

Lecture 7: Random Processes, Part I

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EE210: Probability and Introductory Random Processes
KAIST EE

June 2, 2021

- (1) Introduction of Random Processes
- (2) Bernoulli Processes
- (3) Poisson Processes: Poisson RV and Basic Idea
- (4) Poisson Processes: Definition and Properties
- (5) Playing with Bernoulli and Poisson Processes

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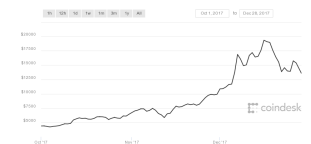
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Roadmap

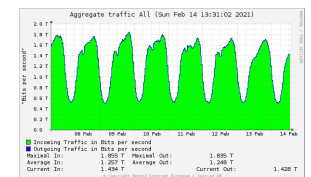
Things that evolve in time

- (1) Introduction of Random Processes
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- Many probabilistic experiments that **evolve in time**
 - Sequence of daily prices of a stock
 - Sequence of scores in football
 - Sequence of failure times of a machine
 - Sequence of traffic loads in Internet
- Random process is a mathematical model for it.



(a) Prices of a cryptocurrency



(b) Internet traffic traces

- A random process is a **sequence** of random variables indexed by **time**.
- Time: discrete or continuous (a modeling choice in most cases)
- Notation
 - $(X_t)_{t \in \mathcal{T}}$ or $(X(t))_{t \in \mathcal{T}}$, where $\mathcal{T} = \mathbb{R}$ (continuous) or $\mathcal{T} = \{0, 1, 2, \dots\}$ (discrete)
 - For the discrete case, we also often use $(X_n)_{n \in \mathbb{Z}_+}$.
 - We will use all of them, unless confusion arises.
- For a **fixed** time t , X_t (or $X(t)$) is a random variable.
- The values that X_t (or $X(t)$) can take: discrete or continuous

L8(1)

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- **Example.** Discrete time RP. $\{X_1, X_2, X_3, \dots\}$.
 - X_i : # of Covid-19 infections at day i in South Korea, which is a random variable.
 - Then, $X_i : \Omega \mapsto \mathbb{R}$, where the sample space Ω is the set of all outcomes.
 - An outcome $\omega \in \Omega$ is a infinite sequence of infections
 - For example, $\omega_1 = (100, 150, 130, \dots)$, $\omega_2 = (200, 300, 400, \dots)$
 - $X_3(\omega_1) = 130$, $X_2(\omega_2) = 300$, $X_1(\omega_2) = 200$, etc.
- **Example.** Continuous time RP. $(X(t))_{t \in \mathbb{R}^+}$
 - $X(t)$: bitcoin price at time t , which is a random variable.
 - Then, $X(t) : \Omega \mapsto \mathbb{R}$.
 - An outcome $\omega \in \Omega$ is a trajectory of prices over $[0, \infty)$
 - $X(3.7, \omega_1) = 3409$, $X(2, \omega_2) = 5000$, $X(7.8, \omega_3) = 2800$, etc.

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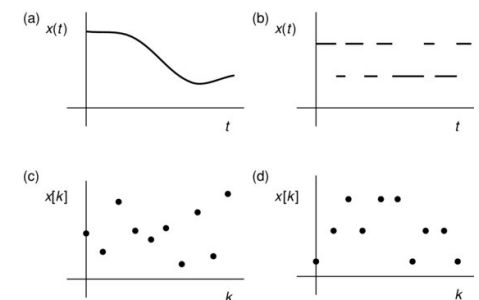
- **Question.** Already studied a sequence (or a collection) of rvs X_1, X_2, \dots, X_n . What's the difference?
 - Assume a discrete time random process for our discussion.
- **Physical difference:** **infinite** sequence of X_1, X_2, \dots .
 - Sample space? set of all outcomes?
 - an outcome: an infinite sequence of sample values x_1, x_2, \dots ,
- **Semantic difference:** Understand i in X_i as time. Also, interesting questions are asked from the random process point of view.
 - **Dependence:** How X_1, X_2, \dots are related to each other as a time series. Prediction of values in the future.
 - **Long-term behavior:** What is the fraction of times that a stock price is above 3000?
 - Other interesting questions, depending on the target random process

L8(1)

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- Types of time and value

- (a) continuous time, continuous value
- (b) continuous time, discrete value
- (c) discrete time, continuous value
- (d) discrete time, discrete value



L8(1)

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- The simplest RP
- discrete time
- $X_i \perp\!\!\!\perp \{X_{i-1}, X_{i-2}, \dots, X_1\}$
- **Bernoulli Process (BP)**
- "today" independent of "past"

Jacob Bernoulli (1654 - 1705),
Swiss



L8(1)

