Lecture 7: Random Processes, Part I

Yi, Yung (이융)

EE210: Probability and Introductory Random Processes
KAIST EE

October 25, 2021

(1) Introduction of Random Processes

(2) Bernoulli Processes

(3) Poisson Processes: Poisson RV and Basic Idea

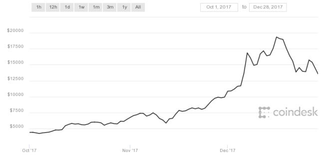
(4) Poisson Processes: Definition and Properties

(5) Playing with Bernoulli and Poisson Processes

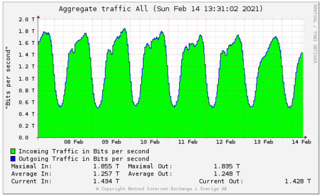
(6) Random Incidence Phenomenon

Roadmap

- Many probabilistic experiments that evolve in time



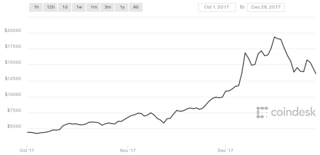
(a) Prices of a crytocurrency



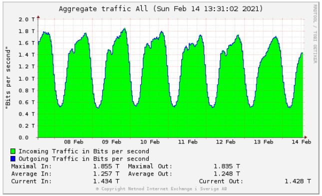
(b) Internet traffic traces

Things that evolve in time

- Many probabilistic experiments that evolve in time
  - Sequence of daily prices of a stock
  - Sequence of scores in football
  - Sequence of failure times of a machine
  - Sequence of traffic loads in Internet



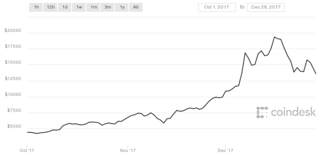
(a) Prices of a crytocurrency



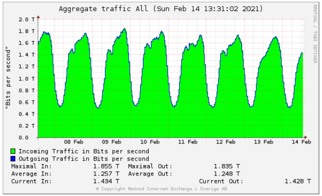
(b) Internet traffic traces

Things that evolve in time

- Many probabilistic experiments that evolve in time
  - Sequence of daily prices of a stock
  - Sequence of scores in football
  - Sequence of failure times of a machine
  - Sequence of traffic loads in Internet
- Random process is a mathematical model for it.



(a) Prices of a crytocurrency



(b) Internet traffic traces

Random Process: Basics (1)

Random Process: Basics (1)

- A random process is a sequence of random variables indexed by time.

Random Process: Basics (1)

- A random process is a sequence of random variables indexed by time.
- Time: discrete or continuous (a modeling choice in most cases)

Random Process: Basics (1)

- A random process is a sequence of random variables indexed by time.
- Time: discrete or continuous (a modeling choice in most cases)
- Notation
  - $(X_t)_{t \in \mathcal{T}}$ or $\Big(X(t)\Big)_{t \in \mathcal{T}}$, where $\mathcal{T} = \mathbb{R}$ (continuous) or $\mathcal{T} = \{0, 1, 2, \dots\}$ (discrete)

Random Process: Basics (1)

- A random process is a sequence of random variables indexed by time.
- Time: discrete or continuous (a modeling choice in most cases)
- Notation
  - $(X_t)_{t\in\mathcal{T}}$ or $\Big(X(t)\Big)_{t\in\mathcal{T}}$, where $\mathcal{T}=\mathbb{R}$ (continuous) or $\mathcal{T}=\{0,1,2,\ldots\}$ (discrete)
  - For the discrete case, we also often use $(X_n)_{n\in\mathbb{Z}_+}$.

Random Process: Basics (1)

- A random process is a sequence of random variables indexed by time.
- Time: discrete or continuous (a modeling choice in most cases)
- Notation
  - $(X_t)_{t \in \mathcal{T}}$ or $\Big(X(t)\Big)_{t \in \mathcal{T}}$, where $\mathcal{T} = \mathbb{R}$ (continuous) or $\mathcal{T} = \{0, 1, 2, \ldots\}$ (discrete)
  - For the discrete case, we also often use $(X_n)_{n \in \mathbb{Z}_+}$.
  - We will use all of them, unless confusion arises.

Random Process: Basics (1)

- A random process is a sequence of random variables indexed by time.
- Time: discrete or continuous (a modeling choice in most cases)
- Notation
  - $(X_t)_{t \in \mathcal{T}}$ or $\Big(X(t)\Big)_{t \in \mathcal{T}}$, where $\mathcal{T} = \mathbb{R}$ (continuous) or $\mathcal{T} = \{0, 1, 2, \ldots\}$ (discrete)
  - For the discrete case, we also often use $(X_n)_{n \in \mathbb{Z}_+}$.
  - We will use all of them, unless confusion arises.
- For a fixed time t, X_t (or $X(t)$) is a random variable.

Random Process: Basics (1)

- A random process is a sequence of random variables indexed by time.
- Time: discrete or continuous (a modeling choice in most cases)
- Notation
  - $(X_t)_{t \in \mathcal{T}}$ or $\Big(X(t)\Big)_{t \in \mathcal{T}}$, where $\mathcal{T} = \mathbb{R}$ (continuous) or $\mathcal{T} = \{0, 1, 2, \ldots\}$ (discrete)
  - For the discrete case, we also often use $(X_n)_{n \in \mathbb{Z}_+}$.
  - We will use all of them, unless confusion arises.
- For a fixed time t, X_t (or $X(t)$) is a random variable.
- The values that X_t (or $X(t)$) can take: discrete or continuous

- Example. Discrete time RP. $\{X_1, X_2, X_3, \ldots, \}$.

Random Processes: Basics (2)

- Example. Discrete time RP. $\{X_1, X_2, X_3, \ldots, \}$.
  - X_i: # of Covid-19 infections at day i in South Korea, which is a random variable.

Random Processes: Basics (2)

- Example. Discrete time RP. $\{X_1, X_2, X_3, \ldots, \}$.
  - X_i: $\#$ of Covid-19 infections at day i in South Korea, which is a random variable.
  - Then, $X_i : \Omega \mapsto \mathbb{R}$, where the sample space Ω is the set of all outcomes.

Random Processes: Basics (2)

- Example. Discrete time RP. $\{X_1, X_2, X_3, \ldots,\}$.
  - X_i: # of Covid-19 infections at day i in South Korea, which is a random variable.
  - Then, $X_i : \Omega \mapsto \mathbb{R}$, where the sample space Ω is the set of all outcomes.
  - An outcome $\omega \in \Omega$ is a infinite sequence of infections

Random Processes: Basics (2)

- Example. Discrete time RP. $\{X_1, X_2, X_3, \ldots, \}$.
  - X_i: # of Covid-19 infections at day i in South Korea, which is a random variable.
  - Then, $X_i : \Omega \mapsto \mathbb{R}$, where the sample space Ω is the set of all outcomes.
  - An outcome $\omega \in \Omega$ is a infinite sequence of infections
  - For example, $\omega_1 = (100, 150, 130, \ldots)$, $\omega_2 = (200, 300, 400, \ldots)$

Random Processes: Basics (2)

- Example. Discrete time RP. $\{X_1, X_2, X_3, \ldots,\}$.
  - X_i: # of Covid-19 infections at day i in South Korea, which is a random variable.
  - Then, $X_i : \Omega \mapsto \mathbb{R}$, where the sample space Ω is the set of all outcomes.
  - An outcome $\omega \in \Omega$ is a infinite sequence of infections
  - For example, $\omega_1 = (100, 150, 130, \ldots)$, $\omega_2 = (200, 300, 400, \ldots)$
  - $X_3(\omega_1) = 130$, $X_2(\omega_2) = 300$, $X_1(\omega_2) = 200$, etc.

Random Processes: Basics (2)

- Example. Discrete time RP. $\{X_1, X_2, X_3, \ldots, \}$.
  - X_i: # of Covid-19 infections at day i in South Korea, which is a random variable.
  - Then, $X_i : \Omega \mapsto \mathbb{R}$, where the sample space Ω is the set of all outcomes.
  - An outcome $\omega \in \Omega$ is a infinite sequence of infections
  - For example, $\omega_1 = (100, 150, 130, \ldots)$, $\omega_2 = (200, 300, 400, \ldots)$
  - $X_3(\omega_1) = 130$, $X_2(\omega_2) = 300$, $X_1(\omega_2) = 200$, etc.
- Example. Continuous time RP. $(X(t))_{t \in \mathbb{R}^+}$

- Example. Discrete time RP. $\{X_1, X_2, X_3, \ldots, \}$.
    - X_i: # of Covid-19 infections at day i in South Korea, which is a random variable.
    - Then, $X_i : \Omega \mapsto \mathbb{R}$, where the sample space Ω is the set of all outcomes.
    - An outcome $\omega \in \Omega$ is a infinite sequence of infections
    - For example, $\omega_1 = (100, 150, 130, \ldots)$, $\omega_2 = (200, 300, 400, \ldots)$
    - $X_3(\omega_1) = 130$, $X_2(\omega_2) = 300$, $X_1(\omega_2) = 200$, etc.
- Example. Continuous time RP. $(X(t))_{t \in \mathbb{R}^+}$
    - $X(t)$: bitcoin price at time t, which is a random variable.

Random Processes: Basics (2)

- Example. Discrete time RP. $\{X_1, X_2, X_3, \ldots, \}$.
  - X_i: # of Covid-19 infections at day i in South Korea, which is a random variable.
  - Then, $X_i : \Omega \mapsto \mathbb{R}$, where the sample space Ω is the set of all outcomes.
  - An outcome $\omega \in \Omega$ is a infinite sequence of infections
  - For example, $\omega_1 = (100, 150, 130, \ldots)$, $\omega_2 = (200, 300, 400, \ldots)$
  - $X_3(\omega_1) = 130$, $X_2(\omega_2) = 300$, $X_1(\omega_2) = 200$, etc.
- Example. Continuous time RP. $(X(t))_{t \in \mathbb{R}^+}$
  - $X(t)$: bitcoin price at time t, which is a random variable.
  - Then, $X(t) : \Omega \mapsto \mathbb{R}$.

KAIST EE

- Example. Discrete time RP. $\{X_1, X_2, X_3, \ldots, \}$.
  - X_i: # of Covid-19 infections at day i in South Korea, which is a random variable.
  - Then, $X_i : \Omega \mapsto \mathbb{R}$, where the sample space Ω is the set of all outcomes.
  - An outcome $\omega \in \Omega$ is a infinite sequence of infections
  - For example, $\omega_1 = (100, 150, 130, \ldots)$, $\omega_2 = (200, 300, 400, \ldots)$
  - $X_3(\omega_1) = 130$, $X_2(\omega_2) = 300$, $X_1(\omega_2) = 200$, etc.
- Example. Continuous time RP. $(X(t))_{t \in \mathbb{R}^+}$
  - $X(t)$: bitcoin price at time t, which is a random variable.
  - Then, $X(t) : \Omega \mapsto \mathbb{R}$.
  - An outcome $\omega \in \Omega$ is a trajectory of prices over $[0, \infty)$

- Example. Discrete time RP. $\{X_1, X_2, X_3, \dots, \}$.
  - X_i: # of Covid-19 infections at day i in South Korea, which is a random variable.
  - Then, $X_i : \Omega \mapsto \mathbb{R}$, where the sample space Ω is the set of all outcomes.
  - An outcome $\omega \in \Omega$ is a infinite sequence of infections
  - For example, $\omega_1 = (100, 150, 130, \dots)$, $\omega_2 = (200, 300, 400, \dots)$
  - $X_3(\omega_1) = 130$, $X_2(\omega_2) = 300$, $X_1(\omega_2) = 200$, etc.
- Example. Continuous time RP. $(X(t))_{t \in \mathbb{R}^+}$
  - $X(t)$: bitcoin price at time t, which is a random variable.
  - Then, $X(t) : \Omega \mapsto \mathbb{R}$.
  - An outcome $\omega \in \Omega$ is a trajectory of prices over $[0, \infty)$
  - $X(3.7, \omega_1) = 3409$, $X(2, \omega_2) = 5000$, $X(7.8, \omega_3) = 2800$, etc.

Random Processes: Our Interest

- Question. Already studied a sequence (or a collection) of rvs $X_1, X_2, \ldots, X_n$. What's the difference?
  - Assume a discrete time random process for our discussion.

Random Processes: Our Interest

- Question. Already studied a sequence (or a collection) of rvs $X_1, X_2, \ldots, X_n$. What's the difference?
  - Assume a discrete time random process for our discussion.
- Physical difference: infinite sequence of $X_1, X_2, \ldots, .$

- Question. Already studied a sequence (or a collection) of rvs $X_1, X_2, \ldots, X_n$. What's the difference?
  - Assume a discrete time random process for our discussion.
- Physical difference: infinite sequence of $X_1, X_2, \ldots, .$
  - Sample space? set of all outcomes?

KAIST EE

- Question. Already studied a sequence (or a collection) of rvs $X_1, X_2, \ldots, X_n$. What's the difference?
  - Assume a discrete time random process for our discussion.
- Physical difference: infinite sequence of $X_1, X_2, \ldots, .$
  - Sample space? set of all outcomes?
  - an outcome: an infinite sequence of sample values $x_1, x_2, \ldots,$

- Question. Already studied a sequence (or a collection) of rvs $X_1, X_2, \ldots, X_n$. What's the difference?
  - Assume a discrete time random process for our discussion.
- Physical difference: infinite sequence of $X_1, X_2, \ldots, .$
  - Sample space? set of all outcomes?
  - an outcome: an infinite sequence of sample values $x_1, x_2, \ldots,$
- Semantic difference: Understand i in X_i as time. Also, interesting questions are asked from the random process point of view.

- Question. Already studied a sequence (or a collection) of rvs $X_1, X_2, \ldots, X_n$. What's the difference?
  - Assume a discrete time random process for our discussion.
- Physical difference: infinite sequence of $X_1, X_2, \ldots, $.
  - Sample space? set of all outcomes?
  - an outcome: an infinite sequence of sample values $x_1, x_2, \ldots,$
- Semantic difference: Understand i in X_i as time. Also, interesting questions are asked from the random process point of view.
  - Dependence: How $X_1, X_2, \ldots$ are related to each other as a time series. Prediction of values in the future.

Random Processes: Our Interest

- Question. Already studied a sequence (or a collection) of rvs $X_1, X_2, \ldots, X_n$. What's the difference?
  - Assume a discrete time random process for our discussion.
- Physical difference: infinite sequence of $X_1, X_2, \ldots, .$
  - Sample space? set of all outcomes?
  - an outcome: an infinite sequence of sample values $x_1, x_2, \ldots,$
- Semantic difference: Understand i in X_i as time. Also, interesting questions are asked from the random process point of view.
  - Dependence: How $X_1, X_2, \ldots$ are related to each other as a time series. Prediction of values in the future.
  - Long-term behavior: What is the fraction of times that a stock price is above 3000?

- Question. Already studied a sequence (or a collection) of rvs $X_1, X_2, \ldots, X_n$. What's the difference?
  - Assume a discrete time random process for our discussion.
- Physical difference: infinite sequence of $X_1, X_2, \ldots, .$
  - Sample space? set of all outcomes?
  - an outcome: an infinite sequence of sample values $x_1, x_2, \ldots,$
- Semantic difference: Understand i in X_i as time. Also, interesting questions are asked from the random process point of view.
  - Dependence: How $X_1, X_2, \ldots$ are related to each other as a time series. Prediction of values in the future.
  - Long-term behavior: What is the fraction of times that a stock price is above 3000?
  - Other interesting questions, depending on the target random process

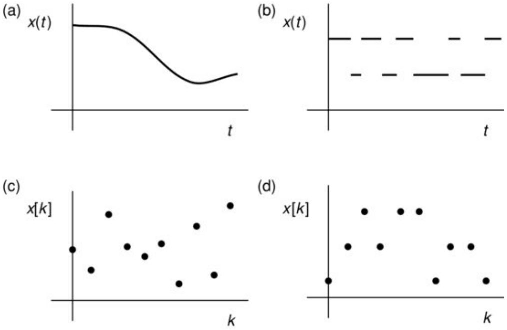
- Types of time and value

(a) continuous time, continuous value

(b) continuous time, discrete value

(c) discrete time, continuous value

(d) discrete time, discrete value

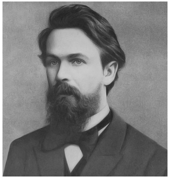


- The simplest RP
- discrete time

Jacob Bernoulli (1654 - 1705), Swiss

Simeon Denis Poisson (1781 - 1840), France

Andrey Markov (1856 - 1922), Russia



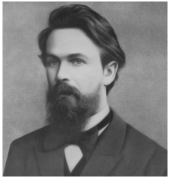
Random Processes in This Course

- The simplest RP
- discrete time
- $X_i \perp\!\!\!\perp \{X_{i-1}, X_{i-2}, \ldots, X_1\}$

Jacob Bernoulli (1654 - 1705), Swiss

Simeon Denis Poisson (1781 - 1840), France

Andrey Markov (1856 - 1922), Russia



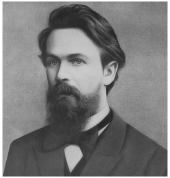
Random Processes in This Course

- The simplest RP
- discrete time
- $X_i \perp\!\!\!\perp \{X_{i-1}, X_{i-2}, \ldots, X_1\}$
- Bernoulli Process (BP)

Jacob Bernoulli (1654 - 1705), Swiss

Simeon Denis Poisson (1781 - 1840), France

Andrey Markov (1856 - 1922), Russia



Random Processes in This Course

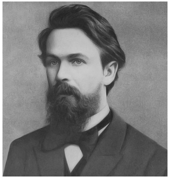
- The simplest RP
- discrete time
- $X_i \perp\!\!\!\perp \{X_{i-1}, X_{i-2}, \dots, X_1\}$
- Bernoulli Process (BP)

Jacob Bernoulli (1654 - 1705), Swiss

- The simplest RP
- Continous time version of BP

Simeon Denis Poisson (1781 - 1840), France

Andrey Markov (1856 - 1922), Russia



Random Processes in This Course

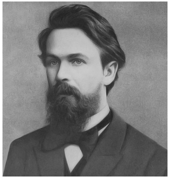
- The simplest RP
- discrete time
- $X_i \perp\!\!\!\perp \{X_{i-1}, X_{i-2}, \ldots, X_1\}$
- Bernoulli Process (BP)

