

ISyE 6739 –Statistical Methods Review of Probability (Chapters 3, 4 & 7)

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List of Topics

- Review of Probability
- Discrete Distributions
 - Binomial
 - Poisson
 - Geometric
 - Negative-Binomial
- Continuous Distributions
 - Exponential
 - Normal
 - Chi-square
 - Student's t
 - F

Review of Probability Distributions

Random Variables

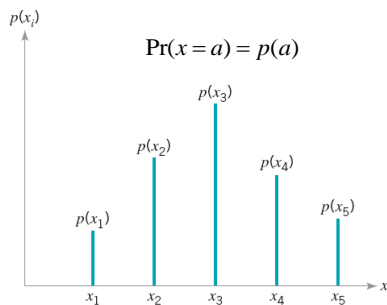
- **Random Experiment:** An experiment that can result in different outcomes, even if repeated in the same manner every time, e.g., rolling a dice, producing a part.
- **Random Variable:** A variable that associates a number with the outcome of a random experiment is called a random variable.
 - X: Compressive strength of each specimen
 - Y: Number of defectives in a sample of n produced parts
- Type of RVs:
 - A **discrete** random variable is a rv with a finite (or countably infinite) range. They are usually integer counts, e.g., number of errors or number of bit errors per 100,000 transmitted (rate).
 - A **continuous** random variable is a rv with an interval (either finite or infinite) of real numbers for its range. Its precision depends on the measuring instrument.

Discrete Distributions

Discrete distributions. When the parameter being measured can only take on certain values, such as the integers 0, 1, 2, . . . , the probability distribution is called a *discrete distribution*.

Examples: Quality of a part (Ok = 0 , Not ok = 1), Number of typos on a typed page (0,1,2, ...)

Probability mass function (pmf): $p(x)$ is pmf $\leftrightarrow \begin{cases} p(x) \geq 0 \\ \sum_{-\infty}^{+\infty} p(x) = 1 \end{cases}$



Examples: X: Number of observed 6 in 2 rolls of a dice

a	0	1	2
Pr(X=a)	$(5/6)^2$	$2(1/6)^1(5/6)^1$	$(1/6)^2$

$$\Pr(X=a) = \binom{2}{a} \left(\frac{1}{6}\right)^a \left(\frac{5}{6}\right)^{2-a}; \quad a=0,1,2$$

Cumulative Distribution Function (CDF)

Cumulative Distribution Function (CDF): The CDF of a discrete random variable X is denoted by $F(X=a)$ and is calculated by

$$F(X=a) = F(a) = \Pr(x \leq a) = \sum_{i=-\infty}^a p(i)$$

Some Properties:

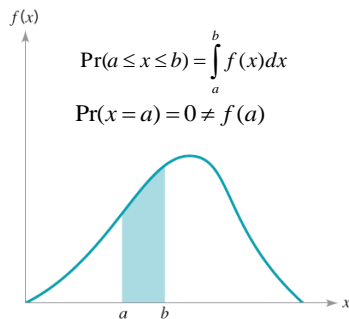
- 1) $0 \leq F(a) \leq 1$
- 2) if $x \leq y$, then $F(x) \leq F(y)$
- 3) $\Pr(a < x \leq b) = \Pr(x \leq b) - \Pr(x \leq a) = F(b) - F(a)$
- 4) $\Pr(a \leq x \leq b) = \Pr(x \leq b) - \Pr(x \leq a-1) = F(b) - F(a-1)$

Continuous Distributions

Continuous distributions. When the variable being measured is expressed on a continuous scale, its probability distribution is called a *continuous distribution*.

Examples: Diameter of piston rings, viscosity of oil, customers' Waiting time

Probability density function (pdf):



$$f(x) \text{ is pdf} \leftrightarrow \begin{cases} f(x) \geq 0 \\ \int_{-\infty}^{+\infty} f(x)dx = 1 \end{cases}$$

Cumulative distribution function (cdf):

$$F(X = a) = F(a) = \Pr(x \leq a) = \int_{-\infty}^a f(x)dx$$

Example: The probability distribution of X is

$$f(x) = 2e^{-kx}, \quad x \geq 0;$$

- (a) Find the appropriate value of k;
- (b) Find the $\Pr(X > 0.5)$.

