Exercise_answers

Franky Zhang

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Part 1

Exercise 4.25

```
# n = 5
approx = c()
for (i in 1:5) {
  approx[i] = (i - (1/3))/(5 + (1/3))
}
iter = 10000
U_5 = matrix(NA, ncol = 5, nrow = iter)
for (n in 1:iter){
  series = runif(5, \min = 0, \max = 1)
  U_5[n, 1] = sort(series)[1]
  U_5[n, 2] = sort(series)[2]
  U_5[n, 3] = sort(series)[3]
  U_5[n, 4] = sort(series)[4]
 U_5[n, 5] = sort(series)[5]
comparison1 = data.frame(rbind(approx, apply(U_5, 2, median)))
rownames(comparison1) = c("approximation", "calculated value")
colnames(comparison1) = c("i = 1", "i = 2", "i = 3", "i = 4", "i = 5")
\# n = 10
U_10 = matrix(NA, ncol = 10, nrow = iter)
for (n in 1:iter){
  series = runif(10, min = 0, max = 1)
  U_10[n, 1] = sort(series)[1]
  U_10[n, 2] = sort(series)[2]
  U_10[n, 3] = sort(series)[3]
  U_10[n, 4] = sort(series)[4]
  U_10[n, 5] = sort(series)[5]
  U_10[n, 6] = sort(series)[6]
  U_10[n, 7] = sort(series)[7]
  U_10[n, 8] = sort(series)[8]
  U_10[n, 9] = sort(series)[9]
  U_10[n, 10] = sort(series)[10]
approx = c()
```

comparison1

```
## i = 1 i = 2 i = 3 i = 4 i = 5

## approximation 0.1250000 0.3125000 0.5000000 0.6875000 0.8750000

## calculated value 0.1344157 0.3166064 0.5006924 0.6898505 0.8721464
```

comparison2

```
## i = 1 i = 2 i = 3 i = 4 i = 5 i = 6 ## approximation 0.06451613 0.1612903 0.2580645 0.3548387 0.4516129 0.5483871 ## calculated value 0.06801226 0.1642092 0.2587907 0.3562169 0.4510261 0.5481125 ## i = 7 i = 8 i = 9 i = 10 ## approximation 0.6451613 0.7419355 0.8387097 0.9354839 ## calculated value 0.6470331 0.7441768 0.8384631 0.9337684
```

I've done the simulation for order statistic. From the results, the approximation equation is true.

Exercise 4.27

(a)

```
print("summary statistic of Jan 1940:")
```

[1] "summary statistic of Jan 1940:"

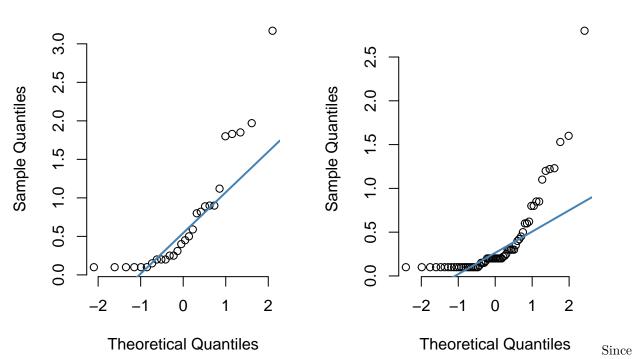
```
summary(Jan_1940)
```

```
## Min. 1st Qu. Median Mean 3rd Qu. Max.
## 0.1000 0.1875 0.4250 0.7196 0.9000 3.1700
```

```
cat("\n")
print("summary statistic of Jul 1940:")
## [1] "summary statistic of Jul 1940:"
summary(Jul_1940)
##
      Min. 1st Qu.
                    Median
                              Mean 3rd Qu.
                                              Max.
    0.1000 0.1000 0.2000 0.3931 0.4275
                                           2.8000
(b)
par(mfrow = c(1, 2))
qqnorm(Jan_1940, pch = 1, frame = FALSE)
qqline(Jan_1940, col = "steelblue", lwd = 2)
qqnorm(Jul_1940, pch = 1, frame = FALSE)
qqline(Jul_1940, col = "steelblue", lwd = 2)
```

Normal Q-Q Plot

Normal Q-Q Plot



the observations are relatively far away from the normal distribution line, normal distribution is no longer be considered. Also, the observations are continuously distributed, consider about the gamma distribution.

