

Lecture 6: Law of Large Numbers and Central Limit Theorem

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EE210: Probability and Introductory Random Processes KAIST EE

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Roadmap



- (1) Weak Law of Large Numbers: Result and Meaning
- (2) Central Limit Theorem: Result and Meaning
- (3) Weak Law of Large Numbers: Proof
 - Inequalities: Markov and Chebyshev
- (4) Central Limit Theorem: Proof
 - Moment Generating Function (MGF)
 - Two most remarkable findings in probability theory

Roadmap



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- (2) Central Limit Theorem: Result and Meaning
- (3) Weak Law of Large Numbers: Proof - Inequalities: Markov and Chebyshev
- (4) Central Limit Theorem: Proof
 - Moment Generating Function (MGF)

L7(1)



• Example 1. *n* students who decides their presence, depending on their feeling. Each student is happy or sad at random, and only happy students will show their presence. How many students will show their presence?

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- X_1, X_2, \ldots, X_n : i.i.d (independent and identically distributed) random variables
- $\mathbb{E}[X_i] = \mu$, $\operatorname{var}[X_i] = \sigma^2$
- Our interest is to understand how the following sum behaves:

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L7(1)



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• Figure out the distribution of S_n . Very challenging. Even just for Z = X + Y, finding the distribution, for example, requires the complex convolution.

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- Easy case: Sum of normal rvs = a normal rv, however, generally very challenging.
- Possible apporach. Take a certain scaling with respect to n that corresponds to a new glass, and investigate the system for large n

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- We call this law of large numbers (LLN).



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- However, M_n is a random variable, which is a function from Ω to \mathbb{R} .
- Need to build up the new concept of convergence for the sequence of rvs.



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- Now, $\{a_n\}$ are just the real numbers, and show that $a_n \to 0$ as $n \to \infty$.
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 - For any $\delta > 0$, there exists $N = N(\delta)$, such that for all $n \geq N$, $|a_n 0| \leq \delta$
- Convergence in probability: $Y_n \xrightarrow{\text{in prob.}} Y$
 - For any $\epsilon > 0$ and for any $\delta > 0$, there exists $N = N(\delta)$, such that for all $n \geq N$, $\mathbb{P}(|Y_n - Y| > \epsilon) < \delta.$

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 - For any $\epsilon > 0$ and for any $\delta > 0$, there exists $N = N(\delta)$, such that for all $n \geq N$,
 - $\mathbb{P}(|Y_n Y| \ge \epsilon) \le \delta.$ For any $\epsilon > 0$, $\mathbb{P}(\{|Y_n Y| \ge \epsilon\}) \xrightarrow{n \to \infty} 0$.

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- For any $\epsilon > 0$, $\mathbb{P}\Big(\{|Y_n \underline{Y}| \ge \epsilon\}\Big) \xrightarrow{n \to \infty} 0$. For any $\epsilon > 0$, $\mathbb{P}\Big(\{|Y_n \underline{a}| \ge \epsilon\}\Big) \xrightarrow{n \to \infty} 0$.
- A special case: when Y = a for some constant $a: Y_n \xrightarrow{\text{in prob.}} a$
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• For any $\epsilon > 0$, $\mathbb{P}\Big(\{|Y_n - \mathbf{a}| \ge \epsilon\}\Big) \xrightarrow{n \to \infty} 0$.



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- Proof. For any $\epsilon > 0$,

$$\mathbb{P}(|Y_n - 0| \ge \epsilon) = \mathbb{P}(X_1 \ge \epsilon, \dots, X_n \ge \epsilon) = \mathbb{P}(X_1 \ge \epsilon) \times \dots \times \mathbb{P}(X_n \ge \epsilon)$$
$$= (1 - \epsilon)^n \xrightarrow{n \to \infty} 0$$



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- Y: exponential rv with the parameter $\lambda=1$ (Remind: $\mathbb{P}(Y>y)=e^{-\lambda y}$)
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$$\mathbb{P}(Y_n = y) = \begin{cases} 1 - \frac{1}{n}, & \text{for } y = 0\\ \frac{1}{n}, & \text{for } y = n^2\\ 0, & \text{otherwise} \end{cases}$$



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• Thus, Y_n converges to 0 in probability.



