

Lecture 3: Random Variable, Part I

Yi, Yung (이용)

EE210: Probability and Introductory Random Processes KAIST EE

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Roadmap



- (1) Random variable: Idea and formal definition
- (2) Popular discrete random variables
- (3) Summarizing random variables: Expectation and Variance
- (4) (Functions of) multiple random variables
- (5) Conditioning for random variables
- (6) Independence for random variables

Roadmap



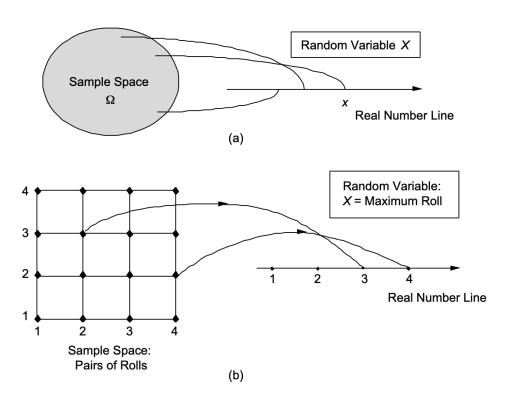
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L3(1) May 13, 2021 3 / 45

Random Variable: Idea



- In reality, many outcomes are numerical , e.g., stock price.
- Even if not, very convenient if we map numerical values to random outcomes, e.g., '0' for male and '1' for female.



(b) Two rolls of tetrahedral dice

L3(1) May 13, 2021 4 / 45

Random Variable: More Formally



- Mathematically, a random variable X is a function which maps from Ω to \mathbb{R} .
- Notation. Random variable X, numerical value x.
- Different random variables can be defined on the same sample space.
- For a fixed value x, we can associate an event that a random variable X has the value x, i.e., $\{\omega \in \Omega \mid X(\omega) = x\}$
- Assume that values x are discrete¹ such as $1, 2, 3, \ldots$. For notational convenience,

$$p_X(x) \triangleq \mathbb{P}(X = x) \triangleq \mathbb{P}\Big(\{\omega \in \Omega \mid X(\omega) = x\}\Big)$$

• For a discrete random variable X, we call $p_X(x)$ probability mass function (PMF).

5 / 45

¹Finite or countably infinite.

Example



- Rolls a dice, $\Omega = \{1, 2, 3, 4, 5, 6\}$
- Define a random variable X=1 for even numbers and X=0 for odd numbers
- Event $A_1 = \{ \omega \in \Omega \mid X(\omega) = 1 \} = \{ 2, 4, 6 \} \subset \Omega$, but simply $A_1 = \{ X = 1 \}$
- Event $A_0 = \{ \omega \in \Omega \mid X(\omega) = 0 \} = \{1, 3, 5\} \subset \Omega$, but simply $A_0 = \{X = 0\}$

• Remember that the random variable X is a function from Ω to $\mathbb R$

L3(1) May 13, 2021 6 / 45

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L3(2) May 13, 2021 7 / 45

Bernoulli X with parameter $p \in [0, 1]$



Only binary values

$$X = egin{cases} 0, & ext{w.p.} & 1-p, \ 1, & ext{w.p.} & p \end{cases}$$

In other words, $p_X(0) = 1 - p$ and $p_X(1) = p$ from our PMF notation.

- Models a trial that results in binary results, e.g., success/failure, head/tail
- Very useful for an indicator rv of an event A. Define a rv $\mathbf{1}_A$ as:

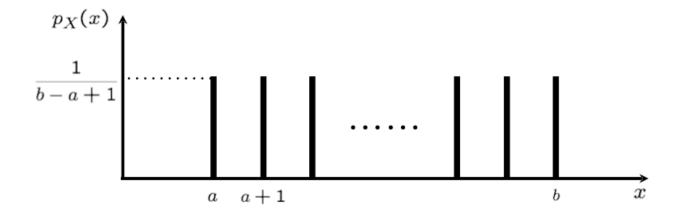
$$\mathbf{1}_A = egin{cases} 1, & ext{if } A ext{ occurs}, \ 0, & ext{otherwise} \end{cases}$$

¹w.p.: with probability
^{L3(2)}

Uniform X with parameter a, b



- integers a, b, where $a \le b$
- Choose a number out of $\Omega = \{a, a+1, \ldots, b\}$ uniformly at random.
- $p_X(i) = \frac{1}{b-a+1}, i \in \Omega$



