Quantifying dynamics of SARS-CoV-2 transmission suggests that epidemic control and avoidance is feasible through instantaneous digital contact tracing: Supplementary Information

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Supplementary Text

Inference of the distribution of generation times

We collected data on 40 source-recipient pairs with known onset of symptoms from public sources. This set of transmission events spans different countries (see Supplementary Figure 4) and social relations (households, family members, colleagues, acquaintances, random contacts). For each transmission event $1 \to 2$, we include the dates of onset of symptoms $t_{s,1}$ and $t_{s,2}$, the intervals of exposure $[e_{1,L}, e_{1,R}]$ and $[e_{2,L}, e_{2,R}]$, and the reporting date T_r . The intervals of exposure must satisfy the conditions $e_{1,R} \le \min(t_{s,1}, e_{2,R})$, $e_{2,L} \ge e_{1,L}$ and $e_{2,R} \le t_{s,2}$.

First, we discretize the incubation and generation time distributions

$$I(j) = \int_{j-0.5}^{j+0.5} i(t) dt, \qquad \Omega(j|\Theta_{\omega}) = \int_{j-0.5}^{j+0.5} \omega(t) dt$$
 (1)

where the incubation time distribution i(t) is taken from [Lauer et al 2020], and Θ_{ω} denotes the parameters of the generation time distribution $\omega(t)$. Then, for a given transmission event $1 \to 2$, we define the likelihood of observing the transmission and the subsequent time of onset of symptoms in the recipient as

$$\mathcal{L}_{\text{trans}}(\Theta_{\omega}) = P[1 \to 2, t_{s,2} | t_{s,1}, \Theta_{\omega}] \tag{2}$$

which can be obtained by summing over all possible infection times $t_{i,1}$ and $t_{i,2}$, resulting in the explicit form

$$\mathcal{L}_{\text{trans}}(\Theta_{\omega}) \propto \sum_{t_{i,1}=e_{1,L}}^{e_{1,R}} e^{-r(T_r - t_{i,1})} I(t_{s,1} - t_{i,1}) \sum_{t_{i,2}=e_{2,L}}^{e_{2,R}} \Omega(t_{i,2} - t_{i,1}|\Theta_{\omega}) I(t_{s,2} - t_{i,2})$$
(3)

up to a multiplicative term independent of $\omega(t)$ which has no effect on likelihood maximisation. The growth rate r corresponds to a doubling time of 5 days for exponentially growing epidemics (China except Hong Kong, Italy, South Korea) while it is assumed to be 0 for countries where local transmission is limited.

For simplicity, we ignore correlations between transmission events. We define the composite likelihood across all 40 pairs (i.e. the approximate global likelihood that neglects correlations) as the product of the likelihoods $\mathcal{L}_{trans}(\Theta_{\omega})$ of all pairs. We then test multiple functional forms for $\omega(t)$ (lognormal, gamma, Weibull), inferring their parameters $\hat{\Theta}_{\omega}$ by Maximum Composite Likelihood (MCL). Finally, we select the functional form based on the lowest AIC. Confidence intervals for the parameters of the distribution are based on likelihood ratios, while confidence intervals for the values of the curve are based on likelihood profiling.

This approach relies not only on the timing of transmissions and onset of symptoms, but it also assumes that transmission pairs themselves are chosen at random from the epidemic. This is not generally the case. To test the robustness of our approach, we relax this assumption by considering also the probability conditional on the occurrence of the transmission $1 \to 2$ at any time:

$$\mathcal{L}_{|\text{trans}}(\Theta_{\omega}) = P[t_{s,2}|1 \to 2, t_{s,1}, \Theta_{\omega}] = \frac{P[1 \to 2, t_{s,2}|t_{s,1}, \Theta_{\omega}]}{P[1 \to 2|t_{s,1}, \Theta_{\omega}]} = \frac{\mathcal{L}_{\text{trans}}(\Theta_{\omega})}{\sum_{t'_{2,s}} P[1 \to 2, t'_{2,s}|t_{s,1}, \Theta_{\omega}]}$$
(4)

This approach does not extract any information from the fact that a transmission actually occurred. The MCL inference uses only the timing of exposure and onset of symptoms. The correspondence between the two approaches is therefore a good test of robustness.

