

#### Lecture 7: Random Processes, Part I

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EE210: Probability and Introductory Random Processes KAIST EE

November 21, 2022

#### Roadmap



- (1) Introduction of Random Processes
- (2) Bernoulli Processes
- (3) Poisson Processes: Poisson RV and Basic Idea
- (4) Poisson Processes: Definition and Properties
- (5) Playing with Bernoulli and Poisson Processes
- (6) Random Incidence Phenomenon

#### Roadmap

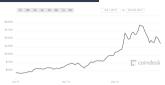


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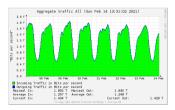
#### Things that evolve in time

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• Many probabilistic experiments that evolve in time



(a) Prices of a crytocurrency



(b) Internet traffic traces

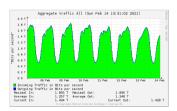
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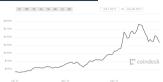


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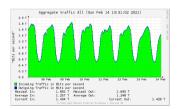
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- Random process is a mathematical model for it.



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- The values that  $X_t$  (or X(t)) can take: discrete or continuous



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  - $X(3.7, \omega_1) = 3409, X(2, \omega_2) = 5000, X(7.8, \omega_3) = 2800, \text{ etc.}$





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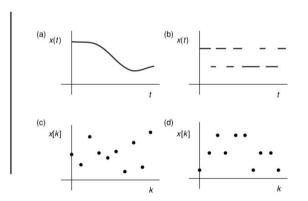
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  - Other interesting questions, depending on the target random process

#### 4 Types of Random Processes



- Types of time and value

- (a) continuous time, continuous value
- (b) continuous time, discrete value
- (c) discrete time, continuous value
- (d) discrete time, discrete value



#### Random Processes in This Course



- The simplest RP
- discrete time

Jacob Bernoulli (1654 - 1705), Swiss



Simeon Denis Poisson (1781 - 1840), France



Andrey Markov (1856 - 1922), Russia





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- $[X(s)]_{s=0}^t \perp [X(s)]_{s=t}^{t+a}$

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- Markov Chain (MC)



KAIST EE

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- Bernoulli Process (BP)
- "today" independent of "past"

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- "today" depends only on "yesterday"



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L8(2)



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# 0011000011011100

- Discrete time, discrete value
- One of the simplest random processes
- A type of "arrival" process

## Bernoulli Process: Questions



Please pause the video and write down the questions that you want to ask about Bernoulli process.

VIDEO PAUSE

Q1.

Q2.

**Q3**.

Q4.

**Q5**.





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- (Q2) # of slots  $T_1$  until the first arrival?
- *T*<sub>1</sub> ∼ Geom(*p*)
- $\mathbb{E}(T_1) = 1/p$ ,  $\text{var}(T_1) = \frac{1-p}{p^2}$



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- *T*<sub>1</sub> is geometric? Memoryless
- Conditioned no first arrival at some time, what's the distribution of the remaining time until the first arrival?
- Still, geometric.



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- $S_n = X_1 + X_2 + \cdots + X_n$
- $S_n \sim \text{Binomial}(n, p)$
- $\mathbb{E}(S_n) = np$ ,  $var(S_n) = np(1-p)$
- This will hold for any n consecutive slots.

(Q2) # of slots  $T_1$  until the first arrival?

- *T*<sub>1</sub> ∼ Geom(*p*)
- $\mathbb{E}(T_1) = 1/p$ ,  $\text{var}(T_1) = \frac{1-p}{p^2}$

- *T*<sub>1</sub> is geometric? Memoryless
- Conditioned no first arrival at some time, what's the distribution of the remaining time until the first arrival?
- Still, geometric.
- But, more than that, as we will see.
   Independence across time slots leads to many useful properties, allowing the quick solution of many problems.



Independence across slots  $\implies$  the fresh-start anytime when I look at the process?

(Q3) 
$$U = X_1 + X_2 \perp \!\!\! \perp V = X_5 + X_6$$
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- If you watch the on-going Bernoulli process(p) from some time n, you still see the same Bernoulli process(p).



