

HOW MANY MODES CAN A CONSTRAINED GAUSSIAN MIXTURE HAVE?

NAVIN KASHYAP AND MANJUNATH KRISHNAPUR

ABSTRACT. We show, by an explicit construction, that a mixture of univariate Gaussians with variance 1 and means in $[-A, A]$ can have $\Omega(A^2)$ modes. This disproves a recent conjecture of Dytso, Yagli, Poor and Shamai [3] who showed that such a mixture can have at most $O(A^2)$ modes and surmised that the upper bound could be improved to $O(A)$. Our result holds even if an additional variance constraint is imposed on the mixing distribution. Extending the result to higher dimensions, we exhibit a mixture of Gaussians in \mathbb{R}^d , with identity covariances and means inside $[-A, A]^d$, that has $\Omega(A^{2d})$ modes.

$$Z = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2}$$

$$= \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}(x-a_i)^2}$$

1. INTRODUCTION

Let X be a random variable with distribution $\mu = p_1\delta_{a_1} + \dots + p_N\delta_{a_N}$ where $-A \leq a_1 < a_2 < \dots < a_N \leq A$ and $p_i > 0$ sum to 1. Throughout this note, Z denotes a standard Gaussian random variable that is independent of X . Then, $Y = X + Z$ has density $f_Y(t) = \sum_{k=1}^N p_k \varphi(t - a_k)$, where $\varphi(t) = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}t^2}$. We want to know the maximum number of modes (local maxima) that f_Y , a mixture of Gaussians with centres (means) a_k constrained to be in $[-A, A]$, can have. Let this quantity be denoted as $m(A)$. The main aim of this note is to give a proof of the following proposition.

Proposition 1. $m(A) = \Omega(A^2)$, i.e., $m(A) \geq c_0 A^2$ for some constant $c_0 > 0$ and all $A > 0$.

Remark 1. It was recently shown by Dytso, Yagli, Poor and Shamai [3, Theorem 6] that $m(A) \leq c_1 A^2$, for some constant $0 < c_1 < \infty$. This, along with our Proposition 1 above, shows that $m(A) = \Theta(A^2)$. In particular, this disproves the conjecture made by Dytso et al. [3, Remark 9] that $m(A) = \Theta(A)$.¹

The motivation for their conjecture was that, via [3, Eqs. (43) and (65)], $2m(A)$ is an upper bound for $N^*(A)$, which is the number of points in the support of the optimal input distribution for an additive white Gaussian noise (AWGN) channel with amplitude constraint A . Thus, one consequence of their conjecture would have been that $N^*(A) = O(A)$. In fact, since they show that $N^*(A) = \Omega(A)$, their conjecture would have implied that $N^*(A) = \Theta(A)$. While our proposition

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¹Independently of us, Polyanskiy and Wu [4] have also obtained a result that effectively disproves this conjecture. They give an example of a random variable X having a density π supported within $[-A, A]$ such that the density, $\pi * \varphi$, of $X + Z$ has $\Omega(A^2)$ modes.

shows that the route via their conjecture is blocked, numerical work does indeed suggest that $N^*(A) = \Theta(A)$. \square

The result of Proposition 1 does not change qualitatively if we further impose a variance constraint on the X in $Y = X + Z$. To be precise, consider now Gaussian mixtures $f_Y(t) = \sum_{k=1}^N p_k \varphi(t - a_k)$, with centres a_k again constrained to be in $[-A, A]$, but additionally requiring the random variable $X \sim \sum_{k=1}^N p_k \delta_{a_k}$ to have variance $\text{var}(X) \leq 1$. (Of course, any constant bound on the variance will do; we take the bound to be 1 for simplicity.) Let $m_{\#}(A)$ denote the maximum number of modes among such mixtures f_Y . We then have the following result.

