Introduction to Deep Learning

5. Maximum Likelihood and Logistic Regression

STAT 157, Spring 2019, UC Berkeley

Mu Li and Alex Smola

courses.d2l.ai/berkeley-stat-157

aws


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Homework

- Please follow the submission instructions on HW3
  - Don’t submit homework1_v1 homework1_a homework1_updated and hope that we figure it out …
  - Please check the PDF that it is readable
- Homework 1 solutions have an example for how to format
- Converting notebooks to PDF

```
ipython nbconvert --to pdf notebook.ipynb
```

Project

- If you haven't found a team yet, do it TODAY
- Email to berkeley-stat-157@googlegroups.com with **Names, Project title, Short abstract (optional)**
- Extended deadline - February 6, 11:59 PM PST
- **Midterm Presentation**
  - Due March 5, 11:59 PM PST (by e-mail)
  - 1-2 slides in PDF, 1-2 pages report (NIPS style file)
  - 4 minute talk
  - Needs to cover **What, How, Why, Novelty**.

Outline

- **Maximum Likelihood**
  - Gauss and means
  - More loss functions (l_1 loss, trimmed mean)
  - Regression revisited
- **Classification**
  - Computing discrete probabilities
  - Likelihood and loss functions
- **Information Theory**


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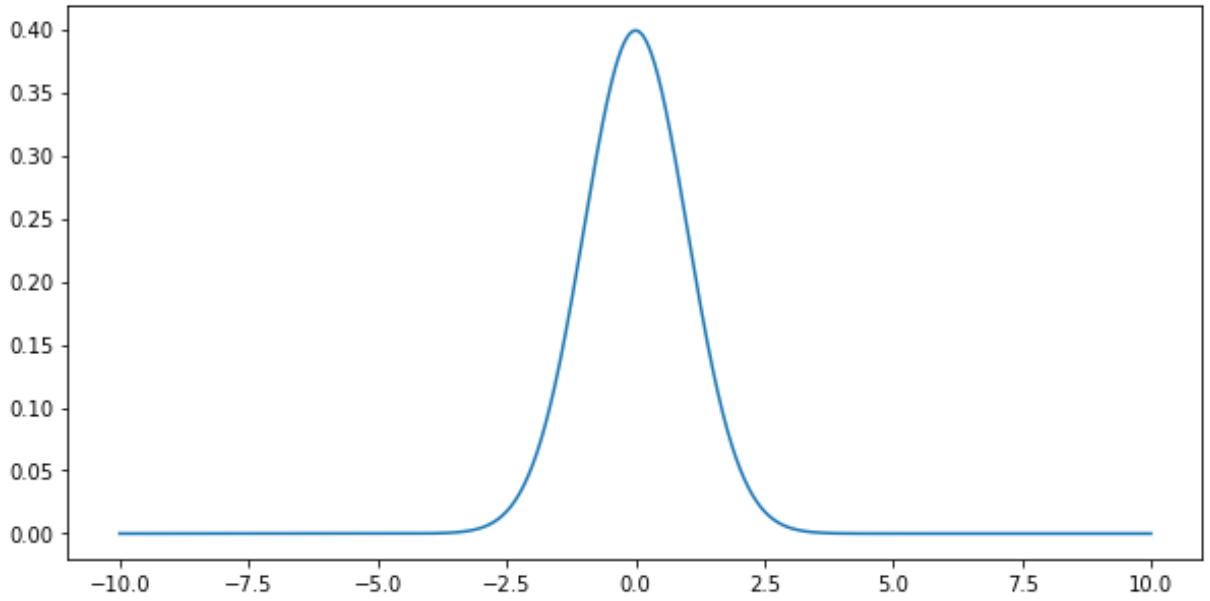
Flashback - Normal Distribution

Density $$p(x) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{(x-\mu)^2}{2\sigma^2}\right)$$



Estimating the parameters in a Gaussian

- **Mean**

$$\mu = \mathbf{E}[x] \text{ hence } \hat{\mu} = \frac{1}{n}\sum_{i=1}^{n} x_i$$

- **Variance**

$$\sigma^2 = \mathbf{E}\left[(x-\mu)^2\right] \text{ hence } \hat{\sigma}^2 = \frac{1}{n}\sum_{i=1}^{n} (x_i - \hat{\mu})^2$$

Estimating the parameters in a Gaussian

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- **Variance**

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Why?

Likelihood

- Observe some data $X = \{x_1, \ldots x_n\}$
- Assume that the data is drawn from a Gaussian

$$p(X; \mu, \sigma^2) = \prod_{i=1}^{n} p(x_i; \mu, \sigma^2) = \prod_{i=1}^{n} \frac{1}{2\pi\sigma^2} \exp\left(-\frac{(x_i - \mu)^2}{2\sigma^2}\right)$$

- **Fitting parameters is maximizing** $p(X; \mu, \sigma^2)$ **wrt.** μ, σ^2 (maximize likelihood that data was generated by model)
- **Practical simplification**

$$\underset{\mu,\sigma^2}{\text{maximize}}\, p(X; \mu, \sigma^2) \iff \underset{\mu,\sigma^2}{\text{minimize}} - \log p(X; \mu, \sigma^2)$$

Maximum Likelihood

- Estimate parameters by finding ones that explain the data

$$\underset{\mu,\sigma^2}{\text{minimize}} - \log p(X; \mu, \sigma^2)$$

- **Decompose likelihood**

$$-\log p(X; \mu, \sigma^2) = \sum_{i=1}^{n} \frac{1}{2} \log(2\pi\sigma^2) + \frac{1}{2\sigma^2}(x_i - \mu)^2 = \frac{n}{2} \log(2\pi\sigma^2) + \frac{1}{2\sigma^2} \sum_{i=1}^{n} (x_i - \mu)^2$$

Minimized for $\mu = \frac{1}{n} \sum_{i=1}^{n} x_i$

Maximum Likelihood

- Estimating the variance

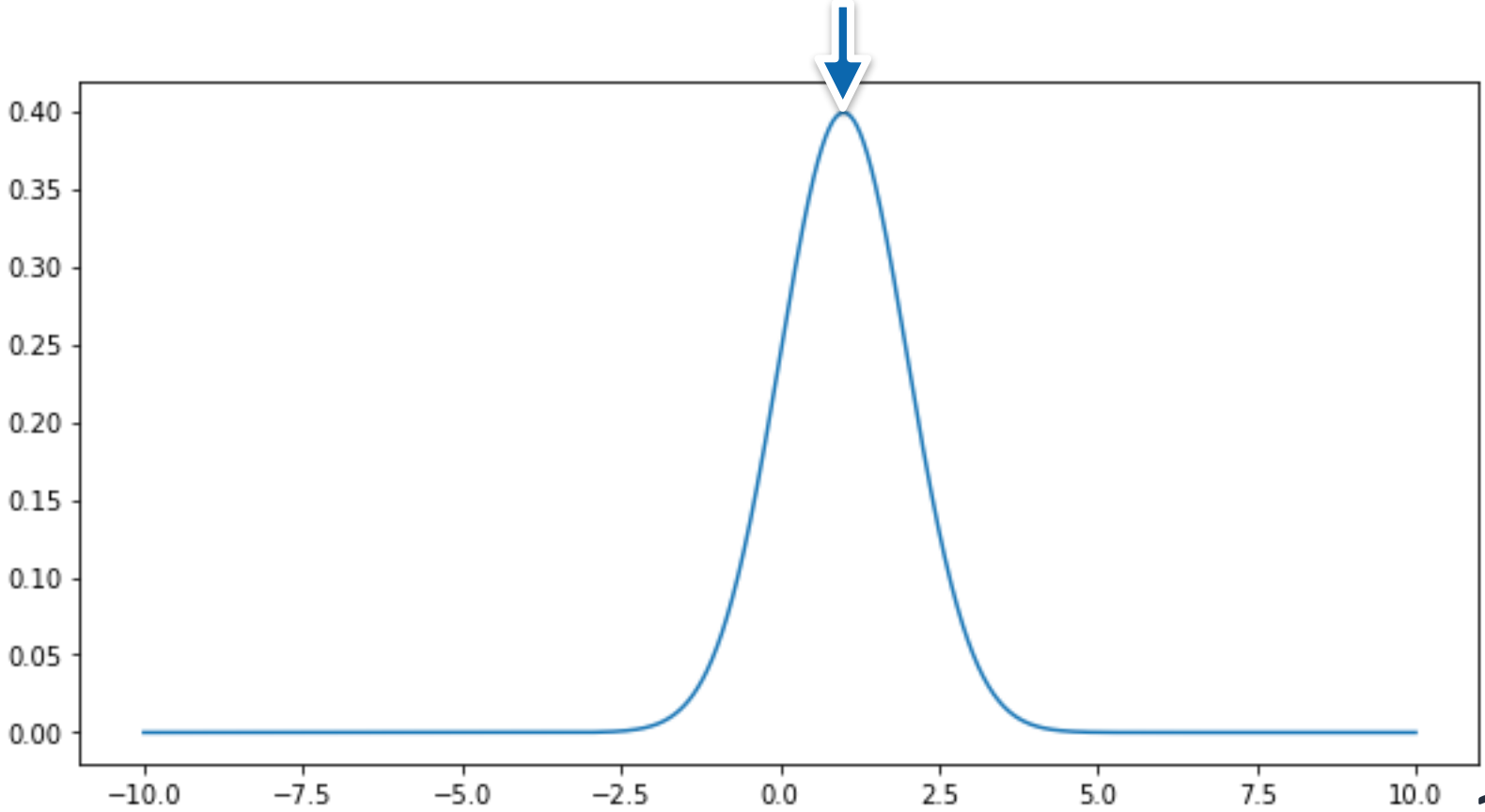
$$\frac{n}{2}\log(2\pi\sigma^2) + \frac{1}{2\sigma^2}\sum_{i=1}^{n}(x_i - \mu)^2$$

