



# Week4 Mixed Models 1: Introduction to Multilevel Models

|        |                           |
|--------|---------------------------|
| Assign |                           |
| Status | Epidemiology and Big Data |

Mixed Model

random effect

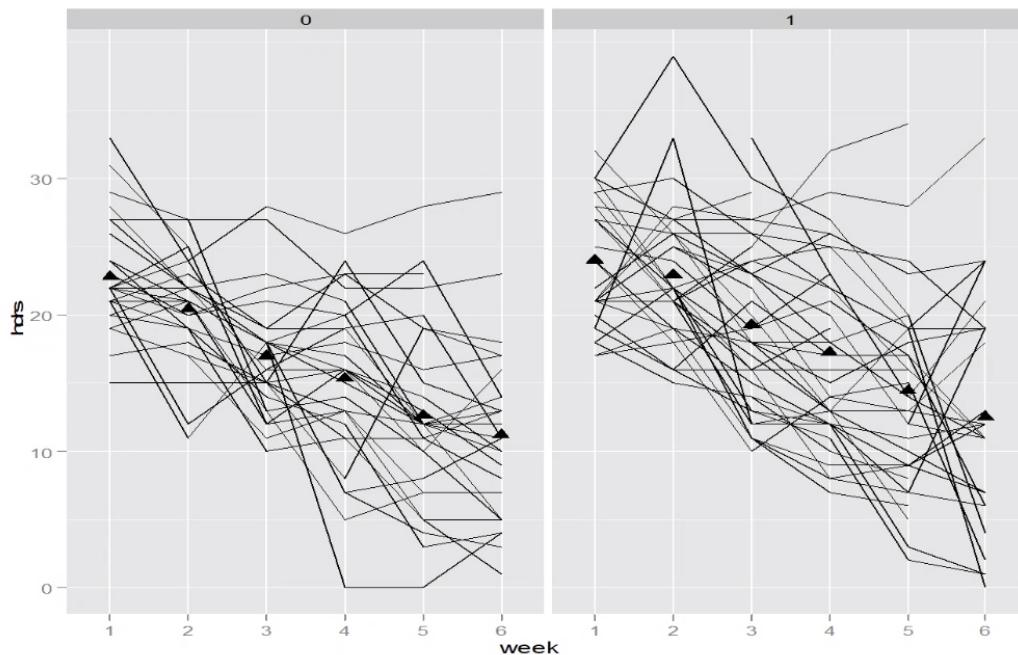
Reisby Data, After several trials

- std increase
- correlation decrease

ex vs indogenous

# Example: Reisby Data

"Spaghetti Plot" (R, using ggplot2)



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Introduction to multilevel data

Example: multilevel data (children within schools)

Example: London Schools

1. linear regression, mean exam per school vs mean LRT ("aggregated data")
2. linear regression, all schools together ("disaggregated data")
3. linear regression per school
4. linear regression, all schools together, regression with main effect and interactions to allow for different intercepts and slopes
5. Linear Mixed Model

Random intercept only vs Random intercept + random slope

The problem, and some possible solutions

The mixed model solution

Adding random effects (random intercept, random slope)