Proposition 2. $m_{\#}(A) = \Omega(A^2)$, i.e., $m_{\#}(A) \geq c_{\#} A^2$ for some constant $c_{\#} > 0$ and all $A > 0$.
↪ variance $\lesssim 1$

Our results extend to higher dimensions without substantial change. Let φ_d denote the standard Gaussian density (zero mean and identity covariance) in \mathbb{R}^d . Let $m_d(A)$ denote the maximum number of modes that the Gaussian mixture density $f(t) = p_1 \varphi_d(t - a_1) + \dots + p_N \varphi_d(t - a_N)$ can have, subject to the constraints that $|a_i| \leq A$ for all i , and $p_i > 0$ sum to 1.

Proposition 3. With the above notation, $m_d(A) \geq c A^{2d}$ for a constant $c > 0$ that is independent of A .

However, we are not aware of a corresponding upper bound. It is worth remarking here that there is considerable interest in counting modes of Gaussian mixtures. For instance, it was conjectured by Sturmfels (see [1, Conjecture 5]) that a Gaussian mixture (with identity covariances, as we have taken) with N components, has at most $\binom{N+d-1}{d}$ modes. In one dimension, this bound reduces to N , which is in fact proved in [5] — see also [2, Section 2.4]. These studies are without any constraint on the centers while the amplitude constraint is a key feature in this paper.

Sketch of the proofs. The main ingredients in our proofs of Propositions 1 and 2 are mixtures of the form

$$\gamma_{a,N}(x) := \frac{1}{2N+1} \sum_{n=-N}^N \varphi(x - an),$$

$$f_Y(t) = \sum_{k=1}^N p_k \varphi(t - a_k),$$

$\frac{1}{2N+1}$ $\frac{c \cdot N}{\sqrt{N}}$

with $a > 0$. This is an equally-weighted mixture of $2N + 1$ Gaussians with centres (means) an , for integers n between $-N$ and N . Fig. 1 illustrates the shape of the unnormalized mixture

$$f_{a,N}(x) := \sum_{n=-N}^N \varphi(x - an).$$

We will show that by choosing $a = \frac{c}{\sqrt{N}}$ for a suitable constant $c > 0$, the resulting unnormalized mixture $f_{a,N}$ has centres in $[-c\sqrt{N}, c\sqrt{N}]$ and at least $N - 1$ modes. Since scaling by a constant has no effect on the number of modes, the same holds for the mixture $\gamma_{a,N}$, which suffices to prove Proposition 1. The proof is elaborated in Section 2.