Jacob Bernoulli (1654 - 1705), Swiss

- The simplest RP
- Continous time version of BP
- $[X(s)]_{s=0}^{t} \perp\!\!\!\perp [X(s)]_{s=t}^{t+a}$

Simeon Denis Poisson (1781 - 1840), France

Andrey Markov (1856 - 1922), Russia



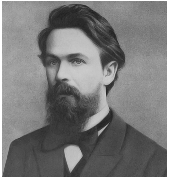
- The simplest RP
- discrete time
- $X_i \perp\!\!\!\perp \{X_{i-1}, X_{i-2}, \ldots, X_1\}$
- Bernoulli Process (BP)

Jacob Bernoulli (1654 - 1705), Swiss

- The simplest RP
- Continous time version of BP
- $[X(s)]_{s=0}^{t} \perp\!\!\!\perp [X(s)]_{s=t}^{t+a}$
- Poisson Process (PP)

Simeon Denis Poisson (1781 - 1840), France

Andrey Markov (1856 - 1922), Russia



- The simplest RP
- discrete time
- $X_i \perp\!\!\!\perp \{X_{i-1}, X_{i-2}, \ldots, X_1\}$
- Bernoulli Process (BP)

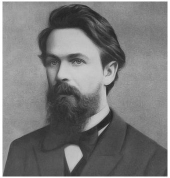
Jacob Bernoulli (1654 - 1705), Swiss

- The simplest RP
- Continous time version of BP
- $[X(s)]_{s=0}^{t} \perp\!\!\!\perp [X(s)]_{s=t}^{t+a}$
- Poisson Process (PP)

Simeon Denis Poisson (1781 - 1840), France

- One-step more general than BP/PP
- discrete time

Andrey Markov (1856 - 1922), Russia



Random Processes in This Course

- The simplest RP
- discrete time
- $X_i \perp\!\!\!\perp \{X_{i-1}, X_{i-2}, \ldots, X_1\}$
- Bernoulli Process (BP)

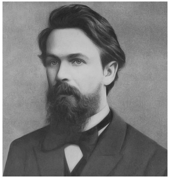
Jacob Bernoulli (1654 - 1705), Swiss

- The simplest RP
- Continous time version of BP
- $[X(s)]_{s=0}^{t} \perp\!\!\!\perp [X(s)]_{s=t}^{t+a}$
- Poisson Process (PP)

Simeon Denis Poisson (1781 - 1840), France

- One-step more general than BP/PP
- discrete time
- X_i depends on X_{i-1}, but $\perp\!\!\!\perp \{X_{i-2}, X_{i-2}, \ldots, X_1\}$

Andrey Markov (1856 - 1922), Russia



Random Processes in This Course

- The simplest RP
- discrete time
- $X_i \perp\!\!\!\perp \{X_{i-1}, X_{i-2}, \ldots, X_1\}$
- Bernoulli Process (BP)

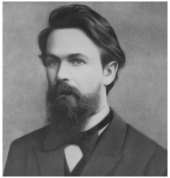
Jacob Bernoulli (1654 - 1705), Swiss

- The simplest RP
- Continous time version of BP
- $[X(s)]_{s=0}^{t} \perp\!\!\!\perp [X(s)]_{s=t}^{t+a}$
- Poisson Process (PP)

Simeon Denis Poisson (1781 - 1840), France

- One-step more general than BP/PP
- discrete time
- X_i depends on X_{i-1}, but $\perp\!\!\!\perp \{X_{i-2}, X_{i-2}, \ldots, X_1\}$
- Markov Chain (MC)

Andrey Markov (1856 - 1922), Russia



- The simplest RP
- discrete time
- $X_i \perp\!\!\!\perp \{X_{i-1}, X_{i-2}, \ldots, X_1\}$
- Bernoulli Process (BP)
- "today" independent of "past"

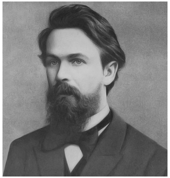
Jacob Bernoulli (1654 - 1705), Swiss

- The simplest RP
- Continous time version of BP
- $[X(s)]_{s=0}^{t} \perp\!\!\!\perp [X(s)]_{s=t}^{t+a}$
- Poisson Process (PP)
- "today" independent of "past"

Simeon Denis Poisson (1781 - 1840), France

- One-step more general than BP/PP
- discrete time
- X_i depends on X_{i-1}, but $\perp\!\!\!\perp \{X_{i-2}, X_{i-2}, \ldots, X_1\}$
- Markov Chain (MC)
- "today" depends only on "yesterday"

Andrey Markov (1856 - 1922), Russia



Roadmap

(1) Introduction of Random Processes

(2) Bernoulli Processes

(3) Poisson Processes: Poisson RV and Basic Idea

(4) Poisson Processes: Definition and Properties

(5) Playing with Bernoulli and Poisson Processes

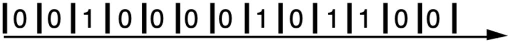
(6) Random Incidence Phenomenon

- At each "minute", we toss a coin with probability of head $0 < p < 1$.

- At each "minute", we toss a coin with probability of head $0 < p < 1$.
  - Sequence of lottery wins/looses
  - Customers (each second) to a bank
  - Clicks (at each time slot) to server

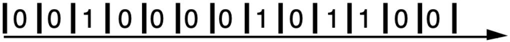
Bernoulli Process

- At each "minute", we toss a coin with probability of head $0 < p < 1$.
  - Sequence of lottery wins/looses
  - Customers (each second) to a bank
  - Clicks (at each time slot) to server
- A sequence of independent Bernoulli trials $X_1, X_2, \ldots,$
  - We call index 1, 2, ... time slots (or simply slots)



Bernoulli Process

- At each "minute", we toss a coin with probability of head $0 < p < 1$.
  - Sequence of lottery wins/looses
  - Customers (each second) to a bank
  - Clicks (at each time slot) to server
- A sequence of independent Bernoulli trials $X_1, X_2, \ldots,$
  - We call index 1, 2, ... time slots (or simply slots)



- Discrete time, discrete value

Bernoulli Process

- At each "minute", we toss a coin with probability of head $0 < p < 1$.
  - Sequence of lottery wins/looses
  - Customers (each second) to a bank
  - Clicks (at each time slot) to server
- A sequence of independent Bernoulli trials $X_1, X_2, \ldots,$
  - We call index 1, 2, ... time slots (or simply slots)

| 0 | 0 | 1 | 0 | 0 | 0 | 0 | 1 | 0 | 1 | 1 | 0 | 0 |

- Discrete time, discrete value
- One of the simplest random processes
- A type of "arrival" process

Please pause the video and write down the questions that you want to ask about Bernoulli process. VIDEO PAUSE

Q1.

Q2.

Q3.

Q4.

Q5.

(Q1) # of arrivals in the first n slots?

(Q1) # of arrivals in the first n slots?

- $S_n = X_1 + X_2 + \cdots + X_n$

(Q1) # of arrivals in the first n slots?

- $S_n = X_1 + X_2 + \cdots + X_n$
- $S_n \sim \text{Binomial}(n, p)$

(Q1) # of arrivals in the first n slots?

- $S_n = X_1 + X_2 + \cdots + X_n$
- $S_n \sim \text{Binomial}(n, p)$
- $\mathbb{E}(S_n) = np$, $\text{var}(S_n) = np(1-p)$

Number of arrivals and Time until the first arrival

(Q1) # of arrivals in the first n slots?

- $S_n = X_1 + X_2 + \cdots + X_n$
- $S_n \sim \text{Binomial}(n, p)$
- $\mathbb{E}(S_n) = np$, $\text{var}(S_n) = np(1-p)$
- This will hold for any n consecutive slots.

(Q2) # of slots T_1 until the first arrival?

(Q1) # of arrivals in the first n slots?

- $S_n = X_1 + X_2 + \cdots + X_n$
- $S_n \sim \text{Binomial}(n, p)$
- $\mathbb{E}(S_n) = np$, $\text{var}(S_n) = np(1-p)$
- This will hold for any n consecutive slots.

(Q2) # of slots T_1 until the first arrival?

- $T_1 \sim \text{Geom}(p)$

Number of arrivals and Time until the first arrival

(Q1) # of arrivals in the first n slots?

- $S_n = X_1 + X_2 + \cdots + X_n$
- $S_n \sim \text{Binomial}(n, p)$
- $\mathbb{E}(S_n) = np$, $\text{var}(S_n) = np(1-p)$
- This will hold for any n consecutive slots.

(Q2) # of slots T_1 until the first arrival?

- $T_1 \sim \text{Geom}(p)$
- $\mathbb{E}(T_1) = 1/p$, $\text{var}(T_1) = \frac{1-p}{p^2}$

(Q1) # of arrivals in the first n slots?

- $S_n = X_1 + X_2 + \cdots + X_n$
- $S_n \sim \text{Binomial}(n, p)$
- $\mathbb{E}(S_n) = np$, $\text{var}(S_n) = np(1-p)$
- This will hold for any n consecutive slots.

(Q2) # of slots T_1 until the first arrival?

- $T_1 \sim \text{Geom}(p)$
- $\mathbb{E}(T_1) = 1/p$, $\text{var}(T_1) = \frac{1-p}{p^2}$

- T_1 is geometric? Memoryless

(Q1) # of arrivals in the first n slots?

- $S_n = X_1 + X_2 + \cdots + X_n$
- $S_n \sim \text{Binomial}(n, p)$
- $\mathbb{E}(S_n) = np$, $\text{var}(S_n) = np(1-p)$
- This will hold for any n consecutive slots.

(Q2) # of slots T_1 until the first arrival?

- $T_1 \sim \text{Geom}(p)$
- $\mathbb{E}(T_1) = 1/p$, $\text{var}(T_1) = \frac{1-p}{p^2}$

- T_1 is geometric? Memoryless
- Conditioned no first arrival at some time, what's the distribution of the remaining time until the first arrival?
- Still, geometric.

(Q1) # of arrivals in the first n slots?

- $S_n = X_1 + X_2 + \cdots + X_n$
- $S_n \sim \text{Binomial}(n, p)$
- $\mathbb{E}(S_n) = np$, $\text{var}(S_n) = np(1-p)$
- This will hold for any n consecutive slots.

(Q2) # of slots T_1 until the first arrival?

- $T_1 \sim \text{Geom}(p)$
- $\mathbb{E}(T_1) = 1/p$, $\text{var}(T_1) = \frac{1-p}{p^2}$

- T_1 is geometric? Memoryless
- Conditioned no first arrival at some time, what's the distribution of the remaining time until the first arrival?
- Still, geometric.
- But, more than that, as we will see. Independence across time slots leads to many useful properties, allowing the quick solution of many problems.

Independence across slots $\Longrightarrow$ the fresh-start anytime when I look at the process?

(Q3) $U = X_1 + X_2 \perp\!\!\!\perp V = X_5 + X_6$?

Memoryless and Fresh-start after Deterministic time n

Independence across slots $\Longrightarrow$ the fresh-start anytime when I look at the process?

(Q3) $U = X_1 + X_2 \perp\!\!\!\perp V = X_5 + X_6$?

- Yes
- Because X_is are independent

Independence across slots $\Longrightarrow$ the fresh-start anytime when I look at the process?

(Q3) $U = X_1 + X_2 \perp\!\!\!\perp V = X_5 + X_6$?

- Yes
- Because X_is are independent

(Q4) After time $n = 6$, I start to look at the process $(X_n)_{n=6}^{\infty}$?

Independence across slots $\Longrightarrow$ the fresh-start anytime when I look at the process?

(Q3) $U = X_1 + X_2 \perp\!\!\!\perp V = X_5 + X_6$?

- Yes
- Because X_is are independent

(Q4) After time $n = 6$, I start to look at the process $(X_n)_{n=6}^{\infty}$?

- $(X_1, \ldots, X_5) \perp\!\!\!\perp (X_n)_{n=6}^{\infty}$

Memoryless and Fresh-start after Deterministic time n

Independence across slots $\Longrightarrow$ the fresh-start anytime when I look at the process?

(Q3) $U = X_1 + X_2 \perp\!\!\!\perp V = X_5 + X_6$?

- Yes
- Because X_is are independent

(Q4) After time $n = 6$, I start to look at the process $(X_n)_{n=6}^{\infty}$?

- $(X_1, \ldots, X_5) \perp\!\!\!\perp (X_n)_{n=6}^{\infty}$
- Fresh-start after a deterministic time n (doesn't matter what happened until $n = 5$.

Memoryless and Fresh-start after Deterministic time n

Independence across slots $\implies$ the fresh-start anytime when I look at the process?

(Q3) $U = X_1 + X_2 \perp\!\!\!\perp V = X_5 + X_6$?

- Yes
- Because X_is are independent

(Q4) After time $n = 6$, I start to look at the process $(X_n)_{n=6}^{\infty}$?

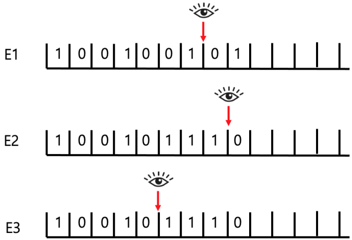
- $(X_1, \ldots, X_5) \perp\!\!\!\perp (X_n)_{n=6}^{\infty}$
- Fresh-start after a deterministic time n (doesn't matter what happened until $n = 5$.
- If you watch the on-going Bernoulli process(p) from some time n, you still see the same Bernoulli process(p).

(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

- This means that the time I start to look at the process is a random variable.

(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

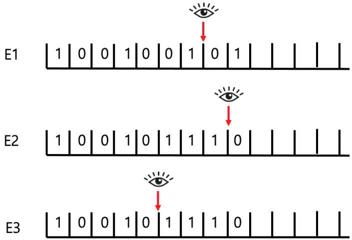
- This means that the time I start to look at the process is a random variable.
- Examples of N



(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

- This means that the time I start to look at the process is a random variable.
- Examples of N

E1. Time of 3rd arrival

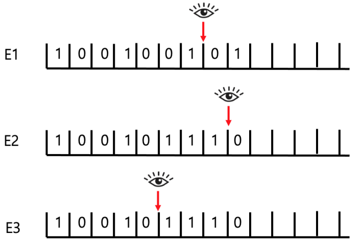


(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

- This means that the time I start to look at the process is a random variable.
- Examples of N

E1. Time of 3rd arrival

E2. First time when 3 consecutive arrivals have been observed



Fresh-start after Random time N (1)

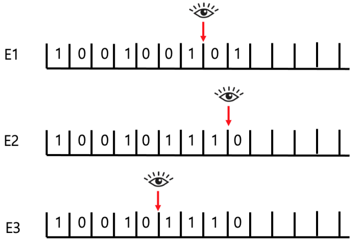
(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

- This means that the time I start to look at the process is a random variable.
- Examples of N

E1. Time of 3rd arrival

E2. First time when 3 consecutive arrivals have been observed

E3. Time just before 3 consecutive arrivals



Fresh-start after Random time N (1)

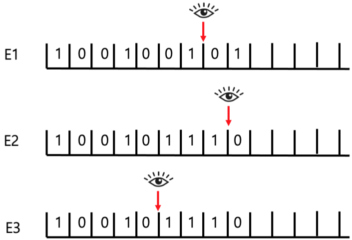
(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

- This means that the time I start to look at the process is a random variable.
- Examples of N

E1. Time of 3rd arrival

E2. First time when 3 consecutive arrivals have been observed

E3. Time just before 3 consecutive arrivals



- Difference of N from n
  - The time when I watch the on-going Bernoulli process is random.
  - N is a random variable, i.e., $N : \Omega \mapsto \mathbb{R}$. What is Ω?

KAIST EE

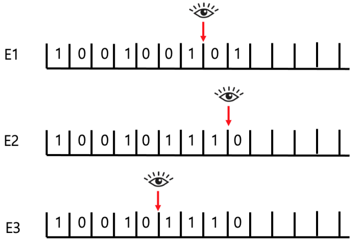
(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

- This means that the time I start to look at the process is a random variable.
- Examples of N

E1. Time of 3rd arrival

E2. First time when 3 consecutive arrivals have been observed

E3. Time just before 3 consecutive arrivals



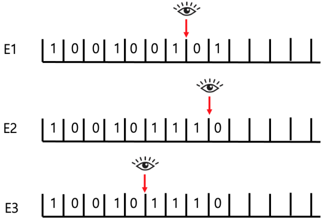
- Difference of N from n
  - The time when I watch the on-going Bernoulli process is random.
  - N is a random variable, i.e., $N : \Omega \mapsto \mathbb{R}$. What is Ω?
- Do we experience the fresh-start for any N? **E1, E2, and E3?**

(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

E1. Time of 3rd arrival

E2. First time when 3 consecutive arrivals have been observed

E3. Time just before 3 consecutive arrivals



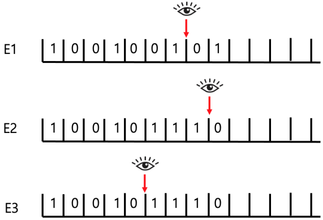
Fresh-start after Random N (2)

(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

E1. Time of 3rd arrival

E2. First time when 3 consecutive arrivals have been observed

E3. Time just before 3 consecutive arrivals



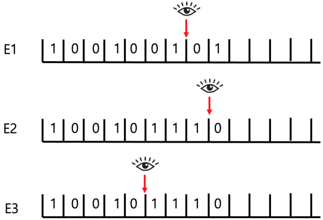
E1. When I watch the process, N has been already determined. Yes

(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

E1. Time of 3rd arrival

E2. First time when 3 consecutive arrivals have been observed

E3. Time just before 3 consecutive arrivals



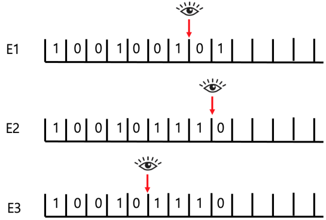
E1. When I watch the process, N has been already determined. Yes

E2. Same as **E1.** Yes

Fresh-start after Random N (2)

(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

- **E1.** Time of 3rd arrival
- **E2.** First time when 3 consecutive arrivals have been observed
- **E3.** Time just before 3 consecutive arrivals



- **E1.** When I watch the process, N has been already determined. Yes
- **E2.** Same as **E1.** Yes
- **E3.** Need the future knowledge. '111' does not become random. No

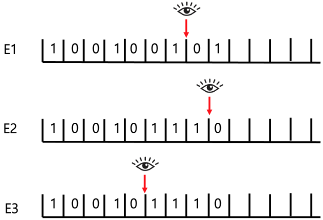
Fresh-start after Random N (2)

(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \ldots)$? Fresh-start even after random N?

