

# Covariance

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# 1 Cholesky decomposition of the covariance

Since the covariance matrix  $P \in \mathbb{R}^{n \times n}$  is symmetric positive-definite, we can apply the Cholesky decomposition:

$$P = LL^T \quad (1)$$

where  $L \in \mathbb{R}^{n \times n}$  is a lower triangular matrix. Let's apply SVD to it.

$$L = USV^T = \sum_{i=1}^n \sigma_i u_i v_i^T \quad (2)$$

$\sigma_i > 0$  ( $i = 1, 2, \dots, n$ ) are the singular values of  $L$ ,  $u_i \in \mathbb{R}^n$  ( $i = 1, 2, \dots, n$ ) are the left singular vectors, and  $v_i \in \mathbb{R}^n$  ( $i = 1, 2, \dots, n$ ) are the right singular vectors of  $L$ .

With this we can express both  $L^T$  and  $L^{-1}$ :

$$L^T = VS^T U^T = VSU^T = \sum_{i=1}^n \sigma_i v_i u_i^T \quad (3)$$

$$L^{-1} = VS^{-1}U^T = \sum_{i=1}^n \frac{1}{\sigma_i} v_i u_i^T \quad (4)$$

Let's express  $P$ :

$$P = LL^T = (USV^T)(VSU^T) = US^2U^T = \sum_{i=1}^n \sigma_i^2 u_i u_i^T \quad (5)$$

This is the eigendecomposition of  $P$ , so its eigenvalues are  $\sigma_i^2$ , its eigenvectors are  $u_i$ .

## 1.1 Numerical example

Consider the following covariance matrix:

$$P = \begin{bmatrix} 2 & 0.5 \\ 0.5 & 1 \end{bmatrix} \quad (6)$$

The Cholesky decomposition is:

$$P = LL^T \quad (7)$$

where

$$L = \begin{bmatrix} 1.4142 & 0 \\ 0.3536 & 0.9354 \end{bmatrix} \quad (8)$$

The SVD of  $L$ :

$$\begin{aligned} L &= \sigma_1 u_1 v_1^T + \sigma_2 u_2 v_2^T \\ &= 1.4856 \cdot \begin{bmatrix} -0.9239 \\ -0.3827 \end{bmatrix} \cdot \begin{bmatrix} -0.9705 & -0.2410 \end{bmatrix} \\ &\quad + 0.8904 \cdot \begin{bmatrix} -0.3827 \\ 0.9239 \end{bmatrix} \cdot \begin{bmatrix} -0.2410 & 0.9705 \end{bmatrix} \end{aligned} \quad (9)$$

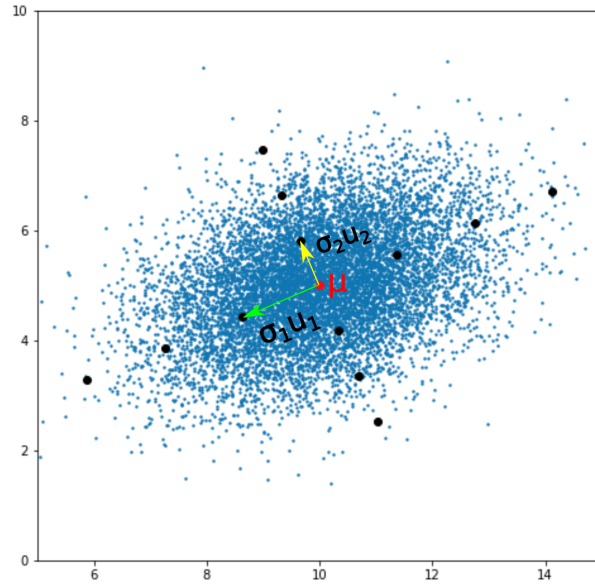


Figure 1: Illustration of the eigenvectors and singular values.

## 2 Standardization

How to standardize a multivariate normal distribution?  $X \sim N(\mu, P) \in \mathbb{R}^n$  is normally distributed with mean  $\mu$ , covariance  $P$ .

$$\boxed{Z = L^{-1}(X - \mu)} \quad (10)$$

Where  $L$  comes from equation (1). It is clear that  $Z$  is normally distributed, since it is a linear transformation of the normal  $X$ . The mean:

$$E(Z) = L^{-1}(E(X) - \mu) = L^{-1}(\mu - \mu) = 0 \quad (11)$$

The covariance:

$$\begin{aligned} P_Z &= \text{Var}(Z) = \text{Var}(L^{-1}(X - \mu)) \\ &= \text{Var}(L^{-1}X) = L^{-1}\text{Var}(X)L^{-T} \\ &= L^{-1}LL^TL^{-T} = I \end{aligned} \quad (12)$$

So we see that  $Z \sim N(0, I)$ , i.e., it is standard normally distributed.