- Take derivatives with respect to it

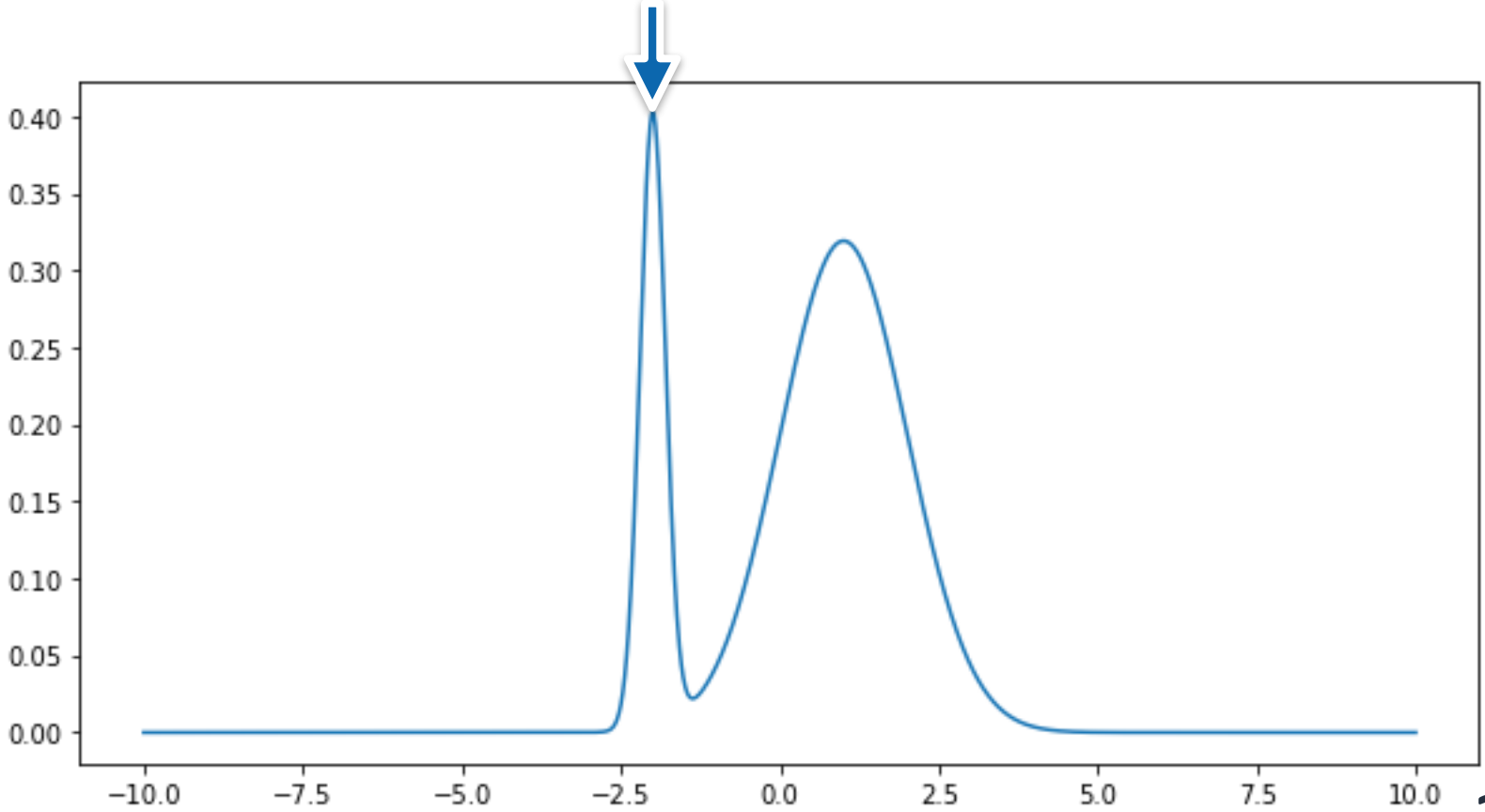
$$\partial_{\sigma^2}[\cdot] = \frac{n}{2\sigma^2} - \frac{1}{2\sigma^4}\sum_{i=1}^{n}(x_i - \mu)^2 = 0$$

