

Chapter 2

Transformations And Expectations

2.1

- (a) $Y = X^3$, $f_X(x) = 42x^5(1-x)$, $0 < x < 1$, $g^{-1}(y) = X^{\frac{1}{3}}$ $f_y = 42x^{\frac{5}{3}}(1-x^{\frac{1}{3}})$
* $\frac{-1}{3}x^{-\frac{4}{3}}$, On integration it comes to 1
- (b) $g^{-1}(y) = \frac{Y-3}{4}$, $f_X(x) = 7e^{-7x} \Rightarrow f_y = 7e^{-7\frac{y-3}{4}}$ $Y = (3 < y < \infty)$
- (c) $Y = X^2$, $f_X(x) = 30x^2(1-x)^2$, $0 < x < 1$ $g^{-1}(y) = X^{\frac{1}{2}}$ $f_y = 20x(1-x^{\frac{1}{2}})$ *
 $\frac{-1}{2}x^{-\frac{3}{2}}$, On integration it comes to 1

2.2

- (a) $Y = X^2 \Rightarrow g^{-1}(x) = x^{\frac{1}{2}}$ $f_y = \frac{1}{2}x^{-\frac{1}{2}}$
- (b) $Y = -\log(X) \Rightarrow g^{-1}(x) = e^{-x} \Rightarrow \frac{(n+m+1)!}{n!m!} e^{-xn}(1-e^{-x})^m * (-e^{-x})$
- (c) $Y = e^x \Rightarrow f_y = \frac{1}{\sigma^2} x^{\frac{-(\log y / \sigma)^2}{2}} * \frac{\log y}{y}$

2.3

$$f_X(x) = \frac{1}{3}(\frac{2}{3})^x \Rightarrow f_X(y) = \frac{1}{3}(\frac{2}{3})^{\frac{y}{1-y}} \quad Y \rightarrow \{0, \frac{1}{2}, \frac{2}{3} \dots\}$$

2.4

$$(a) f(x) = \begin{cases} \frac{1}{2}\lambda e^{-\lambda x} & \text{for } x \geq 0 \\ \frac{1}{2}\lambda e^{\lambda x} & \text{for } x \leq 0 \end{cases}$$

$$\int_{-\infty}^0 \frac{1}{2} \lambda e^{-\lambda x} + \int_0^{\infty} \frac{1}{2} \lambda e^{\lambda x} = \frac{1}{2} + \frac{1}{2} = 1$$

$$(b) \int_{-\infty}^t \frac{1}{2} \lambda e^{-\lambda x} + \int_t^{\infty} \frac{1}{2} \lambda e^{\lambda x} = \frac{t}{2}$$

$$f(x) = \begin{cases} \frac{1}{2} \lambda e^{\lambda t} & \text{for } t \geq 0 \\ 1 - \frac{1}{2} \lambda e^{-\lambda t} & \text{for } t \leq 0 \end{cases}$$

$$(c) P(|X| < t) = 0 \text{ for } t \leq 0 \text{ and } 1 - e^{-\lambda t} \text{ for } t > 0$$

2.5

$$\text{answer} = \frac{1}{\pi} \frac{1}{\sqrt{Y(1-Y)^2}}$$

2.6

$$(a) f_y = \frac{1}{2} e^{-|Y|^{\frac{1}{3}}} \frac{1}{3} Y^{-\frac{2}{3}}$$

$$(b) f_Y(y) = \frac{3}{8} (1-y)^{-\frac{1}{2}} + \frac{3}{8} (1-y)^{\frac{1}{2}}, 0 < y < 1$$

$$(c) f_Y(y) = \frac{3}{16} (1 - (1-y)^{\frac{1}{2}})^2 (1-y)^{-\frac{1}{2}} + \frac{3}{8} (2-y)^2$$

2.7

$$(a), (b) f_y(y) = \begin{cases} \frac{2}{9} \frac{1}{\sqrt{y}} & \text{if } y < 1 \\ \frac{1}{9} + \frac{1}{9} \frac{1}{\sqrt{y}} & \text{if } y \geq 1 \end{cases}$$