(c)

```
fit_Jan <- fitdist(Jan_1940, distr = "gamma", method = "mle")</pre>
fit_Jul <- fitdist(Jul_1940, distr = "gamma", method = "mle")</pre>
mean_Jan <- fit_Jan$estimate[1]/fit_Jan$estimate[2]</pre>
mean_Jul <- fit_Jul$estimate[1]/fit_Jul$estimate[2]</pre>
summary(fit_Jan)
## Fitting of the distribution ' gamma ' by maximum likelihood
## Parameters :
         estimate Std. Error
## shape 1.056222 0.2497495
## rate 1.467650 0.4396202
## Loglikelihood: -18.7616
                              AIC: 41.5232
                                               BIC: 44.18761
## Correlation matrix:
##
             shape
                        rate
## shape 1.0000000 0.7893943
## rate 0.7893943 1.0000000
summary(fit_Jul)
## Fitting of the distribution ' gamma ' by maximum likelihood
## Parameters :
         estimate Std. Error
## shape 1.196419 0.1891196
## rate 3.043403 0.5936302
## Loglikelihood: -3.634886
                               AIC: 11.26977 BIC: 15.58754
## Correlation matrix:
##
             shape
## shape 1.0000000 0.8103948
## rate 0.8103948 1.0000000
```

Answer: For Jan 1940, the maximum log likelihood is -19.8, the estimated value of shape parameter is 1.06 with sd equal to 0.25 and the rate parameter is 1.47 with sd equal to 0.44. The mean parameter is 0.72.

For Jul 1940, the maximum log likelihood is -3.63, the estimated value of shape parameter is 1.196 with sd equal to 0.19 and the rate parameter is 3.04 with sd equal to 0.59. The mean parameter is 0.393.

(d)

```
par(mfrow = c(1, 2))
library(EnvStats)

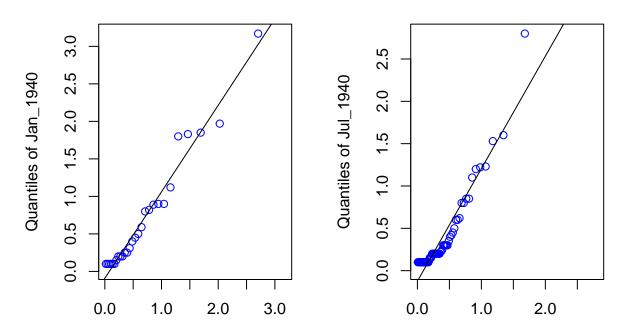
##
## Attaching package: 'EnvStats'

## The following object is masked from 'package:car':
##
## qqPlot

## The following object is masked from 'package:MASS':
##
## boxcox
```

Gamma Q-Q Plot for Jan_1940

Gamma Q-Q Plot for Jul_1940



Quantiles of Gamma(shape = 1.1, scale = (Quantiles of Gamma(shape = 1.2, scale = (

```
detach("package:EnvStats", unload = TRUE)
```

Answer:

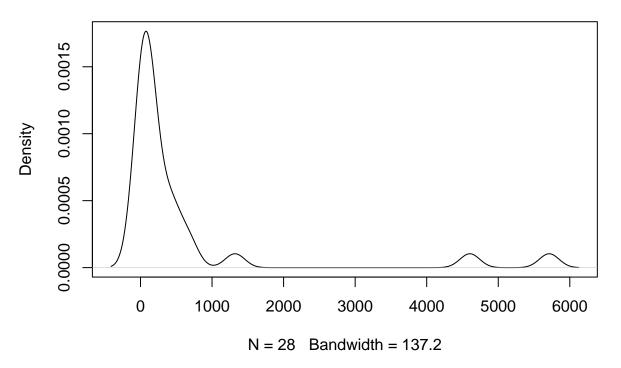
The observations from Jan 1940 and Jul 1940 fit really will with estimated gamma distribution. Thus the adequency of gamma model is good.

