# Probability & Statistics for EECS: Homework #09

Due on Apr 16, 2023 at  $23{:}59$ 

Name: Wang Penghao Student ID: 2021533138

(a) As for the case X, Y are discrete, we have that use the definition of conditioning probability

$$P(Y = y | X = x) = \frac{P(X = x, Y = Y)}{P(X = x)}$$

$$P(X = x | Y = y) = \frac{P(X = x, Y = Y)}{P(Y = y)}$$

So that we have that

$$P(Y = y | X = x) = \frac{P(X = x, Y = Y)}{P(X = x)} = \frac{P(X = x | Y = y)P(Y = y)}{P(X = x)}$$

(b) As for the case X, Y are continuous, we have that use the definition of conditioning probability

$$f_{Y|X}(y|x) = \frac{f_{Y,X}(y,x)}{f_X(x)}$$

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)}$$

So that we have that

$$f_{Y|X}(y|x) = \frac{f_{Y,X}(y,x)}{f_X(x)} = \frac{f_{X|Y}(x|y)f_Y(y)}{f_X(x)}$$

(c) As for the case X is discrete, Y is continuous, we have that use the definition of conditioning probability Denote that  $\varepsilon$  is a small number, then we have that

$$P[Y \in (y - \varepsilon, y + \varepsilon)|X = x] = \frac{P[X = x|Y \in (y - \varepsilon, y + \varepsilon)]P[Y \in (y - \varepsilon, y + \varepsilon)]}{P(X = x)}$$

$$f_Y(y|X = x) = \lim_{\varepsilon \to 0} \frac{P[Y \in (y - \varepsilon, y + \varepsilon)|X = x]}{2\varepsilon}$$

$$= \lim_{\varepsilon \to 0} \frac{P[X = x|Y \in (y - \varepsilon, y + \varepsilon)]\frac{P[Y \in (y - \varepsilon, y + \varepsilon)]}{2\varepsilon}}{P(X = x)}$$

$$= \frac{P(X = x|Y = y)f_Y(y)}{P(X = x)}$$

So that we have that

$$f_Y(y|X = x) = \frac{P(X = x|Y = y)f_Y(y)}{P(X = x)}$$

(d) As for the case X is continuous, Y is discrete, we have that use the definition of conditioning probability Denote that  $\varepsilon$  is a small number, then we have that

$$\begin{split} \lim_{\varepsilon \to 0} P(Y = y | X \in (x - \varepsilon, x + \varepsilon)) &= \lim_{\varepsilon \to 0} \frac{P[X \in (x - \varepsilon, x + \varepsilon) | Y = y] P(Y = y)}{P[X \in (x - \varepsilon, x + \varepsilon)]} \\ &= \lim_{\varepsilon \to 0} \frac{P[X \in (x - \varepsilon, x + \varepsilon) | Y = y]}{\frac{2\varepsilon}{P[X \in (x - \varepsilon, x + \varepsilon)]}} P(Y = y)} \\ &= \frac{f_X(x | Y = y) P(Y = y)}{f_X(x)} \end{split}$$

So we get that 
$$P(Y = y|X = x) = \frac{f_X(x|Y = y)P(Y = y)}{f_X(x)}$$

(a) As for the joint PMF of X, Y, N, we have that the PMF is P(X = x, Y = y, N = n), then as we have that N = X + Y, then only when x + y = n, will the PMF be non-zero. So we have that

$$P(X = x, Y = y, N = n) = P(X = x, Y = y) = (1 - p)^{x} * p * (1 - p)^{y} * p = (1 - p)^{x+y}p^{2}$$

, as we have that x + y = n, so we get that

$$P(X = x, Y = y, N = n) = (1 - p)^n p^2$$

(b) As for the joint PMF os X, N, we have that the PMF is P(X = x, N = n), as only when n = x + y will the PMF be non-zero, so we have that

$$P(X = x, N = n) = P(X = x, Y = n - x) = (1 - p)^{x} p(1 - p)^{n - x} p = (1 - p)^{n} p^{2}$$

(c) As for the conditional PMF of X given N = n, we have that the PMF is

$$P(X = x | N = n) = \frac{P(X = x, N = n)}{P(N = n)}.$$

The numerator is the joint PMF of X and N, which is  $P(X = x, N = n) = (1 - p)^n p^2$ , and the denominator is PMF of N, which is  $P(N = n) = \sum_{x=0}^{n} (1 - p)^n p^2 = (n + 1)(1 - p)^n p^2$ , so we have that

$$P(X = x | N = n) = \frac{P(X = x, N = n)}{P(N = n)} = \frac{(1 - p)^n p^2}{(n + 1)(1 - p)^n p^2} = \frac{1}{n + 1}.$$

where x = 0, 1, 2, ..., n.