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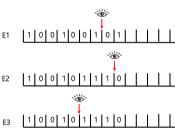
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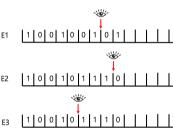




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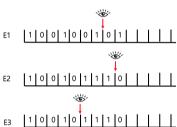
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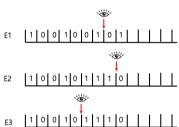
L8(2)

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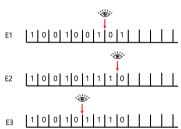


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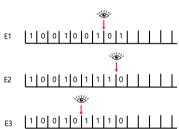
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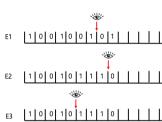
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- Do we experience the fresh-start for any N? E1, E2, and E3?

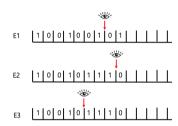


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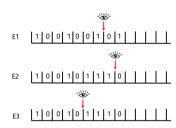
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**E1.** When I watch the process, N has been already determined. Yes



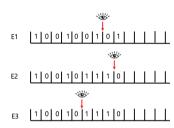
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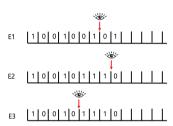
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- The question of N = n? can be answered just from the knowledge about  $X_1, X_2, ..., X_n$ ? Then, Yes! (see pp. 301 for more formal description)



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VIDEO PAUSE

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- Yes. Time when 10 consecutive arrivals have been observed
- No. Time of 2nd arrival in 10 consecutive arrivals



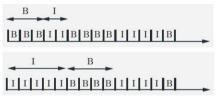
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• Regard an arrival as a server being busy (just for our easy understanding)



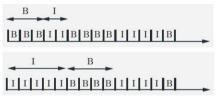
18 / 65

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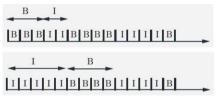


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L8(2)



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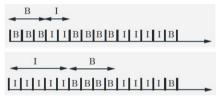


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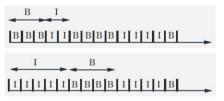
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L8(2)

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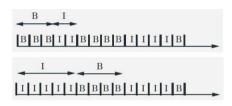


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- $B_1$  is geometric with parameter (1 p)



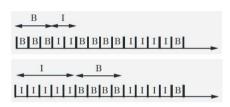


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L8(2)



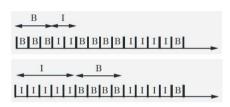
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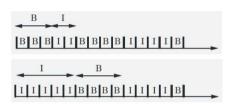




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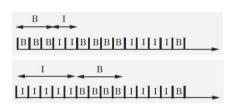




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L8(2)





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- $B_3, B_4, \dots$ ?



• Time of the first arrival  $Y_1 \sim \mathsf{Geom}(p)$ 



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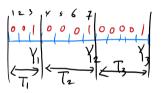
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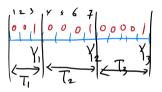


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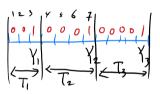


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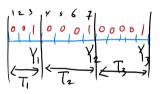


Time of the first arrival Y₁ ~ Geom(p)
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#### VIDEO PALISE



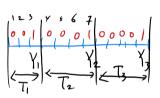


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- After each  $T_k$ , the fresh-start occurs.
- $\{T_i\}$  are i.i.d. and  $\sim \text{Geom}(p)$
- We know  $Y_k$ 's expectation and variance:  $\mathbb{E}[Y_k] = \frac{k}{p}$ ,  $\text{var}[Y_k] = \frac{k(1-p)}{p^2}$ , but its distribution?

# PMF of $Y_k$



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### PMF of $Y_k$



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$$\begin{split} \mathbb{P}(Y_k = t) &= \mathbb{P}\left(X_t = 1 \text{ and } k-1 \text{ arrivals during the first } t-1 \text{ slots}\right) \\ &= \mathbb{P}\left(X_t = 1\right) \cdot \mathbb{P}\left(k-1 \text{ arrivals during the first } t-1 \text{ slots}\right) \\ &= p \times \binom{t-1}{k-1} p^{k-1} (1-p)^{t-k} = \binom{t-1}{k-1} p^k (1-p)^{t-k}, \quad t = k, k+1, \dots \end{split}$$

L8(2)

### Pascal Random Variable with Parameter (k, p)



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• Pascal(1, p) = Geom(p)

#### Roadmap



- (1) Introduction of Random Processes
- (2) Bernoulli Processes
- (3) Poisson Processes: Poisson RV and Basic Idea
- (4) Poisson Processes: Definition and Properties
- (5) Playing with Bernoulli and Poisson Processes



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• A random variable  $S \sim \text{Bin}(n, p)$ : Models the number of successes in a given number n of independent trials with success probability p.

