



# Linear models and generalized linear models

in R

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# What is a linear regression

## Simple linear regression

$$Y_i = \beta_0 + \beta_1 x_i + \epsilon$$
$$\epsilon \sim N(0, \sigma_\epsilon^2)$$

or in distributional notation

$$Y \sim N(\beta_0 + \beta_1 x, \sigma_\epsilon^2)$$

## General linear model

$$Y_i = \beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \dots + \epsilon$$
$$\epsilon \sim N(0, \sigma_\epsilon^2)$$

and

$$Y \sim N(\beta_0 + \beta_1 x_{1i} + \beta_2 x_{2i} + \dots, \sigma^2)$$

# Linear model assumptions

Some are made on the residuals and others on the independent variables. None are made on the (unconditionned) dependent variable.

Residuals are assumed to:

- have a mean of zero
- be independent
- be normally distributed
- be homoscedastic

Independent variables are assumed to:

- have a linear relation with Y
- be measured without error
- to be independent from each other

# Maximum likelihood

Technique used for estimating the parameters of a given distribution, using some observed data

For Example: Population is known to follow a “normal distribution” but “mean” and “variance” are unknown, MLE can be used to estimate them using a limited sample of the population.

# Likelihood vs probability

We maximize the likelihood and make inferences on the probability

## Likelihood

$$L(\textit{parameters}|\textit{data})$$

How likely it is to get those parameters given the data.

## Probability

$$P(\textit{data}|\textit{null parameters})$$

Probability to get the data given the null parameters. Or how probable it is to get those data according to the null model.

# Maximum likelihood approach

$$L(parameters|data) = \prod_{i=1}^n f(data_i|parameters)$$

where  $f$  is the probability density function of your model.

Working with product is more painful than with sum, we can take the log:

$$\ln(L(parameters|data)) = \sum_{i=1}^n \ln(f(data_i|parameters))$$

Need to solve:

$$\frac{\delta \ln(L(parameters|data))}{\delta parameters} = 0$$

For multiple regression, the parameters  $\beta$ s are given by  $\beta = (X^T X)^{-1} X^T y$

Equivalent to minimizing residuals (but you don't want to see the proof)

# Doing linear models in R

Simply use `lm()` function. It works for everything anova, ancova, t-test.

We will use data of sturgeon measurements at different locations in Canada.

```
dat <- read.csv("data/lm_example.csv")
str(dat)
```

```
## 'data.frame':    92 obs. of  4 variables:
## $ year      : int  1978 1978 1978 1978 1978 1978 1978 1978 1978 1978 ...
## $ fklngth   : num  41.9 50.2 50.2 47.3 49.6 ...
## $ locate    : chr   "NELSON"      "LOFW"      "LOFW"      "NELSON"      " ...
## $ age       : int   11 24 23 20 23 20 23 19 17 14 ...
```

# Fitting a model and checking assumptions

First we load the needed packages for:

- data manipulation: tidyverse
- fancy plots: ggplot2
- type III anova: car
- fancy and nicer visual assumptions checks: performance
- formal assumptions tests: lmtest

```
library(car)
library(performance)
library(lmtest)
library(tidyverse)
```



# Data exploration

R Code

Plot

```
ggplot(data = dat, aes(x = age, y = fklength)) +  
  facet_grid(. ~ locate) +  
  geom_point() +  
  stat_smooth(method = lm, se = FALSE) +  
  stat_smooth(se = FALSE, color = "red") +  
  labs(  
    y = "Fork length",  
    x = "Age"  
  )
```

# Creating log10 transform

```
dat <- dat %>%  
  mutate(  
    lage = log10(age),  
    lfk1 = log10(fk1length)  
  )
```

# Data exploration: with log

Code

Plot

```
ggplot(data = dat, aes(x = lage, y = lfk1)) +  
  facet_grid(. ~ locate) +  
  geom_point() +  
  stat_smooth(method = lm, se = FALSE) +  
  stat_smooth(se = FALSE, color = "red") +  
  labs(  
    y = "log 10 Fork length",  
    x = "Log 10 Age"  
  )
```

# Fit the model

```
m1 <- lm(lfkl ~ lage + locate + lage:locate, data = dat)
summary(m1)
```

```
##
## Call:
## lm(formula = lfkl ~ lage + locate + lage:locate, data = dat)
##
## Residuals:
##      Min       1Q   Median       3Q      Max
## -0.09375 -0.01864 -0.00253  0.02090  0.08030
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept)    1.24287    0.04370  28.443  < 2e-16 ***
## lage           0.31431    0.03292   9.546 3.08e-15 ***
## locateNELSON    0.19295    0.06331   3.048  0.00304 **
## lage:locateNELSON -0.14276    0.04902  -2.912  0.00455 **
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 0.02823 on 88 degrees of freedom
## Multiple R-squared:  0.5664,    Adjusted R-squared:  0.5516
## F-statistic: 38.31 on 3 and 88 DF,  p-value: 6.197e-16
```