$$\implies \sigma^2 = \frac{1}{n}\sum_{i=1}^{n}(x_i - \mu)^2$$

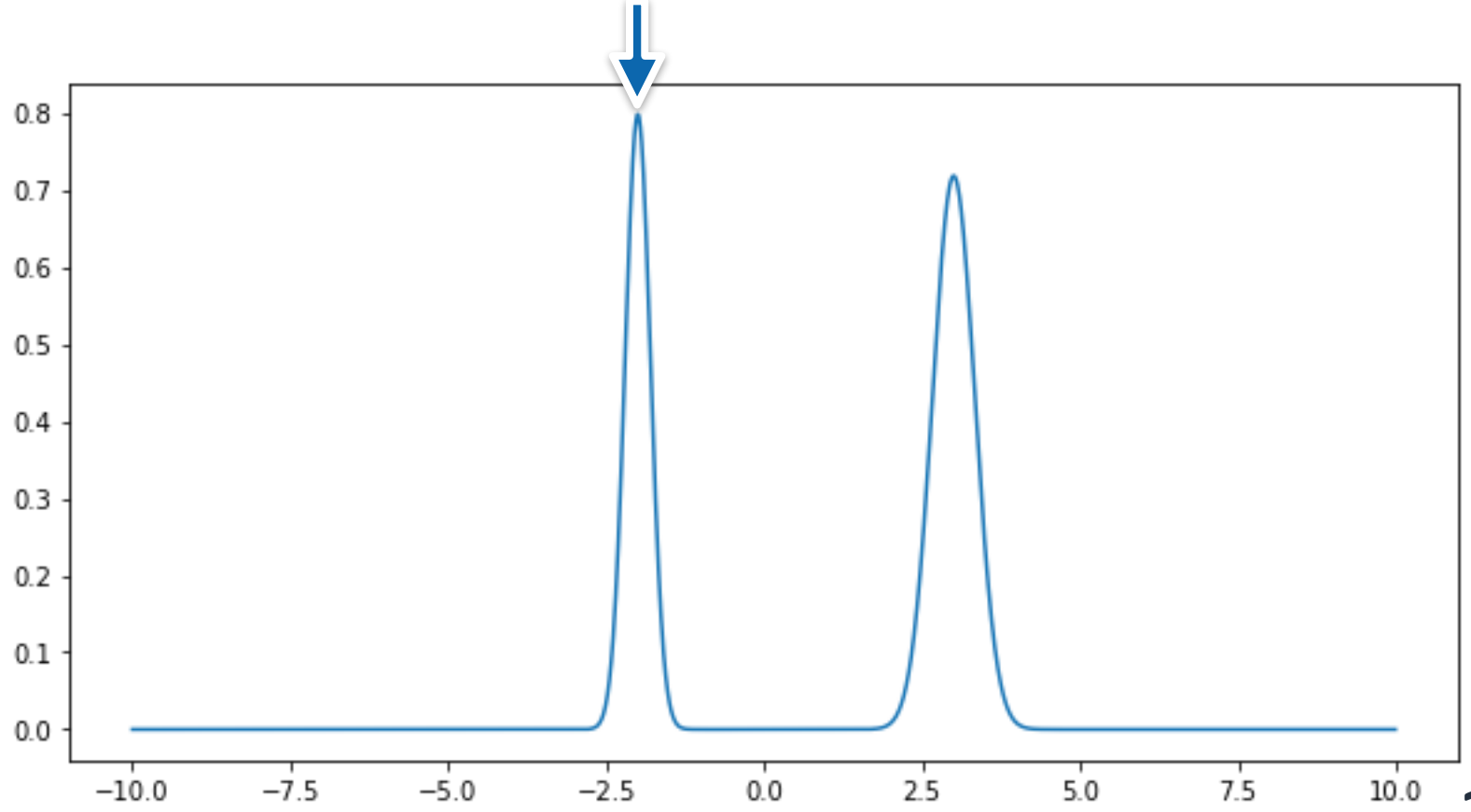
Maximum likelihood estimation



Maximum likelihood estimation



Maximum likelihood estimation



Maximum likelihood estimation

- Data - ‘student didn’t do homework’
- Possible parameters
  - **‘dog ate homework’**
  - **‘abducted by aliens’**
  - **‘too lazy’**
  - **‘sick grandmother’**

- **All parameters explain the data.**

Maximum likelihood estimation

- Data - ‘student didn’t do homework’
- Possible parameters
  - **‘dog ate homework’**
  - **‘abducted by aliens’**
  - **‘too lazy’**
  - **‘sick grandmother’**

Should we believe all?

- **All parameters explain the data.**

Maximum a posteriori estimation

- Posterior Probability

$$p(w \mid X) \propto p(X \mid w)p(w)$$

$$\text{hence} \quad -\log p(w \mid X) = -\log p(X \mid w) - \log p(w) + c$$

- Maximum a Posteriori Estimation

$$\underset{w}{\text{minimize}} \; -\log p(X; w) - \log p(w)$$

- No homework example
  p('no homework'|explanation) = 1 (all explanations work)

lazy student	grandma sick	dog ate it	alien abduction
0.8	0.19	0.0099	0.0001

Maximum a posteriori estimation

- Posterior Probability

$$p(w \mid X) \propto p(X \mid w) p(w)$$

penalty

$$\text{hence} \quad -\log p(w \mid X) = -\log p(X \mid w) - \log p(w) + c$$

- Maximum a Posteriori Estimation

$$\underset{w}{\text{minimize}} \; -\log p(X; w) - \log p(w)$$

- No homework example
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lazy student	grandma sick	dog ate it	alien abduction
0.8	0.19	0.0099	0.0001

What does this have to do with regression?

Regression

- Recall optimization problem

$$\underset{w}{\text{minimize}} \sum_{i=1}^{n} (y_i - f(x_i, w))^2 + \text{penalty}(w)$$

$-\log p(w)$

Does the model work?

Additive Gaussian Noise

- Data generation model

$$y_i = f(x_i, w) + \epsilon_i \text{ where } \epsilon_i \sim \mathcal{N}(0, \sigma^2)$$

- Gaussian Prior $p(w)$ hence $-\log p(w) = \frac{1}{2\bar{\sigma}^2}\|w\|^2 + \text{const}.$

Regression

- Maximum a posteriori

$$\underset{w}{\text{minimize}} \; -\log p(w \mid X, Y)$$

$$\iff \underset{w}{\text{minimize}} \; \frac{1}{2\sigma^2} \sum_{i=1}^{n} (y_i - f(x_i, w))^2 + \frac{1}{2\bar{\sigma}^2} \|w\|^2 + \text{const}.$$

$$\iff \underset{w}{\text{minimize}} \; \frac{1}{2n} \sum_{i=1}^{n} (y_i - f(x_i, w))^2 + \frac{\lambda}{2} \|w\|^2$$

