

CS 70 HW¹ [3]

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Question 1

a) Using the given:

$$\begin{aligned} P(X \geq \alpha) &\leq \frac{E[(X+c)^2]}{(\alpha+c)^2} \\ &= \frac{E[X^2] + 2cE[X] + c^2}{(\alpha+c)^2} \\ &= \frac{(\sigma^2) + (c^2)}{(\alpha+c)^2} \end{aligned}$$

b) To solve this, we need to take the derivative and set it to 0 in respect to c :

$$\frac{d}{dc} \left(\frac{(\sigma^2) + (c^2)}{(\alpha+c)^2} \right) = 0$$

$$\frac{2c(\sigma + c)^2 - 2(\alpha + \gamma)(\sigma^2 + c^2) = 0}{(\alpha + c)^4}$$



$$2c(\sigma + c)^2 - 2(\alpha + \gamma)(\sigma^2 + c^2) = 0$$



$$\alpha c^2 + c(\alpha^2 - \sigma^2) - \sigma^2 \alpha = 0$$



$$c = \frac{\sigma^2}{\alpha}$$

c) i) It isn't always correct because there is losses in the variance we get, or rather, it won't always return so in 0.01

ii) We can get $RV: x$

x : one that's has higher probability than another

$$P(x=0) = 0.75 \quad P(x=1) = 0.25$$

$$W \propto = ?$$

$$\hookrightarrow E[X] = 0.25$$

$$\hookrightarrow \text{Var}_{\text{unum}} = 100 (0.25) (0.15)$$



$$P[X \geq E[X] + \gamma] = \frac{1}{4} > \frac{\text{Var}(X)}{2 \cdot \gamma^2} = 0.12$$

d) i) markov: $P(X \geq s) \leq \frac{E[X]}{s} = \frac{\gamma}{s}$

ii) chebyshev: $P(|X - E[X]| \geq \gamma) \leq \frac{\text{Var}[X]}{\gamma^2} = \frac{1}{2}$

iii) Since $E[X] = E[Y]$

$$\hookrightarrow Y = X - E[X] = X - 3$$

$$\hookrightarrow P(X \geq s) = P(Y \geq \gamma) \leq \frac{\text{Var}[Y]}{\gamma^2 + \text{Var}[Y]} = \frac{1}{3}$$

Gaußsche > Chebyshev'sche Unschärfe in order of tightness
of bounds

Question 2

Variation variance

x_i : number of trials needed to collect the i th monopoly
card given $(i=1)$ Monopoly cards already collected

$$p = (n-i+1)/n$$



$$\text{Var}(Cx_1) = \sum_{i=1}^n \text{Var}(Cx_i)$$

$$= \sum_{i=1}^n \frac{1 - (n-i+1)/n}{\left[(n-i+1)/n \right]^2}$$

$$= \sum_{j=1}^n \frac{1 - j/n}{\left[\sum_{i=j}^n 1/n \right]^2}$$

$$= \sum_{j=1}^n \frac{n(n-j)}{j^2}$$

$$= \sum_{j=1}^n \frac{n^2}{j^2} - \sum_{j=1}^n \frac{n}{j}$$

$$= n^2 \sum_{j=1}^n \frac{1}{j^2} - E[X^2]$$

Question 3

a) $\tan 4:$ $\text{Cov}(X_1, 2\bar{P}_1)$

$$\hookrightarrow \frac{2}{\bar{P}_1} = P_1 \rightarrow 2\bar{P}_1 = \bar{P}$$

$\tan 2:$ $\text{Cov}(4P_2)$

$$\hookrightarrow P_2 = \frac{\bar{P}}{4} \rightarrow 4P_2 = \bar{P}$$

$\tan 3:$ $\sqrt{\frac{6}{P_3}}$ $\text{Cov}(X_3)$

$$P_3 = \frac{6}{\bar{P}^2} \rightarrow \bar{P}^2 = \frac{6}{P_3} \rightarrow \bar{P} = \sqrt{\frac{6}{P_3}}$$

$$\text{b)} \quad \rho(|X_n - N| \geq \varepsilon) \leq \frac{1}{15^2}, \quad \frac{1}{15^2} >$$

$$\hookrightarrow N \geq \frac{\sigma^2}{\varepsilon^2}$$

$$\hookrightarrow N \geq \left(\frac{\bar{P}(4-\bar{P})}{10} \div c^4 \right)$$

$$\hookrightarrow N \geq \left(\frac{\bar{P}(4-\bar{P})}{10 \varepsilon^2 c^4} \right)$$

d) $i=1$

$$\hookrightarrow P[x_1 \sim n_1] = P[n_1 \sim \frac{2}{\pi}]$$

\hookrightarrow Using Chean inequlity \downarrow

$$(1 - \frac{2}{\pi})(1 - \sigma) / \epsilon_{12} \leq 2$$

$$\hookrightarrow N \geq \frac{2}{\sigma} \left(1 - \frac{2}{\pi} \right) / \epsilon_{12}$$

