

Lecture 5: Random Variable, Part III

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EE210: Probability and Introductory Random Processes KAIST EE

MONTH DAY, 2021

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Roadmap



- o Famous discrete random variables used in the community
 - Bernoulli, Uniform, Binomial, Geometric, Poisson, etc.
- o Summarizing a random variable: Expectation and Variance
- o Functions of a single random variable, Functions of multiple random variables
- o Conditioning for random variables, Independence for random variables
- o Continuous random variables
 - Normal, Uniform, Exponential, etc.
- o Bayes' rule for random variables
- (Derived) Distribution of Y = g(X) or Z = g(X, Y)
- Quantifying the degree of dependence between two rvs.
- Conditional expectation/variance
- o (Random) Sum of random variables

- Given the PDF of X, What is the PDF of Y = g(X)?
- Wait! Didn't we cover this topic? No. We covered just $\mathbb{E}[g(X)]$.
- Examples: Y = X, Y = X + 1, $Y = X^2$, etc.
- What are easy or difficult cases?
- Easy cases
 - Discrete
 - Linear: Y = aX + b

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Discrete Case



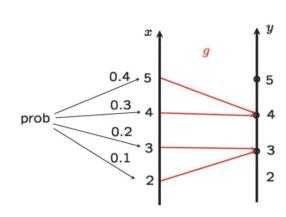
• Take all values of x such that g(x) = y, i.e.,

$$p_Y(y) = \mathbb{P}(g(X) = y)$$

$$= \sum_{x:g(x)=y} p_X(x)$$

$$p_Y(3) = p_X(2) + p_X(3) = 0.1 + 0.2 = 0.3$$

 $p_Y(4) = p_X(4) + p_X(5) = 0.3 + 0.4 = 0.7$





If
$$a > 0$$
, $F_Y(y) = \mathbb{P}(aX + b \le y) = \mathbb{P}(X \le \frac{y - b}{a}) = F_X(\frac{y - b}{a})$

$$\to f_Y(y) = \frac{1}{a} f_X\left(\frac{y - b}{a}\right)$$

If
$$a < 0$$
, $F_Y(y) = \mathbb{P}(aX + b \le y) = \mathbb{P}(X > \frac{y - b}{a}) = 1 - F_X(\frac{y - b}{a})$

$$\to f_Y(y) = -\frac{1}{a}f_X\left(\frac{y - b}{a}\right)$$

Therefore, $f_Y(y) = \frac{1}{|a|} f_X(\frac{y-b}{a})$

Special case. X is normal. Then, Y is also normal, i.e., $Y \sim \textit{N}(a\mu + b, a^2\sigma^2)$

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Generally, Y = g(X)

Step 1. Find the CDF of Y:

$$F_Y(y) = \mathbb{P}(Y \le y) = \mathbb{P}(g(X) \le y)$$

Step 2. Differentiate: $f_Y(y) = \frac{dF_Y}{dy}(y)$

** When Y = g(X) is monotonic, a general formula can be drawn (see the textbook at pp 207)

Ex1. $X \sim uniform[0, 1]$. $Y = \sqrt{X}$.

$$F_Y(y) = \mathbb{P}(\sqrt{X} \le y) = \mathbb{P}(X \le y^2) = y^2$$

$$f_Y(y) = 2y, \quad 0 \le y \le 1$$

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Ex2. $X \sim uniform[0,2]$. $Y = X^3$.

$$F_Y(y) = \mathbb{P}(X^3 \le y) = \mathbb{P}(X \le \sqrt[3]{y}) = \frac{1}{2}y^{1/3}$$

 $f_Y(y) = \frac{1}{6}y^{-2/3}, \quad 0 \le y \le 8$

Ex3. X with $f_X(x)$. $Y = X^2$.

$$F_Y(y) = \mathbb{P}(X^2 \le y) = \mathbb{P}(-\sqrt{y} \le X \le \sqrt{y})$$

$$= F_X(\sqrt{y}) - F_X(-\sqrt{y})$$

$$f_Y(y) = \frac{1}{2\sqrt{y}} f_X(\sqrt{y}) + \frac{1}{2\sqrt{y}} f_X(-\sqrt{y}), \quad y \ge 0$$

Functions of multiple rvs: Y = g(X, Y) (1)



Basically, follow two step approach: (i) CDF and (ii) differentiate.