- The simplest RP
- Continuous time version of BP
- $[X(s)]_{s=0}^t \perp\!\!\!\perp [X(s)]_{s=t}^{t+a}$
- **Poisson Process (PP)**
- "today" independent of "past"

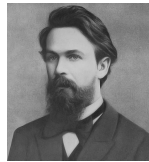
Simeon Denis Poisson (1781 -
1840), France



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- One-step more general than BP/PP
- discrete time
- X_i depends on X_{i-1} , but $\perp\!\!\!\perp \{X_{i-2}, X_{i-3}, \dots, X_1\}$
- **Markov Chain (MC)**
- "today" depends only on "yesterday"

Andrey Markov (1856 - 1922),
Russia



- (1) Introduction of Random Processes
- (2) **Bernoulli Processes**
- (3) Poisson Processes: Poisson RV and Basic Idea
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L8(2)

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- At each "minute", we toss a coin with probability of head $0 < p < 1$.
 - Sequence of lottery wins/loses
 - Customers (each second) to a bank
 - Clicks (at each time slot) to server
- A sequence of **independent Bernoulli trials** X_1, X_2, \dots ,
 - We call index 1, 2, ... **time slots** (or simply slots)

0 0 1 0 0 0 0 1 0 1 1 0 0 →

- Discrete time, discrete value
- One of the simplest random processes
- A type of "arrival" process

Please pause the video and write down the questions that you want to ask about Bernoulli process.

VIDEO PAUSE

Q1.

Q2.

Q3.

Q4.

Q5.

L8(2)

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L8(2)

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(Q1) # of arrivals in the first n slots?

- $S_n = X_1 + X_2 + \dots + X_n$
- $S_n \sim \text{Binomial}(n, p)$
- $\mathbb{E}(S_n) = np$, $\text{var}(S_n) = np(1-p)$
- This will hold for any n consecutive slots.

(Q2) # of slots T_1 until the first arrival?

- $T_1 \sim \text{Geom}(p)$
- $\mathbb{E}(T_1) = 1/p$, $\text{var}(T_1) = \frac{1-p}{p^2}$

- T_1 is geometric? **Memoryless**
- Conditioned no first arrival at some time, what's the distribution of the remaining time until the first arrival?
- Still, geometric.
- But, more than that, as we will see. Independence across time slots leads to many useful properties, allowing the quick solution of many problems.

Independence across slots \implies the fresh-start anytime when I look at the process?(Q3) $U = X_1 + X_2 \perp\!\!\!\perp V = X_5 + X_6$?

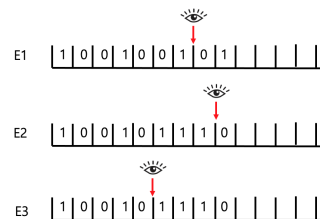
- Yes
- Because X_i s are independent

(Q4) After time $n = 6$, I start to look at the process $(X_n)_{n=6}^\infty$?

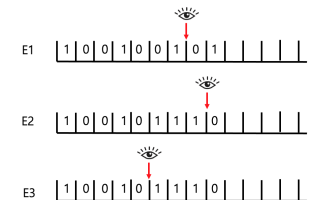
- $(X_1, \dots, X_5) \perp\!\!\!\perp (X_n)_{n=6}^\infty$
- **Fresh-start** after a deterministic time n (doesn't matter what happened until $n = 5$).
- If you watch the on-going Bernoulli process(p) from some time n , you still see the same Bernoulli process(p).

(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \dots)$? Fresh-start even after **random** N ?

- This means that the time I start to look at the process is a random variable.
- Examples of N

E1. Time of 3rd arrival**E2.** First time when 3 consecutive arrivals have been observed**E3.** Time just before 3 consecutive arrivals

- Difference of N from n
 - The time when I watch the on-going Bernoulli process is **random**.
 - N is a random variable, i.e., $N : \Omega \mapsto \mathbb{R}$. What is Ω ?
- Do we experience the fresh-start for any N ? **E1, E2, and E3?**

(Q5) The process $(X_N, X_{N+1}, X_{N+2}, \dots)$? Fresh-start even after random N ?**E1.** Time of 3rd arrival**E2.** First time when 3 consecutive arrivals have been observed**E3.** Time just before 3 consecutive arrivals

- E1.** When I watch the process, N has been already determined. **Yes**
- E2.** Same as **E1**. **Yes**
- E3.** Need the future knowledge. '111' does not become random. **No**

- The question of $N = n$? can be answered just from the knowledge about X_1, X_2, \dots, X_n ? Then, Yes! (see pp. 301 for more formal description)