Example: The product quality characteristic measurement is a random variable X with probability distribution

$$f(x) = \begin{cases} 4(x-11.75) & 11.75 \leq x \leq 12.25 \\ 4(12.75-x) & 12.25 < x \leq 12.75 \end{cases}$$

If the quality specification limit is that X must be less than 12.50, what is the percentage of products to be nonconforming?

- a) 0.125
- b) 0.5
- c) 0.875

Mean, Variance and Moments

- Mean** is a measure of the center of a probability distribution.

$$E(x) = \mu = \sum_{-\infty}^{+\infty} xp(x) \quad \text{Discrete RVs}$$

$$E(x) = \mu = \int_{-\infty}^{+\infty} xf(x)dx \quad \text{Continuous RVs}$$

- Moments** are measures related to the location/shape of probability distribution.

$$E(x^r) = \sum_{-\infty}^{+\infty} x^r p(x) \quad \text{Discrete RVs}$$

$$E(x^r) = \int_{-\infty}^{+\infty} x^r f(x)dx \quad \text{Continuous RVs}$$

- Variance** is a measure of the dispersion or variability of a probability distribution.

$$V(x) = \sigma^2 = \sum_{-\infty}^{+\infty} (x - \mu)^2 p(x) \quad \text{Discrete RVs}$$

$$V(x) = \sigma^2 = \int_{-\infty}^{+\infty} (x - \mu)^2 f(x)dx \quad \text{Continuous RVs}$$

$$V(x) = E(x^2) - [E(x)]^2$$

Example: Find mean and variance for the following RVs

a) X: Number of observed 6 in 2 rolls of a dice

a	0	1	2
Pr(X=a)	$(5/6)^2$	$2(1/6)^1(5/6)^1$	$(1/6)^2$

b) X: exam time limited between 0 and 1 hour.

$$f(x) = \alpha(x)^{\alpha-1}; 0 < x < 1$$

Useful Results on Mean and Variance

If x is a random variable and a is a constant, then

$$E(a+x) = a + E(x)$$

$$E(a \times x) = aE(x)$$

$$V(a+x) = V(x)$$

$$V(a \times x) = a^2 V(x)$$

If x_1, x_2, \dots, x_n are random variables,

$$E(a_1 x_1 + \dots + a_n x_n) = a_1 E(x_1) + \dots + a_n E(x_n)$$

If they are **mutually independent**, and a_1, \dots, a_n are constants

$$V(a_1 x_1 + \dots + a_n x_n) = a_1^2 V(x_1) + \dots + a_n^2 V(x_n)$$

Review of Discrete Distributions

Bernoulli Distribution

$X \sim \text{Bernoulli}(p)$ $x = 0,1$ p : probability of success

Application: To model the outcome of a Bernoulli trial with two possible outcomes (Success = 1, Failure = 0).

Example: The outcome of inspecting part quality;
Nonconforming (Success=1), Conforming (Failure=0)

pmf, mean and variance:

$$P(x) = p^x(1-p)^{1-x}; \quad x = 0,1$$

$$E(x) = p$$

$$\text{Var}(x) = p(1-p)$$

Binomial Distribution

$X \sim \text{Bin}(n, p)$ $x = 0, 1, \dots, n$ p : probability of success

Application: To model the number of successes in n independent Bernoulli trials.

Example: number of nonconforming parts in a random sample of parts of size n .

pmf, mean and variance:

$$P(x) = \binom{n}{x} p^x (1-p)^{n-x}; \quad x = 0, 1, \dots, n$$

$$E(x) = np \qquad \text{Var}(x) = np(1-p)$$

Fact: sum of n independent Bernoulli variables (p) is Binomial (n, p)

Examples

- (1) In a welding process, 10 different points of a battery tab are welded. If the probability of making an acceptable weld is 0.99, calculate :
- a) the probability of making a battery tab with 6 failed welds.
 - b) the probability of making a battery with at least 2 failed welds.

Poisson Distribution

$X \sim \text{Poisson}(\lambda)$ $x = 0, 1, 2, \dots$ λ : **average** number of events in a unit (a.k.a. rate)

Application: To model the number of events in a desired unit

Example: number of scratches on the surface of a refrigerator .
 events unit

pmf, mean and variance:

$$P(x) = \frac{\lambda^x \exp(-\lambda)}{x!}; \quad x = 0, 1, 2, \dots \quad E(x) = \text{Var}(x) = \lambda$$

Fact: sum of n independent Poisson variables (λ) is Poisson with the rate of $(n\lambda)$

Examples

- (2) In a painting shop, on average, there are 5 defects on the body of a painted car. Calculate
- a) the probability of finding a painted body with 10 defects.
 - b) the probability of finding 10 defects on 2 painted bodies.

