

Lecture 4: Random Variable, Part II

Yi, Yung (이용)

EE210: Probability and Introductory Random Processes KAIST EE

MONTH DAY, 2021

Outline



- Continuous Random Variable
- PDF (Probability Density Function)
- CDF (Cumulative Distribution Function)
- Exponential and Normal Distribution
- Joint PDF, Conditional PDF
- · Bayes' rule for continous and even mixed cases

Roadmap



- Famous discrete random variables used in the community
 - Bernoulli, Uniform, Binomial, Geometric, Poisson, etc.
- Summarizing a random variable: Expectation and Variance
- o Functions of a single random variable, Functions of multiple random variables
- Conditioning for random variables, Independence for random variables
- Continuous random variables
 - Normal, Uniform, Exponential, etc.
- Bayes' rule for random variables





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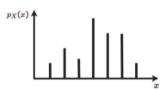
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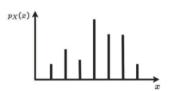
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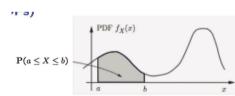
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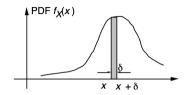


- $\mathbb{P}(a \le X \le b) = \sum_{x:a \le x \le b} p_X(x)$
- $p_X(x) > 0, \sum_{x \in P_X(x)} p_X(x) = 1$



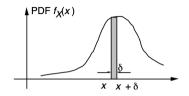
- $\mathbb{P}(a \le X \le b) = \int_a^b f_X(x) dx$ $f_X(x) \ge 0$, $\int_{-\infty}^{\infty} f_X(x) dx = 1$





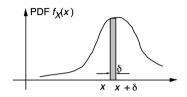
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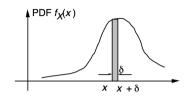




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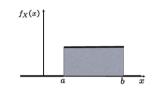
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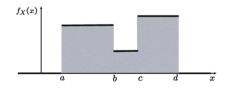




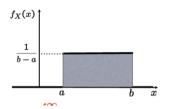
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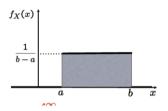




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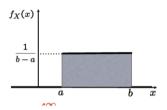




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$$\mathbb{E}[X] = \int_{-\infty}^{\infty} x f_X(x) dx = \int_a^b \frac{x}{b-a} dx = \frac{1}{b-a} \frac{b^2 - a^2}{2} = \frac{b+a}{2}$$

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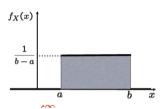




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$$var[X] = \frac{a^2 + ab + b^2}{3} - \frac{a^2 + 2ab + b^2}{4}$$



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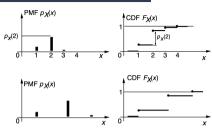


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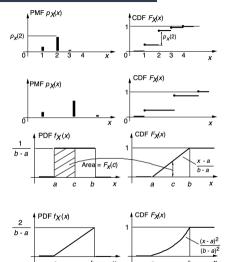


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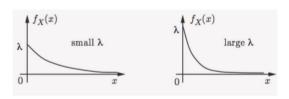
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Now, let's look at famous continuous random variables popularly used in our life.





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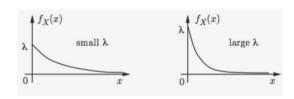




• A rv X is called exponential with λ , if

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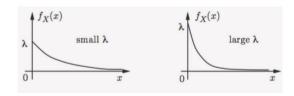
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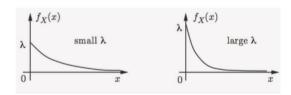
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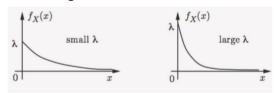
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- (Q) What is the discrete rv which models a waiting time?