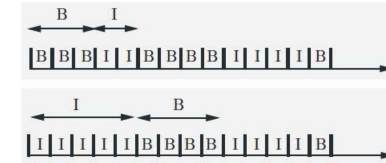
- In probability theory, a random time N is said to be a **stopping time**, if the question of " $N = n$?" can be answered only from the present and the past knowledge of X_1, X_2, \dots, X_n .
- https://en.wikipedia.org/wiki/Stopping_time
- Fresh-start after N in Bernoulli process? **Yes, if N is a stopping time.**
- Please think about two examples of stopping time and not.
 - Yes.** Time when 10 consecutive arrivals have been observed
 - No.** Time of 2nd arrival in 10 consecutive arrivals

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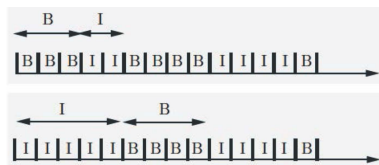
- Regard an arrival as a server being busy (just for our easy understanding)
- First busy period B_1** : starts with the first busy slot and ends just before the first subsequent idle slot



- (Q6) B_1 is a random variable. Distribution of B_1 ?
- N : time of the first busy slot. N is a stopping time?
 - Yes.** Thus, **fresh-start after N** . Because we can answer the question of $N = n$?, just using X_1, X_2, \dots, X_n .
- B_1 is geometric with parameter $(1 - p)$

L8(2)

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- Question.** What about the second busy period B_2 ?
- N : time of the first busy slot of the second busy period. N is a stopping time?
- Yes.** Thus, fresh-start after N .
- Then, B_1 and B_2 are **identically distributed** as $\text{Geom}(1 - p)$.
- B_3, B_4, \dots ?

L8(2)

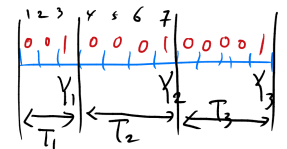
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- Time of the first arrival $Y_1 \sim \text{Geom}(p)$

(Q7) Time of the k -th arrival Y_k ?

VIDEO PAUSE

- $T_k = Y_k - Y_{k-1}$: k -th inter-arrival ($k \geq 2, T_1 = Y_1$)
- $Y_k = T_1 + T_2 + \dots + T_k$.



- After each T_k , the fresh-start occurs.
- $\{T_i\}$ are i.i.d. and $\sim \text{Geom}(p)$
- We know Y_k 's expectation and variance: $\mathbb{E}[Y_k] = \frac{k}{p}$, $\text{var}[Y_k] = \frac{k(1-p)}{p^2}$, but **its distribution?**

L8(2)

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- $Y_k = T_1 + T_2 + \dots + T_k$.
- $\{T_i\}$ are i.i.d. and $\sim \text{Geom}(p)$

$$\begin{aligned}\mathbb{P}(Y_k = t) &= \mathbb{P}(X_k = 1 \text{ and } k-1 \text{ arrivals during the first } t-1 \text{ slots}) \\ &= \mathbb{P}(X_k = 1) \cdot \mathbb{P}(k-1 \text{ arrivals during the first } t-1 \text{ slots}) \\ &= p \times \binom{t-1}{k-1} p^{k-1} (1-p)^{t-k} = \binom{t-1}{k-1} p^k (1-p)^{t-k}, \quad t = k, k+1, \dots\end{aligned}$$

L8(2)

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- In the sequence of Bernoulli trials, the time Y_k of k -th success
- PMF of Y_k

$$\mathbb{P}(Y_k = t) = \begin{cases} \binom{t-1}{k-1} p^k (1-p)^{t-k}, & \text{if } t = k, k+1, \dots \\ 0, & \text{if } t = 1, 2, \dots, k-1 \end{cases}$$

- $\text{Pascal}(1, p) = \text{Geom}(p)$

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L8(3)

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- A random variable $S \sim \text{Bin}(n, p)$: Models the number of successes in a given number n of independent trials with success probability p .

$$p_S(k) = \frac{n!}{(n-k)!k!} p^k (1-p)^{n-k}$$

- Our interest: very large n and very small p , such that $np = \lambda$, i.e., $\lim_{n \rightarrow \infty} p_S(k)$?

$$\begin{aligned}p_S(k) &= \frac{n(n-1) \cdots (n-k+1)}{k!} \cdot \frac{\lambda^k}{n^k} \cdot \left(1 - \frac{\lambda}{n}\right)^{n-k} \\ &= \frac{n}{n} \cdot \frac{(n-1)}{n} \cdots \frac{(n-k+1)}{n} \cdot \frac{\lambda^k}{k!} \cdot \left(1 - \frac{\lambda}{n}\right)^n \cdot \left(1 - \frac{\lambda}{n}\right)^{-k} \xrightarrow{n \rightarrow \infty} e^{-\lambda} \frac{\lambda^k}{k!}\end{aligned}$$