Posterior probability of pre-symptomatic transmission

For each transmission event, the likelihood $\mathcal{L}_{trans}(\Theta_{\omega})$ described above can be decomposed as

$$\mathcal{L}_{trans} = \mathcal{L}_{pre} + \mathcal{L}_{pre} \tag{5}$$

where the pre-symptomatic term \mathcal{L}_{pre} includes only the cases with $t_{i,2} < t_{s,1}$ and the symptomatic term \mathcal{L}_{sym} includes all other cases. We assume a prior probability $p_{pre} = p_{sym} = 0.5$ for pre-symptomatic and symptomatic transmission. Then, for each transmission event, we can obtain the posterior probability of pre-symptomatic transmission via a Bayesian approach:

$$P[\text{presymptomatic transmission}] = \frac{p_{\text{pre}}\mathcal{L}_{\text{pre}}(\hat{\Theta}_{\omega})}{p_{\text{pre}}\mathcal{L}_{\text{pre}}(\hat{\Theta}_{\omega}) + p_{\text{sym}}\mathcal{L}_{\text{sym}}(\hat{\Theta}_{\omega})}$$
(6)

where all likelihoods are evaluated at the parameters Θ_{ω} that maximise the composite likelihood. The distribution of the fraction of transmissions that occurred before onset of symptoms can be estimated by assigning each event as pre-symptomatic or symptomatic at random according to its posterior probability. The empirical distribution of this quantity is obtained from 10,000 random extractions from the posterior, shown in the left panel of Supplementary Figure 5.

$\beta(\tau)$ and the renewal equation

In an epidemic which is growing exponentially, in a deterministic manner, driven by human-to-human¹ transmission, the incidence I(t) can be described by the renewal equation:

$$I(t) = \int_0^\infty I(t-\tau)\beta(\tau)d\tau,\tag{7}$$

¹Equation 7 describes only human-to-human transmission (though this may be indirectly via the environment, as we clarify). Excluding vector-borne diseases, zoonosis events do not scale with the number of people currently infected, and so become a negligible contribution after human-to-human transmission has begun driving exponential growth.

In words, Equation 7 says that the incidence now is set by the rate at which people were infected at all previous times, weighted by how infectious those people are now. $\beta(\tau)$ is the mean rate at which an individual infects others a time τ after being infected itself. Here we take $\beta(\tau)$ to be independent of the stage of the epidemic (calendar time t): we neglect depletion of susceptible individuals through acquired immunity, changing contact patterns etc. over the timescale of the data informing our estimations of $\beta(\tau)$. After $\beta(\tau)$ has been determined, we may consider how to change it through interventions to reduce infectiousness. If one's direct and indirect infectiousness is certain to be zero after having been infected for a time T say, we only need consider the previous time window T of the epidemic – replacing the upper limit of the integral in Equation 7 by T for convenience. We take T to be infinite for generality, with $\beta(\tau)$ tending to zero at large times. Substituting into Equation 7 an exponentially growing incidence, $I(t) = I_0 e^{rt}$, gives the condition

$$1 = \int_0^\infty e^{-r\tau} \beta(\tau) d\tau, \tag{8}$$

 $\beta(\tau)$ can be written as the product of two things: R_0 and the unit-normalised function $w(\tau)$

$$\beta(\tau) = R_0 w(\tau), \text{ with } R_0 = \int_0^\infty \beta(\tau) d\tau$$
 (9)

 $w(\tau)$ is the generation time distribution – the probability density function for the time between an individual becoming infected and their subsequent onward transmission events. R_0 is the basic reproduction number. If the exponential growth rate r and the generation time distribution $w(\tau)$ have been estimated, R_0 is determined by Equation 8, i.e.

$$R_0 = 1/\int_0^\infty e^{-r\tau} w(\tau) d\tau, \tag{10}$$

We decompose $\beta(\tau)$ without any loss of generality into the following distinct contributions:

- Direct transmissions from asymptomatic individuals those who never develop symptoms. The degree to which individuals show symptoms is of course a continuum, but a threshold can be defined for clinical purposes (i.e. sub-clinical and clinical infections) or for epidemiological purposes. We define P_a as the proportion of such individuals among all infected individuals, and $\beta_a(\tau)$ as their mean infectiousness at age-of-infection τ .
- Direct transmissions from pre-symptomatic individuals (currently without symptoms, but who will develop symptoms later). We define $\beta_p(\tau)$ as the mean infectiousness of these individuals at age-of-infection τ , conditional upon their being pre-symptomatic, which has probability $1 s(\tau)$ where $s(\tau)$ is the cumulative distribution function of the incubation period distribution.
- Direct transmissions from symptomatic individuals (including those who have stopped showing symptoms, in general, if infectiousness may outlast symptoms), with infectiousness $\beta_s(\tau)$ conditional on having started symptoms.
- Indirect transmission via the environment. We define $\beta_e(\tau)$ as the mean rate of contaminating one's environment (with the mean being over all asymptomatic, pre-symptomatic and symptomatic individuals at age-of-infection τ). Let E(l) be the rate at which contaminated environment infects new individuals a time lag l after having been contaminated. The environmentally mediated infectiousness of an individual infected a time τ ago is given by the total effect of their previous environmental contamination now: $\int_{l=0}^{\tau} \beta_e(\tau-l)E(l)d\tau$.

We therefore have, in general,

$$\beta(\tau) = P_a \beta_a(\tau) + (1 - P_a)(1 - s(\tau))\beta_p(\tau) + (1 - P_a)s(\tau)\beta_s(\tau) + \int_{l=0}^{\tau} \beta_e(\tau - l)E(l)dl$$
 (11)

Integrating each of these terms separately gives their respective contribution to R_0 :

$$R_0 = R_a + R_p + R_s + R_e (12)$$

We make the following simplifying assumptions about the contributions described above, compared to the general case.

- Asymptomatic individuals are assumed to have an infectiousness proportional to that of symptomatic individuals: $\beta_a(\tau) = x_a \beta_s(\tau)$.
- Pre-symptomatic individuals are assumed to have an infectiousness equal to that of symptomatic individuals at the same age of infection: $\beta_p(\tau) = \beta_s(\tau)$
- The rate at which individuals contaminate their environment, $\beta_e(\tau)$ is assumed to be proportional to the direct infectiousness of symptomatic individuals $\beta_s(\tau)$. The proportionality constant can be absorbed into the function E(l) which multiplies $\beta_e(\tau)$, so that we have $\beta_e(\tau) = \beta_s(\tau)$.

The $\beta(\tau)$ we consider is therefore

$$\beta(\tau) = \underbrace{P_a x_a \beta_s(\tau)}_{\text{asymptomatic}} + \underbrace{(1 - P_a)(1 - s(\tau))\beta_s(\tau)}_{\text{pre-symptomatic}} + \underbrace{(1 - P_a)s(\tau)\beta_s(\tau)}_{\text{symptomatic}} + \underbrace{\int_{l=0}^{\tau} \beta_s(\tau - l)E(l)dl}_{\text{environmental}}$$
(13)

Derivation of the impact of interventions

We follow the approach and notation of (Fraser, Riley et al. 2004). One difference in notation is that we use $s(\tau)$ here to denote the probability of having started showing symptoms, corresponding to their $(1 - S(\tau))$.