2.8

$$(a) F_x^{-1}(y) = -\ln(1-y)$$

$$(b) F_x^{-1}(y) = \begin{cases} \ln 2y & 0 < y < \frac{1}{2} \\ \frac{1}{2} & y = \frac{1}{2} \\ 1 - \ln(2-2y) & 0 < y < \frac{1}{2} \end{cases}$$

$$(c) F_x^{-1}(y) = \begin{cases} \ln 4y & 0 < y \leq \frac{1}{4} \\ -\ln(4-4y) & \frac{3}{4} < y < 1 \end{cases}$$

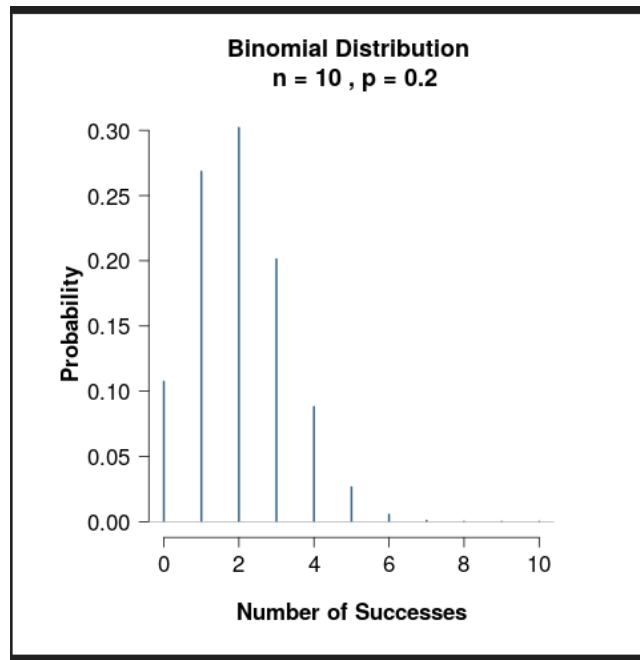


Figure 2.1:

2.9

Logically cdf $F_x(x) \sim \text{Uniform}(0,1)$ $F_x^{-1}(x) = \begin{cases} 0 & -\infty < x \leq 1 \\ \frac{(x-1)^2}{4} & 1 < x < 3 \\ 1 & 3 \leq \infty \end{cases}$

2.10

Refer to figure 2.1

logically $P(Y > y) \geq P([U > y])$

$P(Y < y) \leq P([U < y])$

for some $x = x_0$ $P(Y > y) = P([U > y])$ but at next point ϵ , $P(Y > y)$ will be stagnant since its a discrete probability. But $P([U > y])$ will eventually be lifted some points up. So the two cases are justified

2.11

(a) $f_X(x) = \frac{1}{\sqrt{2\pi}} e^{-x^2/2}$
 $\int_{-\infty}^{\infty} x^2 \frac{1}{\sqrt{2\pi}} e^{-x^2/2} dx$

$$\Rightarrow Y = X^2$$

$$f_y(y) = 2\left(\frac{1}{\sqrt{2\pi}}e^{-\frac{y^2}{2}}\left|\frac{1}{\sqrt{2\pi y}}\right|\right)$$

Integrating gives 1

$$(b)f_X(x) = \frac{1}{\sqrt{2\pi}}e^{-x^2/2} \quad EY = \frac{2}{\pi} \quad EY_2 = 1$$

$$(EY) = 1 - \frac{2}{\pi}$$

2.12

skipping

2.13

$$f_X(X) = p(1-p)^x + (1-p)p^x$$

$$EX = p(1-p)\left[\frac{1}{p^2} + \frac{1}{(1-p)^2}\right]$$

2.14

(doubt) $EX = \int_0^\infty xf_X(x)dx$ Integration by parts

$$EX = [xF_X(x)]_0^\infty - \int_0^\infty F_X(x)dx \quad \text{Now let say the identity is true}$$

$$\int_0^\infty [1 - F_X(x)]dx = \int_0^\infty dx - \int_0^\infty F_X(x)dx$$

So 1st term doesn't make any sense

2.15

$$\int_{-\infty}^\infty xf_1(x)dx + \int_{-\infty}^\infty xf_2(x)dx = \int_{-\infty}^\infty x(f_1(x) + f_2(x))dx = \int_{-\infty}^\infty x(\min(f_1(x), f_2(x)) + \max(f_1(x), f_2(x)))dx = (X \vee Y) + (X \wedge Y) = X + Y$$