Geometric Distribution

$X \sim \text{Geo}(p)$ $x = 1, 2, \dots$ p : probability of success

Application: To model the number of Bernoulli trials until the first success

Example: number of inspected part until the first nonconforming part is found

pmf, mean and variance:

$$P(x) = p(1-p)^{x-1}; \quad x = 1, 2, \dots$$
$$E(x) = \frac{1}{p}$$
$$\text{Var}(x) = \frac{1}{p} \left(\frac{1}{p} - 1 \right)$$

Fact: the number of Bernoulli trials between two successes also follows a Geometric distribution (because of Memoryless property)

Negative-Binomial (Pascal) Distribution

$X \sim \text{NB}(r, p)$ $x = r, r+1, \dots$ p : probability of success

Application: To model the number of Bernoulli trials until r successes occur

Example: number of inspected part until the r th nonconforming part is found

pmf, mean and variance:

$$P(x) = \binom{x-1}{r-1} p^r (1-p)^{x-r}; \quad x = r, r+1, \dots$$
$$E(x) = \frac{r}{p}$$
$$\text{Var}(x) = \frac{r}{p} \left(\frac{1}{p} - 1 \right)$$

Fact: sum of r independent Geometric variables (p) is Negative-Binomial (r, p)

Examples

In a welding process, probability of making an acceptable weld is 0.99, calculate :

- a) the probability of making at least 5 welds until a defective weld is made.
- b) the probability of making the 3rd defective weld in the 5th battery tab.

Review of Continuous Distributions

Exponential Distribution

$$X \sim \text{Exp}(\lambda) \quad x \geq 0$$

$$f(x) = \lambda \exp(-\lambda x); \quad x \geq 0$$

Example: waiting time of customers,
time to failure of a machine

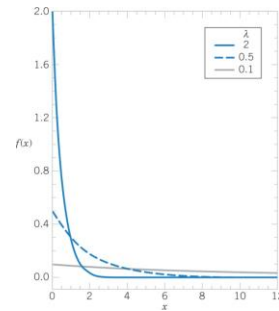
CDF, mean and variance:

$$E(x) = \frac{1}{\lambda}$$

$$\text{Var}(x) = \frac{1}{\lambda^2}$$

$$F(a) = \Pr(X \leq a) = 1 - \exp(-a\lambda)$$

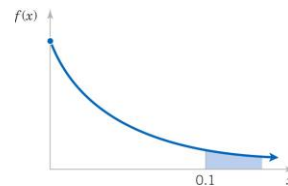
Memoryless property: $\Pr(x > a + b | x > a) = \Pr(x > b)$



Example 4-21: Computer Usage-1

In a large corporate computer network, time to user log-ons to the system can be modeled as exponential distribution with a mean of 0.04 hours. What is the probability that there are no log-ons in the next 6 minutes (0.1 hours)?

Let X denote the time in hours from the start of the interval until the first log-on.



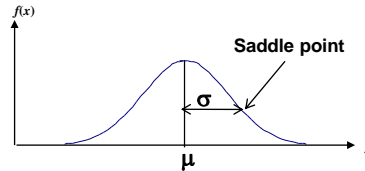
what is the interval of time such that the probability that no log-on occurs during the interval is 0.90?

Normal Distribution

$$X \sim N(\mu, \sigma^2); \quad -\infty < x < +\infty$$

$$f(x; \mu, \sigma) = \frac{1}{\sqrt{2\pi}\sigma} e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

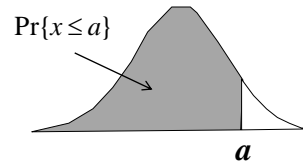
$$E(x) = \mu \quad \text{Var}(x) = \sigma^2$$



Fact: If x_1, x_2 are independently normally distributed variables, then $y=x_1+x_2$ also follows the normal distribution, i.e. $y \sim N(\mu_1+\mu_2, \sigma_1^2 + \sigma_2^2)$

$$\Pr\{x \leq a\} = \int_{-\infty}^a \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right\} dx$$

?