Exercise 4.39

step 1

```
dat = c(0.4, 1.0, 1.9, 3.0, 5.5, 8.1, 12.1, 25.6,
 115.0,
             119.5,
                       154.5,
                               157.0,
                                         175.0,
 419.0,
             423.0,
                       440.0,
                               655.0,
                                         680.0,
  50.0,
              56.0,
                        70.0,
                               115.0,
 179.0,
             180.0,
                       406.0,
1320.0,
            4603.0,
                      5712.0)
plot(density(dat))
```

density.default(x = dat)



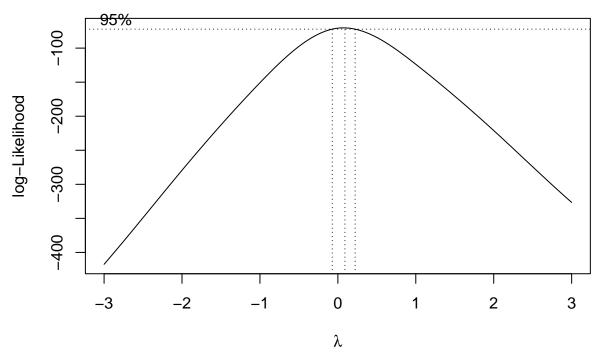
```
shapiro.test(dat)
```

```
##
## Shapiro-Wilk normality test
##
## data: dat
## W = 0.45173, p-value = 3.763e-09
```

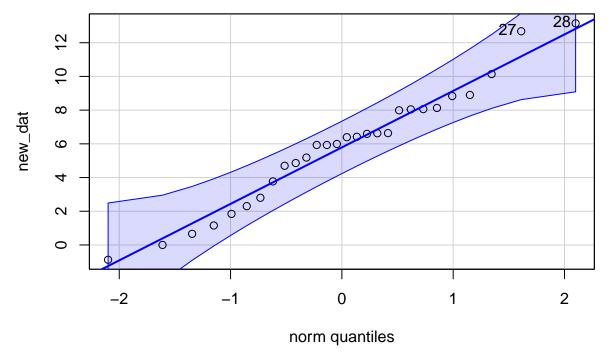
Do Shaprio test to raw data, the p-value is 3.763e-09, which is far below the significance level. The result rejects the null hypothesis of normality, thus it is necessary to conduct Box Cox transformation.

step 2

```
full_model = lm(dat~1)
library(MASS)
bc = boxcox(full_model, lambda = seq(-3, 3))
```



```
lambda = bc$x[which(bc$y == max(bc$y))]
# the best transformation lambda = 0.09090909
new_dat = ((dat^lambda-1)/lambda)
qqPlot(new_dat, dist = "norm")
```



[1] 28 27

Answer:

The best transformation lambda = 0.09090909, and after transformation, all observation points are in the normal distribution confidence interval.

step 3

```
# Conduct Shaprio test again
shapiro.test(new_dat)

##
## Shapiro-Wilk normality test
##
## data: new_dat
## W = 0.9724, p-value = 0.6462
```

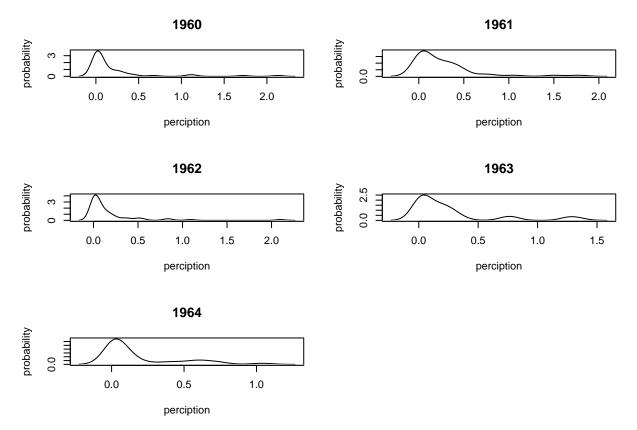
The p-value of Shaprio test = 0.6462, which is far higher than the significance level, thus the test passes.