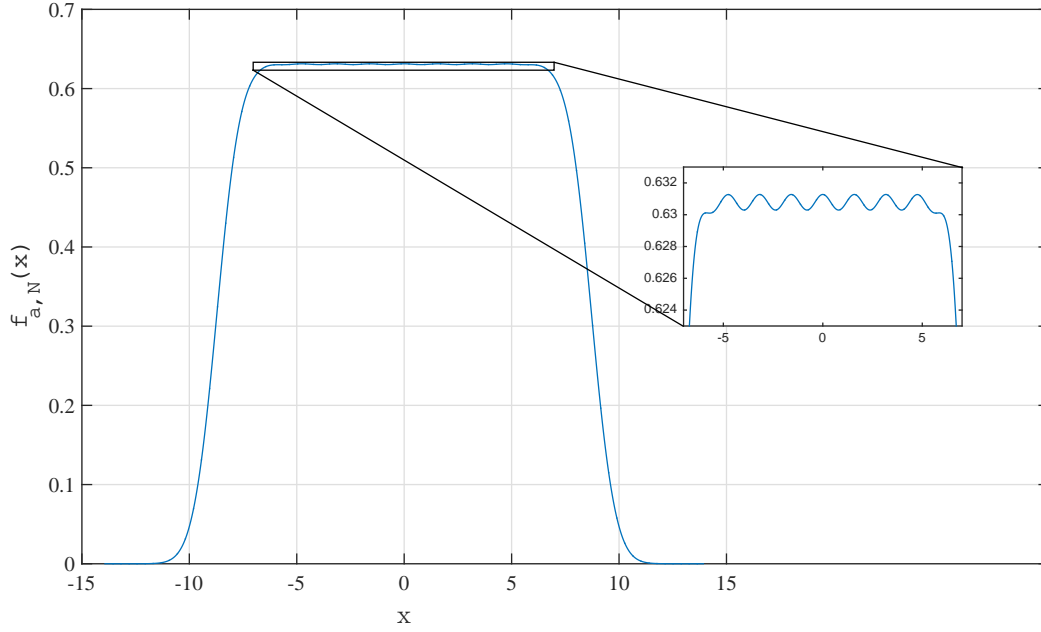


FIGURE 1. A plot of $f_{a,N}(x) = \sum_{n=-N}^N \varphi(x - an)$ for $N = 5$ and $a = 2\sqrt{\pi/N}$.

For Proposition 2, we work with the mixture

$$\begin{aligned} \Gamma_{\alpha; a, N}(x) &:= (1 - 2\alpha) \varphi(x) + \alpha \gamma_{a, N}(x + 2aN) + \alpha \gamma_{a, N}(x - 2aN) \\ &= (1 - 2\alpha) \varphi(x) + \frac{\alpha}{2N+1} \sum_{n=-3N}^{-N} \varphi(x - an) + \frac{\alpha}{2N+1} \sum_{n=N}^{3N} \varphi(x - an), \end{aligned} \quad (1)$$

where $a = \frac{c}{\sqrt{N}}$ is as above, and $\alpha \in (0, \frac{1}{2})$. This is a Gaussian mixture with centres at 0 and $\pm an$, $n = N, N+1, \dots, 3N$, weighted by $1 - 2\alpha$ and $\frac{\alpha}{2N+1}$, respectively. It is easy to check that by taking $\alpha \sim \frac{1}{N}$, we can get the underlying random variable X to have variance at most 1. We will, moreover, show that for this choice of α , the mixture $\Gamma_{\alpha; a, N}$ has $\Omega(N)$ modes. Since $\Gamma_{\alpha; a, N}$ has all its centres within $[-3c\sqrt{N}, 3c\sqrt{N}]$, this will prove Proposition 2. The detailed proof is in Section 2.

The proof of Proposition 3 is entirely analogous to that of Proposition 1, and uses a mixture with equal weights and centers at ak , where $k = (k_1, \dots, k_d) \in \mathbb{Z}^d$ with $-N \leq k_i \leq N$, for appropriately chosen a and N (the right choices turn out to be $a = 1/A$ and $N = A^2$). Details are in Section 3.

2. PROOF OF PROPOSITION 1 AND PROPOSITION 2

Our analysis is based on the fact that, for any $a > 0$, the unnormalized mixture $f_{a,N}$ is a truncation of the infinite series

$$f_a(x) := \sum_{n \in \mathbb{Z}} \varphi(x - an).$$

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$= \frac{1}{\sqrt{2\pi}} e^{-\frac{(x-an)^2}{2}}$

$\dots e^{-\frac{(x+2)^2}{2}} + e^{-\frac{x^2}{2}} + e^{-\frac{(x-2)^2}{2}} + \dots$

Note that f_a is well-defined and periodic with period a . By standard real-analysis arguments, f_a is continuous on \mathbb{R} .

We first obtain an estimate for $h_a := f_a(0) - f_a(\frac{a}{2})$, which we will use in our proofs.