E1. Time of 3rd arrival

E2. First time when 3 consecutive arrivals have been observed

E3. Time just before 3 consecutive arrivals



E1. When I watch the process, N has been already determined. Yes

E2. Same as **E1.** Yes

E3. Need the future knowledge. '111' does not become random. No

- The question of $N = n$? can be answered just from the knowledge about $X_1, X_2, \ldots, X_n$? Then, Yes! (see pp. 301 for more formal description)

- In probability theory, a random time N is said to be a stopping time, if the question of "$N = n$?" can be answered only from the present and the past knowledge of $X_1, X_2, \ldots, X_n$.

- In probability theory, a random time N is said to be a stopping time, if the question of "$N = n$?" can be answered only from the present and the past knowledge of $X_1, X_2, \ldots, X_n$.
- `https://en.wikipedia.org/wiki/Stopping_time`

Stopping Time

- In probability theory, a random time N is said to be a stopping time, if the question of "$N = n$?" can be answered only from the present and the past knowledge of $X_1, X_2, \ldots, X_n$.
- https://en.wikipedia.org/wiki/Stopping_time
- Fresh-start after N in Bernoulli process? Yes, if N is a stopping time.

- In probability theory, a random time N is said to be a stopping time, if the question of "$N = n$?" can be answered only from the present and the past knowledge of $X_1, X_2, \ldots, X_n$.
- https://en.wikipedia.org/wiki/Stopping_time
- Fresh-start after N in Bernoulli process? Yes, if N is a stopping time.
- Please think about two examples of stopping time and not. VIDEO PAUSE

- In probability theory, a random time N is said to be a stopping time, if the question of "$N = n$?" can be answered only from the present and the past knowledge of $X_1, X_2, \ldots, X_n$.
- https://en.wikipedia.org/wiki/Stopping_time
- Fresh-start after N in Bernoulli process? Yes, if N is a stopping time.
- Please think about two examples of stopping time and not. VIDEO PAUSE
  - Yes. Time when 10 consecutive arrivals have been observed

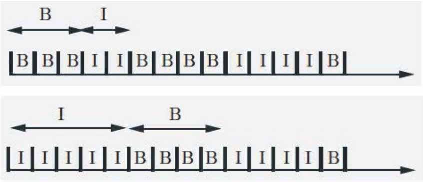
Stopping Time

- In probability theory, a random time N is said to be a stopping time, if the question of "$N = n$?" can be answered only from the present and the past knowledge of $X_1, X_2, \ldots, X_n$.
- `https://en.wikipedia.org/wiki/Stopping_time`
- Fresh-start after N in Bernoulli process? Yes, if N is a stopping time.
- Please think about two examples of stopping time and not. VIDEO PAUSE
  - Yes. Time when 10 consecutive arrivals have been observed
  - No. Time of 2nd arrival in 10 consecutive arrivals

Distribution of Busy Periods (1)

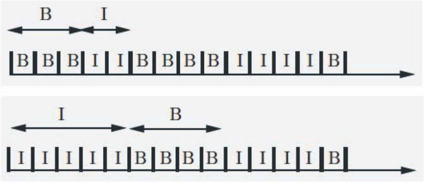
- Regard an arrival as a server being busy (just for our easy understanding)

- Regard an arrival as a server being busy (just for our easy understanding)
- First busy period B_1: starts with the first busy slot and ends just before the first subsequent idle slot



Distribution of Busy Periods (1)

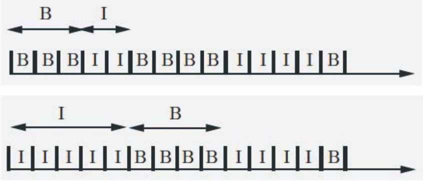
- Regard an arrival as a server being busy (just for our easy understanding)
- First busy period B_1: starts with the first busy slot and ends just before the first subsequent idle slot



- (Q6) B_1 is a random variable. Distribution of B_1?

Distribution of Busy Periods (1)

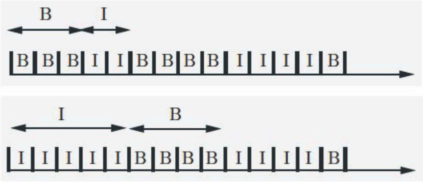
- Regard an arrival as a server being busy (just for our easy understanding)
- First busy period B_1: starts with the first busy slot and ends just before the first subsequent idle slot



- (Q6) B_1 is a random variable. Distribution of B_1?
- N: time of the first busy slot. N is a stopping time?

Distribution of Busy Periods (1)

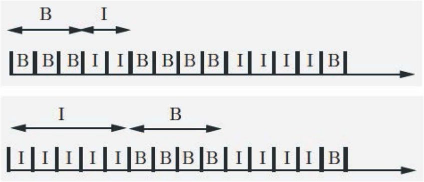
- Regard an arrival as a server being busy (just for our easy understanding)
- First busy period B_1: starts with the first busy slot and ends just before the first subsequent idle slot



- (Q6) B_1 is a random variable. Distribution of B_1?
- N: time of the first busy slot. N is a stopping time?
  - Yes. Thus, fresh-start after N. Because we can answer the question of $N = n$?, just using $X_1, X_2, \ldots, X_n$.

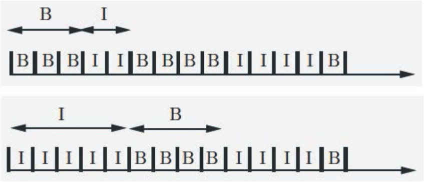
Distribution of Busy Periods (1)

- Regard an arrival as a server being busy (just for our easy understanding)
- First busy period B_1: starts with the first busy slot and ends just before the first subsequent idle slot



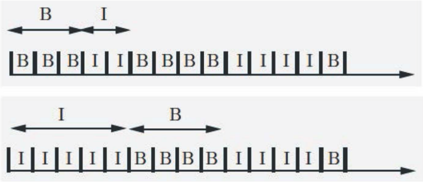
- (Q6) B_1 is a random variable. Distribution of B_1?
- N: time of the first busy slot. N is a stopping time?
  - Yes. Thus, fresh-start after N. Because we can answer the question of $N = n$?, just using $X_1, X_2, \ldots, X_n$.
- B_1 is geometric with parameter $(1 - p)$

Distribution of Busy Periods (2)

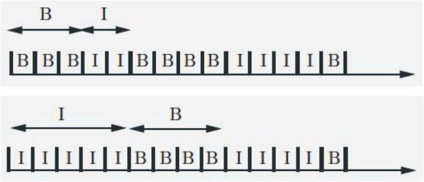


- Question. What about the second busy period B_2?

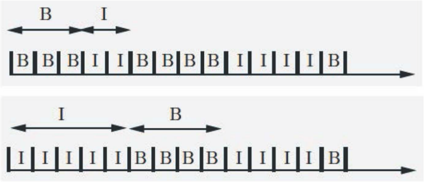
Distribution of Busy Periods (2)



- Question. What about the second busy period B_2?
- N: time of the first busy slot of the second busy period. N is a stopping time?

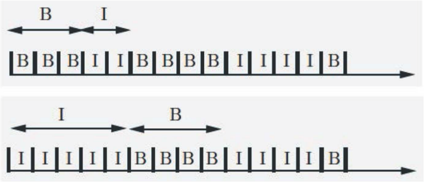


- Question. What about the second busy period B_2?
- N: time of the first busy slot of the second busy period. N is a stopping time?
- Yes. Thus, fresh-start after N.



- Question. What about the second busy period B_2?
- N: time of the first busy slot of the second busy period. N is a stopping time?
- Yes. Thus, fresh-start after N.
- Then, B_1 and B_2 are identically distributed as $\text{Geom}(1-p)$.

Distribution of Busy Periods (2)



- Question. What about the second busy period B_2?
- N: time of the first busy slot of the second busy period. N is a stopping time?
- Yes. Thus, fresh-start after N.
- Then, B_1 and B_2 are identically distributed as $\text{Geom}(1-p)$.
- $B_3, B_4, \ldots$?

Time of k-th arrival

- Time of the first arrival $Y_1 \sim \text{Geom}(p)$

Time of k-th arrival

- Time of the first arrival $Y_1 \sim \mathrm{Geom}(p)$

  (Q7) Time of the k-th arrival Y_k?

VIDEO PAUSE

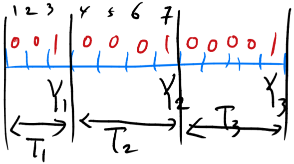
Time of k-th arrival

- Time of the first arrival $Y_1 \sim \text{Geom}(p)$

  (Q7) Time of the k-th arrival Y_k?

VIDEO PAUSE

- $T_k = Y_k - Y_{k-1}$: k-th inter-arrival ($k \geq 2$, $T_1 = Y_1$)



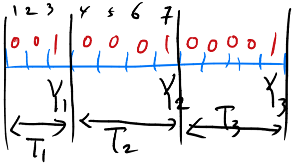
Time of k-th arrival

- Time of the first arrival $Y_1 \sim \text{Geom}(p)$

  (Q7) Time of the k-th arrival Y_k?

VIDEO PAUSE

- $T_k = Y_k - Y_{k-1}$: k-th inter-arrival ($k \geq 2$, $T_1 = Y_1$)
- $Y_k = T_1 + T_2 + \ldots + T_k$.

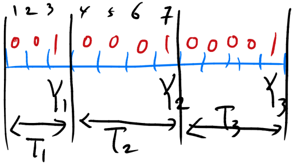


- Time of the first arrival $Y_1 \sim \text{Geom}(p)$

  (Q7) Time of the k-th arrival Y_k?

VIDEO PAUSE

- $T_k = Y_k - Y_{k-1}$: k-th inter-arrival ($k \geq 2$, $T_1 = Y_1$)
- $Y_k = T_1 + T_2 + \ldots + T_k$.



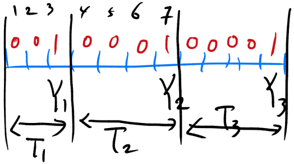
- After each T_k, the fresh-start occurs.

- Time of the first arrival $Y_1 \sim \text{Geom}(p)$

(Q7) Time of the k-th arrival Y_k?

VIDEO PAUSE

- $T_k = Y_k - Y_{k-1}$: k-th inter-arrival ($k \geq 2,\ T_1 = Y_1$)
- $Y_k = T_1 + T_2 + \ldots + T_k$.



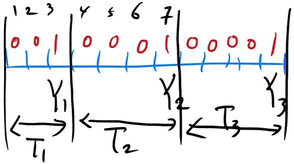
- After each T_k, the fresh-start occurs.
- $\{T_i\}$ are i.i.d. and $\sim \text{Geom}(p)$

- Time of the first arrival $Y_1 \sim \text{Geom}(p)$

  (Q7) Time of the k-th arrival Y_k?

VIDEO PAUSE

- $T_k = Y_k - Y_{k-1}$: k-th inter-arrival ($k \geq 2$, $T_1 = Y_1$)
- $Y_k = T_1 + T_2 + \ldots + T_k$.



- After each T_k, the fresh-start occurs.
- $\{T_i\}$ are i.i.d. and $\sim \text{Geom}(p)$
- We know Y_k's expectation and variance: $\mathbb{E}[Y_k] = \frac{k}{p}$, $\text{var}[Y_k] = \frac{k(1-p)}{p^2}$, but its distribution?

PMF of Y_k

- $Y_k = T_1 + T_2 + \ldots + T_k$.
- $\{T_i\}$ are i.i.d. and $\sim \text{Geom}(p)$

PMF of Y_k

- $Y_k = T_1 + T_2 + \ldots + T_k$.
- $\{T_i\}$ are i.i.d. and $\sim \text{Geom}(p)$

  $\mathbb{P}(Y_k = t) = \mathbb{P}\,(X_k = 1 \text{ and } k-1 \text{ arrivals during the first } t-1 \text{ slots})$

PMF of Y_k

- $Y_k = T_1 + T_2 + \ldots + T_k$.
- $\{T_i\}$ are i.i.d. and $\sim \text{Geom}(p)$

$$
\begin{aligned}
\mathbb{P}(Y_k = t) &= \mathbb{P}\left(X_k = 1 \text{ and } k-1 \text{ arrivals during the first } t-1 \text{ slots}\right) \\
&= \mathbb{P}\left(X_k = 1\right) \cdot \mathbb{P}\left(k-1 \text{ arrivals during the first } t-1 \text{ slots}\right)
\end{aligned}
$$

PMF of Y_k

- $Y_k = T_1 + T_2 + \ldots + T_k$.
- $\{T_i\}$ are i.i.d. and $\sim \text{Geom}(p)$

$$\begin{aligned}\mathbb{P}(Y_k = t) &= \mathbb{P}\left(X_k = 1 \text{ and } k-1 \text{ arrivals during the first } t-1 \text{ slots}\right) \\ &= \mathbb{P}\left(X_k = 1\right) \cdot \mathbb{P}\left(k-1 \text{ arrivals during the first } t-1 \text{ slots}\right) \\ &= p \times \binom{t-1}{k-1} p^{k-1}(1-p)^{t-k} = \binom{t-1}{k-1} p^k (1-p)^{t-k}, \quad t = k, k+1, \ldots\end{aligned}$$

Pascal Random Variable with Parameter (k, p)

- In the sequence of Bernoulli trials, the time Y_k of k-th success

Pascal Random Variable with Parameter (k, p)

- In the sequence of Bernoulli trials, the time Y_k of k-th success
- PMF of Y_k

$$\mathbb{P}(Y_k = t) = \begin{cases} \binom{t-1}{k-1} p^k (1-p)^{t-k}, & \text{if} \quad t = k, k+1, \ldots \\ 0, & \text{if} \quad t = 1, 2, \ldots k-1 \end{cases}$$

Pascal Random Variable with Parameter (k, p)

- In the sequence of Bernoulli trials, the time Y_k of k-th success
- PMF of Y_k

$$\mathbb{P}(Y_k = t) = \begin{cases} \binom{t-1}{k-1} p^k (1-p)^{t-k}, & \text{if} \quad t = k, k+1, \ldots \\ 0, & \text{if} \quad t = 1, 2, \ldots k-1 \end{cases}$$

- $\text{Pascal}(1, p) = \text{Geom}(p)$

Roadmap

(1) Introduction of Random Processes

(2) Bernoulli Processes

(3) Poisson Processes: Poisson RV and Basic Idea

(4) Poisson Processes: Definition and Properties

(5) Playing with Bernoulli and Poisson Processes

Background: Poisson rv X with parameter λ (1)

- A random variable $S \sim \text{Bin}(n, p)$: Models the number of successes in a given number n of independent trials with success probability p.

$$p_S(k) = \frac{n!}{(n-k)!k!} p^k (1-p)^{n-k}$$

Background: Poisson rv X with parameter λ (1)

- A random variable $S \sim \text{Bin}(n, p)$: Models the number of successes in a given number n of independent trials with success probability p.

$$p_S(k) = \frac{n!}{(n-k)!k!} p^k (1-p)^{n-k}$$

- Our interest: very large n and very small p, such that $np = \lambda$, i.e., $\lim_{n \to \infty} p_S(k)$?

$$p_S(k) =$$

Background: Poisson rv X with parameter λ (1)

- A random variable $S \sim \text{Bin}(n, p)$: Models the number of successes in a given number n of independent trials with success probability p.

$$p_S(k) = \frac{n!}{(n-k)!k!} p^k (1-p)^{n-k}$$

- Our interest: very large n and very small p, such that $np = \lambda$, i.e., $\lim_{n \to \infty} p_S(k)$?

$$p_S(k) = \frac{n(n-1)\cdots(n-k+1)}{k!} \cdot \frac{\lambda^k}{n^k} \cdot \left(1 - \frac{\lambda}{n}\right)^{n-k}$$

KAIST EE

- A random variable $S \sim \text{Bin}(n, p)$: Models the number of successes in a given number n of independent trials with success probability p.

$$p_S(k) = \frac{n!}{(n-k)!k!} p^k (1-p)^{n-k}$$

- Our interest: very large n and very small p, such that $np = \lambda$, i.e., $\lim_{n \to \infty} p_S(k)$?

$$\begin{aligned} p_S(k) &= \frac{n(n-1)\cdots(n-k+1)}{k!} \cdot \frac{\lambda^k}{n^k} \cdot \left(1 - \frac{\lambda}{n}\right)^{n-k} \\ &= \frac{n}{n} \cdot \frac{(n-1)}{n} \cdots \frac{(n-k+1)}{n} \cdot \frac{\lambda^k}{k!} \cdot \left(1 - \frac{\lambda}{n}\right)^{n} \cdot \left(1 - \frac{\lambda}{n}\right)^{-k} \end{aligned}$$

- A random variable $S \sim \text{Bin}(n, p)$: Models the number of successes in a given number n of independent trials with success probability p.

$$p_S(k) = \frac{n!}{(n-k)!k!} p^k (1-p)^{n-k}$$

- Our interest: very large n and very small p, such that $np = \lambda$, i.e., $\lim_{n \to \infty} p_S(k)$?

$$\begin{aligned} p_S(k) &= \frac{n(n-1)\cdots(n-k+1)}{k!} \cdot \frac{\lambda^k}{n^k} \cdot \left(1 - \frac{\lambda}{n}\right)^{n-k} \\ &= \frac{n}{n} \cdot \frac{(n-1)}{n} \cdots \frac{(n-k+1)}{n} \cdot \frac{\lambda^k}{k!} \cdot \left(1 - \frac{\lambda}{n}\right)^{n} \cdot \left(1 - \frac{\lambda}{n}\right)^{-k} \xrightarrow{n \to \infty} e^{-\lambda} \frac{\lambda^k}{k!} \end{aligned}$$

Background: Poisson rv X with parameter λ (2)

Background: Poisson rv X with parameter λ (2)

- A Poisson random variable Z with parameter λ takes nonnegative integer values, whose PMF is:

$$p_Z(k) = e^{-\lambda}\frac{\lambda^k}{k!}, \quad k = 0, 1, \ldots$$

Background: Poisson rv X with parameter λ (2)

- A Poisson random variable Z with parameter λ takes nonnegative integer values, whose PMF is:

$$p_Z(k) = e^{-\lambda}\frac{\lambda^k}{k!}, \quad k = 0, 1, \ldots$$

- Infinitely many slots (n) with the infinitely small slot duration (thus infinitely small success probabilty $p = \lambda/n$)

Background: Poisson rv X with parameter λ (2)

- A Poisson random variable Z with parameter λ takes nonnegative integer values, whose PMF is:
$$p_Z(k) = e^{-\lambda}\frac{\lambda^k}{k!}, \quad k = 0, 1, \ldots$$
- Infinitely many slots (n) with the infinitely small slot duration (thus infinitely small success probabilty $p = \lambda/n$)
- $\mathbb{E}(Z) = \lambda$ (because $\lambda = np$ is the mean of binomial rv)

Background: Poisson rv X with parameter λ (2)

- A Poisson random variable Z with parameter λ takes nonnegative integer values, whose PMF is:

$$p_Z(k) = e^{-\lambda}\frac{\lambda^k}{k!}, \quad k = 0, 1, \ldots$$

- Infinitely many slots (n) with the infinitely small slot duration (thus infinitely small success probabilty $p = \lambda/n$)

- $\mathbb{E}(Z) = \lambda$ (because $\lambda = np$ is the mean of binomial rv)

- $\text{var}(Z) = \lambda$ (because $np(1-p)$ is the variance of binomial rv)

Example: Poisson Approximation

- A packet consisting of a string of n symbols is transmitted over a noisy channel.