Modeling Waiting Time? A Discrete Twin (1)



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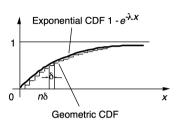
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- *n*-th system: $X_{geo}^n(p_n)$ with CDF $F_{geo}^n(\cdot)$



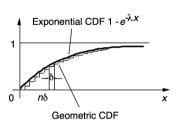
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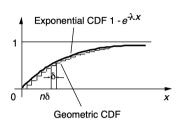


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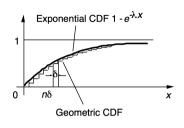
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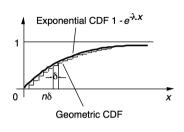
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- As $n \to \infty$, the slot length $\delta \to 0$ thus $p_n \to 0$
- The CDF values of exponential and *n*-th geometric rvs become equal whenever $x = \delta, 2\delta, 3\delta, \ldots$, i.e.,

$$F_{exp}(n\delta) = F_{geo}^n(n), \quad n = 1, 2, \dots$$





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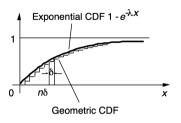
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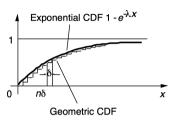
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- As n grows, the number of slots grows, but the success probability over one slot decreases, so that everything is balanced up.
- As *n* grows, $F_{geo}^n(n)$ approaches $F_{exp}(n\delta)$.

Normal (also called Gaussian) Random Variable



Why important?

- Central limit theorem (중심극한정리)
 - One of the most remarkable findings in the probability theory

Convenient analytical properties

· Modeling aggregate noise with many small, independent noise terms

Normal: PDF, Expectation, Variance



• Standard Normal N(0,1)

$$f_X(x) = \frac{1}{\sqrt{2\pi}}e^{-x^2/2}$$

- $\mathbb{E}[X] = 0$
- var[X] = 1

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• Standard Normal N(0,1)

$$f_X(x) = \frac{1}{\sqrt{2\pi}}e^{-x^2/2}$$

• General Normal
$$N(\mu, \sigma^2)$$

$$f_X(x) = \frac{1}{\sigma\sqrt{2\pi}}e^{-(x-\mu)^2/2\sigma^2}$$
• $\mathbb{E}[X] = \mu$
• $\operatorname{var}[X] = \sigma^2$

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Need to check:

- a legitimate PDF or not
- expectation/variance

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• Linear transformation preserves normality

Linear transformation of Normal

If $X \sim \textit{Norm}(\mu, \sigma^2)$, then for $a \neq 0$ and b $Y = aX + b \sim \textit{Norm}(a\mu + b, a^2\sigma^2)$.



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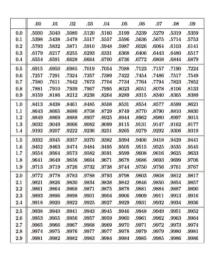
- Thus, every normal rv can be standardized: If $X \sim \textit{Norm}(\mu, \sigma^2)$, then $Y = \frac{\mathsf{X} \mu}{\sigma} \sim \textit{Norm}(0, 1)$
- Thus, we can make the table which records the following CDF values:

$$\Phi(y) = \mathbb{P}(Y \leq y) = \mathbb{P}(Y < y) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{y} e^{-t^2/2} dt$$

Example



 Annual snowfall X is modeled as Norm(60, 20²). What is the probability that this year's snowfall is at least 80 inches?



Example



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- $Y = \frac{X-60}{20}$.

