#### SURVIVAL ANALYSIS WORKSHOP

# JOOYOUNG LEE DEPARTMENT OF APPLIED STATISTICS CHUNG-ANG UNIVERSITY

SEOUL MEDICAL CENTER
JANUARY 06, 2023

#### Instructor

#### Jooyoung Lee

- Department of Applied Statistics, Chung-Ang University
- Email: jooylee@cau.ac.kr
- Slides: https://github.com/BLDACAU/survivalSMC2023

#### **OVERVIEW**

- The objective of this course is to understand statistical methods for survival analysis.
- Application of methods will be presented using the National Health Insurance Sharing Service (NHISS) big data.

#### SCHEDULE

Week 1: Introduction to Survival Analysis

Week 2: Clinical Study using NHISS I

Week 3: Introduction to Competing Risks Model

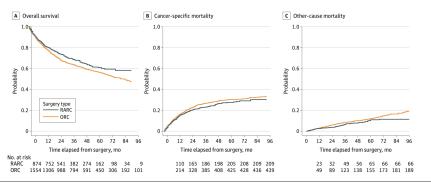
Week 4: Clinical Study using NHISS II



#### Week 1 Materials

- An introduction to censoring, truncation and examples
- Survival curves, risk set tables, and the Kaplan-Meier estimator
- Log-rank test
- Cox proportional hazard regression models

#### EXAMPLE



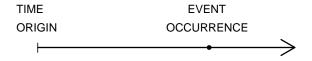
ORC indicates open radical cystectomy; RARC, robot-assisted radical cystectomy.

Figure 1: Figure from Mortezavi, et al., JAMA Network Open 2022

#### TIME-TO-EVENT DATA

Terminology: Survival time, failure time, lifetime, time-to-event data **Time-to-event data**: the time to event measured from some particular starting point.

The definition includes a **time origin**, a **time scale**, and a definition of **the event** of interest.



- Time origin: entry into study, randomization, birth
- Time scale: years, months, weeks, age, mileage
- Event: death, disease occurrence, progression of tumor, hospitalization, recovery

#### THE PURPOSE OF TIME-TO-EVENT DATA ANALYSIS

- Estimate a event time distribution of a population
- Compare the event time distributions of two or more groups
- Evaluate the effect of risk factors on the event time

#### FEATURES OF TIME-TO-EVENT DATA

Time-to-event data consist of a mixture of complete and incomplete observations.

**Censoring**: Censoring arises when the value of a response measurement (event time) is only partially known.

- Right censoring
- Left censoring
- Interval censoring

**Truncation**: Truncation is termed to describe a condition by which subjects are screened or excluded from the study population.

Left truncation

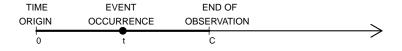
#### NOTATION

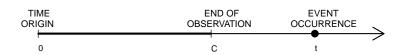
- T: Event time (survival/failure time) of an individual from a population
- *C*: Censoring time
- X: Observed time
- $\delta$ : Censoring indicator

#### RIGHT CENSORED DATA

Right censoring: The true unobserved event time is to the right of the censoring time. Right censoring may occur due to no event before the study ends, loss to follow-up or withdrawal from the study.

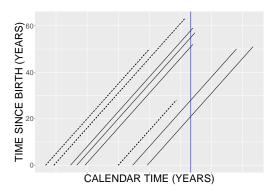
$$X = \min(T, C)$$
, and  $\delta = I(T \le C)$ 





#### LEFT TRUNCATED DATA

Left truncation arises when there are particular selection conditions for recruiting study participants.



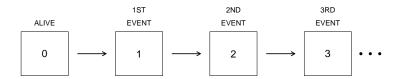
### Univariate Survival data



# library(survival) head(ovarian)

	futime	fustat	age	resid.ds	rx	ecog.ps
1	59	1	72.3315	2	1	1
2	115	1	74.4932	2	1	1
3	156	1	66.4658	2	1	2
4	421	0	53.3644	2	2	1
5	431	1	50.3397	2	1	1
6	448	0	56.4301	1	1	2

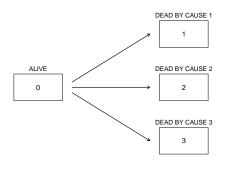
#### RECURRENT EVENT DATA



	id	start	stop	status	tstatus	enum	trt
1	1	0	122	1	1	1	1
2	2	0	122	0	1	1	1
3	3	0	3	1	1	1	1
4	3	3	88	1	2	2	1
5	3	88	122	0	2	3	1

#### Competing Risks data

Competing risks data arise when an individual is at risk of more than one mutually exclusive event.

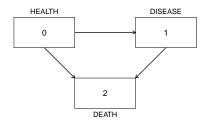


```
library(survival); data(pbc)
pbc[1:5, c(1,2,3,4)]

id time status trt
1 1 400 2 1
2 2 4500 0 1
3 3 1012 2 1
4 4 1925 2 1
5 5 1504 1 2
```

#### MULTISTATE MODEL

Multistate model is a model for a process where subjects move among a finite number of states.



```
library(mstate); data(ebmt3)
ebmt3[1:5, c(1,2,3,4, 5)]
id notice protect reference research
```

	id	prtime	prstat	${\tt rfstime}$	${\tt rfsstat}$
1	1	23	1	744	0
2	2	35	1	360	1
3	3	26	1	135	1
4	4	22	1	995	0
5	5	29	1	422	1

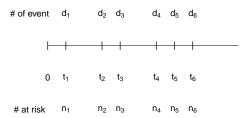
#### EVENT TIME DISTRIBUTION

Response (T): Time to event (survival/failure time) of an individual from a population.