Lemma 4. *For any $a > 0$, we have*

$$\frac{4}{a} e^{-\frac{2\pi^2}{a^2}} \leq h_a \leq \frac{4}{a} e^{-\frac{2\pi^2}{a^2}} \left(1 - e^{-\frac{2\pi^2}{a^2}}\right)^{-1}.$$

Proof. We prove the lower bound first. By the Poisson summation formula², for any $x \in \mathbb{R}$,

$$f_a(x) = \sum_{n \in \mathbb{Z}} \varphi\left(a\left(\frac{x}{a} - n\right)\right) = \frac{1}{a} \sum_{n \in \mathbb{Z}} e^{-\frac{2\pi^2 n^2}{a^2}} e^{2\pi i n \frac{x}{a}}, \quad (2)$$

from which we get

$$f_a(0) = \frac{1}{a} \sum_{n \in \mathbb{Z}} e^{-\frac{2\pi^2 n^2}{a^2}} > \frac{1}{a} > \frac{1}{a} \sum_{n \in \mathbb{Z}} (-1)^n e^{-\frac{2\pi^2 n^2}{a^2}} = f_a\left(\frac{a}{2}\right).$$

In particular, we have

$$\begin{aligned} h_a &= \frac{1}{a} \sum_{n \in \mathbb{Z}} e^{-\frac{2\pi^2 n^2}{a^2}} - \frac{1}{a} \sum_{n \in \mathbb{Z}} (-1)^n e^{-\frac{2\pi^2 n^2}{a^2}} \\ &= \frac{2}{a} \sum_{\substack{n \in \mathbb{Z}, \\ n \text{ odd}}} e^{-\frac{2\pi^2 n^2}{a^2}} \\ &= \frac{4}{a} \sum_{\substack{n > 0, \\ n \text{ odd}}} e^{-\frac{2\pi^2 n^2}{a^2}} \\ &> \frac{4}{a} e^{-\frac{2\pi^2}{a^2}}. \end{aligned}$$

For the upper bound, consider

$$|f_a(x) - \frac{1}{a}| \leq \frac{1}{a} \sum_{n \neq 0} e^{-\frac{2\pi^2 n^2}{a^2}} \leq \frac{2e^{-\frac{2\pi^2}{a^2}}}{a \left(1 - e^{-\frac{2\pi^2}{a^2}}\right)},$$

the first inequality arising from (2), and the second inequality being obtained by replacing n^2 by n to get a geometric series. Thus,

$$h_a = |f_a(0) - \frac{1}{a}| + |f_a(\frac{a}{2}) - \frac{1}{a}| \leq \frac{4e^{-\frac{2\pi^2}{a^2}}}{a \left(1 - e^{-\frac{2\pi^2}{a^2}}\right)}, \quad (3)$$

$a > c > b$
 $a-b = |a-c| + |b-c|$

which is the claimed upper bound. □

²With the notation $\hat{f}(\lambda) = \int f(x) e^{-2\pi i \lambda x} dx$, we have $\sum_{n \in \mathbb{Z}} f(x+n) = \sum_{n \in \mathbb{Z}} \hat{f}(n) e^{2\pi i n x}$.

probably because the imaginary terms cancel out and the reals are less than 1.

Thus, for $a \ll 1$, we have $h_a \approx \frac{4}{a} \exp(-\frac{2\pi^2}{a^2})$. We actually need only the lower bound on h_a for our arguments.

Remark 2. A minor modification in the above proof shows that the bounds in Lemma 4 in fact apply to $\bar{h}_a = \max(f_a) - \min(f_a)$ as well. Indeed, the lower bound is obvious, since $\bar{h}_a \geq h_a$. For the upper bound, we observe that if x^* and x_* achieve the maximum and minimum, respectively, of f_a , then $\bar{h}_a = |f_a(x^*) - \frac{1}{a}| + |f_a(x_*) - \frac{1}{a}|$, so that the upper bound in (3) still holds.

It is clear from (2) that $f_a(0) > f_a(x)$ for all $x \in [-\frac{a}{2}, \frac{a}{2}]$, since there is non-trivial cancellation in the terms of the series unless x is an integer multiple of a . By the fact that f_a has period a , we see that na is a strict maximum of f_a in the interval $I_{a,n} := [na - \frac{a}{2}, na + \frac{a}{2}]$ for any $n \in \mathbb{Z}$. We wish argue that $f_{a,N}$ also has local maxima within those intervals $I_{a,n}$ that are contained in $[-\frac{1}{2}aN, \frac{1}{2}aN]$. For this, we will need the simple lemma stated next.