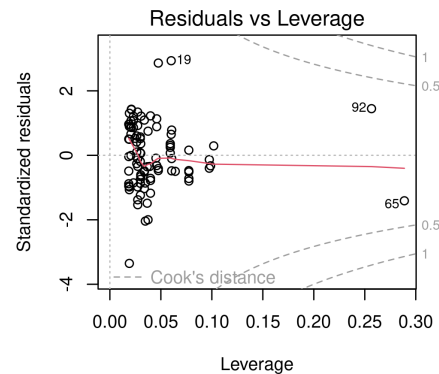
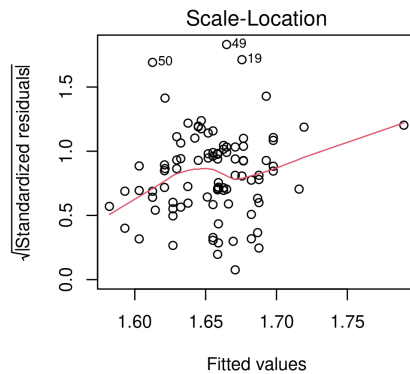
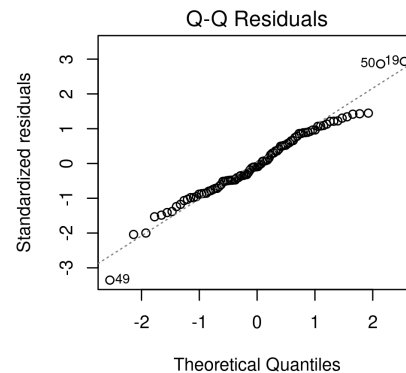
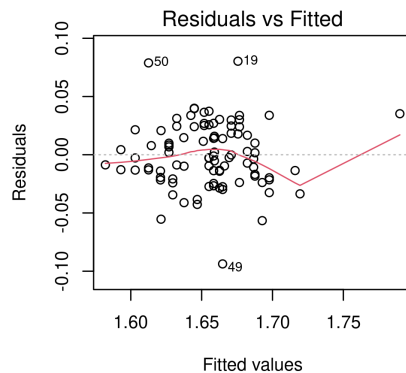
# Anova for factors

```
Anova(m1, type = 3)
```

```
## Anova Table (Type III tests)
##
## Response: lfk1
##           Sum Sq Df  F value    Pr(>F)
## (Intercept) 0.64467  1 809.0107 < 2.2e-16 ***
## lage        0.07262  1  91.1310 3.079e-15 ***
## locate      0.00740  1   9.2901 0.003042 **
## lage:locate 0.00676  1   8.4815 0.004546 **
## Residuals   0.07012 88
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

# Assumptions (classic plot)

```
par(mfrow = c(2, 2))  
plot(m1)
```

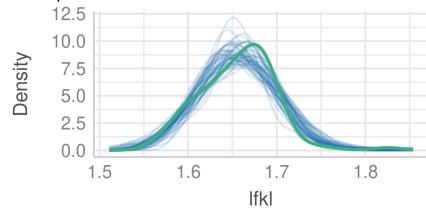


# Assumptions (Nicer plot)

```
check_model(m1)
```

Posterior Predictive Check

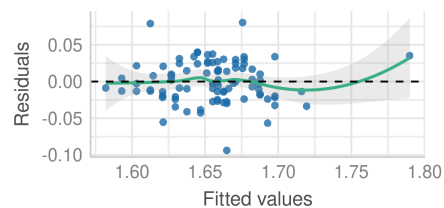
Model-predicted lines should resemble observed data



— Observed data — Model-predicted data

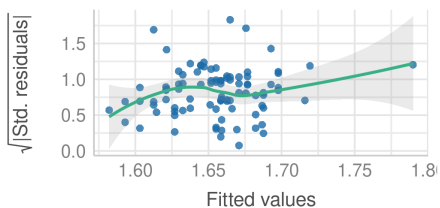
Linearity

Reference line should be flat and horizontal



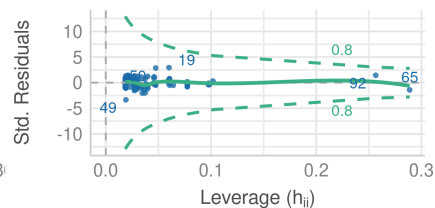
Homogeneity of Variance

Reference line should be flat and horizontal



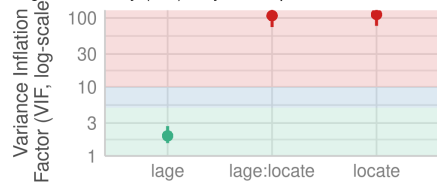
Influential Observations

Points should be inside the contour lines



Collinearity

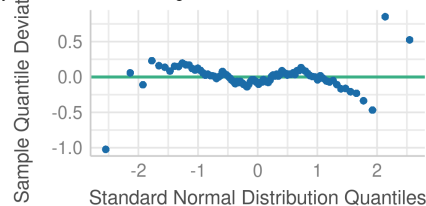
High collinearity (VIF) may inflate parameter uncertainty



— Low ( $< 5$ ) — High ( $\geq 10$ )

Normality of Residuals

Dots should fall along the line



# Formal tests

## Normality of residuals

```
shapiro.test(residuals(m1))
```

```
##  
##      Shapiro-Wilk normality test  
##  
## data:  residuals(m1)  
## W = 0.97639, p-value = 0.09329
```



# Formal tests

## Heteroscedasticity

```
bptest(m1)
```

```
##  
##      studentized Breusch-Pagan test  
##  
## data:  m1  
## BP = 1.8366, df = 3, p-value = 0.607
```

# Formal tests

## Linearity

