Instructions for Paper Submissions to AISTATS 2023: Supplementary Materials

THORETICAL DERIVATIONS 1

Proof that $\hat{\sigma}^2(z_1, z_2)$ is an unbiased estimator of $\sigma^2_*(z_1, z_2)$

In Section .. we use that $\hat{\sigma}^2(z_1,z_2)$ is an unbiased estimator of $\sigma^2_*(z_1,z_2) = \frac{\int_{z_1}^{z_2} \mathbb{E}_{X_c|X_s=z} \left[(f^s(z,X_c) - \mu(z_1,z_2))^2 \right] \partial z}{z_2 - z_1}$, under the assumption that the points are uniformly distributed in $[z_1,z_2)$, i.e., $x_s^i \sim \mathcal{U}(z_1,z_2)$.

Proof We have made the hypothesis that the points \mathbf{x}^i are coming from the distibution $\mathcal{S} = p(X_c|X_s)p(X_s) =$ $p(X_c|X_s)\mathcal{U}(x_s;z_1,z_2) = \frac{p(\mathbf{x}_c|z)}{z_2-z_1}$ uniformly distributed in the interval $[z_1,z_2)$ in terms

$$\mathbb{E}_{\mathcal{S}}[\hat{\sigma}^{2}(z_{1}, z_{2})] = \mathbb{E}_{\mathcal{S}}\left[\frac{1}{|\mathcal{S}|} \sum_{i=1}^{|\mathcal{S}|} (f^{s}(\mathbf{x}^{i}) - \mu(z_{1}, z_{2}))^{2}\right]$$
(1)

$$= \frac{1}{|S|} \sum_{i=1}^{|S|} \mathbb{E}_{S}[(f^{s}(\mathbf{x}^{i}) - \mu(z_{1}, z_{2}))^{2}]$$
 (2)

$$= \frac{1}{|S|} \sum_{i=1}^{|S|} \mathbb{E}_{S}[(f^{s}(\mathbf{x}^{i}) - \mu(z_{1}, z_{2}))^{2}]$$
(3)

(4)

$$\sigma_*^2(z_1, z_2) = \frac{\int_{z_1}^{z_2} \mathbb{E}_{X_c \mid X_s = z} \left[(f^s(z, X_c) - \mu(z_1, z_2))^2 \right] \partial z}{z_2 - z_1}$$
 (5)

$$= \mathbb{E}_{z \sim \mathcal{U}(z_1, z_2)} \mathbb{E}_{X_c \mid X_s = z} \left[\left(f^s(z, X_c) - \mu(z_1, z_2) \right)^2 \right]$$
 (6)

$$= \mathbb{E}_{z \sim \mathcal{U}(z_1, z_2)} \mathbb{E}_{X_c \mid X_s = z} \left[\left(f^s(z, X_c) - \mu(z_1, z_2) \right)^2 \right]$$

$$= \mathbb{E}_X \left[\left(f^s(X) - \mu(z_1, z_2) \right)^2 \right]$$
(6)
$$= \mathbb{E}_X \left[\left(f^s(X) - \mu(z_1, z_2) \right)^2 \right]$$
(7)

Proof Of Theorem 3.1 1.2

If we define (a) the residual $\rho(z)$ as the difference between the expected effect at z and the bin-effect, i.e $\rho(z)=\mu(z)$ $\mu(z_1,z_2)$ and (b) $\mathcal{E}(z_1,z_2)$ as the mean squared residual of the bin, i.e. $\mathcal{E}(z_1,z_2)=\frac{\int_{z_1}^{z_2} \rho^2(z)\partial z}{z_2-z_1}$, then it holds that:

$$\sigma_*^2(z_1, z_2) = \sigma^2(z_1, z_2) + \mathcal{E}^2(z_1, z_2)$$
(8)

Proof

$$\sigma_{*}^{2}(z_{1}, z_{2}) = \frac{1}{z_{2} - z_{1}} \int_{z_{1}}^{z_{2}} \mathbb{E}_{X_{c}|z} \left[(f^{s}(z, X_{c}) - \mu(z_{1}, z_{2}))^{2} \right] \partial z$$

$$= \frac{1}{z_{2} - z_{1}} \int_{z_{1}}^{z_{2}} \mathbb{E}_{X_{c}|z} \left[(f^{s}(z, X_{c}) - \mu(z) + \rho(z))^{2} \right] \partial z$$

$$= \frac{1}{z_{2} - z_{1}} \int_{z_{1}}^{z_{2}} \mathbb{E}_{X_{c}|z} \left[(f^{s}(z, X_{c}) - \mu(z))^{2} + \rho(z)^{2} + 2f^{s}(z, X_{c})\mu(z)) \right]$$