Ex1. $X, Y \sim uniform[0, 1]$, and $X \perp \!\!\!\perp Y$. Z = max(X, Y).

*
$$\mathbb{P}(X \le z) = \mathbb{P}(Y \le z) = z, \ z \in [0, 1].$$

$$F_Z(z) = \mathbb{P}(\max(X, Y) \le z) = \mathbb{P}(X \le z, Y \le z)$$
 $= \mathbb{P}(X \le z)\mathbb{P}(Y \le z) = z^2$
 $f_Z(z) = \begin{cases} 2z, & \text{if } 0 \le z \le 1 \\ 0, & \text{otherwise} \end{cases}$

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Functions of multiple rvs: Y = g(X, Y) (2)



Basically, follows two step approach: (i) CDF and (ii) differentiate.

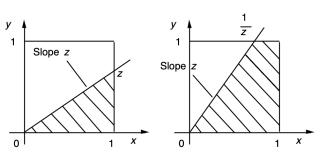
Ex2. $X, Y \sim uniform[0, 1]$, and $X \perp \!\!\! \perp Y$. Z = Y/X.

$$F_Z(z) = \mathbb{P}(Y/X \le z)$$

$$= \begin{cases} z/2, & 0 \le z \le 1 \\ 1 - 1/2z, & z > 1 \\ 0, & \text{otherwise} \end{cases}$$

$$f_Z(z) = egin{cases} 1/2, & 0 \leq z \leq 1 \ 1/(2z^2), & z > 1 \ 0, & ext{otherwise} \end{cases}$$

- Depending on the value of z, two cases need to be considered separately.



Functions of multiple rvs: Z = X + Y, $X \perp \!\!\!\perp Y$

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- A very basic case with many applications
- Assume that $X, Y \in \mathbb{Z}$

$$p_{Z}(z) = \mathbb{P}(X + Y = z)$$

$$= \sum_{\{(x,y): x+y=z\}} \mathbb{P}(X = x, Y = y)$$

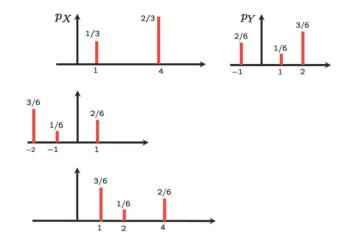
$$= \sum_{x} \mathbb{P}(X = x, Y = z - x)$$

$$= \sum_{x} \mathbb{P}(X = x) \mathbb{P}(Y = z - x)$$

$$= \sum_{x} p_{X}(x) p_{Y}(z - x)$$

- $p_Z(z)$ is called **convolution** of the PMFs of X and Y.

- Interpretation (for a given z)
- (i) Flip (horizontally) $p_Y(y)$ ($p_Y(-x)$)
- (ii) Put it underneath $p_X(x)$ $(p_Y(-x+z))$



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Y = X + Y, $X \perp \!\!\!\perp Y$: Continuous

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• Same logic as the discrete case

$$f_Z(z) = \int_{-\infty}^{\infty} f_X(x) f_Y(z - x) dx$$

- Very special, but useful case
 - X and Y are normal.

Sum of two independent normal rvs

$$X \sim \textit{N}(\mu_{x}, \sigma_{x}^{2})$$
 and $Y \sim \textit{N}(\mu_{x}, \sigma_{x}^{2})$
Then, $X + Y \sim \textit{N}(\mu_{x} + \mu_{y}, \sigma_{x}^{2} + \sigma_{y}^{2})$

- Why normal rvs are used to model the sum of random noises.
- (Extension) The sum of finitely many independent normals is also normal.