Prelude: describing epidemic growth with no intervention

Let $Y(t, \tau, \tau')$ be the number² of individuals at time t who were infected at a time $t - \tau$ by individuals who were in turn infected at a time $t - \tau'$. Y satisfies $Y(t + dt, \tau + dt, \tau' + dt) = Y(t, \tau, \tau')$, 'translational invariance', because incrementing all three arguments by exactly the same value means following exactly the same cohort of individuals through to a different moment in time. Equivalently,

$$\frac{\partial Y(t,\tau,\tau')}{\partial t} + \frac{\partial Y(t,\tau,\tau')}{\partial \tau} + \frac{\partial Y(t,\tau,\tau')}{\partial \tau'} = 0 \tag{14}$$

In the absence of any intervention, $Y(t, \tau, \tau')$ satisfies the generalised Kermack-McKendrick equations (also referred to as the Von Foerster equations):

$$Y(t,0,\tau) = \beta(\tau) \int_{\tau'=\tau}^{\infty} Y(t,\tau,\tau') d\tau'$$
(15)

²Technically Y is a double-density: one must integrate it over a range of τ values and a range of τ' values to get an actual number of people at time t.

In words, equation 15 says that the incidence of newly infected individuals at time t due to individuals who were themselves infected a time τ ago is given by the current infectiousness of those individuals infected a time τ ago, multiplied by how many of those individuals there are: integrating their number over all possible times that their infector was infected. Since the number of infected individuals is self-renewing, we anticipate exponential dependence on t. Together with translational invariance, this implies the general form

$$Y(t,\tau,\tau') = y(\tau'-\tau)e^{r(t-\tau)}$$
(16)

Substituting equation 16 into equation 15 gives

$$y(\tau)e^{rt} = \beta(\tau)\int_{\tau'=\tau}^{\infty} y(\tau'-\tau)e^{r(t-\tau)}d\tau'$$
(17)

$$y(\tau) = e^{-r\tau} \beta(\tau) \int_{\tau'=\tau}^{\infty} y(\tau'-\tau) d\tau'$$
 (18)

$$=e^{-r\tau}\beta(\tau)\int_{\rho=0}^{\infty}y(\rho)d\rho\tag{19}$$

The solution to equation 19 is, for any value of the constant y_0 ,

$$y(\tau) = y_0 e^{-r\tau} \beta(\tau)$$
, with the constraint $\int_{\rho=0}^{\infty} e^{-r\rho} \beta(\rho) d\rho = 1$ (20)

Decomposing $\beta(\tau)$ into two factors – its integral R_0 , and the unit-normalised generation time interval $w(\tau)$ – the constraint above gives the relationship between r, R_0 and $w(\tau)$. Substituting this solution for $y(\tau)$ back into equation 16 gives

$$Y(t,\tau,\tau') = y_0 e^{-r(\tau'-\tau)} \beta(\tau'-\tau) e^{r(t-\tau)}$$
(21)

$$= y_0 e^{r(t-\tau')} \beta(\tau' - \tau) \tag{22}$$

The impact of case isolation, contact tracing and quarantine

In the presence of case isolation of efficacy ϵ_I and contact tracing plus quarantine of efficacy ϵ_T , equation 15 is modified (see Fraser, Riley et al. 2004) to

$$Y(t,0,\tau) = \beta(\tau) \left(1 - \epsilon_I s(\tau)\right) \int_{\tau}^{\infty} \left(1 - \epsilon_T + \epsilon_T \frac{1 - s(\tau')}{1 - s(\tau' - \tau)}\right) Y(t,\tau,\tau') d\tau'$$
 (23)

Note that we take the upper limit of the integral here to be ∞ rather than t (so that the epidemic is solved exactly by exponential growth, instead of beginning according to specific boundary conditions and then tending toward exponential growth as the boundary conditions are forgotten). Translational invariance together with exponential growth with t imply the same general form as previously – equation 16 – but with a different functional form for $y(\tau)$. Substituting equation 16 into equation 23 gives the 'next-generation' equation for y:

$$y(\tau) = e^{-r\tau} \beta(\tau) \left(1 - \epsilon_I s(\tau) \right) \int_{\tau'=\tau}^{\infty} \left(1 - \epsilon_T + \epsilon_T \frac{1 - s(\tau')}{1 - s(\tau' - \tau)} \right) y(\tau' - \tau) d\tau'$$
 (24)

$$= e^{-r\tau} \beta(\tau) \left(1 - \epsilon_I s(\tau)\right) \int_{\rho=0}^{\infty} \left(1 - \epsilon_T \frac{s(\rho + \tau) - s(\rho)}{1 - s(\rho)}\right) y(\rho) d\rho \tag{25}$$

