

CS 70 HW #9 -

Sunday

worked alone

1) Monty Hall Revenge

a) switching in this scenario will help your odds of getting the car.

probability of picking car with 1st picks = $\frac{1}{n}$
↓

The host will always remove a non car option
that wasn't picked by player

↓

New probability of picking car with switching
picks is the chance of not picking the
car in the 1st picks times the chance
of picking it at this point

$$\hookrightarrow \underbrace{\frac{n-1}{n}}_{\text{chance w/ 1st picks}} \times \underbrace{\frac{1}{n-1-1}}_{\text{chance picking car after host}} = \frac{1}{n} \times \frac{n-1}{n-2}$$

chance w/ 1st picks
car w/ 1st picks

chance picking
car after host
eliminates bad option

$$\frac{1}{n} \times \frac{n-1}{n-2} \geq \frac{1}{n}$$


switch


non switch

 This shows that switch will never
hurt your odds

b) Switching will never hurt your odds in this scenario

(building from part a)

Probability of picking car with 1st pick = $\frac{1}{n}$



The host will always remove a non car door that wasn't picked by player, this time by $n-2$ doors



New probability of picking cur with switch
 picks is the chance of not picking the
 cur in the 1st picks times the chance
 of picking it at this point

$$\hookrightarrow \underbrace{\frac{n-1}{n}}_{\text{chance w/ picks}} \times \underbrace{\frac{1}{n-1-(n-2)}}_{\text{chance picking cur after last elements but option}} = \frac{1}{n} \times \frac{n-1}{1}$$

chance w/ picks
 cur w/ 1st picks

chance picking
 cur after last
 elements but option



$$\frac{1}{n} \times (n-1) \geq \frac{1}{n}$$

$\underbrace{}$

switch

$\underbrace{}$

non switch



This shows that switches will never
 hurt you $\sigma \geq 1$

c) Again based from a) and b)

probability of
picking car with = $\frac{k}{n}$

1st picks



The host will always remove a non-prize door
that wasn't picked by player, this
time by 5 doors (JL wins)



New probability of picking car with switch
 picks is the chance of not picking the
 car in the 1st picks times the chance
 of picking it at this point

$$\hookrightarrow \underbrace{\frac{n-k}{n}}_{\text{chance w/ picks}} \times \underbrace{\frac{k}{n-1-j}}_{\text{chance picking car after last elements but option}}$$

chance w/ picks
 car w/ 1st picks

chance picking
 car after last
 elements but option



We want to max 15 and min 5 in
 order for the highest ratio.

$$\hookrightarrow 15 \geq n-2$$

$$5 \leq \min(n - 15)$$

2) Men sprays Truth

a) $P = \frac{1}{3}$ $P_{\text{not } A} = \frac{3}{4}$

In other words:

$$P[A] \text{ Men sprays Truth} = \frac{3}{4}$$

$$P[B] : \text{Last on hands} = \frac{1}{3}$$

$$P[\bar{A}] = \frac{1}{4}$$

$$P[\bar{B}] = \frac{2}{3}$$

$$\hookrightarrow P[A|B] = \frac{P[B|A] P(A)}{P(B)}$$

↓ Bayes theorem

i)

$$\frac{\left(\frac{1}{3}\right)\left(\frac{3}{4}\right)}{\left(\frac{1}{3}\right)\left(\frac{2}{3}\right) + \left(\frac{2}{3}\right)\left(\frac{1}{4}\right)}$$

$$\hookrightarrow \frac{\frac{1}{4}}{\frac{1}{4} + \frac{1}{12}} \rightarrow \frac{1}{4} \times \frac{12}{16} = \frac{3}{5}$$

ii)
iii)

the numbers didn't change, it went up. Because the man was 112 to 75%, this increases our confidence that the man isn't lying given that he said he is being re-faulted, thus the intention is problematic.

b)

$$P(A) = \frac{1}{6} \quad P(B) = \frac{3}{4}$$

In other words:

$$P[A] \text{ Men speaks Truth} = \frac{3}{4}$$

$$P[B] : \text{Lies on } b = \frac{1}{6}$$

$$P[\bar{A}] = \frac{1}{4}$$

$$P[\bar{B}] = \frac{5}{6}$$

$$\hookrightarrow P[A|B] = \frac{P[B|A] P[A]}{P[B]}$$

↓ Bayes theorem

i)

$$\frac{\left(\frac{1}{6}\right)\left(\frac{3}{4}\right)}{\left(\frac{1}{6}\right)\left(\frac{3}{4}\right) + \left(\frac{5}{6}\right)\left(\frac{1}{4}\right)}$$

$$\hookrightarrow \frac{\frac{1}{8}}{\frac{1}{8} + \frac{5}{24}} \rightarrow \frac{\frac{1}{8}}{\frac{1}{3}} = \cancel{\frac{3}{8}}$$

iii) the pusher will dig deeper, it went up. Because the man was 1121 → 75 in, this increases our confidence that the man isn't lying given that he said he is taller than fourth, thus the killer is probably/vb.