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Making a Metric of Dependence Degree

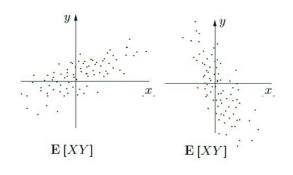


- Goal: Given two rvs X and Y, assign some number that quantifies the degree of their dependence
- Reqs.
 - a) Increases (resp. decreases) as they become more (resp. less) dependent.
 - b) 0 when they are independent.
 - c) Shows the direction of dependence by + and -
 - d) Always bounded by some numbers, e.g., [-1,1]
- Good engineers: Good at making good metrics
 - Metric of how our society is economically polarized
 - A lot of metrics in our professional sports leagues (baseball, basketball, etc)
 - Cybermetrics in MLB (Major League Baseball):
 http://m.mlb.com/glossary/advanced-stats

OK. Let's Design!



- Simple case: $\mathbb{E}[X] = \mu_X = 0$ and $\mathbb{E}[Y] = \mu_Y = 0$
- Dependent: Positive (If $X \uparrow$, $Y \uparrow$) or Negative (If $X \uparrow$, $Y \downarrow$)
- What about $\mathbb{E}[XY]$? Seems good.
 - $\circ \mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y] = 0 \text{ when } X \perp \!\!\!\perp Y$
 - More data points (thus increases) when xy > 0 (both positive or negative)



(Q) What about $\mathbb{E}[X + Y]$?

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What If $\mu_X \neq 0, \mu_Y \neq 0$?



• Solution: Centering. $X \to X - \mu_X$ and $Y \to Y - \mu_Y$

Covariance

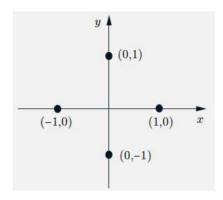
$$\operatorname{\mathsf{cov}}(X,Y) = \mathbb{E} ig[(X - \mathbb{E}[X]) \cdot (Y - \mathbb{E}[Y]) ig]$$

- After some algebra, $cov(X, Y) = \mathbb{E}[XY] \mathbb{E}[X]\mathbb{E}[Y]$
- $X \perp \!\!\!\perp Y \Longrightarrow \operatorname{cov}(X,Y) = 0$
- $cov(X, Y) = 0 \Longrightarrow X \perp \!\!\!\perp Y$? NO.
- When cov(X, Y) = 0, we say that X and Y are uncorrelated.

Example: cov(X, Y) = 0, but not independent



- $p_{X,Y}(1,0) = p_{X,Y}(0,1) = p_{X,Y}(-1,0) = p_{X,Y}(0,-1) = 1/4.$
- $\mathbb{E}[X] = \mathbb{E}[Y] = 0$, and $\mathbb{E}[XY] = 0$. So, $\mathsf{cov}(X,Y) = 0$
- Are they independent? No, because if X = 1, then we should have Y = 0.



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Some Properties



$$cov(X,X)=0$$

$$cov(aX + b, Y) = \mathbb{E}[(aX + b)Y] - \mathbb{E}[aX + b]\mathbb{E}[Y] = a \cdot cov(X, Y)$$

$$\mathsf{cov}(X,Y+Z) = \mathbb{E}[X(Y+Z)] - \mathbb{E}[X]\mathbb{E}[Y+Z] = \mathsf{cov}(X,Y) + \mathsf{cov}(X,Z)$$

$$\mathsf{var}[X+Y] = \mathbb{E}[(X+Y)^2] - (\mathbb{E}[X+Y])^2 = \mathsf{var}[X] + \mathsf{var}[Y] - 2\mathsf{cov}(X,Y)$$

Example: The hat problem in Lecture 3. Remember?



- n people throw their hats in a box and then pick one at random
- X: number of people with their own hat
- (Q) var[X]
- Key step 1. Define a rv $X_i = 1$ if i selects own hat and 0 otherwise. Then, $X = \sum_{i=1}^{n} X_i$.
- Key step 2. Are X_i s are independent?
- $X_i \sim Bernoulli(1/n)$. Thus, $\mathbb{E}[X_i] = 1/n$ and $var[X_i] = \frac{1}{n}(1 \frac{1}{n})$