$$= \frac{1}{z_{2} - z_{1}} \int_{z_{1}}^{z_{2}} \left(\underbrace{\mathbb{E}_{X_{c}|z} \left[(f^{s}(z, X_{c}) - \mu(z))^{2} \right]}_{\sigma^{2}(z)} + \underbrace{\mathbb{E}_{X_{c}|z} \left[\rho^{2}(z) \right]}_{\rho^{2}(z)} + 2(\underbrace{\mathbb{E}_{X_{c}|z} \left[(f^{s}(z, X_{c})) - \mu(z) \rho(z) \right)}_{\mu(z)} \right) \partial z$$

$$= \underbrace{\frac{1}{z_{2} - z_{1}} \int_{z_{1}}^{z_{2}} \sigma^{2}(z) \partial z}_{\sigma^{2}(z)} + \underbrace{\frac{1}{z_{2} - z_{1}} \int_{z_{1}}^{z_{2}} \rho^{2}(z) \partial z}_{\varepsilon^{2}(z, z_{1})} + \mathcal{E}^{2}(z_{1}, z_{2}) + \mathcal{E}^{2}(z_{1}, z_{2})$$

$$(12)$$

1.3 Proof Of Corollary

If a bin-splitting Z minimizes the accumulated error, then it also minimizes $\sum_{k=1}^K \sigma_*^2(z_1,z_2) \Delta z_k$

We want to show that

$$\mathcal{Z}^* = \arg\min_{\mathcal{Z}} \sum_{k=1}^K \sigma_*^2(z_{k-1}, z_k) \Delta z_k \Leftrightarrow \mathcal{Z}^* = \arg\min_{\mathcal{Z}} \sum_{k=1}^K \mathcal{E}^2(z_{k-1}, z_k) \Delta z_k$$

Proof

$$\mathcal{Z}^* = \arg\min_{\mathcal{Z}} \sum_{k=1}^K \sigma_*^2(z_{k-1}, z_k) \Delta z_k$$
(14)

$$= \arg\min_{\mathcal{Z}} \left[\sum_{k=1}^{K} (\sigma^{2}(z_{k-1}, z_{k}) + \mathcal{E}^{2}(z_{k-1}, z_{k})) \Delta z_{k} \right]$$
 (15)

$$= \arg\min_{\mathcal{Z}} \left[\sum_{k=1}^{K} \left(\frac{\Delta z_k}{\Delta z_k} \int_{z_{k-1}}^{z_k} \sigma^2(z) \partial z \right) + \mathcal{E}^2(z_{k-1}, z_k) \Delta z_k \right) \right]$$
(16)

$$= \arg\min_{\mathcal{Z}} \left| \underbrace{\int_{z_0}^{z_K} \sigma^2(z) \partial z}_{\text{independent of } \mathcal{Z}} + \sum_{k=1}^K \mathcal{E}^2(z_{k-1}, z_k) \Delta z_k) \right|$$
 (17)

$$= \arg\min_{\mathcal{Z}} \sum_{k=1}^{K} \mathcal{E}^2(z_{k-1}, z_k) \Delta z_k \tag{18}$$

2 Dynamic Programming

For achieving a computationally-grounded solution we set a threshold K_{max} on the maximum number of bins which also discretizes the solution space. The width of the bin can take discrete values that are multiple of the minimum step $u = \frac{x_{s,max} - x_{s,min}}{K_{max}}$. For defining the solution, we use two indexes. The index $i \in \{0,\ldots,K_{max}\}$ denotes the point (z_i) and the index $j \in \{0,\ldots,K_{max}\}$ denotes the position of the j-th multiple of the minimum step, i.e., $x_j = x_{s,min} + j \cdot u$. The recursive cost function T(i,j) is the cost of setting $z_i = x_j$:

$$\mathcal{T}(i,j) = \min_{l \in \{0,...,K_{max}\}} \left[\mathcal{T}(i-1,l) + \mathcal{B}(x_l, x_j) \right]$$
(19)

where $\mathcal{T}(0,j)$ equals zero if j=0 and ∞ in any other case. $\mathcal{B}(x_l,x_j)$ denotes the cost of creating a bin with limits $[x_l,x_j)$:

$$\mathcal{B}(x_l, x_j) = \begin{cases} \infty, & \text{if } x_j > x_l \text{ or } |\mathcal{S}_{(x_j, x_l)}| < N \\ 0, & \text{if } x_j = x_l \\ \hat{\sigma}^2(x_j, x_l), & \text{if } x_j \le x_l \end{cases}$$

$$(20)$$

The optimal solution is given by solving $\mathcal{L} = \mathcal{T}(K_{max}, K_{max})$ and keeping track of the sequence of steps.

3 EXPERIMENTS APPEARED AT THE MAIN PAPER

- 3.1 Simulation 1
- 3.2 Simulation 2
- 3.3 Real World Example
- 4 Further Experimentation