Standard Normal Distribution

$$X \sim N(\mu, \sigma^2); \quad -\infty < x < +\infty$$

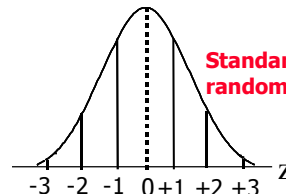
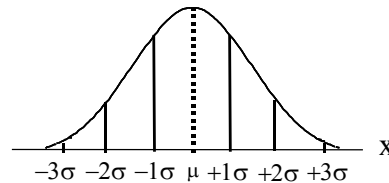
$$f(x; \mu, \sigma^2) = \frac{1}{\sqrt{2\pi}\sigma} \exp\left\{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2\right\}$$

Map any X into Z

$$Z = \frac{X - \mu}{\sigma}$$

$$Z \sim N(0, 1^2); \quad -\infty < z < +\infty$$

$$f(z; \mu = 0, \sigma^2 = 1) = \frac{1}{\sqrt{2\pi}} \exp\left(-\frac{z^2}{2}\right)$$



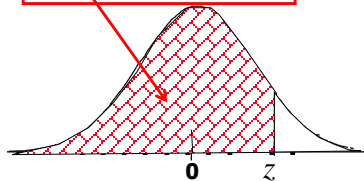
Standard normal random variable

Excel Function: $F(x) = \text{NORMDIST}(x, \mu, \sigma, \text{true})$

$\Pr(\mu - \sigma \leq x \leq \mu + \sigma) = 68.26\%$
 $\Pr(\mu - 2\sigma \leq x \leq \mu + 2\sigma) = 95.46\%$
 $\Pr(\mu - 3\sigma \leq x \leq \mu + 3\sigma) = 99.73\%$

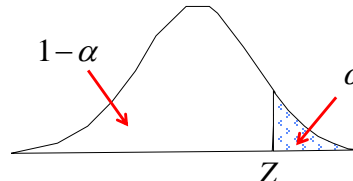
Z-Values and CDF

$$\Phi(z) = P(Z \leq z)$$



Input: real number (z)
Output: probability $\Phi(z)$

$$Z_\alpha = \Phi^{-1}(1 - \alpha) \Leftrightarrow P(Z > Z_\alpha) = \alpha$$



Input: probability α
Output: real number Z_α

$$\Phi(1.645) = P(Z < 1.645) = 0.95 \Leftrightarrow z_{0.05} = \Phi^{-1}(0.95) = 1.645$$

Useful equations:

$$\Phi(Z_\alpha) = P(Z < Z_\alpha) = 1 - \alpha \Rightarrow \alpha = 1 - \Phi(Z_\alpha)$$

$$\Phi(-z) = 1 - \Phi(z)$$

$$Z_{1-\alpha} = -Z_\alpha$$

Example $X \sim N(40, 5^2)$

$$p(x \leq 37.9) =$$

$$p(x \geq a) = 0.3783$$

$$\Phi(z) = \int_{-\infty}^z \frac{1}{\sqrt{2\pi}} e^{-u^2/2} du$$



Appendix A: Table III

z	0.00	0.01	0.02	0.03	0.04	z
0.0	0.50000	0.50399	0.50798	0.51197	0.51595	0.0
0.1	0.53983	0.54379	0.54776	0.55172	0.55567	0.1
0.2	0.57926	0.58317	0.58706	0.59095	0.59483	0.2
0.3	0.61791	0.62172	0.62551	0.62930	0.63307	0.3
0.4	0.65542	0.65910	0.66276	0.66640	0.67003	0.4
0.5	0.69146	0.69497	0.69847	0.70194	0.70540	0.5
0.6	0.72575	0.72907	0.73237	0.73565	0.73891	0.6
0.7	0.75803	0.76115	0.76424	0.76730	0.77035	0.7

Example: Three shafts are made and assembled in a linkage. The length of each shaft, in centimeters, is distributed as follows:

Shaft 1: $N \sim (75, 0.09)$

Shaft 2: $N \sim (60, 0.16)$

Shaft 3: $N \sim (25, 0.25)$

Assume the shafts' length are independent to each other:

(a) What is the distribution of the linkage?

(b) What is the probability that the linkage will be longer than 160.5 cm?