Hence by rearranging we get the equation

2.16

$$\int_0^\infty ae^{-\lambda t} + (1-a)e^{-\mu t} = \frac{a}{\lambda} + \frac{1-a}{\mu}$$

2.17

$$(a) m = 2^{\frac{1}{3}}$$

$$(b) m = 0$$

2.18

$E|X-a| = \int_{-\infty}^{\infty} |x-a|f(x) = \int_{-\infty}^a (-x+a)f(x) + \int_a^{\infty} (x-a)f(x)$. On differentiating,
 $\frac{dE|x-a|}{da} = \int_{-\infty}^a f(x) - \int_a^{\infty} f(x) = 0 \Rightarrow \int_{-\infty}^a f(x) = \int_a^{\infty} f(x)$

2.19

$E|X-a|^2 = \int_{-\infty}^{\infty} (x-a)^2 f(x)dx = \frac{dE}{dx} = \int_{-\infty}^{\infty} 2(x-a)f(x)dx \Rightarrow \int_{-\infty}^{\infty} xf(x) = a \int_{-\infty}^{\infty} f(x)dx = a(1) \Rightarrow EX = a$

2.20

$$EX = \sum_{k=0}^{k=\infty} k(1-p)^k p = \frac{1}{p^2}$$

2.21

$$Eg(X) = \int_{-\infty}^{\infty} g(X)f_X(x)dx = \int_{-\infty}^{\infty} yf_X(g^{-1}(y))\left|\frac{dg^{-1}(y)}{dy}\right|dy = \int_{-\infty}^{\infty} yf_y(y)dy = EY$$

2.22

- (a) Integration leads to 1.
 (b) $EX = \frac{2\beta}{\sqrt{\pi}}$, $\text{Var}(X) = \beta^2\left(\frac{3}{2} - \frac{4}{\pi}\right)$

2.23

$$\begin{aligned} f_y(y) &= \frac{1}{2}y^{-\frac{1}{2}} \\ EY &= \frac{1}{3} \\ EY^2 &= \frac{1}{5} \\ \text{Var}(y) &= \frac{4}{45} \end{aligned}$$

2.24

- (a), (b), (c) same as above

2.25

Let $A = \int_0^X f_X(x)dx$ Now $\int_0^{-X} f_X(x)dx \Rightarrow Y = -X \Rightarrow dY = -dX \Rightarrow \int_{-X}^0 f_X(x)dx \Rightarrow -\int_Y^0 f_Y(y)dy = A$
Hence symmetrical

2.26

$M_X(t) = \int_{-\infty}^{\infty} e^{tx} f_X(x)dx$
 $M_X(-t) = \int_{-\infty}^{\infty} e^{-tx} f_X(x)dx \Rightarrow \int_{-\infty}^{\infty} e^{t(-x)} f_X(-x)dx. \text{ Let } j = -x \Rightarrow dj = -dx \Rightarrow \int_{\infty}^{-\infty} e^{tj} f_X(j)dj = M_X(t)$

