

Lecture 6: Law of Large Numbers and Central Limit Theorem

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EE210: Probability and Introductory Random Processes KAIST EE

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Roadmap



- (1) Weak Law of Large Numbers: Result and Meaning
- (2) Central Limit Theorem: Result and Meaning
- (3) Weak Law of Large Numbers: ProofInequalities: Markov and Chebyshev
- (4) Central Limit Theorem: Proof
 - Moment Generating Function (MGF)
 - Two most remarkable findings in probability theory

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- (1) Weak Law of Large Numbers: Result and Meaning
- (2) Central Limit Theorem: Result and Meaning
- (3) Weak Law of Large Numbers: Proof
 Inequalities: Markov and Chebyshev
- (4) Central Limit Theorem: Proof
 - Moment Generating Function (MGF)



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L7(1)



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- X_1, X_2, \dots, X_n : i.i.d (independent and identically distributed) random variables
- $\mathbb{E}[X_i] = \mu$, $\text{var}[X_i] = \sigma^2$
- Our interest is to understand how the following sum behaves:

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L7(1)



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- Easy case: Sum of normal rvs = a normal rv, however, generally very challenging.
- Possible apporach. Take a certain scaling with respect to n that corresponds to a new glass, and investigate the system for large n



• Consider the sample mean, and try to understand how S_n behaves:

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- We call this law of large numbers (LLN).



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- However, M_n is a random variable, which is a function from Ω to \mathbb{R} .
- Need to build up the new concept of convergence for the sequence of rvs.



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- Now, $\{a_n\}$ are just the real numbers, and show that $a_n \to 0$ as $n \to \infty$.
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 - For any $\delta > 0$, there exists $N = N(\delta)$, such that for all $n \geq N$, $|a_n 0| \leq \delta$
- Convergence in probability: $Y_n \xrightarrow{\text{in prob.}} Y$
 - For any $\epsilon > 0$ and for any $\delta > 0$, there exists $N = N(\delta)$, such that for all $n \ge N$, $\mathbb{P}(|Y_n Y| \ge \epsilon) \le \delta$.

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 - For any $\epsilon > 0$ and for any $\delta > 0$, there exists $N = N(\delta)$, such that for all $n \ge N$, $\mathbb{P}(|Y_n Y| \ge \epsilon) \le \delta$.
 - $\mathbb{P}(|Y_n Y| \ge \epsilon) \le \delta.$ For any $\epsilon > 0$, $\mathbb{P}(\{|Y_n Y| \ge \epsilon\}) \xrightarrow{n \to \infty} 0$.



- For any $\epsilon > 0$, $\mathbb{P}\left(\{|Y_n Y| \ge \epsilon\}\right) \xrightarrow{n \to \infty} 0$. For any $\epsilon > 0$, $\mathbb{P}\left(\{|Y_n a| \ge \epsilon\}\right) \xrightarrow{n \to \infty} 0$.
- A special case: when Y = a for some constant $a: Y_n \xrightarrow{\text{in prob.}} a$
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• For any $\epsilon > 0$, $\mathbb{P}(\{|Y_n - \mathbf{a}| \ge \epsilon\}) \xrightarrow{n \to \infty} 0$.



- For any $\epsilon>0,\,\mathbb{P}\left(\{|Y_n-\mathbf{a}|\geq\epsilon\}\right)\xrightarrow{n\to\infty}0.$
- A sequence of iid rvs $X_n \sim \mathcal{U}[0,1]$, and let

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- Proof. For any $\epsilon > 0$,

$$\mathbb{P}(|Y_n - 0| \ge \epsilon) = \mathbb{P}(X_1 \ge \epsilon, \dots, X_n \ge \epsilon) = \mathbb{P}(X_1 \ge \epsilon) \times \dots \times \mathbb{P}(X_n \ge \epsilon)$$
$$= (1 - \epsilon)^n \xrightarrow{n \to \infty} 0$$

Example $\overline{2}$: Convergence in Probability



• For any $\epsilon>0,\,\mathbb{P}\left(\{|Y_{n}-\textbf{\textit{a}}|\geq\epsilon\}\right)\xrightarrow{n\to\infty}0.$



- For any $\epsilon > 0$, $\mathbb{P}\left(\{|Y_n \mathbf{a}| \ge \epsilon\}\right) \xrightarrow{n \to \infty} 0$.
- Y: exponential rv with the parameter $\lambda=1$ (Remind: $\mathbb{P}(Y>y)=e^{-\lambda y}$)
- a sequence of rvs $Y_n = Y/n$ (note that these are dependent)



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- For any $\epsilon > 0$, $\mathbb{P}\left(\{|Y_n \mathbf{a}| \geq \epsilon\}\right) \xrightarrow{n \to \infty} 0$.
- Consider a sequence of rvs Y_n with the following distribution:

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• Thus, Y_n converges to 0 in probability.