Adding fixed effects (school- and child-level) to the model

Interpretation of mixed models

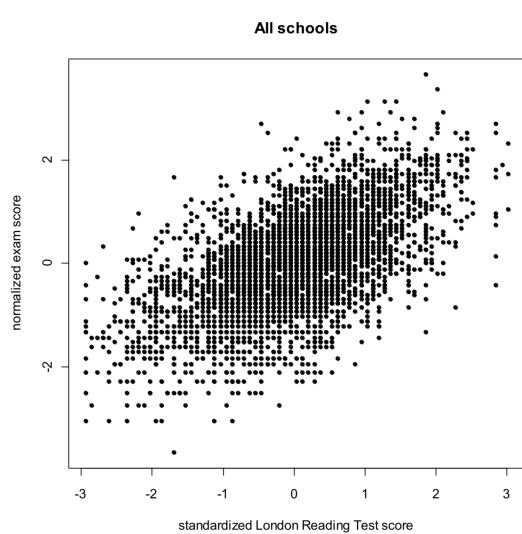
Summary

## Introduction to multilevel data

### Example: multilevel data (children within schools)

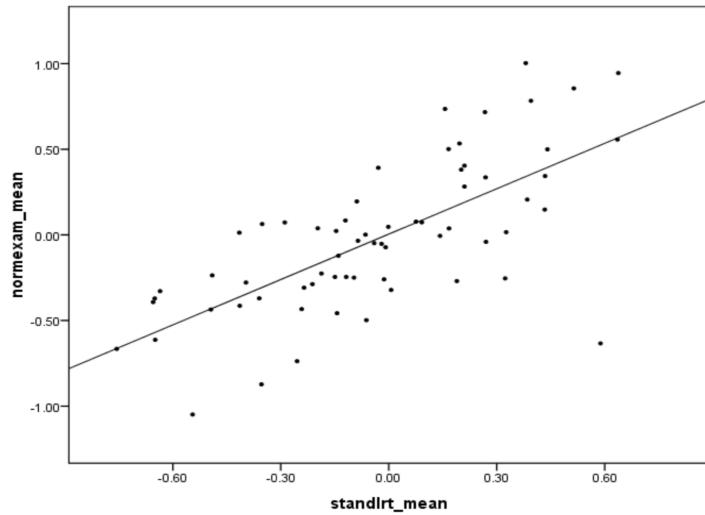
- Hierarchical structure of data
- Variation at all levels
- “Units” within a level expected to be correlated
- Variables can be measured at different level

### Example: London Schools



#### Variables in dataset:

- School ID
- Student ID
- **Normalised exam score**  
(outcome variable)
- **Standardised LR test score**
- Student gender
- School gender
- School average of intake score
- Student level Verbal Reasoning  
(VR) score category at intake
- Category of students' intake  
score (averaged)