```
resettest(m1, power = 2:3, type = "fitted", data = dat)
```

```
##  
##      RESET test  
##  
## data:  m1  
## RESET = 1.6953, df1 = 2, df2 = 86, p-value = 0.1896
```

# Generalized linear models

# Generalized linear models

An extension to **linear models**

GLM expresses the transformed conditional expectation of the dependent variable  $Y$  as a linear combination of the regression variables  $X$

Model has 3 components

- a dependent variable  $Y$  with a response distribution to model it:  
**Gaussian, Binomial, Bernoulli, Poisson, negative binomial, zero-inflated ..., zero-truncated ..., ...**
- linear predictors (or independent variables)

$$\eta = \beta_0 + \beta_1 X_1 + \dots + \beta_k X_k$$

- a link function such that

$$E(Y|X) = \mu = g^{-1}(\eta)$$

# Dependent variable

- when continuous and follows *conditional* normal distribution, called **Linear regression**
- Binary outcomes (success/failure), follows a *Binomial distribution*, called **Logistic regression**
- Count data (number of events), follows a *Poisson*, called **Poisson regression**

# Classic link functions

- Identity link (form used in linear regression models)

$$g(\eta) = \mu$$

- Log link (used when  $\mu$  cannot be negative, *e.g.* Poisson data)

$$g(\eta) = \log(\mu)$$

Logit link (used when  $\mu$  is bounded between 0 and 1, *e.g.* binary data)

$$g(\eta) = \log \left( \frac{\mu}{1 - \mu} \right)$$

# Linear regression

- Y: continuous
- Response distribution: Gaussian
- Link function: identity

$$g(\eta) = \mu$$
$$\mu(X_1, \dots, X_k) = \beta_0 + \beta_1 X_1 + \dots + \beta_k X_k$$

# Logistic regression

- Y: binary or proportion
- Response distribution: Binomial or bernoulli
- Link function: logit

$$g(\eta) = \ln \left( \frac{\mu}{1 - \mu} \right)$$
$$\mu(X_1, \dots, X_k) = \frac{1}{1 + e^{-(\beta_0 + \beta_1 X_1 + \dots + \beta_k X_k)}}$$



# Poisson regression

- Y: discrete variable (integers)
- Response distribution: Poisson or Negative binomial
- Link function: natural logarithm

$$g(\eta) = \ln(\mu)$$
$$\mu(X_1, \dots, X_k) = e^{\beta_0 + \beta_1 X_1 + \dots + \beta_k X_k}$$

# Model assumptions

- Easy answer none or really few
- More advanced answer I am not sure, it is complicated
- Just check residuals I as usual
- Technically only 3 assumption:
  - **Variance is a function of the mean specific to the distribution used**
  - observations are independent
  - linear relation on the latent scale

GLMs do not care if the residual errors are Gaussian as long as the specified mean-variance relationship is satisfied by the data

- what about DHaRMA ? It's complicated

# Choosing a link function

A link function should map the structural component from  $(-\infty, \infty)$  to the distribution interval (e.g. (0,1) for binomial)

So number of link function possible is extremely large.

Choice of **link** function heavily influenced by field tradition

For binomial models

- **logit** assume modelling probability of an observation to be one
- **probit** assume binary outcome from a hidden gaussian variable (*i.e.* threshold model)
- **logit** & **probit** are really similar, both are symmetric but **probit** tapers faster. **logit** coefficient easier to interpret directly
- **cologlog** not-symmetrical

# Logistic regression

# Data

Here is some data to play with from a study on bighorn sheep.

We will look at the relation between reproduction and age

Loading and tweaking the data

```
mouflon0 <- read.csv("data/mouflon.csv")
mouflon <- mouflon0 %>%
  arrange(age) %>%
  mutate(
    reproduction = case_when(
      age >= 13 ~ 0,
      age <= 4 ~ 1,
      .default = reproduction
    )
  )
```

# First plot

Code

Plot

```
bubble <- data.frame(  
  age = rep(2:16, 2),  
  reproduction = rep(0:1, each = 15),  
  size = c(table(mouflon$age, mouflon$reproduction))  
) %>%  
  mutate(size = ifelse(size == 0, NA, size))  
ggplot(  
  bubble,  
  aes(x = age, y = reproduction, size = size)  
) +  
  geom_point(alpha = 0.8) +  
  scale_size(range = c(.1, 20), name = "Nb individuals")
```

# Fitting the logistic regression

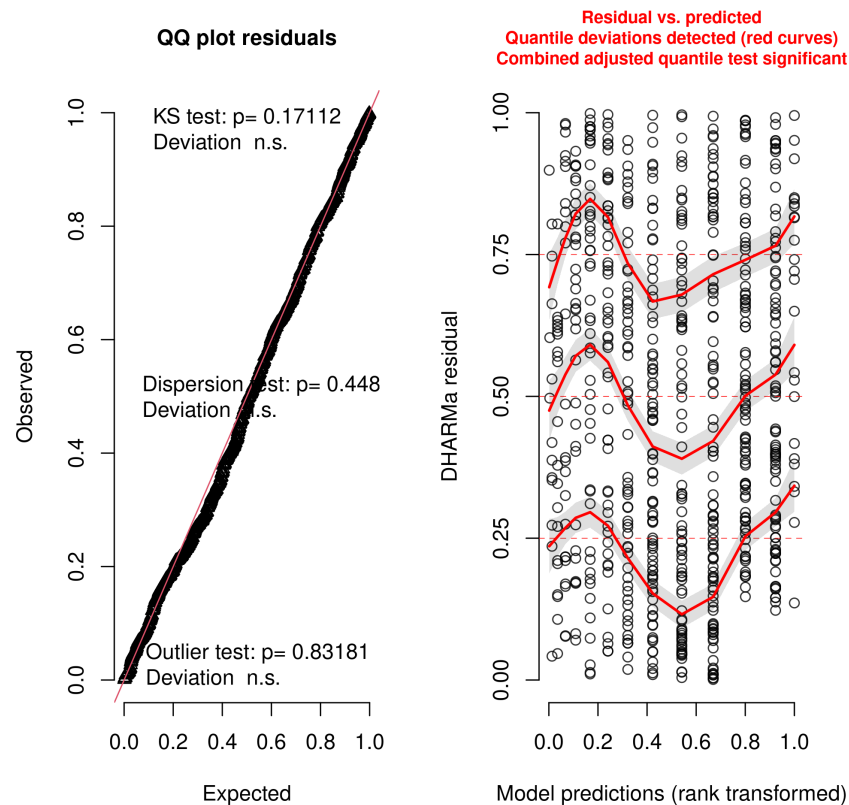
```
m1 <- glm(reproduction ~ age,  
  data = mouflon,  
  family = binomial  
)  
summary(m1)
```