Implement this


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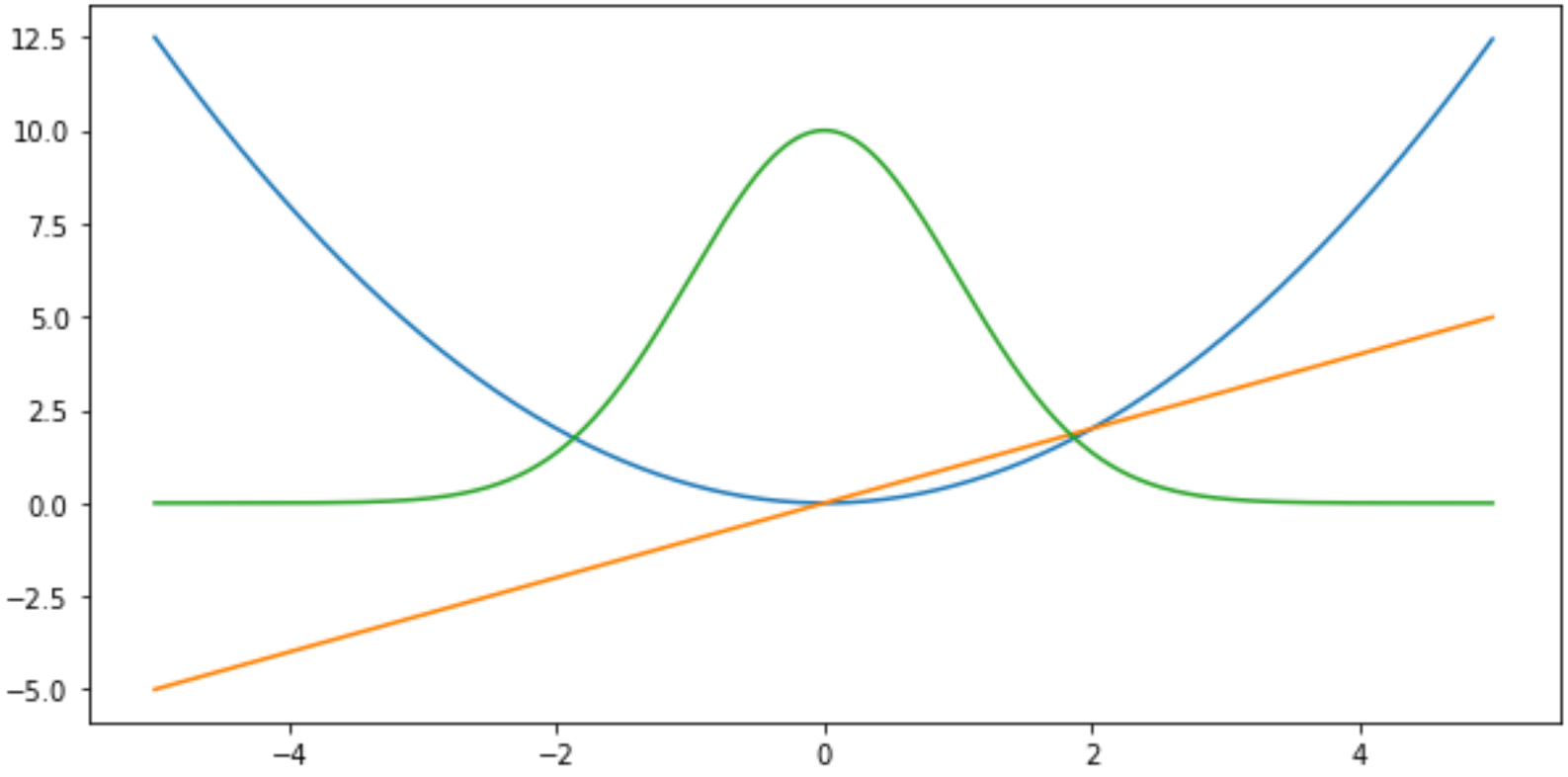
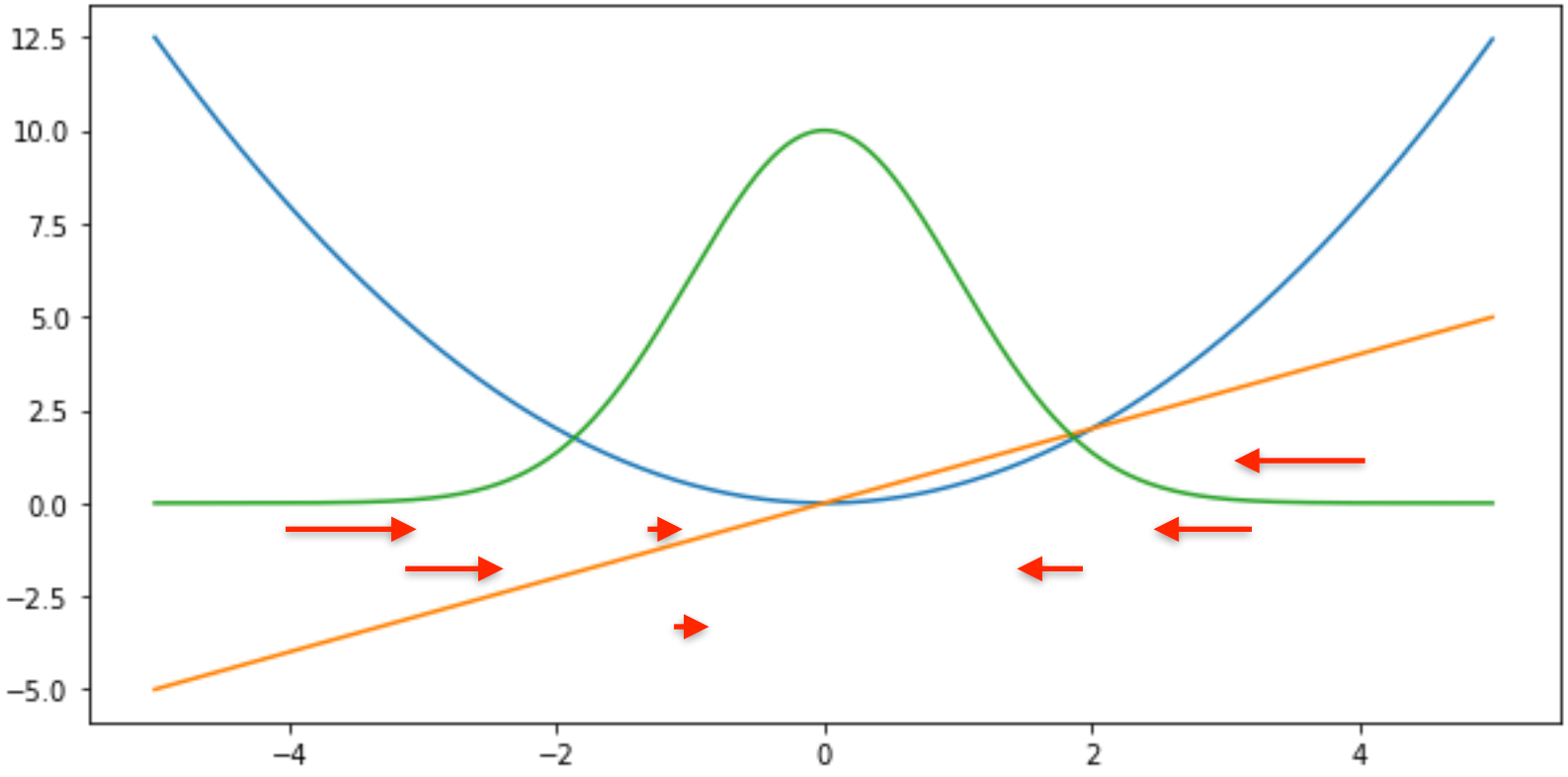
L2 Loss

