

# Biostatistics Homework 5

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## 1. One-sided confidence interval 1

In this question, we are interested in one sided confidence intervals. Thus we want to find:

$$P(\mu \leq b) = 0.95$$

Since the normal distribution is symmetric, and our population luckily follows a normal distribution with a known standard deviation, the one sided confidence interval for the lower tail and the upper bound will be just symmetric around  $\bar{X}$ . The margin of error can be calculated with a z score.

```
qnorm(0.95)
```

```
## [1] 1.644854
```

We get from R that when  $z = 1.645$ ,  $P(Z < z) = 0.95$ . We also know that:

$$Z = \frac{\bar{X} - \mu}{\sigma/\sqrt{n}}$$

Thus the one sided confidence interval of the lower tail will be C.I. =  $(-\infty, \bar{X} + 1.645 \frac{\sigma}{\sqrt{n}}]$ .

Similarly, the one sided confidence interval of the upper tail will be C.I. =  $[\bar{X} - 1.645 \frac{\sigma}{\sqrt{n}}, +\infty)$ .

## 2. One-sided confidence interval 2

Different from question 1, although our population still follows a normal distribution, this time, we don't know anything about its variance. Thus, we need to substitute the population variance with the sample variance. Our random variable  $\frac{\bar{X} - \mu}{S/\sqrt{n}}$  follows a T-distribution. We want to find:

$$P(\mu \leq b) = 0.95$$

The T-distribution is also symmetric, so the upper tail and the lower tail is symmetric around  $\bar{X}$ . The margin of error can be calculated from the t score.

If we want to calculate the confidence interval for the lower tail, the t score is  $t_{0.05, v-1}$ . We also know that:

$$t = \frac{\bar{X} - \mu}{s/\sqrt{n}}$$

Thus the confidence interval for the lower tail will be C.I. =  $(-\infty, \bar{X} + t_{0.05, v-1} \frac{s}{\sqrt{n}}]$ .

Similarly, the confidence interval for the upper tail will be C.I. =  $[\bar{X} - t_{0.05, v-1} \frac{s}{\sqrt{n}}, +\infty) = [\bar{X} - t_{0.05, v-1} \frac{s}{\sqrt{n}}, +\infty)$ .

**3.**

T

F

[0.7421, 0.8229] (See 3 supplement)

F

T

**3 supplement** We know from the problem that this is a test for proportion. Thus the confidence interval is:

```
qnorm(0.975)
```

```
## [1] 1.959964
```

```
313/400
```

```
## [1] 0.7825
```

$$\text{C.I.} = [0.7825 - 1.96\sqrt{\frac{(1 - 0.7825) \times 0.7825}{400}}, 0.7825 + 1.96\sqrt{\frac{(1 - 0.7825) \times 0.7825}{400}}]$$

Which is [0.7421, 0.8229].

## 4 Null and Alternative Hypothesis

4.1)

$H_0$  : The rate of female student hadn't increased.  $\pi \leq 7\%$

$H_1$  : The rate of female student increased.  $\pi > 7\%$

4.2)

$H_0$  : The new technique is not significantly better or worse.  $\pi = 75\%$

$H_1$  : The new technique is significantly better or worse.  $\pi \neq 75\%$

4.3)

$H_0$  : The machine is not dispensing too much in medium drinks.  $\mu \leq 530\text{ml}$

$H_1$  : The machine is dispensing too much in medium drinks.  $\mu > 530\text{ml}$

4.4)

$H_0$  : The proportion of people who received the flu shot had not changed.  $\pi = 48\%$

$H_1$  : The proportion of people who received the flu shot had changed.  $\pi \neq 48\%$

## 5 The $p$ -value

F

F

T

F

F

## 6 Blue M&M's candies

6.1)

D

6.2)

$H_0$ : The plain M&M's candies do contain 24% blue ones.  $\pi = 24\%$

$H_1$ : The plain M&M's candies does not contain 24% blue ones.  $\pi \neq 24\%$

6.3)

The test statistic, which should be a z-statistic, because we are testing on a proportion, and  $np > 10, nq > 10$ .

$$z = \frac{p - \pi}{\sqrt{\frac{(1-\pi)\pi}{n}}} = \frac{18.4\% - 24\%}{\sqrt{\frac{(1-24\%) \times 24\%}{500}}} \approx -2.932$$

6.4)

The population follows a binomial distribution. Our sample satisfies  $np > 10$  and  $n(1-p) > 10$ . Thus the test statistic approximately follows a standard normal distribution.