```
##  
## Call:  
## glm(formula = reproduction ~ age, family = binomial, data = mouflon)  
##  
## Coefficients:  
##              Estimate Std. Error z value Pr(>|z|)  
## (Intercept)  3.19921    0.25417   12.59  <2e-16 ***  
## age         -0.36685    0.03287  -11.16  <2e-16 ***  
## ---  
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1  
##  
## (Dispersion parameter for binomial family taken to be 1)  
##  
##    Null deviance: 928.86  on 715  degrees of freedom  
## Residual deviance: 767.51  on 714  degrees of freedom  
## (4 observations deleted due to missingness)  
## AIC: 771.51  
##  
## Number of Fisher Scoring iterations: 4
```

# Checking assumptions

```
simulationOutput <- simulateResiduals(m1)  
plot(simulationOutput)
```

DHARMa residual

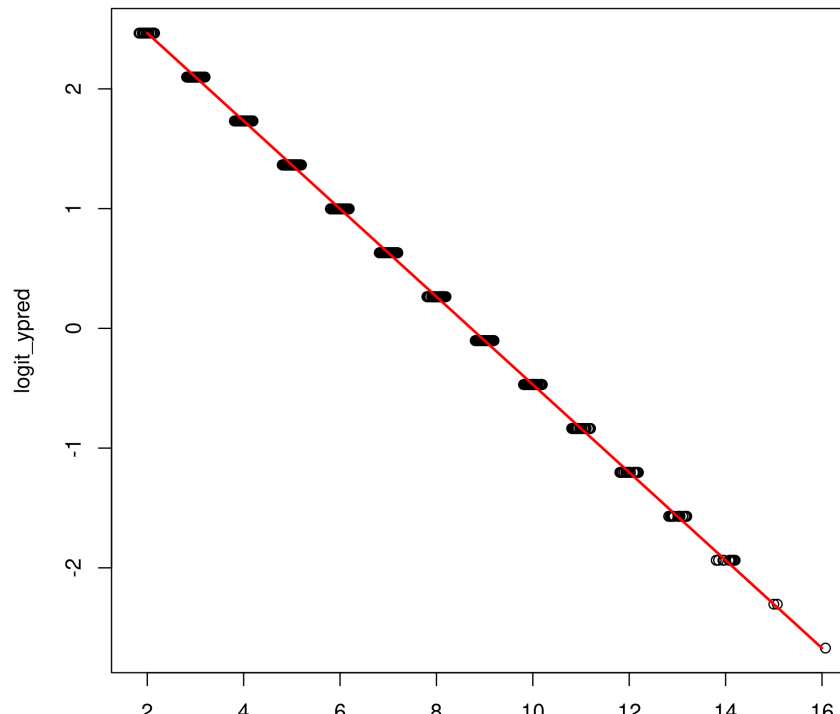




# Plotting predictions

plotting the model prediction on the link (latent) scale

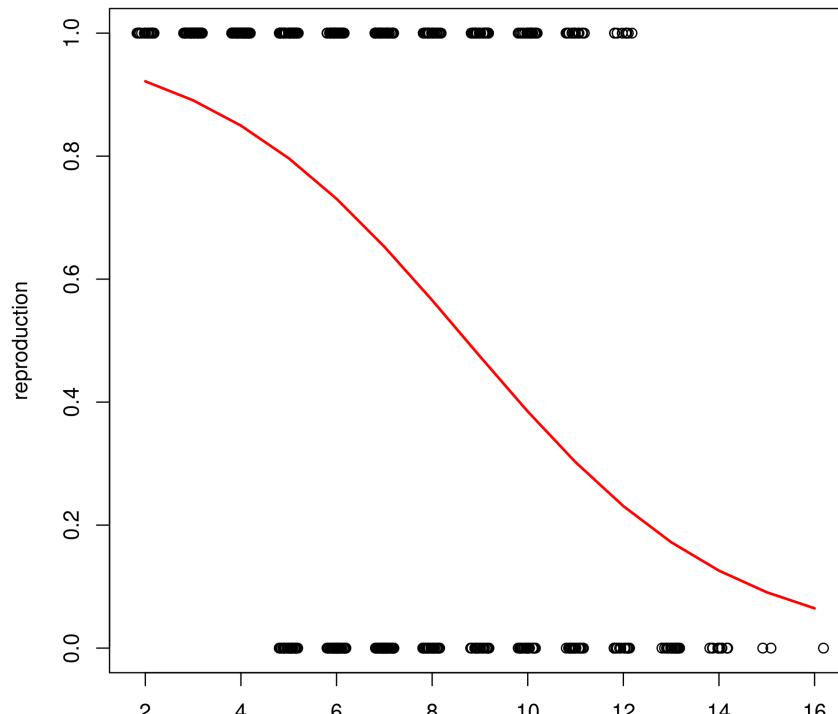
```
mouflon$logit_ypred <- 3.19921 - 0.36685 * mouflon$age  
plot(logit_ypred ~ jitter(age), mouflon)  
points(mouflon$age, mouflon$logit_ypred, col = "red", type = "l", lwd = 2)
```



# Plotting predictions

plotting on the observed scale

```
mouflon$ypred <- exp(mouflon$logit_ypred) / (1 + exp(mouflon$logit_ypred))  
plot(reproduction ~ jitter(age), mouflon)  
points(mouflon$age, mouflon$ypred, col = "red", type = "l", lwd = 2)
```



# Plotting predictions

but it can be much simpler

Code

Plot