$$l(y, y') = \frac{1}{2}(y - y')^2$$

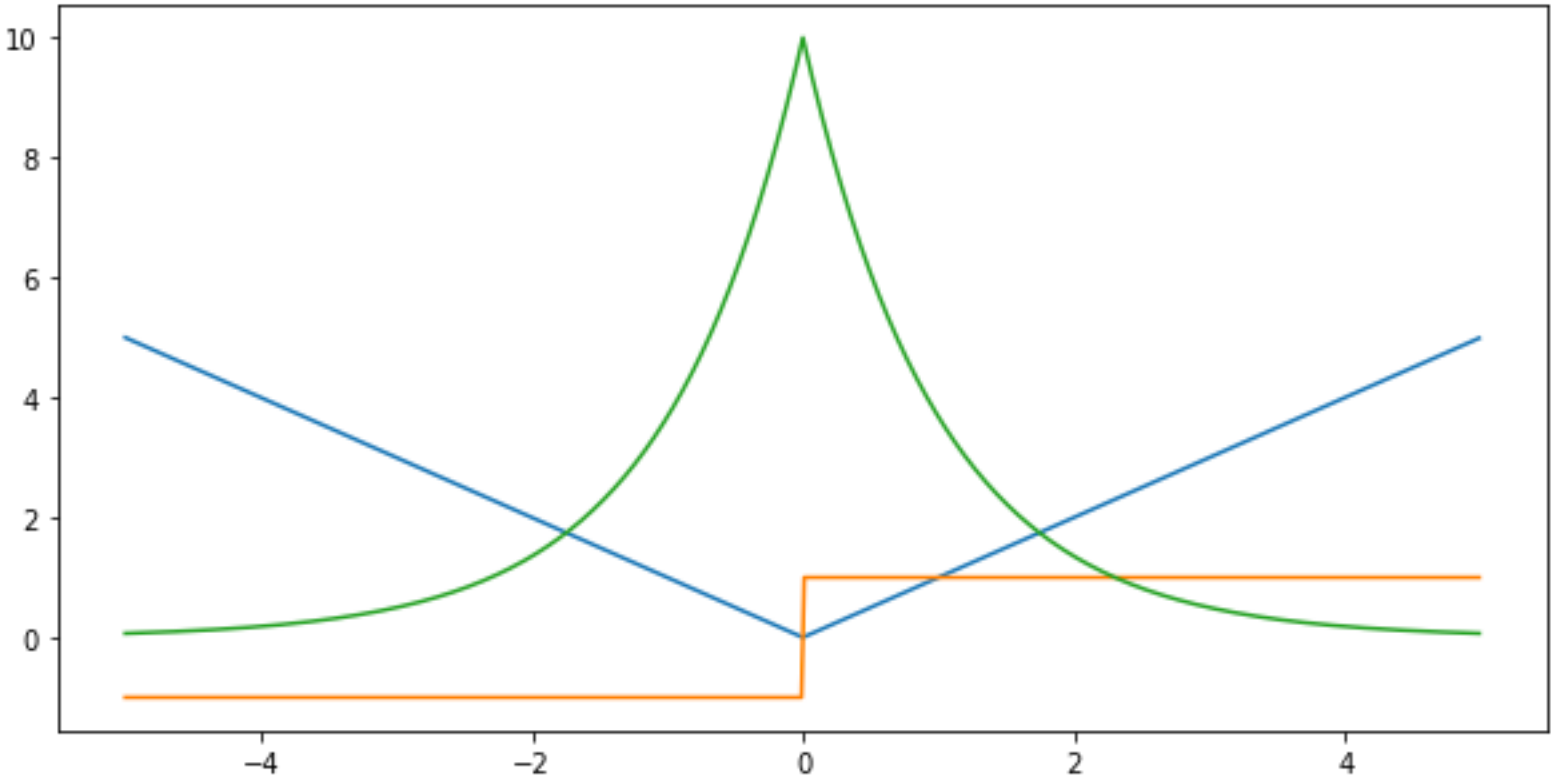


L2 Loss - mean $l(y, y') = \frac{1}{2}(y - y')^2$

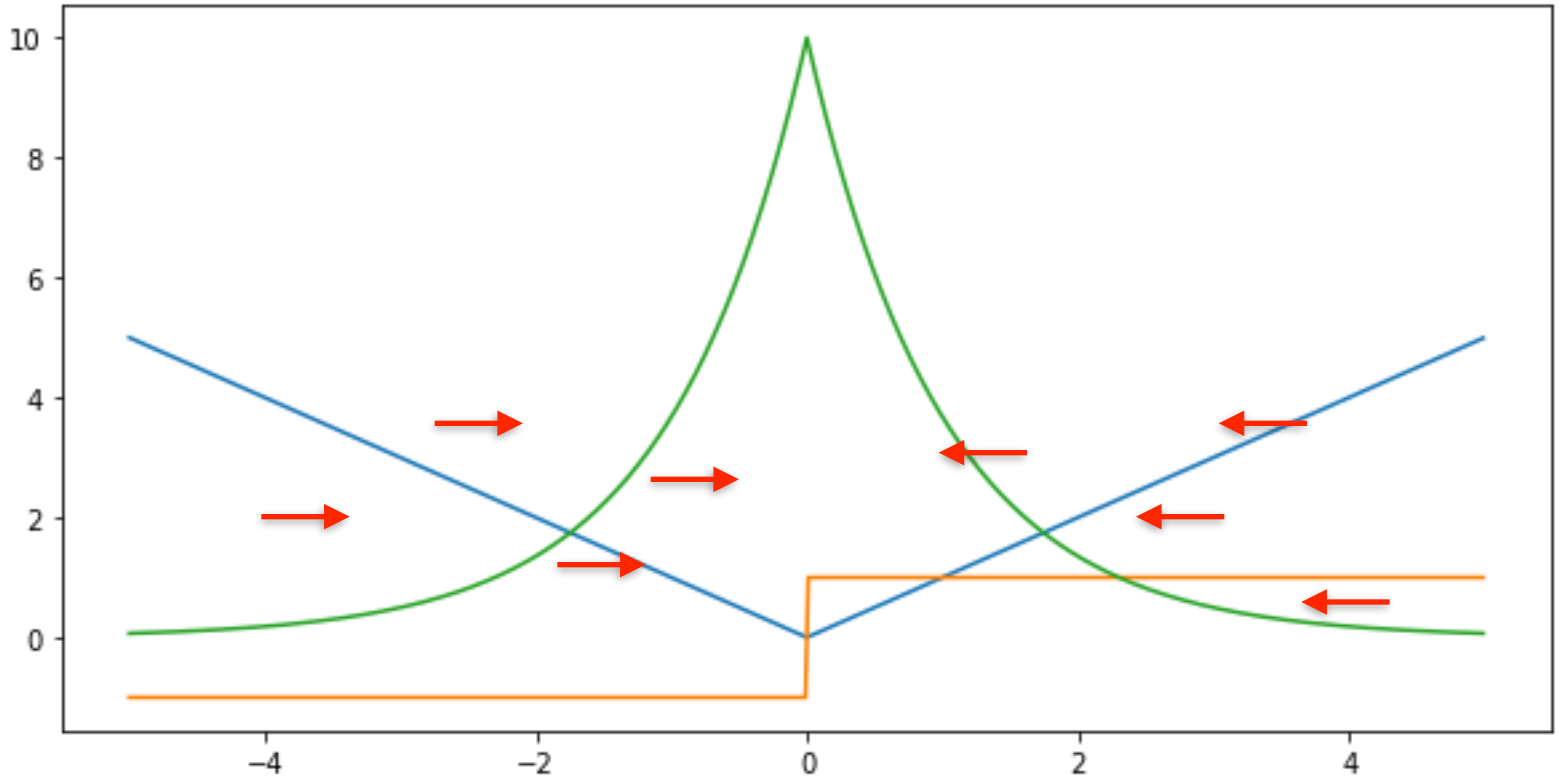


L1 Loss

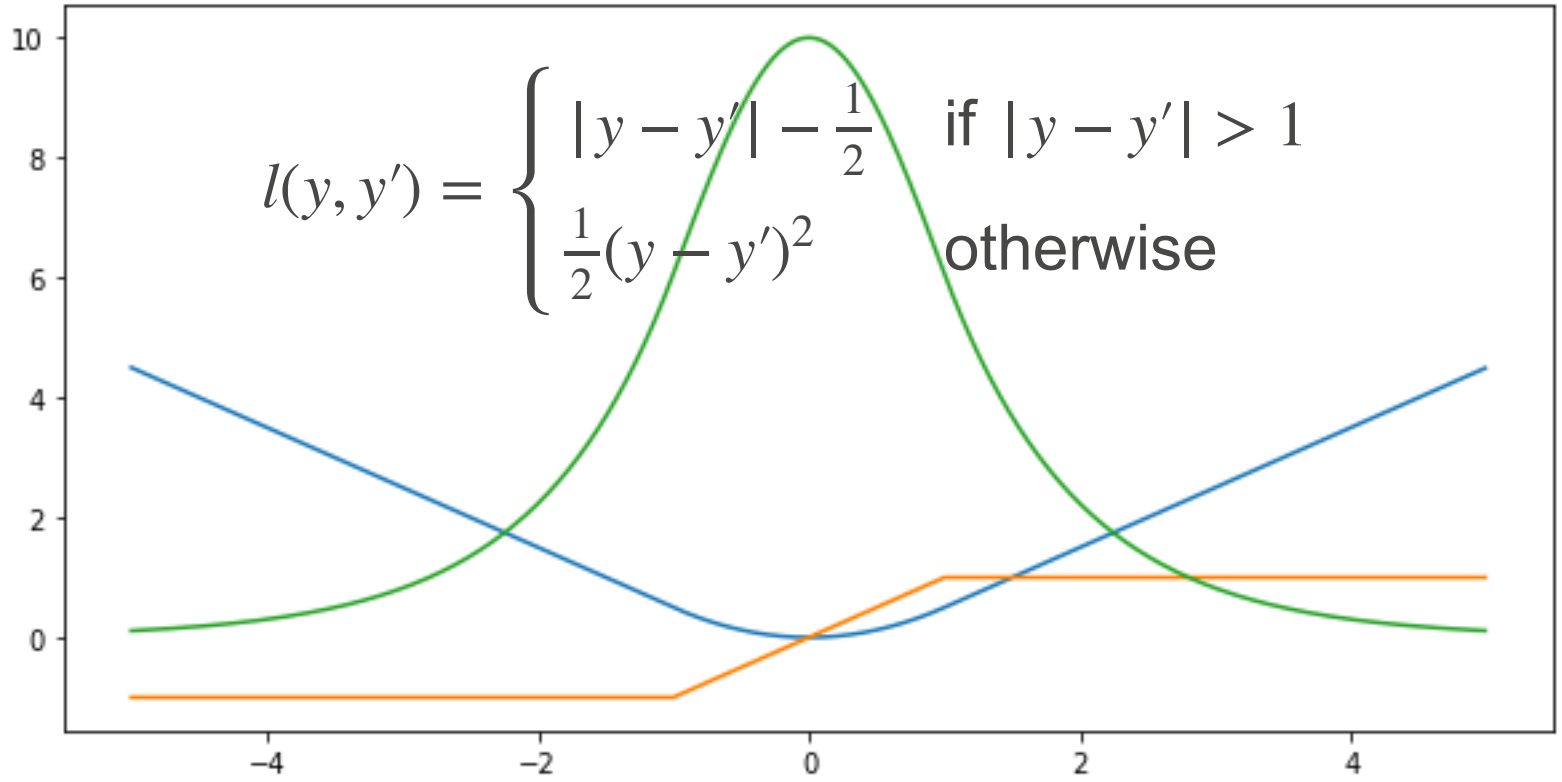
$$l(y, y') = |y - y'|$$



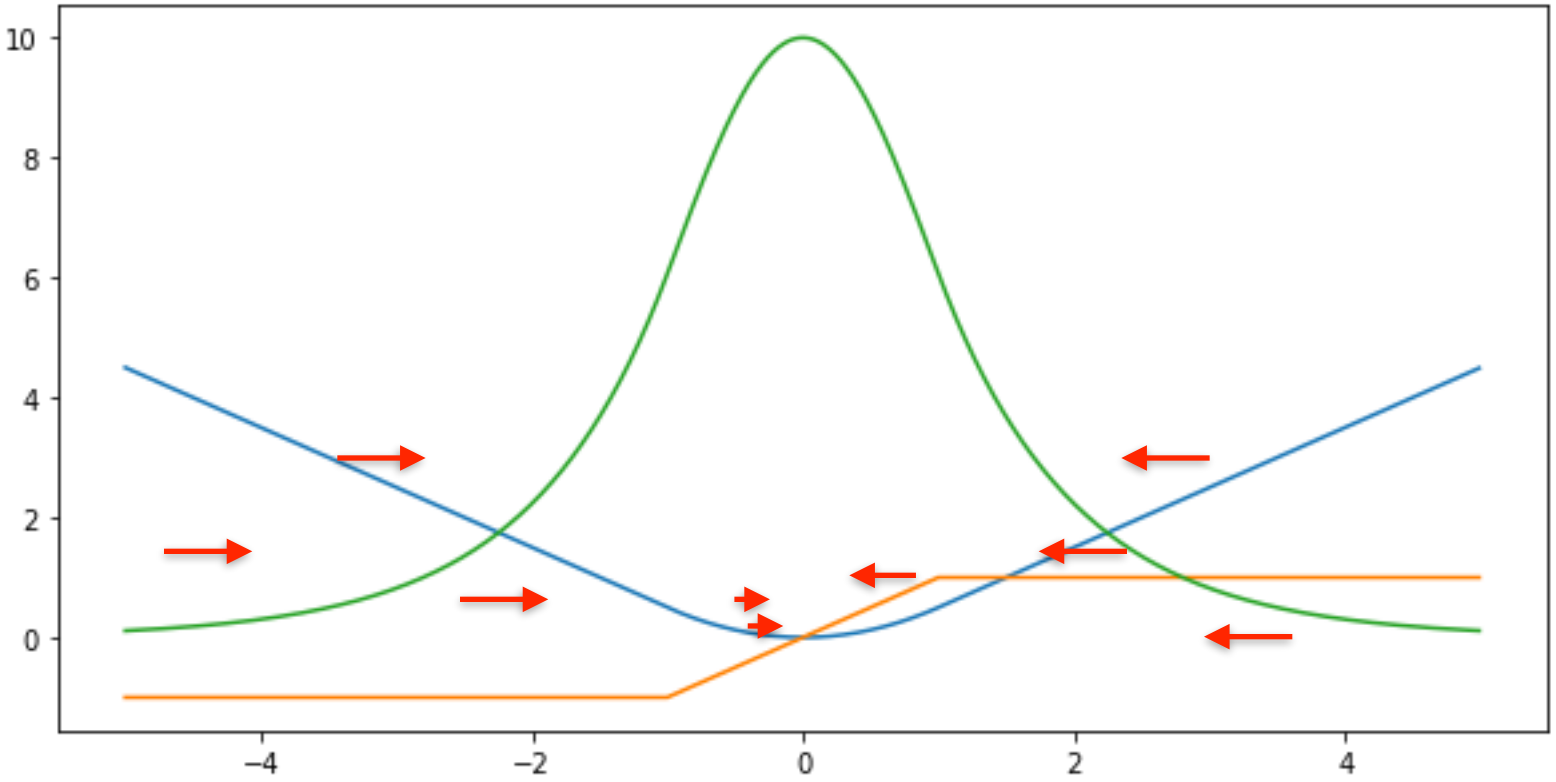
L1 Loss - median $l(y, y') = |y - y'|$



Huber's Robust Loss



Huber's Robust Loss - trimmed mean

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Regression vs. Classification