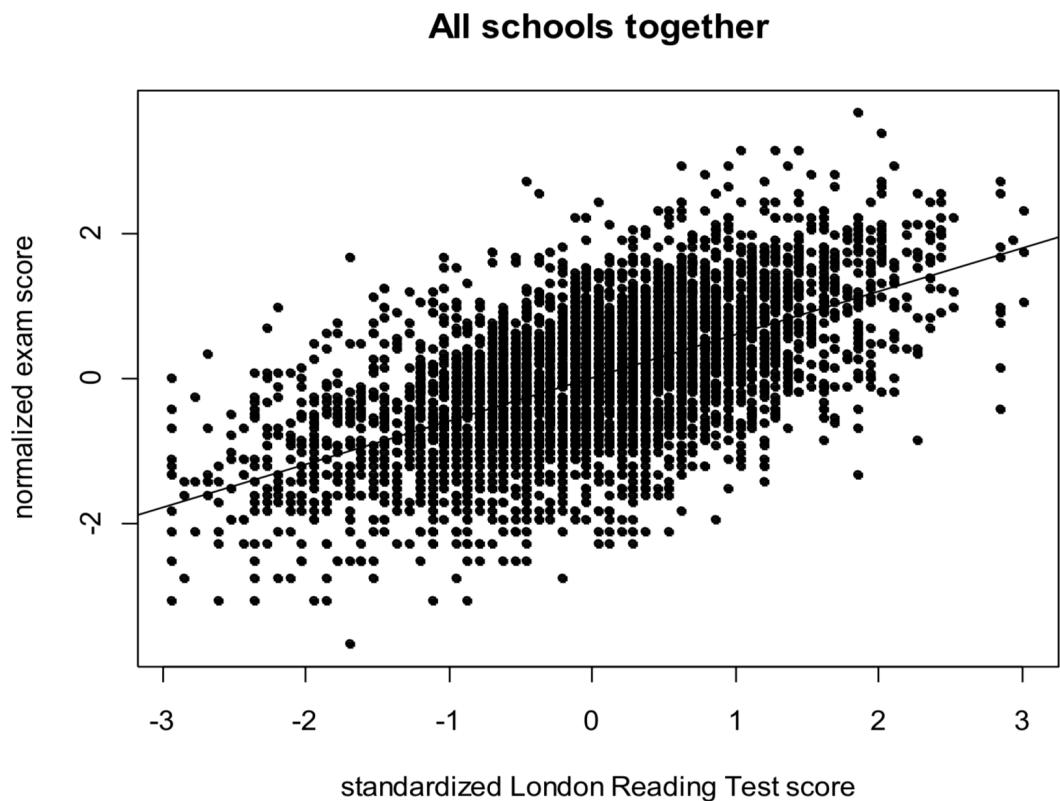


## 1. linear regression, mean exam per school vs mean LRT ("aggregated data")

- estimate for intercept: 0.005 (se 0.040)
- estimate for slope: 0.884 (se 0.116)

### Disadvantages:

- every school (regardless of sample size) given equal weight
- $N = 65$
- school-level variables possible, but not child-level variables
- we can only make inference at school level, not child-level
- possibility of “ecological fallacy”