Lemma 5. Let g be a continuous function such that $|f_a - g| < \frac{1}{2}h_a$ on a subset $S \subseteq \mathbb{R}$. Then, g has a local maximum in the interior of any interval $I_{a,n}$ that is contained within S .

Proof. Recall that $I_{a,n} = [na - \frac{a}{2}, na + \frac{a}{2}]$, for $n \in \mathbb{Z}$. If $|f_a - g| < \frac{1}{2}h_a$ holds on $I_{a,n}$, then we have

$$\begin{aligned} g(na) - g(na - \frac{a}{2}) &= (g(na) - f_a(na)) + (f_a(na) - f_a(na - \frac{a}{2})) + (f_a(na - \frac{a}{2}) - g(na - \frac{a}{2})) \\ &> (-\frac{1}{2}h_a) + h_a + (-\frac{1}{2}h_a) \\ &= 0. \end{aligned}$$

min value is $\frac{h_a}{2}$

Hence, $g(na) > g(na - \frac{a}{2})$. Analogously, $g(na) > g(na + \frac{a}{2})$. Therefore, the global maximum of g in $I_{n,a}$ is attained at an interior point. In particular, g has a local maximum strictly between $na - \frac{a}{2}$ and $na + \frac{a}{2}$. \square

We now have the facts necessary to furnish proofs of Propositions 1 and 2.

Proof of Proposition 1. We apply Lemma 5 with $g = f_{a,N}$. Note first that

$$\begin{aligned} |f_a(x) - f_{a,N}(x)| &= \frac{1}{\sqrt{2\pi}} \sum_{n:|n|>N} e^{-\frac{1}{2}(an-x)^2} \\ &\leq \frac{1}{\sqrt{2\pi}} \sum_{n:|n|>N} e^{-\frac{1}{2}(a|n|-|x|)^2} \quad (\text{since } |an-x| \geq |a|n| - |x|) \\ &= \frac{2}{\sqrt{2\pi}} \sum_{n>N} e^{-\frac{1}{2}(an-|x|)^2} \\ &= \frac{2}{\sqrt{2\pi}} e^{-\frac{1}{2}(aN-|x|)^2} \sum_{n>N} e^{-\frac{1}{2}a(n-N)(a(N+n)-2|x|)} \\ &\leq \frac{2}{\sqrt{2\pi}} e^{-\frac{1}{2}(aN-|x|)^2} \sum_{n>N} e^{-a(n-N)(aN-|x|)} \end{aligned}$$

Now take $|x| \leq \frac{1}{2}aN$ to get

$$\begin{aligned} |f_a(x) - f_{a,N}(x)| &\leq \frac{2}{\sqrt{2\pi}} e^{-\frac{1}{8}a^2N^2} \sum_{n>N} e^{-\frac{1}{2}a^2N(n-N)} \\ &= \frac{2}{\sqrt{2\pi}} e^{-\frac{1}{8}a^2N^2} \frac{e^{-\frac{1}{2}a^2N}}{1 - e^{-\frac{1}{2}a^2N}}. \end{aligned} \quad (4)$$

If we take $a = \frac{2\sqrt{\pi}}{\sqrt{N}}$ and $S = [-\frac{1}{2}aN, \frac{1}{2}aN] = [-\sqrt{\pi N}, \sqrt{\pi N}]$, then (4) holds for all $x \in S$, so that

$$|f_a(x) - f_{a,N}(x)| \leq C_0 e^{-\frac{1}{2}\pi N} \quad (5)$$

with $C_0 = \frac{2}{\sqrt{2\pi}} \left(\frac{e^{-2\pi}}{1 - e^{-2\pi}} \right)$. On the other hand, from the lower bound for h_a in Lemma 4, we have