• Models complete ignorance (I don't know anything about X)

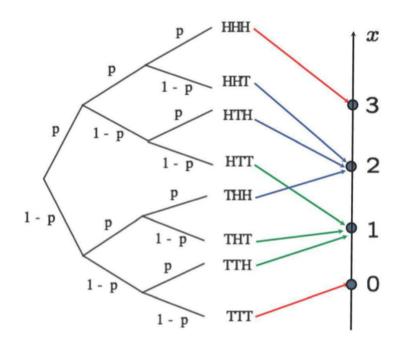
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Binomial X with parameter n, p



- Models the number of successes in a given number of independent trials
- *n* independent trials, where one trial has the success probability *p*.

$$p_X(k) = \binom{n}{k} p^k (1-p)^{n-k}$$



 $[\]binom{1}{k} = \frac{n!}{k!(n-k)!}$, which we read 'n choose k'.

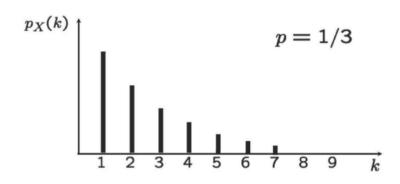
Geometric X with parameter p



- Infinitely many independent Bernoulli trials, where each trial has success probability p
- Random variable: number of trials until the first success.

$$p_X(k) = (1-p)^{k-1}p$$

 Models waiting times until something happens.



L3(2) May 13, 2021 11 / 45

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L3(3) May 13, 2021 12 / 45

Expectation/Mean



Average

Definition

$$\mathbb{E}[X] = \sum_{x} x p_X(x)$$

- $p_X(x)$: relative frequency of value x (trials with x/total trials)
- Example. Bernoulli rv with p

$$\mathbb{E}[X] = 1 \times p + 0 \times (1-p) = p = p_X(1)$$

L3(3) May 13, 2021 13 / 45

Properties of Expectation



Not very surprising. Easy to prove using the definition.

$$\circ$$
 If $X \geq 0$, $\mathbb{E}[X] \geq 0$.

• If
$$a \le X \le b$$
, $a \le \mathbb{E}[X] \le b$.

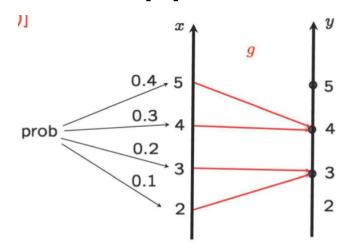
• For a constant c, $\mathbb{E}[c] = c$.

L3(3) May 13, 2021 14 / 45

Expectation of a function of a RV



- For a rv X, Y = g(X) is also a r.v.
- $\mathbb{E}[Y] = \mathbb{E}[g(X)] = \sum_{x} g(x) \rho_X(x)$
- Compute $\mathbb{E}[Y]$ for the following:



$$4 \times (0.4 + 0.3) + 3 \times (0.1 + 0.2)$$

= $2.8 + 0.9 = 3.7$

Linearity of Expectation

$$\mathbb{E}[aX+b]=a\mathbb{E}[X]+b$$

L3(3) May 13, 2021 15 / 45

Variance



- Measures how much the spread of a PMF is.
- What about $\mathbb{E}[X \mu]$, where $\mu = \mathbb{E}[X]$? Zero
- Then, what about $\mathbb{E}[(X \mu)^2]$?

Variance, Standard Deviation

$$var[X] = \mathbb{E}[(X - \mu)^2]$$

$$\sigma_X = \sqrt{\operatorname{var}[X]}$$

L3(3) May 13, 2021 16 / 45

Variance: Useful Property



•
$$\operatorname{var}[X] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$

 $\operatorname{var}[X] = \mathbb{E}[X^2 - 2\mu X + \mu^2]$
 $= \mathbb{E}[X^2] - 2\mu \mathbb{E}[X] + \mu^2 = \mathbb{E}[X^2] - \mu^2$

•
$$Y = X + b$$
, $var[Y] = var[X]$
 $var[Y] = \mathbb{E}[(X + b)^2] - (\mathbb{E}[X + b])^2$

•
$$Y = aX$$
, $var[Y] = a^2 var[X]$
 $var[Y] = \mathbb{E}[a^2X^2] - (a\mathbb{E}[X])^2$