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• Why "Weak"? There exists a stronger version, which we call "strong" law of large numbers. We will not cover the strong law of large numbers in this class.



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- The proof requires some knowledge about useful inequalities, which we will cover later.

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- For example, assume that a large number of identically distributed noises come to you. Then, you can roughly approximate it as $(n \times \text{average noise})$
- Provides an interpretation of expectations (as well as probabilities) in terms of a long sequence of identical independent experiments. For example, what is the probability of head of a coin? Toss 1000 times, and count the number of heads.

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- What's α for our magic?
- The answer is $\frac{1}{2}$

L7(2)



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- Interestingly, it converges to some well-known random variable.
 - Need a new concept of convergence: "convergence in distribution"

L7(2)



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• Another type of convergence of rvs



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$$\mathbb{P}(Y_n \leq y) \xrightarrow{n \to \infty} \mathbb{P}(Y \leq y)$$

- Another type of convergence of rvs
- · Comparison with convergence in probability?



• Consider a sequence of rvs $(Y_n)_{n=1,2,...}$ and a rv Y.

Convergence in Distribution: $Y_n \xrightarrow{\text{in dist.}} Y$

For every y,

$$\mathbb{P}(Y_n \leq y) \xrightarrow{n \to \infty} \mathbb{P}(Y \leq y)$$

- Another type of convergence of rvs
- Comparison with convergence in probability?
 - Convergence in probability

 Convergence in distribution, but the reverse is not true.
 - The proof is beyond what this class covers, but it will be interesting to find an example that shows convergence in distribution, which is not convergence in probability.

L7(2)

Example: in Distribution, but not in Probability



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- What about convergence in probability?

$$\mathbb{P}(|X_n - X| \ge \epsilon) = \mathbb{P}(|X_n - 1 + X_n| \ge \epsilon) = \mathbb{P}(|2X_n - 1| \ge \epsilon)$$

= $\mathbb{P}(1 \ge \epsilon)$ (because $|2X_n - 1| = 1$)

L7(2)



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• We can find ϵ small enough so that the above does not go to zero.

L7(2)

Central Limit Theorem: Formalism



•
$$S_n = X_1 + X_2 + \cdots + X_n$$
, $Z_n = \frac{S_n - n\mu}{\sigma\sqrt{n}}$

Central Limit Theorem

 Z_n convergens to Z in distribution, where $Z \sim \mathcal{N}(0,1)$.

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Central Limit Theorem

 Z_n convergens to Z in distribution, where $Z \sim \mathcal{N}(0,1)$.

- Very surprising!
- Irrespecitive of the distribution of X_i , Z is normal.

LLG vs. CLT: Different Scaling Glasses



• For simplicity, assume that $\mathbb{E}(X_i)=0$ and $\mathrm{var}(X_i)=1,\,i=1,2,\ldots,n$

LLG vs. CLT: Different Scaling Glasses



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- Law of Large Numbers

Scaling S_n by 1/n, you go to a deterministic world.

LLG vs. CLT: Different Scaling Glasses



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Scaling S_n by 1/n, you go to a deterministic world.

Central Limit Theorem

Scaling S_n by $1/\sqrt{n}$, you still stay at the random world, but not an arbitrary random world. That's the normal random world, not depending on the distribution of each X_i .

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$$Z_n = rac{\mathcal{S}_n - n \mu}{\sigma \sqrt{n}}, \qquad \quad \mathbb{P}(Z_n \leq z) \xrightarrow{n o \infty} \mathbb{P}(Z \leq z), \,\, Z \sim \mathcal{N}(0,1)$$

¹Only unique mode. A single maximum or minimum.



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- How large should n be?

