

CS215 Assignment 3

Question 2

Atishay Jain - 210050026

Gohil Megh Hiteshkumar - 210050055

November 1, 2022

1 Introduction

For each $N \in \{5, 10, 20, 40, 60, 80, 100, 500, 10^3, 10^4\}$, we repeated the experiment for $M \geq 100$ times. The experiment includes generating a data sample x_1, x_2, \dots, x_N of N points from a Uniform Distribution on $[0, 1]$ and Transform this data x to generate a transformed data sample y , where for each datum the transformed data is defined as -

$$y = -\frac{1}{\lambda} \log(x)$$

where λ is a parameter, and $\lambda = 5$ is used for above data generation.

Now, the transformed data y has some distribution with parameter λ . We find its analytical form and given this distribution, we have to estimate λ using different estimation methods. Finally, we calculated Relative errors between true value of λ and its estimates $\hat{\lambda}$. Relative Error calculated as -

$$Error = \frac{|\hat{\lambda} - \lambda_{true}|}{\lambda_{true}}$$

2 Analytical Form of Transformed Data

We have transformation function as -

$$y = g(x) = -\frac{1}{\lambda} \log(x)$$

Since, $g(x)$ is a monotonically decreasing function for $x > 0$, the PDF of y can be derived by applying Transformation of Random Variable formula as follows -

$$x = g^{-1}(y) = e^{-\lambda y}$$
$$P_Y(y) = P_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right| \quad (1)$$

As x is a Uniform Distribution in $[0, 1]$,

$$P_X(x) = \begin{cases} 1 & , 0 \leq x \leq 1 \\ 0 & , otherwise \end{cases}$$

Putting $P_X(x)$ in eq.(1) and solving, we get -

$$P_Y(y) = \begin{cases} \lambda e^{-\lambda y} & , y \geq 0 \\ 0 & , otherwise \end{cases}$$

3 Maximum Likelihood Estimate of λ

For ML Estimate of λ , given data y_1, y_2, \dots, y_N from distribution $P_Y(y)$, we need to find $\hat{\lambda}^{ML}$ which maximizes the Likelihood function. Solving it for $y \geq 0$ (as $P_Y(y)$ is 0 for $y < 0$) -

$$\begin{aligned} L(\lambda; y_1, \dots, y_N) &= P_Y(y_1, y_2, \dots, y_N | \lambda) \\ &= \prod_{i=1}^N P_Y(y_i | \lambda) \\ &= \lambda^N e^{-\lambda \sum_{i=1}^N y_i} \end{aligned}$$

For simplifying calculation, we are taking \log of Likelihood function and maximizing it wrt. λ . As \log is a monotonically increasing function, maximizing \log of Likelihood is equivalent to maximizing Likelihood function wrt. λ

$$\log(L(\mu; x_i, \dots, x_N)) = N \log(\lambda) - \lambda \sum_{i=1}^N y_i$$

Differentiating above expression wrt. λ and equating it is to zero -

$$\begin{aligned} \frac{d}{d\lambda} \left(N \log(\lambda) - \lambda \sum_{i=1}^N y_i \right) &= 0 \\ \frac{N}{\lambda} - \sum_{i=1}^N y_i &= 0 \\ \therefore \hat{\lambda}^{ML} &= \frac{N}{\sum_{i=1}^N y_i} \end{aligned}$$

4 Maximum Posterior Mean Estimate of λ

- Joint Likelihood function is same as $P_Y(y_1, y_2, \dots, y_N | \lambda)$ as calculated for MLE above -

$$P_Y(y_1, y_2, \dots, y_N | \lambda) = \begin{cases} \lambda^N e^{-\lambda \sum_{i=1}^N y_i} & , y_i \geq 0 \forall i \\ 0 & , otherwise \end{cases}$$

- Using a Gamma Prior for λ with parameters α and β -

$$P(\lambda | \alpha, \beta) = \frac{\beta^\alpha \lambda^{\alpha-1} e^{-\beta\lambda}}{\Gamma(\alpha)}$$

- So, Posterior Function is -

$$\begin{aligned} P(\lambda | y_1, y_2, \dots, y_N) &= \frac{P_Y((y_1, y_2, \dots, y_N | \lambda) P(\lambda)}{P(y_1, y_2, \dots, y_N)} \\ &= \frac{P_Y((y_1, y_2, \dots, y_N | \lambda) P(\lambda)}{\int_{\lambda} P_Y((y_1, y_2, \dots, y_N | \lambda) P(\lambda) \cdot d\lambda} \\ &= \frac{\lambda^N e^{-\lambda \sum_{i=1}^N y_i} \lambda^{\alpha-1} e^{-\beta\lambda}}{\int_{\lambda} \lambda^N e^{-\lambda \sum_{i=1}^N y_i} \lambda^{\alpha-1} e^{-\beta\lambda} \cdot d\lambda} \end{aligned}$$

Therefore, the Posterior Mean is -

$$\begin{aligned} E_{P(\lambda|y_1, y_2, \dots, y_N)}[\lambda] &= \int_{\lambda} \lambda P(\lambda|y_1, y_2, \dots, y_N) \cdot d\lambda \\ &= \frac{\int_0^{\infty} \lambda^{N+\alpha} e^{-\lambda(\beta + \sum_{i=1}^N y_i)} \cdot d\lambda}{\int_0^{\infty} \lambda^{N+\alpha-1} e^{-\lambda(\beta + \sum_{i=1}^N y_i)} \cdot d\lambda} \end{aligned}$$

Using the formula for Gamma function of $\int_0^{\infty} x^{\alpha-1} e^{-\beta x} \cdot dx = \frac{\Gamma(\alpha)}{\beta^{\alpha}}$,

$$E_{P(\lambda|y_1, y_2, \dots, y_N)}[\lambda] = \frac{\Gamma(N + \alpha + 1) / (\beta + \sum_{i=1}^N y_i)^{N+\alpha+1}}{\Gamma(N + \alpha) / (\beta + \sum_{i=1}^N y_i)^{N+\alpha}}$$

Using expression of Gamma function $\Gamma(x) = (x - 1)!$ and solving further, we get the formula for **Posterior Mean** as -

$$\hat{\lambda}^{PosteriorMean} = E_{P(\lambda|y_1, y_2, \dots, y_N)}[\lambda] = \frac{N + \alpha}{(\sum_{i=1}^N y_i) + \beta}$$

5 Boxplot of Relative Errors

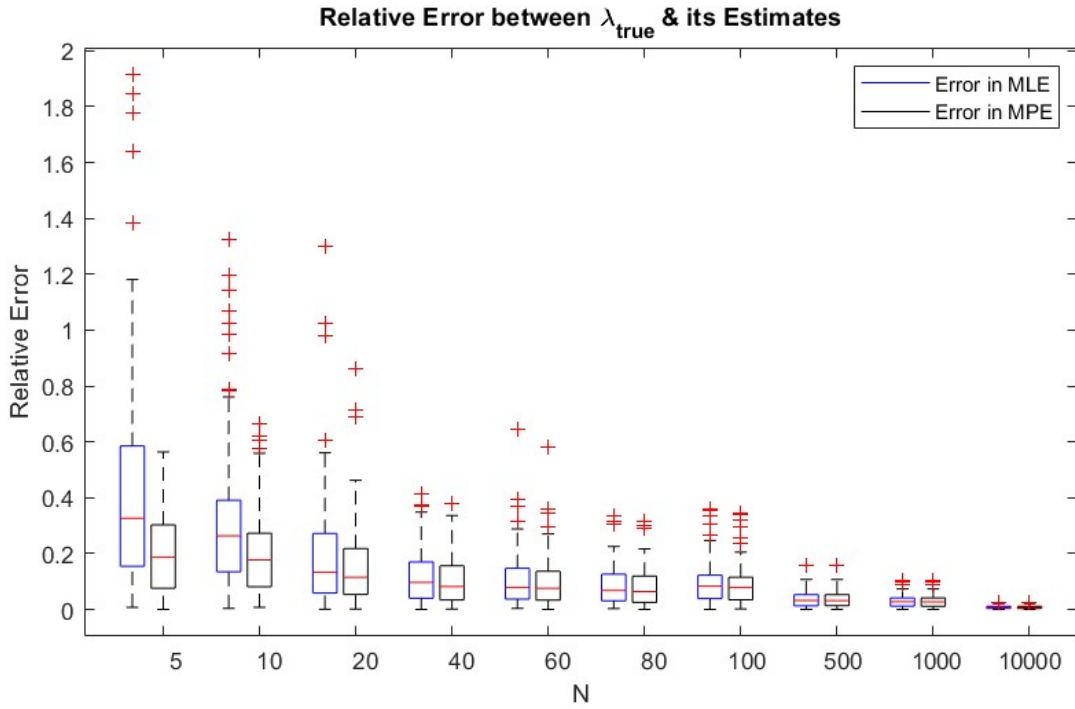


Figure 1: Boxplot of Relative errors in both the estimates for each N

The above Boxplot shows the Relative errors between true value and estimated values of the parameter λ . The Blue ones are the Boxplot for Maximum Likelihood Estimation and the Black ones are for Maximum Posterior Estimation of λ . As seen from the plot, the error in values estimated using Posterior Mean are relatively lesser.

6 Interpretation of Plot

6.1 What happens when N increases?

- We can see that the spread distribution of error decreases as the data size i.e. N increases.
- The median of distribution of error decreases as the data size increases.
- We can observe that the differences in the spread of errors for both the estimates are more at small values of N
- But as N increases, all the spread are almost nearly equal and are very small, i.e. both estimates tend to true mean at larger values of N .

6.2 Preferable Estimate

- As we can see from the box-plot graph at large values of N , the box-plot are nearly same for both estimates.
- So it means that both of the estimates give almost same estimates at large values of N . So any one of them is preferable at large values of N .
- But at lower values of N , we can see that the spread of error follows the trend -

Maximum Likelihood estimate $>$ Maximum Posterior Mean estimate.

- Hence considering all values of N , Preferable estimate is the **Maximum Posterior Estimate** for Posterior Mean with Prior taken as Gamma function.