```
dat_predict <- data.frame(  
  age = seq(min(mouflon$age), max(mouflon$age), length = 100)  
) %>%  
  mutate(  
    reproduction = predict(m1, type = "response", newdata = dat_predict)  
  )  
  
ggplot(mouflon, aes(x = age, y = reproduction)) +  
  geom_jitter(height = 0.01) +  
  geom_line(data = dat_predict, aes(x = age, y = reproduction))
```

# Your turn

we can do the same things with more complex models

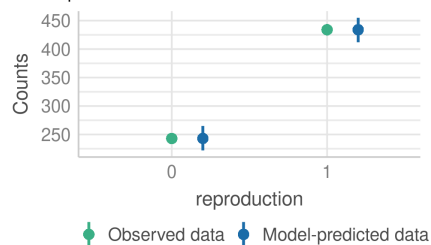
```
m2 <- glm(  
  reproduction ~ age + mass_sept + as.factor(sex_lamb) +  
    mass_gain + density + temp,  
  data = mouflon,  
  family = binomial  
)
```

# check model

```
check_model(m2)
```

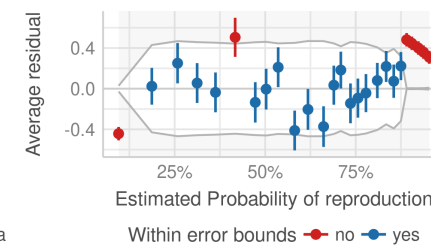
## Posterior Predictive Check

Model-predicted intervals should include observed data



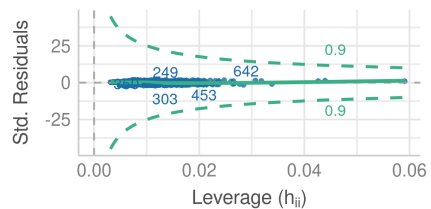
## Binned Residuals

Points should be within error bounds



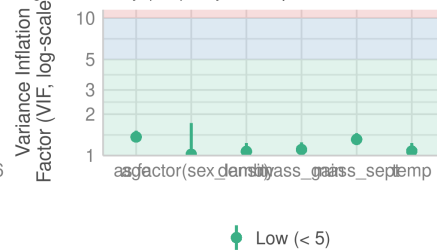
## Influential Observations

Points should be inside the contour lines



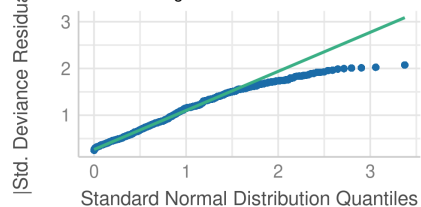
## Collinearity

High collinearity (VIF) may inflate parameter uncertainty



## Normality of Residuals

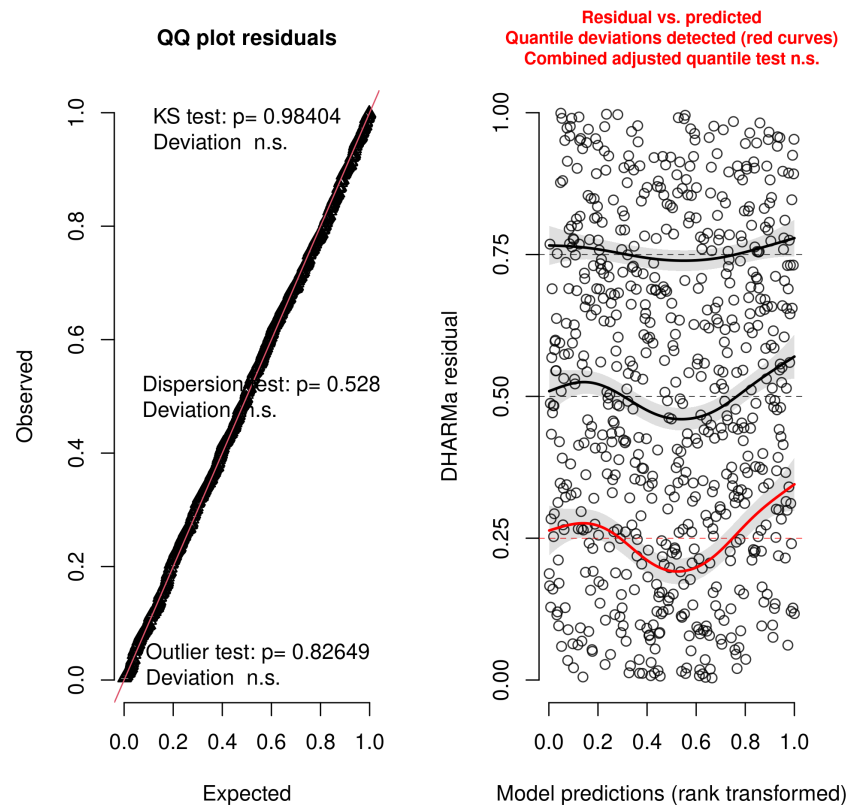
Points should fall along the line



# with DHaRMA