$$h_a \geq 2 \sqrt{\frac{N}{\pi}} e^{-\frac{1}{2}\pi N}.$$

As $C_0 < \frac{2}{\sqrt{\pi}} \left(\frac{e^{-2\pi}}{1 - e^{-2\pi}} \right)$, we have for all $N \geq 1$, $C_0 e^{-\frac{1}{2}\pi N} < \left(\frac{e^{-2\pi}}{1 - e^{-2\pi}} \right) h_a$, and consequently,

$$|f_a(x) - f_{a,N}(x)| < \left(\frac{e^{-2\pi}}{1 - e^{-2\pi}} \right) h_a \quad \text{for all } x \in S.$$

Since $\frac{e^{-2\pi}}{1 - e^{-2\pi}} \approx 0.0019$, the conclusion of Lemma 5 holds, i.e., $f_{a,N}$ has a local maximum in the interior of each of the intervals $I_{a,n}$ contained in $S = [-\frac{1}{2}aN, \frac{1}{2}aN]$. There are at least $N - 1$ such intervals $I_{a,n}$, and hence, $f_{a,N}$ has at least $N - 1$ local maxima within S . Thus, we conclude that the Gaussian mixture $\gamma_{a,N} = \frac{1}{2N+1} f_{a,N}$ (with $a = \frac{2\sqrt{\pi}}{\sqrt{N}}$), which has all its centres inside $[-2\sqrt{\pi N}, 2\sqrt{\pi N}]$, has at least $N - 1$ modes (within $S = [-\sqrt{\pi N}, \sqrt{\pi N}]$). Choosing $N = A^2$ proves Proposition 1. \square

Proof of Proposition 2. Consider $\Gamma_{\alpha;a,N}$ as defined in (1), with $a = \frac{2\sqrt{\pi}}{\sqrt{N}}$ as in the proof of Proposition 1. This is the density of $Y = X + Z$, where $Z \sim \mathcal{N}(0, 1)$ is independent of $X \sim (1 - 2\alpha)\delta_0 + \frac{\alpha}{2N+1} \sum_{n=N}^{3N} (\delta_{-an} + \delta_{an})$. We then have

$$\begin{aligned} \text{var}(X) &= \frac{\alpha}{2N+1} \sum_{n=N}^{3N} 2(an)^2 \\ &\leq \frac{2\alpha a^2}{2N+1} \sum_{n=1}^{3N} n^2 \\ &= \frac{2\alpha a^2}{2N+1} \left(\frac{3N(3N+1)(6N+1)}{6} \right) \\ &\leq \alpha a^2 (3N)(3N+1) \\ &= 12\pi(3N+1)\alpha \quad (\text{using } a = \frac{2\sqrt{\pi}}{\sqrt{N}}) \end{aligned}$$

Hence, setting $\alpha = \frac{1}{12\pi(3N+1)}$, we obtain $\text{var}(X) \leq 1$.

We will next show that, with a and α as above, $\Gamma_{\alpha;a,N}$ has $\Omega(N)$ modes. This suffices to prove the proposition, since $\Gamma_{\alpha;a,N}$ is a Gaussian mixture with all of its centres in $[-6\sqrt{\pi N}, 6\sqrt{\pi N}]$.

It is easy to check that $\Gamma_{\alpha;a,N}$ has a mode at 0. We will show that, when N is sufficiently large, $\Gamma_{\alpha;a,N}$ has at least $N - 1$ modes in each of the intervals $[-5\sqrt{\pi N}, -3\sqrt{\pi N}]$ and $[3\sqrt{\pi N}, 5\sqrt{\pi N}]$. By symmetry, it is enough to show this for the interval $[3\sqrt{\pi N}, 5\sqrt{\pi N}]$. For this, we use Lemma 5 with $g = \left(\frac{2N+1}{\alpha}\right)\Gamma_{\alpha;a,N}$. For this choice of g , we have