redefining the integration variable to be $\rho = \tau' - \tau$ for convenience. Hence, the growth rate after the interventions corresponds to the value of r for which the functional linear 'next-generation' operator

$$\mathcal{N}_r y = e^{-r\tau} \beta(\tau) \left(1 - \epsilon_I s(\tau) \right) \int_0^\infty \left(1 - \epsilon_T \frac{s(\rho + \tau) - s(\rho)}{1 - s(\rho)} \right) y(\rho) d\rho \tag{26}$$

has the largest eigenvalue equal to 1. In other words, given the eigenvalue equation $\mathcal{N}_r y = \nu_r^{max} y$, r is determined by $\nu_r^{max} = 1$, in a functional generalization of the Euler-Lotka equation.

If R_0 is unknown, but the generation time distribution $\omega(\tau) = \beta(\tau)/R_0$ is known, then we consider the operator

$$\mathcal{O} = \mathcal{N}_0 / R_0 = \omega(\tau) \left(1 - \epsilon_I s(\tau) \right) \int_0^\infty \left(1 - \epsilon_T \frac{s(\rho + \tau) - s(\rho)}{1 - s(\rho)} \right) (\cdot) d\rho \tag{27}$$

By construction, this operator has maximal eigenvalue $1/R_0$ if and only if the value of R_0 corresponds to a growth rate r=0 after interventions. Then, the inverse of the largest eigenvalue of \mathcal{O} is precisely the maximum value of R_0 for which $R \leq 1$ in the presence of interventions.

This approach includes potential transmission from a fraction P_a of completely asymptomatic individuals. Denote by x_a the relative infectiousness of those individuals (which could be different from 1). For the purpose of this model, this is effectively equivalent to having $P_a x_a$ infectious individuals with relative infectiousness 1. Hence, the effective fraction of asymptomatic individuals is $P_a x_a/(1 - P_a + P_a x_a)$. Then, the fraction of infected individuals that have already shown symptoms a time τ after infection corresponds to

$$s(\tau) = \frac{1 - P_a}{1 - P_a + P_a x_a} \int_0^{\tau} i(\tau') d\tau'$$
 (28)

where $i(\tau)$ is the incubation time distribution for individuals that will eventually become symptomatic.

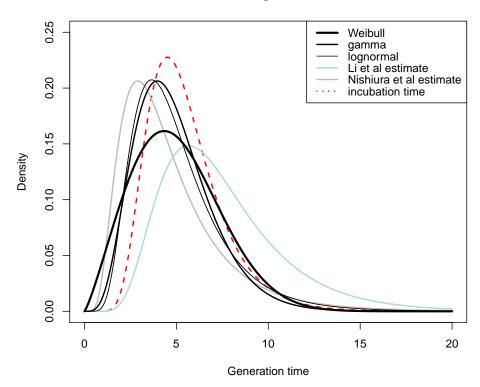
This approach can be modified to include environmental transmission as well. As a first approximation, neglecting the delays due to the persistence of SARS-CoV-2 in the environment, environmental transmission events can be treated as cases of untraceable transmission. Isolation and quarantine are effective in preventing environmental spread, since they stop the infected individual from releasing the virus in the environment. (The only relevant exception is environment transmission within isolation facilities, which is not prevented by isolation - quite the opposite.) We assume here that all environmental transmission events correspond simply to untraceable transmissions, and that the rate of viral shedding in the environment is proportional to the person-to-person infectiousness. Then, the theory developed above can be applied by replacing the efficacy of contact tracing $\epsilon_T \to \epsilon_T (1 - R_e/R_0)$, where the factor $1 - R_e/R_0$ corresponds to the fraction of traceable transmissions and ϵ_T is the efficacy per traceable transmission. Hence, the results above can be generalised to include environmental transmission by considering the next-generation operator

$$\mathcal{N}_r = e^{-r\tau} \beta(\tau) \left(1 - \epsilon_I s(\tau) \right) \int_0^\infty \left(1 - \epsilon_T \left(1 - \frac{R_e}{R_0} \right) \frac{s(\rho + \tau) - s(\rho)}{1 - s(\rho)} \right) (\cdot) d\rho \tag{29}$$