$i=2$

$$P[x_2 \sim n_2] = P[n_2 \sim \frac{6}{\pi^2}]$$

\hookrightarrow Using Chean inequlity \downarrow

$$(1 - \sigma/2)(1 - \sigma_2) / \epsilon_{12} \leq 2$$

$$\hookrightarrow N \geq \frac{\sigma^2}{6} \left(1 - \frac{\sigma^2}{6} \right) / \epsilon_{12}$$

or from this, we can see that tasknum 1 has the lowest error tasknum 2 has the highest N value

Question 4

a) Because x is uniform

$$\hookrightarrow f_{X \sim U} = \begin{cases} 1/2 & 0 \leq x \leq 2 \\ 0 & \text{otherwise} \end{cases} \quad \text{PDF}$$

$$\text{CDF} \quad F(x) = P[X \leq x] = \begin{cases} 0; & x < 0 \\ x/2; & 0 \leq x \leq 2 \\ 1; & x \geq 2 \end{cases}$$

Now looking at y

$$\hookrightarrow y = u_x^2 + 1$$

$$\text{CDF: } P(X \leq \frac{\sqrt{y-1}}{2}) = F_y(y)$$

$$\hookrightarrow F_y(t) = \begin{cases} 0; & y < 1 \\ \frac{\sqrt{y-1}}{2}; & 1 \leq y \leq 18 \\ 1; & y > 18 \end{cases}$$

PDF(y)!

$$f_y(y) = \frac{d}{dy}(F_y(y))$$

$$= \begin{cases} \frac{1}{4\sqrt{y-1}}; & 1 \leq y \leq 18 \\ 0; & \text{otherwise} \end{cases}$$

Variance / Erwartung:

$$E(x) = \frac{0+1}{2} = 1$$

$$\text{Var}(x) = \frac{1}{12} (2-1)^2 \rightarrow \frac{1}{3}$$

Für \tilde{x} :

$$E(\tilde{x}) = E(4x^2 + 1) \rightarrow 16E(x^2) + E(1)$$

↳ 16(1) + 1

$$\hookrightarrow 17$$

$$\text{Var}(\tilde{x}) = \text{Var}(4x^2 + 1) = 18 \text{ Var}(x)$$

$$\hookrightarrow 6 = \text{Var}(\tilde{x})$$

ii) $f_{x,y}(x,y) = \begin{cases} cx + y & ; x \in \{1,2\} \\ 0 & ; y \in \{0,2\} \\ & ; \text{otherwise} \end{cases}$

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{x,y}(x,y) dx dy = 1$$

$$\hookrightarrow \int_{-\infty}^{\infty} \left(2cx + \frac{1}{2} \right) dx = 1$$

$$\hookrightarrow 2c \left(\frac{x^2}{2} \right)_1 + \left(\frac{x}{2} \right)_1^2 = 1$$

$$\hookrightarrow 2c \left(2 - \frac{1}{2} \right) + \left(1 - \frac{1}{2} \right) = 1$$

$$\hookrightarrow c = \frac{1}{2}$$

$$\hookrightarrow c = \frac{1}{6}$$

~~x and y are independent when we will take the joint shown outcome in intersection~~

$$c) X \sim \text{Exp}(\lambda)$$

$$\text{which means } \lambda = \gamma$$

$$i) Y = \lceil X \rceil \rightarrow \text{Geom}(n) \text{ where } p = 1 - e^{-\lambda}$$

$$ii) \text{ Thus } P_X(Y=1) = (1-p)^{\lfloor 0 \rfloor} p$$

↳ when $p = 1 - e^{-\lambda}$ so in this case, Geom distribution
and it is geometric distribution

$$d) X_i \sim \text{Exp}(\lambda_i) \text{ for } i \text{ mutually independent}$$

$$P\left[\min(X_1, \dots, X_n) > x\right]$$

$$\prod_{i=1}^n \exp(-x\lambda_i)$$

$$\Rightarrow \exp\left(-x \sum_{i=1}^n \lambda_i\right)$$

$$\min(X_1, \dots, X_n) \sim \text{Exp}\left(\sum_{i=1}^n \lambda_i\right)$$

Question 5

on x : $P[x \in (a, b)] \geq 0$; $a, b \in \mathbb{R}$

$\hookrightarrow x \sim N(0, 1)$ would be an example that fits the more

$$\hookrightarrow P[x \in (a, b)] = F(b) - F(a) \geq 0 \quad \text{if } b > a$$

Since this is an increasing function, we know it

will invertibly fullfill the given conditions

b) The function is increasing due to the following:

$$f(x+c) > f(x); \text{ for any } c$$

$$= P[x < x+c] - P[x < x]$$

$$= P[x < x] + P[x + (x, x+c)] - P[x < x]$$

$$= P[x + (x, x+c)] > 0$$

\hookrightarrow For $+w$, we can see that the function

\hookrightarrow is strictly increasing

c) $\text{Unif}(\Theta, \cdot)$

$$F^{-1}(v) = Y$$

$$\begin{aligned} \hookrightarrow F_Y(y) &= P[Y \leq y] \\ &= P[F^{-1}(v) \leq y] \\ &= P[F(F^{-1}(v)) \leq F(y)] \end{aligned}$$

$$= P[v \leq F(y)]$$

$$= F(y) = P[X \leq y]$$

$$= F_X(y)$$

\hookrightarrow From this, we can see that the joint comes out like this

\exists $x \in \text{domain}$ van de variabelen

wy $P[x = x_i] = p_i$

i) gevraag uit $(0, 1)$

(1) $y = \begin{cases} x_1 & \text{if } 0 \leq v < p_1 \\ x_2 & \text{if } p_1 \leq v < p_1 + p_2 \\ \vdots & \vdots \\ x_n & \text{if } p_1 + \dots + p_{n-1} \leq v < p_1 + \dots + p_n \end{cases}$

↳ $y = x_i$ wy probabilit $= P[p_1 + \dots + p_i]$

↳ $y = x$

→ By this method, we can find x

Question b)

$$\begin{aligned}
 a) P[Y_i < x_i] &= \int_{t=0}^{\infty} P[X_i = t \wedge Y_i < x_i] \\
 &= \int_{t=0}^{\infty} P[X_i = t] P[Y_i < x_i] \\
 &= \int_{t=0}^{\infty} \lambda e^{-\lambda t} (1 - e^{-Nt}) \\
 &= \int_{t=0}^{\infty} \lambda e^{-\lambda t} - \lambda e^{-(\lambda + N)t} + \\
 &= \int_{t=0}^{\infty} \lambda e^{-\lambda t} - \int_{t=0}^{\infty} \lambda e^{-(\lambda + N)t} + \\
 &= 1 - \frac{\lambda}{\lambda + N} \\
 &= \frac{N}{N + \lambda}
 \end{aligned}$$

$$b) P(D >) \stackrel{D}{=} P(X > 20 + 1 | X \geq 20)$$

\downarrow Bayes Rule

$$\begin{aligned}
 P(X > 20 + 1 | X \geq 20) &= \frac{P(X > 20 + 1)}{P(X \geq 20)} \\
 &\geq \frac{e^{N(20+1)}}{e^{20N}} \\
 &= e^N
 \end{aligned}$$

L) CDF Distribution: $(1 - e^{-\lambda t})^n$

* It is exponential(1) distribution w/ parameter $n\lambda$

c) Whenever among first n events until t null happens:

$$Z = \min(X, Y)$$

$$P(Z > t) = P(X > t \cap Y > t)$$

$$= P(X > t) P(Y > t)$$

$$= (1 - P(X \leq t))(1 - P(Y \leq t))$$

$$= (1 - (1 - e^{-\lambda t}))(1 - (1 - e^{-\mu t}))$$

$$= e^{-\lambda t} e^{-\mu t}$$

$$= e^{-(\lambda + \mu)t}$$

L) CDF follows for Z : $P(Z \leq t) = 1 - e^{-(\lambda + \mu)t}$

L) Z exponential w/ parameter $n\lambda + \mu$

d) if $t > 0$

\hookrightarrow if $x_1 + x_2 \leq t \rightarrow x_1, x_2 \geq 0 \rightarrow \begin{array}{l} x_1 \leq t \\ x_2 \leq t - x_1 \end{array}$

\downarrow Substitution

$P(T \leq t) =$

$$\begin{aligned} P(x_1 \leq t, x_2 \leq t - x_1) &= \int_0^t \int_0^{t-x_1} \lambda \exp(-\lambda x_1) (\lambda \exp(-\lambda x_2))^2 dx_1 dx_2 \\ &= \lambda^2 \int_0^t \exp(-\lambda x_1) \cdot \frac{1 - \exp(-\lambda(t - x_1))}{\lambda} dx_1 \\ &= \lambda \int_0^t \exp(-\lambda x_1) - \exp(-\lambda(t - x_1)) dx_1 = \lambda \left[\frac{1 - \exp(-\lambda t)}{\lambda} - \exp(-\lambda t) \right] \end{aligned}$$

Dif. (continuity) \Rightarrow CDF

$$F_V(t) = \frac{d}{dt} P(T \leq t) = \lambda \exp(-\lambda t) - \lambda \exp(-\lambda t) + \lambda^2 t \exp(-\lambda t)$$

$$= \lambda^2 t \exp(-\lambda t)$$