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Weak law of large numbers

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- The proof requires some knowledge about useful inequalities, which we will cover later.



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- For example, assume that a large number of identically distributed noises come to you. Then, you can roughly approximate it as $(n \times \text{average noise})$
- Provides an interpretation of expectations (as well as probabilities) in terms of a long sequence of identical independent experiments. For example, what is the probability of head of a coin? Toss 1000 times, and count the number of heads.

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- What's α for our magic?
- The answer is $\frac{1}{2}$



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 - Need a new concept of convergence: "convergence in distribution"

L7(2)



• Consider a sequence of rvs $(Y_n)_{n=1,2,...}$ and a rv Y.

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- · Comparison with convergence in probability?



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Convergence in Distribution: $Y_n \xrightarrow{\text{in dist.}} Y$

For every y,

$$\mathbb{P}(Y_n \leq y) \xrightarrow{n \to \infty} \mathbb{P}(Y \leq y)$$

- Another type of convergence of rvs
- Comparison with convergence in probability?
 - Convergence in probability

 Convergence in distribution, but the reverse is not true.
 - The proof is beyond what this class covers, but it will be interesting to find an example that shows convergence in distribution, which is not convergence in probability.

Example: in Distribution, but not in Probability



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- What about convergence in probability?

$$\mathbb{P}(|X_n - X| \ge \epsilon) = \mathbb{P}(|X_n - 1 + X_n| \ge \epsilon) = \mathbb{P}(|2X_n - 1| \ge \epsilon)$$

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• We can find ϵ small enough so that the above does not go to zero.

Central Limit Theorem: Formalism



•
$$S_n = X_1 + X_2 + \cdots + X_n$$
, $Z_n = \frac{S_n - n\mu}{\sigma\sqrt{n}}$

Central Limit Theorem

 Z_n convergens to Z in distribution, where $Z \sim \mathcal{N}(0,1)$.

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- Very surprising!
- Irrespecitive of the distribution of X_i , Z is normal.

LLG vs. CLT: Different Scaling Glasses



• For simplicity, assume that $\mathbb{E}(X_i) = 0$ and $\text{var}(X_i) = 1, i = 1, 2, \dots, n$

LLG vs. CLT: Different Scaling Glasses



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Scaling S_n by 1/n, you go to a deterministic world.

Central Limit Theorem

Scaling S_n by $1/\sqrt{n}$, you still stay at the random world, but not an arbitrary random world. That's the normal random world, not depending on the distribution of each X_i .

L7(2)



$$Z_n = rac{\mathcal{S}_n - n \mu}{\sigma \sqrt{n}}, \qquad \quad \mathbb{P}(Z_n \leq z) \xrightarrow{n o \infty} \mathbb{P}(Z \leq z), \,\, Z \sim \mathcal{N}(0,1)$$

¹Only unique mode. A single maximum or minimum.



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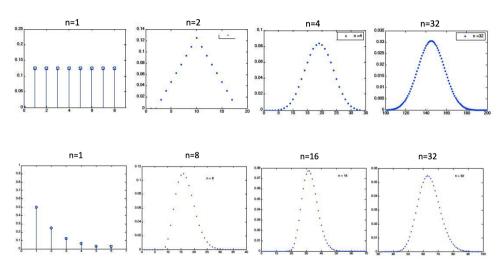
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- How large should n be?
 - A moderate n (20 or 30) usually works, which is the power of CLT.
 - If X_i resembles a normal rv more, smaller n works: symmetry and unimodality¹

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CLT: Examples of Required *n*







 $\mathbb{P}(S_n \leq a) \approx b$: Given two parameters, find the third

• Package weights X_i : iid exponential $\lambda=1/2$ ($\mu=1/\lambda=2$ and $\sigma^2=1/\lambda^2=4$)



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$$\mathbb{P}(S_{100} \geq 210) = \mathbb{P}\Big[\frac{S_{100} - 100 \cdot 2}{2\sqrt{100}} \geq \frac{210 - 200}{20}\Big] = \mathbb{P}(Z_{100} \geq 0.5)$$

L7(2)



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L7(2)



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$$\mathbb{P}(S_{100} \geq a) = \mathbb{P}\Big[\frac{S_{100} - 100 \cdot 2}{2\sqrt{100}} \geq \frac{a - 200}{20}\Big] = \mathbb{P}(Z_{100} \geq \frac{a - 200}{20})$$



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• The value of a such that $\Phi(\frac{a-200}{20}) = 0.95$? $\frac{a-200}{20} = 1.645$ and a = 232.9

L7(2)



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• The value of *n* such that $\frac{210-2n}{2\sqrt{n}} = 1.645$? n = 89

Roadmap



- (1) Weak Law of Large Numbers: Result and Meaning
- (2) Central Limit Theorem: Result and Meaning
- (3) Weak Law of Large Numbers: Proof
 Inequalities: Markov and Chebyshev
- (4) Central Limit Theorem: Proof
 - Moment Generating Function (MGF)

L7(3)



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If $X \geq 0$ and a > 0, then $\mathbb{P}(X \geq a) \leq \frac{\mathbb{E}(X)}{a}$.

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Thus,
$$a \cdot \mathbb{P}(X \geq a) \leq \mathbb{E}[X]$$
.

L7(3)





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L7(3) October 26, 2021



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L7(3)



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- Knowing the variance helps
- Both bounds are the ones that bound the probability of rare events.

Back to WLLN Proof



$$M_n = \frac{S_n}{n} = \frac{X_1 + X_2 + \dots X_n}{n}$$

Weak law of large numbers

 M_n converges to μ in probability.

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Comparison: WLLN vs. CLT



We ask the same question, and try to answer it, using WLLN or CLT.

See how the answers become different.



• p: fraction of voters who support "Yung".



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- Interview n randomly selected voters and record the result in $M_n = \frac{X_1 + ... + X_n}{n}$ which is an estimate of p, where the Bernoulli rv $X_i = 1$ if i-th interviewee answers "yes", and 0 otherwise.



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•
$$\epsilon = 0.1$$
 and $\frac{1}{4n\epsilon^2} \leq 0.25 \implies n \geq 100$

• Question. What is *n* so that the probability that our estimate is incorrect by more than 0.01 is no larger than 0.05?

L7(3)



- p: fraction of voters who support "Yung".
- Interview n randomly selected voters and record the result in $M_n = \frac{X_1 + ... + X_n}{n}$ which is an estimate of p, where the Bernoulli rv $X_i = 1$ if i-th interviewee answers "yes", and 0 otherwise.
- $\mathbb{P}(|M_n-p|\geq\epsilon)\leq rac{\sigma^2}{n\epsilon^2}=rac{p(1-p)}{n\epsilon^2}\leq rac{1}{4n\epsilon^2}$ (because $p(1-p)\leq 1/4$)
- Question. What is *n* so that the probability that our estimate is incorrect by more than 0.1 is no larger than 0.25?

•
$$\epsilon = 0.1$$
 and $\frac{1}{4n\epsilon^2} \leq 0.25 \implies n \geq 100$

• Question. What is *n* so that the probability that our estimate is incorrect by more than 0.01 is no larger than 0.05?

$$\epsilon = 0.01$$
 and $\frac{1}{4nc^2} \le 0.05 \implies n \ge 50000$

L7(3)



$$\mathbb{P}(|M_n - p| \ge \epsilon) = =$$
 $\leq =$



$$\mathbb{P}(|M_n - p| \ge \epsilon) = \mathbb{P}\left[\left|\frac{S_n - np}{n}\right| \ge \epsilon\right] =$$

$$\le \qquad \qquad =$$



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$$\leq \mathbb{P}\left[\left|\frac{S_n - np}{\sigma\sqrt{n}}\right| \ge 2\epsilon\sqrt{n}\right] = \qquad \text{(because } \sigma = \sqrt{p(1-p)} \le 1/2\text{)}$$



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L7(3) October 26, 2021



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L7(3)



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Compare: 50,000 from LLN vs. 9604 from CLT

L7(3)

Roadmap



- (1) Weak Law of Large Numbers: Result and Meaning
- (2) Central Limit Theorem: Result and Meaning
- (3) Weak Law of Large Numbers: Proof
 Inequalities: Markov and Chebyshev
- (4) Central Limit Theorem: Proof
 - Moment Generating Function (MGF)

L7(4)

Moment Generating Function (MGF)



• For a rv X, it is a kind of transform

Moment Generating Function (MGF)



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• If the context is clear, we omit X and use just M(s).