- Regression estimates a continuous value
- Classification predicts a discrete category

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MNIST: classify hand-written digits
(10 classes)

Regression vs. Classification

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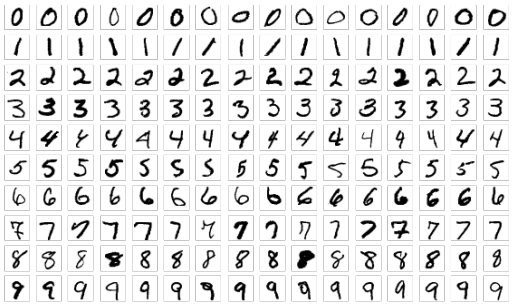
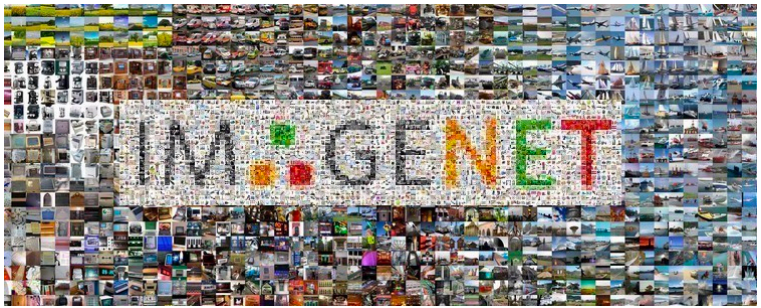
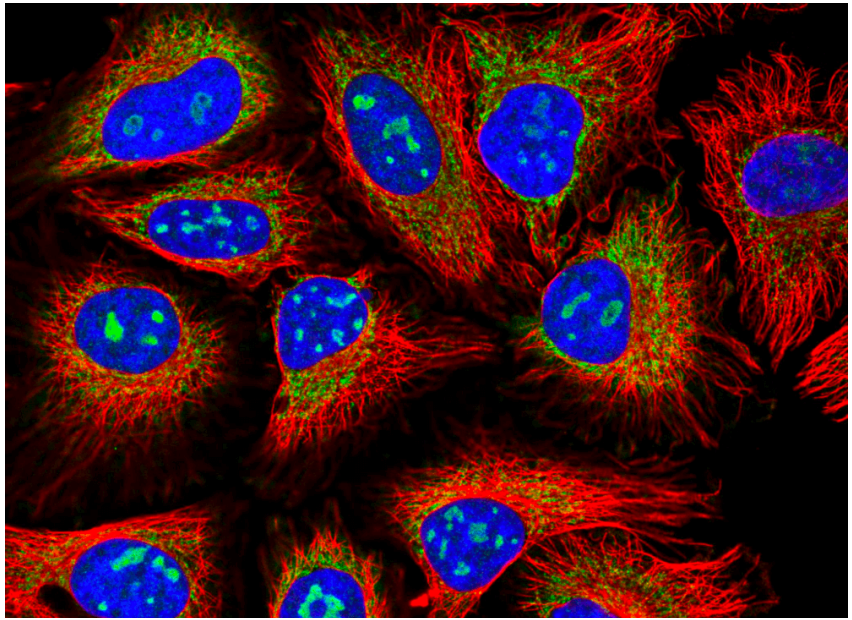