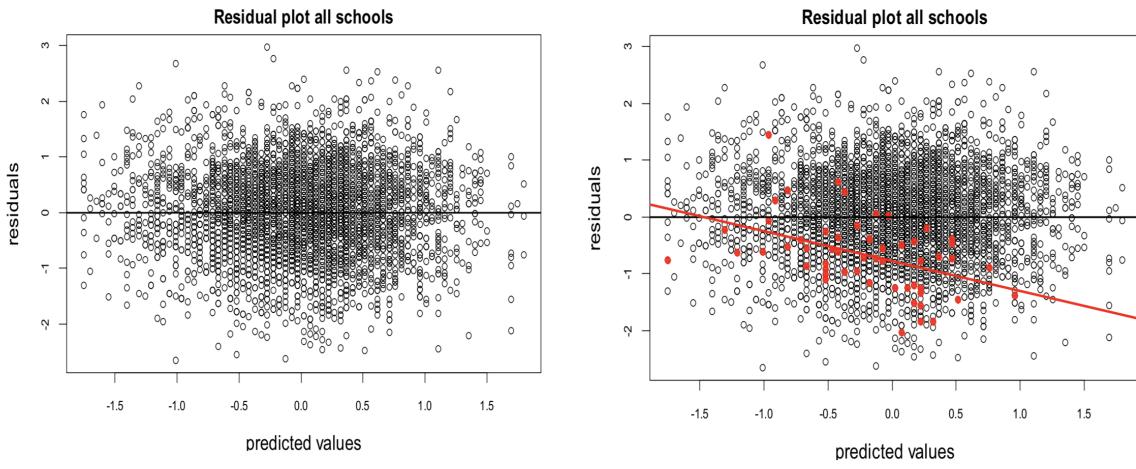


## 2. linear regression, all schools together ("disaggregated data")

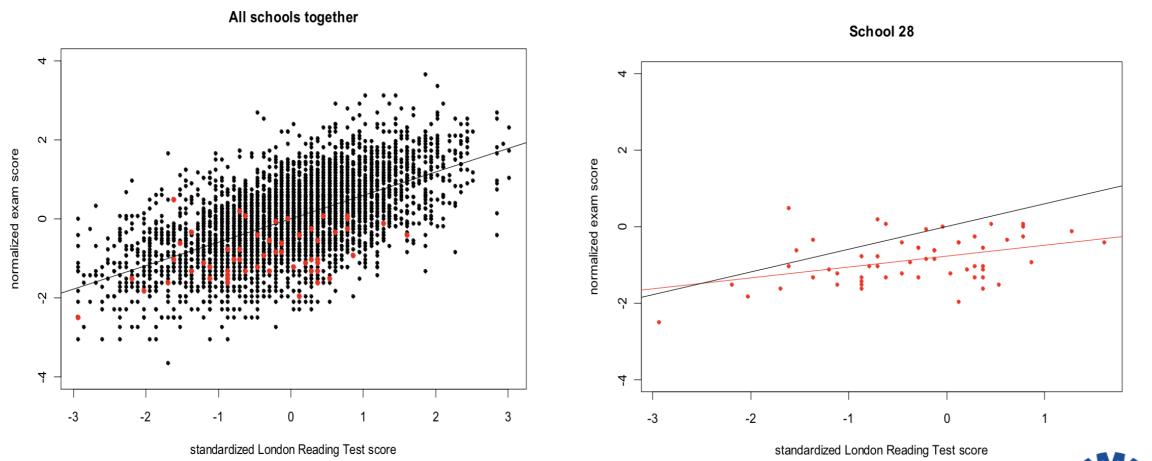
- estimate for intercept: - 0.001 (se 0.013)
- estimate for slope: 0.595 (se 0.013)

### Disadvantages:

- inflates sample size, especially for level-2 variables
  - SE's of level-2 variables tend to be underestimated → p-values too small, CI's too narrow (type I error inflated)
  - SE's of level-1 variables may be over or underestimated
- ignore correlated residuals (correlation of children within schools)



?



If we split the School\_28 separately, we will find that the slope is totally different

### 3. linear regression per school

summary intercepts:

- mean=-0.068;sd=0.519;sem=0.064

summary slopes:

- mean=0.425;sd=0.939;sem=0.116

### Disadvantages:

- 65 different regressions, how to combine the results?
  - mean slope: every school has equal weight
  - standard error of parameter estimate correct?
- child-level variables possible, but not school-level variables



## 4. linear regression, **all schools together**, regression with **main effect** and **interactions** to allow for different intercepts and slopes

### Advantage over previous analysis:

- now we can include both child- and school-level variables
- residuals probably normally distributed (with a constant variance?) around individual lines

### **Disadvantages:**

- We wanted 1 intercept and 1 slope for LRT, but:
- 65 schools, so 1 reference category and 64 estimates for intercepts (main effects per school) + 64 estimates for interactions (slopes per school)!
- Which school is the reference?
  - We can't generalize beyond these 65 schools
  - This model uses 128 extra df for all those intercepts & slopes

