

# Sequential Monte Carlo methods

## Lecture 15 – General Sequential Monte Carlo

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## Outline – Lecture 15

**Aim:** Show how SMC can be used for a much wider class of problems than inference in state-space models.

### Outline:

1. Summary of day 3
2. Examples of probabilistic models
3. General SMC formulation
4. Locally optimal proposals

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## Summary of day 3

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Simulate a Markov chain which is designed in such a way that its stationary distribution coincides with the target distribution.

An MCMC sampler generates the Markov chain  $\{x[m]\}_{m=1}^M$  by:

- **Initialize:** set  $x[1]$  arbitrarily.
- **For**  $m = 2$  **to**  $M$ : sample  $x[m] \sim \kappa(x[m-1], x^*)$ .

$\kappa(x, x^*)$  is a **Markov kernel** on  $\mathcal{X}$ , i.e. a conditional distribution for the next state  $x^*$  given the current state  $x$ .

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## Summary of day 3

### Algorithm 1 Pseudo-marginal Metropolis Hastings

1. **Initialize** ( $m = 1$ ): Set  $\theta[1]$  and run a particle filter for  $\hat{z}[1]$ .

2. **For**  $m = 2$  **to**  $M$ , **iterate**:

a. Sample  $\theta' \sim q(\theta | \theta[m-1])$ .

b. Sample  $\hat{z}' \sim \psi(z | \theta', y_{1:T})$  (i.e. run a particle filter).

c. With probability

$$\alpha = \min \left( 1, \frac{\hat{z}' p(\theta')}{\hat{z}[m-1] p(\theta[m-1])} \frac{q(\theta[m-1] | \theta')}{q(\theta' | \theta[m-1])} \right)$$

set  $\{\theta[m], \hat{z}[m]\} \leftarrow \{\theta', \hat{z}'\}$  (accept candidate sample) and

with prob.  $1 - \alpha$  set  $\{\theta[m], \hat{z}[m]\} \leftarrow \{\theta[m-1], \hat{z}[m-1]\}$  (reject candidate sample).

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## Summary of day 3

**Gibbs kernel:** This procedure defines a Markov kernel  $\kappa(x, x^*)$  with stationary distribution  $\pi(x)$ .

Particle Gibbs kernel: A Markov kernel  $\kappa_{N,\theta}(x_{0:T}, x_{0:T}^*)$  on  $\mathcal{X}^{T+1}$ .

**Particle Gibbs:** Run a particle filter, but at each time step

- sample only  $N - 1$  particles in the standard way.
- set the  $N$ th particle deterministically:  $x_t^N = x_t$  and  $a_t^N = N$ .
- At final time  $t = T$ , output  $x_{0:T}^* = x_{0:T}^b$  with  $b \sim \mathcal{C}(\{w_T^i\}_{i=1}^N)$

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## Summary of day 3

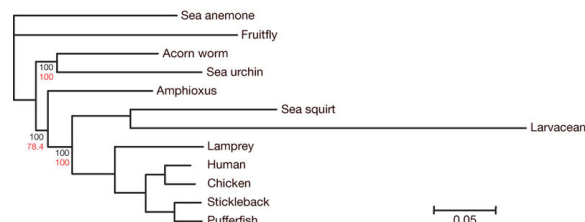
- The algorithm stochastically “maps”  $x_{0:T}$  into  $x_{0:T}^*$ .
- Implicitly defines a Markov kernel  $\kappa_{N,\theta}(x_{0:T}, x_{0:T}^*)$  on  $\mathcal{X}^{T+1}$   
— **the particle Gibbs kernel**.

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## Examples of probabilistic models

## Phylogenetic trees

A phylogenetic (evolutionary) tree shows the inferred evolutionary relationships among various species based upon similarities and differences in their physical or genetic characteristics.