Chi-square Distribution

$$X \sim \chi^2(\nu)$$

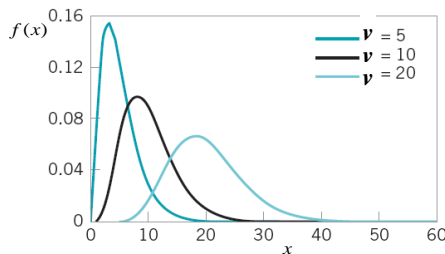
Degrees of freedom

$$E(X) = \nu$$

$$Var(X) = 2\nu$$

$$f(x; \nu) = \frac{1}{2^{\nu/2} \Gamma(\nu/2)} \exp\left(-\frac{x}{2}\right) x^{\left(\frac{\nu}{2}-1\right)}$$

$$\Gamma(a) = (a-1)!; a \in \mathbb{Z}^+$$



The Chi-square pdf is not easily integrable (except for $\nu=2$). The CDF table (Appendix A, Table IV) is usually used.

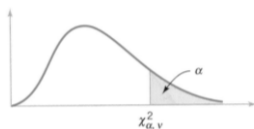


Table IV Percentage Points $\chi^2_{\alpha, \nu}$ of the Chi-Squared Distribution

$\nu \backslash \alpha$.995	.990	.975	.950	.900	.500	.100	.050	.025	.010	.005
1	.004	.005	.008	.010	.015	.455	2.706	3.841	5.024	6.635	7.879
2	.010	.015	.020	.025	.035	1.385	4.605	5.991	7.378	9.210	10.597
3	.016	.020	.025	.030	.040	2.366	6.251	7.879	9.348	11.345	12.838
4	.020	.025	.030	.035	.045	3.347	7.779	9.488	11.143	13.277	14.860
5	.024	.029	.035	.040	.050	4.351	9.236	11.070	12.833	15.086	16.750
6	.027	.031	.037	.042	.052	5.349	10.645	12.592	14.449	16.812	18.548
7	.029	.033	.039	.044	.054	6.344	12.017	14.067	16.013	18.475	20.278
8	.031	.035	.041	.046	.056	7.344	13.362	15.508	17.535	20.090	21.955
9	.032	.036	.042	.047	.057	8.344	14.682	16.919	19.023	21.667	23.589
10	.033	.037	.043	.048	.058	9.348	15.987	18.307	20.483	23.210	25.188

Chi-square Distribution

- **Fact 1:** If Z_1, \dots, Z_k are *independent, standard normal* random variables, then the sum of their squares is distributed according to the chi-square distribution with k degrees of freedom.

$$X = \sum_{i=1}^k Z_i^2 \sim \chi^2(k)$$

- **Fact 2:** If X_i 's ($i = 1, 2, \dots, n$) are *independent, Chi-square* random variables with ν_i 's degrees of freedom, then

$$Y = \sum_{i=1}^n X_i \sim \chi^2\left(\sum_{i=1}^n \nu_i\right)$$

Example

- Calculate the following quantities:

1. $\chi_{0.05,5}^2$

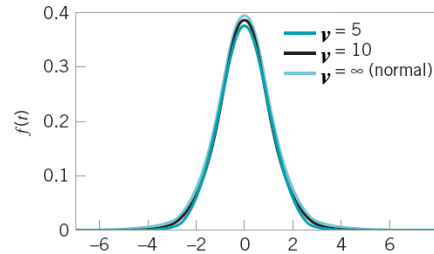
2. $\Pr\left\{\sum_{i=1}^8 \left(\frac{X_i - \mu}{\sigma}\right)^2 > 1.34\right\}$

Student's t Distribution

$$f(t; \nu) = \frac{1}{\sqrt{\pi\nu}} \frac{\Gamma(\frac{\nu+1}{2})}{\Gamma(\nu/2)} \left(1 + \frac{t^2}{\nu}\right)^{-\left(\frac{\nu+1}{2}\right)}$$

$$E(t) = 0$$

$$Var(t) = \frac{\nu}{\nu-2}; \nu > 2$$



The student's t pdf is not easily integrable (except for $\nu = 1$). The CDF table (Appendix A, Table V) is usually used.