Example: Variance of a Bernoulli rv (p)

$$\mu = \mathbb{E}[X] = 1 \times p + 0 \times (1 - p) = p$$

$$\mathbb{E}[X^2] = 1 \times p + 0 \times (1 - p) = p$$

$$\text{var}[X] = \mathbb{E}[X^2] - \mu^2 = p - p^2$$

$$= p(1 - p)$$

L3(3) May 13, 2021 17 / 45

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L3(4) May 13, 2021 18 / 45

Joint PMF



• Joint PMF. For two random variables X, Y, consider two events $\{X = x\}$ and $\{Y = y\}$, and

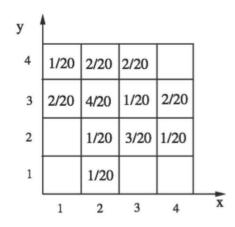
$$p_{X,Y}(x,y) \triangleq \mathbb{P}(\{X=x\} \cap \{Y=y\})$$

- $\sum_{x} \sum_{y} p_{X,Y}(x,y) = 1$
- Marginal PMF.

$$p_X(x) = \sum_{y} p_{X,Y}(x,y),$$
$$p_Y(y) = \sum_{x} p_{X,Y}(x,y)$$

Example.

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$$p_{X,Y}(1,3) = 2/20$$

$$p_X(4) = 2/20 + 1/20 = 3/20$$

$$\mathbb{P}(X = Y) = 1/20 + 4/20 + 3/20 = 8/20$$

L3(4)

Functions of Multiple RVs



• Consider a rv Z = g(X, Y). (Ex) X + Y, $X^2 + Y^2$. Then, PMF of Z is:

$$p_Z(z) = \mathbb{P}(g(X, Y) = z) = \sum_{(x,y):g(x,y)=z} p_{X,Y}(x,y)$$

• Similarly,

$$\mathbb{E}[Z] = \mathbb{E}[g(X,Y)] = \sum_{x} \sum_{y} g(x,y) p_{X,Y}(x,y)$$

L3(4) May 13, 2021 20 / 45

Linearity of Expectation for Multiple RVs



- Remember: $\mathbb{E}[aX + b] = a\mathbb{E}[X] + b$
- Similarly,

$$\mathbb{E}[X+Y] = \mathbb{E}[X] + \mathbb{E}[Y]$$

(easy to prove, using the definition.)

- $\mathbb{E}[X_1 \ldots + X_n] = \mathbb{E}[X_1] + \ldots + \mathbb{E}[X_n]$
- $\mathbb{E}[2X+3Y-Z]=2\mathbb{E}[X]+3\mathbb{E}[Y]-\mathbb{E}[Z]$

- Example. Mean of a binomial rv Y with (n, p)
- Y: number of successes in n Bernoulli trials with p
- $Y = X_1 + ... X_n$, where X_i is a Bernoulli rv.
- $\mathbb{E}[Y] = n\mathbb{E}[X_i] = n\mathbb{P}(X_i = 1) = np$

Message. When some rv X is written as a linear combination of other rvs, X becomes easy to handle.

L3(4) May 13, 2021 21 / 45

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L3(5) May 13, 2021 22 / 45

Conditional PMF: Conditioning on an event



Remember two probability laws: $\mathbb{P}(\cdot)$ and $\mathbb{P}(\cdot|A)$ for an event A.

•
$$p_X(x) \triangleq \mathbb{P}(X=x)$$

•
$$\mathbb{E}[X] = \sum_{x} x p_X(x)$$

•
$$\mathbb{E}[g(X)] = \sum_{x} g(x) p_X(x)$$

•
$$var[X] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$

•
$$p_{X|A}(x) \triangleq \mathbb{P}(X = x|A)$$

• $\mathbb{E}[X|A] \triangleq \sum_{x} x p_{X|A}(x)$

•
$$\mathbb{E}[X|A] \triangleq \sum_{x} x p_{X|A}(x)$$

•
$$\mathbb{E}[g(X)|A] \triangleq \sum_{x} g(x) p_{X|A}(x)$$

•
$$\operatorname{var}[X|A] \triangleq \mathbb{E}[X^2|A] - (\mathbb{E}[X|A])^2$$

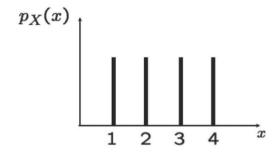
• (Note) $p_{X|A}(x)$, $\mathbb{E}[X|A]$, $\mathbb{E}[g(X)|A]$, and var[X|A] are all just notations!