Numerically, the eigenfunction corresponding to the largest eigenvalue can be found by iterating the operator \mathcal{N}_r . Up to a constant factor, the exact solution is given by $\lim_{k\to\infty} \mathcal{N}_r^k y_0$ where y_0 is an arbitrary initial condition; in practice, $\mathcal{N}_r^k y_0$ for large enough k provides a good approximation to the solution, with exponentially fast convergence (at a rate given by the ratio of second- and first-most dominant eigenvalues per application of the operator). We found that fewer than 10 iterations provided an excellent approximation in the parameter space and for the distributions and values of R_0 considered in this paper.

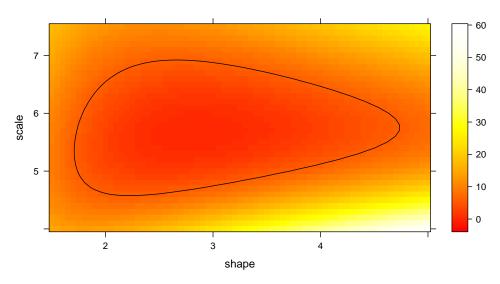
Supplementary Figures

Distribution of generation times



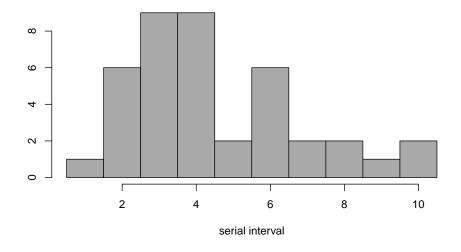
Supplementary Figure 1: Same as Figure 1 left panel in Main Text, but the generation time distribution is inferred from the likelihood function conditional on the occurrence of the transmission events.



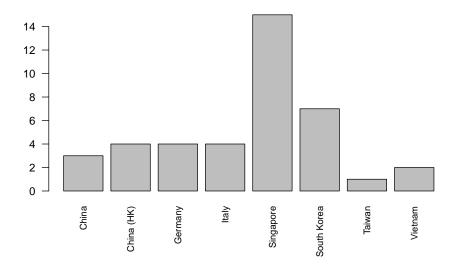


Supplementary Figure 2: Contour of the 95% confidence interval for the shape and scale parameters of the Weibull distribution of generation times. The colours represent the $-2\Delta \log \mathcal{L}$ test statistic (lower values correspond to parameters for which the observed data is more likely).

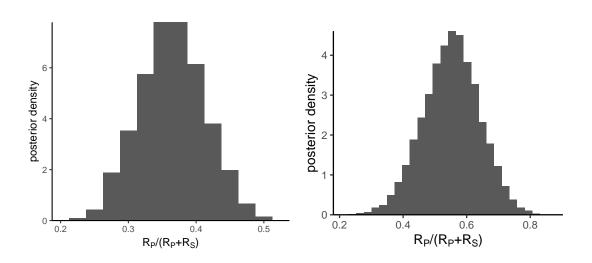
empirical distribution of serial intervals



Supplementary Figure 3: Distribution of serial intervals (i.e. time from onset of symptoms in source to onset of symptoms in recipient) from the 40 transmission pairs analysed in this study.

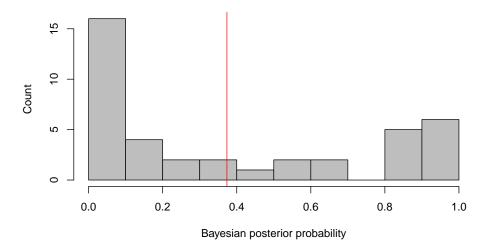


Supplementary Figure 4: Geographical distribution of transmission occurrence among the 40 source-recipient pairs analysed in this study.