### 3 Mahalanobis distance and Chi square

The Mahalanobis (squared) distance is defined as

$$d_M^2(x; \mu, P) = (x - \mu)^T P^{-1} (x - \mu) \quad (13)$$

where  $\mu, x \in \mathbb{R}^n$ ,  $P \in \mathbb{R}^{n \times n}$  is a symmetric positive-definite matrix (the covariance).

Assuming that  $X$  is normally distributed with  $\mu$  mean, and  $P$  (nonsingular) covariance, how is  $d^2(X; \mu, P)$  distributed?

We know that  $Z^T Z \sim \chi_n^2$ , where  $Z$  comes from the standardization of  $X$ , equation (10).

$$\begin{aligned} Z^T Z &= (X - \mu)^T L^{-T} L^{-1} (X - \mu) \\ &= (X - \mu)^T (LL^T)^{-1} (X - \mu) \\ &= (X - \mu)^T P^{-1} (X - \mu) \\ &= d_M^2(X; \mu, P) \\ &\sim \chi_n^2 \end{aligned} \quad (14)$$

So the Mahalanobis squared distance  $d_M^2(X; \mu, P)$  is Chi-squared distributed with  $n$  degrees of freedom.

## 4 How to visualize Covariance

### 4.1 Points $c^2$ from the mean

How to find those points that are in a specific Mahalanobis distance from the mean ( $\mu$ )?

Construct the following vector ( $\alpha_1, \alpha_2 \in \mathbb{R}$ ):

$$x = \mu + \sum_{i=1}^n \alpha_i \sigma_i u_i \quad (15)$$

The Mahalanobis squared distance is:

$$\begin{aligned} d_M^2(x; \mu, P) &= (x - \mu)^T P^{-1} (x - \mu) \\ &= (x - \mu)^T L^{-T} L^{-1} (x - \mu) \\ &= \|L^{-1}(x - \mu)\|^2 \\ &= \left\| \sum_{i=1}^n \alpha_i \sigma_i L^{-1} u_i \right\|^2 \\ &= \left\| \sum_{i=1}^n \alpha_i \sigma_i \frac{1}{\sigma_i} v_i \right\|^2 \\ &= \left\| \sum_{i=1}^n \alpha_i v_i \right\|^2 \\ &= \sum_{i=1}^n \alpha_i^2 \end{aligned} \quad (16)$$

We used equation (4) to compute  $L^{-1} u_i$ . So if we seek vectors  $x$  for which the Mahalanobis squared distance is  $c^2$ , then we can construct it by picking a point on the  $n$ -dimensional sphere with radius  $c$ , i.e.,

$$\sum_{i=1}^n \alpha_i^2 = c^2 \quad (17)$$

Once we have a specific  $\alpha$ , we can construct  $x$ . For this  $x(\alpha)$ ,  $d_M^2(x(\alpha); \mu, P) = c^2$

### 4.2 How to find $c^2$ ?

Say I want to find a  $c^2$  for which it is 95% probable that  $d_M^2(X; \mu, P) < c^2$ , where  $X \sim N(\mu, P)$ . This is easy, since we know that  $d_M^2(X; \mu, P) \sim \chi_n^2$ .

We need the quantile function (or percent point function) of the Chi-squared distribution.

$$c^2 = Q_{\chi_n^2}(0.95) \quad (18)$$

In 2 dimensions,  $Q_{\chi_n^2}(0.95) = 5.99$ .

### 4.3 Example

In 2 dimensions, we have  $\alpha_1$  and  $\alpha_2$ . With  $t \in (0, 2\pi)$

$$\alpha_1 = c \cdot \sin(t) \quad (19)$$

$$\alpha_2 = c \cdot \cos(t) \quad (20)$$

The following figure shows the ellipse for this scenario.

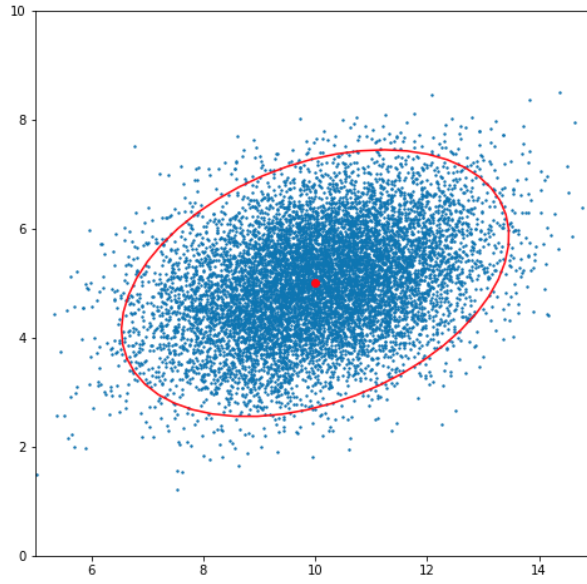


Figure 2: ellipse of the 95% confidence region.