L3(5) May 13, 2021 23 / 45

Example: Conditional PMF



$$A = \{X \ge 2\}$$



$$\mathbb{E}[X] = \frac{1}{4} (1 + 2 + 3 + 4) = 2.5$$

$$var[X] = \mathbb{E}[X^2] - (\mathbb{E}[X])^2$$
$$= \frac{1}{4}(1+2^2+3^2+4^2) - 2.5^2$$

$$\mathbb{E}[X|A] = \frac{1}{3}(2+3+4) = 3$$

$$var[X|A] = \mathbb{E}[X^2|A] - (\mathbb{E}[X|A])^2$$
$$= \frac{1}{3}(2^2 + 3^2 + 4^2) - 3^2 = 2/3$$

L3(5) May 13, 2021 24 / 45

Conditional PMF: Conditioning on a RV



What do we mean by "conditioning on a rv"? Consider $A = \{Y = y\}$ for a rv Y.

•
$$p_{X|A}(x) \triangleq \mathbb{P}(X = x|A)$$

•
$$\mathbb{E}[X|A] \triangleq \sum_{x} x p_{X|A}(x)$$

•
$$\mathbb{E}[g(X)|A] \triangleq \sum_{x} g(x) p_{X|A}(x)$$

•
$$\operatorname{var}[X|A] \triangleq \mathbb{E}[X^2|A] - (\mathbb{E}[X|A])^2$$

•
$$p_{X|Y}(x|y) \triangleq \mathbb{P}(X = x|Y = y)$$

• $\mathbb{E}[X|Y = y] \triangleq \sum_{x} x p_{X|Y}(x|y)$

•
$$\mathbb{E}[X|Y=y] \triangleq \sum_{x} x p_{X|Y}(x|y)$$

•
$$\mathbb{E}[g(X)|Y=y] \triangleq \sum_{x} g(x) p_{X|Y}(x|y)$$

•
$$\operatorname{var}[X|Y = y] \triangleq \mathbb{E}[X^2|Y = y] - (\mathbb{E}[X|Y = y])^2$$

L3(5) May 13, 2021 25 / 45

Conditional PMF



Conditional PMF

$$p_{X|Y}(x|y) \triangleq \mathbb{P}(X=x|Y=y) = \frac{p_{X,Y}(x,y)}{p_{Y}(y)}$$

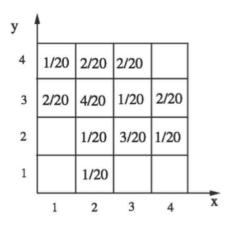
for y such that $p_Y(y) > 0$.

- $\bullet \quad \sum_{x} p_{X|Y}(x|y) = 1$
- Multiplication rule

$$p_{X,Y}(x,y) = p_Y(y)p_{X|Y}(x|y)$$
$$= p_X(x)p_{Y|X}(y|x)$$

• $p_{X,Y,Z}(x, y, z) = p_X(x)p_{Y|X}(y|x)p_{Z|X,Y}(z|x, y)$

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$$p_{X|Y}(2|2) = \frac{1}{1+3+1}$$

$$p_{X|Y}(3|2) = \frac{3}{1+3+1}$$

$$\mathbb{E}[X|Y=3] = 1(2/9)+2(4/9)+3(1/9)+4(2/9)$$

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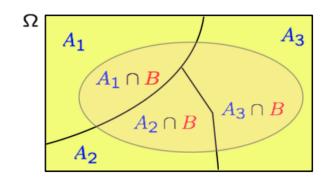
Remind: Total Probability Theorem (from Lecture 2)

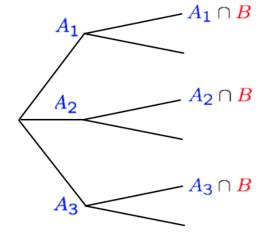


- Partition of Ω into A_1, A_2, A_3
- Known: $\mathbb{P}(A_i)$ and $\mathbb{P}(B|A_i)$
- What is $\mathbb{P}(B)$?