Example: Poisson Approximation

- A packet consisting of a string of n symbols is transmitted over a noisy channel.
- Each symbol: errorneous transimission with probability of 0.0001, independent of other symbols. Incorrect transmission is when at least one symbol is in error.

Example: Poisson Approximation

- A packet consisting of a string of n symbols is transmitted over a noisy channel.
- Each symbol: erroneous transimission with probability of 0.0001, independent of other symbols. Incorrect transmission is when at least one symbol is in error.
- Question. How small should n be in order for the probability of incorrect transmission to be less than 0.001?

Example: Poisson Approximation

- A packet consisting of a string of n symbols is transmitted over a noisy channel.
- Each symbol: erroneous transimission with probability of 0.0001, independent of other symbols. Incorrect transmission is when at least one symbol is in error.
- Question. How small should n be in order for the probability of incorrect transmission to be less than 0.001?
- p is very small, and n is reasonably large $\rightarrow$ Poisson approximation

Example: Poisson Approximation

- A packet consisting of a string of n symbols is transmitted over a noisy channel.
- Each symbol: errorneous transimission with probability of 0.0001, independent of other symbols. Incorrect transmission is when at least one symbol is in error.
- Question. How small should n be in order for the probability of incorrect transmission to be less than 0.001?
- p is very small, and n is reasonably large $\rightarrow$ Poisson approximation
- Prob. of incorrect transmission $=$
  $1 - \mathbb{P}(\text{no symbol error}) = 1 - e^{-\lambda} = 1 - e^{-0.0001n} < 0.0001$

Example: Poisson Approximation

- A packet consisting of a string of n symbols is transmitted over a noisy channel.
- Each symbol: errorneous transimission with probability of 0.0001, independent of other symbols. Incorrect transmission is when at least one symbol is in error.
- Question. How small should n be in order for the probability of incorrect transmission to be less than 0.001?
- p is very small, and n is reasonably large $\rightarrow$ Poisson approximation
- Prob. of incorrect transmission $=$ $1 - \mathbb{P}(\text{no symbol error}) = 1 - e^{-\lambda} = 1 - e^{-0.0001n} < 0.0001$
- $n < \frac{-\ln 0.999}{0.0001} = 10.005$

- Remind. Geometric vs. Exponential L4(3)
  - Two rvs with memoryless property
  - continuous system = discrete system with infintely many slots whose duration is infinitley small.

L4(3)

- Remind. Geometric vs. Exponential
  - Two rvs with memoryless property
  - continuous system = discrete system with infintely many slots whose duration is infinitley small.
- Independence between what happens in a different time region

Design of Continuous Analogue of Bernoulli Process

L4(3)

- Remind. Geometric vs. Exponential
  - Two rvs with memoryless property
  - continuous system = discrete system with infintely many slots whose duration is infinitley small.
- Independence between what happens in a different time region
- Memoryless and fresh-start property

- Remind. Geometric vs. Exponential L4(3)
  - Two rvs with memoryless property
  - continuous system = discrete system with infintely many slots whose duration is infinitley small.
- Independence between what happens in a different time region
- Memoryless and fresh-start property
- Choose one of discrete or continuous versions for our modeling convenience.

L4(3)

- Remind. Geometric vs. Exponential
  - Two rvs with memoryless property
  - continuous system = discrete system with infintely many slots whose duration is infinitley small.
- Independence between what happens in a different time region
- Memoryless and fresh-start property
- Choose one of discrete or continuous versions for our modeling convenience.
- Question. How do we design the continuous analog of Bernoulli process?

L4(3)

- Remind. Geometric vs. Exponential
  - Two rvs with memoryless property
  - continuous system = discrete system with infintely many slots whose duration is infinitley small.
- Independence between what happens in a different time region
- Memoryless and fresh-start property
- Choose one of discrete or continuous versions for our modeling convenience.
- Question. How do we design the continuous analog of Bernoulli process?

  \- Key idea: Making it as a $\boxed{}$ of a sequence of Bernoulli processes

KAIST EE

- Remind. Geometric vs. Exponential L4(3)
  - Two rvs with memoryless property
  - continuous system = discrete system with infintely many slots whose duration is infinitley small.
- Independence between what happens in a different time region
- Memoryless and fresh-start property
- Choose one of discrete or continuous versions for our modeling convenience.
- Question. How do we design the continuous analog of Bernoulli process?

  - Key idea: Making it as a $\boxed{\text{limit}}$ of a sequence of Bernoulli processes

Design of Continuous Analogue of Bernoulli Process

L4(3)

- Remind. Geometric vs. Exponential
  - Two rvs with memoryless property
  - continuous system = discrete system with infintely many slots whose duration is infinitley small.
- Independence between what happens in a different time region
- Memoryless and fresh-start property
- Choose one of discrete or continuous versions for our modeling convenience.
- Question. How do we design the continuous analog of Bernoulli process?

  - Key idea: Making it as a $\boxed{\text{limit}}$ of a sequence of Bernoulli processes
- Need a "modeling sense" to make this possible. It's a good practice for engineers!
- VIDEO PAUSE

Key Design Idea to Develop a Continuous Twin (1)

- Continuous twin

Key Design Idea to Develop a Continuous Twin (1)

- Continuous twin
  - Key point: Understand the number of arrivals over a given interval $[0, \tau]$.

Key Design Idea to Develop a Continuous Twin (1)

- Continuous twin
  - Key point: Understand the number of arrivals over a given interval $[0, \tau]$.
  - Assume that it has some arrival rate λ ($\#$ of arrivals/unit time).

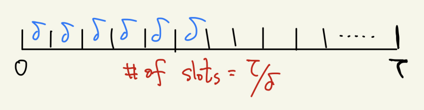
Key Design Idea to Develop a Continuous Twin (1)

- Continuous twin
  - Key point: Understand the number of arrivals over a given interval $[0, \tau]$.
  - Assume that it has some arrival rate λ ($\#$ of arrivals/unit time).
  - We know how to handle Bernoulli process with discrete time slots.

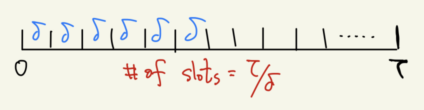
Key Design Idea to Develop a Continuous Twin (1)

KAIST EE

- Continuous twin
  - Key point: Understand the number of arrivals over a given interval $[0, \tau]$.
  - Assume that it has some arrival rate λ ($\#$ of arrivals/unit time).
  - We know how to handle Bernoulli process with discrete time slots.
- Divide $[0, \tau]$ into slots whose length $= \delta$. Then, $n = \#$ of slots $= \frac{\tau}{\delta}$.

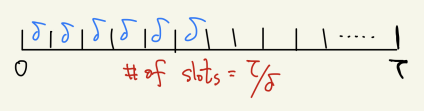


- Continuous twin
  - Key point: Understand the number of arrivals over a given interval $[0, \tau]$.
  - Assume that it has some arrival rate λ ($\#$ of arrivals/unit time).
  - We know how to handle Bernoulli process with discrete time slots.
- Divide $[0, \tau]$ into slots whose length $= \delta$. Then, $n = \#$ of slots $= \frac{\tau}{\delta}$.



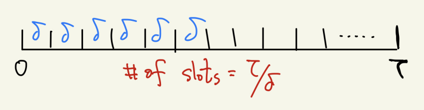
- What's the limit as $\delta \to 0$ (equivalently, $n \to \infty$)

Key Design Idea to Develop a Continuous Twin (2)



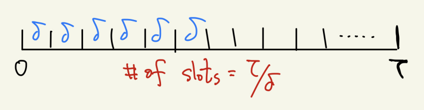
- Now, our design idea: during one time slot of length δ,

Key Design Idea to Develop a Continuous Twin (2)



- Now, our design idea: during one time slot of length δ,

$\mathbb{P}(1 \text{ arrival}) \propto$ arrival rate and slot length

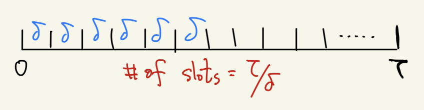


- Now, our design idea: during one time slot of length δ,

$\mathbb{P}$(1 arrival) $\propto$ arrival rate and slot length

$\mathbb{P}$($\geq$2 arrivals) $\propto$ something, but very small

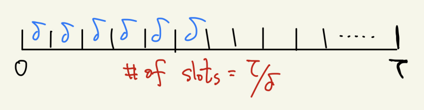
Key Design Idea to Develop a Continuous Twin (2)



- Now, our design idea: during one time slot of length δ,

$$
\begin{aligned}
\mathbb{P}(1 \text{ arrival}) &\propto \text{arrival rate and slot length} \\
\mathbb{P}(\geq 2 \text{ arrivals}) &\propto \text{something, but very small} \\
&\quad \text{for small sloth length} \\
\mathbb{P}(0 \text{ arrival}) &= 1 - \mathbb{P}(1 \text{ arrival or } \geq 2 \text{ arrivals})
\end{aligned}
$$

Key Design Idea to Develop a Continuous Twin (2)



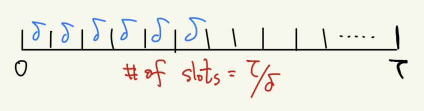
- Now, our design idea: during one time slot of length δ,

$$\mathbb{P}(\text{1 arrival}) \propto \text{arrival rate and slot length}$$

$$\mathbb{P}(\geq 2 \text{ arrivals}) \propto \text{something, but very small for small sloth length}$$

$$\mathbb{P}(\text{0 arrival}) = 1 - \mathbb{P}(\text{1 arrival or } \geq 2 \text{ arrivals})$$

$$\mathbb{P}(\text{1 arrival}) = \lambda\delta$$

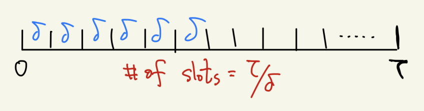


- Now, our design idea: during one time slot of length δ,

$$\begin{aligned}
\mathbb{P}(\text{1 arrival}) &\propto \text{arrival rate and slot length} \\
\mathbb{P}(\geq 2 \text{ arrivals}) &\propto \text{something, but very small} \\
&\quad \text{for small sloth length} \\
\mathbb{P}(\text{0 arrival}) &= 1 - \mathbb{P}(\text{1 arrival or } \geq 2 \text{ arrivals})
\end{aligned}$$

$$\begin{aligned}
\mathbb{P}(\text{1 arrival}) &= \lambda\delta \\
\mathbb{P}(\geq 2 \text{ arrivals}) &= 0
\end{aligned}$$

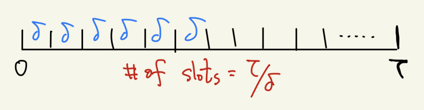
Key Design Idea to Develop a Continuous Twin (2)



- Now, our design idea: during one time slot of length δ,

$$
\begin{aligned}
\mathbb{P}(1 \text{ arrival}) &\propto \text{arrival rate and slot length} \\
\mathbb{P}(\geq 2 \text{ arrivals}) &\propto \text{something, but very small} \\
&\ \text{for small sloth length} \\
\mathbb{P}(0 \text{ arrival}) &= 1 - \mathbb{P}(1 \text{ arrival or } \geq 2 \text{ arrivals})
\end{aligned}
$$

$$
\begin{aligned}
\mathbb{P}(1 \text{ arrival}) &= \lambda\delta \\
\mathbb{P}(\geq 2 \text{ arrivals}) &= 0 \\
\mathbb{P}(0 \text{ arrival}) &= 1 - \lambda\delta
\end{aligned}
$$

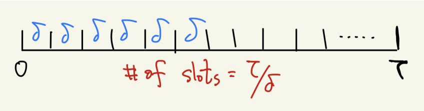


- Now, our design idea: during one time slot of length δ,

$$\mathbb{P}(1 \text{ arrival}) \propto \text{arrival rate and slot length}$$
$$\mathbb{P}(\geq 2 \text{ arrivals}) \propto \text{something, but very small for small sloth length}$$
$$\mathbb{P}(0 \text{ arrival}) = 1 - \mathbb{P}(1 \text{ arrival or } \geq 2 \text{ arrivals})$$

$$\mathbb{P}(1 \text{ arrival}) = \lambda\delta + o(\delta)$$
$$\mathbb{P}(\geq 2 \text{ arrivals}) = \quad + o(\delta)$$
$$\mathbb{P}(0 \text{ arrival}) = 1 - \lambda\delta + o(\delta)$$

Key Design Idea to Develop a Continuous Twin (2)



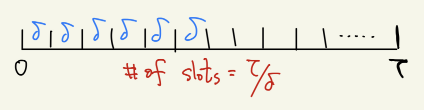
- Now, our design idea: during one time slot of length δ,

$$
\begin{aligned}
\mathbb{P}(1 \text{ arrival}) &\propto \text{arrival rate and slot length} \\
\mathbb{P}(\geq 2 \text{ arrivals}) &\propto \text{something, but very small} \\
&\quad\ \text{for small sloth length} \\
\mathbb{P}(0 \text{ arrival}) &= 1 - \mathbb{P}(1 \text{ arrival or } \geq 2 \text{ arrivals})
\end{aligned}
$$

$$
\begin{aligned}
\mathbb{P}(1 \text{ arrival}) &= \lambda\delta + o(\delta) \\
\mathbb{P}(\geq 2 \text{ arrivals}) &= \quad + o(\delta) \\
\mathbb{P}(0 \text{ arrival}) &= 1 - \lambda\delta + o(\delta)
\end{aligned}
$$

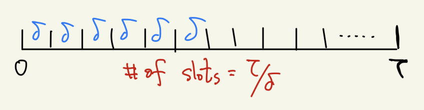
- $o(\delta)$: some function that goes to zero faster than δ.
  - Thus, for very small δ, $o(\delta)$ becomes negligible, compared to δ.
  - Example: $o(\delta) = \delta^{\alpha}$, where any $\alpha > 1$

Key Design Idea to Develop a Continuous Twin (3)



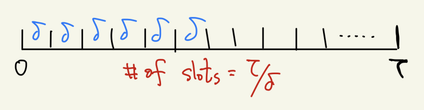
- Our interest: probability of k arrivals over $[0, \tau]$

Key Design Idea to Develop a Continuous Twin (3)



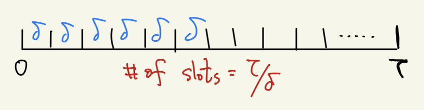
- Our interest: probability of k arrivals over $[0, \tau]$
- Given "small" δ, # of arrivals $\sim$ Bin(n, p), where $n = \tau/\delta$ and $p = \lambda\delta$

Key Design Idea to Develop a Continuous Twin (3)

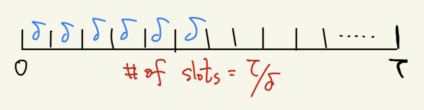


- Our interest: probability of k arrivals over $[0, \tau]$
- Given "small" δ, # of arrivals $\sim$ Bin(n, p), where $n = \tau/\delta$ and $p = \lambda\delta$
- As $\delta \to 0$, $np = \tau/\delta \times \lambda\delta = \lambda\tau$.

Key Design Idea to Develop a Continuous Twin (3)



- Our interest: probability of k arrivals over $[0, \tau]$
- Given "small" δ, # of arrivals $\sim$ Bin(n, p), where $n = \tau/\delta$ and $p = \lambda\delta$
- As $\delta \to 0$, $np = \tau/\delta \times \lambda\delta = \lambda\tau$.
- # of arrivals over $[0, \tau]$, $\sim$ Poisson$(\lambda\tau)$



- Our interest: probability of k arrivals over $[0, \tau]$
- Given "small" δ, # of arrivals $\sim$ Bin(n, p), where $n = \tau/\delta$ and $p = \lambda\delta$
- As $\delta \to 0$, $np = \tau/\delta \times \lambda\delta = \lambda\tau$.
- # of arrivals over $[0, \tau]$, $\sim$ Poisson$(\lambda\tau)$
- This is a continuous twin process of Bernous process, which we call Poisson process.

Roadmap

(1) Introduction of Random Processes

(2) Bernoulli Processes

(3) Poisson Processes: Poisson RV and Basic Idea

(4) Poisson Processes: Definition and Properties

(5) Playing with Bernoulli and Poisson Processes

(6) Random Incidence Phenomenon

Poisson Process: Definition (1)

An arrival process is called a Poisson process with rate λ, if the following are satisfied:

An arrival process is called a Poisson process with rate λ, if the following are satisfied:

- (Independence) Let N_τ be the number of arrivals over the interval $[0, \tau]$. For any $\tau_1, \tau_2 > 0$, $N_{s+\tau_1} - N_s$ is independent of $N_{t+\tau_2} - N_t$, if $t > s + \tau_1$.
  - The number of arrivals over two disjoint intervals are independent.

An arrival process is called a Poisson process with rate λ, if the following are satisfied:

- (Independence) Let N_τ be the number of arrivals over the interval $[0, \tau]$. For any $\tau_1, \tau_2 > 0$, $N_{s+\tau_1} - N_s$ is independent of $N_{t+\tau_2} - N_t$, if $t > s + \tau_1$.
  - The number of arrivals over two disjoint intervals are independent.
- (Time homogeneity) For any s, the distribution of $N_{s+\tau} - N_s$ is equal to that of N_τ.
  - N_τ becomes the number of arrivals over any interval of length τ.

An arrival process is called a Poisson process with rate λ, if the following are satisfied:

- (Independence) Let N_τ be the number of arrivals over the interval $[0, \tau]$. For any $\tau_1, \tau_2 > 0$, $N_{s+\tau_1} - N_s$ is independent of $N_{t+\tau_2} - N_t$, if $t > s + \tau_1$.
  - The number of arrivals over two disjoint intervals are independent.
- (Time homogeneity) For any s, the distribution of $N_{s+\tau} - N_s$ is equal to that of N_τ.
  - N_τ becomes the number of arrivals over any interval of length τ.
- (Small interval probability) Let $\mathbb{P}(k, \tau) = \mathbb{P}(N_\tau = k)$, which satisfy:

$$\mathbb{P}(0, \tau) = 1 - \lambda\tau + o(\tau)$$

$$\mathbb{P}(1, \tau) = \lambda\tau + o_1(\tau)$$

$$\mathbb{P}(k, \tau) = o_k(\tau) \quad \text{for } k = 2, 3, \ldots, \quad \text{where} \quad \lim_{\tau \to 0} \frac{o(\tau)}{\tau} = 0, \quad \lim_{\tau \to 0} \frac{o_k(\tau)}{\tau} = 0$$

Poisson Process: Definition (2)

An arrival process is called a Poisson process with rate λ, if the following are satisfied:

- (Independence) Let N_τ be the number of arrivals over the interval $[0, \tau]$. For any $\tau_1, \tau_2 > 0$, $N_{s+\tau_1} - N_s$ is independent of $N_{t+\tau_2} - N_t$, if $t > s + \tau_1$.
  - The number of arrivals over two disjoint intervals are independent.
- (Time homogeneity) For any s, the distribution of $N_{s+\tau} - N_s$ is equal to that of N_τ.
  - N_τ becomes the number of arrivals over any interval of length τ.
- (Distribution of N_τ) N_τ is the Poisson rv with parameter $\lambda\tau$, i.e., if we let $\mathbb{P}(k, \tau) = \mathbb{P}(N_\tau = k)$, we have:

$$\mathbb{P}(k, \tau) = e^{\lambda\tau} \frac{(\lambda\tau)^k}{k!}, \quad k = 0, 1, 2, \dots$$

Poisson Process: $\mathbb{P}(k, \tau)$, N_τ, and T

(Q1) Number of arrivals of any interval with length $\tau \sim \text{Poisson}(\lambda\tau)$, i.e.,

$$\mathbb{P}(k,\tau) = e^{-\lambda\tau}\frac{(\lambda\tau)^k}{k!}, \quad k = 0, 1, 2, \ldots$$

Poisson Process: $\mathbb{P}(k, \tau)$, N_τ, and T

(Q1) Number of arrivals of any interval with length $\tau \sim$ Poisson($\lambda\tau$), i.e.,

$$\mathbb{P}(k, \tau) = e^{-\lambda\tau}\frac{(\lambda\tau)^k}{k!}, \quad k = 0, 1, 2, \ldots$$

$$\mathbb{E}(N_\tau) = \lambda\tau \text{ and } \mathrm{var}(N_\tau) = \lambda\tau$$

Poisson Process: $\mathbb{P}(k,\tau)$, N_τ, and T

(Q1) Number of arrivals of any interval with length $\tau \sim \text{Poisson}(\lambda\tau)$, i.e.,

$$\mathbb{P}(k,\tau) = e^{-\lambda\tau}\frac{(\lambda\tau)^k}{k!}, \quad k = 0,1,2,\ldots$$

$$\mathbb{E}(N_\tau) = \lambda\tau \text{ and } \mathrm{var}(N_\tau) = \lambda\tau$$

(Q2) Time of first arrival T

Poisson Process: $\mathbb{P}(k,\tau)$, N_τ, and T

(Q1) Number of arrivals of any interval with length $\tau \sim \text{Poisson}(\lambda\tau)$, i.e.,

$$\mathbb{P}(k,\tau) = e^{-\lambda\tau}\frac{(\lambda\tau)^k}{k!}, \quad k = 0, 1, 2, \ldots$$

$$\mathbb{E}(N_\tau) = \lambda\tau \text{ and } \text{var}(N_\tau) = \lambda\tau$$

(Q2) Time of first arrival T

$$F_T(t) = \mathbb{P}(T \le t) = 1 - \mathbb{P}(T > t) = 1 - P(0,t) = 1 - e^{-\lambda t}$$

Poisson Process: $\mathbb{P}(k, \tau)$, N_τ, and T

KAIST EE

(Q1) Number of arrivals of any interval with length $\tau \sim \text{Poisson}(\lambda\tau)$, i.e.,

$$\mathbb{P}(k, \tau) = e^{-\lambda\tau}\frac{(\lambda\tau)^k}{k!}, \quad k = 0, 1, 2, \ldots$$

$$\mathbb{E}(N_\tau) = \lambda\tau \text{ and } \text{var}(N_\tau) = \lambda\tau$$

(Q2) Time of first arrival T

$$F_T(t) = \mathbb{P}(T \le t) = 1 - \mathbb{P}(T > t) = 1 - P(0, t) = 1 - e^{-\lambda t}$$

$$f_T(t) = \frac{dF_T(t)}{dt} = \lambda e^{-\lambda t}, \quad t \ge 0.$$

(Q1) Number of arrivals of any interval with length $\tau \sim \text{Poisson}(\lambda\tau)$, i.e.,

$$\mathbb{P}(k,\tau) = e^{-\lambda\tau}\frac{(\lambda\tau)^k}{k!}, \quad k = 0, 1, 2, \ldots$$

$$\mathbb{E}(N_\tau) = \lambda\tau \text{ and } \text{var}(N_\tau) = \lambda\tau$$

(Q2) Time of first arrival T

$$F_T(t) = \mathbb{P}(T \le t) = 1 - \mathbb{P}(T > t) = 1 - P(0,t) = 1 - e^{-\lambda t}$$

$$f_T(t) = \frac{dF_T(t)}{dt} = \lambda e^{-\lambda t}, \quad t \ge 0.$$

- $T \sim \text{Exp}(\lambda)$. Thus, $\mathbb{E}(T) = 1/\lambda$ and $\text{var}(T) = 1/\lambda^2$
  - Continuous twin of geometric rv in Bernoulli process
  - Memoryless

- Receive emails according to a Poisson process at rate $\lambda = 5$ messages/hour
- Mean and variance of mails received during a day
- $\mathbb{P}$[one new message in the next hour]
- $\mathbb{P}$[exactly two msgs during each of the next three hours]

Poisson Process: Example

- Receive emails according to a Poisson process at rate $\lambda = 5$ messages/hour
- Mean and variance of mails received during a day
  \- $5 * 24 = 120$
- $\mathbb{P}$[one new message in the next hour]
- $\mathbb{P}$[exactly two msgs during each of the next three hours]

KAIST EE

- Receive emails according to a Poisson process at rate $\lambda = 5$ messages/hour
- Mean and variance of mails received during a day
  - 5* $24 = 120$
- $\mathbb{P}$[one new message in the next hour]
  - $\mathbb{P}(1,1) = \dfrac{(5 \cdot 1)^1 e^{-5 \cdot 1}}{1!} = 5e^{-5}$
- $\mathbb{P}$[exactly two msgs during each of the next three hours]

Poisson Process: Example

- Receive emails according to a Poisson process at rate $\lambda = 5$ messages/hour
- Mean and variance of mails received during a day
  - $5 * 24 = 120$
- $\mathbb{P}$[one new message in the next hour]
  - $\mathbb{P}(1,1) = \dfrac{(5 \cdot 1)^1 e^{-5 \cdot 1}}{1!} = 5e^{-5}$
- $\mathbb{P}$[exactly two msgs during each of the next three hours]
  - $\left(\dfrac{5^2 e^{-5}}{2!}\right)^3$

- Remind. Similar property for Bernoulli processes, but here no time slots.

Memoryless and Fresh-start Property

- Remind. Similar property for Bernoulli processes, but here no time slots.
- Fresh-start at determinsitic time t: Start watching at time t, then you see the Poisson process, independent of what has happened in the past.

Memoryless and Fresh-start Property

- Remind. Similar property for Bernoulli processes, but here no time slots.
- Fresh-start at determinsitic time t: Start watching at time t, then you see the Poisson process, independent of what has happened in the past.
- Fresh-start at random time T: Similarly, it holds.
  - For example, when you start watching at random time T_1 (time of first arrival).
  - Generally, it holds when T is a stopping time.

- Remind. Similar property for Bernoulli processes, but here no time slots.
- Fresh-start at determinsitic time t: Start watching at time t, then you see the Poisson process, independent of what has happened in the past.
- Fresh-start at random time T: Similarly, it holds.
  - For example, when you start watching at random time T_1 (time of first arrival).
  - Generally, it holds when T is a stopping time.

(Q3) The k-th arrival time Y_k?

- Remind. Similar property for Bernoulli processes, but here no time slots.
- Fresh-start at determinsitic time t: Start watching at time t, then you see the Poisson process, independent of what has happened in the past.
- Fresh-start at random time T: Similarly, it holds.
  - For example, when you start watching at random time T_1 (time of first arrival).
  - Generally, it holds when T is a stopping time.

(Q3) The k-th arrival time Y_k?

- k-th inter-arrival time $T_k = Y_k - Y_{k-1}$, $k \geq 2$, and $T_1 = Y_1$.

KAIST EE

- Remind. Similar property for Bernoulli processes, but here no time slots.
- Fresh-start at determinsitic time t: Start watching at time t, then you see the Poisson process, independent of what has happened in the past.
- Fresh-start at random time T: Similarly, it holds.
  - For example, when you start watching at random time T_1 (time of first arrival).
  - Generally, it holds when T is a stopping time.

(Q3) The k-th arrival time Y_k?

- k-th inter-arrival time $T_k = Y_k - Y_{k-1}$, $k \geq 2$, and $T_1 = Y_1$.
- $Y_k = T_1 + T_2 + \cdots + T_k$ is sum of i.i.d. exponential rvs.

- Remind. Similar property for Bernoulli processes, but here no time slots.
- Fresh-start at determinsitic time t: Start watching at time t, then you see the Poisson process, independent of what has happened in the past.
- Fresh-start at random time T: Similarly, it holds.
  - For example, when you start watching at random time T_1 (time of first arrival).
  - Generally, it holds when T is a stopping time.

(Q3) The k-th arrival time Y_k?

- k-th inter-arrival time $T_k = Y_k - Y_{k-1}$, $k \geq 2$, and $T_1 = Y_1$.
- $Y_k = T_1 + T_2 + \cdots + T_k$ is sum of i.i.d. exponential rvs.
- $\mathbb{E}[Y_k] = k/\lambda$ and $\text{var}[Y_k] = k/\lambda^2$, but what is the distribution of Y_k?

PDF of Y_k

PDF of Y_k

- For a given δ, $\boxed{}$: prob. of k-th arrival over $[y, y+\delta]$.

PDF of Y_k

- For a given δ, $\boxed{\delta \cdot f_{Y_k}(y)}$: prob. of k-th arrival over $[y, y+\delta]$.

PDF of Y_k

- For a given δ, $\boxed{\delta \cdot f_{Y_k}(y)}$: prob. of k-th arrival over $[y, y+\delta]$.
- When δ is small, only one arrival occurs. Thus,

$\delta \cdot f_{Y_k}(y) =$

PDF of Y_k

- For a given δ, $\boxed{\delta \cdot f_{Y_k}(y)}$: prob. of k-th arrival over $[y, y+\delta]$.
- When δ is small, only one arrival occurs. Thus,

  $$\delta \cdot f_{Y_k}(y) = \mathbb{P}(\text{an arrival over } [y, y+\delta]) \times \mathbb{P}(k-1 \text{ arrivals before } y)$$

PDF of Y_k

- For a given δ, $\boxed{\delta \cdot f_{Y_k}(y)}$: prob. of k-th arrival over $[y, y+\delta]$.
- When δ is small, only one arrival occurs. Thus,

$$\begin{aligned}\delta \cdot f_{Y_k}(y) &= \mathbb{P}\left(\text{an arrival over } [y, y+\delta]\right) \times \mathbb{P}\left(k-1 \text{ arrivals before } y\right) \\ &\approx \lambda\delta \times \mathbb{P}(k-1, y) = \lambda\delta \times \frac{\lambda^{k-1} y^{k-1} e^{-\lambda y}}{(k-1)!}\end{aligned}$$

PDF of Y_k

- For a given δ, $\boxed{\delta \cdot f_{Y_k}(y)}$: prob. of k-th arrival over $[y, y+\delta]$.
- When δ is small, only one arrival occurs. Thus,

$$\delta \cdot f_{Y_k}(y) = \mathbb{P}\left(\text{an arrival over } [y, y+\delta]\right) \times \mathbb{P}\left(k-1 \text{ arrivals before } y\right)$$

$$\approx \lambda\delta \times \mathbb{P}(k-1, y) = \lambda\delta \times \frac{\lambda^{k-1} y^{k-1} e^{-\lambda y}}{(k-1)!}$$

$$f_{Y_k}(y) = \frac{\lambda^k y^{k-1} e^{-\lambda y}}{(k-1)!}, \quad y \geq 0.$$

KAIST EE

- For a given δ, $\boxed{\delta \cdot f_{Y_k}(y)}$: prob. of k-th arrival over $[y, y+\delta]$.
- When δ is small, only one arrival occurs. Thus,

$$\delta \cdot f_{Y_k}(y) = \mathbb{P}\left(\text{an arrival over } [y, y+\delta]\right) \times \mathbb{P}\left(k-1 \text{ arrivals before } y\right)$$

$$\approx \lambda\delta \times \mathbb{P}(k-1, y) = \lambda\delta \times \frac{\lambda^{k-1} y^{k-1} e^{-\lambda y}}{(k-1)!}$$

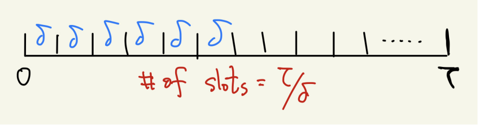
$$f_{Y_k}(y) = \frac{\lambda^k y^{k-1} e^{-\lambda y}}{(k-1)!}, \quad y \geq 0.$$

- This is called Erlang rv.

An Erlang random variable Z with parameter (k, λ) has the following pdf:

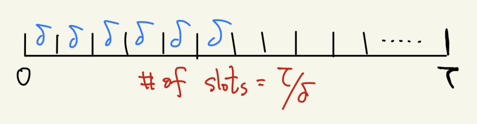
$$f_Z(z) = \frac{\lambda^k z^{k-1} e^{-\lambda z}}{(k-1)!}, \quad z \geq 0$$

- $n = \tau/\delta$, $p = \lambda\delta$, $np = \lambda\tau$



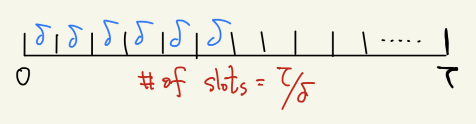
	Bernoulli process	Poisson process
time of arrival	Discrete	Continuous
PMF of # of arrivals		
Interarrival time		
Time of k-th arrival		
Arrival rate		

- $n = \tau/\delta$, $p = \lambda\delta$, $np = \lambda\tau$



	Bernoulli process	Poisson process
time of arrival	Discrete	Continuous
PMF of # of arrivals	Binomial	Poisson
Interarrival time		
Time of k-th arrival		
Arrival rate		

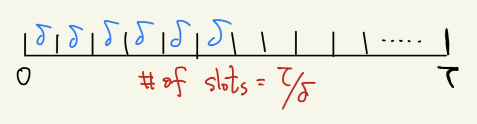
- $n = \tau/\delta$, $p = \lambda\delta$, $np = \lambda\tau$



	Bernoulli process	Poisson process
time of arrival	Discrete	Continuous
PMF of # of arrivals	Binomial	Poisson
Interarrival time	Geometric	Exponential
Time of k-th arrival		
Arrival rate		

Poisson Process vs. Bernoulli Process

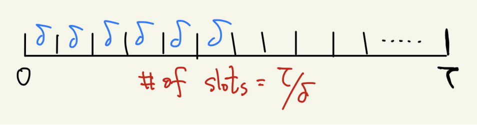
- $n = \tau/\delta$, $p = \lambda\delta$, $np = \lambda\tau$



	Bernoulli process	Poisson process
time of arrival	Discrete	Continuous
PMF of # of arrivals	Binomial	Poisson
Interarrival time	Geometric	Exponential
Time of k-th arrival	Pascal	Erlang
Arrival rate		

Poisson Process vs. Bernoulli Process

- $n = \tau/\delta$, $p = \lambda\delta$, $np = \lambda\tau$



	Bernoulli process	Poisson process
time of arrival	Discrete	Continuous
PMF of # of arrivals	Binomial	Poisson
Interarrival time	Geometric	Exponential
Time of k-th arrival	Pascal	Erlang
Arrival rate	p/per slot	λ/unit time

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0, 2)$

Example: Poisson Fishing (Problem 10, page 329)

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0, 2)$
Method 2: $\mathbb{P}(T_1 > 2)$

Example: Poisson Fishing (Problem 10, page 329)

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0, 2)$
Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$

Example: Poisson Fishing (Problem 10, page 329)

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0, 2)$
Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$

Method 1: $\mathbb{P}(0, 2)(1 - \mathbb{P}(0, 3))$

- Catching fish: Poisson process $\lambda = 0.6$/hour.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0, 2)$
Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$

Method 1: $\mathbb{P}(0, 2)(1 - \mathbb{P}(0, 3))$
Method 2: $\mathbb{P}(2 < T_1 < 5)$

Example: Poisson Fishing (Problem 10, page 329)

KAIST EE

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0, 2)$
Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$

Method 1: $\mathbb{P}(0, 2)(1 - \mathbb{P}(0, 3))$
Method 2:$\mathbb{P}(2 < T_1 < 5)$

(Q3) $\mathbb{P}(\text{Catch at least two fish})$

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0, 2)$
Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$

Method 1: $\mathbb{P}(0, 2)(1 - \mathbb{P}(0, 3))$
Method 2:$\mathbb{P}(2 < T_1 < 5)$

(Q3) $\mathbb{P}(\text{Catch at least two fish})$

Method 1:$\sum_{k=2}^{\infty} = 1 - \mathbb{P}(0, 2) - \mathbb{P}(1, 2)$

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0, 2)$
Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$

Method 1: $\mathbb{P}(0, 2)(1 - \mathbb{P}(0, 3))$
Method 2:$\mathbb{P}(2 < T_1 < 5)$

(Q3) $\mathbb{P}(\text{Catch at least two fish})$

Method 1:$\sum_{k=2}^{\infty} = 1 - \mathbb{P}(0, 2) - \mathbb{P}(1, 2)$
Method 2: $\mathbb{P}(Y_2 \leq 2)$

Example: Poisson Fishing (Problem 10, page 329)

KAIST EE

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0,2)$
Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$

Method 1: $\mathbb{P}(0,2)(1 - \mathbb{P}(0,3))$
Method 2:$\mathbb{P}(2 < T_1 < 5)$

(Q3) $\mathbb{P}(\text{Catch at least two fish})$

Method 1:$\sum_{k=2}^{\infty} = 1 - \mathbb{P}(0,2) - \mathbb{P}(1,2)$
Method 2: $\mathbb{P}(Y_2 \leq 2)$

(Q4) $\mathbb{E}[\text{future fi. time}|\text{already fished for 3h}]$

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$
Method 1: $\mathbb{P}(0,2)$
Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$
Method 1: $\mathbb{P}(0,2)(1 - \mathbb{P}(0,3))$
Method 2:$\mathbb{P}(2 < T_1 < 5)$

(Q3) $\mathbb{P}(\text{Catch at least two fish})$
Method 1:$\sum_{k=2}^{\infty} = 1 - \mathbb{P}(0,2) - \mathbb{P}(1,2)$
Method 2: $\mathbb{P}(Y_2 \leq 2)$

(Q4) $\mathbb{E}[\text{future fi. time}|\text{already fished for 3h}]$
Fresh-start. So,
$\mathbb{E}[exp(\lambda)] = 1/\lambda = 1/0.6$

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$
Method 1: $\mathbb{P}(0,2)$
Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$
Method 1: $\mathbb{P}(0,2)(1 - \mathbb{P}(0,3))$
Method 2:$\mathbb{P}(2 < T_1 < 5)$

(Q3) $\mathbb{P}(\text{Catch at least two fish})$
Method 1:$\sum_{k=2}^{\infty} = 1 - \mathbb{P}(0,2) - \mathbb{P}(1,2)$
Method 2: $\mathbb{P}(Y_2 \leq 2)$

(Q4) $\mathbb{E}[\text{future fi. time}|\text{already fished for 3h}]$
Fresh-start. So,
$\mathbb{E}[exp(\lambda)] = 1/\lambda = 1/0.6$

(Q5) $\mathbb{E}[\text{F=total fishing time}]$

Example: Poisson Fishing (Problem 10, page 329)

KAIST EE

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0, 2)$
Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$

Method 1: $\mathbb{P}(0, 2)(1 - \mathbb{P}(0, 3))$
Method 2:$\mathbb{P}(2 < T_1 < 5)$

(Q3) $\mathbb{P}(\text{Catch at least two fish})$

Method 1:$\sum_{k=2}^{\infty} = 1 - \mathbb{P}(0, 2) - \mathbb{P}(1, 2)$
Method 2: $\mathbb{P}(Y_2 \le 2)$

(Q4) $\mathbb{E}[\text{future fi. time}|\text{already fished for 3h}]$
Fresh-start. So,
$\mathbb{E}[exp(\lambda)] = 1/\lambda = 1/0.6$

(Q5) $\mathbb{E}[\text{F=total fishing time}]$

$$
\begin{aligned}
2 + \mathbb{E}[F - 2] &= 2 + \mathbb{P}(F = 2) \cdot 0 + \\
&\quad \mathbb{P}(F > 2) \cdot \mathbb{E}[F - 2 | F > 2] \\
&= 2 + \mathbb{P}(0, 2) \cdot \frac{1}{\lambda}
\end{aligned}
$$

Example: Poisson Fishing (Problem 10, page 329)

KAIST EE

- Catching fish: Poisson process $\lambda = 0.6/\text{hour}$.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0,2)$
Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$

Method 1: $\mathbb{P}(0,2)(1-\mathbb{P}(0,3))$
Method 2:$\mathbb{P}(2 < T_1 < 5)$

(Q3) $\mathbb{P}(\text{Catch at least two fish})$

Method 1:$\sum_{k=2}^{\infty} = 1 - \mathbb{P}(0,2) - \mathbb{P}(1,2)$
Method 2: $\mathbb{P}(Y_2 \le 2)$

(Q4) $\mathbb{E}[\text{future fi. time}|\text{already fished for 3h}]$
Fresh-start. So,
$\mathbb{E}[exp(\lambda)] = 1/\lambda = 1/0.6$

(Q5) $\mathbb{E}[\text{F=total fishing time}]$

$$
\begin{aligned}
2 + \mathbb{E}[F-2] = 2 + \mathbb{P}(F=2)\cdot 0 + \\
\mathbb{P}(F>2)\cdot \mathbb{E}[F-2|F>2] \\
= 2 + \mathbb{P}(0,2)\cdot \frac{1}{\lambda}
\end{aligned}
$$

(Q6) $\mathbb{E}[\text{number of fish}] = \lambda \cdot 2 + \mathbb{P}(0,2)\cdot 1$

Roadmap

(1) Introduction of Random Processes

(2) Bernoulli Processes

(3) Poisson Processes: Poisson RV and Basic Idea

(4) Poisson Processes: Definition and Properties

(5) Playing with Bernoulli and Poisson Processes

(6) Random Incidence Phenomenon

KAIST EE

Alternative Description of the Bernoulli Process

1. Start with a sequence of independent geometric random variables T_1, $T_2, \ldots$, with common parameter p, and let these stand for the interarrival times.
2. Record a success (or arrival) at times T_1, $T_1 + T_2$, $T_1 + T_2 + T_3$, etc.

Alternative Description of the Poisson Process

1. Start with a sequence of independent exponential random variables $T_1, T_2, \ldots$, with common parameter λ, and let these stand for the interarrival times.
2. Record an arrival at times T_1, $T_1 + T_2$, $T_1 + T_2 + T_3$, etc.

- It has been observed that after a rainy day, the number of days until the next rain $\sim \text{Geom}(p)$, independent of the past.

Example

- It has been observed that after a rainy day, the number of days until the next rain $\sim \text{Geom}(p)$, independent of the past.
- Question. What is the probability that it rains on both the 5th and the 8th day of the month?

Example

- It has been observed that after a rainy day, the number of days until the next rain $\sim \text{Geom}(p)$, independent of the past.
- Question. What is the probability that it rains on both the 5th and the 8th day of the month?
- Approach 1: Handling this problem directly with the geometric PMFs is very tedious and complex.

- It has been observed that after a rainy day, the number of days until the next rain $\sim$ Geom(p), independent of the past.
- Question. What is the probability that it rains on both the 5th and the 8th day of the month?
- Approach 1: Handling this problem directly with the geometric PMFs is very tedious and complex.
- Approach 2: Rainy days is a Bernoulli process with arrival. probability p.

Example

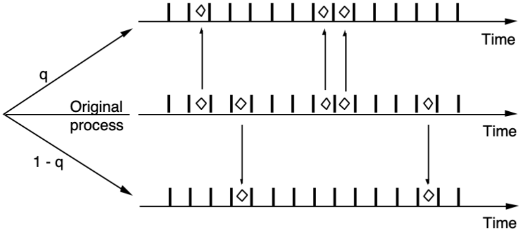
- It has been observed that after a rainy day, the number of days until the next rain $\sim$ Geom(p), independent of the past.
- Question. What is the probability that it rains on both the 5th and the 8th day of the month?
- Approach 1: Handling this problem directly with the geometric PMFs is very tedious and complex.
- Approach 2: Rainy days is a Bernoulli process with arrival. probability p.
- Thus, the answer is p^2.

Coding of Random Arrivals

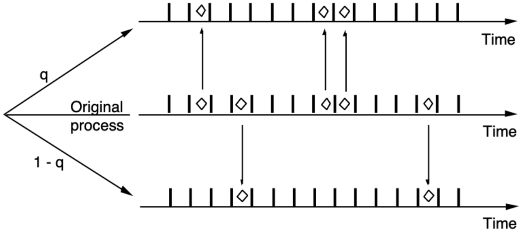
- Question. How to make software codes of Bernoulli process with p and Poisson process with λ
- Inter-arrival times are very handy.
- Bernoulli process with p: Obtain a sequence of random values following the geometric distribution with parameter p.
- Poisson process with λ: Obtain a sequence of random values following the exponential distribution with parameter λ.

- Bernoulli random variable: $\text{Bern}(p)$
- Bernoulli process: $\text{BP}(p)$
- Poisson random variable: $\text{Poisson}(\lambda)$
- Poisson process: $\text{PP}(\lambda)$

- Split BP(p) into two processes. Whenever there is an arrival, keep it w.p. q and discard it w.p. $(1-q)$.

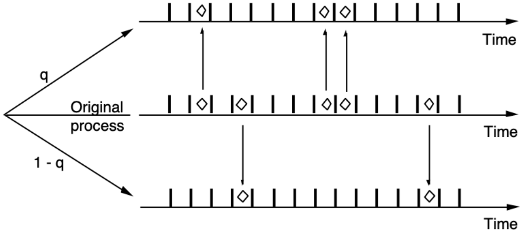


- Split BP(p) into two processes. Whenever there is an arrival, keep it w.p. q and discard it w.p. $(1-q)$.
- Split decisions are independent of arrivals.



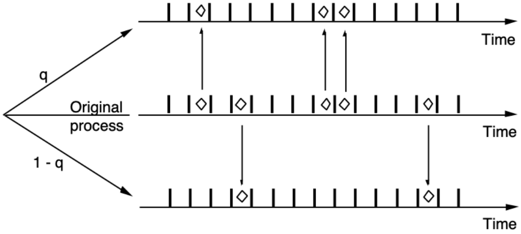
Split: Bernoulli Process

- Split BP(p) into two processes. Whenever there is an arrival, keep it w.p. q and discard it w.p. $(1 - q)$.
- Split decisions are independent of arrivals.
- Question. What are the two split processes?

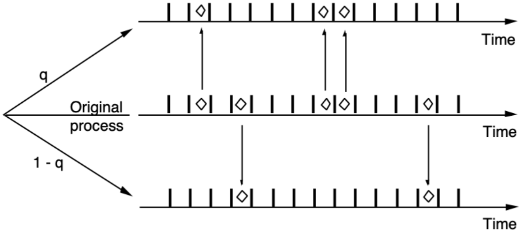


Split: Bernoulli Process

- Split BP(p) into two processes. Whenever there is an arrival, keep it w.p. q and discard it w.p. $(1-q)$.
- Split decisions are independent of arrivals.
- Question. What are the two split processes?
- BP(pq) and BP($p(1-q)$). Why?

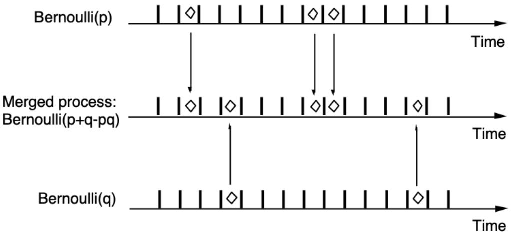


- Split BP(p) into two processes. Whenever there is an arrival, keep it w.p. q and discard it w.p. $(1-q)$.
- Split decisions are independent of arrivals.
- Question. What are the two split processes?
- BP(pq) and BP($p(1-q)$). Why?
- Are they independent? No.

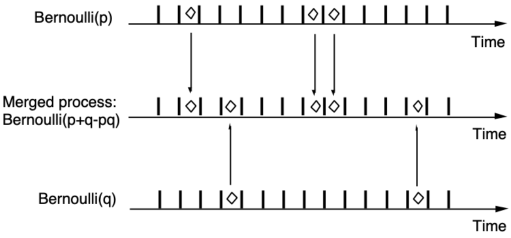


Merge: Bernoulli Process

- Merge BP(p) and BP(q) into one process.

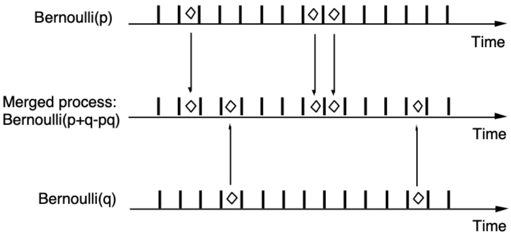


- Merge BP(p) and BP(q) into one process.
- Collided arrival is regarded just one arrival in the merged process



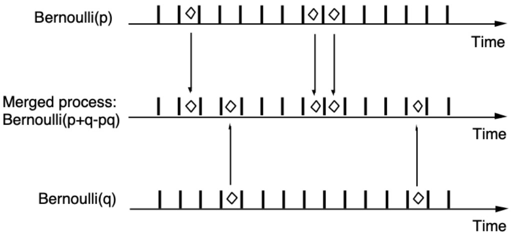
Merge: Bernoulli Process

- Merge BP(p) and BP(q) into one process.
- Collided arrival is regarded just one arrival in the merged process
- Probability of having at least one arrival: $1 - (1-p)(1-q) = p + q - pq$



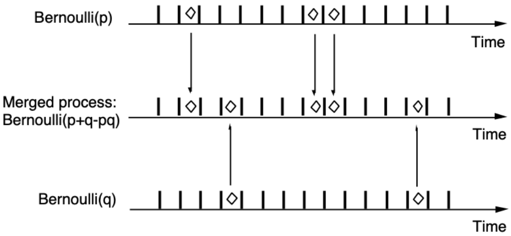
Merge: Bernoulli Process

- Merge BP(p) and BP(q) into one process.
- Collided arrival is regarded just one arrival in the merged process
- Probability of having at least one arrival: $1 - (1-p)(1-q) = p + q - pq$
- Merged process: BP($p + q - pq$)



Merge: Bernoulli Process

- Merge BP(p) and BP(q) into one process.
- Collided arrival is regarded just one arrival in the merged process
- Probability of having at least one arrival: $1-(1-p)(1-q)=p+q-pq$
- Merged process: $\text{BP}(p+q-pq)$
- $\mathbb{P}(\text{arrival from proc. } 1 \mid \text{arrival in the merged proc.}) = \frac{p}{p+q-pq}$



Split: Poisson Process

- Split a Poisson process $\mathrm{PP}(\lambda)$ into two processes by keeping each arrival w.p. p and discarding it w.p. $(1-p)$

Split: Poisson Process

- Split a Poisson process PP(λ) into two processes by keeping each arrival w.p. p and discarding it w.p. $(1-p)$
- Question. What are the split processes?

- Split a Poisson process PP(λ) into two processes by keeping each arrival w.p. p and discarding it w.p. $(1 - p)$
- Question. What are the split processes?
- Let's focus on the process that we keep

Split: Poisson Process

- Split a Poisson process PP(λ) into two processes by keeping each arrival w.p. p and discarding it w.p. $(1 - p)$
- Question. What are the split processes?
- Let's focus on the process that we keep
- Independence and time-homogeneity? Yes

Split: Poisson Process

- Split a Poisson process PP(λ) into two processes by keeping each arrival w.p. p and discarding it w.p. $(1-p)$
- Question. What are the split processes?
- Let's focus on the process that we keep
- Independence and time-homogeneity? Yes
- Small interval probability over δ-interval

Split: Poisson Process

- Split a Poisson process PP(λ) into two processes by keeping each arrival w.p. p and discarding it w.p. $(1-p)$
- Question. What are the split processes?
- Let's focus on the process that we keep
- Independence and time-homogeneity? Yes
- Small interval probability over δ-interval
  - $\mathbb{P}(1 \text{ arrival}) = p\lambda\delta + p \cdot o(\delta) = p\lambda\delta + o(\delta)$
  - $\mathbb{P}(2 \text{ arrivals}) = p \cdot o(\delta) = o(\delta)$
  - $\mathbb{P}(0 \text{ arrival}) = 1 - p\lambda\delta - p \cdot o(\delta) - p \cdot o(\delta) = 1 - p\lambda\delta + o(\delta)$
- PP(λp) and PP($\lambda(1-p)$)

Merge: Poisson Process (1)

- Merge from $\mathrm{PP}(\lambda_1)$ and $\mathrm{PP}(\lambda_2)$

Merge: Poisson Process (1)

- Merge from PP(λ_1) and PP(λ_2)
- Independence and time-homogeneity? Yes

Merge: Poisson Process (1)

- Merge from $\mathrm{PP}(\lambda_1)$ and $\mathrm{PP}(\lambda_2)$
- Independence and time-homogeneity? Yes
- Small interval probabilty over δ-interval (ignoring $o(\delta)$ for small δ)

  $\mathbb{P}(0 \text{ arrival}) \approx (1-\lambda_1\delta)(1-\lambda_2\delta) \approx 1-(\lambda_1+\lambda_2)\delta$

  $\mathbb{P}(1 \text{ arrival}) \approx (\lambda_1\delta)(1-\lambda_2\delta)+\lambda_2\delta(1-\lambda_1\delta) \approx (\lambda_1+\lambda_2)\delta$

Merge: Poisson Process (1)

- Merge from $\text{PP}(\lambda_1)$ and $\text{PP}(\lambda_2)$
- Independence and time-homogeneity? Yes
- Small interval probabilty over δ-interval (ignoring $o(\delta)$ for small δ)

  $\mathbb{P}(0 \text{ arrival}) \approx (1-\lambda_1\delta)(1-\lambda_2\delta) \approx 1-(\lambda_1+\lambda_2)\delta$

  $\mathbb{P}(1 \text{ arrival}) \approx (\lambda_1\delta)(1-\lambda_2\delta)+\lambda_2\delta(1-\lambda_1\delta) \approx (\lambda_1+\lambda_2)\delta$
- Merged process: $\text{PP}(\lambda_1+\lambda_2)$

Merge: Poisson Process (2)

- Red: $PP(\lambda_1)$ and Blue: $PP(\lambda_2)$

Merge: Poisson Process (2)

- Red: PP(λ_1) and Blue: PP(λ_2)
- $\mathbb{P}$(from red | arrival at time t in the merged proc.)?

Merge: Poisson Process (2)

- Red: PP(λ_1) and Blue: PP(λ_2)
- $\mathbb{P}$(from red | arrival at time t in the merged proc.)? $\dfrac{\lambda_1\delta}{\lambda_1\delta + \lambda_2\delta} = \dfrac{\lambda_1}{\lambda_1 + \lambda_2}$

Merge: Poisson Process (2)

- Red: PP(λ_1) and Blue: PP(λ_2)
- $\mathbb{P}$(from red | arrival at time t in the merged proc.)? $\dfrac{\lambda_1\delta}{\lambda_1\delta + \lambda_2\delta} = \dfrac{\lambda_1}{\lambda_1 + \lambda_2}$
- Consider an event $A_k = \{k\text{-th arrival in the merged proc. is red}\}$

- Red: PP(λ_1) and Blue: PP(λ_2)
- $\mathbb{P}$(from red | arrival at time t in the merged proc.)? $\dfrac{\lambda_1\delta}{\lambda_1\delta + \lambda_2\delta} = \dfrac{\lambda_1}{\lambda_1 + \lambda_2}$
- Consider an event $A_k = \{k\text{-th arrival in the merged proc. is red}\}$
  - $\mathbb{P}(A_k)$?

Merge: Poisson Process (2)

- Red: PP(λ_1) and Blue: PP(λ_2)
- $\mathbb{P}$(from red | arrival at time t in the merged proc.)? $\dfrac{\lambda_1\delta}{\lambda_1\delta + \lambda_2\delta} = \dfrac{\lambda_1}{\lambda_1 + \lambda_2}$
- Consider an event $A_k = \{k\text{-th arrival in the merged proc. is red}\}$
  - $\mathbb{P}(A_k)$? $\dfrac{\lambda_1}{\lambda_1 + \lambda_2}$

Merge: Poisson Process (2)

- Red: PP(λ_1) and Blue: PP(λ_2)
- $\mathbb{P}$(from red | arrival at time t in the merged proc.)? $\dfrac{\lambda_1\delta}{\lambda_1\delta + \lambda_2\delta} = \dfrac{\lambda_1}{\lambda_1 + \lambda_2}$
- Consider an event $A_k = \{k\text{-th arrival in the merged proc. is red}\}$
  - $\mathbb{P}(A_k)$? $\dfrac{\lambda_1}{\lambda_1 + \lambda_2}$
  - $A_1, A_2, \ldots$ are independent: origins (red or blue) of arrivals in the merged proc. are indepdnent

- Red: $\mathrm{PP}(\lambda_1)$ and Blue: $\mathrm{PP}(\lambda_2)$
- $\mathbb{P}$(from red | arrival at time t in the merged proc.)? $\dfrac{\lambda_1\delta}{\lambda_1\delta + \lambda_2\delta} = \dfrac{\lambda_1}{\lambda_1 + \lambda_2}$
- Consider an event $A_k = \{k\text{-th arrival in the merged proc. is red}\}$
  - $\mathbb{P}(A_k)$? $\dfrac{\lambda_1}{\lambda_1 + \lambda_2}$
  - $A_1, A_2, \ldots$ are independent: origins (red or blue) of arrivals in the merged proc. are indepdnent
- $\mathbb{P}$(k out of first 10 arrivals are red)?

KAIST EE

- Red: PP(λ_1) and Blue: PP(λ_2)
- $\mathbb{P}$(from red | arrival at time t in the merged proc.)? $\dfrac{\lambda_1\delta}{\lambda_1\delta + \lambda_2\delta} = \dfrac{\lambda_1}{\lambda_1 + \lambda_2}$
- Consider an event $A_k = \{k\text{-th arrival in the merged proc. is red}\}$
  - $\mathbb{P}(A_k)$? $\dfrac{\lambda_1}{\lambda_1 + \lambda_2}$
  - $A_1, A_2, \ldots$ are independent: origins (red or blue) of arrivals in the merged proc. are indepdnent
- $\mathbb{P}$(k out of first 10 arrivals are red)? $\binom{10}{k}\left(\dfrac{\lambda_1}{\lambda_1+\lambda_2}\right)^k\left(\dfrac{\lambda_2}{\lambda_1+\lambda_2}\right)^{10-k}$

1. Competing exponentials
2. Sum of independent Poisson rvs
3. Poisson arrivals during and Exponential interval

Competing Exponential (1)

- Two independent light bulbs have life times $T_a \sim \text{Exp}(\lambda_a)$ and $T_b \sim \text{Exp}(\lambda_b)$.

- Two independent light bulbs have life times $T_a \sim \text{Exp}(\lambda_a)$ and $T_b \sim \text{Exp}(\lambda_b)$.

(Q) Distribution of $Z = \min\{T_a, T_b\}$, the first time when a bulb burns out?

Competing Exponential (1)

- Two independent light bulbs have life times $T_a \sim \text{Exp}(\lambda_a)$ and $T_b \sim \text{Exp}(\lambda_b)$.

(Q) Distribution of $Z = \min\{T_a, T_b\}$, the first time when a bulb burns out?

- Approach 1
  - $\mathbb{P}(Z \geq z) = \mathbb{P}(T_a \geq z)\mathbb{P}(T_b \geq z) = e^{-\lambda_a z} e^{-\lambda_b z} = e^{-(\lambda_a + \lambda_b)z}$

Competing Exponential (1)

- Two independent light bulbs have life times $T_a \sim \text{Exp}(\lambda_a)$ and $T_b \sim \text{Exp}(\lambda_b)$.

(Q) Distribution of $Z = \min\{T_a, T_b\}$, the first time when a bulb burns out?

- Approach 1
  - $\mathbb{P}(Z \geq z) = \mathbb{P}(T_a \geq z)\mathbb{P}(T_b \geq z) = e^{-\lambda_a z} e^{-\lambda_b z} = e^{-(\lambda_a + \lambda_b)z}$
  - Thus, $Z \sim \text{Exp}(\lambda_a + \lambda_b)$

Competing Exponential (1)

- Two independent light bulbs have life times $T_a \sim \text{Exp}(\lambda_a)$ and $T_b \sim \text{Exp}(\lambda_b)$.

(Q) Distribution of $Z = \min\{T_a, T_b\}$, the first time when a bulb burns out?

- Approach 1
  - $\mathbb{P}(Z \geq z) = \mathbb{P}(T_a \geq z)\mathbb{P}(T_b \geq z) = e^{-\lambda_a z} e^{-\lambda_b z} = e^{-(\lambda_a + \lambda_b) z}$
  - Thus, $Z \sim \text{Exp}(\lambda_a + \lambda_b)$
- Approach 2
  - T_a and T_b are the first arrival times of two Poisson processes of λ_a and λ_b, respectively.

Competing Exponential (1)

- Two independent light bulbs have life times $T_a \sim \text{Exp}(\lambda_a)$ and $T_b \sim \text{Exp}(\lambda_b)$.

(Q) Distribution of $Z = \min\{T_a, T_b\}$, the first time when a bulb burns out?

- Approach 1
  - $\mathbb{P}(Z \geq z) = \mathbb{P}(T_a \geq z)\mathbb{P}(T_b \geq z) = e^{-\lambda_a z} e^{-\lambda_b z} = e^{-(\lambda_a + \lambda_b)z}$
  - Thus, $Z \sim \text{Exp}(\lambda_a + \lambda_b)$
- Approach 2
  - T_a and T_b are the first arrival times of two Poisson processes of λ_a and λ_b, respectively.
  - Z is the first arrival time of merged Poisson process $(\lambda_a + \lambda_b)$.

Competing Exponential (1)

- Two independent light bulbs have life times $T_a \sim \text{Exp}(\lambda_a)$ and $T_b \sim \text{Exp}(\lambda_b)$.

(Q) Distribution of $Z = \min\{T_a, T_b\}$, the first time when a bulb burns out?

- Approach 1
    - $\mathbb{P}(Z \geq z) = \mathbb{P}(T_a \geq z)\mathbb{P}(T_b \geq z) = e^{-\lambda_a z} e^{-\lambda_b z} = e^{-(\lambda_a + \lambda_b) z}$
    - Thus, $Z \sim \text{Exp}(\lambda_a + \lambda_b)$
- Approach 2
    - T_a and T_b are the first arrival times of two Poisson processes of λ_a and λ_b, respectively.
    - Z is the first arrival time of merged Poisson process $(\lambda_a + \lambda_b)$.
    - Thus, $Z \sim \text{Exp}(\lambda_a + \lambda_b)$

- Three independent light bulbs have life time $T \sim \text{Exp}(\lambda)$.

(Q) $\mathbb{E}$[time until the last bulb burns out]?

Competing Exponential (2)

- Three independent light bulbs have life time $T \sim \text{Exp}(\lambda)$.

(Q) $\mathbb{E}$[time until the last bulb burns out]?

- Understanding from the merged Poisson process

Competing Exponential (2)

- Three independent light bulbs have life time $T \sim \text{Exp}(\lambda)$.

(Q) $\mathbb{E}$[time until the last bulb burns out]?

- Understanding from the merged Poisson process
  - T_1: time until the first burn-out, T_2: time until the second burn-out, T_3: time until the third burn-out
  - $\mathbb{E}(T_1 + T_2 + T_3)$?

- Three independent light bulbs have life time $T \sim \text{Exp}(\lambda)$.

(Q) $\mathbb{E}$[time until the last bulb burns out]?

- Understanding from the merged Poisson process
  - T_1: time until the first burn-out, T_2: time until the second burn-out, T_3: time until the third burn-out
  - $\mathbb{E}(T_1 + T_2 + T_3)$?
  - $\text{PP}(3\lambda) \xrightarrow{\text{1st burn out}} \text{PP}(2\lambda) \xrightarrow{\text{2st burn out}} \text{PP}(\lambda)$
  - $T_1 \sim \text{Exp}(3\lambda)$, $T_2 \sim \text{Exp}(2\lambda)$, $T_3 \sim \text{Exp}(\lambda)$

Competing Exponential (2)

- Three independent light bulbs have life time $T \sim \text{Exp}(\lambda)$.

(Q) $\mathbb{E}[\text{time until the last bulb burns out}]$?

- Understanding from the merged Poisson process
  - T_1: time until the first burn-out, T_2: time until the second burn-out, T_3: time until the third burn-out
  - $\mathbb{E}(T_1 + T_2 + T_3)$?
  - $\text{PP}(3\lambda) \xrightarrow{\text{1st burn out}} \text{PP}(2\lambda) \xrightarrow{\text{2st burn out}} \text{PP}(\lambda)$
  - $T_1 \sim \text{Exp}(3\lambda),\ T_2 \sim \text{Exp}(2\lambda),\ T_3 \sim \text{Exp}(\lambda)$

$$\mathbb{E}\Big[T_1 + T_2 + T_3\Big] = \frac{1}{3\lambda} + \frac{1}{2\lambda} + \frac{1}{\lambda}$$

Sum of Independent Poisson rvs

- Two independent rvs X and Y, where $X \sim \text{Poisson}(\mu)$ and $Y \sim \text{Poisson}(\nu)$.

- Two independent rvs X and Y, where $X \sim \text{Poisson}(\mu)$ and $Y \sim \text{Poisson}(\nu)$.
- Question. Distribution of $X + Y$? Complex convolution, but any other easy way?

Sum of Independent Poisson rvs

- Two independent rvs X and Y, where $X \sim \text{Poisson}(\mu)$ and $Y \sim \text{Poisson}(\nu)$.
- Question. Distribution of $X + Y$? Complex convolution, but any other easy way?
- Poisson process perspective
  - X: number of arrivals of PP(1) over a time interval of length μ
  - Y: number of arrivals of PP(1) over a time interval of length ν
  - Two intervals do not overlap and located in a consecutive manner $\Longrightarrow$ $X \perp\!\!\!\perp Y$

KAIST EE

- Two independent rvs X and Y, where $X \sim \text{Poisson}(\mu)$ and $Y \sim \text{Poisson}(\nu)$.
- Question. Distribution of $X + Y$? Complex convolution, but any other easy way?
- Poisson process perspective
  - X: number of arrivals of PP(1) over a time interval of length μ
  - Y: number of arrivals of PP(1) over a time interval of length ν
  - Two intervals do not overlap and located in a consecutive manner $\implies X \perp\!\!\!\perp Y$
- Distribution of $X + Y$: the number of arrivals of PP(1) over a time interval of length $\mu + \nu$

Sum of Independent Poisson rvs

- Two independent rvs X and Y, where $X \sim \text{Poisson}(\mu)$ and $Y \sim \text{Poisson}(\nu)$.
- Question. Distribution of $X + Y$? Complex convolution, but any other easy way?
- Poisson process perspective
  - X: number of arrivals of PP(1) over a time interval of length μ
  - Y: number of arrivals of PP(1) over a time interval of length ν
  - Two intervals do not overlap and located in a consecutive manner $\implies X \perp\!\!\!\perp Y$
- Distribution of $X + Y$: the number of arrivals of PP(1) over a time interval of length $\mu + \nu$
- Thus, $X + Y \sim \text{Poisson}(\mu + \nu)$

Poisson arrivals during Exponential Interval (1)

- Problem 24, pp. 335
- Consider PP(λ) and an independent rv $T \sim \text{Exp}(\nu)$

Poisson arrivals during Exponential Interval (1)

- Problem 24, pp. 335
- Consider PP(λ) and an independent rv $T \sim \text{Exp}(\nu)$
- Question. Distribution of N_T?

Poisson arrivals during Exponential Interval (1)

- Problem 24, pp. 335
- Consider PP(λ) and an independent rv $T \sim \text{Exp}(\nu)$
- Question. Distribution of N_T?
- Approach 1: Total probability theorem

$$\mathbb{P}(N_T = k) =$$

Poisson arrivals during Exponential Interval (1)

- Problem 24, pp. 335
- Consider PP(λ) and an independent rv $T \sim \text{Exp}(\nu)$
- Question. Distribution of N_T?
- Approach 1: Total probability theorem

$$\mathbb{P}(N_T = k) = \int_0^\infty \mathbb{P}(N_T = k | T = \tau) f_T(\tau) d\tau = \int_0^\infty \mathbb{P}(N_\tau = k) f_T(\tau) d\tau$$

Poisson arrivals during Exponential Interval (1)

- Problem 24, pp. 335
- Consider PP(λ) and an independent rv $T \sim \text{Exp}(\nu)$
- Question. Distribution of N_T?
- Approach 1: Total probability theorem

$$\mathbb{P}(N_T = k) = \int_0^\infty \mathbb{P}(N_T = k | T = \tau) f_T(\tau) d\tau = \int_0^\infty \mathbb{P}(N_\tau = k) f_T(\tau) d\tau$$

- Very tedious and not very intuitive.

Poisson arrivals during Exponential Interval (2)

- Consider PP(λ) and an independent rv $T \sim \text{Exp}(\nu)$

Poisson arrivals during Exponential Interval (2)

- Consider PP(λ) and an independent rv $T \sim$ Exp(ν)
- Consider another PP(ν), and let's view T as the first arrival time in PP(ν).

- Consider $\text{PP}(\lambda)$ and an independent rv $T \sim \text{Exp}(\nu)$
- Consider another $\text{PP}(\nu)$, and let's view T as the first arrival time in $\text{PP}(\nu)$.
- Now, consider the merged process of $\text{PP}(\lambda)$ and $\text{PP}(\nu)$.
  - $\mathbb{P}\Big[\text{from PP}(\lambda)|\text{arrival}\Big] = \frac{\lambda}{\lambda+\nu}$ and $\mathbb{P}\Big[\text{from PP}(\nu)|\text{arrival}\Big] = \frac{\nu}{\lambda+\nu}$

Poisson arrivals during Exponential Interval (2)

- Consider PP(λ) and an independent rv $T \sim \text{Exp}(\nu)$
- Consider another PP(ν), and let's view T as the first arrival time in PP(ν).
- Now, consider the merged process of PP(λ) and PP(ν).
  - $\mathbb{P}\Big[\text{from PP}(\lambda)|\text{arrival}\Big] = \frac{\lambda}{\lambda+\nu}$ and $\mathbb{P}\Big[\text{from PP}(\nu)|\text{arrival}\Big] = \frac{\nu}{\lambda+\nu}$
- K: number of total arrivals until we get the first arrival from PP(ν).

Poisson arrivals during Exponential Interval (2)

- Consider PP(λ) and an independent rv $T \sim$ Exp(ν)
- Consider another PP(ν), and let's view T as the first arrival time in PP(ν).
- Now, consider the merged process of PP(λ) and PP(ν).
    - $\mathbb{P}\Big[\text{from PP}(\lambda)|\text{arrival}\Big] = \frac{\lambda}{\lambda+\nu}$ and $\mathbb{P}\Big[\text{from PP}(\nu)|\text{arrival}\Big] = \frac{\nu}{\lambda+\nu}$
- K: number of total arrivals until we get the first arrival from PP(ν).
    - Then, $K \sim$ Geom($\frac{\nu}{\lambda+\nu}$).