• For $i \neq j$,

$$\begin{aligned} \text{cov}(X_i, X_j) &= \mathbb{E}[X_i X_j] - \mathbb{E}[X_i] \mathbb{E}[X_j] \\ &= \mathbb{P}(X_i = 1 \text{ and } X_j = 1) - \frac{1}{n^2} \\ &= \mathbb{P}(X_i = 1) \mathbb{P}(X_j = 1 | X_i = 1) - \frac{1}{n^2} \\ &= \frac{1}{n} \frac{1}{n-1} - \frac{1}{n^2} = \frac{1}{n^2(n-1)} \end{aligned}$$

$$\operatorname{var}[X] = \operatorname{var}\left[\sum X_i\right]$$

$$= \sum \operatorname{var}[X_i] + \sum_{i \neq j} \operatorname{cov}(X_i, X_j)$$

$$= n \frac{1}{n} (1 - \frac{1}{n}) + n(n - 1) \frac{1}{n^2(n - 1)} = 1$$

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Bounding the metric: Correlation Coefficient



- Reqs. a), b), and c) satisfied.
 - d) Always bounded by some numbers, e.g., $\left[-1,1\right]$
- Dimensionless metric. How? Normalization, but by what?

Correlation Coefficient

$$\rho(X,Y) = \mathbb{E}\left[\frac{(X - \mu_X)}{\sigma_X} \cdot \frac{Y - \mu_Y}{\sigma_Y}\right] = \frac{\text{cov}(X,Y)}{\sqrt{\text{var}[X]\text{var}[Y]}}$$

- $-1 \le \rho \le 1$
- $|\rho| = 1 \Longrightarrow X \mu_X = c(Y \mu_Y)$ (linear relation, VERY related)

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A Special Random Variable

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Consider a rv Y, such that

$$Y = \begin{cases} 0, & \text{w.p. } 1/4 \\ 1, & \text{w.p. } 1/4 \\ 2, & \text{w.p. } 1/2 \end{cases}$$

• If $h(y) = y^2$, then a new rv h(Y) is:

$$h(Y) = \begin{cases} 0, & \text{w.p. } 1/4 \\ 1, & \text{w.p. } 1/4 \\ 4, & \text{w.p. } 1/2 \end{cases}$$

ullet Consider other rv X, such that

$$g(y) = \mathbb{E}[X|Y = y] = \begin{cases} 3, & \text{if } y = 0 \\ 8, & \text{if } y = 1 \\ 9, & \text{if } y = 2 \end{cases}$$

• Then, a rv g(Y) is:

$$g(Y) = \begin{cases} 3, & \text{w.p. } 1/4 \\ 8, & \text{w.p. } 1/4 \\ 9, & \text{w.p. } 1/2 \end{cases}$$

- The rv g(Y) looks special, so let's notate it with some fancy one.
- What about? $X_{exp}(Y)$, $\mathbb{E}[X_Y]$, $\mathbb{E}_X[Y]$?

Conditional Expectation $\mathbb{E}[X|Y]$



Conditional Expectation

A random variable $g(Y) = \boxed{\mathbb{E}[X|Y]}$, called conditional expectation of X given Y, takes the value $g(y) = \mathbb{E}[X|Y = y]$, if Y happens to take the value y.

- A function of Y
- A random variable
- Thus, having a distribution, expectation, variance, all the things that a random variable has
- Often confusing because of the notation

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Expectation of $\mathbb{E}[X|Y]$



Expectation of Conditional Expectation

$$\mathbb{E}ig[\mathbb{E}[X|Y]ig]=\mathbb{E}[X],$$
 Law of iterated expectations

Proof.

$$\mathbb{E}\Big[\mathbb{E}[X|Y]\Big] = \sum_{y} \mathbb{E}[X|Y = y] p_{Y}(y)$$
$$= \mathbb{E}[X]$$

Examples and Meaning



- Stick of length I
- Uniformly break at point Y, and break what is left uniformly at point X.
- $\mathbb{E}[X|Y=y]=y/2$
- $\mathbb{E}[X|Y] = Y/2$
- $\mathbb{E}[X] = \mathbb{E}[\mathbb{E}[X|Y]] = \mathbb{E}[Y/2] = \frac{1}{2}\frac{I}{2} = I/4$

- Forecasts on sales: calculating expected value, given any available information
- X : February sales
- Forecast in the beg. of the year: $\mathbb{E}[X]$
- End of Jan. new information Y = y (Jan. sales) Revised forecast: $\mathbb{E}[X|Y = y]$ Revised forecast $\neq \mathbb{E}[X]$
- Law of iterated expectations $\mathbb{E}[\text{revised forecast}] = \text{original one}$

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Conditional Variance var[X|Y]



$$var[X] = \mathbb{E}[(X - \mathbb{E}[X])^2]$$

$$var[X|Y = y] = \mathbb{E}[(X - \mathbb{E}[X|Y = y])^2|Y = y]$$

Conditional Variance

A random variable g(Y) = var[X|Y] and called conditional variance of X given Y takes the value g(y) = var[X|Y = y], if Y happens to take the value y.

- A function of Y
- A random variable
- Thus, having a distribution, expectation, variance, all the things that a random variable has

Expectation and Variance of $\mathbb{E}[X|Y]$ and var[X|Y]



	$\mathbb{E}[X Y]$	var[X Y]
Expectation	$\mathbb{E}\Big[\mathbb{E}(X Y)\Big]$	$\mathbb{E}\Big[var(X Y)\Big]$
Variance	$varigl[\mathbb{E}(X Y)igr]$	var[var(X Y)]

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Law of Total Variance



Law of total variance

$$\mathsf{var}[X] = \mathbb{E}\Big[\mathsf{var}(X|Y)\Big] + \mathsf{var}[\mathbb{E}(X|Y)]$$

Proof.

$$\operatorname{var}(X|Y) = \mathbb{E}[X^{2}|Y] - (\mathbb{E}[X|Y])^{2}$$

$$\mathbb{E}\left[\operatorname{var}(X|Y)\right] = \mathbb{E}[X^{2}] - \mathbb{E}\left[(\mathbb{E}[X|Y])^{2}\right]$$
(1)

$$\operatorname{var}\left[\mathbb{E}(X|Y)\right] = \mathbb{E}\left[\left(\mathbb{E}[X|Y]\right)^{2}\right] - \left(\mathbb{E}\left[\mathbb{E}(X|Y)\right]\right)^{2} = \mathbb{E}\left[\left(\mathbb{E}[X|Y]\right)^{2}\right] - \left(\mathbb{E}[X]\right)^{2} \tag{2}$$

$$(1) + (2) = \mathbb{E}[X^2] + (\mathbb{E}[X])^2 = \text{var}[X]$$

Sum of a random number of rvs



- N: number of stores visited (random)
- X_i : money spent in store i, independent of other X_j and N, X_i s are identically distributed with $\mathbb{E}[X_i] = \mu$
- $Y = X_1 + X_2 + \dots X_N$. What are $\mathbb{E}[Y]$ and var[Y]?
- $\mathbb{E}[Y] = \mathbb{E}[\mathbb{E}[Y|N]] = \mathbb{E}[N\mathbb{E}[X_i]] = \mathbb{E}[N]\mathbb{E}[X_i] = \mu\mathbb{E}[N]$
- $\operatorname{var}[Y] = \mathbb{E}\left[\operatorname{var}(Y|N)\right] + \operatorname{var}\left[\mathbb{E}(Y|N)\right] = \mathbb{E}[N]\operatorname{var}[X_i] \mu^2\operatorname{var}[N]$ $\operatorname{var}(\mathbb{E}[Y|N]) = \operatorname{var}(N\mu) = \mu^2\operatorname{var}[N]$ $\operatorname{var}[Y|N] = N\operatorname{var}[X_i]$ $\mathbb{E}[\operatorname{var}(Y|N)] = \mathbb{E}[N\operatorname{var}[X_i]] = \mathbb{E}[N]\operatorname{var}[X_i]$

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Questions?

Review Questions



- 1) What are the key steps to get the derived distributions of Y = g(X) or Z = g(X, Y)?
- 2) How can we compute the distribution of Z + X + Y when X and Y are independent?
- 3) What are covariance and correlation coefficient? Why do we need them?
- 4) Please explain the concepts of conditional expectation and conditional variance.