Description: The conditional PMF of X given N=n is a uniform distribution, which is  $P(X=x|N=n)=\frac{1}{n+1}$ . The event P(X=x) is a Geom distribution, while the event N=n is actually a negative binomial distribution, which denote the fail times before the second success. So the conditional PMF of X given N=n is  $\frac{1}{n+1}$ , which denote that the first success between the first and the second success is uniformly distributed.

(a) To verify that the conditional distribution of X given X > c is the same as the distribution of c + X, firstly we can find the corresponding CDF of X given X > c, which is  $P(X \le x | X > c) = \frac{P(c < X \le x)}{P(X > c)} = \frac{F(x) - F(c)}{1 - F(c)}$ . As  $X \sim Expo(\lambda)$ , so we have  $F(x) = 1 - e^{-\lambda x}$ . So, the  $P(X \le x | X > c) = \frac{e^{-\lambda c} - e^{-\lambda x}}{e^{-\lambda c}} = \frac{1 - e^{-\lambda(x-c)}}{1 - e^{-\lambda(x-c)}}$ .

As for the CDF of c + X, we have that  $P(c+X \le x) = P(X \le x-c) = 1 - e^{-\lambda(x-c)}$ . So we have that  $P(X \le x|X > c) = P(c+X \le x)$ . So that the conditional CDF of X given X > c is the same as the c + X.

(b) As for the CDF of X given X < c, we have that for x < c,  $P(X \le x|X < c) = \frac{P(X \le x, X < c)}{P(X < c)} = \frac{P(X \le x)}{P(X < c)} = \frac{1 - e^{-\lambda x}}{1 - e^{-\lambda c}}$  As for the PDF, we have that  $f(x|X < c) = (P(X \le x|X < c))' = (\frac{P(X \le x)}{P(X < c)})' = (\frac{1 - e^{-\lambda x}}{1 - e^{-\lambda c}})' = \frac{\lambda e^{-\lambda x}}{1 - e^{-\lambda c}}$  for x < c, as for  $x \ge c$ , PDF is zero.

As we have that  $U_1, U_2, U_3$  be i.i.d. Unif(0, 1), and let  $L = min(U_1, U_2, U_3)$ ,  $M = max(U_1, U_2, U_3)$ 

- (a) 1. As for the marginal CDF of M is  $F_M(m) = P(M \le m) = P(U_1 \le m, U_2 \le m, U_3 \le m) = P(U_1 \le m)P(U_2 \le m)P(U_3 \le m) = m^3$ . For  $m \in [0, 1]$ 
  - 2. As for the marginal PDF of M, we have that  $f_M(m) = (F_M(m))' = (m^3)' = 3m^2$ . For  $m \in [0,1]$
  - 3. As for the joint CDF of M and L, firstly we consider the event  $L > l, M \le m$ , which is easy to calculate, that is  $P(L > l, M \le m) = (m-l)^3$ , we have that  $P(L \le l, M \le m) = P(M \le m) P(L > l, M \le m) = m^3 (m-l)^3$  for  $m \ge l$  and that  $m, l \in [0, 1]$ .
  - 4. As for the joint PDF of M and L, we have that  $f(l,m) = \frac{\partial^2 P(L \leq l, M \leq m)}{\partial l \partial m} = 6(m-l)$  for  $m, l \in [0, 1]$  and that  $m \geq l$ .
- (b) As for the conditional PDF of M given L, firstly we have that  $P(L > l) = P(U_1 \ge l, U_2 \ge l, U_3 \ge l) = (1 l)^3$ . Then we have that  $P(L \le l) = 1 P(L > l) = 1 (1 l)^3$ , then we get that  $f_L(l) = 3(1 l)^2$ , where  $l \in [0, 1]$ . So we get  $f_{M|L}(m|l) = \frac{f(l, m)}{f_L(l)} = \frac{6(m l)}{3(1 l)^2} = \frac{2(m l)}{(1 l)^2}$ , where  $m, l \in [0, 1]$  and that  $m \ge l$ .