L8(3)

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- A Poisson random variable Z with parameter λ takes nonnegative integer values, whose PMF is:

$$p_Z(k) = e^{-\lambda} \frac{\lambda^k}{k!}, \quad k = 0, 1, \dots$$

- Infinitely many slots (n) with the infinitely small slot duration (thus infinitely small success probability $p = \lambda/n$)
- $\mathbb{E}(Z) = \lambda$ (because $\lambda = np$ is the mean of binomial rv)
- $\text{var}(Z) = \lambda$ (because $np(1-p)$ is the variance of binomial rv)

L8(3)

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- A packet consisting of a string of n symbols is transmitted over a noisy channel.
- Each symbol: erroneous transmission with probability of 0.0001, independent of other symbols. Incorrect transmission is when at least one symbol is in error.
- Question.** How small should n be in order for the probability of incorrect transmission to be less than 0.001?
- p is very small, and n is reasonably large \rightarrow Poisson approximation
- Prob. of incorrect transmission =
 $1 - \mathbb{P}(\text{no symbol error}) = 1 - e^{-\lambda} = 1 - e^{-0.0001n} < 0.0001$
- $n < \frac{-\ln 0.999}{0.0001} = 10.005$

L8(3)

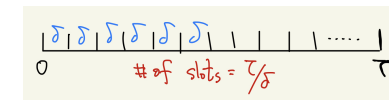
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- Remind.** Geometric vs. Exponential L4(3)
 - Two rvs with memoryless property
 - continuous system = discrete system with infinitely many slots whose duration is infinitely small.
- Independence between what happens in a different time region
- Memoryless and fresh-start property
- Choose one of discrete or continuous versions for our modeling convenience.
- Question.** How do we design the continuous analog of Bernoulli process?
 - Key idea: Making it as a limit of a sequence of Bernoulli processes
- Need a “modeling sense” to make this possible. It’s a good practice for engineers!
- VIDEO PAUSE

L8(3)

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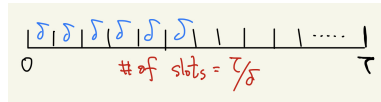
- Continuous twin
 - Key point: Understand the number of arrivals over a given interval $[0, \tau]$.
 - Assume that it has some arrival rate λ (# of arrivals/unit time).
 - We know how to handle Bernoulli process with discrete time slots.
- Divide $[0, \tau]$ into slots whose length = δ . Then, $n = \#$ of slots = $\frac{\tau}{\delta}$.



- What’s the limit as $\delta \rightarrow 0$ (equivalently, $n \rightarrow \infty$)

L8(3)

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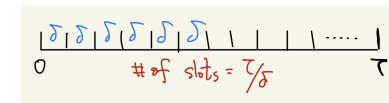


- Now, our design idea: during one time slot of length δ ,

$\mathbb{P}(1 \text{ arrival}) \propto \text{arrival rate and slot length}$
 $\mathbb{P}(\geq 2 \text{ arrivals}) \propto \text{something, but very small}$
 for small slot length
 $\mathbb{P}(0 \text{ arrival}) = 1 - \mathbb{P}(1 \text{ arrival or } \geq 2 \text{ arrivals})$

$$\begin{aligned}
 \mathbb{P}(1 \text{ arrival}) &= \lambda\delta + o(\delta) \\
 \mathbb{P}(\geq 2 \text{ arrivals}) &= o(\delta) \\
 \mathbb{P}(0 \text{ arrival}) &= 1 - \lambda\delta + o(\delta)
 \end{aligned}$$

- $o(\delta)$: some function that goes to zero faster than δ .
 - Thus, for very small δ , $o(\delta)$ becomes negligible, compared to δ .
 - Example: $o(\delta) = \delta^\alpha$, where any $\alpha > 1$



- Our interest: probability of k arrivals over $[0, \tau]$
- Given "small" δ , # of arrivals $\sim \text{Bin}(n, p)$, where $n = \tau/\delta$ and $p = \lambda\delta$
- As $\delta \rightarrow 0$, $np = \tau/\delta \times \lambda\delta = \lambda\tau$.
- # of arrivals over $[0, \tau]$, $\sim \text{Poisson}(\lambda\tau)$
- This is a continuous twin process of Bernoulli process, which we call **Poisson process**.