$$\begin{aligned} |f_a(x) - g(x)| &= \left| \sum_{n: |n| < N \text{ or } |n| > 3N} \varphi(x - an) - \left(\frac{1-2\alpha}{\alpha}\right)(2N+1)\varphi(x) \right| \\ &\leq \sum_{n: |n| < N \text{ or } |n| > 3N} \varphi(x - an) + \left(\frac{1-2\alpha}{\alpha}\right)(2N+1)\varphi(x) \\ &\leq \sum_{n < N \text{ or } n > 3N} \varphi(x - an) + \left(\frac{2N+1}{\alpha}\right)\varphi(x) \end{aligned} \quad (6)$$

Consider the first term in (6) above. Writing $x' = x - 2aN$, we have

$$\begin{aligned} \sum_{n < N \text{ or } n > 3N} \varphi(x - an) &= \sum_{n < N \text{ or } n > 3N} \varphi(x' - a(n - 2N)) \\ &= \sum_{n < -N \text{ or } n > N} \varphi(x' - an) \\ &= |f_a(x') - f_{a,N}(x')| \\ &\leq C_0 e^{-\frac{1}{2}\pi N} \end{aligned}$$

for $|x'| \leq \frac{1}{2}aN$ and $C_0 = \frac{2}{\sqrt{2\pi}} \left(\frac{e^{-2\pi}}{1-e^{-2\pi}}\right)$, by (5) in the proof of Proposition 1. Thus, for $|x - 2aN| \leq \frac{1}{2}aN$, i.e., for $x \in [3\sqrt{\pi N}, 5\sqrt{\pi N}]$, we see that the first term in (6) is bounded above by $C_0 e^{-\frac{1}{2}\pi N}$.

Turning our attention to the second term in (6), we first observe that $\frac{2N+1}{\alpha} \leq C'_0 N^2$ for some constant C'_0 . Thus,

$$\left(\frac{2N+1}{\alpha}\right)\varphi(x) \leq C'_0 N^2 \varphi(x) \leq \frac{1}{\sqrt{2\pi}} C'_0 N^2 e^{-\frac{9}{2}\pi N},$$

for $x \geq 3\sqrt{\pi N}$.

Combining these bounds, we obtain that for $x \in [3\sqrt{\pi N}, 5\sqrt{\pi N}]$,

$$|f_a(x) - g(x)| \leq C_0 e^{-\frac{1}{2}\pi N} + \frac{1}{\sqrt{2\pi}} C'_0 N^2 e^{-\frac{9}{2}\pi N} \leq 2C_0 e^{-\frac{1}{2}\pi N}$$

when N is sufficiently large. As shown in the proof of Proposition 1, $C_0 e^{-\frac{1}{2}\pi N} < \left(\frac{e^{-2\pi}}{1-e^{-2\pi}}\right) h_a$. Consequently, when N is sufficiently large, for $x \in [3\sqrt{\pi N}, 5\sqrt{\pi N}]$, we have

$$|f_a(x) - g(x)| < \left(\frac{2e^{-2\pi}}{1-e^{-2\pi}}\right) h_a < 0.004 h_a.$$

Then, applying Lemma 4, we obtain that, for all sufficiently large N , the function $g = (\frac{2N+1}{\alpha})\Gamma_{\alpha;a,N}$ has at least $N-1$ modes within the interval $[3\sqrt{\pi N}, 5\sqrt{\pi N}]$. This naturally holds for $\Gamma_{\alpha;a,N}$ as well, thus proving the proposition. \square