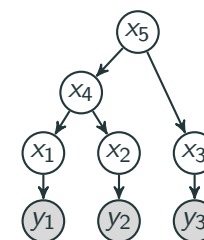
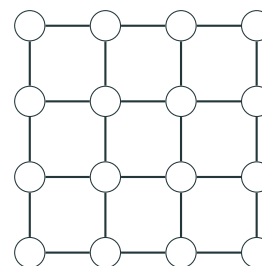


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## Probabilistic graphical models

A **probabilistic graphical model** (PGM) is a probabilistic model where a graph  $\mathcal{G} = (\mathcal{V}, \mathcal{E})$  represents the conditional independency structure between random variables,

1. a set of **vertices**  $\mathcal{V}$  (nodes) represents the random variables
2. a set of **edges**  $\mathcal{E}$  containing elements  $(i, j) \in \mathcal{E}$  connecting a pair of nodes  $(i, j) \in \mathcal{V} \times \mathcal{V}$



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## Gaussian process state-space model

The Gaussian process (GP) is a **non-parametric** and **probabilistic** model for nonlinear functions.

**Non-parametric** means that it does not rely on any particular parametric functional form to be postulated.

$$\begin{aligned} X_t &= f(X_{t-1}) + V_t, & \text{s.t. } f(X) &\sim \mathcal{GP}(0, \kappa_{\eta, f}(x, x')), \\ Y_t &= g(X_t) + E_t, & \text{s.t. } g(X) &\sim \mathcal{GP}(0, \kappa_{\eta, g}(x, x')). \end{aligned}$$

The model functions  $f$  and  $g$  are assumed to be realizations from Gaussian process priors and  $V_t \sim \mathcal{N}(0, Q)$ ,  $E_t \sim \mathcal{N}(0, R)$ .

**Task:** Compute the posterior  $p(f, g, Q, R, \eta, x_{0:T} | y_{1:T})$ .

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## General SMC formulation

## Model specification

SMC can be used to approximate a **sequence** of probability distributions on a sequence of probability spaces of **increasing dimension**.

Let  $\{\pi_k(\mathbf{x}_{1:k})\}_{k \geq 1}$  be an arbitrary sequence of target distributions

$$\pi_k(\mathbf{x}_{1:k}) = \frac{\tilde{\pi}_k(\mathbf{x}_{1:k})}{Z_k}$$

- The domain of  $\mathbf{x}_k$  is  $\mathcal{X}_k$ , and  $\mathcal{X}_{1:k} = \mathcal{X}_k \times \mathcal{X}_{1:k-1}$  for all  $k$ .
- $\tilde{\pi}_k(\mathbf{x}_{1:k})$  can be evaluated pointwise.
- The normalizing constant  $Z_k$  may be unknown.

Common tasks:

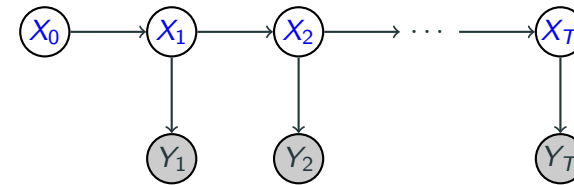
1. Approximate the normalization constant  $Z_k$ .
2. Approximate  $\pi_k(\mathbf{x}_k)$  and compute  $\int \phi(\mathbf{x}_k) \pi_k(\mathbf{x}_k) d\mathbf{x}_k$ .

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## ex) State space model

The sequence of target distributions  $\{\pi_k(\mathbf{x}_{0:k})\}_{k=1}^n$  can be constructed in **many** different ways.