Table V Percentage Points $t_{\alpha, \nu}$ of the t Distribution

$\nu \backslash \alpha$.40	.25	.10	.05	.025	.01	.005	.0025	.001	.0005
1	.325	1.000	3.078	6.314	12.706	31.821	63.657	127.32	318.31	636.62
2	.289	.816	1.886	2.920	4.303	6.965	9.925	14.089	23.326	31.598
3	.277	.765	1.638	2.353	3.182	4.541	5.841	7.453	10.213	12.924
4	.271	.741	1.533	2.132	2.776	3.747	4.604	5.598	7.173	8.610
5	.267	.727	1.476	2.015	2.571	3.365	4.032	4.773	5.893	6.869
6	.265	.718	1.440	1.943	2.447	3.143	3.707	4.317	5.208	5.959
7	.263	.711	1.415	1.895	2.365	2.998	3.499	4.029	4.785	5.408
8	.262	.706	1.397	1.860	2.306	2.896	3.355	3.833	4.501	5.041
9	.261	.703	1.383	1.833	2.262	2.821	3.250	3.690	4.297	4.781
10	.260	.700	1.372	1.812	2.228	2.764	3.169	3.581	4.144	4.587

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Student's t Distribution

- **Fact 1:** If
 - Z is a *standard normal* ($N(0,1)$) random variable
 - V has a *chi-square distribution* with ν degrees of freedom;
 - Z and V are *independent*.

Then the following ratio follows a student's t

$$\frac{Z}{\sqrt{V/\nu}} \sim t(\nu)$$

- **Fact 2:** the student's t distribution tends to standard normal distribution as ν increases ($\nu > 30$)

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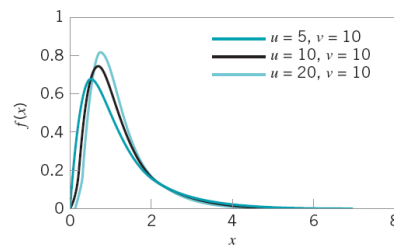
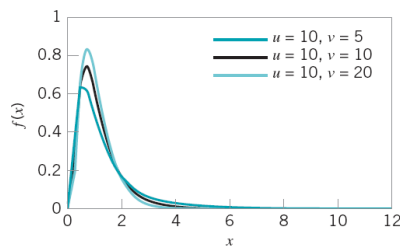
Example

- Calculate the following quantities:

1. $t_{0.05,5}$

2. $\Pr\left(\frac{Z}{\sqrt{\left(\frac{X - \mu_x}{\sigma_x}\right)^2}} > 1\right)$ Z follows a standard normal dist. and X follows a normal dist. and they are independent

F Distribution



$$f(x) = \frac{\Gamma\left(\frac{u+v}{2}\right)\left(\frac{u}{v}\right)^{u/2}}{\Gamma\left(\frac{u}{2}\right)\Gamma\left(\frac{v}{2}\right)} \frac{x^{(u/2)-1}}{\left[\left(\frac{u}{v}\right)x + 1\right]^{(u+v)/2}}, 0 < x < \infty$$

F Distribution

The F pdf is not easily integrable. The CDF table (Appendix A, Table VI) is usually used.

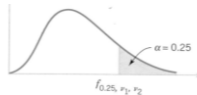


Table VI Percentage Points f_{α, v_1, v_2} of the F Distribution

		f_{α, v_1, v_2}															
v_2	v_1	Degrees of freedom for the numerator (v_1)															
		1	2	3	4	5	6	7	8	9	10	12	15	20			
1	1	5.83	7.50	8.20	8.58	8.82	8.98	9.10	9.19	9.26	9.32	9.41	9.49	9.58			
2	1	2.57	3.00	3.15	3.23	3.28	3.31	3.34	3.35	3.37	3.38	3.39	3.41	3.43			
3	1	2.02	2.28	2.36	2.39	2.41	2.42	2.43	2.44	2.44	2.44	2.45	2.46	2.46			
4	1	1.81	2.00	2.05	2.06	2.07	2.08	2.08	2.08	2.08	2.08	2.08	2.08	2.08			
5	1	1.69	1.85	1.88	1.89	1.89	1.89	1.89	1.89	1.89	1.89	1.89	1.89	1.89			
6	1	1.62	1.76	1.78	1.79	1.79	1.78	1.78	1.78	1.77	1.77	1.77	1.76	1.76			
7	1	1.57	1.70	1.72	1.72	1.71	1.71	1.70	1.70	1.70	1.69	1.68	1.68	1.67			
8	1	1.54	1.66	1.67	1.66	1.66	1.65	1.64	1.64	1.63	1.63	1.62	1.62	1.61			
9	1	1.51	1.62	1.63	1.63	1.62	1.61	1.60	1.60	1.59	1.59	1.58	1.57	1.56			
10	1	1.49	1.60	1.60	1.59	1.59	1.58	1.57	1.56	1.56	1.55	1.54	1.53	1.52			