```
simulationOutput <- simulateResiduals(m2)  
plot(simulationOutput)
```

DHARMA residual



# Poisson regression

# Data

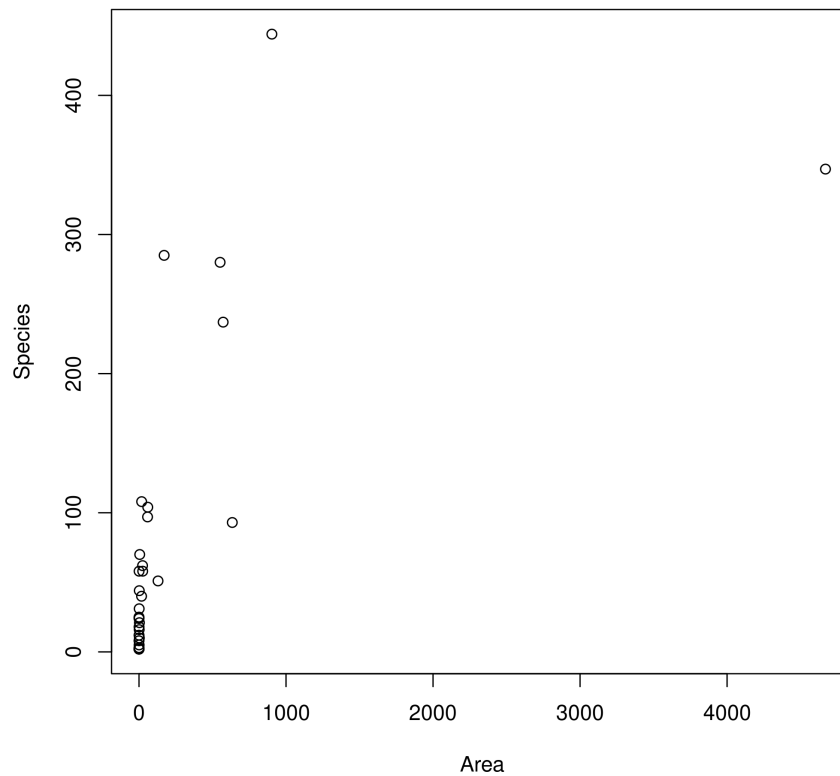
data on galapagos islands species richness

Fit 3 models:

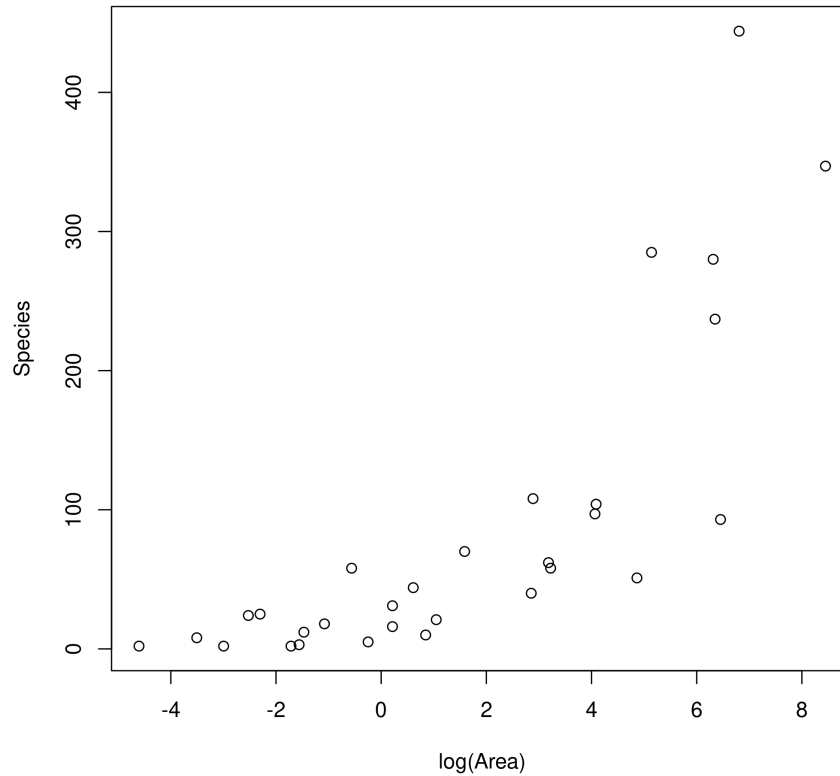
- model of total number of species
- model of proportion of endemics to total
- model of species density



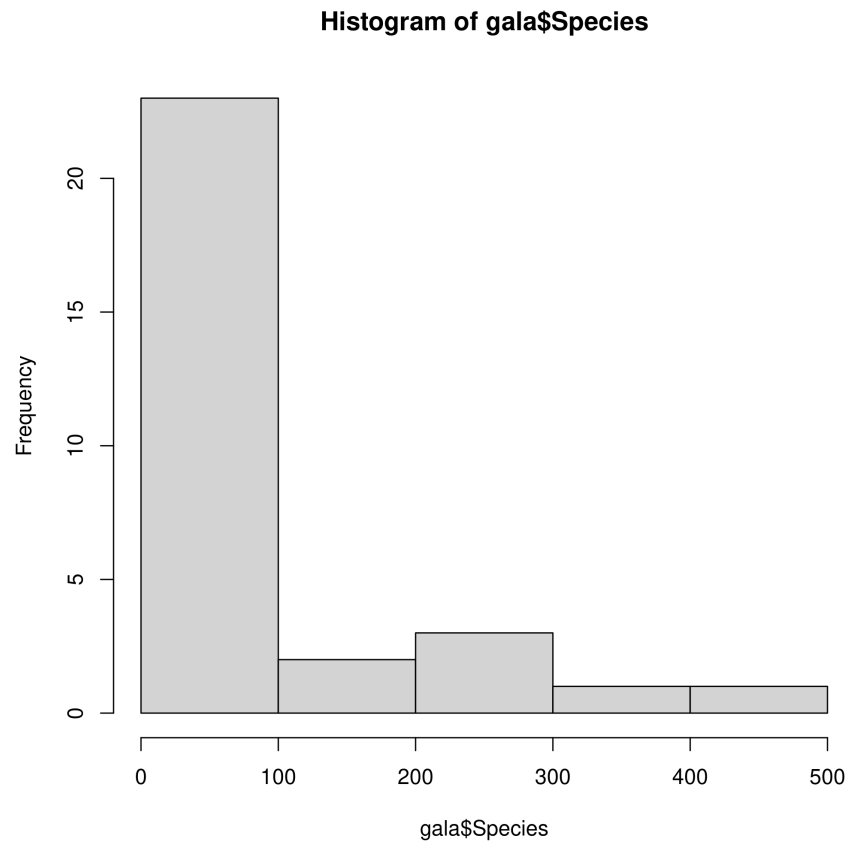
```
gala <- read.csv("data/gala.csv")  
plot(Species ~ Area, gala)
```