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An arrival process is called a **Poisson process** with rate λ , if the following are satisfied:

- (Independence)** Let N_τ be the number of arrivals over the interval $[0, \tau]$. For any $\tau_1, \tau_2 > 0$, $N_{s+\tau_1} - N_s$ is independent of $N_{t+\tau_2} - N_t$, if $t > s + \tau_1$.
 - The number of arrivals over two disjoint intervals are independent.
- (Time homogeneity)** For any s , the distribution of $N_{s+\tau} - N_s$ is equal to that of N_τ .
 - N_τ becomes the number of arrivals over any interval of length τ .
- (Small interval probability)** Let $\mathbb{P}(k, \tau) = \mathbb{P}(N_\tau = k)$, which satisfy:

$$\begin{aligned}
 \mathbb{P}(0, \tau) &= 1 - \lambda\tau + o(\tau) \\
 \mathbb{P}(1, \tau) &= \lambda\tau + o_1(\tau) \\
 \mathbb{P}(k, \tau) &= o_k(\tau) \quad \text{for } k = 2, 3, \dots, \quad \text{where } \lim_{\tau \rightarrow 0} \frac{o(\tau)}{\tau} = 0, \quad \lim_{\tau \rightarrow 0} \frac{o_k(\tau)}{\tau} = 0
 \end{aligned}$$

An arrival process is called a **Poisson process** with rate λ , if the following are satisfied:

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 - The number of arrivals over two disjoint intervals are independent.
- **(Time homogeneity)** For any s , the distribution of $N_{s+\tau} - N_s$ is equal to that of N_τ .
 - N_τ becomes the number of arrivals over any interval of length τ .
- **(Distribution of N_τ)** N_τ is the Poisson rv with parameter $\lambda\tau$, i.e., if we let $\mathbb{P}(k, \tau) = \mathbb{P}(N_\tau = k)$, we have:

$$\mathbb{P}(k, \tau) = e^{-\lambda\tau} \frac{(\lambda\tau)^k}{k!}, \quad k = 0, 1, 2, \dots$$

L8(4)

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(Q1) Number of arrivals of any interval with length $\tau \sim \text{Poisson}(\lambda\tau)$, i.e.,

$$\mathbb{P}(k, \tau) = e^{-\lambda\tau} \frac{(\lambda\tau)^k}{k!}, \quad k = 0, 1, 2, \dots$$

$$\mathbb{E}(N_\tau) = \lambda\tau \text{ and } \text{var}(N_\tau) = \lambda\tau$$

(Q2) Time of first arrival T

$$F_T(t) = \mathbb{P}(T \leq t) = 1 - \mathbb{P}(T > t) = 1 - P(0, t) = 1 - e^{-\lambda t}$$

$$f_T(t) = \frac{dF_T(t)}{dt} = \lambda e^{-\lambda t}, \quad t \geq 0.$$

- $T \sim \text{Exp}(\lambda)$. Thus, $\mathbb{E}(T) = 1/\lambda$ and $\text{var}(T) = 1/\lambda^2$
 - Continuous twin of geometric rv in Bernoulli process
 - Memoryless

L8(4)

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- Receive emails according to a Poisson process at rate $\lambda = 5$ messages/hour
- Mean and variance of mails received during a day
 - $5 \cdot 24 = 120$
- $\mathbb{P}[\text{one new message in the next hour}]$
 - $\mathbb{P}(1, 1) = \frac{(5 \cdot 1)^1 e^{-5 \cdot 1}}{1!} = 5e^{-5}$
- $\mathbb{P}[\text{exactly two msgs during each of the next three hours}]$
 - $\left(\frac{5^2 e^{-5}}{2!}\right)^3$

L8(4)

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- **Remind.** Similar property for Bernoulli processes, but here no time slots.
- **Fresh-start at deterministic time t :** Start watching at time t , then you see the Poisson process, independent of what has happened in the past.
- **Fresh-start at random time T :** Similarly, it holds.
 - For example, when you start watching at random time T_1 (time of first arrival).
 - Generally, it holds when T is a **stopping time**.

(Q3) The k -th arrival time Y_k ?

- k -th inter-arrival time $T_k = Y_k - Y_{k-1}$, $k \geq 2$, and $T_1 = Y_1$.
- $Y_k = T_1 + T_2 + \dots + T_k$ is sum of i.i.d. exponential rvs.
- $\mathbb{E}[Y_k] = k/\lambda$ and $\text{var}[Y_k] = k/\lambda^2$, but what is the distribution of Y_k ?