Ex1) Let $p_X(x)$ is given as:

$$p_X(x) = \begin{cases} 1/2, & \text{if } x = 2\\ 1/6, & \text{if } x = 3\\ 1/3, & \text{if } x = 5 \end{cases}$$

$$M(s) = \mathbb{E}(e^{sX}) = \frac{1}{2}e^{2s} + \frac{1}{6}e^{3s} + \frac{1}{3}e^{5s}$$



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Ex2)
$$X \sim \exp(\lambda), f_X(x) = \lambda e^{-\lambda x}, x \ge 0$$

$$M(s) = \lambda \int_0^\infty e^{sx} e^{-\lambda x} dx$$

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$$Y = aX + b$$
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$$= e^{\frac{s^2}{2}} \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-\frac{(y-s)^2}{2}} dy$$

 $=e^{s^2/2}$ (because it is the pdf of $\mathcal{N}(s,1)$)

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• Question. MGF of $\mathcal{N}(\mu,\sigma^2)$?

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L7(4)



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- 2. Similarly, $M''(0) = \mathbb{E}[X^2]$
- $3. \left. \frac{d^n}{ds^n} M(s) \right|_{s=0} = \mathbb{E}[X^n]$
- 4. MGF provides a convenient way of generating moments. That's why it is called moment generating function.



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- The first and the second moments are:

$$M'(s) = \frac{\lambda}{(\lambda - s)^2} \rightarrow \mathbb{E}(X) = M'(0) = 1/\lambda$$
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Example¹



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• Thus, $var(X) = 2/\lambda^2 - 1/\lambda^2 = 1/\lambda^2$

Inversion Property



Inversion Property

The transform $M_X(s)$ associated with a random variable X uniquely determines the CDF of X, assuming that $M_X(s)$ is finite for all s in some interval [-a,a], where a is a positive number.

- In easy words, we can recover the distribution if we know the MGF.
- Thus, each rv has its own unique MGF.



• Given the following MGF of rv X, what is the distribution of X?

$$M(s) = \frac{1}{4}e^{-s} + \frac{1}{2} + \frac{1}{8}e^{4s} + \frac{1}{8}e^{5s}$$



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- We can see that

$$p_X(-1) = \frac{1}{4}, \ p_X(0) = \frac{1}{2}, \ p_X(4) = \frac{1}{8}, \ p_X(5) = \frac{1}{8}$$



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L7(4)



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- X is a geometric rv with parameter p





- Without loss of generality, assume $\mathbb{E}(X_i) = 0$ and $\text{var}(X_i) = 1$
- $Z_n = \frac{S_n}{\sqrt{n}} = \frac{X_1 + X_2 + \dots X_n}{\sqrt{n}}$



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$$\begin{split} \mathbb{E}\Big[e^{\mathsf{s} S_n/\sqrt{n}}\Big] &= \mathbb{E}\Big[e^{\mathsf{s} X_1/\sqrt{n}}\Big] \times \dots \times \mathbb{E}\Big[e^{\mathsf{s} X_n/\sqrt{n}}\Big] \\ &= \left(\mathbb{E}\Big[e^{\mathsf{s} X_1/\sqrt{n}}\Big]\right)^n = \end{split}$$



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• For simplicity, let $M(\cdot) = M_{X_1}(\cdot)$



• M(0) = 1, M'(0) = 0, M''(0) = 1



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L7(4)



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- For convenience, do the change of variable $y = \frac{1}{\sqrt{n}}$. Then, $\lim_{y \to 0} \frac{\log M(ys)}{y^2}$
- If we apply l'hopital's rule twice (please check), we get

$$\lim_{y\to 0}\frac{\log M(ys)}{y^2}=\frac{s^2}{2}$$



Questions?

L7(4)

Review Questions



- 1) Explain the meaning of Markov inequality and Chebyshev inequality.
- 2) What are the practical values of LLN and CLT?
- 3) Explain LLN and CLT from the scaling perspective.
- 4) Why do we need different concepts of convergence for random variables?
- 5) Explain what is convergence in probability.
- 6) Explain what is convergence in distribution.
- 7) Why is MGF (Moment Generating Function) useful?
- 8) Prove CLT using MGF.