Total Probability Theorem

$$\mathbb{P}(B) = \sum_{i} \mathbb{P}(A_i) \mathbb{P}(B|A_i)$$





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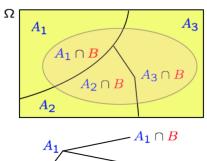
Total Probability Theorem: $B = \{X = x\}$

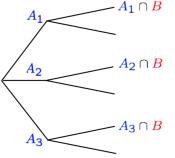


• Partition of Ω into A_1, A_2, A_3

Total Probability Theorem

$$p_X(x) = \sum_i \mathbb{P}(A_i)\mathbb{P}(X = x|A_i) = \sum_i \mathbb{P}(A_i)p_{X|A_i}(x)$$





L3(5) May 13, 2021 28 / 45

Total Expectation Theorem for $\{A_i\}$



• Partition of Ω into A_1, A_2, A_3

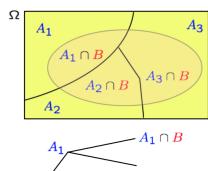
Total Probability Theorem

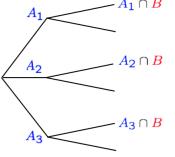
$$p_X(x) = \sum_i \mathbb{P}(A_i)\mathbb{P}(X = x|A_i) = \sum_i \mathbb{P}(A_i)p_{X|A_i}(x)$$

Total Expectation Theorem

$$\mathbb{E}[X] = \sum_{i} \mathbb{P}(A_i) \mathbb{E}[X|A_i]$$

• Weighted average of expectations from A_i 's perspective





L3(5) May 13, 2021 29 / 45

Total Expectation Theorem for $\{Y = y\}$



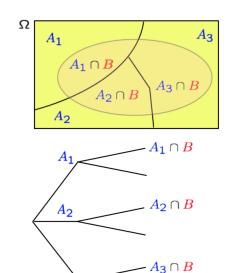
• Partition of Ω into A_1, A_2, A_3

Total Expectation Theorem

$$\mathbb{E}[X] = \sum_{i} \mathbb{P}(A_{i}) \mathbb{E}[X|A_{i}]$$

Total Expectation Theorem

$$\mathbb{E}[X] = \sum_{y} \mathbb{P}(\frac{Y}{Y} = y) \mathbb{E}[X|\frac{Y}{Y} = y] = \sum_{y} p_{Y}(y) \mathbb{E}[X|\frac{Y}{Y} = y]$$



L3(5) May 13, 2021 30 / 45

Example 1: Total Expectation Theorem



- Question. What is $\mathbb{E}(X)$?
- (1) Just using the definition of expectation,

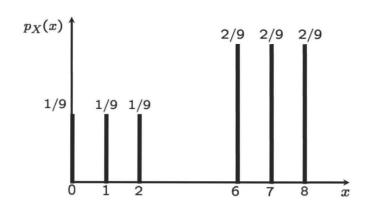
$$\mathbb{E}[X] = \frac{1}{9}(0+1+2) + \frac{2}{9}(6+7+8)$$
$$= \frac{3+12+14+16}{9} = 5$$

(2) Let's use TET, for which consider

$$A_1 = \{X \in \{0, 1, 2\}\}, A_2 = \{X \in \{6, 7, 8\}\}$$

$$\mathbb{E}[X] = \sum_{i=1,2} \mathbb{P}(A_i)\mathbb{E}[X|A_i]$$

$$= 1/3 \cdot 1 + 2/3 \cdot 7 = 5$$



L3(5) May 13, 2021 31 / 45

Example 2: Mean of Geometric rv



- Write softwares over and over, and each time w.p. p of working correctly (independent from previous programs).
- X: number of trials until the program works correctly.
- (Q) $\mathbb{E}(X)$?
- X is a geometric rv
- Direct computation is boring.

$$\mathbb{E}[X] = \sum_{k=1}^{\infty} k(1-p)^{k-1}p = p + 2(1-p)p + 3(1-p)^2p + \cdots$$

Total expectation theorem and a notion of memorylessness helps a lot.

