

Logistic regression

- Classification setup
- Logistic regression is the classification counterpart of linear regression.
 - we focus on binary classification, i.e. we assume $X \in \mathbb{R}^d$, $Y = \{\text{yes}, \text{no}\}$

For example,

- the two classes can be associated with whether an individual will default on their credit card payments.

- The attributes may be the person's annual income & monthly credit card balance.
 - ($d = 2$ in this case).

Probability prediction

- The default labels fall into one of the two categories, i.e. $Y \in \{\text{yes}, \text{no}\}$
- Idea: rather than modelling Y directly, logistic regression models $p(x) = \text{Prob}(Y = \text{yes} | X)$.
i.e. the probability that $Y \in \text{yes}$ (given x)
- Let $\tilde{X} = [1, X]^T$ with $X \in \mathbb{R}^d$, then

$$F(x, \lambda) = p_\lambda(x) = \sigma(\tilde{X}^T \lambda) = \sigma\left(\lambda_0 + \sum_{i=1}^d \lambda_i X_i\right)$$

Where; $\sigma: \mathbb{R} \rightarrow [0, 1]$ is the logistic function.

The sigmoid function

• The logistic function

$$\sigma(x) = \frac{1}{1+e^{-x}} = \frac{e^x}{1+e^x}$$

- is nonnegative and bounded.
- is monotone

i.e. - its derivative $\sigma'(x) = \sigma(x)(1-\sigma(x))$

- is positive everywhere

- maps from the real line to the interval $[0,1]$ & it can be naturally interpreted as a probability

$$\lim_{x \rightarrow \infty} \sigma(x) = 1 \quad \text{and} \quad \lim_{x \rightarrow -\infty} \sigma(x) = 0$$

Odds

- The odds of the conditional probability $\text{Prob}(Y = \text{yes} | X) = p_\lambda(X)$ are defined as

$$\frac{p_\lambda(X)}{1 - p_\lambda(X)} = e^{\tilde{x}^T \lambda}$$

and their logarithms,

i.e. the log-odds of $p_\lambda(X)$, are linear in the attribute
i.e.

$$\log \frac{p(X)}{1 - p(X)} = \tilde{x}^T \lambda$$

- Note that: $1 - p_\lambda(X) = \text{Prob}(Y = \text{no} | X)$

Maximum likelihood estimation

- Let $\mathcal{D} = \{(x_n, y_n) \in \mathbb{R}^d \times \{\text{yes}, \text{no}\}\}_{n=1}^N$ be a training data set.
- The interpretation of $F(X, \lambda) = \sigma(\tilde{x}^T \lambda) \in [0, 1]$
 - as the conditional probability of observing $Y = \text{yes}$ given X
 - allows you to estimate λ by
 - maximizing the likelihood of \mathcal{D} .

The ML estimate on \mathcal{D} is the parameter that maximises the probability of observing \mathcal{D} .

Conditional likelihood maximisation

- As the attributes are assumed to be always known,
 - we focus on maximizing the probability of observing the labels $y_1, \dots, y_N \in \mathcal{D}$
- More precisely, we let

$$\hat{\lambda} = \arg \max_{\lambda} \mathcal{L}(\mathcal{D}, \lambda)$$

where:

$$\mathcal{L}(\mathcal{D}, \lambda) = \text{Prob}(y_1, \dots, y_N | x_1, \dots, x_N, \lambda)$$

$$= \prod_{i=1}^N \mathbb{1}[y_i = \text{yes}] F(x_i, \lambda) + \mathbb{1}[y_i = \text{no}] (1 - F(x_i, \lambda))$$

Log-likelihood minimization

- An numerically easier but equivalent problem is to minimize the negative of the logarithm of $L(D, \lambda)$,
i.e. to let

$$\hat{\lambda} = \arg \min_{\lambda \in \mathbb{R}^{d+1}} J(D, \lambda) = -\log(L(D, \lambda))$$

$$\hat{\lambda} = -\log \left(\prod_{i=1}^N 1[y_i = \text{yes}] F(x_i, \lambda) + 1[y_i = \text{no}] (1 - F(x_i, \lambda)) \right)$$

- In particular, using standard property of the logarithm we have

$$J(D, \lambda) = \sum_{i=1}^N \log(y_i F(x_i, \lambda) + (1 - y_i)(1 - F(x_i, \lambda)))$$

where;

- the representation numerical of the nominal variables,
i.e. {yes, no} \rightarrow {1, 0}

- allows us to rewrite the i^{th} factor in $L(D, \lambda)$ as

$$y_i F(x_i, \lambda) + (1 - y_i)(1 - F(x_i, \lambda))$$

- which is differentiable on λ .

Example :