3. PROOF OF PROPOSITION 3

Since the proof is entirely analogous to that of Proposition 1, we shall only sketch the modifications needed and omit the details. For $a > 0$ and integer $N \geq 1$ and define the functions

$$f_a(x) = \sum_{n \in \mathbb{Z}^d} \varphi_d(x - na),$$

$$f_{a,N}(x) = \sum_{n \in Q_N} \varphi_d(x - na),$$

where $Q_N = \{n \in \mathbb{Z}^d : -N \leq n_i \leq N \text{ for } 1 \leq i \leq d\}$. By the Poisson summation formula on \mathbb{R}^d with respect to the lattice \mathbb{Z}^d , we get

$$f_a(x) = \frac{1}{a^d} \sum_{p \in \mathbb{Z}^d} e^{-\frac{1}{2a^2}|p|^2 + \frac{2\pi i}{a}\langle p, x \rangle}$$

$$= \frac{1}{a^d} \left(1 + 2e^{-\frac{1}{2a^2}} \sum_{j=1}^d \cos(2\pi x_j/a) + O(e^{-\frac{2}{a^2}}) \right)$$

where the big-O term includes the contribution of all p with $|p| \geq 2$. Since $\cos(2\pi t) \leq 1 - 8t^2$ for any $t \in \mathbb{R}$, we see that when $|x| = \frac{a}{2}$,

$$f_a(x) \leq \frac{1}{a^d} \left(1 + 2e^{-\frac{1}{2a^2}} \sum_{j=1}^d \left(1 - \frac{8}{a^2} x_j^2 \right) + O(e^{-\frac{2}{a^2}}) \right)$$

$$= \frac{1}{a^d} \left(1 + 2(d-2)e^{-\frac{1}{2a^2}} + O(e^{-\frac{2}{a^2}}) \right).$$

Since $f_a(0) = \frac{1}{a^d} (1 + 2de^{-1/(2a^2)} + O(e^{-2/a^2}))$, we see that

$$f_a(0) - \sup_{|x|=\frac{a}{2}} f_a(x) = \frac{1}{a^d} (4e^{-\frac{1}{2a^2}} + O(e^{-\frac{2}{a^2}}))$$

which is at least $h_a := \frac{3}{a^d} e^{-\frac{1}{2a^2}}$, for small enough a . By periodicity, in each cube of the form $na + [-\frac{1}{2}a, \frac{1}{2}a]^d$, the graph of f_a has a hill with peak at na and having height at least h_a . Further,

$$|f_a(x) - f_{a,N}(x)| = \frac{1}{(2\pi)^{\frac{d}{2}}} \sum_{n \in \mathbb{Z}^d \setminus Q_N} e^{-\frac{1}{2a^2}|(x+na)|^2}$$

$$= O(e^{-\frac{1}{8}a^2N^2}) \quad \text{for } |x| \leq \frac{1}{2}aN.$$

Now take $a = \frac{c}{\sqrt{N}}$ to see that for suitable c, c' ,

$$\sup_{|x| \leq c'\sqrt{N}} |f_a(x) - f_{a,N}(x)| < \frac{1}{2}h_a.$$

Therefore, the function $f_{a,N}$ has a local maximum in each cube of the form $na + [-\frac{1}{2}a, \frac{1}{2}a]^d$ that is contained inside the larger cube $[-c'\sqrt{N}, c'\sqrt{N}]^d$. This is because the perturbation is too small to wash away the local maximum of f_a located at na . The number of such cubes is about $(2c'\sqrt{N}/a)^d$, which is $\Theta(N^d)$.

Taking $N = \sqrt{A}$ gives us a function $f_{a,N}$ (with $a = c/A$) that is a mixture of Gaussians with centers in Q_A and having $\Theta(A^{2d})$ modes. This was the claim of Proposition 3.

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NAVIN KASHYAP, DEPARTMENT OF ELECTRICAL COMMUNICATION ENGINEERING, INDIAN INSTITUTE OF SCIENCE, BANGALORE, INDIA.

E-mail address: `nkashyap@iisc.ac.in`

MANJUNATH KRISHNAPUR, DEPARTMENT OF MATHEMATICS, INDIAN INSTITUTE OF SCIENCE, BANGALORE, INDIA.

E-mail address: `manju@iisc.ac.in`