2.27

- (a) gaussian, $|x|$, x^2
- (b) $\int_{-\infty}^a f_X(x) = \int_a^{\infty} f_X(x) \Rightarrow x = a - y \Rightarrow \int_0^{\infty} f_X(a - y)dy = \int_{-\infty}^a f_X(a - y)dy \Rightarrow a$ is the median
- (c) Wrong sol $\int_{-\infty}^{\infty} x f_X(x)dx = \int_{-\infty}^a x f_X(x)dx + \int_a^{\infty} x f_X(x)dx$
let for first part $x \rightarrow a - x$ and second part $x \rightarrow x - a$
 $-\int_{\infty}^0 ((a-x)f_X(a-x))dx + \int_0^{\infty} ((x-a)f_X(x-a))dx = \int_0^{\infty} ((a-x)f_X(a-x))dx + \int_0^{\infty} ((x-a)f_X(x-a))dx = \int_0^{\infty} ((a-x)f_X(a-x))dx + \int_0^{\infty} ((x-a)f_X(a-x))dx = 0$
symmetric function
- (d) It has a monotonic slope. Hence it cant have a symmetric pdf
- (e) $EX = 1$, median = $\ln 2$

2.28

- (a) Gaussian distribution
- (b) uniform distribution
- (c) Let assume that the point of symmetry is not the modal point. Since the function is symmetric there will be another x for which the mode value will be same. This is in contradiction as our function is unimodal. Hence its symmetric about the mode.
- (d) for $a \geq x \geq y$ where $a = 0$ and $f(0) \geq x \geq y$.

2.29

- (a) Let think logically $x-a$ will shift the a^{th} to zero. Hence symmetrical.
- (b) $\alpha_3 = 2$

- (c) (i) $\alpha_4 = 3$ (ii) $\alpha_4 = \frac{9}{5}$ (iii) $\alpha_4 = 6$

2.30

- (a) $EX^2 = n(n-1)p^2 + np$. $EX = np$ So $E(X(X-1)) = n(n-1)p^2$ (binomial)
 $EX^2 = \lambda^2 + \lambda EX = \lambda E(X(X-1)) = \lambda^2$
 (b) $\text{Var}(\text{binomial}) = n(n-1)p^2 - np^2$ $\text{Var}(\text{binomial}) = \lambda^2 + \lambda - \lambda^2 = \lambda$
 (c) I won't be able to solve

2.31

- (a) $M_X(t) = \frac{e^{ct}-1}{ct}$
 (b) $M_X(t) = \frac{2}{c^2}((c-1)e^c - c + 1)$
 (c) $M_X(t) = \frac{4e^{\alpha t}}{4-\beta^2 t^2}$
 (d) $P(X=x) = \sum_{x=0}^{\infty} \binom{r+x-1}{x} p^r (1-p)^x dx \Rightarrow$
 $M_X(t) = \sum_{x=0}^{\infty} e^{tx} \binom{r+x-1}{x} p^r (1-p)^x dx \Rightarrow \sum_{x=0}^{\infty} e^{tx} \binom{r+x-1}{x} p^r ((1-p)e^t)^x (1 - ((1-p)e^t))^r dx = 1 \Rightarrow M_X(t) = \left(\frac{p}{(1 - ((1-p)e^t))} \right)^r$

2.32

$M_X(0) = 0$, But $M_X(0) = 1$ as the distribution of pmf should be 1. So no

2.33

$\frac{d}{dt} S(t) = \frac{M'_X(t)}{M_X(t)}$ Putting $t=0$ we get $M_X(0) = 1$, $M'_X = EX$, Same go by differentiating divide rule and you will get
 $(M_X^t)^2 \Rightarrow EX^2 - (EX)^2$

2.34

- (a) $M_X(x) \sum_{x=0}^{\infty} \frac{e^{tx} e^{-\lambda} \lambda^x}{x!}$ Using Taylor series $e^{-\lambda} * e^{\lambda e^t}$ Hence we derive the $M_X(t)$. (Rest is Maths) $EX = \lambda$ $EX^2 = \lambda^2 + \lambda$ $\text{Var}(x) = \lambda$
 (b) $M_X(X) \sum_{x=0}^{\infty} p(1-p)^x$ (Think in terms of Geometric Mean.) $M_X(t) = \frac{p}{1-(1-p)e^t}$
 $EX = \frac{1-p}{p}$, $EX^2 = \frac{p(1-p)+2(1-p)^2}{p^2}$ $\text{Var}(x) = \frac{1-p}{p^2}$
 (c) Yes it's an mgf (Some complicated mathematics) $EX = \mu$ $EX^2 = \mu^2 + \sigma^2$.