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 - A moderate n (20 or 30) usually works, which is the power of CLT.

L7(2)

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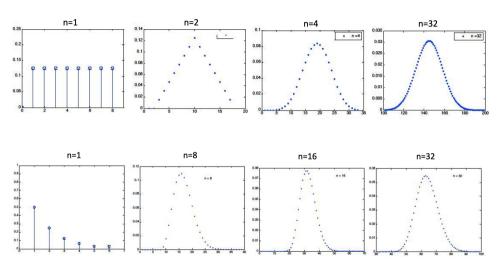
$$-S_n = n\mu + Z_n \sigma \sqrt{n}$$

- How large should n be?
 - A moderate n (20 or 30) usually works, which is the power of CLT.
 - If X_i resembles a normal rv more, smaller n works: symmetry and unimodality¹

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CLT: Examples of Required *n*







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 $\mathbb{P}(S_n \leq a) \approx b$: Given two parameters, find the third

• Package weights X_i : iid exponential $\lambda=1/2$ ($\mu=1/\lambda=2$ and $\sigma^2=1/\lambda^2=4$)



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- Load container with n = 100 packages

$$\mathbb{P}(S_{100} \geq 210) = \mathbb{P}\Big[\frac{S_{100} - 100 \cdot 2}{2\sqrt{100}} \geq \frac{210 - 200}{20}\Big] = \mathbb{P}(Z_{100} \geq 0.5)$$

L7(2)



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L7(2)



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• The value of a such that $\Phi(\frac{a-200}{20}) = 0.95$? $\frac{a-200}{20} = 1.645$ and a = 232.9

L7(2)



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• The value of *n* such that $\frac{210-2n}{2\sqrt{n}} = 1.645$? n = 89

Roadmap



- (1) Weak Law of Large Numbers: Result and Meaning
- (2) Central Limit Theorem: Result and Meaning
- (3) Weak Law of Large Numbers: Proof
 Inequalities: Markov and Chebyshev
- (4) Central Limit Theorem: Proof
 - Moment Generating Function (MGF)

L7(3)



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If $X \geq 0$ and a > 0, then $\mathbb{P}(X \geq a) \leq \frac{\mathbb{E}(X)}{a}$.

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 \neg



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Thus,
$$a \cdot \mathbb{P}(X \geq a) \leq \mathbb{E}[X]$$
.

L7(3)



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$$\mathbb{P}\Big(|X-\mu| \geq c\Big) = \mathbb{P}\Big((X-\mu)^2 \geq c^2\Big) \leq \frac{\mathbb{E}\Big[(X-\mu)^2\Big]}{c^2} = \frac{\mathsf{var}(X)}{c^2}$$

L7(3)



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- Both bounds are the ones that bound the probability of rare events.

Back to WLLN Proof



$$M_n = \frac{S_n}{n} = \frac{X_1 + X_2 + \dots X_n}{n}$$

Weak law of large numbers

 M_n converges to μ in probability.

Proof. For any given $\epsilon > 0$,

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Proof. For any given $\epsilon > 0$,

$$\mathbb{P}\Big(|M_n - \mu| \ge \epsilon\Big) \le \frac{\operatorname{var}(M_n)}{\epsilon^2} = \frac{\sigma^2}{n\epsilon^2} \xrightarrow{n \to \infty} 0$$

Comparison: WLLN vs. CLT



We ask the same question, and try to answer it, using WLLN or CLT.

See how the answers becomes different.



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- Interview n randomly selected voters and record the result in $M_n = \frac{X_1 + ... + X_n}{n}$ which is an estimate of p, where the Bernoulli rv $X_i = 1$ if i-th interviewee answers "yes", and 0 otherwise.



