

17s1: COMP9417 Machine Learning and Data Mining

Lectures: Linear Models for Regression

Topic: Questions from lecture topics

Version: with answers

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Introduction

Some questions and exercises from the second week's course lectures, covering aspects of learning linear models (models “linear in the parameters”) for regression, i.e., numeric prediction, tasks.

Question 1 A *univariate linear regression model* is a linear equation $y = a + bx$. Learning such a model requires fitting it to a sample of training data so as to minimize the error function $E(a, b) = \sum_{i=1}^n (y_i - (a + bx_i))^2$. To find the best parameters a and b that minimize this error function we need to find the error *gradients* $\frac{\partial E(a, b)}{\partial a}$ and $\frac{\partial E(a, b)}{\partial b}$. So derive these expressions by taking partial derivatives, divide by n (the number of (x, y) data points in the training sample) set them to zero, and solve for a and b .

Answer

In the following note that \bar{z} denotes the mean of z .

$$\begin{aligned}
\frac{\partial E(a,b)}{\partial a} &= \frac{\partial}{\partial a} \sum_{i=1}^n (y_i - (a + bx_i))^2 \\
&= \sum_{i=1}^n 2(y_i - (a + bx_i)) \left(\frac{\partial}{\partial a} y_i - \frac{\partial}{\partial a} (a + bx_i) \right) \\
&= \sum_{i=1}^n 2(y_i - (a + bx_i)) - \left(\frac{\partial}{\partial a} a + \frac{\partial}{\partial a} bx_i \right) \\
&= \sum_{i=1}^n 2(y_i - (a + bx_i)) - 1 \\
&= -2 \sum_{i=1}^n (y_i - (a + bx_i)) \\
0 &= \frac{-2}{n} \left(\sum_{i=1}^n (y_i - (a + bx_i)) \right) \\
0 &= \frac{1}{n} \left(\sum_{i=1}^n (y_i - (a + bx_i)) \right) \\
a &= \frac{1}{n} \sum_{i=1}^n y_i - b \frac{1}{n} \sum_{i=1}^n x_i \\
a &= \bar{y} - b\bar{x}
\end{aligned}$$

Similarly, the expression for b is as follows (the full expansion once the derivative is found can be skipped).

$$\begin{aligned}
\frac{\partial E(a,b)}{\partial b} &= \frac{\partial}{\partial b} \sum_{i=1}^n (y_i - (a + bx_i))^2 \\
&= \sum_{i=1}^n 2(y_i - (a + bx_i)) \left(\frac{\partial}{\partial b} y_i - \frac{\partial}{\partial b} (a + bx_i) \right) \\
&= \sum_{i=1}^n 2(y_i - (a + bx_i)) - \left(\frac{\partial}{\partial b} a + \frac{\partial}{\partial b} bx_i \right) \\
&= -2 \sum_{i=1}^n (y_i - (a + bx_i)) x_i \\
b &= \frac{\frac{1}{n} \sum_{i=1}^n x_i y_i - \frac{1}{n} \sum_{i=1}^n y_i \frac{1}{n} \sum_{i=1}^n x_i}{\frac{1}{n} \sum_{i=1}^n x_i^2 - \left(\frac{1}{n} \sum_{i=1}^n x_i \right)^2}
\end{aligned}$$

Looking at this, the denominator is recognisable as the variance of x , and the whole expression turns out to be: $\frac{\text{covariance of } x \text{ and } y}{\text{variance of } x}$.

Question 2 A linear regression model is represented by the linear equation $y = a + bx$. Show that the mean point (\bar{x}, \bar{y}) must line on the regression line.

Answer

In Question 1 we derived the expression $a = \bar{y} - b\bar{x}$. Suppose we ask the question: what is the value of the output y for the linear regression model when the input is \bar{x} , i.e., the mean of x ?

We formulate the regression model like this, then substitute the expression for a :

$$\begin{aligned} y &= a + b\bar{x} \\ &= \bar{y} - b\bar{x} + b\bar{x} \\ &= \bar{y} \end{aligned}$$

which completes the proof.

Question 3 *Mean Square Error*, or MSE, of an estimator such as a regression model can be decomposed. Show that $\text{MSE} = (\text{variance}) + (\text{bias})^2$.

Background Suppose we have a repeatable setting in which a regression model is trained and has its predictions tested to determine error. Let f (standing for $f(x)$) be the actual value and y (standing for \hat{y}) be the predicted value. Then, to simplify notation, the MSE can be written as follows:

$$\text{MSE} = E[f - y]^2$$

You will need some properties of the expected value, or expectation operator $E[\]$ (information on these properties is available on the web, e.g., http://en.wikipedia.org/wiki/Expected_value).

The key point about the expectation is that it provides a way to describe characteristics of the distribution of possible values of random variables (or functions of random variables). The most well-known is the mean μ , or expected value, of a random variable, which is defined as the sum of *all* possible values of the random variable, weighted by the probability of that value occurring:

$$E[X] = \sum_{i=1}^N x_i p(x_i) = \mu$$

Analogously, the expected value of a function of one (or more) random variable(s) is the sum of the possible values of the function for each outcome, weighted by the probability of that outcome occurring. For example, $E[Y] = E[f(X)]$, for the function $Y = f(X)$:

$$E[f(X)] = \sum_{i=1}^N f(x_i) p(x_i)$$

Here are some properties following from this definition.

$$E[X + c] = E[X] + c \quad \text{for } c \text{ a constant}$$

$$E[X + Y] = E[X] + E[Y] \quad \text{where } X \text{ and } Y \text{ are random variables defined on the same probability space}$$

$$E[aX] = aE[X] \quad \text{for } a \text{ not a random variable}$$

These can be used, for example, to decompose the expectation of a linear model $E[a+bX] = a+bE[X]$.
Some further properties of the expectation you may need to use:

$$E[fE[y]] = fE[y] \quad \text{for } f \text{ a deterministic function, } y \text{ a random variable}$$

$$E[E[y]] = E[y]$$

$$E[E[y]^2] = E[y]^2$$

$$E[yf] = fE[y] \quad \text{again, since } f \text{ is deterministic, and } y \text{ a random variable}$$

$$E[yE[y]] = E[y]^2 \quad \text{by the definition of expectation.}$$

Answer

Suppose we have a repeatable setting in which a regression model is trained and has its predictions tested to determine error. Let f (standing for $f(x)$) be the actual value and y (standing for \hat{y}) be the predicted value. Then the MSE can be decomposed as follows:

$$\begin{aligned}
\text{MSE} &= E[f - y]^2 \\
&= E[(f - E[y] + E[y] - y)^2] \\
&= E[(f - E[y] + E[y] - y)(f - E[y] + E[y] - y)] \\
&= E[f^2 - fE[y] + fE[y] - fy \\
&\quad - fE[y] + E[y]^2 - E[y]^2 + yE[y] \\
&\quad + fE[y] - E[y]^2 + E[y]^2 - yE[y] \\
&\quad - yf + yE[y] - yE[y] + y^2] \\
&= E[y]^2 - 2fE[y] + f^2 + E[y^2 - 2yE[y] + E[y]^2] + \\
&\quad E[fE[y] + fE[y] - fy - fy - E[y]^2 - E[y]^2 + yE[y] + yE[y]]
\end{aligned}$$

We can see that this derivation starts with a small “trick”, by adding and subtracting the expectation $E[y]$. Then follows the rather tedious expansion of the terms. By re-arranging we find the expressions for the **Bias**² and **Variance**.

$$\begin{aligned}
\text{Variance} &= E[(y - E[y])^2] \\
&= E[y^2 - 2yE[y] + E[y]^2]
\end{aligned}$$

and

$$\begin{aligned}
\text{Bias}^2 &= (E[y] - f)^2 \\
&= E[y]^2 - 2fE[y] + f^2
\end{aligned}$$

However, we still have something left over at the end:

$$E[2fE[y] - 2fy - 2E[y]^2 + 2yE[y]]$$

which can be written as:

$$2E[fE[y] - fy - E[y]^2 + yE[y]]$$

and then:

$$2(E[fE[y]] - E[fy] - E[E[y]^2] + E[yE[y]])$$

How do we get rid of this ? If a term in the expectation is deterministic it can be brought out. So we get $2E[fE[y] - fy - E[y]^2]$, which expands to

$$2[E[fE[y]] - E[E[y]^2] - E[yf] + E[yE[y]]]$$

Now $E[fE[y]] = fE[y]$, since f is deterministic. Also, since $E[E[z]] = E[z]$ we get $E[E[y]^2] = E[y]^2$. Lastly, $E[yf] = fE[y]$ since f is deterministic, and $E[yE[y]] = E[y]^2$, by the definition of expectation. Thus

$$\begin{aligned}
&5 \\
&2[fE[y] - fE[y] - E[y]^2 + E[y]^2] \\
&= 2(0) \\
&= 0
\end{aligned}$$