Poisson arrivals during Exponential Interval (2)

- Consider PP(λ) and an independent rv $T \sim \text{Exp}(\nu)$
- Consider another PP(ν), and let's view T as the first arrival time in PP(ν).
- Now, consider the merged process of PP(λ) and PP(ν).
  - $\mathbb{P}\Big[\text{from PP}(\lambda)|\text{arrival}\Big] = \frac{\lambda}{\lambda+\nu}$ and $\mathbb{P}\Big[\text{from PP}(\nu)|\text{arrival}\Big] = \frac{\nu}{\lambda+\nu}$
- K: number of total arrivals until we get the first arrival from PP(ν).
  - Then, $K \sim \text{Geom}(\frac{\nu}{\lambda+\nu})$.
- Let L be the number of arrivals from PP(λ) until we get the first arrival from PP(ν).

$$p_L(l) = p_K(l+1) = \left(\frac{\nu}{\lambda+\nu}\right)\left(\frac{\lambda}{\lambda+\nu}\right)^l, \quad l = 0, 1, \ldots$$

Approximation of Binomial: Poisson vs. Normal

- Consider $S_n \sim \text{Binomial}(n, p)$. Then, $S_n = \sum_{i=1}^{n} X_i$, where $X_i \sim \text{Bern}(p)$.

- Consider $S_n \sim \text{Binomial}(n, p)$. Then, $S_n = \sum_{i=1}^{n} X_i$, where $X_i \sim \text{Bern}(p)$.
- Poisson approximation. $\text{Poisson}(np)$ is a good approximation of S_n

- Consider $S_n \sim \text{Binomial}(n, p)$. Then, $S_n = \sum_{i=1}^{n} X_i$, where $X_i \sim \text{Bern}(p)$.
- Poisson approximation. $\text{Poisson}(np)$ is a good approximation of S_n
- Normal approximation. $\sigma\sqrt{n}Z + n\mu$ is a good approximation of S_n

- Consider $S_n \sim \text{Binomial}(n, p)$. Then, $S_n = \sum_{i=1}^n X_i$, where $X_i \sim \text{Bern}(p)$.
- Poisson approximation. Poisson(np) is a good approximation of S_n
    - Holds when $np = \lambda$ $(n \to \infty,\ p \to 0)$, i.e., X_i's behavior changes over n.
- Normal approximation. $\sigma\sqrt{n}Z + n\mu$ is a good approximation of S_n

Approximation of Binomial: Poisson vs. Normal

- Consider $S_n \sim \text{Binomial}(n, p)$. Then, $S_n = \sum_{i=1}^{n} X_i$, where $X_i \sim \text{Bern}(p)$.
- Poisson approximation. $\text{Poisson}(np)$ is a good approximation of S_n
  - Holds when $np = \lambda$ ($n \to \infty$, $p \to 0$), i.e., X_i's behavior changes over n.
- Normal approximation. $\sigma\sqrt{n}Z + n\mu$ is a good approximation of S_n
  - Holds for a fixed p ($n \to \infty$), i.e., X_i's behavior does not change over n.

Approximation of Binomial: Poisson vs. Normal

- Consider $S_n \sim \text{Binomial}(n, p)$. Then, $S_n = \sum_{i=1}^n X_i$, where $X_i \sim \text{Bern}(p)$.
- Poisson approximation. $\text{Poisson}(np)$ is a good approximation of S_n
  - Holds when $np = \lambda$ ($n \to \infty$, $p \to 0$), i.e., X_i's behavior changes over n.
- Normal approximation. $\sigma\sqrt{n}Z + n\mu$ is a good approximation of S_n
  - Holds for a fixed p ($n \to \infty$), i.e., X_i's behavior does not change over n.
- In practice, an actual numbers of n and p are given, so which approximation is good under what situation?

KAIST EE

- Consider $S_n \sim \text{Binomial}(n, p)$. Then, $S_n = \sum_{i=1}^{n} X_i$, where $X_i \sim \text{Bern}(p)$.
- Poisson approximation. $\text{Poisson}(np)$ is a good approximation of S_n
  - Holds when $np = \lambda$ $(n \to \infty,\ p \to 0)$, i.e., X_i's behavior changes over n.
- Normal approximation. $\sigma\sqrt{n}Z + n\mu$ is a good approximation of S_n
  - Holds for a fixed p $(n \to \infty)$, i.e., X_i's behavior does not change over n.
- In practice, an actual numbers of n and p are given, so which approximation is good under what situation?
  - $p = 1/100$, $n = 100$: $np = 1$, very asymmetric X_i, small p $\implies$ Poisson

KAIST EE

- Consider $S_n \sim \text{Binomial}(n, p)$. Then, $S_n = \sum_{i=1}^{n} X_i$, where $X_i \sim \text{Bern}(p)$.
- Poisson approximation. Poisson(np) is a good approximation of S_n
  - Holds when $np = \lambda$ ($n \to \infty$, $p \to 0$), i.e., X_i's behavior changes over n.
- Normal approximation. $\sigma\sqrt{n}Z + n\mu$ is a good approximation of S_n
  - Holds for a fixed p ($n \to \infty$), i.e., X_i's behavior does not change over n.
- In practice, an actual numbers of n and p are given, so which approximation is good under what situation?
  - $p = 1/100$, $n = 100$: $np = 1$, very asymmetric X_i, small $p \implies$ Poisson
  - $p = 1/3$, $n = 100$: large, reasonly symmetric p, at least moderate $n \implies$ Normal

Approximation of Binomial: Poisson vs. Normal

- Consider $S_n \sim \text{Binomial}(n, p)$. Then, $S_n = \sum_{i=1}^{n} X_i$, where $X_i \sim \text{Bern}(p)$.
- Poisson approximation. $\text{Poisson}(np)$ is a good approximation of S_n
  - Holds when $np = \lambda$ $(n \to \infty,\ p \to 0)$, i.e., X_i's behavior changes over n.
- Normal approximation. $\sigma\sqrt{n}Z + n\mu$ is a good approximation of S_n
  - Holds for a fixed p $(n \to \infty)$, i.e., X_i's behavior does not change over n.
- In practice, an actual numbers of n and p are given, so which approximation is good under what situation?
  - $p = 1/100$, $n = 100$: $np = 1$, very asymmetric X_i, small $p \implies$ Poisson
  - $p = 1/3$, $n = 100$: large, reasonly symmetric p, at least moderate $n \implies$ Normal
  - $p = 1/100$, $n = 10,000$: small p, but large $n \implies$ Both Poisson and Normal

(1) Introduction of Random Processes

(2) Bernoulli Processes

(3) Poisson Processes: Poisson RV and Basic Idea

(4) Poisson Processes: Definition and Properties

(5) Playing with Bernoulli and Poisson Processes

(6) Random Incidence Phenomenon

Example: Survey of Utilization of Town Buses

- What we want to survey: How available are town buses in a city?

Example: Survey of Utilization of Town Buses

- What we want to survey: How available are town buses in a city?
- Two approaches

- What we want to survey: How available are town buses in a city?

- Two approaches

  M1. (1) select a few buses at random, and (2) calculate the average number of riders in the selected bus

Example: Survey of Utilization of Town Buses

- What we want to survey: How available are town buses in a city?
- Two approaches

  M1. (1) select a few buses at random, and (2) calculate the average number of riders in the selected bus

  M2. (1) select a few bus riders at random, (2) look at the buses that they rode, and (3) calculate the average number of riders in those buses

Example: Survey of Utilization of Town Buses

- What we want to survey: How available are town buses in a city?

- Two approaches

  M1. (1) select a few buses at random, and (2) calculate the average number of riders in the selected bus

  M2. (1) select a few bus riders at random, (2) look at the buses that they rode, and (3) calculate the average number of riders in those buses

- Which is correct?

Example: Survey of Utilization of Town Buses

- What we want to survey: How available are town buses in a city?

- Two approaches

  M1. (1) select a few buses at random, and (2) calculate the average number of riders in the selected bus

  M2. (1) select a few bus riders at random, (2) look at the buses that they rode, and (3) calculate the average number of riders in those buses

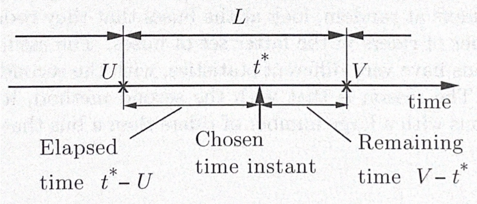
- Which is correct?

- (i) M1 $=$ M2? (ii) M1 $>$ M2? (iii) M1 $<$ M2?

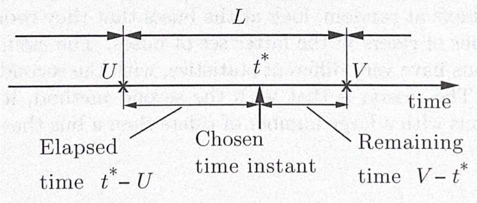
Random Incidence Phenomemon (1)

- We know: in PP(λ), inter-arrival time $\sim$ Exp(λ)

- We know: in PP(λ), inter-arrival time $\sim$ Exp(λ)
- Fix a time instant $t^\star$, and consider the length L of the inter-arrival interval that constains $t^\star$.



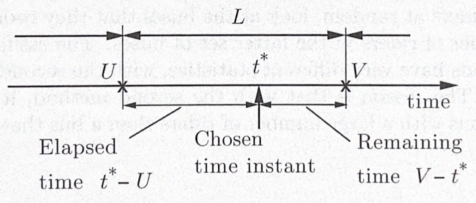
- We know: in PP(λ), inter-arrival time $\sim$ Exp(λ)
- Fix a time instant $t^\star$, and consider the length L of the inter-arrival interval that constains $t^\star$.



- Practical context: Yung shows up at the bus station at some arbitrary time $t^\star$ and records the time from the previous bus arrival (U) until the next bus arrival (V)

Random Incidence Phenomemon (1)

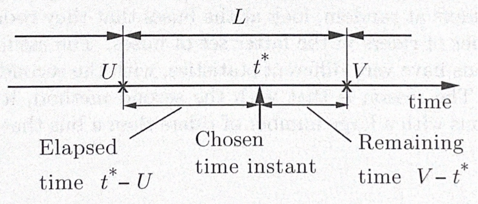
- We know: in PP(λ), inter-arrival time $\sim$ Exp(λ)
- Fix a time instant $t^\star$, and consider the length L of the inter-arrival interval that constains $t^\star$.

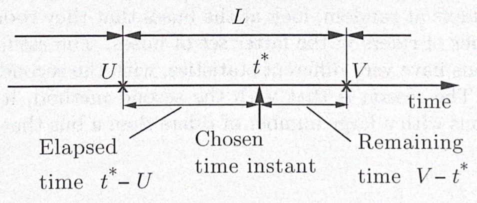


- Practical context: Yung shows up at the bus station at some arbitrary time $t^\star$ and records the time from the previous bus arrival (U) until the next bus arrival (V)
- Question. What is the distribution of L?

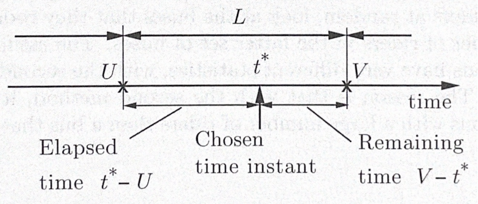
VIDEO PAUSE

Random Incidence Phenomemon (2)

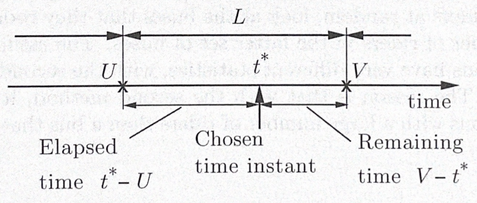




- $t^\star$ is not random, so "random incidence" may be confusing and misleading.



- $t^\star$ is not random, so "random incidence" may be confusing and misleading.
- Assumption. For simplicity, $t^\star$ is large enough that we must have an arrival before $t^\star$ $(U > 0)$

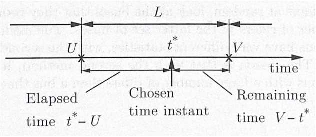


- $t^\star$ is not random, so "random incidence" may be confusing and misleading.
- Assumption. For simplicity, $t^\star$ is large enough that we must have an arrival before $t^\star$ ($U > 0$)
- One might superficially argue that $L \sim \text{Exp}(\lambda)$, but it is NOT.

$$\boxed{L = (t^\star - U) + (V - t^\star)}$$

- $V - t^\star$:

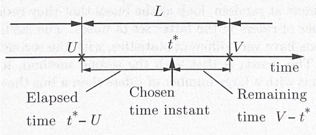
- $t^\star - U$:



$$\boxed{L = (t^\star - U) + (V - t^\star)}$$

- $V - t^\star$: Due to the memoryless property and the fresh-restart,

- $t^\star - U$:

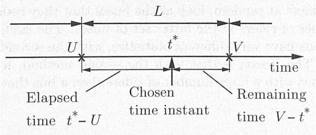


$$L = (t^\star - U) + (V - t^\star)$$

- $V - t^\star$: Due to the memoryless property and the fresh-restart,

  Thus, $V - t^\star \sim \mathrm{Exp}(\lambda)$

- $t^\star - U$:

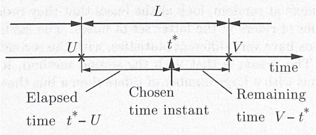


$$\boxed{L = (t^\star - U) + (V - t^\star)}$$

- $V - t^\star$: Due to the memoryless property and the fresh-restart,

  Thus, $\quad V - t^\star \sim \text{Exp}(\lambda)$

- $t^\star - U$: If we run the PP(λ) backwards in time, it remains Poisson. Why?



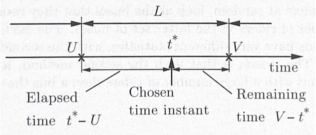
$$L = (t^\star - U) + (V - t^\star)$$

- $V - t^\star$: Due to the memoryless property and the fresh-restart,

  Thus, $\quad V - t^\star \sim \text{Exp}(\lambda)$

- $t^\star - U$: If we run the PP(λ) backwards in time, it remains Poisson. Why? More formally,

$$\mathbb{P}(t^\star - U > x) = \mathbb{P}(\text{no arrivals over } [t^\star - x, t^\star])$$
$$= e^{-\lambda x} = \mathbb{P}(T_{\text{inter}} > x)$$



$$L = (t^\star - U) + (V - t^\star)$$

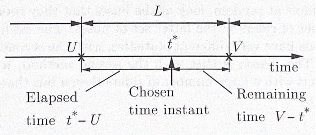
- $V - t^\star$: Due to the memoryless property and the fresh-restart,

  Thus, $\quad V - t^\star \sim \text{Exp}(\lambda)$

- $t^\star - U$: If we run the PP(λ) backwards in time, it remains Poisson. Why? More formally,

  $$\mathbb{P}(t^\star - U > x) = \mathbb{P}(\text{no arrivals over } [t^\star - x, t^\star])$$
  $$= e^{-\lambda x} = \mathbb{P}(T_{\text{inter}} > x)$$

  Thus, $\quad t^\star - U \sim \text{Exp}(\lambda)$



Distribution of L (2)

$$\boxed{L = (t^\star - U) + (V - t^\star)}$$

Distribution of L (2)

$$\boxed{L = (t^\star - U) + (V - t^\star)}$$

- $L = X_1 + X_2$, where $X_1, X_2 \sim \mathrm{Exp}(\lambda)$

Distribution of L (2)

$$\boxed{L = (t^\star - U) + (V - t^\star)}$$

- $L = X_1 + X_2$, where $X_1, X_2 \sim \mathrm{Exp}(\lambda)$
- Time until we have two arrivals in $\mathrm{PP}(\lambda)$

Distribution of L (2)

$$\boxed{L = (t^\star - U) + (V - t^\star)}$$

- $L = X_1 + X_2$, where $X_1, X_2 \sim \text{Exp}(\lambda)$
- Time until we have two arrivals in PP(λ)
- Erlang random variable with parameter $(2, \lambda)$, i.e.,

page 37

$$f_L(l) = \lambda^2 \cdot l \cdot e^{-\lambda l}, \quad l \geq 0$$

Distribution of L (2)

$$\boxed{L = (t^\star - U) + (V - t^\star)}$$

- $L = X_1 + X_2$, where $X_1, X_2 \sim \text{Exp}(\lambda)$
- Time until we have two arrivals in PP(λ)
- Erlang random variable with parameter $(2, \lambda)$, i.e.,

$$f_L(l) = \lambda^2 \cdot l \cdot e^{-\lambda l}, \quad l \geq 0$$

page 37

- Mean $= 2/\lambda$

Distribution of L (2)

$$\boxed{L = (t^{\star} - U) + (V - t^{\star})}$$

- $L = X_1 + X_2$, where $X_1, X_2 \sim \text{Exp}(\lambda)$
- Time until we have two arrivals in $\text{PP}(\lambda)$
- Erlang random variable with parameter $(2, \lambda)$, i.e., page 37

$$f_L(l) = \lambda^2 \cdot l \cdot e^{-\lambda l}, \quad l \geq 0$$

- Mean $= 2/\lambda$
- Why not $\text{Exp}(\lambda)$?

$$\boxed{L = (t^\star - U) + (V - t^\star)}$$

- $L = X_1 + X_2$, where $X_1, X_2 \sim \text{Exp}(\lambda)$
- Time until we have two arrivals in $\text{PP}(\lambda)$
- Erlang random variable with parameter $(2, \lambda)$, i.e., page 37

$$f_L(l) = \lambda^2 \cdot l \cdot e^{-\lambda l}, \quad l \geq 0$$

- Mean $= 2/\lambda$
- Why not $\text{Exp}(\lambda)$? An observer who arrives at an arbitrary time is more likely to fall in a large rather than a small interarrival interval.

KAIST EE

- Two Approaches

  M1. (1) select a few buses at random, and (2) calculate the average number of riders in the selected bus

  M2. (1) select a few bus riders at random, (2) look at the buses that they rode, and (3) calculate the average number of riders in those buses

- (i) M1 $=$ M2? (ii) M1 $>$ M2? (iii) M1 $<$ M2?
- Answer: M1 $<$ M2
- More likely to select a bus with a large number of riders than a bus that is near-empty.

Questions?

1) Explain what is random process. Why do we need such a concept?
2) Explain Bernoulli processes. When do we use Bernoulli processes?
3) Explain Poisson processes. When do we use Poisson processes?
4) What are the relations between those two processces? What features do they share?
5) In both processces, ho do we compute (i) number of arrivals over a given interval of time, (ii) time until the first arrival, (iii) time until k-th arrival?
6) What is Pascal rvs and Erlang rvs?
7) What is the "stopping time" and how is it related to fresh-restart?
8) See many examples on how merging and splitting of BP and PP can be used for intuitive soloving of many problems.