```
plot(Species ~ log(Area), gala)
```



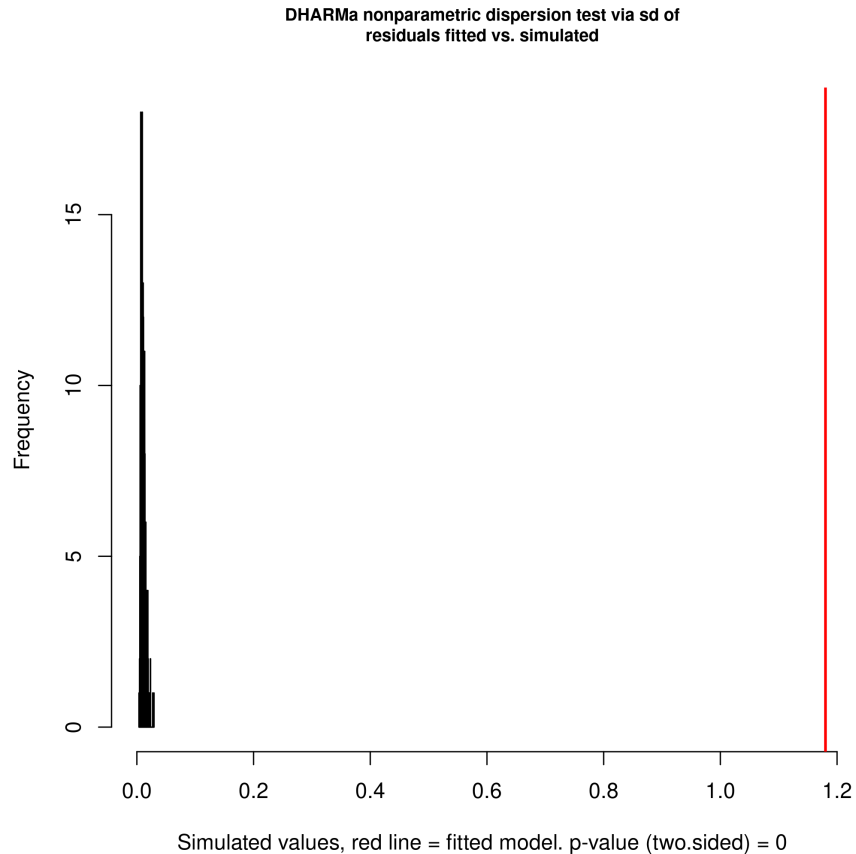
```
hist(gala$Species)
```



```
modpl <- glm(Species ~ Area + Elevation + Nearest, family = poisson, gala)
summary(modpl)
```

```
##
## Call:
## glm(formula = Species ~ Area + Elevation + Nearest, family = poisson,
##      data = gala)
##
## Coefficients:
##              Estimate Std. Error z value Pr(>|z|)
## (Intercept)  3.548e+00  3.933e-02  90.211  < 2e-16 ***
## Area        -5.529e-05  1.890e-05  -2.925  0.00344 **
## Elevation    1.588e-03  5.040e-05  31.502  < 2e-16 ***
## Nearest      5.921e-03  1.466e-03   4.039  5.38e-05 ***
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
##
## (Dispersion parameter for poisson family taken to be 1)
##
##      Null deviance: 3510.7  on 29  degrees of freedom
## Residual deviance: 1797.8  on 26  degrees of freedom
## AIC: 1966.7
##
## Number of Fisher Scoring iterations: 5
```

```
res <- simulateResiduals(modpl)
testDispersion(res)
```

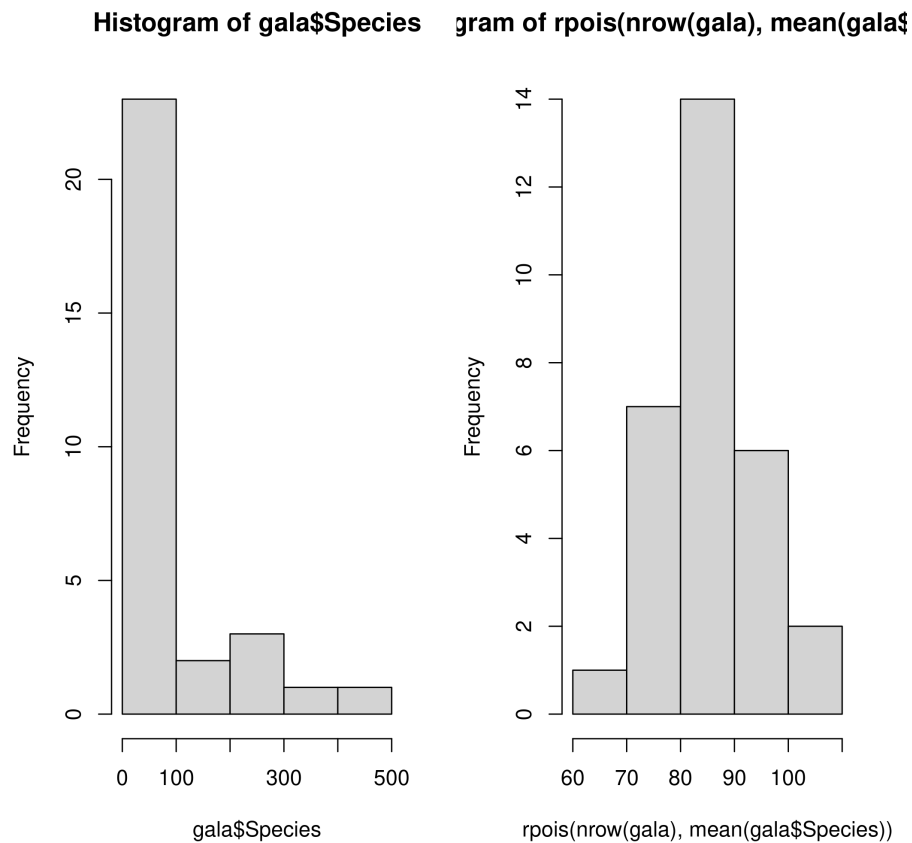