## **5. Linear Mixed Model**

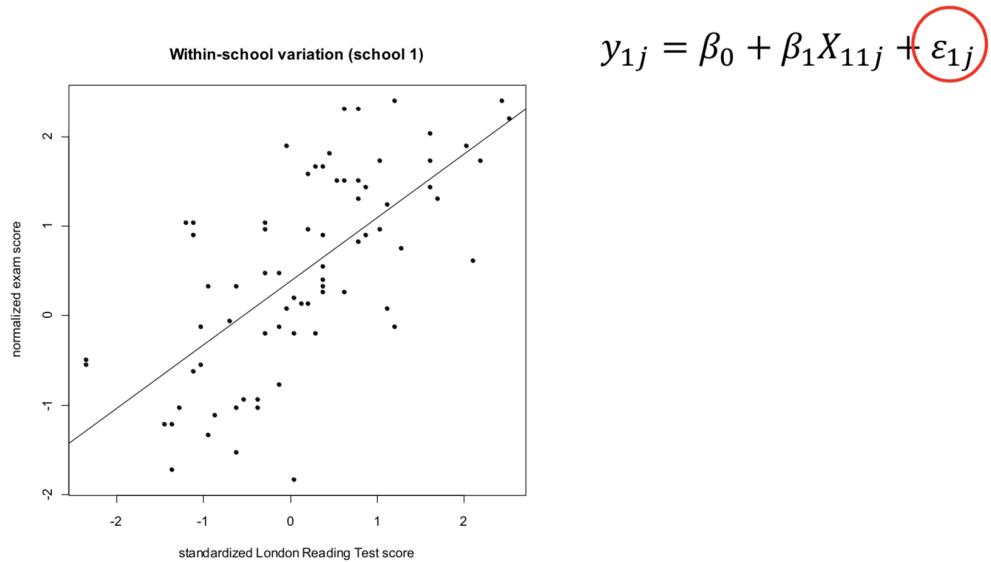
### **Advantages:**

- sample size correct, account for correlation of children within schools
  - so: correct SE's/p-values/CI's
- no need for 64 main effects and interactions
  - differences between schools captured one or more 'variance components'
- both child-level and school-level variables simultaneously
  - so: inference for both children and schools
  - interactions between child- and school-level variables possible
- examine variation at different levels
- models work well in presence of missing outcomes (longitudinal)

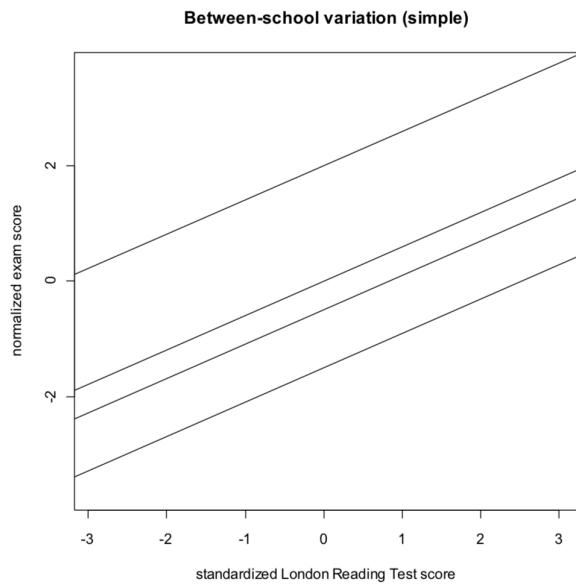
**Mixed Models = random effects models (inaccurately) = random coefficient models**

- Mixed models = fixed effects + random effects
- Some variables (or: their coefficients) can be included as both "fixed" (of interest) and "random" (random variation across the **level-2 units**)

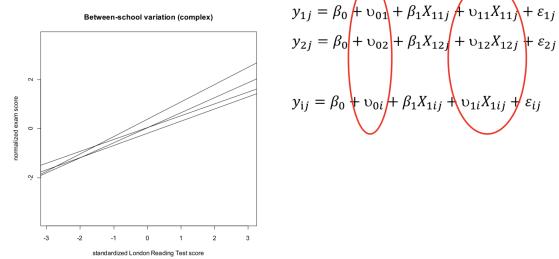
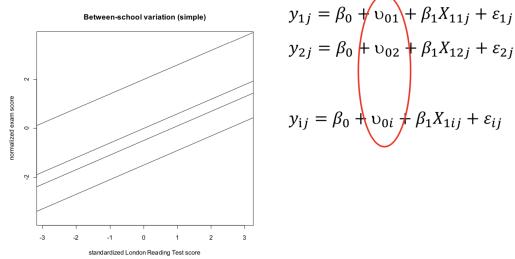
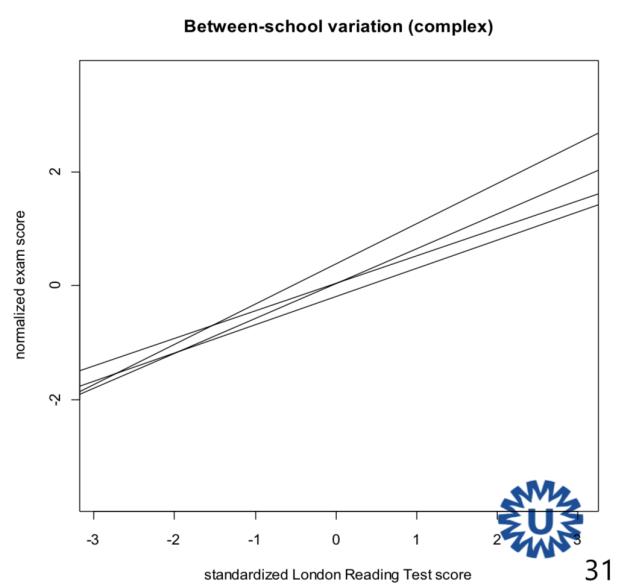
## Random intercept only vs Random intercept + random slope



