

## ISyE 6739 – Group Activity 9

### solutions

1. (a) `n = length(data1$Score_Diff)`  
`t.a <- t.test(data1$Score_Diff, alternative = "two.sided", mu = 0, conf.level = 0.95)`  
`t.a$conf.int`  
 $\Rightarrow$  0.95% confidence interval for the mean of score difference is  $(-5.611, 7.968)$ . Therefore, we can conclude that on average the score difference is 0 (because  $-5.611 < 0 < 7.968$ ).

- (b) The sample size  $n_{new}$  for which the estimation error is 0.1 is

$$n_{new} = \left( \frac{Z_{\alpha/2}s}{E} \right)^2 = \left( \frac{1.96 \cdots 17.5}{0.1} \right)^2 \approx 117778.$$

here  $s$  is a standard deviation.

- (c) `t.b <- t.test(data1$Attend, alternative = "greater", mu = 0, sigma.x = NULL, sigma.y = NULL, conf.level = 0.95)`  
`t.b$conf.int`  
 $\Rightarrow$  0.95% confidence interval for the attendance is  $(6322.4, +\infty)$ . It is more appropriate to use a lower bound because we are interested in increasing the attendance. Also, attendance is already bounded from above by the number of seats.

2. `data2 <- c(.9765, .9961, 1.0, .9922, .9961, 1.0, .9922, .9843, .9804, 1.0)`  
`df <- length(data2) - 1`  
`L = var(data2) * df / qchisq(0.05/2, df, lower.tail = FALSE)`  
`U = var(data2) * df / qchisq(1 - 0.05/2, df, lower.tail = FALSE)`  
`# CI for the variance`  
`c(lower = L, upper = U)`  
 $\Rightarrow$  0.95% confidence interval for the variance of the pixel values is  $(3.47 \cdot 10^{-5}, 2.442 \cdot 10^{-4})$ .  
`# CI for the std deviation`  
`c(lower = L^.5, upper = U^.5)`

$\Rightarrow$  0.95% confidence interval for the standard deviation of the pixel values is  $(0.00589, 0.01562)$ .

3. Fun size packets of M&M's in general contains 10 candies. There are 1 red and 2 yellow candies in the first packet, 1 red and 1 yellow candies in the second one.

- (a) Point estimates for the proportions of red and yellow candies:

$$\hat{p}_{red} = \frac{1+1}{10+10} = 0.1, \quad \hat{p}_{yellow} = \frac{2+1}{10+10} = 0.15.$$

- (b) The confidence interval for population proportion is

$$\hat{p} - Z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}} \leq p \leq \hat{p} + Z_{\alpha/2} \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

Then 95% confidence interval for the proportions of red and yellow candies are respectively

$$\begin{aligned} -0.0315 &\leq p_{red} \leq 0.2315, \\ -0.00649 &\leq p_{yellow} \leq 0.30649. \end{aligned}$$

Based on two fun size packets, we can conclude that on average proportion of yellow candies is greater than proportion of red ones.

(c)

$$n_{new} = \left( \frac{Z_{\alpha/2}}{E} \right)^2 \hat{p}_{red}(1 - \hat{p}_{red}) \approx 3458$$

$$n_{new} = \left( \frac{Z_{\alpha/2}}{E} \right)^2 \cdot 0.25 \approx 9604.$$

4.  $X \sim Exp(\lambda)$ ,  $E[X] = \frac{1}{\lambda}$ ,  $Var[X] = \frac{1}{\lambda^2}$ . We know that the MLE for parameter  $\lambda$  is  $\hat{\lambda} = \frac{1}{\bar{X}}$ . Let  $\theta = \frac{1}{\lambda}$ . Then by the invariance property of MLE

$$\hat{\theta} = \frac{1}{\hat{\lambda}} = \bar{X},$$

$$Var[\hat{\theta}] = Var[\bar{X}] = \frac{1}{n\lambda^2}$$

We can approximate  $Var[\hat{\theta}]$  by setting  $\lambda \approx \hat{\lambda}$ :