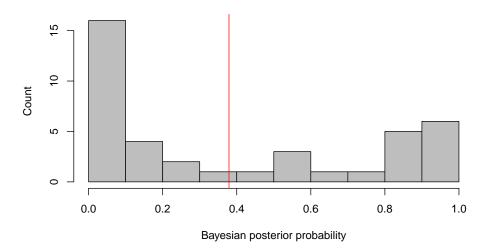
Supplementary Figure 5: Uncertainty on the contribution of pre-symptomatic transmissions to all transmissions from infected individuals that will eventually show symptoms. Left: posterior distribution from the 40 transmission pairs analysed in this study. Right: posterior distribution from the infectiousness model approach.

Probability of pre-symptomatic transmission



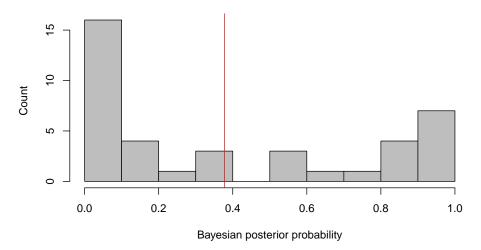
Supplementary Figure 6: Same as Figure 1 right panel in Main Text, but assuming a gamma distribution of generation times.

Probability of pre-symptomatic transmission

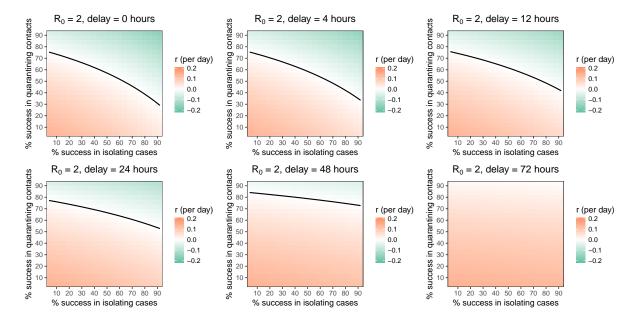


Supplementary Figure 7: Same as Figure 1 right panel in Main Text, but assuming a lognormal distribution of generation times.

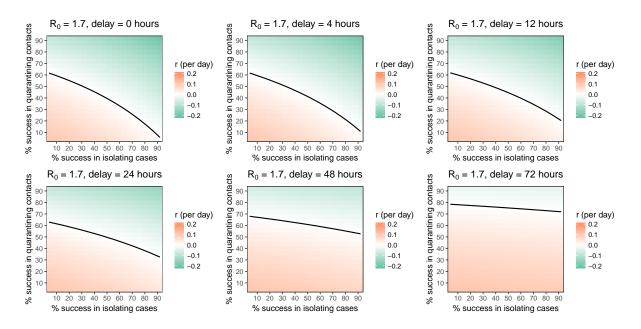
Probability of pre-symptomatic transmission



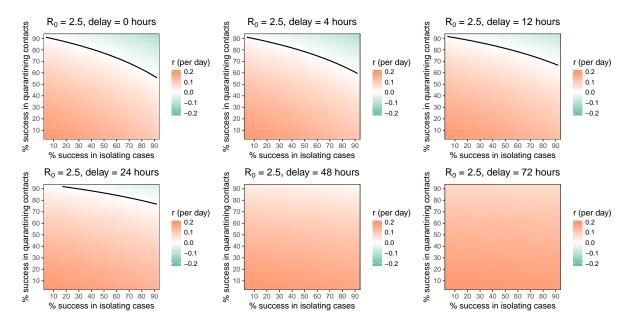
Supplementary Figure 8: Same as Figure 1 right panel in Main Text, but the posterior probabilities of presymptomatic transmission are inferred from the likelihood function conditional on the occurrence of the transmission events.



