(x)$ have the same expected value, but
- $\hat{f}_{avg}(x)$ has smaller variance:

$$Var(\hat{f}_{avg}) = \frac{1}{B^2} Var \left(\sum_{b=1}^{B} \hat{f}_b(x) \right)$$
$$= \frac{1}{B} Var \left(\hat{f}_1(x) \right)$$

Averaging Independent Prediction Functions

Using

$$\hat{f}_{avg} = \frac{1}{B} \sum_{b=1}^{B} \hat{f}_b$$

seems like a win.

- But in practice we don't have B independent training sets...
- Instead, we can use the bootstrap....

Bagging

Bagging

- Draw B bootstrap samples D^1, \ldots, D^B from original data \mathfrak{D} .
- Let $\hat{f}_1, \hat{f}_2, \dots, \hat{f}_B$ be the decision functions for each set.
- The bagged decision function is a combination of these:

$$\hat{f}_{\mathsf{avg}}(x) = \mathsf{Combine}\left(\hat{f}_1(x), \hat{f}_2(x), \dots, \hat{f}_B(x)\right)$$

- How might we combine
 - decision functions for regression?
 - binary class predictions?
 - binary probability predictions?
 - multiclass predictions?
- Bagging proposed by Leo Breiman (1996).

Bagging for Regression

- Draw B bootstrap samples D^1, \ldots, D^B from original data \mathfrak{D} .
- Let $\hat{f}_1, \hat{f}_2, \dots, \hat{f}_B : \mathcal{X} \to \mathbf{R}$ be the predictions functions for each set.
- Bagged prediction function is given as

$$\hat{f}_{bag}(x) = \frac{1}{B} \sum_{b=1}^{B} \hat{f}_{b}(x).$$

- If bootstrap samples were independent draws from P,
 - $\hat{f}_{\mathsf{bag}}(x)$ would have the same expectation as $\hat{f}_1(x)$, but
 - $\hat{f}_{\text{bag}}(x)$ would have smaller variance.
- Empirically: Often get a similar effect for bagging.

Out-of-Bag Error Estimation

- Each bagged predictor is trained on about 63% of the data.
- Remaining 37% are called out-of-bag (OOB) observations.
- For ith training point, let

$$S_i = \{b \mid D^b \text{ does not contain } i\text{th point}\}.$$

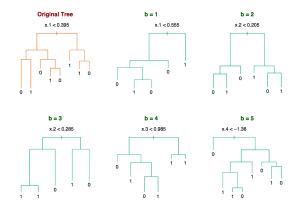
• The OOB prediction on x_i is

$$\hat{f}_{OOB}(x_i) = \frac{1}{|S_i|} \sum_{b \in S_i} \hat{f}_b(x).$$

- The OOB error is a good estimate of the test error.
- For large enough B, OOB error is like cross validation.

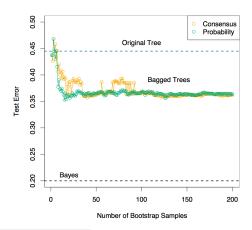
Bagging Trees

- Input space $\mathfrak{X}=\mathbf{R}^5$ and output space $\mathfrak{Y}=\{-1,1\}.$
- Sample size N = 30 (simulated data)



Bagging Trees

 Two ways to combine classifications: consensus class or average probabilities.



From ESL Figure 8.10

Terms "Bias" and "Variance" in Casual Usage

- ullet Restricting the hypothesis space \mathcal{F} "biases" the fit
 - towards a simpler model and
 - away from the best possible fit of the training data.
- Full, unpruned decision trees have very little bias.
- Pruning decision trees introduces a bias.
- Variance describes how much the fit changes across different random training sets.
- Decision trees are found to be high variance.

Conventional Wisdom on When Bagging Helps

- Bagging does nothing to eliminate bias.
- Hope is that bagging reduces variance.
- General sentiment is that bagging helps most when
 - Relatively unbiased base predictions
 - High variance
 - e.g. small changes in training set can cause large changes in predictions
- I'm not aware of solid theory on this...
- Empirical observation
 - Bagging trees works well.
 - Trees have high variance and low bias.
 - QED?

Random Forests

Recall the Motivating Principal of Bagging

- Averaging $\hat{f}_1, \ldots, \hat{f}_B$ reduces variance, if they're based on i.i.d. samples.
- Bootstrap samples are not indepedendent.
- This probably limits the amount of variance reduction we can get.
- Would be nice to reduce the dependence between \hat{f}_i 's...

Variance of a Mean of Correlated Variables

• For Z, Z_1, \ldots, Z_n i.i.d. with $\mathbb{E}Z = \mu$ and $\text{Var}Z = \sigma^2$,

$$\mathbb{E}\left[\frac{1}{n}\sum_{i=1}^{n}Z_{i}\right] = \mu \qquad \operatorname{Var}\left[\frac{1}{n}\sum_{i=1}^{n}Z_{i}\right] = \frac{\sigma^{2}}{n}.$$

- What if Z's are correlated?
- Suppose $\forall i \neq j$, $Corr(Z_i, Z_j) = \rho$. Then

$$\operatorname{Var}\left[\frac{1}{n}\sum_{i=1}^{n}Z_{i}\right]=\rho\sigma^{2}+\frac{1-\rho}{n}\sigma^{2}.$$

• For large n, the $\rho\sigma^2$ term dominates – limits benefit of averaging.

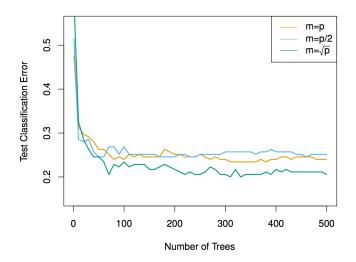
Random Forest

Main idea of random forests

Use **bagged decision trees**, but modify the tree-growing procedure to reduce the correlation between trees.

- Key step in random forests:
 - When constructing each tree node, restrict choice of splitting variable to a randomly chosen subset of features of size *m*.
- Typically choose $m \approx \sqrt{p}$, where p is the number of features.
- Can choose *m* using cross validation.

Random Forest: Effect of m size



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Random Forest: Effect of m size

 See movie in Criminisi et al's PowerPoint: http://research.microsoft.com/en-us/um/people/antcrim/ ACriminisi_DecisionForestsTutorial.pptx