	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
0.0	.5000	.5040	.5080	.5120	.5160	.5199	.5239	.5279	.5319	.5359
0.1	.5398	.5438	.5478	.5517	.5557	.5596	.5636	.5675	.5714	.5753
0.2	.5793	.5832	.5871	.5910	.5948	.5987	.6026	.6064	.6103	.6141
0.3	.6179	.6217	.6255	.6293	.6331	.6368	.6406	.6443	.6480	.6517
0.4	.6554	.6591	.6628	.6664	.6700	.6736	.6772	.6808	.6844	.6879
0.5	.6915	.6950	.6985	.7019	.7054	.7088	.7123	.7157	.7190	.7224
0.6	.7257	.7291	.7324	.7357	.7389	.7422	.7454	.7486	.7517	.7549
0.7	.7580	.7611	.7642	.7673	.7704	.7734	.7764	.7794	.7823	.7852
0.8	.7881	.7910	.7939	.7967	.7995	.8023	.8051	.8078	.8106	.8133
0.9	.8159	.8186	.8212	.8238	.8264	.8289	.8315	.8340	.8365	.8389
1.0	.8413	.8438	.8461	.8485	.8508	.8531	.8554	.8577	.8599	.8621
1.1	.8643	.8665	.8686	.8708	.8729	.8749	.8770	.8790	.8810	.8830
1.2	.8849	.8869	.8888	.8907	.8925	.8944	.8962	.8980	.8997	.9015
1.3	.9032	.9049	.9066	.9082	.9099	.9115	.9131	.9147	.9162	.9177
1.4	.9192	.9207	.9222	.9236	.9251	.9265	.9279	.9292	.9306	.9319
1.5	.9332	.9345	.9357	.9370	.9382	.9394	.9406	.9418	.9429	.9441
1.6	.9452	.9463	.9474	.9484	.9495	.9505	.9515	.9525	.9535	.9545
1.7	.9554	.9564	.9573	.9582	.9591	.9599	.9608	.9616	.9625	.9633
1.8	.9641	.9649	.9656	.9664	.9671	.9678	.9686	.9693	.9699	.9706
1.9	.9713	.9719	.9726	.9732	.9738	.9744	.9750	.9756	.9761	.9767
2.0	.9772	.9778	.9783	.9788	.9793	.9798	.9803	.9808	.9812	.9817
2.1	.9821	.9826	.9830	.9834	.9838	.9842	.9846	.9850	.9854	.9857
2.2	.9861	.9864	.9868	.9871	.9875	.9878	.9881	.9884	.9887	.9890
2.3	.9893	.9896	.9898	.9901	.9904	.9906	.9909	.9911	.9913	.9916
2.4	.9918	.9920	.9922	.9925	.9927	.9929	.9931	.9932	.9934	.9936
2.5	.9938	.9940	.9941	.9943	.9945	.9946	.9948	.9949	.9951	.9952
2.6	.9953	.9955	.9956	.9957	.9959	.9960	.9961	.9962	.9963	.9964
2.7	.9965	.9966	.9967	.9968	.9969	.9970	.9971	.9972	.9973	.9974
2.8	.9974	.9975	.9976	.9977	.9977	.9978	.9979	.9979	.9980	.9981
2.9	.9981	.9982	.9982	.9983	.9984	.9984	.9985	.9985	.9986	.9986

Example^l



- Annual snowfall X is modeled as Norm(60, 20²). What is the probability that this year's snowfall is at least 80 inches?
- $Y = \frac{X-60}{20}$. $\mathbb{P}(X \ge 80) = \mathbb{P}(Y \ge \frac{80-60}{20})$ $= \mathbb{P}(Y \ge 1) = 1 - \Phi(1)$ = 1 - 0.8413 = 0.1587

	.00	.01	.02	.03	.04	.05	.06	.07	.08	.09
0.0	.5000	.5040	.5080	.5120	.5160	.5199	.5239	.5279	.5319	.5359
0.1	.5398	.5438	.5478	.5517	.5557	.5596	.5636	.5675	.5714	.5753
0.2	.5793	.5832	.5871	.5910	.5948	.5987	.6026	.6064	.6103	.6141
0.3	.6179	.6217	.6255	.6293	.6331	.6368	.6406	.6443	.6480	.6517
0.4	.6554	.6591	.6628	.6664	.6700	.6736	.6772	.6808	.6844	.6879
0.5	.6915	.6950	.6985	.7019	.7054	.7088	.7123	.7157	.7190	.7224
0.6	.7257	.7291	.7324	.7357	.7389	.7422	.7454	.7486	.7517	.7549
0.7	.7580	.7611	.7642	.7673	.7704	.7734	.7764	.7794	.7823	.7852
0.8	.7881	.7910	.7939	.7967	.7995	.8023	.8051	.8078	.8106	.8133
0.9	.8159	.8186	.8212	.8238	.8264	.8289	.8315	.8340	.8365	.8389
1.0	.8413	.8438	.8461	.8485	.8508	.8531	.8554	.8577	.8599	.8621
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Roadmap



- Famous discrete random variables used in the community
 - Bernoulli, Uniform, Binomial, Geometric, Poisson, etc.
- Summarizing a random variable: Expectation and Variance
- Functions of a single random variable, Functions of multiple random variables
- Conditioning for random variables, Independence for random variables
- Continuous random variables
 - Normal, Uniform, Exponential, etc.
- Bayes' rule for random variables
- ** Continuous counterparts are intuitively understandable. So, we will be quick at reviewing them.