$$\text{Var}(x) = \sigma^2$$

2.35

$$\begin{aligned} \text{(a)} \quad & \int_0^\infty x^r * \frac{1}{\sqrt{2\pi x}} * e^{-(\log x)^2/2} = e^{r^2/2} \\ \text{(b)} \quad & e^{r^2/2 - 2\pi^2} \end{aligned}$$

2.36

$$\begin{aligned} & \int_0^\infty \frac{e^{tx}}{\sqrt{2\pi x}} e^{-(\log x)^2/2} dx \\ &= \lim_{x \rightarrow \infty} e^{tx - \log^2(x)} \\ &= \lim_{x \rightarrow \infty} \frac{e^{tx}}{e^{\log^2(x)}} \text{ Taking log} \\ &= \lim_{x \rightarrow \infty} \frac{tx}{\log^2(x)} \text{ On solving } \lim_{x \rightarrow \infty} tx/2 \rightarrow \infty \end{aligned}$$

2.37

$$\begin{aligned} \text{(a)} \quad & \\ \text{(b)} \quad & \\ \text{(c)} \quad & \text{(i) } M_X(t) = e^{K_1} \\ & \text{ii } M_X(t) = e^{K_2} \end{aligned}$$

$$\text{(d) make transformation } e^x$$

2.38

$$\begin{aligned} \text{(a)} \quad & M_X(t) = [p/(1 - (1 - p)e^t)]^r \\ \text{(b)} \quad & M_X(t) = [p/(1 - (1 - p)e^2 pt)]^r \text{ On solving } [\frac{1}{1-2t}]^r \end{aligned}$$

2.39

$$\begin{aligned} \text{(a)} \quad & e^{-\lambda x} \\ \text{(b)} \quad & \frac{-1}{\lambda^2} \\ \text{(c)} \quad & \frac{1}{t^2} \\ \text{(d)} \quad & \frac{1}{(1-t)^2} \end{aligned}$$

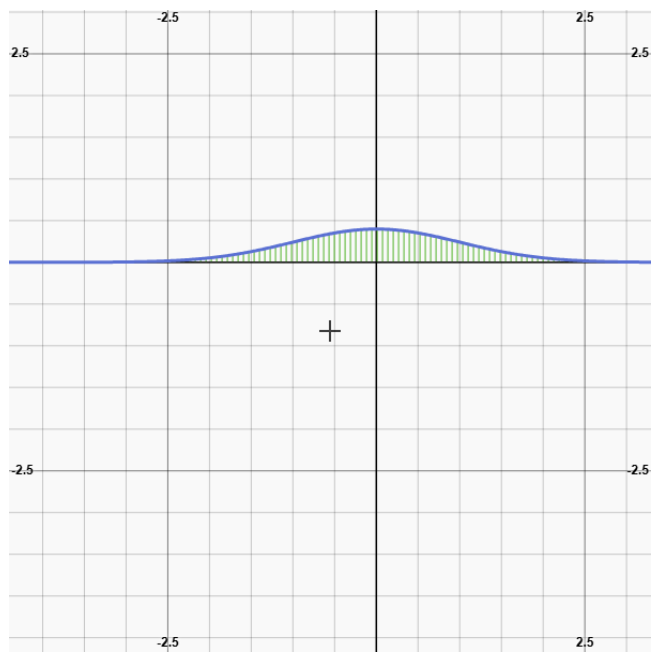


Figure 2.2:

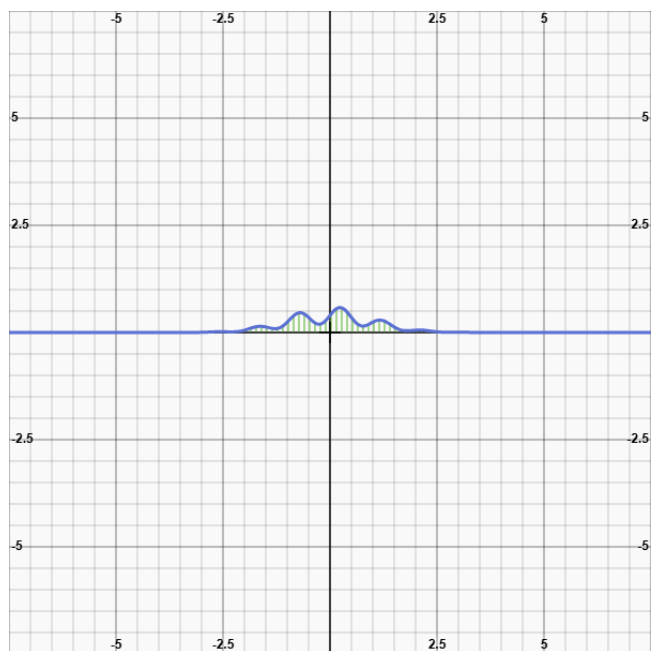


Figure 2.3:

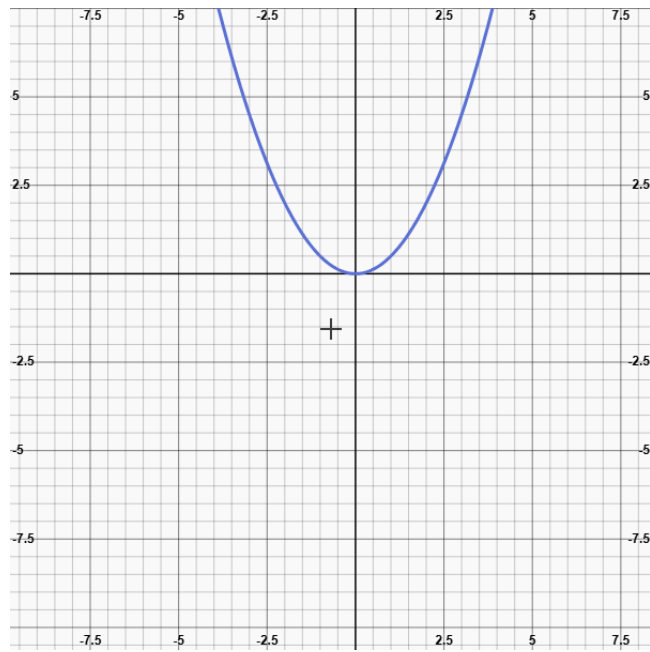


Figure 2.4:

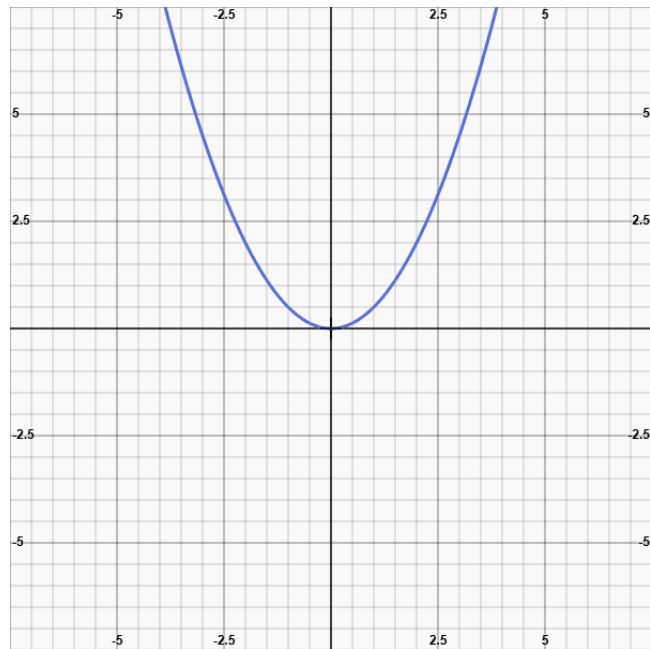


Figure 2.5: