

The IPython Notebook: open, reproducible scientific computing

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Abstract

While computing has become a foundation of all it is challenging for researchers .

1 Introduction

2 The lifecycle of research

3 The IPython Notebook

3.1 Web application

3.2 Notebook document format

3.3 Installation

4 Collaboration

5 Broader ecosystem

6 Future directions

Dexy Snippets

The notebook `notebooks/Chapter1_Introduction.ipynb` uses format 3.0.

Here is an example of a markdown cell:

3What is the mean of λ_1 **given** we know τ is less than 45. That is, suppose we have new information as we know for certain that the change in behaviour occurred before day 45. What is the expected value of λ_1 now? (You do not need to redo the PyMC part, just consider all instances where ‘ $\tau_{samples} < 45$ ’.)

Here is the same cell in a Verbatim block:

3\. What is the mean of λ_1 **given** we know τ is less than 45. That is,

Here is the same cell presented as highlighted markup:

3\. What **is** the mean of λ_1 **given** we know τ **is** less than **45**. That **is**,

Here are the worksheets in the notebook:

- notebooks/Chapter1_Introduction.ipynb--ws-0

This worksheet has 47 cells and 59 documents based on these cells.

- notebooks/Chapter1_Introduction.ipynb--ws-0-cell-0.md

Here are the contents of notebooks/Chapter1_Introduction.ipynb--ws-0-cell-0.md:

Probabilistic Programming

=====

and Bayesian Methods **for** Hackers

=====

#####Version **0.1**

Welcome to *Bayesian Methods **for** Hackers*. The full Github repository, and addi

- notebooks/Chapter1_Introduction.ipynb--ws-0-cell-1.md

Here are the contents of notebooks/Chapter1_Introduction.ipynb--ws-0-cell-1.md:

Chapter **1**

=====

- notebooks/Chapter1_Introduction.ipynb--ws-0-cell-2.md

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The Philosophy of Bayesian Inference

> You are a skilled programmer, but bugs still slip into your code. After a par

If you think **this** way, then congratulations, you already are a Bayesian practit

- notebooks/Chapter1_Introduction.ipynb--ws-0-cell-3.md

Here are the contents of notebooks/Chapter1_Introduction.ipynb--ws-0-cell-3.md:

###The Bayesian state of mind

Bayesian inference differs from more traditional statistical inference by prese

The Bayesian world-view interprets probability as measure of *believability in

For this to be clearer, we consider an alternative interpretation of probabilit

Bayesians, on the other hand, have a more intuitive approach. Bayesians interpr

Notice in the paragraph above, I assigned the belief (probability) measure to a

- I flip a coin, and we both guess the result. We would both agree, assuming th
- Your code either has a bug in it or not, but we do not know for certain whic
- A medical patient is exhibiting symptoms x , y and z . There are a numbe

This philosophy of treating beliefs as probability is natural to humans. We emp

To align ourselves with traditional probability notation, we denote our belief

John Maynard Keynes, a great economist and thinker, said "When the facts change

- 1\ . $P(A)$: \;\;\\$ the coin has a 50 percent chance of being heads. $P(A \mid X)$: \;\;\
- 2\ . $P(A)$: \;\;\\$ This big, complex code likely has a bug in it. $P(A \mid X)$: \;
- 3\ . $P(A)$: \;\;\\$ The patient could have any number of diseases. $P(A \mid X)$: \;\;\\$

It's clear that in each example we did not completely discard the prior belief

By introducing prior uncertainty about events, we are already admitting that an

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###Bayesian Inference in Practice

If frequentist and Bayesian inference were programming functions, with inputs

For example, in our debugging problem above, calling the frequentist function w

```
> *YES*, with probability 0.8; *NO*, with probability 0.2
```

This `is` very different from the answer the frequentist `function` returned. Notice

```
####Incorporating evidence
```

As we acquire more and more instances of evidence, our prior belief `is` *washed

Denote N as the number of instances of evidence we possess. As we gather an *

One may think that `for` large N , one can be indifferent between the two techni

```
> Sample sizes are never large. If  $N$  is too small to get a sufficiently-prec
```

```
### Are frequentist methods incorrect then?
```

```
**No.**
```

Frequentist methods are still useful or state-of-the-art `in` many areas. Tools 1

```
#### A note on *Big Data*
```

Paradoxically, big data's predictive analytic problems are actually solved by r

The much more difficult analytic problems involve *medium data* and, especially

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```
### Our Bayesian framework
```

We are interested `in` beliefs, which can be interpreted `as` probabilities by thin

Secondly, we observe our evidence. To `continue` our buggy-code example: `if` our c

```
\begin{align}
```

```
P( A | X ) = & \frac{ P(X | A) P(A) } { P(X) } \\\[5pt]
```

```
& \propto P(X | A) P(A)\;; (\propto \text{is proportional to } )
```

```
\end{align}
```

The above formula `is` not unique to Bayesian inference: it `is` a mathematical fac

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Example: Mandatory coin-flip example

Every statistics text must contain a coin-flipping example, I'll use it here to

We begin to flip a coin, and record the observations: either $\$H\$$ or $\$T\$$. This i

Below we plot a sequence of updating posterior probabilities as we observe incr

– notebooks/Chapter1_Introduction.ipynb--ws-0-cell-7-input.py

Here are the contents of notebooks/Chapter1_Introduction.ipynb--ws-0-cell-7-input.py:

```
"""
```

The book uses a custom matplotlibrc file, which provides the unique styles for
If executing this book, and you wish to use the book's styling, provided are tw

1. Overwrite your own matplotlibrc file with the rc-file provided in the bo
See <http://matplotlib.org/users/customizing.html>
2. Also in the styles is bmh_matplotlibrc.json file. This can be used to u
in only this notebook. Try running the following code:

```
import json
s = json.load( open("../styles/bmh_matplotlibrc.json") )
matplotlib.rcParams.update(s)
```

```
"""
```

#the code below can be passed over, as it is currently not important.

```
%pylab inline
```

```
figsize( 11, 9)
```

```
import scipy.stats as stats
```

```
dist = stats.beta
```

```
n_trials = [0,1,2,3,4,5,8,15, 50, 500]
```

```
data = stats.bernoulli.rvs(0.5, size = n_trials[-1] )
```

```
x = np.linspace(0,1,100)
```

```
for k, N in enumerate(n_trials):
```

```
    sx = subplot( len(n_trials)/2, 2, k+1)
```

```
    plt.xlabel("$p$, probability of heads") if k in [0,len(n_trials)-1] else No
```

```
    plt.setp(sx.get_yticklabels(), visible=False)
```

```
    heads = data[:N].sum()
```

```
    y = dist.pdf(x, 1 + heads, 1 + N - heads )
```

```
    plt.plot( x, y, label= "observe %d tosses,\n %d heads"%(N,heads) )
```

```
    plt.fill_between( x, 0, y, color="#348ABD", alpha = 0.4 )
```

```

plt.vlines( 0.5, 0, 4, color = "k", linestyle = "--", lw=1 )

leg = plt.legend()
leg.get_frame().set_alpha(0.4)
plt.autoscale(tight = True)

plt.suptitle( "Bayesian updating of posterior probabilities",
              y = 1.02,
              fontsize = 14);

plt.tight_layout()

```

– notebooks/Chapter1_Introduction.ipynb--ws-0-cell-7-output-0.txt
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– notebooks/Chapter1_Introduction.ipynb--ws-0-cell-8.md
Here are the contents of notebooks/Chapter1_Introduction.ipynb--ws-0-cell-8.md:
The posterior probabilities are represented by the curves, and our confidence interval is

Notice that the plots are not always *peaked* at 0.5. There is no reason it should be.

The next example is a simple demonstration of the mathematics of Bayesian inference.

– notebooks/Chapter1_Introduction.ipynb--ws-0-cell-9.md
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#####Example: Bug, or just sweet, unintended feature?

Let A denote the event that our code has *no bugs* in it. Let X denote the event that the code passes our debug tests.

We are interested in $P(A|X)$, i.e. the probability of no bugs, given our debug tests.

What is $P(X | A)$, i.e., the probability that the code passes X tests *given* that it has no bugs?

$P(X)$ is a little bit trickier: The event X can be divided into two possibilities: $X = A \cup \bar{A}$.

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$$\begin{aligned}
P(X) &= P(X \text{ and } A) + P(X \text{ and } \bar{A}) \\
&= P(X|A)P(A) + P(X | \bar{A})P(\bar{A}) \\
&= P(X|A)p + P(X | \bar{A})(1-p)
\end{aligned}$$

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We have already computed $P(X|A)$ above. On the other hand, $P(X | \sim A)$ is

```
\begin{align}
P(A | X) &= \frac{1 \cdot p}{1 \cdot p + 0.5 (1-p)} \\
&= \frac{2 p}{1+p}
\end{align}
```

This is the posterior probability. What does it look like as a function of our

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```
figsize(12.5,4)
p = np.linspace( 0,1, 50)
plt.plot( p, 2*p/(1+p), color = "#348ABD", lw = 3 )
#plt.fill_between( p, 2*p/(1+p), alpha = .5, facecolor = ["#A60628"])
plt.scatter( 0.2, 2*(0.2)/1.2, s = 140, c = "#348ABD" )
plt.xlim( 0, 1)
plt.ylim( 0, 1)
plt.xlabel( "Prior,  $P(A) = p$ " )
plt.ylabel( "Posterior,  $P(A|X)$ , with  $P(A) = p$ " )
plt.title( "Are there bugs in my code?" );
```

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– notebooks/Chapter1_Introduction.ipynb--ws-0-cell-13.md

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We can see the biggest gains if we observe the X tests passed when the prior

Recall that the prior is a probability: p is the prior probability that there

Similarly, our posterior is also a probability, with $P(A | X)$ the probability

– notebooks/Chapter1_Introduction.ipynb--ws-0-cell-14-input.py

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```
figsize( 12.5, 4 )
colours = ["#348ABD", "#A60628"]

prior = [0.20, 0.80]
posterior = [1./3, 2./3]
plt.bar( [0,.7], prior ,alpha = 0.70, width = 0.25, \
        color = colours[0], label = "prior distribution",
        lw = "3", edgecolor = colours[0])
```

```

plt.bar( [0+0.25,.7+0.25], posterior ,alpha = 0.7, \
         width = 0.25, color = colours[1],
         label = "posterior distribution",
         lw = "3", edgecolor = colours[1])

plt.xticks( [0.20,.95], ["Bugs Absent", "Bugs Present"] )
plt.title("Prior and Posterior probability of bugs present, prior = 0.2")
plt.ylabel("Probability")
plt.legend(loc="upper left");

```

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– notebooks/Chapter1_Introduction.ipynb--ws-0-cell-15.md
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Notice that after we observed X occur, the probability of bugs being absent is

This was a very simple example of Bayesian inference and Bayes rule. Unfortunately

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Probability Distributions

Let's quickly recall what a probability distribution **is**:

Let Z be some random variable.

We can divide random variables into three classifications:

- **Z is discrete**: Discrete random variables may only assume values on a finite or countable set.
- **Z is continuous**: Continuous random variable can take on arbitrarily many values.
- **Z is mixed**: Mixed random variables assign probabilities to both discrete and continuous values.

Discrete Case

If Z is discrete, then its distribution is called a **probability mass function**.

$$P(Z = k) = \frac{\lambda^k e^{-\lambda}}{k!}, \quad k=0,1,2, \dots$$

What is λ ? It is called the parameter, and it describes the shape of the distribution.

Unlike λ , which can be any positive number, the value k in the above

If a random variable Z has a Poisson mass distribution, we denote **this** by writing

$$Z \sim \text{Poi}(\lambda)$$

One very useful property of the Poisson random variable, given we know λ

$$E[Z] = \lambda$$

We will use **this** property often, so it's something useful to remember. Below we

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```
figsize( 12.5, 4)
```

```
import scipy.stats as stats
```

```
a = np.arange( 16 )
```

```
poi = stats.poisson
```

```
lambda_ = [1.5, 4.25 ]
```

```
plt.bar( a, poi.pmf( a, lambda_[0]), color=colours[0],
        label = "$\lambda = %.1f$" % lambda_[0], alpha = 0.60,
        edgecolor = colours[0], lw = "3")
```

```
plt.bar( a, poi.pmf( a, lambda_[1]), color=colours[1],
        label = "$\lambda = %.1f$" % lambda_[1], alpha = 0.60,
        edgecolor = colours[1], lw = "3")
```

```
plt.xticks( a + 0.4, a )
```

```
plt.legend()
```

```
plt.ylabel("probability of  $k$ ")
```

```
plt.xlabel(" $k$ ")
```

```
plt.title("Probability mass function of a Poisson random variable; differing \
 $\lambda$  values");
```

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###Continuous Case

Instead of a probability mass **function**, a continuous random variable has a **probability density function**

$$f_Z(z | \lambda) = \lambda e^{-\lambda} \frac{\lambda^z}{z!}, \quad z \geq 0$$

Like the Poisson random variable, an exponential random variable can only take

When a random variable Z has an exponential distribution with parameter λ

$Z \sim \text{Exp}(\lambda)$

Given a specific λ , the expected value of an exponential random variable

$E[Z] = \frac{1}{\lambda}$

– notebooks/Chapter1_Introduction.ipynb--ws-0-cell-19-input.py

Here are the contents of notebooks/Chapter1_Introduction.ipynb--ws-0-cell-19-input.py

```
a = np.linspace(0,4, 100)
expo = stats.expon
lambda_ = [0.5, 1]

for l,c in zip(lambda_,colours):
    plt.plot( a, expo.pdf( a, scale=1./l), lw=3,
              color=c, label = "$\lambda = %.1f"%l)
    plt.fill_between( a, expo.pdf( a, scale=1./l), color=c, alpha = .33)

plt.legend()
plt.ylabel("PDF at $z$")
plt.xlabel("$z$")
plt.title("Probability density function of an Exponential random variable;\ndiffering $\lambda$");
```

– notebooks/Chapter1_Introduction.ipynb--ws-0-cell-19-output-0.png

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– notebooks/Chapter1_Introduction.ipynb--ws-0-cell-20.md

Here are the contents of notebooks/Chapter1_Introduction.ipynb--ws-0-cell-20.md:

###But what is λ ?

****This question is what motivates statistics**.** In the real world, λ is

Bayesian inference is concerned with *beliefs* about what λ is. Rather

This might seem odd at first: after all, λ is fixed, it is not (necessa

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Example: Inferring behaviour from text-message data

Let's try to model a more interesting example, concerning text-message rates:

> You are given a series of text-message counts from a user of your system. Th

- notebooks/Chapter1_Introduction.ipynb--ws-0-cell-22-input.py
Here are the contents of notebooks/Chapter1_Introduction.ipynb--ws-0-cell-22-input.py
- ```

figsize(12.5, 3.5)
count_data = np.loadtxt("data/txtdata.csv")
n_count_data = len(count_data)
plt.bar(np.arange(n_count_data), count_data, color = "#348ABD")
plt.xlabel("Time (days)")
plt.ylabel("count of text-msgs received")
plt.title("Did the user's texting habits change over time?")
plt.xlim(0, n_count_data);

```
- notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-22-output-0.png  
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  - notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-23.md  
Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-23.md:
- Before we begin, with respect to the plot above, would you say there was a change during the time period?

How can we start to model this? Well, as I conveniently already introduced, a Poisson process

\$\$ C\_i \sim \text{Poisson}(\lambda) \$\$

We are not sure about what the  $\lambda$  parameter is though. Looking at the change in the data

How can we mathematically represent this? We can think, that at some later date

\$\$

$$\lambda = \begin{cases} \lambda_1 & \text{if } t < \tau \\ \lambda_2 & \text{if } t \geq \tau \end{cases}$$

\$\$

If, in reality, no sudden change occurred and indeed  $\lambda_1 = \lambda_2$ ,

We are interested in inferring the unknown  $\lambda$ s. To use Bayesian inference

$$\begin{aligned} & \lambda_1 \sim \text{Exp}(\alpha) \\ & \lambda_2 \sim \text{Exp}(\alpha) \end{aligned}$$

$\alpha$  is called a \*hyper-parameter\*, or a \*parent-variable\*, literally a par

$$\frac{1}{N} \sum_{i=0}^N C_i \approx E[\lambda; \alpha] = \frac{1}{70}$$

Alternatively, and something I encourage the reader to try, is to have two priors

What about  $\tau$ ? Well, due to the randomness, it is too difficult to pick out

```
\begin{align}
& \tau \sim \text{DiscreteUniform}(1,70) \\
& \Rightarrow P(\tau = k) = \frac{1}{70}
\end{align}
```

So after all this, what does our overall prior for the unknown variables look like

Introducing our first hammer: PyMC  
-----

PyMC is a Python library for programming Bayesian analysis [3]. It is a fast, w

We will model the above problem using the PyMC library. This type of programmin

B. Cronin [5] has a very motivating description of probabilistic programming:

> Another way of thinking about this: unlike a traditional program, which onl

Due to its poorly understood title, I'll refrain from using the name \*probabili

The PyMC code is easy to follow along: the only novel thing should be the synta

– notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-24-input.py

Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-24-input.py

```
import pymc as mc
```

```
alpha = 1.0/count_data.mean() #recall count_data is
 #the variable that holds our txt counts
```

```
lambda_1 = mc.Exponential("lambda_1", alpha)
lambda_2 = mc.Exponential("lambda_2", alpha)
```

```
tau = mc.DiscreteUniform("tau", lower = 0, upper = n_count_data)
```

– notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-25.md

- Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-25.md:
- In the above code, we create the PyMC variables corresponding to  $\lambda_1$ ,  $\lambda_2$ ;
- notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-26-input.py  
Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-26-input.py:  

```
print "Random output:", tau.random(),tau.random(), tau.random()
```
  - notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-26-output-0.txt  
Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-26-output-0.txt:
  - notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-27-input.py  
Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-27-input.py:  

```
@mc.deterministic
def lambda_(tau = tau, lambda_1 = lambda_1, lambda_2 = lambda_2):
 out = np.zeros(n_count_data)
 out[:tau] = lambda_1 #lambda before tau is lambda1
 out[tau:] = lambda_2 #lambda after tau is lambda2
 return out
```
  - notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-28.md  
Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-28.md:  
This code is creating a new function `lambda_`, but really we think of it as a
  - notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-29-input.py  
Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-29-input.py:  

```
observation = mc.Poisson("obs", lambda_, value = count_data, observed = True)

model = mc.Model([observation, lambda_1, lambda_2, tau])
```
  - notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-30.md  
Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-30.md:  
The variable `observation` combines our data, `count_data`, with our proposed distribution.
  - notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-31-input.py  
Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-31-input.py:  

```
Mysterious code to be explained in Chapter 3.
mcmc = mc.MCMC(model)
mcmc.sample(40000, 10000, 1)
```
  - notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-31-output-0.txt  
Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-31-output-0.txt:
  - notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-31-output-1.txt  
Here are the contents of notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-31-output-1.txt:

```

- notebooks/Chapter1_Introduction.ipynb--ws-0-cell-32-input.py
Here are the contents of notebooks/Chapter1_Introduction.ipynb--ws-0-cell-32-input.py

lambda_1_samples = mcmc.trace('lambda_1')[:]
lambda_2_samples = mcmc.trace('lambda_2')[:]
tau_samples = mcmc.trace('tau')[:]

- notebooks/Chapter1_Introduction.ipynb--ws-0-cell-33-input.py
Here are the contents of notebooks/Chapter1_Introduction.ipynb--ws-0-cell-33-input.py

figsize(12.5, 10)
#histogram of the samples:

ax = plt.subplot(311)
ax.set_autoscaley_on(False)

plt.hist(lambda_1_samples, histtype='stepfilled', bins = 30, alpha = 0.85,
 label = "posterior of λ_1 ", color = "#A60628", normed = True)
plt.legend(loc = "upper left")
plt.title(r"Posterior distributions of the variables $\lambda_1, \lambda_2, \tau$ ")
plt.xlim([15,30])
plt.xlabel(" λ_2 value")
plt.ylabel("probability")

ax = plt.subplot(312)
ax.set_autoscaley_on(False)

plt.hist(lambda_2_samples, histtype='stepfilled', bins = 30, alpha = 0.85,
 label = "posterior of λ_2 ", color="#7A68A6", normed = True)
plt.legend(loc = "upper left")
plt.xlim([15,30])
plt.xlabel(" λ_2 value")
plt.ylabel("probability")

plt.subplot(313)

w = 1.0/ tau_samples.shape[0] * np.ones_like(tau_samples)
plt.hist(tau_samples, bins = n_count_data, alpha = 1,
 label = r"posterior of τ ",
 color="#467821", weights=w, rwidth =2.)
plt.xticks(np.arange(n_count_data))

plt.legend(loc = "upper left");
plt.ylim([0,.75])
plt.xlim([35, len(count_data)-20])

