

Statistical Methods for Analysis with Missing Data

Lecture 11 setup: examples of Gibbs sampler, data augmentation for proper multiple imputation, MICE

Mauricio Sadinle

Department of Biostatistics

W UNIVERSITY *of* WASHINGTON

Previous Lectures

- ▶ Gibbs sampling
- ▶ Data augmentation to handle missing data in Bayesian inference
- ▶ Multiple imputation as a Monte Carlo approximation of proper Bayesian procedure
- ▶ Uncongeniality generally leads to invalidity of inferences based on Rubin's combining rules
- ▶ MICE: practical implementation of multiple imputation that builds on Gibbs sampling ideas

Today's Lecture

R implementations of

- ▶ Gibbs sampler
- ▶ Data augmentation for proper multiple imputation
- ▶ MICE

Outline

Example of Gibbs Sampler

Data Augmentation for Proper Multiple Imputation

MICE

Summary

Bhattacharyya's Distribution

Consider real-valued random variables X and Y having a joint distribution with density¹

$$p_{X,Y}(x,y) = \exp \left\{ [1, x, x^2] \begin{bmatrix} m_{00} & m_{01} & m_{02} \\ m_{10} & m_{11} & m_{12} \\ m_{20} & m_{21} & m_{22} \end{bmatrix} \begin{bmatrix} 1 \\ y \\ y^2 \end{bmatrix} \right\},$$

where either

(a) $m_{22} = m_{21} = m_{12} = 0$; $m_{20}, m_{02} < 0$; $m_{11}^2 < 4m_{20}m_{02}$;

(b) $m_{22} < 0$, $4m_{22}m_{02} > m_{12}^2$, $4m_{22}m_{20} > m_{21}^2$.

m_{00} is determined by the other m_{ij} 's so that $p_{X,Y}$ integrates to 1.

¹Distribution credited to Anil Kumar Bhattacharyya, who was a professor at the Indian Statistical Institute. See, e.g.,

Bhattacharyya's Distribution

From $p_{X,Y}(x,y)$ it is easy to see that

$$p_{X|Y}(x|y) \propto \frac{1}{\sigma_X(y)} \exp \left\{ -\frac{[x - \mu_X(y)]^2}{2\sigma_X^2(y)} \right\},$$

where

$$\mu_X(y) = -\frac{m_{10} + m_{11}y + m_{12}y^2}{2(m_{20} + m_{21}y + m_{22}y^2)},$$

and

$$\sigma_X^2(y) = -\frac{1}{2(m_{20} + m_{21}y + m_{22}y^2)}$$

Bhattacharyya's Distribution

And analogously, it is easy to see that

$$p_{Y|x}(y|x) \propto \frac{1}{\sigma_Y(x)} \exp \left\{ -\frac{[y - \mu_Y(x)]^2}{2\sigma_Y^2(x)} \right\},$$

where

$$\mu_Y(x) = -\frac{m_{01} + m_{11}x + m_{21}x^2}{2(m_{02} + m_{12}x + m_{22}x^2)},$$

and

$$\sigma_Y^2(x) = -\frac{1}{2(m_{02} + m_{12}x + m_{22}x^2)}$$

Bhattacharyya's Distribution

- ▶ In fact, Bhattacharyya's distribution characterizes *all* bivariate distributions with normal conditionals²
- ▶ Gibbs sampler to draw from $p_{X,Y}$ is easy to implement:
 - ▶ Choose starting point $(x^{(0)}, y^{(0)})$
 - ▶ At iteration t draw

$$X^{(t)} \sim \text{Normal}[\mu_X(y^{(t-1)}), \sigma_X^2(y^{(t-1)})]$$

$$Y^{(t)} \sim \text{Normal}[\mu_Y(x^{(t)}), \sigma_Y^2(x^{(t)})]$$

²Arnold, Castillo and Sarabia (Statistical Science, 2001):

R Time!