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$$\circ \ \epsilon = 0.1 \ \text{and} \ \frac{1}{4n\epsilon^2} \le 0.25 \implies n \ge 100$$



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$$\epsilon = 0.1$$
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L7(3)



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L7(3)



$$\mathbb{P}(|M_n - p| \ge \epsilon) = =$$
 $\leq =$



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 \le



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• Compare: 50,000 from LLN vs. 9604 from CLT

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Roadmap



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- (1) Weak Law of Large Numbers: Result and Meaning
- (2) Central Limit Theorem: Result and Meaning
- (3) Weak Law of Large Numbers: Proof
 Inequalities: Markov and Chebyshev
- (4) Central Limit Theorem: Proof
 - Moment Generating Function (MGF)

L7(4)

Moment Generating Function (MGF)



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• If the context is clear, we omit X and use just M(s).



Ex1) Let $p_X(x)$ is given as:

$$p_X(x) = \begin{cases} 1/2, & \text{if } x = 2\\ 1/6, & \text{if } x = 3\\ 1/3, & \text{if } x = 5 \end{cases}$$

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$$= \lambda \frac{e^{(s-\lambda)x}}{s-\lambda} \bigg|_{0}^{\infty} \quad \text{(if } s < \lambda \text{)} = \frac{\lambda}{\lambda - s}$$



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$$Y = aX + b$$
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L7(4)

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$$M(s) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} e^{sy} dy + \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{y^2}{2} + sy} dy$$
$$= e^{\frac{s^2}{2}} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-(y-s)^2} dy$$

 $=e^{s^2/2}$ (because it is the pdf of $\mathcal{N}(s,1)$)

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• Question. MGF of $\mathcal{N}(\mu, \sigma^2)$?

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- $3. \left. \frac{d^n}{ds^n} M(s) \right|_{s=0} = \mathbb{E}[X^n]$
- 4. MGF provides a convenient way of generating moments. That's why it is called moment generating function.



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- The first and the second moments are:

$$M'(s) = \frac{\lambda}{(\lambda - s)^2} \rightarrow \mathbb{E}(X) = M'(0) = 1/\lambda$$
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Example¹



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• Thus, $var(X) = 2/\lambda^2 - 1/\lambda^2 = 1/\lambda^2$

Inversion Property



Inversion Property

The transform $M_X(s)$ associated with a random variable X uniquely determines the CDF of X, assuming that $M_X(s)$ is finite for all s in some interval [-a,a], where a is a positive number.

- In easy words, we can recover the distribution if we know the MGF.
- Thus, each rv has its own unique MGF.



• Given the following MGF of rv X, what is the distribution of X?

$$M(s) = \frac{1}{4}e^{-s} + \frac{1}{2} + \frac{1}{8}e^{4s} + \frac{1}{8}e^{5s}$$



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- We can see that

$$p_X(-1) = \frac{1}{4}, \ p_X(0) = \frac{1}{2}, \ p_X(4) = \frac{1}{8}, \ p_X(5) = \frac{1}{8}$$



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- X is a geometric rv with parameter p



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$$\begin{split} \mathbb{E}\Big[e^{sS_n/\sqrt{n}}\Big] &= \mathbb{E}\Big[e^{sX_1/\sqrt{n}}\Big] \times \dots \times \mathbb{E}\Big[e^{sX_n/\sqrt{n}}\Big] \\ &= \left(\mathbb{E}\Big[e^{sX_1/\sqrt{n}}\Big]\right)^n = \end{split}$$



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• For simplicity, let $M(\cdot) = M_{X_1}(\cdot)$



•
$$M(0) = 1$$
, $M'(0) = 0$, $M''(0) = 1$



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- If we apply l'hopital's rule twice (please check), we get

$$\lim_{y\to 0}\frac{\log M(ys)}{y^2}=\frac{s^2}{2}$$



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Questions?

L7(4) May 13, 2021

Review Questions



- 1) What's the practical value of LLN and CLT?
- 2) Explain LLN and CLT from the scaling perspective.
- 3) Why are LLN and CLT great?
- 4) Why do we need different concepts of convergence for random variables?
- 5) Explain what is convergence in probability.
- 6) Explain what is convergence in distribution.
- 7) Why is MGF useful?