L8(4)

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- For a given δ , $\delta \cdot f_{Y_k}(y)$: prob. of k -th arrival over $[y, y + \delta]$.
- When δ is small, only one arrival occurs. Thus,
 $\delta \cdot f_{Y_k}(y) = \mathbb{P}(\text{an arrival over } [y, y + \delta]) \times \mathbb{P}(k-1 \text{ arrivals before } y)$

$$\approx \lambda \delta \times \mathbb{P}(k-1, y) = \lambda \delta \times \frac{\lambda^{k-1} y^{k-1} e^{-\lambda y}}{(k-1)!}$$

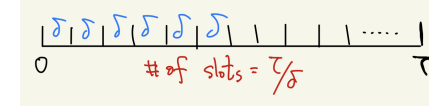
$$f_{Y_k}(y) = \frac{\lambda^k y^{k-1} e^{-\lambda y}}{(k-1)!}, \quad y \geq 0.$$

- This is called **Erlang** rv.

An Erlang random variable Z with parameter (k, λ) has the following pdf:

$$f_Z(z) = \frac{\lambda^k z^{k-1} e^{-\lambda z}}{(k-1)!}, \quad z \geq 0$$

$$- n = \tau/\delta, \quad p = \lambda\delta, \quad np = \lambda\tau$$



	Bernoulli process	Poisson process
time of arrival	Discrete	Continuous
PMF of # of arrivals	Binomial	Poisson
Interarrival time	Geometric	Exponential
Time of k -th arrival	Pascal	Erlang
Arrival rate	p /per slot	λ /unit time

Example: Poisson Fishing (Problem 10, page 329)

Roadmap

- Catching fish: Poisson process $\lambda = 0.6$ /hour.
- Fish for 2 hours. Within 2 hours, if he catches at least one fish, then he stops at 2 hours. Otherwise, he fishes until one fish is caught.

(Q1) $\mathbb{P}(\text{fishing time} > 2 \text{ hours})$

Method 1: $\mathbb{P}(0, 2)$

Method 2: $\mathbb{P}(T_1 > 2)$

(Q2) $\mathbb{P}(2 < \text{fishing time} < 5)$

Method 1: $\mathbb{P}(0, 2)(1 - \mathbb{P}(0, 3))$

Method 2: $\mathbb{P}(2 < T_1 < 5)$

(Q3) $\mathbb{P}(\text{Catch at least two fish})$

Method 1: $\sum_{k=2}^{\infty} = 1 - \mathbb{P}(0, 2) - \mathbb{P}(1, 2)$

Method 2: $\mathbb{P}(Y_2 \leq 2)$

(Q4) $\mathbb{E}[\text{future fi. time} | \text{already fished for 3h}]$

Fresh-start. So,

$$\mathbb{E}[\exp(\lambda)] = 1/\lambda = 1/0.6$$

(Q5) $\mathbb{E}[F = \text{total fishing time}]$

$$2 + \mathbb{E}[F - 2] = 2 + \mathbb{P}(F = 2) \cdot 0 +$$

$$\mathbb{P}(F > 2) \cdot \mathbb{E}[F - 2 | F > 2]$$

$$= 2 + \mathbb{P}(0, 2) \cdot \frac{1}{\lambda}$$

(Q6) $\mathbb{E}[\text{number of fish}] = \lambda \cdot 2 + \mathbb{P}(0, 2) \cdot 1$

(1) Introduction of Random Processes

(2) Bernoulli Processes

(3) Poisson Processes: Poisson RV and Basic Idea

(4) Poisson Processes: Definition and Properties

(5) **Playing with Bernoulli and Poisson Processes**

Alternative Description of the Bernoulli Process

1. Start with a sequence of independent geometric random variables T_1, T_2, \dots , with common parameter p , and let these stand for the interarrival times.
2. Record a success (or arrival) at times $T_1, T_1 + T_2, T_1 + T_2 + T_3$, etc.

Alternative Description of the Poisson Process

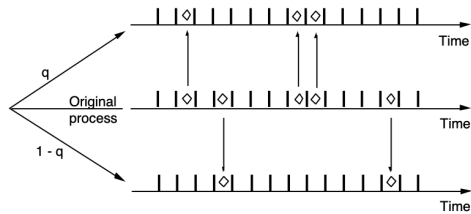
1. Start with a sequence of independent exponential random variables T_1, T_2, \dots , with common parameter λ , and let these stand for the interarrival times.
2. Record an arrival at times $T_1, T_1 + T_2, T_1 + T_2 + T_3$, etc.

- It has been observed that after a rainy day, the number of days until the next rain $\sim \text{Geom}(p)$, independent of the past.
- **Question.** What is the probability that it rains on both the 5th and the 8th day of the month?
- **Approach 1:** Handling this problem directly with the geometric PMFs is very tedious and complex.
- **Approach 2:** Rainy days is a Bernoulli process with arrival probability p .
- Thus, the answer is p^2 .