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L8(3)

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$$p_{\mathcal{S}}(k) = \frac{n(n-1)\cdots(n-k+1)}{k!} \cdot \frac{\lambda^k}{n^k} \cdot \left(1 - \frac{\lambda}{n}\right)^{n-k}$$

L8(3)

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• Our interest: very large n and very small p, such that  $np = \lambda$ , i.e.,  $\lim_{n \to \infty} p_S(k)$ ?

$$p_{S}(k) = \frac{n(n-1)\cdots(n-k+1)}{k!} \cdot \frac{\lambda^{k}}{n^{k}} \cdot \left(1 - \frac{\lambda}{n}\right)^{n-k}$$
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L8(3)



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L4(3)

27 / 65

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- Need a "modeling sense" to make this possible. It's a good practice for engineers!
- VIDEO PAUSE



Continuous twin



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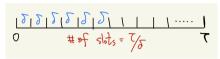
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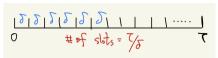
• What's the limit as  $\delta \to 0$  (equivalently,  $n \to \infty$ )





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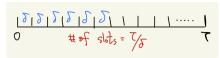




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- $o(\delta)$ : some function that goes to zero faster than  $\delta$ .
  - Thus, for very small  $\delta$ ,  $o(\delta)$  becomes negligible, compared to  $\delta$ .
  - Example:  $o(\delta) = \delta^{\alpha}$ , where any  $\alpha > 1$



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L8(3)

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- This is a continuous twin process of Bernous process, which we call Poisson process.

L8(3)

#### Roadmap



- (1) Introduction of Random Processes
- (2) Bernoulli Processes
- (3) Poisson Processes: Poisson RV and Basic Idea
- (4) Poisson Processes: Definition and Properties
- (5) Playing with Bernoulli and Poisson Processes
- (6) Random Incidence Phenomenon



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| An arriva | l process is | s called a | Poisson | process | with ra | te $\lambda$ , if $^{\dagger}$ | he follo | wing a | re satisfied |
|-----------|--------------|------------|---------|---------|---------|--------------------------------|----------|--------|--------------|
|           |              |            |         |         |         |                                |          |        |              |
|           |              |            |         |         |         |                                |          |        |              |
|           |              |            |         |         |         |                                |          |        |              |
|           |              |            |         |         |         |                                |          |        |              |
|           |              |            |         |         |         |                                |          |        |              |
|           |              |            |         |         |         |                                |          |        |              |



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An arrival process is called a Poisson process with rate  $\lambda$ , if the following are satisfied:

- (Independence) Let  $N_{\tau}$  be the number of arrivals over the interval  $[0, \tau]$ . For any  $\tau_1, \tau_2 > 0, N_{s+\tau_1} - N_s$  is independent of  $N_{t+\tau_2} - N_t$ , if  $t > s + \tau_1$ .
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- (Time homogeneity) For any s, the distribution of  $N_{s+\tau} N_s$  is equal to that of  $N_{\tau}$ .
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  - $\circ N_{\tau}$  becomes the number of arrivals over any interval of length  $\tau$ .
- (Small interval probability) Let  $\mathbb{P}(k,\tau) = \mathbb{P}(N_{\tau} = k)$ , which satisfy:

$$\mathbb{P}(0,\tau) = 1 - \lambda \tau + o(\tau)$$

$$\mathbb{P}(1,\tau) = \lambda \tau + o_1(\tau)$$

$$\mathbb{P}(k,\tau) = o_k(\tau) \quad \text{for } k = 2,3,\ldots, \quad \text{where} \quad \lim_{\tau \to 0} \frac{o(\tau)}{\tau} = 0, \quad \lim_{\tau \to 0} \frac{o_k(\tau)}{\tau} = 0$$