- A random variate of the F-distribution arises as the ratio of two chi-squared variates: $\frac{U_1/d_1}{U_2/d_2}$ where
 - U_1 and U_2 have *chi-square distributions* with d_1 and d_2 degrees of freedom respectively, and
 - U_1 and U_2 are independent.

Example

- Calculate the following quantities:

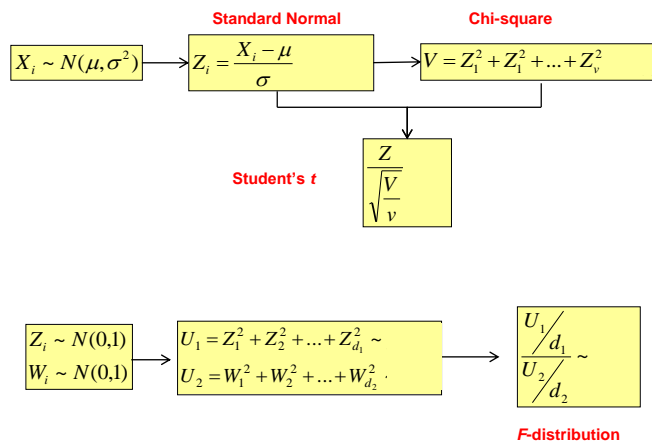
1. $F_{0.25, 6, 8}$

2. $\Pr\left(\frac{\sum_{i=1}^2 \left(\frac{X_i - \mu_x}{\sigma_x}\right)^2}{\sum_{i=1}^2 \left(\frac{Y_i - \mu_y}{\sigma_y}\right)^2} > 3\right)$

X and Y are independent and follow a normal dist.

Excel Function: $\text{f}(f_0)=\text{fdist}(f_0, df_1, df_2)$

Summary



Useful Fact About Sampling Distributions

- Fact 1:** A Chi-square random variable with 2 dof is equivalent to an exponential random variable with $\lambda = 0.5$.

$$f(x) = \lambda e^{-\lambda x}, \quad x \geq 0;$$

- Fact 2:** A t random variable with 1 dof is equivalent to a Cauchy random variable

$$f(t) = \frac{1}{\pi} \times \frac{1}{1+t^2}$$

- Fact 3:** An F random variable with 2 and 2 dof's

$$f \sim F_{2,2} \rightarrow g(f) = \frac{1}{(1+f)^2}$$

- Fact 4:** if $f \sim F(v_1, v_2)$, then $1/f \sim F(v_2, v_1)$

$$F_{1-\alpha, v_1, v_2} = \frac{1}{F_{\alpha, v_1, v_2}}$$

- Fact 5:** if $t \sim \text{students}'-t(v)$, then $t^2 \sim F(1, v)$

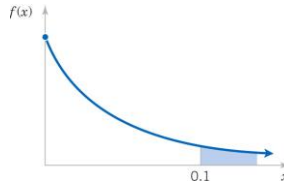
$$F_{\alpha, 1, v} = t_{\alpha/2, v}^2$$

Example 4-21: Computer Usage-1

In a large corporate computer network, time to user log-ons to the system can be modeled as exponential distribution with a mean of 0.04 hours. What is the probability that there are no log-ons in the next 6 minutes (0.1 hours)?

Let X denote the time in hours from the start of the interval until the first log-on.

$$\begin{aligned} P(X > 0.1) &= \int_{0.1}^{\infty} 25e^{-25x} dx = e^{-25(0.1)} \\ &= 1 - F(0.1) = 0.082 \end{aligned}$$



what is the interval of time such that the probability that no log-on occurs during the interval is 0.90?

$$\begin{aligned} P(X > x) &= e^{-25x} = 0.90, \quad -25x = \ln(0.90) \\ x &= \frac{-0.10536}{-25} = 0.00421 \text{ hour} = 0.253 \text{ minute} \end{aligned}$$