Jointly Continuous

Two continuous rvs are if a non-negative function $f_{X,Y}(x,y)$

(called joint PDF) satisfies: for $\boxed{\text{every}}$ subset B of the two dimensional plane,

$$\mathbb{P}((X,Y)\in B)=\iint_{(x,y)\in B}f_{X,Y}(x,y)dxdy$$



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1. The joint PDF is used to calculate probabilities

$$\mathbb{P}((X,Y)\in B)=\iint_{(X,Y)\in B}f_{X,Y}(x,y)dxdy$$

Our particular interest: $B = \{(x, y) \mid a \le x \le b, c \le y \le d\}$





2. The marginal PDFs of X and Y are from the joint PDF as:

$$f_X(x) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy, \quad f_Y(y) = \int_{-\infty}^{\infty} f_{X,Y}(x,y) dx$$



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3. The joint CDF is defined by $F_{X,Y}(x,y) = \mathbb{P}(X \le x, Y \le y)$, and determines the joint PDF as:

$$f_{X,Y}(x,y) = \frac{\partial^2 F_{x,y}}{\partial x \partial y}(x,y)$$



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4. A function g(X, Y) of X and Y defines a new random variable, and

$$\mathbb{E}[g(X,Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x,y) f_{X,Y}(x,y) dxdy$$

Continuous: Conditional PDF given an event



* Conditional PDF, given an event

* Conditional PDF, given $X \in B$

Continuous: Conditional PDF given an event



- * Conditional PDF, given an event
- $f_X(x) \cdot \delta \approx \mathbb{P}(x \le X \le x + \delta)$ $f_{X|A}(x) \cdot \delta \approx \mathbb{P}(x \le X \le x + \delta|A)$

* Conditional PDF, given $X \in B$

Continuous: Conditional PDF given an event



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- $\mathbb{P}(X \in B) = \int_B f_X(x) dx$ $\mathbb{P}(X \in B|A) = \int_B f_{X|A}(x) dx$

Note: A is an event, but B is a subset that includes the possible values which can be taken by the rv X.

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•
$$\int f_{X|A}(x) = 1$$

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* Conditional PDF, given $X \in B$

$$\mathbb{P}(x \le X \le x + \delta | X \in B) \approx f_{X|X \in B}(x) \cdot \delta$$

$$f_{X|X\in B}(x) = \begin{cases} 0, & \text{if } x \notin B \\ \frac{f_X(x)}{\mathbb{P}(B)}, & \text{if } x \in B \end{cases}$$

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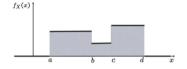
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(Q) In the discrete, we consider the event $\{X = x\}$, not $\{X \in B\}$. Why?



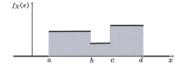
$$A = \left\{ \frac{a+b}{2} \le X \le b \right\}$$

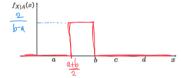






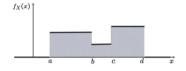
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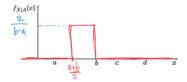






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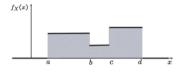


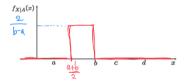
•
$$\mathbb{E}[X] = \int x f_X(x) dx$$

 $\mathbb{E}[X|A] = \int x f_{X|A}(x) dx$



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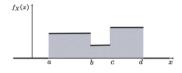
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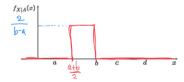
•
$$\mathbb{E}[g(X)] = \int g(x) f_X(x) dx$$

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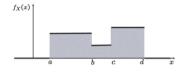
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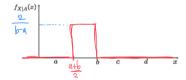
$$\mathbb{E}[X|A] =$$

$$\mathbb{E}[X^2|A] =$$



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•
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$$\mathbb{E}[g(X)] = \int g(x) f_X(x) dx$$

 $\mathbb{E}[g(X)|A] = \int g(x) f_{X|A}(x) dx$

$$\mathbb{E}[X|A] = \int_{(a+b)/2}^{b} x \frac{2}{b-a} dx = \frac{a}{4} + \frac{3b}{4}$$

$$\mathbb{E}[X^{2}|A] = \int_{(a+b)/2}^{b} x^{2} \frac{2}{b-a} dx =$$



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- Thus, expected to be memeoryless.