ImageNet: classify nature objects (1000 classes)



Classification Tasks at Kaggle

Classify human protein microscope images into 28 categories



```
0.  Nucleoplasm
1.  Nuclear membrane
2.  Nucleoli
3.  Nucleoli fibrillar
4.  Nuclear speckles
5.  Nuclear bodies
6.  Endoplasmic reticu
7.  Golgi apparatus
8.  Peroxisomes
9.  Endosomes
10.  Lysosomes
11.  Intermediate fila
12.  Actin filaments
13.  Focal adhesion si
14.  Microtubules
15.  Microtubule ends
16.  Cytokinetic bridg
```

https://www.kaggle.com/c/human-protein-atlas-image-classification

Classification Tasks at Kaggle

Classify malware into 9 categories

https://www.kaggle.com/c/malware-classification

Classification Tasks at Kaggle

Classify toxic Wikipedia comments into 7 categories

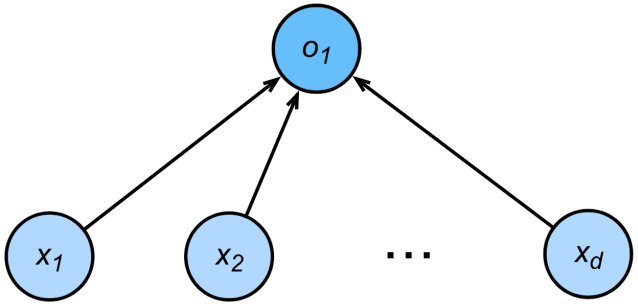
comment_text	toxic	severe_toxic	obsc
Explanation\nWhy the edits made under my usern...	0	0	0
D'aww! He matches this background colour I'm s...	0	0	0
Hey man, I'm really not trying to edit war. It...	0	0	0
"\nMore\nI can't make any real suggestions on ...	0	0	0
You, sir, are my hero. Any chance you remember...	0	0	0

https://www.kaggle.com/c/jigsaw-toxic-comment-classification-challenge

From Regression to Multi-class Classification

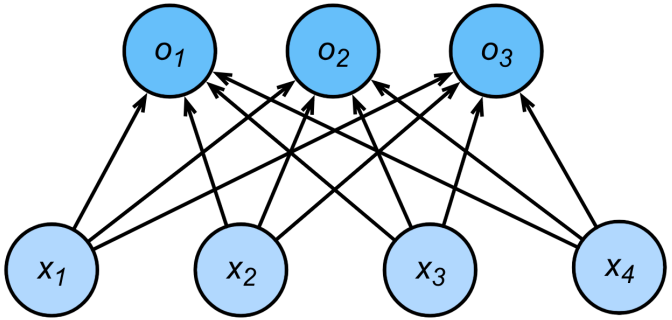
Regression

- Single continuous output
- Natural scale in $\mathbb{R}$
- Loss given e.g. in terms of difference $y - f(x)$



Classification

- Multiple classes, typically multiple outputs
- Score *should* reflect confidence …



From Regression to Multi-class Classification

Square Loss

- One hot encoding per class

$$\mathbf{y} = [y_1, y_2, \ldots, y_n]^\top$$

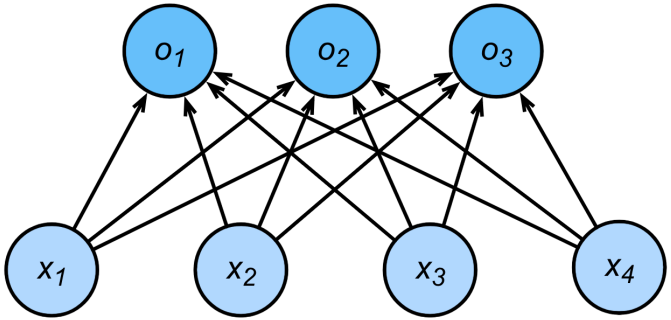
$$y_i = \begin{cases} 1 \text{ if } i = y \\ 0 \text{ otherwise} \end{cases}$$

- Train with squared loss
- Largest output wins

$$\hat{y} = \underset{i}{\operatorname{argmax}}\, o_i$$

Classification

- Multiple classes, typically multiple outputs
- Score *should* reflect confidence …



From Regression to Multi-class Classification

Uncalibrated Scale

- One output per class
- Largest output wins

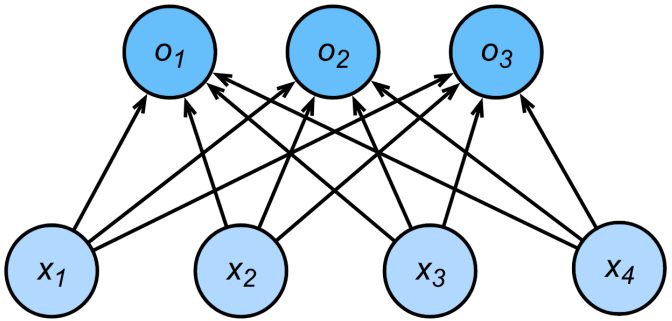
$$\hat{y} = \operatorname*{argmax}_i o_i$$

- Want that correct class is recognized confidently (**large margin**)

$$o_y - o_i \geq \Delta(y, i)$$

Classification

- Multiple classes, typically multiple outputs
- Score *should* reflect confidence …



From Regression to Multi-class Classification

Calibrated Scale

- Output matches probabilities (nonnegative, sums to 1)

$$p(y \mid o) = \text{softmax}(o)$$

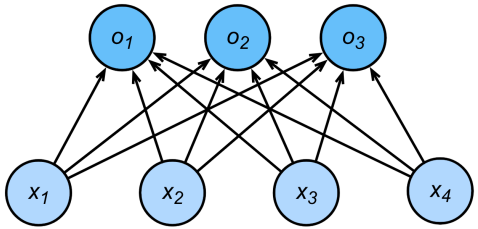
$$= \frac{\exp(o_y)}{\sum_i \exp(o_i)}$$

- Negative log-likelihood

$$-\log p(y \mid y) = \log \sum_i \exp(o_i) - o_y$$

Classification

- Multiple classes, typically multiple outputs
- Score *should* reflect confidence …



Softmax and Cross-Entropy Loss

- Negative log-likelihood (for given label y)