$$Var[\hat{\theta}] = \frac{(\bar{X})^2}{n}.$$

Then the confidence interval for parameter  $\theta$  is

$$\hat{\theta} - Z_{\alpha/2} \sqrt{Var[\hat{\theta}]} \leq \theta \leq \hat{\theta} + Z_{\alpha/2} \sqrt{Var[\hat{\theta}]}$$

$$\bar{X} - Z_{\alpha/2} \frac{\bar{X}}{\sqrt{n}} \leq \frac{1}{\lambda} \leq \bar{X} + Z_{\alpha/2} \frac{\bar{X}}{\sqrt{n}}$$

where  $s$  is a sample standard deviation,  $n$  is a sample size.  $\Rightarrow$  If  $\bar{X} - Z_{\alpha/2} \frac{\bar{X}}{\sqrt{n}} > 0$  then the confidence interval for parameter  $\lambda$  is

$$\frac{1}{\bar{X} + Z_{\alpha/2} \frac{\bar{X}}{\sqrt{n}}} \leq \frac{1}{\lambda} \leq \frac{1}{\bar{X} - Z_{\alpha/2} \frac{\bar{X}}{\sqrt{n}}}$$

otherwise

$$\frac{1}{\lambda} \geq \frac{1}{\bar{X} + Z_{\alpha/2} \frac{\bar{X}}{\sqrt{n}}}$$

Now we have that  $n = 100$ ,  $\bar{X} = 10$ . Then 95% confidence interval for parameter  $\lambda$  is

$$0.0836 \leq \lambda \leq 0.1244.$$

5.  $X_i \sim Beta(\alpha, 1)$

$$\Rightarrow E[X_i] = \frac{\alpha}{\alpha + 1}, \quad Var(X_i) = \frac{\alpha}{(\alpha + 1)^2(\alpha + 2)}.$$

By MOM:

$$\bar{X} = \frac{\alpha}{\alpha + 1}$$

$$\Rightarrow \hat{\alpha} = \frac{\bar{X}}{1 - \bar{X}}, \tag{1}$$

$$\widehat{Var(X_i)} = \frac{\hat{\alpha}}{(\hat{\alpha} + 1)^2(\hat{\alpha} + 2)} = \frac{\bar{X}(1 - \bar{X})^2}{2 - \bar{X}}.$$

Then CI for the mean is

$$\hat{X} - Z_{\alpha/2} \sqrt{\frac{\widehat{Var(X)}}{n}} \leq \frac{\hat{\alpha}}{1 + \hat{\alpha}} \leq \hat{X} + Z_{\alpha/2} \sqrt{\frac{\widehat{Var(X)}}{n}}$$

$$\Rightarrow 1 - \hat{X} + Z_{\alpha/2} \sqrt{\frac{\widehat{Var(X)}}{n}} \geq \frac{1}{1 + \hat{\alpha}} \geq 1 - \hat{X} - Z_{\alpha/2} \sqrt{\frac{\widehat{Var(X)}}{n}}, \tag{2}$$

If  $1 - \hat{X} - Z_{\alpha/2} \sqrt{\frac{\widehat{\text{Var}}(X)}{n}} > 0$  then CI for parameter  $\alpha$  is

$$-1 + \frac{1}{1 - \hat{X} + Z_{\alpha/2} \sqrt{\frac{\widehat{\text{Var}}(X)}{n}}} \leq \hat{\alpha} \leq -1 + \frac{1}{1 - \hat{X} - Z_{\alpha/2} \sqrt{\frac{\widehat{\text{Var}}(X)}{n}}}$$

otherwise

$$\hat{\alpha} \geq -1 + \frac{1}{1 - \hat{X} + Z_{\alpha/2} \sqrt{\frac{\widehat{\text{Var}}(X)}{n}}}$$