L8(4)



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- (Distribution of  $N_{\tau}$ )  $N_{\tau}$  is the Poisson rv with parameter  $\lambda \tau$ , i.e., if we let  $\mathbb{P}(k,\tau) = \mathbb{P}(N_{\tau} = k)$ , we have:

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L8(4)

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### Poisson Process: $\mathbb{P}(k,\tau)$ , $N_{\tau}$ , and T



## Poisson Process: $\mathbb{P}(k, au),\ extstyle{N}_{ au},\ extstyle{and}\ \ \overline{T}$



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- $T \sim \text{Exp}(\lambda)$ . Thus,  $\mathbb{E}(T) = 1/\lambda$  and  $\text{var}(T) = 1/\lambda^2$ 
  - Continuous twin of geometric rv in Bernoulli process
  - Memoryless



- Receive emails according to a Poisson process at rate  $\lambda=5$  messages/hour
- Mean and variance of mails received during a day

P[one new message in the next hour]

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• k-th inter-arrival time  $T_k = Y_k - Y_{k-1}, k \ge 2$ , and  $T_1 = Y_1$ .



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- Fresh-start at determinsitic time *t*: Start watching at time *t*, then you see the Poisson process, independent of what has happened in the past.
- Fresh-start at random time T: Similarly, it holds.
  - For example, when you start watching at random time  $T_1$  (time of first arrival).
  - Generally, it holds when T is a stopping time.

(Q3) The k-th arrival time  $Y_k$ ?

- k-th inter-arrival time  $T_k = Y_k Y_{k-1}$ ,  $k \ge 2$ , and  $T_1 = Y_1$ .
- $Y_k = T_1 + T_2 + \cdots + T_k$  is sum of i.i.d. exponential rvs.

#### Memoryless and Fresh-start Property



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- $Y_k = T_1 + T_2 + \cdots + T_k$  is sum of i.i.d. exponential rvs.
- $\mathbb{E}[Y_k] = k/\lambda$  and  $\text{var}[Y_k] = k/\lambda^2$ , but what is the distribution of  $Y_k$ ?



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• For a given  $\delta$ , | : prob. of k-th arrival over  $[y, y + \delta]$ .



• For a given  $\delta$ ,  $\delta \cdot f_{Y_k}(y)$ : prob. of k-th arrival over  $[y, y + \delta]$ .



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This is called Erlang rv.

An Erlang random variable Z with parameter  $(k, \lambda)$  has the following pdf:

$$f_Z(z) = \frac{\lambda^k z^{k-1} e^{-\lambda z}}{(k-1)!}, \quad z \ge 0$$



$$- n = \tau/\delta, \ p = \lambda\delta, \ np = \lambda\tau$$

$$0 \qquad \text{# of slots} = \sqrt{\delta}$$

|                         | Bernoulli process | Poisson process |
|-------------------------|-------------------|-----------------|
| time of arrival         | Discrete          | Continuous      |
| PMF of $\#$ of arrivals |                   |                 |
| Interarrival time       |                   |                 |
| Time of $k$ -th arrival |                   |                 |
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- 
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(Q4) 
$$\mathbb{E}[\text{future fi. time}|\text{already fished for 3h}]$$
  
Fresh-start. So,  
 $\mathbb{E}[\exp(\lambda)] = 1/\lambda = 1/0.6$   
(Q5)  $\mathbb{E}[\text{F=total fishing time}]$   
 $2 + \mathbb{E}[F - 2] = 2 + \mathbb{P}(F = 2) \cdot 0 + \mathbb{P}(F > 2) \cdot \mathbb{E}[F - 2|F > 2]$   
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L8(4)

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(Q6)  $\mathbb{E}[\text{number of fish}] = \lambda \cdot 2 + \mathbb{P}(0,2) \cdot 1$ 

November 21, 2022

#### Roadmap



- (1) Introduction of Random Processes
- (2) Bernoulli Processes
- (3) Poisson Processes: Poisson RV and Basic Idea
- (4) Poisson Processes: Definition and Properties
- (5) Playing with Bernoulli and Poisson Processes
- (6) Random Incidence Phenomenon

#### Description via Inter-arrival Times



#### Alternative Description of the Bernoulli Process

- 1. Start with a sequence of independent geometric random variables  $T_1$ ,  $T_2$ ,..., with common parameter p, and let these stand for the interarrival times.
- 2. Record a success (or arrival) at times  $T_1$ ,  $T_1 + T_2$ ,  $T_1 + T_2 + T_3$ , etc.