6.5)

There might be multiple reasons. I list some of my explanations below:

- Li Lei is very lucky and got weird M&M's candies.
- Li Lei bought the 500 candies from the same production batch, and they are not of good quality.
- The M&M might change their policy on the proportion of blue candies.
- Li Lei might have counting mistakes (e.g. He is colour blind)

6.6)

The margin of error is:

$$z_{0.025} \sqrt{\frac{(1-p)p}{n}} = 1.96 \times \sqrt{\frac{(1-0.184) \times 0.184}{500}} = 0.03$$

So the two sided 95% confidence interval is:

$$\text{C.I.} = [15.4\%, 21.4\%]$$

6.7

```
pnorm(-2.932) * 2
```

```
## [1] 0.003367867
```

We can see that the  $p$ -value is  $0.003 < 0.01$ . Thus, Li Lei should reject the null hypothesis.

## 7 Renal Disease

7.1)

C

7.2)

$H_0$ : The mean serum-creatinine level in the people who have taken the antibiotic is not significantly different from the normal people.  $\mu = 1.0\text{mg/dL}$

$H_1$ : The mean serum-creatinine level in the people who have taken the antibiotic is significantly different from the normal people.  $\mu \neq 1.0\text{mg/dL}$

7.3)

Since we are comparing between the sample and the whole normal population, and the antibiotic does not change the dispersion of the serum-creatinine level, we can assume that a z statistic should be calculated.

$$z = \frac{\bar{x} - \mu}{\sigma/\sqrt{n}} = \frac{1.2 - 1.0}{0.4/\sqrt{12}} = \sqrt{3} = 1.732$$

7.4)

If  $H_0$  was true, as mentioned above, the test statistics will follow a standard normal distribution.

7.5)

The margin of error is:

$$\varepsilon = z_{0.025} \times \frac{\sigma}{\sqrt{n}} = 1.96 \times \frac{0.4}{\sqrt{12}} \approx 0.2263$$

So the confidence interval is C.I. = [0.9736, 1.4263].

7.6)

```
pnorm(1.732, lower.tail = F) * 2
```

```
## [1] 0.08327356
```

The p-value is 0.08 which is not smaller than 0.05, I think I will not reject  $H_0$ .

## 8 Serum Cholesterol Level

8.1)

From lecture 26, we now that the standard deviation in the sample is 46mg/100ml. We also know that we want a significance level of  $\alpha = 0.05$ . We can calculate the rejection threshold:

$$x_{\text{reject}} = \bar{x} + z_{0.05} \times \frac{\sigma}{\sqrt{n}} = 180 + 1.645 \times \frac{46}{\sqrt{50}} \approx 190.7$$

So the rejection threshold is 190.7mg/100ml.

The power,  $1 - \beta$ , can be calculated from  $\beta$ , which is the possibility of not rejecting the null hypothesis when the null hypothesis is not true. First we calculate the test statistics for the population:

$$z = \frac{x_{\text{reject}} - \mu}{\sigma/\sqrt{n}} = \frac{190.7 - 200}{46/\sqrt{50}} \approx -1.43$$

Then we can calculate the probability of  $P(\text{not reject } H_0 | H_0 \text{ is false})$  which is  $\beta$ .

```
pnorm(-1.43)
```

```
## [1] 0.07635851
```

So  $\beta = 0.076$ . The power is  $1 - \beta = 0.924$ .

## 8.2)

Because I don't want a step by step calculation, so I'm just going to derive the general formula.

The test statistic we calculate are  $z_\alpha$  and  $z_{1-\beta} = -z_\beta$ . From the previous example, we can see that:

$$-z_\beta = \frac{\bar{x} + z_\alpha \times \frac{\sigma}{\sqrt{n}} - \mu}{\sigma/\sqrt{n}}$$

With a bit of fumbling around, we can get the equation:

$$n = \left( \frac{(z_\alpha + z_\beta)\sigma}{\mu - \bar{x}} \right)^2$$

So in our case,  $\lceil n \rceil = 101$ , so the required sample size is 101.

## 8.3)

B

## 9 Type I & II errors

### 9.1)

B

### 9.2)

C

### 9.3)

B

### 9.4)

D (But I think C is also ok, because we don't want to waste a lot of money to clean the pool.)