L3(5) May 13, 2021 32 / 45

Memoryless Property: Motivating Example





No bus arrives



How long do I have to wait? Probability of waiting for more than n mins?

$$\mathbb{P}(X > n + m | X > m)$$

Yung arrives

0

m mins

A bus left at time 0



0

time



Lin arrives

How long do I have to wait? Probability of waiting for more than n mins?

$$\mathbb{P}(X > n)$$

L3(5) May 13, 2021 33 / 45

Background: Memoryless Property



- Some random variable often does not have memory.
- Definition. A random variable X is called $\boxed{\text{memoryless}}$ if, for any $n,m\geq 0,$ $\mathbb{P}(X>n+m|X>m)=\mathbb{P}(X>n)$
- Meaning. Conditioned on X > m, X m's distribution is the same as the original X.

$$\mathbb{P}(X-m>n|X>m)=\mathbb{P}(X>n)$$

L3(5) May 13, 2021 34 / 45

Background: Memoryless Property of Geometric RVs



- Theorem. Any geometric random variable is memoryless.
- Remind. Geometric rv X with parameter p

$$\mathbb{P}(X=k)=(1-p)^{k-1}p, \quad \mathbb{P}(X>k)=\sum_{i=k+1}^{\infty}(1-p)^{i-1}p=(1-p)^k$$

Proof.

$$\mathbb{P}(X > n + m | X > m) = \frac{\mathbb{P}(X > n + m \text{ and } X > m)}{\mathbb{P}(X > m)} = \frac{\mathbb{P}(X > n + m)}{\mathbb{P}(X > m)}$$
$$= \frac{(1 - p)^{n+m}}{(1 - p)^m} = (1 - p)^n = \mathbb{P}(X > n)$$

• Meaning. Conditioned on X > m, X - m is geometric with the same parameter.

L3(5) May 13, 2021 35 / 45

Back to Example 2: Mean of Geometric rv



•
$$A_1=\{X=1\}$$
 (first try is success), $A_2=\{X>1\}$ (first try is failure).
$$\mathbb{E}[X]=1+\mathbb{E}[X-1]$$

$$=1+\mathbb{P}(A_1)\mathbb{E}[X-1|X=1]+\mathbb{P}(A_2)\mathbb{E}[X-1|X>1]$$
 (from TET)
$$=1+(1-p)\mathbb{E}[X]$$
 (from memorylessness)

• Thus,
$$\mathbb{E}[X] = \frac{1}{p}$$

L3(5) May 13, 2021 36 / 45

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Independence, Conditional Independence



Two events

$$\mathbb{P}(A \cap B) = \mathbb{P}(A) \cdot \mathbb{P}(B)$$

$$\mathbb{P}(A \cap B | C) = \mathbb{P}(A | C) \cdot \mathbb{P}(B | C)$$

A rv and an event

$$\mathbb{P}(\{X = x\} \cap B) = \mathbb{P}(X = x) \cdot \mathbb{P}(B), \text{ for all } x$$

$$\mathbb{P}(\{X = x\} \cap B | C) = \mathbb{P}(X = x | C) \cdot \mathbb{P}(B | C), \text{ for all } x$$

Two rvs

$$\mathbb{P}(\{X=x\} \cap \{Y=y\}) = \mathbb{P}(X=x) \cdot \mathbb{P}(Y=y), \text{ for all } x, y$$
$$p_{X,Y}(x,y) = p_X(x) \cdot p_Y(y)$$

$$\mathbb{P}(\{X=x\} \cap \{Y=y\} | Z=z) = \mathbb{P}(X=x | Z=z) \cdot \mathbb{P}(Y=y | Z=z), \text{ for all } x, y$$
$$p_{X,Y|Z}(x,y) = p_{X|Z}(x) \cdot p_{Y|Z}(y)$$

L3(6) May 13, 2021 38 / 45

Example



• *X* ⊥⊥ *Y*?

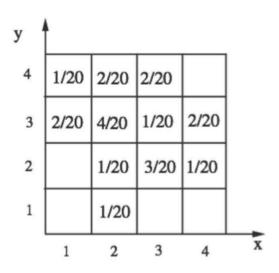
$$p_{X,Y}(1,1) = 0$$
 $p_X(1) = 3/20$
 $p_Y(1) = 1/20$

• $X \perp \!\!\! \perp Y | \{X \le 2 \text{ and } Y \ge 3\}$?

VIDEO PAUSE

Y = 4 (1/3)	1/9	2/9
Y = 3 (2/3)	2/9	4/9
	X = 1 (1/3)	X = 2 (2/3)

- Yes.