- **Question.** How to make software codes of Bernoulli process with p and Poisson process with λ
- Inter-arrival times are very handy.
- **Bernoulli process with p :** Obtain a sequence of random values following the **geometric distribution** with parameter p .
- **Poisson process with λ :** Obtain a sequence of random values following the **exponential distribution** with parameter λ .

- Bernoulli random variable: **Bern(p)**
- Bernoulli process: **BP(p)**
- Poisson random variable: **Poisson(λ)**
- Poisson process: **PP(λ)**

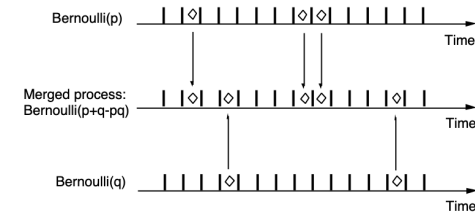
- Split $BP(p)$ into two processes. Whenever there is an arrival, keep it w.p. q and discard it w.p. $(1 - q)$.
- Split decisions are independent of arrivals.
- Question.** What are the two split processes?
- $BP(pq)$ and $BP(p(1 - q))$. Why?
- Are they independent? **No.**



L8(5)

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- Merge $BP(p)$ and $BP(q)$ into one process.
- Collided arrival is regarded just one arrival in the merged process
- Probability of having at least one arrival: $1 - (1 - p)(1 - q) = p + q - pq$
- Merged process: $BP(p + q - pq)$
- $\mathbb{P}(\text{arrival from proc. 1} \mid \text{arrival in the merged proc.}) = \frac{p}{p + q - pq}$



L8(5)

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- Split a Poisson process $PP(\lambda)$ into two processes by keeping each arrival w.p. p and discarding it w.p. $(1 - p)$
- Question.** What are the split processes?
- Let's focus on the process that we keep
- Independence and time-homogeneity? **Yes**
- Small interval probability over δ -interval
 - $\mathbb{P}(1 \text{ arrival}) = p\lambda\delta + p \cdot o(\delta) = p\lambda\delta + o(\delta)$
 - $\mathbb{P}(2 \text{ arrivals}) = p \cdot o(\delta) = o(\delta)$
 - $\mathbb{P}(0 \text{ arrival}) = 1 - p\lambda\delta - p \cdot o(\delta) - p \cdot o(\delta) = 1 - p\lambda\delta + o(\delta)$
- $PP(\lambda p)$ and $PP(\lambda(1 - p))$

L8(5)

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- Merge from $PP(\lambda_1)$ and $PP(\lambda_2)$
- Independence and time-homogeneity? **Yes**
- Small interval probability over δ -interval (ignoring $o(\delta)$ for small δ)
 - $\mathbb{P}(0 \text{ arrival}) \approx (1 - \lambda_1\delta)(1 - \lambda_2\delta) \approx 1 - (\lambda_1 + \lambda_2)\delta$
 - $\mathbb{P}(1 \text{ arrival}) \approx (\lambda_1\delta)(1 - \lambda_2\delta) + \lambda_2\delta(1 - \lambda_1\delta) \approx (\lambda_1 + \lambda_2)\delta$
- Merged process: $PP(\lambda_1 + \lambda_2)$

L8(5)

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- **Red**: $PP(\lambda_1)$ and **Blue**: $PP(\lambda_2)$
- $\mathbb{P}(\text{from red} \mid \text{arrival at time } t \text{ in the merged proc.})? \frac{\lambda_1 \delta}{\lambda_1 \delta + \lambda_2 \delta} = \frac{\lambda_1}{\lambda_1 + \lambda_2}$
- Consider an event $A_k = \{k\text{-th arrival in the merged proc. is red}\}$
 - $\mathbb{P}(A_k)? \frac{\lambda_1}{\lambda_1 + \lambda_2}$
 - A_1, A_2, \dots are independent: origins (red or blue) of arrivals in the merged proc. are independent
- $\mathbb{P}(k \text{ out of first 10 arrivals are red})? \binom{10}{k} \left(\frac{\lambda_1}{\lambda_1 + \lambda_2} \right)^k \left(\frac{\lambda_2}{\lambda_1 + \lambda_2} \right)^{10-k}$

L8(5)

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1. Competing exponentials
2. Sum of independent Poisson rvs
3. Poisson arrivals during and Exponential interval

L8(5)

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- Two independent light bulbs have life times $T_a \sim \text{Exp}(\lambda_a)$ and $T_b \sim \text{Exp}(\lambda_b)$.
- (Q) Distribution of $Z = \min\{T_a, T_b\}$, the first time when a bulb burns out?
 - **Approach 1**
 - $\mathbb{P}(Z \geq z) = \mathbb{P}(T_a \geq z)\mathbb{P}(T_b \geq z) = e^{-\lambda_a z} e^{-\lambda_b z} = e^{-(\lambda_a + \lambda_b)z}$
 - Thus, $Z \sim \text{Exp}(\lambda_a + \lambda_b)$
 - **Approach 2**
 - T_a and T_b are the first arrival times of two Poisson processes of λ_a and λ_b , respectively.
 - Z is the first arrival time of merged Poisson process $(\lambda_a + \lambda_b)$.
 - Thus, $Z \sim \text{Exp}(\lambda_a + \lambda_b)$

L8(5)

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- Three independent light bulbs have life time $T \sim \text{Exp}(\lambda)$.