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 Geom(p), independent of the past.



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- Thus, the answer is  $p^2$ .

#### Coding of Random Arrivals



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- Question. How to make software codes of Bernoulli process with p and Poisson process with  $\lambda$
- Inter-arrival times are very handy.
- Bernoulli process with p: Obtain a sequence of random values following the geometric distribution with parameter p.
- Poisson process with  $\lambda$ : Obtain a sequence of random values following the exponential distribution with parameter  $\lambda$ .

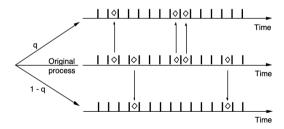
#### Notations In the Rest of These Slides



- Bernoulli random variable: Bern(p)
- Bernoulli process: BP(p)
- Poisson random variable: Poisson( $\lambda$ )
- Poisson process:  $PP(\lambda)$



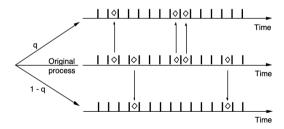
• Split BP(p) into two processes. Whenever there is an arrival, keep it w.p. q and discard it w.p. (1-q).



L8(5)

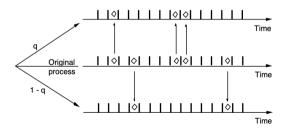


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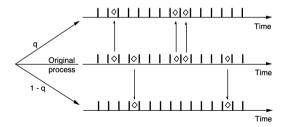
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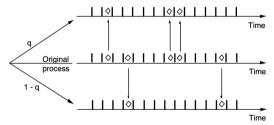


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- BP(pq) and BP(p(1-q)). Why?





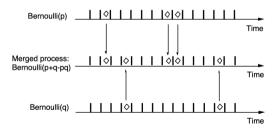
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- $\mathsf{BP}(pq)$  and  $\mathsf{BP}(p(1-q))$ . Why?
- Are they independent? No.



L8(5)

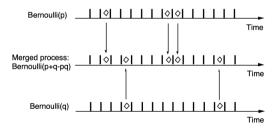


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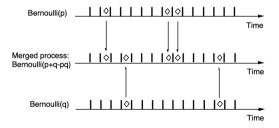


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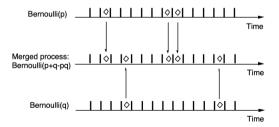


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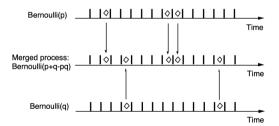


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- $\mathbb{P}(\text{arrival from proc. 1} \mid \text{arrival in the merged proc.}) = \frac{p}{p+q-pq}$





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• Split a Poisson process  $PP(\lambda)$  into two processes by keeping each arrival w.p. p and discarding it w.p. (1-p)



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- Question. What are the split processes?
- Let's focus on the process that we keep
- Independence and time-homogeneity? Yes



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• 
$$\mathbb{P}(1 \text{ arrival}) = p\lambda\delta + p \cdot o(\delta) = p\lambda\delta + o(\delta)$$

• 
$$\mathbb{P}(2 \text{ arrivals}) = p \cdot o(\delta) = o(\delta)$$

• 
$$\mathbb{P}(0 \text{ arrival}) = 1 - p\lambda\delta - p \cdot o(\delta) - p \cdot o(\delta) = 1 - p\lambda\delta + o(\delta)$$

•  $PP(\lambda p)$  and  $PP(\lambda(1-p))$ 

L8(5)