Expectation and Variance



Always true.

$$\mathbb{E}[X+Y] = \mathbb{E}[X] + \mathbb{E}[Y]$$

- Generally, $\mathbb{E}[g(X,Y)]
 eq g(\mathbb{E}[X],\mathbb{E}[Y])$
- However, if $X \perp \!\!\! \perp Y$,

$$\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$$
 $\mathbb{E}[g(X)h(Y)] = \mathbb{E}[g(X)]\mathbb{E}[g(Y)]$

• Proof.

$$\mathbb{E}[g(X)h(Y)] = \sum_{x} \sum_{y} g(x)h(y)p_{X,Y}(x,y)$$
$$= \sum_{x} g(x)p_{X}(x) \sum_{y} h(y)p_{Y}(y)$$

- Always true. $var[aX] = a^2 var[X]$, var[X + a] = var[X]
- Generally, $var[X + Y] \neq var[X] + var[Y]$ (next slide)
- However, if $X \perp \!\!\! \perp Y$, var[X + Y] = var[X] + var[Y]
- Practice.

$$\circ X = Y \Longrightarrow var[X + Y] = 4var[X]$$

$$X = -Y \Longrightarrow var[X + Y] = 0$$

L3(6)

$var[X + Y] \neq var[X] + var[Y]$



• Why not generally true?

$$var[X + Y] = \mathbb{E}[(X + Y)^{2}] - (\mathbb{E}[X + Y])^{2}$$

$$= \mathbb{E}[X^{2} + Y^{2} + 2XY] - ((\mathbb{E}[X])^{2} + (\mathbb{E}[Y])^{2} + 2\mathbb{E}[X]\mathbb{E}[Y])$$

$$= var[X] + var[Y] + 2(\mathbb{E}[XY] - \mathbb{E}[X]\mathbb{E}[Y])$$

- \circ $X \perp\!\!\!\perp Y$ is a sufficient condition for $\mathbb{E}[XY] = \mathbb{E}[X]\mathbb{E}[Y]$
- Also, a necessary condition? we will see later, when we study covariance.

L3(6) May 13, 2021 41 / 45

Example: The hat problem (1)



- n people throw their hats in a box and then pick one at random
- X: number of people with their own hat
- $\mathbb{E}[X]$? var[X]?
- All permutations are equally likely as 1/n!. Thus, this equals to picking one hat at a time.
- Key step 1. Define a rv $X_i = 1$ if i selects its own hat and 0 otherwise.

$$X = \sum_{i=1}^{n} X_i.$$

• $\{X_i\}, i = 1, 2, ..., n$: identically distributed (from symmetry)

L3(6) May 13, 2021 42 / 45

Example: The hat problem (2)



- $\mathbb{E}[X] = n\mathbb{E}[X_1] = n\mathbb{P}(X_1 = 1) = n \times \frac{1}{n} = 1.$
- Key step 2. Are X_i s are independent? If yes, easy to get var(X).
- Assume n=2. Then, $X_1=1 \to X_2=1$, and $X_1=0 \to X_2=0$. Thus, dependent.

$$\operatorname{\mathsf{var}}(X) = \mathbb{E}[X^2] - (\mathbb{E}[X])^2 = \mathbb{E}\Big[\sum_i X_i^2 + \sum_{i,j:i \neq j} X_i X_j\Big] - (\mathbb{E}[X])^2$$

$$\mathbb{E}[X_i^2] = \mathbb{E}[X_1^2] = 1 \times \frac{1}{n} + 0 \times \frac{n-1}{n} = \frac{1}{n}$$

$$\mathbb{E}[X_i X_j] = \mathbb{E}[X_1 X_2] = 1 \times \mathbb{P}(X_1 X_2 = 1) = \mathbb{P}(X_1 = 1) \mathbb{P}(X_2 = 1 | X_1 = 1), \quad (i \neq j)$$

•
$$\mathbb{E}[X^2] = n\mathbb{E}[X_1^2] + n(n-1)\mathbb{E}[X_1X_2] = n\frac{1}{n} + n(n-1)\frac{1}{n(n-1)} = 2$$

•
$$var(X) = 2 - 1 = 1$$

L3(6) May 13, 2021 43 / 45



Questions?

L3(6) May 13, 2021 44 / 45

Review Questions



- 1) What is Random Variable? Why is it useful?
- 2) What is PMF (Probability Mass Function)?
- 3) Explain Bernoulli, Binomial, Poisson, Geometric rvs, when they are used and what their PMFs are.
- 4) What are joint and marginal PMFs?
- 5) Describe and explain the total probability/expectation theorem for random variables?
- 6) When is it useful to use total probability/expectation theorem?
- 7) What is conditional independence?

L3(6) May 13, 2021 45 / 45