Open file `Lecture11code.R`, part 1

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Example: Multivariate Normal

- Distribution of the data

$$\mathbf{Z} = \{Z_i\}_{i=1}^n \mid \mu, \Lambda \stackrel{i.i.d.}{\sim} \text{Normal}(\mu, \Lambda^{-1})$$

where $Z_i \in \mathbb{R}^K$, μ is the vector of means, Λ^{-1} is the covariance matrix, and Λ is the inverse covariance matrix (the *precision matrix*)

- Conjugate prior is constructed in two steps

$$\begin{aligned}\mu \mid \Lambda &\sim \text{Normal}(\mu_0, (\kappa_0 \Lambda)^{-1}) \\ \Lambda &\sim \text{Wishart}(v_0, W_0)\end{aligned}$$

Joint distribution of (μ, Λ) is called *Normal-Wishart*. The parameterization is such that $E(\Lambda) = v_0 W_0$

Example: Multivariate Normal

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Example: Multivariate Normal

Posterior is also Normal-Wishart

$$\mu \mid \Lambda, \mathbf{z} \sim \text{Normal}(\mu', (\kappa' \Lambda)^{-1})$$

$$\Lambda \mid \mathbf{z} \sim \text{Wishart}(v', W')$$

where

$$\mu' = (\kappa_0 \mu_0 + n \bar{z}) / \kappa'$$

$$\kappa' = \kappa_0 + n$$

$$v' = v_0 + n$$

$$W' = \{W_0^{-1} + n[\hat{\Sigma} + \frac{\kappa_0}{\kappa'}(\bar{z} - \mu_0)(\bar{z} - \mu_0)^T]\}^{-1}$$

$$\bar{z} = \sum_{i=1}^n z_i / n$$

$$\hat{\Sigma} = \sum_{i=1}^n (z_i - \bar{z})(z_i - \bar{z})^T / n$$

Example: Multivariate Normal

HW3: write down and implement a data augmentation algorithm under ignorability and multivariate normality

R Time!

Open file `Lecture11code.R`, part 2

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Multivariate Imputation by Chained Equations

Multivariate Imputation by Chained Equations (MICE)³ is an ad-hoc multiple imputation procedure that builds on Gibbs sampling ideas

- ▶ If each Y_1, \dots, Y_K is subject to missingness, we can posit K different regression models

$$p_1(y_1 \mid y_{-1}, \theta_1)$$

$$p_2(y_2 \mid y_{-2}, \theta_2)$$

$$\vdots$$

$$p_K(y_K \mid y_{-K}, \theta_K)$$

- ▶ θ_k : parameters of the k th conditional distribution
- ▶ $y_{-k} = (y_1, \dots, y_{k-1}, y_{k+1}, \dots, y_K)$
- ▶ Key idea: use these models to sequentially impute, one variable at a time. Repeat this over a number of iterations

³<https://www.jstatsoft.org/article/view/v045i03/v45i03.pdf>

Multivariate Imputation by Chained Equations

The MICE algorithm:

- ▶ Initialize the algorithm by randomly imputing the missing values of each variable/column by observed values of that variable/column.

Denote this initial completed data as $\mathbf{y}_1^{(0)}, \dots, \mathbf{y}_K^{(0)}$

- ▶ Run a pseudo Gibbs/Data Augmentation sampler, with t th iteration:

$$\theta_1^{(t)} \sim p_1(\theta_1 \mid \mathbf{y}_{1(r_1)}, \mathbf{y}_2^{(t-1)}, \dots, \mathbf{y}_K^{(t-1)}) \propto p_1(\theta_1) \prod_{i:r_{i1}=1} p_1(y_{i1} \mid y_{i2}^{(t-1)}, \dots, y_{iK}^{(t-1)}, \theta_1)$$

$$y_{i1}^{(t)} \sim p_1(y_1 \mid y_{i2}^{(t-1)}, \dots, y_{iK}^{(t-1)}, \theta_1^{(t)}), \text{ for all missing } y_{i1}$$

$$\vdots$$

$$\theta_K^{(t)} \sim p_K(\theta_K \mid \mathbf{y}_{K(r_K)}, \mathbf{y}_1^{(t)}, \dots, \mathbf{y}_{K-1}^{(t)}) \propto p_K(\theta_K) \prod_{i:r_{iK}=1} p_K(y_{iK} \mid y_{i1}^{(t)}, \dots, y_{i,K-1}^{(t)}, \theta_K)$$

$$y_{iK}^{(t)} \sim p_K(y_K \mid y_{i1}^{(t)}, \dots, y_{i,K-1}^{(t)}, \theta_K^{(t)}), \text{ for all missing } y_{iK}$$

- ▶ Iterate for a number of times

R Time!

Open file Lecture11code.R, part 3

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Summary

Main take-aways from today's lecture:

- ▶ Example of Gibbs sampler
- ▶ Data Augmentation and Proper Multiple Imputation using 'norm' package
- ▶ Multivariate Imputation by Chained Equations with the 'mice' package

Next lecture:

- ▶ Inverse Probability Weighting