- The survival function: S(t) = P(T > t).
- The probability density function of T: f(t) = -dS(t)/dt.
- The hazard function:  $\lambda(t) = \lim_{h \to 0^+} P(t \le T < t + h | T \ge t)/h$ .
- The cumulative hazard function:  $\Lambda(t) = \int_0^t \lambda(s) ds$ .

#### Risk Table

- $\{(X_i, \delta_i), i = 1, ..., n\}$ : a sample of n right-censored survival data.
- $k(k \le n)$  distinct event times  $t_1 < t_2 < \cdots < t_k$ .
- $n_j$ : the number of individuals at risk at  $t_j$ . The number of individuals with no event and uncensored "just before" time t.
- d<sub>j</sub>: the number of individuals who experience the event at t<sub>j</sub>.



A risk table displays the number at risk  $n_i$  at each time point.

# Estimating Survival Curves: the Kaplan-Meier estimator

The Kaplan-Meier estimator  $\widehat{S}(t)$  is

$$\widehat{S}(t) = \prod_{j:t_j \le t} \left(1 - \frac{d_j}{n_j}\right)$$

# Comparing the Survival of Groups: The Log-Rank Test

We want to test whether the survival curves are the same for two groups.

- $O_j = d_{1j}$ : the observed number of events for group 1 at time  $t_j$
- $E_j = n_{1j}d_j/n_j$ : the expected number of events for group 1 at time  $t_j$
- $V_j=rac{n_{1j}n_{2j}d_j(n_j-d_j)}{n_j^2(n_j-1)}$ : the variance of the observed number of events for group 1 at time  $t_j$

The logrank test statistic is

$$Z = rac{\sum_{j=1}^{k} (O_j - E_j)}{\sqrt{\sum_{j=1}^{k} V_j}} \sim N(0, 1)$$
 under  $H_0$ 

# Example

0

1

3

4

8

At 
$$t = 3$$

Group 0

Group 1

Total

At 
$$t = 8$$

Group 0	1	1	2
Group 1	0	3	3
Total	1	4	5

At t=9

#### EXAMPLE

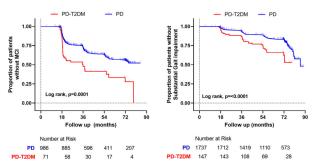


FIG. 3. Timeline for the development of mild cognitive impairment (MCI) and substantial gait impairment, comparing Parkinson's disease (PD) cases with and without type 2 diabetes mellitus (T2DM). Kaplan-Meier curves show the significantly shorter time to develop both of these complications in patients with T2DM. (Color figure can be viewed at wileyonlinelibrary.com)

Figure 2: Figure from Athauda, et al., Movement Disorder 2022

#### Cox Proportional Hazard Regression Model

The Cox (1972) proposed the proportional hzard regression model:

$$\lambda(t|\mathbf{z}) = \lambda_0(t) \exp(\mathbf{z}'\boldsymbol{\beta})$$

- A vector of covariates  $\mathbf{z} = (z_1, \dots, z_p)^T$
- A baseline hazard function  $\lambda_0(t)$ .

#### Hazard Ratio

- The estimated hazard ratio (HR) =  $\exp(\hat{eta}_j)$
- A 95% CI for HR =  $\exp(\hat{eta}_j \pm 1.96s.e.(\hat{eta}_j))$
- p-value =  $P(|Z| > |z_0|)$ , where  $z_0 = \hat{\beta}_j/s.e.(\hat{\beta}_j)$ ,  $Z \sim N(0,1)$  under  $H_0: \beta_j = 0$ .

# Stratified Cox Regression Models

- Often times we are interested in assessing the effects of exposures while adjusting for other factors in a regression analysis.
- We can stratify on the factor we want to control for if we are not interested in the hazard ratio for this factor to avoid the proportional hazard assumption.

A stratified Cox regression model is formed by

$$\lambda_k(t|\mathbf{z}_k) = \lambda_{k0}(t) \exp(\mathbf{z}_k'\boldsymbol{\beta}),$$

where  $\lambda_{k0}(t)$  is the baseline hazard for stratum k and  $\beta$  represents common covariate effects.

## Extended Cox Models: Time-dependent Covariates

Consider the study for transplant on survival.

- The Cox model would compare the survival distributions between those without a transplant (ever) to those with a transplant.
- A transplant status was determined at the end of the study and this value characterizes a subject for the entire follow-up period.
- Patients who died early had less time available to get transplants.
- Therefore, the survival estimates are overestimated for the patients who had transplants, and we call this problem as "immortal bias".

# Extended Cox Models: Time-dependent Covariates

- As the status of transplant changes over time, we consider a model with a time-dependent indicator of whether a patient had a transplant at each point.
- We compare the risk of an event between the patients who had already transplants and who had not yet transplants at each event time.
- We recalculate the number at risk for each group at each event time.

A path of the covariate upto t:  $\bar{\mathbf{z}}_i(t) = \{\mathbf{z}_i(u), 0 \le u \le t\}$ .

We model

$$\lambda(t|\bar{\mathbf{z}}(t)) = \lambda_0(t) \exp(\mathbf{z}(t)'\beta).$$

#### Reference

- Mortezavi, A., Crippa, A., Kotopouli, M. I., Akre, O., Wiklund, P., & Hosseini, A. (2022). Association of Open vs Robot-Assisted Radical Cystectomy With Mortality and Perioperative Outcomes Among Patients With Bladder Cancer in Sweden. JAMA network open, 5(4), e228959-e228959.
- Athauda, D., Evans, J., Wernick, A., Virdi, G., Choi, M. L., Lawton, M., ... & Gandhi, S. (2022). The impact of type 2 diabetes in Parkinson's disease. Movement Disorders, 37(8), 1612-1623.