Random intercept only:



Random intercept + random slope:



## Interlude: some notation

- level-1 (child) model:  $y_{ij} = b_{0i} + b_{1i} \cdot x_{1ij} + \varepsilon_{ij}$
- level-2 (school) model:  $b_{0i} = \beta_0 + v_{0i}$ ,  $b_{1i} = \beta_1 + v_{1i}$
- combine the two:  $y_{ij} = \beta_0 + v_{0i} + \beta_1 \cdot x_{1ij} + v_{1i} \cdot x_{1ij} + \varepsilon_{ij}$ 
  - rewrite:  $y_{ij} = (\beta_0 + v_{0i}) + (\beta_1 + v_{1i}) \cdot x_{1ij} + \varepsilon_{ij}$
- $y_{ij}$  : outcome (exam score) for  $j^{\text{th}}$  child in  $i^{\text{th}}$  school
- $x_{1ij}$ : 1st explanatory var (LRT score) at level 1 ( $j^{\text{th}}$  child in  $i^{\text{th}}$  school)
- $\beta_0, \beta_1, \dots$  : regression coefficients for overall effects of explanatory vars (“fixed effects”)
- $v_{0i}$  : individual effect of  $i^{\text{th}}$  school on intercept (“random effect”)
- $v_{1i}$  : individual effect of  $i^{\text{th}}$  school on slope (for LRT) (“random effect”)
- $\varepsilon_{ij}$  : level-1 residual ( $j^{\text{th}}$  child in  $i^{\text{th}}$  school)



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- A random intercept per school allows schools to have different intercepts
- vice versa, A random slope  $\Rightarrow$  different slope
- All variables of interest are added as fixed (The target should be fixed)
- Model assumptions:
  - $\varepsilon_{ij} \sim N(0, \sigma_e^2)$ ;  $v_{0i} \sim N(0, \sigma_{v0}^2)$ ;  $v_{1i} \sim N(0, \sigma_{v1}^2)$
  - $\varepsilon_{ij}$  independent
  - $\text{cov}(v_{0i}, v_{1i}) = \sigma_{v01}$
  - $\text{cov}(\varepsilon_{ij}, v_{0i}) = \text{cov}(\varepsilon_{ij}, v_{1i}) = 0$

The parameter  $\mu$  is the [mean](#) or [expectation](#) of the distribution (and also its [median](#) and [mode](#)), while the parameter  $\sigma$  is its [standard deviation](#).<sup>[1]</sup> The [variance](#) of the distribution is  $\sigma^2$

- normal distribution mean = 0, variance = 1
- **(mean, standard deviation, variance)**

## The problem, and some possible solutions

### The mixed model solution

#### Adding random effects (random intercept, random slope)

#### Adding fixed effects (school- and child-level) to the model

### Interpretation of mixed models

### Summary

## Q&A

p18

sem?

child params

LRT? = London Reading test

level2 unit

variable of interest = 哪些是你關注的變量 (目標變量)

## Mixed Models: what is a “random effect”?

- A random intercept per school allows schools to have different intercepts
- A random effect for LRT per school allows the effect of LRT on exam score to differ per school (“random slope for LRT” = different slope for exam-LRT relation for each school)
- Random effect (“slope”) can also be for a categorical variable
  - difference between boys and girls on exam score could differ per school
  - treatment effect on an outcome can be thought to vary per center in a multi-center study
- All variables of interest are added as fixed
- Depending on theory, none/one/some fixed variables may also be modelled as random



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Random effect (“slope”) can also be for a categorical variable, (intercept can be too? )

