

Lecture 7: Random Processes, Part I

Yi, Yung (이용)

EE210: Probability and Introductory Random Processes KAIST EE

June 2, 2021

Roadmap



- (1) Introduction of Random Processes
- (2) Bernoulli Processes
- (3) Poisson Processes: Poisson RV and Basic Idea
- (4) Poisson Processes: Definition and Properties
- (5) Playing with Bernoulli and Poisson Processes

Roadmap

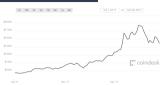


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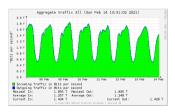
Things that evolve in time

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Many probabilistic experiments that evolve in time



(a) Prices of a crytocurrency

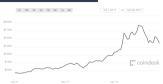


(b) Internet traffic traces

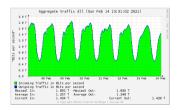
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 - Sequence of daily prices of a stock
 - Sequence of scores in football
 - Sequence of failure times of a machine
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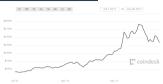


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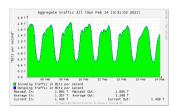
Things that evolve in time



- Many probabilistic experiments that evolve in time
 - Sequence of daily prices of a stock
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- Random process is a mathematical model for it.



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 - $X(3.7, \omega_1) = 3409, X(2, \omega_2) = 5000, X(7.8, \omega_3) = 2800, \text{ etc.}$





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 - Long-term behavior: What is the fraction of times that a stock price is above 3000?
 - Other interesting questions, depending on the target random process

L8(1)

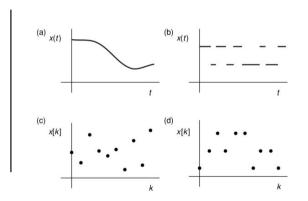
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4 Types of Random Processes



- Types of time and value

- (a) continuous time, continuous value
- (b) continuous time, discrete value
- (c) discrete time, continuous value
- (d) discrete time, discrete value



Random Processes in This Course



- The simplest RP
- discrete time

Jacob Bernoulli (1654 - 1705), Swiss



Simeon Denis Poisson (1781 - 1840), France



Andrey Markov (1856 - 1922), Russia





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- $X_i \perp \{X_{i-1}, X_{i-2}, \dots, X_1\}$

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- X_i depends on X_{i-1}, but $\perp \!\!\!\perp \{X_{i-2}, X_{i-2}, \ldots, X_1\}$
- Markov Chain (MC)



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- Bernoulli Process (BP)
- "today" independent of "past"

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- "today" depends only on "yesterday"



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L8(2)



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- Discrete time, discrete value
- One of the simplest random processes
- A type of "arrival" process

L8(2)

Bernoulli Process: Questions



Please pause the video and write down the questions that you want to ask about Bernoulli process.

VIDEO PAUSE

Q1.

Q2.

Q3.

Q4.

Q5.





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• $T_1 \sim \mathsf{Geom}(p)$



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- Still, geometric.
- But, more than that, as we will see.
 Independence across time slots leads to many useful properties, allowing the quick solution of many problems.



Independence across slots \implies the fresh-start anytime when I look at the process?

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$$U = X_1 + X_2 \perp \!\!\! \perp V = X_5 + X_6$$
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(Q4) After time n = 6, I start to look at the process $(X_n)_{n=6}^{\infty}$?



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Independence across slots \implies the fresh-start anytime when I look at the process?

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L8(2)



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- Fresh-start after a deterministic time n (doesn't matter what happened until n = 5.
- If you watch the on-going Bernoulli process(p) from some time n, you still see the same Bernoulli process(p).

L8(2)



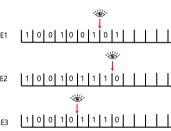
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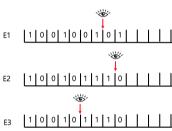




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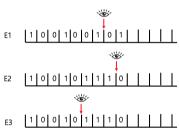
E1. Time of 3rd arrival





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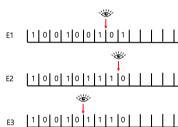
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Fresh-start after Random time N(1)



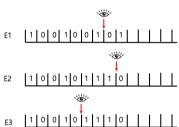
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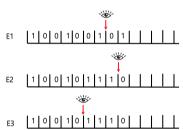


- Difference of N from n
 - The time when I watch the on-going Bernoulli process is random.
 - N is a random variable, i.e., $N: \Omega \mapsto \mathbb{R}$. What is Ω ?

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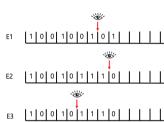
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- Difference of *N* from *n*
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- Do we experience the fresh-start for any N? E1, E2, and E3?

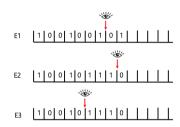


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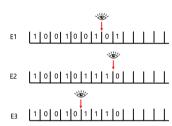
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E1. When I watch the process, N has been already determined. Yes



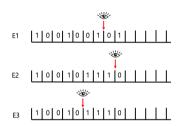
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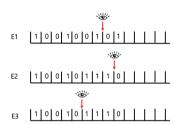


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<u>Fresh-start after</u> Random *N* (2)



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- **E1.** When I watch the process, N has been already determined. Yes
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- E3. Need the future knowledge. '111' does not become random. No
- The question of N = n? can be answered just from the knowledge about X_1, X_2, \dots, X_n ? Then, Yes! (see pp. 301 for more formal description)



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L8(2) June 2, 2021



• In probability theory, a random time N is said to be a stopping time, if the question of "N = n?" can be answered only from the present and the past knowledge of X_1, X_2, \ldots, X_n .



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VIDEO PAUSE

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- Yes. Time when 10 consecutive arrivals have been observed
- No. Time of 2nd arrival in 10 consecutive arrivals



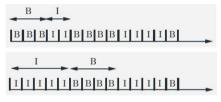
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• Regard an arrival as a server being busy (just for our easy understanding)



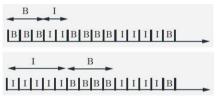
18 / 1

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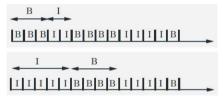


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L8(2)



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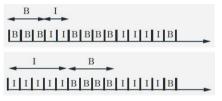


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L8(2)



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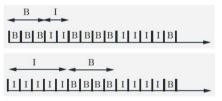
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L8(2)

18 / 1



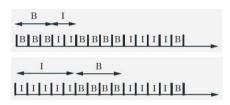
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L8(2)



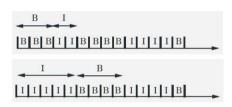


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L8(2)



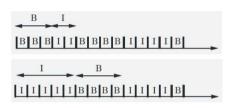
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L8(2) June 2, 2021



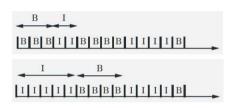


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L8(2)



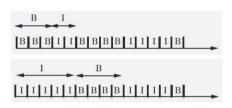
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L8(2) June 2, 2021





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- B_3, B_4, \dots ?

L8(2)



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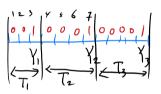
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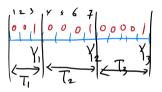


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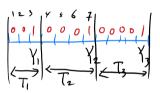


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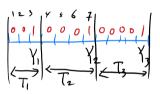
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VIDEO PALISE



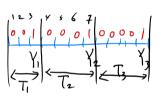


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- We know Y_k 's expectation and variance: $\mathbb{E}[Y_k] = \frac{k}{p}$, $\text{var}[Y_k] = \frac{k(1-p)}{p^2}$, but its distribution?

PMF of Y_k



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= $\mathbb{P}(X_k = 1) \cdot \mathbb{P}(k - 1 \text{ arrivals during the first } t - 1 \text{ slots})$

L8(2)

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PMF of Y_{k}



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$$\begin{split} \mathbb{P}(Y_k = t) &= \mathbb{P}\left(X_k = 1 \text{ and } k-1 \text{ arrivals during the first } t-1 \text{ slots}\right) \\ &= \mathbb{P}\left(X_k = 1\right) \cdot \mathbb{P}\left(k-1 \text{ arrivals during the first } t-1 \text{ slots}\right) \\ &= p \times \binom{t-1}{k-1} p^{k-1} (1-p)^{t-k} = \binom{t-1}{k-1} p^k (1-p)^{t-k}, \quad t = k, k+1, \dots \end{split}$$

L8(2)

Pascal Random Variable with Parameter (k, p)



• In the sequence of Bernoulli trials, the time Y_k of k-th success

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• Pascal(1, p) = Geom(p)

Roadmap



- (1) Introduction of Random Processes
- (2) Bernoulli Processes
- (3) Poisson Processes: Poisson RV and Basic Idea
- (4) Poisson Processes: Definition and Properties
- (5) Playing with Bernoulli and Poisson Processes



• A random variable $S \sim \text{Bin}(n, p)$: Models the number of successes in a given number n of independent trials with success probability p.

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L8(3)





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- $var(Z) = \lambda$ (because np(1-p) is the variance of binomial rv)



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• Remind. Geometric vs. Exponential

L4(3)

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- Need a "modeling sense" to make this possible. It's a good practice for engineers!
- VIDEO PAUSE



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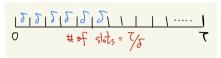
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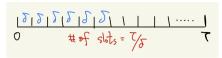
• What's the limit as $\delta \to 0$ (equivalently, $n \to \infty$)





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 $\begin{array}{c} \mathbb{P}(1 \text{ arrival}) \propto \text{arrival rate and slot length} \\ \mathbb{P}(\geq \! 2 \text{ arrivals}) \propto \text{something, but very small} \\ \text{ for small sloth length} \\ \mathbb{P}(0 \text{ arrival}) = 1 - \mathbb{P}(1 \text{ arrival or } \geq \! 2 \text{ arrivals}) \end{array}$



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L8(3) June 2, 2021



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- $o(\delta)$: some function that goes to zero faster than δ .
 - Thus, for very small δ , $o(\delta)$ becomes negligible, compared to δ .
 - Example: $o(\delta) = \delta^{\alpha}$, where any $\alpha > 1$



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- # of arrivals over $[0, \tau]$, $\sim \mathsf{Poisson}(\lambda \tau)$
- This is a continuous twin process of Bernous process, which we call Poisson process.

Roadmap



- (1) Introduction of Random Processes
- (2) Bernoulli Processes
- (3) Poisson Processes: Poisson RV and Basic Idea
- (4) Poisson Processes: Definition and Properties
- (5) Playing with Bernoulli and Poisson Processes



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An arriva	l process	is called	a Poisson	process	with ra	te λ , if t	the followir	g are satisfie	d:



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- (Independence) Let N_{τ} be the number of arrivals over the interval $[0,\tau]$. For any $\tau_1,\tau_2>0,\ N_{s+\tau_1}-N_s$ is independent of $N_{t+\tau_2}-N_t$, if $t>s+\tau_1$.
 - The number of arrivals over two disjoint intervals are independent.



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- (Time homogeneity) For any s, the distribution of $N_{s+\tau} N_s$ is equal to that of N_{τ} .
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- (Time homogeneity) For any s, the distribution of $N_{s+\tau} N_s$ is equal to that of N_{τ} .
 - \circ N_{τ} becomes the number of arrivals over any interval of length τ .
- (Small interval probability) Let $\mathbb{P}(k,\tau) = \mathbb{P}(N_{\tau} = k)$, which satisfy:

$$\mathbb{P}(0,\tau) = 1 - \lambda \tau + o(\tau)$$

$$\mathbb{P}(1,\tau) = \lambda \tau + o_1(\tau)$$

$$\mathbb{P}(k,\tau) = o_k(\tau)$$
 for $k = 2,3,\ldots$, where $\lim_{\tau \to 0} \frac{o(\tau)}{\tau} = 0$, $\lim_{\tau \to 0} \frac{o_k(\tau)}{\tau} = 0$



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 - \circ N_{τ} becomes the number of arrivals over any interval of length τ .
- (Distribution of N_{τ}) N_{τ} is the Poisson rv with parameter $\lambda \tau$, i.e., if we let $\mathbb{P}(k,\tau) = \mathbb{P}(N_{\tau} = k)$, we have:

$$\mathbb{P}(k,\tau) = e^{\lambda \tau} \frac{(\lambda \tau)^k}{k!}, \quad k = 0, 1, 2, \dots$$

L8(4)

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(Q1) Number of arrivals of any interval with length $\tau \sim \text{Poisson}(\lambda \tau)$, i.e.,

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- $T \sim \text{Exp}(\lambda)$. Thus, $\mathbb{E}(T) = 1/\lambda$ and $\text{var}(T) = 1/\lambda^2$
 - Continuous twin of geometric rv in Bernoulli process
 - Memoryless



- Receive emails according to a Poisson process at rate $\lambda=5$ messages/hour
- Mean and variance of mails received during a day

• $\mathbb{P}[\text{one new message in the next hour}]$

 \bullet $\mathbb{P}[\text{exactly two msgs during each of the next three hours}]$



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 - Generally, it holds when T is a stopping time.



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Memoryless and Fresh-start Property



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- $\mathbb{E}[Y_k] = k/\lambda$ and $\text{var}[Y_k] = k/\lambda^2$, but what is the distribution of Y_k ?

L8(4)



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• For a given δ , : prob. of k-th arrival over $[y, y + \delta]$.

PDF of $\overline{Y_k}$



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L8(4)



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• This is called Erlang rv.

An Erlang random variable Z with parameter (k, λ) has the following pdf:

$$f_{Z}(z) = \frac{\lambda^{k} z^{k-1} e^{-\lambda z}}{(k-1)!}, \quad z \ge 0$$

L8(4)



-
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, $p = \lambda\delta$, $np = \lambda\tau$

$$0 \qquad \text{# of slots} = \sqrt{\delta}$$

	Bernoulli process	Poisson process
time of arrival	Discrete	Continuous
PMF of $\#$ of arrivals		
Interarrival time		
Time of k -th arrival		
Arrival rate		



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L8(4)



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38 / 1

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Interarrival time	Geometric	Exponential
Time of k -th arrival	Pascal	Erlang
Arrival rate		



$$-n = \tau/\delta, \ p = \lambda\delta, \ np = \lambda\tau$$

$$0 \qquad \text{# of slots} = \frac{7}{\delta}$$

	Bernoulli process	Poisson process
time of arrival	Discrete	Continuous
PMF of $\#$ of arrivals	Binomial	Poisson
Interarrival time	Geometric	Exponential
Time of k -th arrival	Pascal	Erlang
Arrival rate	p/per slot	$\lambda/$ unit time

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(Q4)
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Fresh-start. So,
 $\mathbb{E}[\exp(\lambda)] = 1/\lambda = 1/0.6$
(Q5) $\mathbb{E}[\text{F=total fishing time}]$
 $2 + \mathbb{E}[F - 2] = 2 + \mathbb{P}(F = 2) \cdot 0 + \mathbb{P}(F > 2) \cdot \mathbb{E}[F - 2|F > 2]$
 $= 2 + \mathbb{P}(0, 2) \cdot \frac{1}{\lambda}$



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(Q6) $\mathbb{E}[\mathsf{number of fish}] = \lambda \cdot 2 + \mathbb{P}(0,2) \cdot 1$

Roadmap



- (1) Introduction of Random Processes
- (2) Bernoulli Processes
- (3) Poisson Processes: Poisson RV and Basic Idea
- (4) Poisson Processes: Definition and Properties
- (5) Playing with Bernoulli and Poisson Processes

Description via Inter-arrival Times



Alternative Description of the Bernoulli Process

- 1. Start with a sequence of independent geometric random variables T_1 , T_2 , ..., with common parameter p, and let these stand for the interarrival times.
- 2. Record a success (or arrival) at times T_1 , $T_1 + T_2$, $T_1 + T_2 + T_3$, etc.

Alternative Description of the Poisson Process

- 1. Start with a sequence of independent exponential random variables T_1, T_2, \ldots , with common parameter λ , and let these stand for the interarrival times.
- 2. Record an arrival at times T_1 , $T_1 + T_2$, $T_1 + T_2 + T_3$, etc.



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- Approach 2: Rainy days is a Bernoulli process with arrival. probability p.
- Thus, the answer is p^2 .

Coding of Random Arrivals



- Question. How to make software codes of Bernoulli process with $\it p$ and Poisson process with $\it \lambda$
- Inter-arrival times are very handy.
- Bernoulli process with p: Obtain a sequence of random values following the geometric distribution with parameter p.
- Poisson process with λ : Obtain a sequence of random values following the exponential distribution with parameter λ .

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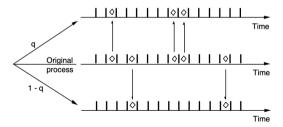
Notations In the Rest of These Slides



- Bernoulli random variable: Bern(p)
- Bernoulli process: BP(p)
- Poisson random variable: Poisson(λ)
- Poisson process: $PP(\lambda)$



• Split BP(p) into two processes. Whenever there is an arrival, keep it w.p. q and discard it w.p. (1-q).

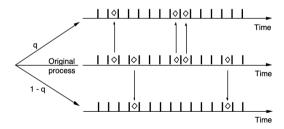


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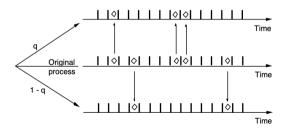


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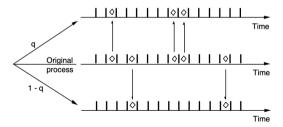


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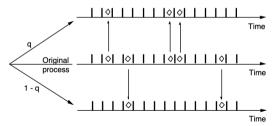


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- BP(pq) and BP(p(1-q)). Why?



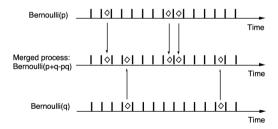


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- Question. What are the two split processes?
- $\mathsf{BP}(pq)$ and $\mathsf{BP}(p(1-q))$. Why?
- Are they independent? No.



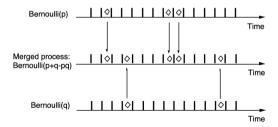


• Merge BP(p) and BP(q) into one process.



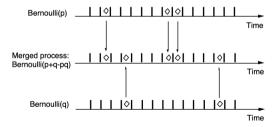


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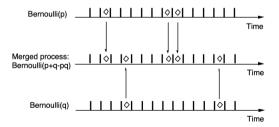


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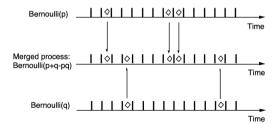


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- $\mathbb{P}(\text{arrival from proc. 1} \mid \text{arrival in the merged proc.}) = \frac{p}{p+q-pq}$





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•
$$\mathbb{P}(1 \text{ arrival}) = p\lambda\delta + p \cdot o(\delta) = p\lambda\delta + o(\delta)$$

•
$$\mathbb{P}(2 \text{ arrivals}) = p \cdot o(\delta) = o(\delta)$$

•
$$\mathbb{P}(0 \text{ arrival}) = 1 - p\lambda\delta - p \cdot o(\delta) - p \cdot o(\delta) = 1 - p\lambda\delta + o(\delta)$$

• PP(λp) and PP($\lambda(1-p)$)



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- Small interval probabilty over δ -interval (ignoring $o(\delta)$ for small δ) $\mathbb{P}(0 \text{ arrival}) \approx (1 \lambda_1 \delta)(1 \lambda_2 \delta) \approx 1 (\lambda_1 + \lambda_2)\delta$ $\mathbb{P}(1 \text{ arrival}) \approx (\lambda_1 \delta)(1 \lambda_2 \delta) + \lambda_2 \delta(1 \lambda_1 \delta) \approx (\lambda_1 + \lambda_2)\delta$

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- Merged process: $PP(\lambda_1 + \lambda_2)$



• Red: $PP(\lambda_1)$ and Blue: $PP(\lambda_2)$



- Red: $PP(\lambda_1)$ and Blue: $PP(\lambda_2)$
- $\mathbb{P}(\text{from red} \mid \text{arrival at time } t \text{ in the merged proc.})$?

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- Red: $PP(\lambda_1)$ and Blue: $PP(\lambda_2)$
- $\mathbb{P}(\text{from red} \mid \text{arrival at time } t \text{ in the merged proc.})? \frac{\lambda_1 \delta}{\lambda_1 \delta + \lambda_2 \delta} = \frac{\lambda_1}{\lambda_1 + \lambda_2}$

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 - $\mathbb{P}(A_k)$? $\frac{\lambda_1}{\lambda_1 + \lambda_2}$
 - \circ A_1, A_2, \ldots are independent: origins (red or blue) of arrivals in the merged proc. are indepdnent
- P(k out of first 10 arrivals are red)?



- Red: $PP(\lambda_1)$ and Blue: $PP(\lambda_2)$
- $\mathbb{P}(\text{from red} \mid \text{arrival at time } t \text{ in the merged proc.})? \frac{\lambda_1 \delta}{\lambda_1 \lambda_2 \lambda_2 \lambda_3} = \frac{\lambda_1}{\lambda_1 \lambda_2 \lambda_3}$
- Consider an event $A_k = \{k$ -th arrival in the merged proc. is red
 - $\mathbb{P}(A_k)$? $\frac{\lambda_1}{\lambda_1 + \lambda_2}$
 - \circ A_1, A_2, \ldots are independent: origins (red or blue) of arrivals in the merged proc. are indepdnent
- $\mathbb{P}(k \text{ out of first } 10 \text{ arrivals are red})?$ $\binom{10}{k} \left(\frac{\lambda_1}{\lambda_1 + \lambda_2}\right)^k \left(\frac{\lambda_2}{\lambda_1 + \lambda_2}\right)^{10-k}$

Using Poisson Processes for Intuitive Problem Solving



- 1. Competing exponentials
- 2. Sum of independent Poisson rvs
- 3. Poisson arrivals during and Exponential interval



• Two independent light bulbs have life times $T_a \sim \text{Exp}(\lambda_a)$ and $T_b \sim \text{Exp}(\lambda_b)$.



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- (Q) Distribution of $Z = \min\{T_a, T_b\}$, the first time when a bulb burns out?
- Approach 1
 - $P(Z \ge z) = \mathbb{P}(T_a \ge z)\mathbb{P}(T_b \ge z) = e^{-\lambda_a z}e^{-\lambda_b z} = e^{-(\lambda_a + \lambda_b)z}$



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- (Q) Distribution of $Z = \min\{T_a, T_b\}$, the first time when a bulb burns out?
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$$\mathbb{E}\Big[T_1+T_2+T_3\Big]=\frac{1}{3\lambda}+\frac{1}{2\lambda}+\frac{1}{\lambda}$$



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• Very tedious and not very intuitive.



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 - Then, $K \sim \text{Geom}(\frac{\nu}{\lambda + \nu})$.
- Let L be the number of arrivals from $PP(\lambda)$ until we get the first arrival from $PP(\nu)$.

$$p_L(I) = p_K(I+1) = \left(\frac{\nu}{\lambda+\nu}\right) \left(\frac{\lambda}{\lambda+\nu}\right)^I, \quad I = 0, 1, \dots$$



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 - p = 1/100, n = 10,000: small p, but large $n \implies Both Poisson and Normal$



Questions?

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Review Questions



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