(Q) $\mathbb{E}[\text{time until the last bulb burns out}]?$

- Understanding from the merged Poisson process
 - T_1 : time until the first burn-out, T_2 : time until the second burn-out, T_3 : time until the third burn-out
 - $\mathbb{E}(T_1 + T_2 + T_3)?$
 - $PP(3\lambda) \xrightarrow{\text{1st burn out}} PP(2\lambda) \xrightarrow{\text{2nd burn out}} PP(\lambda)$
 - $T_1 \sim \text{Exp}(3\lambda), T_2 \sim \text{Exp}(2\lambda), T_3 \sim \text{Exp}(\lambda)$

$$\mathbb{E}[T_1 + T_2 + T_3] = \frac{1}{3\lambda} + \frac{1}{2\lambda} + \frac{1}{\lambda}$$

L8(5)

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- Two independent rvs X and Y , where $X \sim \text{Poisson}(\mu)$ and $Y \sim \text{Poisson}(\nu)$.
- **Question.** Distribution of $X + Y$? Complex convolution, but any other easy way?
- Poisson process perspective
 - X : number of arrivals of PP(1) over a time interval of length μ
 - Y : number of arrivals of PP(1) over a time interval of length ν
 - Two intervals do not overlap and located in a consecutive manner $\implies X \perp\!\!\!\perp Y$
- Distribution of $X + Y$: the number of arrivals of PP(1) over a time interval of length $\mu + \nu$
- Thus, $X + Y \sim \text{Poisson}(\mu + \nu)$

L8(5)

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- Problem 24, pp. 335
- Consider PP(λ) and an independent rv $T \sim \text{Exp}(\nu)$
- **Question.** Distribution of N_T ?

- **Approach 1:** Total probability theorem

$$\mathbb{P}(N_T = k) = \int_0^\infty \mathbb{P}(N_T = k | T = \tau) f_T(\tau) d\tau = \int_0^\infty \mathbb{P}(N_\tau = k) f_T(\tau) d\tau$$

- Very tedious and not very intuitive.

L8(5)

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- Consider PP(λ) and an independent rv $T \sim \text{Exp}(\nu)$
- Consider another PP(ν), and let's view T as the first arrival time in PP(ν).
- Now, consider the merged process of PP(λ) and PP(ν).
 - $\mathbb{P}[\text{from PP}(\lambda) | \text{arrival}] = \frac{\lambda}{\lambda + \nu}$ and $\mathbb{P}[\text{from PP}(\nu) | \text{arrival}] = \frac{\nu}{\lambda + \nu}$
- K : number of total arrivals until we get the first arrival from PP(ν).
 - Then, $K \sim \text{Geom}(\frac{\nu}{\lambda + \nu})$.
- Let L be the number of arrivals from PP(λ) until we get the first arrival from PP(ν).

$$p_L(l) = p_K(l+1) = \left(\frac{\nu}{\lambda + \nu}\right) \left(\frac{\lambda}{\lambda + \nu}\right)^l, \quad l = 0, 1, \dots$$

L8(5)

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- Consider $S_n \sim \text{Binomial}(n, p)$. Then, $S_n = \sum_{i=1}^n X_i$, where $X_i \sim \text{Bern}(p)$.
- **Poisson approximation.** Poisson(np) is a good approximation of S_n
 - Holds when $np = \lambda$ ($n \rightarrow \infty, p \rightarrow 0$), i.e., X_i 's behavior changes over n .
- **Normal approximation.** $\sigma\sqrt{n}Z + n\mu$ is a good approximation of S_n
 - Holds for a fixed p ($n \rightarrow \infty$), i.e., X_i 's behavior does not change over n .
- In practice, an actual numbers of n and p are given, so which approximation is good under what situation?
 - $p = 1/100, n = 100$: $np = 1$, very asymmetric X_i , small $p \implies$ **Poisson**
 - $p = 1/3, n = 100$: large, reasonably symmetric p , at least moderate $n \implies$ **Normal**
 - $p = 1/100, n = 10,000$: small p , but large $n \implies$ **Both Poisson and Normal**

L8(5)

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Questions?

1)