The most basic construction arises from **chain-structured graphs**, such as the state space model (SSM).



$$p(\mathbf{x}_{0:t} | \mathbf{y}_{1:t}) = \frac{\tilde{\pi}_t(\mathbf{x}_{0:t})}{p(\mathbf{y}_{1:t})}$$

$$Z_t = \int \tilde{\pi}_t(\mathbf{x}_{0:t}) d\mathbf{x}_{0:t}$$

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## General Sequential Monte Carlo

Sequential Monte Carlo approximates

$$\pi_k(\mathbf{x}_{0:k}) \approx \sum_{i=1}^N w_k^i \delta_{\mathbf{x}_{0:k}^i}(\mathbf{x}_{0:k}).$$

The weighted particle populations  $\{\mathbf{x}_{0:k}^i, w_k^i\}_{i=1}^N$  are generated sequentially for  $k = 1, 2, \dots$

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## General Sequential Monte Carlo



Assume that we have obtained  $\{\mathbf{x}_{0:k-1}^i, w_{k-1}^i\}_{i=1}^N$

**Resampling:** Sample  $\mathbf{a}_k^i$  with  $\mathbb{P}(\mathbf{a}_k^i = j) = \nu_{k-1}^j$ ,  $j = 1, \dots, N$ .

**Propagation:**  $\mathbf{x}_k^i \sim q_k(\mathbf{x}_k | \mathbf{x}_{1:k-1}^{\mathbf{a}_k^i})$  and  $\mathbf{x}_{0:k}^i = (\mathbf{x}_{0:k-1}^{\mathbf{a}_k^i}, \mathbf{x}_k^i)$ .

**Weighting:**  $w_k^i \propto \frac{w_{k-1}^{\mathbf{a}_k^i}}{\nu_{k-1}^{\mathbf{a}_k^i}} \frac{\tilde{\pi}_k(\mathbf{x}_{0:k}^i)}{\tilde{\pi}_{k-1}(\mathbf{x}_{0:k-1}^{\mathbf{a}_k^i}) q_k(\mathbf{x}_k^i | \mathbf{x}_{0:k-1}^{\mathbf{a}_k^i})}$ .

The result is a new weighted set of particles  $\{\mathbf{x}_{0:k}^i, w_k^i\}_{i=1}^N$ .

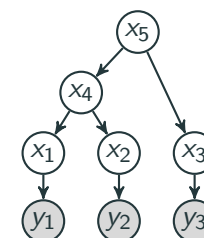
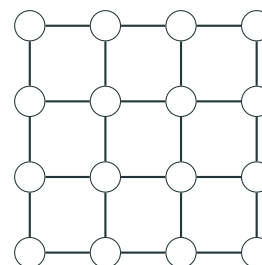
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## SMC for probabilistic graphical models

## Recall – Probabilistic graphical models

A **probabilistic graphical model** (PGM) is a probabilistic model where a graph  $\mathcal{G} = (\mathcal{V}, \mathcal{E})$  represents the conditional independency structure between random variables,

1. a set of **vertices**  $\mathcal{V}$  (nodes) represents the random variables
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## Key idea

SMC methods are used to approximate a **sequence of probability distributions** on a sequence of spaces of increasing dimension.

### Key idea:

1. Introduce a **sequential decomposition** of the PGM.
2. Each **subgraph** induces an intermediate target dist.
3. Apply SMC to the sequence of intermediate target dist.

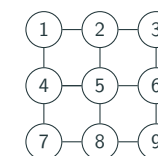
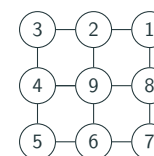
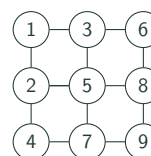
Using an artificial sequence of intermediate target distributions for an SMC method is a powerful (quite possibly underutilized) idea.

**Key question:** Exactly how do we define  $\{\tilde{\pi}_k(\mathbf{x}_{1:k})\}_{k \geq 1}$ ?

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## ex) Illustrating possible graph decomposition

Using a 2D lattice model from statistical physics,  $\mathbf{x} \in (-\pi, \pi]$ .



$$p(\mathbf{x}_{\mathcal{V}}) \propto e^{-\beta H(\mathbf{x}_{\mathcal{V}})}, \quad H(\mathbf{x}_{\mathcal{V}}) = - \sum_{(i,j) \in \mathcal{E}} J_{ij} \cos(\mathbf{x}_i - \mathbf{x}_j),$$

The **intermediate sequence** of target distributions can be chosen

$$\tilde{\pi}_k(\mathbf{x}_{\mathcal{L}_k}) \propto \tilde{\pi}_{k-1}(\mathbf{x}_{\mathcal{L}_{k-1}}) e^{\kappa(\mathbf{x}_{\mathcal{L}_{k-1}}) \cos(\mathbf{x}_k - \mu(\mathbf{x}_{\mathcal{L}_{k-1}}))}.$$

$\mathcal{L}_k$  – index to the nodes in the  $k^{\text{th}}$  intermediate target  $\tilde{\pi}_k(\mathbf{x}_{\mathcal{L}_k})$ .