```

```
plt.xlabel("τ (in days)")
plt.ylabel("probability");
```

– notebooks/Chapter1\_Introduction.ipynb--ws-0-cell-33-output-0.png

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### Interpretation

Recall that the Bayesian methodology returns a *distribution*, hence we now have

Also notice that the posterior distributions for the  $\lambda$ 's do not look like

Our analysis also returned a distribution for what  $\tau$  might be. Its posterior

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### Why would I want samples from the posterior, anyways?

We will deal with this question for the remainder of the book, and it is an und

In the code below, we are calculating the following: Let  $i$  index samples from

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```
figsize(12.5, 5)
tau_samples, lambda_1_samples, lambda_2_samples contain
N samples from the corresponding posterior distribution
N = tau_samples.shape[0]
expected_texts_per_day = np.zeros(n_count_data)
for day in range(0, n_count_data):
 # ix is a bool index of all tau samples corresponding to
 # the switchpoint occurring prior to value of 'day'
 ix = day < tau_samples
 # Each posterior sample corresponds to a value for tau.
 # for each day, that value of tau indicates whether we're "before"
 # (in the lambda1 "regime") or
 # "after" (in the lambda2 "regime") the switchpoint.
 # by taking the posterior sample of lambda1/2 accordingly, we can average
 # over all samples to get an expected value for lambda on that day.
 # As explained, the "message count" random variable is Poisson distributed,
 # and therefore lambda (the poisson parameter) is the expected value of "me
 expected_texts_per_day[day] = (lambda_1_samples[ix].sum()
 + lambda_2_samples[~ix].sum()) /N
```

```
plt.plot(range(n_count_data), expected_texts_per_day, lw =4, color = "#E24A33",
 label = "expected number of text-messages recieved")
plt.xlim(0, n_count_data)
plt.xlabel("Day")
plt.ylabel("Expected # text-messages")
plt.title("Expected number of text-messages received")
plt.ylim(0, 50)
plt.bar(np.arange(len(count_data)), count_data, color = "#348ABD", alpha = 0.5,
 label="observed texts per day")
```

```
plt.legend(loc="upper left");
```

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Our analysis shows strong support for believing the user's behavior did change

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##### Exercises

1\ . Using 'lambda\_1\_samples' and 'lambda\_2\_samples', what is the mean of the p

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2\ . What is the expected percentage increase in text-message rates? 'hint: co

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#type your code here.

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3\ . What is the mean of  $\lambda_1$  \*\*given\*\* we know  $\tau$  is less than 45.

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#type your code here.



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```
from IPython.core.display import HTML
def css_styling():
 styles = open("../styles/custom.css", "r").read()
 return HTML(styles)
css_styling()
```

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