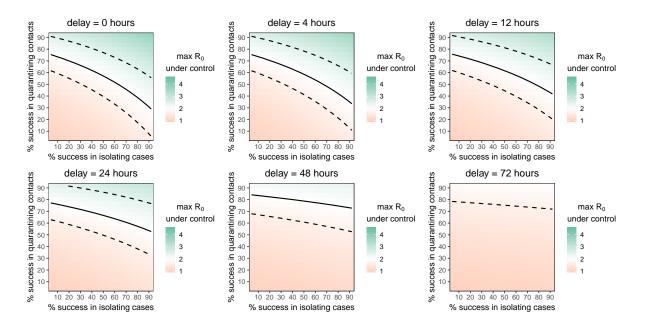
Supplementary Figure 9: As Figure 4 in the main text, but varying the delay (from initiation of symptoms to isolation and quarantine of contacts) between the six panels. Instead of showing uncertainty in the position of r=0 due to uncertainty in R_0 , as in main text Figure 4, we show results for each different value of R_0 in different plots: see Supplementary Figures 10 and 13. The fraction of transmissions that are environmentally mediated is fixed at 10%, and the fraction from asymptomatic individuals at 6.25% (corresponding to our central estimates of $P_a=40\%$, $x_a=0.1$), as in the main text.



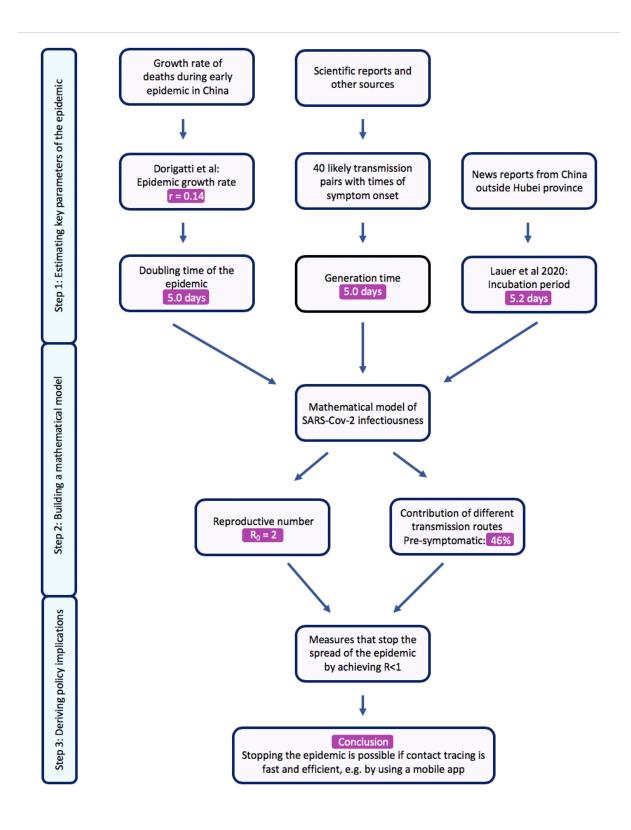
Supplementary Figure 10: As Supplementary Figure 9, but using the lower bound R_0 value of 1.7 instead of the central estimate of 2.



Supplementary Figure 11: As Supplementary Figure 9, but using the upper bound R_0 value of 2.5 instead of the central estimate of 2.



Supplementary Figure 12: Heat map plot showing the maximum possible R_0 that could be contained (by reduction to less than 1) for a pathogen with the same generation time distribution that we have inferred for SARS-CoV-2, as a function of the success rate of isolation of symptomatic cases (x axis) and the success rate of contact tracing (y axis). Contours of constant R_0 show the different combinations of the two success rates that would be able to control the epidemic for an R_0 of that value. The solid black line is such a contour for our central estimate of R_0 for SARS-CoV-2, namely $R_0 = 2$; this line therefore shows the threshold for epidemic control. The dashed lines show the contours of the upper and lower bounds of the CI for R_0 (1.7 - 2.5), i.e. these indicate one source of uncertainty in the epidemic control threshold. Each of the six panels shows results for a different delay from initiation of symptoms to isolation of the case and quarantine of their contacts (the same delay for both). The fraction of transmissions that are environmentally mediated is fixed at 10%, and the fraction from asymptomatic individuals at 6.25% (corresponding to our central estimates of $P_a = 40\%$, $x_a = 0.1$), as in the main text.



Supplementary Figure 13: A graphical summary of our analysis.