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$$\mathbb{P}(0 \text{ arrival}) \approx (1 - \lambda_1 \delta)(1 - \lambda_2 \delta) \approx 1 - (\lambda_1 + \lambda_2)\delta$$

$$\mathbb{P}(1 \text{ arrival}) \approx (\lambda_1 \delta)(1 - \lambda_2 \delta) + \lambda_2 \delta(1 - \lambda_1 \delta) \approx (\lambda_1 + \lambda_2)\delta$$

$$\mathbb{P}(1 ext{ arrival}) pprox (\lambda_1 \delta) (1 - \lambda_2 \delta) + \lambda_2 \delta (1 - \lambda_1 \delta) pprox (\lambda_1 + \lambda_2) \delta$$

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- Small interval probabilty over  $\delta$ -interval (ignoring  $o(\delta)$  for small  $\delta$ )  $\mathbb{P}(0 \text{ arrival}) \approx (1 \lambda_1 \delta)(1 \lambda_2 \delta) \approx 1 (\lambda_1 + \lambda_2)\delta$   $\mathbb{P}(1 \text{ arrival}) \approx (\lambda_1 \delta)(1 \lambda_2 \delta) + \lambda_2 \delta(1 \lambda_1 \delta) \approx (\lambda_1 + \lambda_2)\delta$
- Merged process:  $PP(\lambda_1 + \lambda_2)$



• Red:  $PP(\lambda_1)$  and Blue:  $PP(\lambda_2)$ 



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# Using Poisson Processes for Intuitive Problem Solving



- 1. Competing exponentials
- 2. Sum of independent Poisson rvs
- 3. Poisson arrivals during and Exponential interval

November 21, 2022



• Two independent light bulbs have life times  $T_a \sim \text{Exp}(\lambda_a)$  and  $T_b \sim \text{Exp}(\lambda_b)$ .



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  - $PP(3\lambda) \xrightarrow{1st \text{ burn out}} PP(2\lambda) \xrightarrow{2st \text{ burn out}} PP(\lambda)$
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$$\mathbb{E}\Big[T_1+T_2+T_3\Big]=\frac{1}{3\lambda}+\frac{1}{2\lambda}+\frac{1}{\lambda}$$



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- Thus,  $X + Y \sim \mathsf{Poisson}(\mu + \nu)$



- Problem 24, pp. 335
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$$\mathbb{P}(N_T = k) =$$



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Very tedious and not very intuitive.



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- Now, consider the merged process of  $PP(\lambda)$  and  $PP(\nu)$ .
  - $\quad \circ \ \mathbb{P}\Big[\mathsf{from} \ \mathsf{PP}(\lambda)|\mathsf{arrival}\Big] = \tfrac{\lambda}{\lambda + \nu} \ \mathsf{and} \ \mathbb{P}\Big[\mathsf{from} \ \mathsf{PP}(\nu)|\mathsf{arrival}\Big] = \tfrac{\nu}{\lambda + \nu}$



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$$p_L(I) = p_K(I+1) = \left(\frac{\nu}{\lambda+\nu}\right) \left(\frac{\lambda}{\lambda+\nu}\right)^I, \quad I = 0, 1, \dots$$



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- In practice, an actual numbers of n and p are given, so which approximation is good under what situation?
  - p = 1/100, p = 100: p = 1, very asymmetric  $X_i$ , small  $p \implies Poisson$
  - p = 1/3, n = 100: large, reasonly symmetric p, at least moderate  $n \implies Normal$
  - p = 1/100, n = 10,000: small p, but large  $n \implies Both Poisson and Normal$

#### Roadmap



- (1) Introduction of Random Processes
- (2) Bernoulli Processes
- (3) Poisson Processes: Poisson RV and Basic Idea
- (4) Poisson Processes: Definition and Properties
- (5) Playing with Bernoulli and Poisson Processes
- (6) Random Incidence Phenomenon

#### Example: Survey of Utilization of Town Buses



What we want to survey: How available are town buses in a city?

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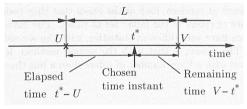


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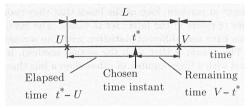
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- Fix a time instant  $t^*$ , and consider the length L of the inter-arrival interval that constains  $t^*$ .





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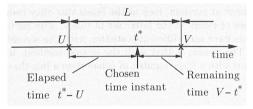
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• Practical context: Yung shows up at the bus station at some arbitrary time  $t^*$  and records the time from the previous bus arrival (U) until the next bus arrival (V)



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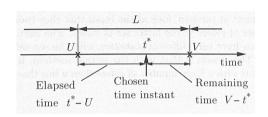


- Practical context: Yung shows up at the bus station at some arbitrary time  $t^*$  and records the time from the previous bus arrival (U) until the next bus arrival (V)
- Question. What is the distribution of L?

VIDEO PAUSE



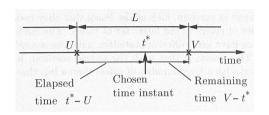
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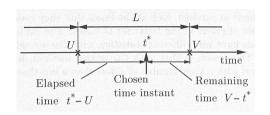


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•  $t^*$  is not random, so "random incidence" may be confusing and misleading.

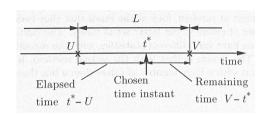




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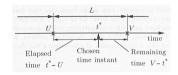
- t\* is not random, so "random incidence" may be confusing and misleading.
- Assumption. For simplicity,  $t^*$  is large enough that we must have an arrival before  $t^*$  (U>0)
- One might superficially argue that  $L \sim \text{Exp}(\lambda)$ , but it is NOT.



$$L = (t^{\star} - U) + (V - t^{\star})$$

•  $V-t^*$ :

•  $t^{\star} - U$ :

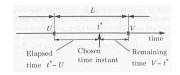




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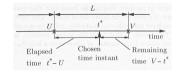


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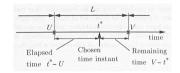
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  - Thus,  $V t^* \sim \mathsf{Exp}(\lambda)$
- $t^* U$ : If we run the PP( $\lambda$ ) backwards in time, it remains Poisson. Why?





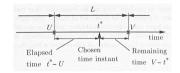
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•  $t^* - U$ : If we run the PP( $\lambda$ ) backwards in time, it remains Poisson. Why? More formally,

$$\mathbb{P}(t^* - U > x) = \mathbb{P}(\text{no arrivals over } [t^* - x, t^*])$$
  
=  $e^{-\lambda x} = \mathbb{P}(T_{\text{inter}} > x)$ 





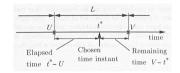
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Thus.  $t^* - U \sim \text{Exp}(\lambda)$ 





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•  $L = X_1 + X_2$ , where  $X_1, X_2 \sim \mathsf{Exp}(\lambda)$ 



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- Time until we have two arrivals in  $PP(\lambda)$



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- Time until we have two arrivals in  $PP(\lambda)$
- Erlang random variable with parameter  $(2, \lambda)$ , i.e.,

$$f_L(I) = \lambda^2 \cdot I \cdot e^{-\lambda I}, \quad I \ge 0$$

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$$f_L(I) = \lambda^2 \cdot I \cdot e^{-\lambda I}, \quad I \ge 0$$

- Mean =  $2/\lambda$
- Why not  $Exp(\lambda)$ ? An observer who arrives at an arbitrary time is more likely to fall in a large rather than a small interarrival interval.

#### Back to Survey of Utilization of Town Buses



- Two Approaches
  - M1. (1) select a few buses at random, and (2) calculate the average number of riders in the selected bus
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- (i) M1 = M2? (ii) M1 > M2? (iii) M1 < M2?
- Answer: M1 < M2
- More likely to select a bus with a large number of riders than a bus that is near-empty.



# Questions?

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#### **Review Questions**



- 1) Explain what is random process. Why do we need such a concept?
- 2) Explain Bernoulli processes. When do we use Bernoulli processes?
- 3) Explain Poisson processes. When do we use Poisson processes?
- 4) What are the relations between those two processces? What features do they share?
- 5) In both processces, ho do we compute (i) number of arrivals over a given interval of time, (ii) time until the first arrival, (iii) time until k-th arrival?
- 6) What is Pascal rvs and Erlang rvs?
- 7) What is the "stopping time" and how is it related to fresh-restart?
- 8) See many examples on how merging and splitting of BP and PP can be used for intuitive soloving of many problems.

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