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$$\mathbb{P}(X>n+m|X>m)=\frac{\mathbb{P}(X>n+m)}{\mathbb{P}(X>m)}=\frac{e^{-\lambda(n+m)}}{e^{-\lambda m}}=e^{-\lambda n}=\mathbb{P}(X>n)$$



Partition of Ω into A_1, A_2, A_3, \dots

* Discrete case

* Continuous case



Partition of Ω into A_1, A_2, A_3, \dots

* Discrete case

Total Probability Theorem

$$p_X(x) = \sum_i \mathbb{P}(A_i)\mathbb{P}(X = x|A_i)$$
$$= \sum_i \mathbb{P}(A_i)p_{X|A_i}(x)$$

Total Expectation Theorem

$$\mathbb{E}[X] = \sum_{i} \mathbb{P}(A_i) \mathbb{E}[X|A_i]$$

* Continuous case



Partition of Ω into A_1, A_2, A_3, \dots

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Total Probability Theorem

$$\rho_X(x) = \sum_i \mathbb{P}(A_i) \mathbb{P}(X = x | A_i)$$

$$= \sum_i \mathbb{P}(A_i) \rho_{X|A_i}(x)$$

Total Expectation Theorem

$$\mathbb{E}[X] = \sum_{i} \mathbb{P}(A_i) \mathbb{E}[X|A_i]$$

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Total Probability Theorem

$$f_X(x) = \sum_i \mathbb{P}(A_i) f_{X|A_i}(x)$$



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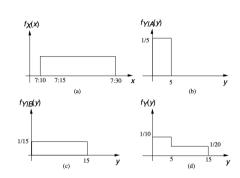
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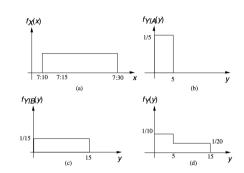


- Your train's arrival every quarter hour (0, 15min, 30min, 45min).
- Your arrival \sim uniform(7:10, 7:30) am.



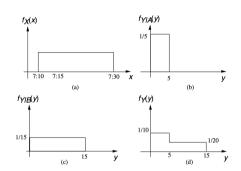


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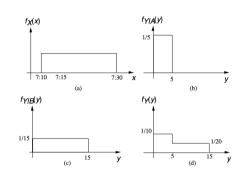




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$$A = \{7:10 \le X \le 7:15\}$$

$$B = \{7:15 < X < 7:30\}$$

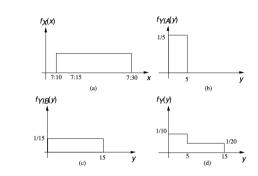




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$$f_Y(y) = \mathbb{P}(A)f_{Y|A}(y) + \mathbb{P}(B)f_{Y|B}(y)$$
 for $0 \le y \le 5$

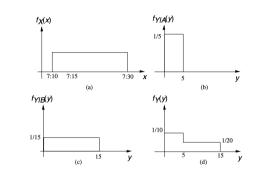
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$$f_{Y}(y) = \mathbb{P}(A)f_{Y|A}(y) + \mathbb{P}(B)f_{Y|B}(y)$$

$$f_{Y}(y) = \frac{1}{4}\frac{1}{5} + \frac{3}{4}\frac{1}{15} = \frac{1}{10}, \text{ for } 0 \le y \le 5$$

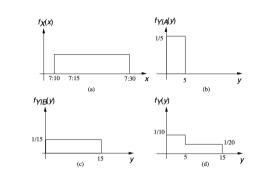
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- X : your arrival time, Y : waiting time.
- The value of X makes a different waiting time. So, consider two events:

$$A = \{7:10 \le X \le 7:15\}$$

$$B = \{7:15 \le X \le 7:30\}$$



$$f_Y(y) = \mathbb{P}(A)f_{Y|A}(y) + \mathbb{P}(B)f_{Y|B}(y)$$

$$f_Y(y) = \frac{1}{4}\frac{1}{5} + \frac{3}{4}\frac{1}{15} = \frac{1}{10}, \quad \text{for } 0 \le y \le 5$$

$$f_Y(y) = \frac{1}{4}0 + \frac{3}{4}\frac{1}{15} = \frac{1}{20}, \quad \text{for } 5 < y \le 15$$