$$-\log p(y \mid o) = \log \sum_i \exp(o_i) - o_y$$

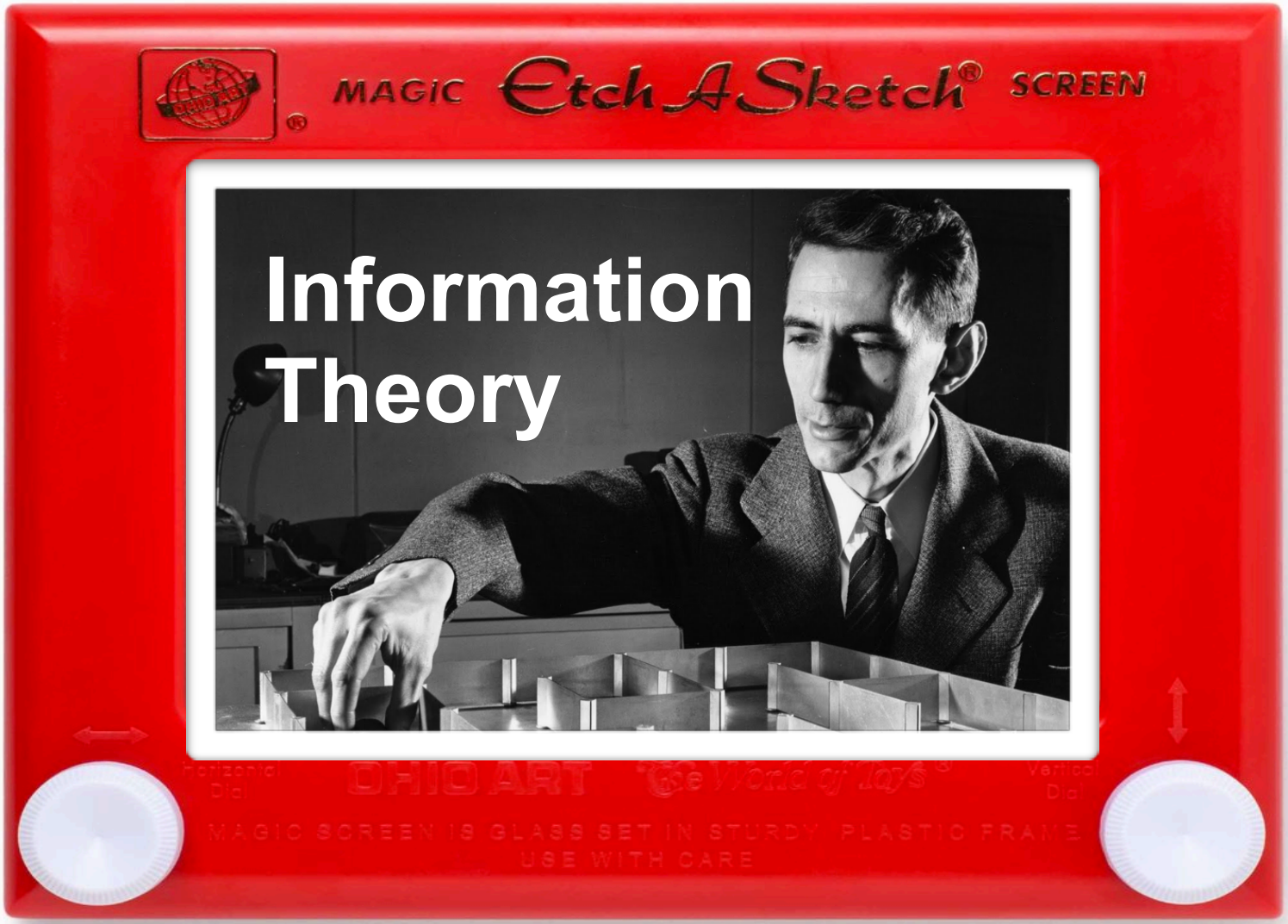
- Cross-Entropy Loss (for probability distribution y)

$$l(y, o) = \log \sum_i \exp(o_i) - y^\top o$$

- **Gradient**

Difference between true and estimated probability

$$\partial_o l(y, o) = \frac{\exp(o)}{\sum_i \exp(o_i)} - y$$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Entropy

- Data source producing observations $x_1 \ldots x_n$
- **How much 'information' is in this source?**
  - Tossing a fair coin - at each step the surprise is whether it's heads or tails
  - Rolling a fair dice - we have 1 out of 6 outcomes. This should be *more* surprising than the dice
  - Picture of a white wall vs. picture of a football stadium (the football stadium should have more information)
- **Measure is minimum number of bits needed**

Entropy

- Data source producing data $x_1 \dots x_n$ with probability $p(x)$
- **Definition**

$$H[p] = -\sum_j p_j \log p_j$$

- **Coding theorem**
  Entropy is lower bound on bits (or rather nats - base e)

$$2^a = e^b \text{ hence } a \log 2 = b \text{ hence bits} = \frac{H[p]}{\log 2}$$

- Entropy is concave since $p \log p$ is convex

$$H[\lambda p + (1-\lambda)q] \geq \lambda H[p] + (1-\lambda)H[q]$$

Entropy (binary form)

- Fair coin (p = 0.5)

$$H[p] = -0.5 \cdot \log_2 0.5 - 0.5 \cdot \log_2 0.5 = 1 \text{ bit}$$

- Biased coin (p = 0.9)

$$H[p] = -0.9 \cdot \log_2 0.9 - 0.1 \cdot \log_2 0.1 = 0.47 \text{ bit}$$

- Dungeons and Dragons (20-sided dice)

$$H[p] = -\log_2 \frac{1}{20} = 4.32 \text{ bit}$$

Kraft Inequality

- **Prefix Code**
  - Map x to code c(x) with length l(x)
  - No c(x) is the prefix for any c(x'),
    e.g. STAT and STATISTICS cannot both be codewords

$$\begin{cases} a \to 0 \\ b \to 01 \\ c \to 011 \\ d \to 0111 \end{cases} \qquad \begin{cases} a \to 0 \\ b \to 10 \\ c \to 110 \\ d \to 111 \end{cases}$$

Kraft Inequality

- **Inequality**

$$1 \geq \sum_x 2^{-l(x)} \text{ if and only if prefix code}$$

- **Proof**

  Generate random string. Probability that it's a codeword

$$1 \geq \Pr(\text{collision}) = \sum_x \Pr(x \text{ is a hit}) = \sum_x 2^{-l(x)}$$

  For converse explicitly construct prefix code recursively

  - Pick set of {x} with smallest l(x) and generate code
  - Use leftovers and break them up into sets of weight $2^{-l(x)}$
  - Give each of them prefix and rescale by $2^{l(x)}$

Kraft Inequality

- Forward part

$$\begin{cases} a \to 0 & \frac{1}{2} \\ b \to 10 & \frac{1}{4} \\ c \to 110 & \frac{1}{8} \\ d \to 11110 & \frac{1}{32} \end{cases}$$

110001010

collision

- Backwards part lengths (1, 2, 3, 5)
- Pick 1
  - Use code ‘0’ for it
  - Use prefix ‘1’ for the rest
  - Remaining set is (1, 2, 4)
  - Pick 1
    - Use code ‘0’ for it (thus ’10’)
    - Use prefix ‘1’ for the rest
    - Remaining set is (1,3)

Coding Theorem Proof

- Entropy is lower bound on number of bits

$$\text{bits}(p) = \frac{H[p]}{\log 2} = -\sum_j p_j \log_2 p_j$$

- Generate prefix code with length $l(x) = \lceil \log_2 p(x) \rceil$
  This is within 1 bit of optimal code
  Kraft inequality shows that such a thing exists.

$$\sum_x 2^{-\lceil -\log_2 p(x) \rceil} \leq \sum_x 2^{\log_2 p(x)} = \sum_x p(x) = 1$$

- Combine data in k-tuples to encode (within 1/k bit of optimal)
- Optimality proof via KL divergence (omitted here)

Kullback-Leibler Divergence

- Distance between distributions (e.g. truth & estimate)

  Number of extra bits when using the wrong code

$$D[p\|q] = \int dp(x) \log \frac{p(x)}{q(x)} = \int dp(x) \left[-\log q(x)\right] - \left[-\log p(x)\right]$$

Inefficient bits

Optimal bits

- Nonnegativity of KL Divergence

$$D[p\|p] = \int dp(x) \log \frac{p(x)}{p(x)} = 0$$

Jensen Inequality

$$D[p\|q] = -\int dp(x) \log \frac{q(x)}{p(x)} \geq -\log \int dp(x) \frac{q(x)}{p(x)} = 0$$

Back to the Cross Entropy Loss

- Cross entropy loss

$$l(y, o) = \log \sum_i \exp(o_i) - y^\top o$$

- Kullback Leiber divergence

$$D(\text{softmax}(o) \| q) = \sum_i q_i \log q_i - q_i \log \text{softmax}(o)_i$$

$$= -H[q] + \log \sum_i \exp(o_i) - \sum_i q_i o_i$$

Independent of o

Extended Reading

- Cover and Thomas (Elements of Information Theory)
  dl.acm.org/citation.cfm?id=1146355
- Information theory course (Entropy primer)
  slpl.cse.nsysu.edu.tw/cpchen/courses/ita/l1_entropy.pdf
- David MacKay (Information Theory and Learning)
  www.inference.org.uk/itprnn/book.html

- Conditional Entropy, Mutual Information, Exponential Families, Maximum Entropy Estimation

Summary

- **Maximum Likelihood**
  - Gauss and means
  - More loss functions (l_1 loss, trimmed mean)
  - Regression revisited
- **Classification**
  - Computing discrete probabilities
  - Likelihood and loss functions
- **Information Theory**