Part 2

Data preparation

```
data = readxl::read_xlsx("Illinois_rain_1960-1964.xlsx", col_names = TRUE)
percip_1960 = data$^1960^
percip_1961 = data$^1961^
percip_1962 = data$^1962^
percip_1963 = data$^1963^
percip_1964 = data$^1964^
percip_1960 = percip_1960[!is.na(percip_1960)]
percip_1961 = percip_1961[!is.na(percip_1961)]
percip_1962 = percip_1962[!is.na(percip_1962)]
percip_1963 = percip_1963[!is.na(percip_1963)]
percip_1964 = percip_1964[!is.na(percip_1964)]
percip = list()
percip[[1]] = percip_1960; percip[[2]] = percip_1961; percip[[3]] = percip_1962; percip[[4]] = percip_1
percip[[5]] = percip_1964
names(percip) = c("1960", "1961", "1962", "1963", "1964")
```

Distribution check



For all 5 years, the density plots are not symmetric and show a heavy tail pattern. Thus we decide to apply gamma distribution.

Fit gamma distribution

```
percip_param <- matrix(NA, ncol = 4, nrow = 5)</pre>
colnames(percip_param) <- c("shape", "rate", "shape_sd", "rate_sd")</pre>
for (i in 1:5) {
  fit <- fitdist(percip[[i]], distr = "gamma", method = "mle")</pre>
  percip_param[i, 1] <- as.numeric(fit$estimate)[1]</pre>
  percip_param[i, 2] <- as.numeric(fit$estimate)[2]</pre>
  percip_param[i, 3] <- as.numeric(fit$sd)[1]</pre>
  percip_param[i, 4] <- as.numeric(fit$sd)[2]</pre>
rownames(percip_param) <- names(percip)</pre>
percip_param <- data.frame(percip_param)</pre>
percip_param <- percip_param %>% mutate(mean = shape/rate,
                                            per_shape_sd = shape_sd/shape,
                                            per_rate_sd = rate_sd/rate)
percip_param
##
             shape
                                                          mean per_shape_sd per_rate_sd
                        rate
                                shape_sd
                                            rate_sd
```

```
## shape rate shape_sd rate_sd mean per_shape_sd per_rate_sd ## 1960 0.3542986 1.608142 0.05789003 0.4702240 0.2203155 0.1633933 0.2924020 ## 1961 0.5783901 2.103720 0.09880101 0.5371665 0.2749369 0.1708207 0.2553413 ## 1962 0.4130575 2.235782 0.06332713 0.5775788 0.1847485 0.1533131 0.2583341
```

```
## 1963 0.5283565 2.013200 0.10192874 0.5984650 0.2624461 0.1929166 0.2972705 ## 1964 0.4454876 2.380694 0.08347842 0.7306273 0.1871251 0.1873866 0.3068967
```

The previous summary table is estimated parameters and means for 5 years. Estimated mean values for year 1962 and 1964 are obviously lower than other year (around 0.185). The standard errors for shape parameters are around 0.05~0.09 and for rate parameters are around 0.5~0.6. For year 1964 and 1963, the estimations are relatively not condident.

Wet years and dry years

The highest mean value appears in 1961 around 0.27, which can be considered as a wet year. The lowest mean value appears in 1962 approximately 0.18 thus this year is relatively dry. According to the 'Water Resource Research' by Huff in 1964 (Table 5, 6), the rainfall perception seems no relationship with storm moving.

generalizability

Since we only have five years data, the conclusion is not solid enough. On the one hand, the distribution estimations are with great variability. On the other hand, the strom analysis contains 10 years data only.

About further direction, I will look into more perception and storm papers to come out more generalizable conclusion.