3) Mario's Coins

- a) No, it is dependent. Since we are picking coins without replacement, and that all the coins have a different probability of Heads, this tells us that the last coin we pick will be dependent on the last coin we pick.
- b) No it isn't independent. Y_1 causes no such info on which coin is picked when directly influences Y_2 .

$$(1) \quad \frac{1}{4}, \quad \frac{1}{2}, \quad \frac{3}{4}$$

(1st 2nd 3rd)

↳ $P[3_{rd} \text{ wins} | 1st \text{ 2 wins}]$

$$= \frac{P\{A|1^{\text{st}} \text{ 3 wins}\}}{P[1^{\text{st}} \text{ 2 wins}]}$$

$$= \frac{\frac{1}{4} \left(\frac{1}{2}\right) \left(\frac{3}{4}\right)}{\sum_{\text{outcomes}} P\{1^{\text{st}} \text{ 2 wins} | \text{ outcome}\} P[\text{outcome}]}$$

$$= \frac{\frac{1}{4} \left(\frac{1}{2}\right) \left(\frac{3}{4}\right)}{\frac{1}{6} \left(\frac{1}{4} \times \frac{1}{2}\right) + \left(\frac{1}{4}\right) \left(\frac{3}{4}\right) + \left(\frac{1}{2}\right) \left(\frac{3}{4}\right) + \left(\frac{1}{4}\right) \left(\frac{3}{4}\right) \left(\frac{1}{2}\right)}$$

$$\text{↳ } \frac{3}{72} \times \frac{48}{16} > \frac{1}{22}$$

4) Symmetric Mixing

a) They are not independent. Since A_1 and A_2 are drawn without replacement, this means that A_2 is dependent on the outcome of A_1 , whether it's red or blue. Thus, they aren't independent.

b) Probability that Ruchael wins = Number wins / $P = N$

$$\hookrightarrow 1 - 2p = \text{total}$$

$\hookrightarrow \binom{8}{4}$ ways to have 4 wins out of $\binom{12}{4}$ total ways

$$\hookrightarrow \frac{\binom{8}{4} \binom{4}{2}}{\binom{12}{4}} = \frac{18}{35} \rightarrow 1 - 2p = \frac{18}{35}$$

$$\hookrightarrow p = \frac{17}{70}$$

$$c) \left(\frac{1}{2}\right)\left(\frac{3}{7}\right)\left(\frac{1}{3}\right)\left(\frac{1}{3}\right) = \frac{1}{60}$$

1st 2nd 3rd 4th

choose 4

R R R R

2) By symmetry, $P[\gamma_{11}] = P[1st]$

↳ Since 8 red numbers on 12 numbers
factual $\sim \frac{2}{3}$

e) we know 1s number on 4 dozen

↳ $\frac{15}{48}$ by symmetry

f) Only way to loss if all blue

$$\text{↳ } \frac{4}{12} \times \frac{3}{11} \times \frac{4}{10} \times \frac{2}{9} = \frac{8}{495}$$

↳ same that 1st rule up is $\frac{8}{12}$

$$\text{↳ } \frac{8}{495} \times \frac{3}{2} = \frac{4}{165} \quad \begin{array}{l} \text{same 1st} \\ \text{blue losses sum} \\ \text{1st rule up} \end{array}$$

$$\text{↳ complement} = \frac{161}{165} \quad \begin{array}{l} \text{same 1st} \\ \text{blue wins sum} \\ \text{1st rule up} \end{array}$$

5) Lookie Jav

$X = \# \text{ coins in } 1 \text{ jar}$

$Y = \# \text{ coins in other jar}$

→ After every step, (X, Y) will enter

where $(X-1, Y)$ or $(X, Y-1)$ with equal chance of either happening

↓

symmetric random walk in the plane starting at $(x, y) = (n, n)$

↳ suppose walk ends at $(k, 0)$, then previous position was $(k+1, 1)$.

To go from $(n, n) \rightarrow (k, 0)$ need

$$(n-k) + (n-1) = 2n - k + 1 \text{ steps}$$

of probability $\frac{1}{2}$ each.

↳ Exact Probability: $\left(\frac{1}{2}\right)^{2n-k+1}$

↓

$$\binom{2n-k}{n}$$

* Because of symmetry, we can add or remove the first term.

↳ Adding the $\frac{1}{2}$ term of final sign to $(k, 0)$,

We get the following

$$P\{ (k, 0) \} = \binom{2n-k}{n} \frac{1}{2^{15-2n}}$$

where $k \in \{0, \dots, n\}\}$

6.1 Testing Model Plans

a) A model performs n trials, each with independent probability p of success.

$$X_1 \sim \text{Binom}(n, p)$$

\downarrow

$$\Pr[X=k] = \binom{n}{k} p^k (1-p)^{n-k} \text{ for } 0 \leq k \leq n$$

b) Each model plan involves n probabilities of p^1 or survival between

$$X_2 \sim \text{Binom}(n, p^1)$$

\downarrow

$$\Pr[X=k] = \binom{n}{k} p^1 k (1-p^1)^{n-k} \text{ for } 0 \leq k \leq n$$

c) By calculating means n trials, we see when my probability of p^+ of surviving t days

$$X_+ \sim \text{Binom}(n, p^+)$$

\downarrow

$$\Pr[X=k] = \binom{n}{k} p^+ k (1-p^+)^{n-k} \text{ for } 0 \leq k \leq n$$

2) We consider the complement

↳ noticing that we must return after $\lceil \log \rceil$

↳ various parts, $P[x_t = 0] = (1-p^+)^n$



Probability of loss $1 - (1-p^+)^n$

c) NO, not invariant

↳ $P[A_1 \cap B_1]$ claim that first $\lceil \log \rceil$ nodes

$$\text{crash} = (1-p)^n. \text{ However, since}$$

$$P[A_1] P[B_1] = (1-p)^n, \text{ we consider}$$

two other very different situations

f) $P[A_1 \cap B_1]$ claim that lot, two nodes crash.

$$(1-p)^2 p^{n-2}. P[A_1]$$
 claim that lot

comes an extra $1 - n - 1$ node more

$$\text{comes, so } P[A_2] = (1-p) \binom{n-1}{1} (1-p)^{n-2}$$

$$\text{So from this, } n = \frac{1}{1-p} + 1$$

9) No, not independent

So if $k_1 = 0$ and $k_2 = 1$, then the solution would not be possible because there can't be 1 when at least 1 less than zero is zero plus one. This tells us that the 2 must not be independent.