```
##
## DHARMA nonparametric dispersion test via sd of residuals fitted vs.
## simulated
##
## data: simulationOutput
## dispersion = 110.22, p-value < 2.2e-16
```

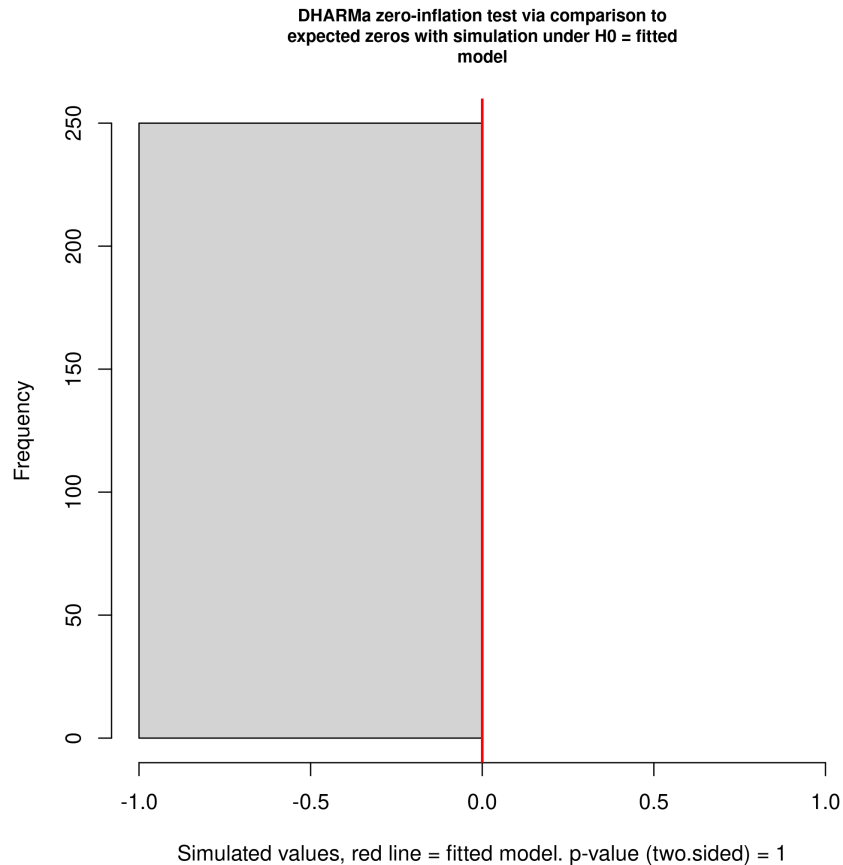
```
c(mean(gala$Species), var(gala$Species))
```

```
## [1] 85.23333 13140.73678
```

```
par(mfrow=c(1,2))  
hist(gala$Species)  
hist(rpois(nrow(gala), mean(gala$Species)))
```

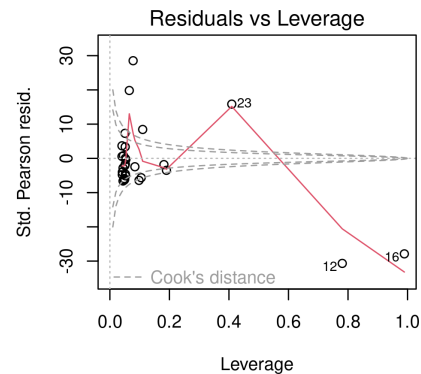
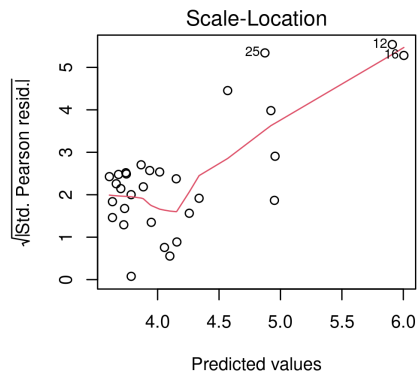
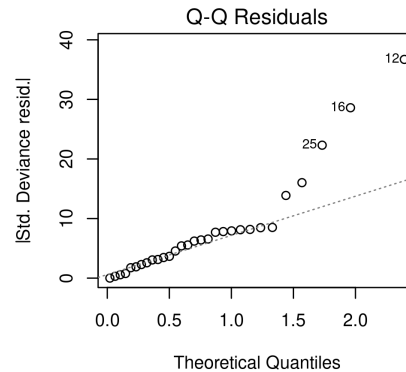
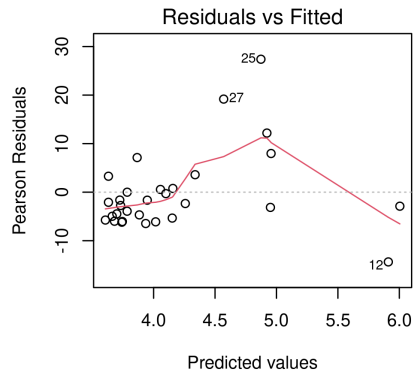


```
testZeroInflation(res)
```



```
##  
## DHARMA zero-inflation test via comparison to expected zeros with  
## simulation under  $H_0$  = fitted model  
##  
## data: simulationOutput  
## ratioObsSim = NaN, p-value = 1  
## alternative hypothesis: two.sided
```

```
par(mfrow = c(2, 2))
plot(modpl)
```





# Happy coding