•
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$$f_{X,Y}(x,y) = f_Y(y) \cdot f_{X|Y}(x|y)$$

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• Total prob./exp. theorem.

$$f_X(x) = \int_{-\infty}^{\infty} f_Y(y) f_{X|Y}(x|y) dy$$

$$\mathbb{E}[X|Y = y] = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$$

$$\mathbb{E}[X] = \int_{-\infty}^{\infty} f_Y(y) \mathbb{E}[X|Y = y] dy$$



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• Independence.

$$f_{X,Y}(x,y) = f_X(x)f_Y(y)$$
, for all x and y



- Break a stick of length / twice
 - first break at $X \sim uniform[0.l]$
 - second break at $Y \sim \textit{uniform}[0, X]$



- Break a stick of length / twice
 - first break at $X \sim uniform[0.1]$
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Using the TET,

$$\mathbb{E}[Y] = \int_0^I \frac{1}{I} \mathbb{E}[Y|X = x] dx$$
$$= \int_0^I \frac{1}{I} \frac{x}{2} dx = \frac{I}{4}$$

Example: Stick-breaking (Ch 3. Prob 21)



- Break a stick of length / twice
 - first break at $X \sim uniform[0.1]$
 - second break at $Y \sim uniform[0, X]$
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$$= \int_0^I \frac{1}{I} \frac{x}{2} dx = \frac{I}{4}$$

 f_X(x) and f_{Y|X}(y|x) seems easy to compute. Thus,

$$f_{X,Y}(x,y) = f_X(x)f_{Y|X}(y|x) = \frac{1}{l} \cdot \frac{1}{x}$$

You can do many other things with the joint PDF.

Roadmap



- Famous discrete random variables used in the community
 - Bernoulli, Uniform, Binomial, Geometric, Poisson, etc.
- Summarizing a random variable: Expectation and Variance
- o Functions of a single random variable, Functions of multiple random variables
- Conditioning for random variables, Independence for random variables
- Continuous random variables
 - Normal, Uniform, Exponential, etc.
- Bayes' rule for random variables

Bayes Rule for Continuous



- X: state/cause/original value $\rightarrow Y$: result/resulting action/noisy measurement
- Model: $\mathbb{P}(X)$ (prior) and $\mathbb{P}(Y|X)$ (cause o result)
- Inference: $\mathbb{P}(X|Y)$?

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• Inference of K given Y

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K: discrete, Y: continuous

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Inference of discrete K given continuous Y:

$$p_{K|Y}(k|y) = \frac{p_K(k)f_{Y|K}(y|k)}{f_Y(y)}, \quad f_Y(y) = \sum_{k'} p_K(k')f_{Y|K}(y|k')$$

• K: -1, +1, original signal, equally likely. $p_K(1) = 1/2, p_K(-1) = 1/2$.



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- Your received signal = 0.7. What's your guess about the original signal?



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- Your received signal = 0.7. What's your guess about the original signal?
- Your received signal = -0.2. What's your guess about the original signal?



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- Your received signal = 0.7. What's your guess about the original signal? +1
- Your received signal = -0.2. What's your guess about the original signal? -1



• $Y|K=1 \sim N(1,1)$ and $Y|K=-1 \sim N(-1,1)$.



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$$f_{Y|K}(y|k) = \frac{1}{\sqrt{2\pi}}e^{-\frac{1}{2}(y-k)^2}, \quad k = 1, -1$$



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• Probability that K = 1, given Y = y? After some algebra,

$$ho_{K|Y}(1|y) = rac{1}{1 + e^{-2y}}$$

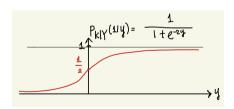


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Questions?

Review Questions



- 1) What is PDF and CDF?
- 2) Why do we need CDF?
- 3) What are joint/marginal/conditional PDFs?
- Explain memorylessness of exponential random variables.
- 5) Explain the version of Bayes' rule for continuous and mixed random variables.