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## SMC for graphical models – algorithm

### Algorithm SMC for graphical models

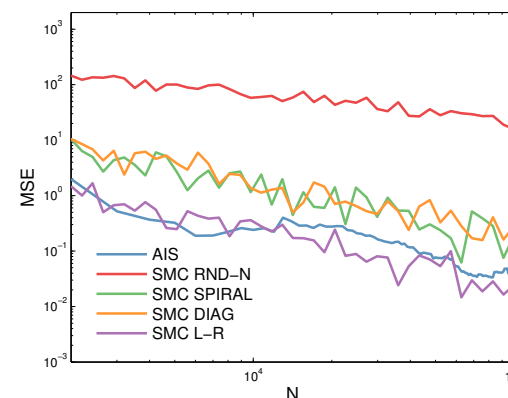
1. **Initialize** ( $k = 1$ ):
  - (a) Draw  $x_{\mathcal{L}_1}^i \sim q_1(\cdot)$ .
  - (b) Set  $w_1^i = W_1(x_{\mathcal{L}_1}^i)$ .
2. **For**  $k = 2$  **to**  $K$  **do**:
  - (a) **Resampling**: Draw  $a_k^i$ ,  $\mathbb{P}(a_k^i = j) = \tilde{w}_{k-1}^j / \sum_l \tilde{w}_{k-1}^l$ .
  - (b) **Propagation**: Draw  $\xi_k^i \sim q_k(\cdot | x_{\mathcal{L}_{k-1}}^{a_k^i})$ , set  $x_{\mathcal{L}_k}^i = x_{\mathcal{L}_{k-1}}^{a_k^i} \cup \xi_k^i$ .
  - (c) **Weighting**: Set  $\tilde{w}_k^i = W_k(x_{\mathcal{L}_k}^i)$ .
3. **End**

$\mathcal{L}_k$  – index to the nodes in the  $k^{\text{th}}$  intermediate target  $\tilde{\pi}_k(x_{\mathcal{L}_k})$ .

$\xi_k^i$  – nodes added at step  $k$ .

Also provides an unbiased estimate of the **normalizing constant**! 16/19

## ex) Classical XY-model



This model is borrowed from

John M. Kosterlitz and David J. Thouless. **Ordering, metastability and phase transitions in two-dimensional systems.** *J. of Physics C: Solid State Physics*, 6(7):1181–1203, 1973.

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## What about stability?

**Recall:** for a state-space model we need exponential forgetting for the particle filter to be stable.

### The same is true in the general case!

If there are **strong** and **long-ranging** dependencies among the variables  $X_{1:k}$  under the distribution  $\pi_k$ , then the asymptotic variance of SMC may be exponential in  $k$ .

However,

- In many applications we *do* have fast enough forgetting (though, it can be difficult to verify theoretically)
- Even if this is not the case, SMC can give good results for moderate values of  $k$

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## Further reading

SMC for phylogenetic trees:

Alexandre Bouchard-Côté. **Sequential Monte Carlo (SMC) for Bayesian phylogenetics.** *Bayesian phylogenetics: methods, algorithms, and applications*, 163–186, 2014.

SMC for graphical models:

Christian A. Naesseth, Fredrik Lindsten and Thomas B. Schön. **Sequential Monte Carlo methods for graphical models.** *Advances in Neural Information Processing Systems (NIPS)*, Montreal, Canada, December, 2014.

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