Firstly, we denote that q = 1 - p

(a) As we have that L = min(X, Y), M = max(X, Y), then we have that for l < m, we have  $P(L = l, M = m) = P(X = l, Y = m) + P(X = m, Y = l) = q^l p q^m p + q^m p q^l p = 2p^2 q^{l+m}$ 

for l=m, we have  $P(L=l,M=m)=P(X=l,Y=l)=q^lpq^lp=p^2q^{2l}$ 

for l > m, we have P(L = l, M = m) = P(X = l, Y = m) + P(X = m, Y = l) = 0

So we get that the joint PMF of L and M is

$$P(L = l, M = m) = \begin{cases} 2p^2q^{l+m} & l < m \\ p^2q^{2l} & l = m \\ 0 & l > m \end{cases}$$

However, as we have that the joint PMF of X and Y depend on l and m's size, it is clearly that the joint can't be get by multiplying the function of l and m, so L and M are no independent.

- (b) As for the marginal distribution of L
  - (1) By using the joint PMF, we have that

$$\begin{split} P(L=l) &= \sum_{m=l}^{\infty} P(L=l, M=m) \\ &= p^2 q^{2l} + \sum_{m=l+1}^{\infty} 2p^2 q^{l+m} \\ &= p^2 q^{2l} + 2p^2 q^l \sum_{m=l+1}^{\infty} q^m \\ &= p^2 q^{2l} + 2p^2 q^l q^{l+1} \left(\frac{1-q^{\infty}}{1-q}\right) \\ &= p^2 q^{2l} + 2p^2 q^l q^{l+1} \frac{1}{p} \\ &= p^2 q^{2l} + 2pq^{2l+1} \end{split}$$

(2) Also we we can use story, that is, consider 2 independent Bernuoilli independent sequence, at the time n, at least one of the 2 sequences has a success, then, we have that the probability of the event is  $1 - (1 - p)^2$ , so we have that  $P(L = l) = (1 - (1 - q^2))^l (1 - q^2) = q^{2l} (1 - q^2)$ , as we have that  $(p+q)^2 = 1$ , so  $1 - q^2 = p^2 + 2pq$ , so the  $P(L = l) = q^{2l}(p^2 + 2pq) = p^2q^{2l} + 2pq^{2l+1}$ .

So, both using joint PMF and using story, can we find the marginal distribution of L.

(c) As we have that L = min(X, Y), M = max(X, Y), then we have that L + M = X + Y, then EL + EM = E(L + M) = E(X + Y) = E(X) + E(Y), as X and Y i.i.d. Geom(p), we have that E(X) = E(Y) = E(X) + E(X) +

$$2\frac{q}{p} = \frac{2q}{p}$$
 Then, we find the  $E(L)$ , as  $P(L=l) = p^2q^{2l} + 2pq^{2l+1}$ , so

$$E(L) = \sum_{l=0}^{\infty} lP(L = l)$$

$$= (p^2 + 2pq) \sum_{l=1}^{\infty} lq^{2l}$$

$$q^2 E(L) = (p^2 + 2pq) \sum_{l=1}^{\infty} lq^{2l+2}$$

$$E(L) - q^2 E(l) = (p^2 + 2pq) \sum_{l=1}^{\infty} lq^{2l+2}$$

$$= (1 - q^2) \frac{q^2}{1 - q^2}$$

$$E(L) = \frac{q^2}{1 - q^2}$$

So we get that  $EM = \frac{2q}{p} - EL = \frac{2q}{p} - \frac{q^2}{1 - q^2} = \frac{(1 - p)(3 - p)}{p(2 - p)}$ 

(d) As for the joint PMF of L and M-L, we have that for  $k \geq 0$ 

$$P(L = l, M - L = k) = P(L = l, M = k + l),$$

when k > 0, we have

$$\begin{split} P(L=l, M-L=k) &= P(X=l, Y=k+l) + P(X=k+l, Y=l) \\ &= q^l p q^{k+l} p + q^{k+l} p q^l p \\ &= 2 p^2 q^{2l+k}. \end{split}$$

When k=0, we have that  $P(L=l,M-L=k)=p^2q^{2l+k}$ . As for the P(L=l), we have  $P(L=l)=(1-q^2)q^{2l}$ As for P(M-L=k) For k>0, we have that

$$P(M - L = k) = \sum_{l=0}^{\infty} P(L = l, M = l + k) P(L = l) = \frac{2p^2q^k}{1 - q^2},$$

where P(M-L=k)P(L=l)=P(L=l,M-L=k), L and M-L are independent and for k=0, we have that  $P(M-L=k)=P(M=L)=\sum_{l=0}^{\infty}P(M=L=l)=\frac{p^2q^k}{1-q^2}.$  So we have for k=0, P(L=l,M-l=k)=P(M-L=k)P(L=l). In conclusion, L and M - L are independent.