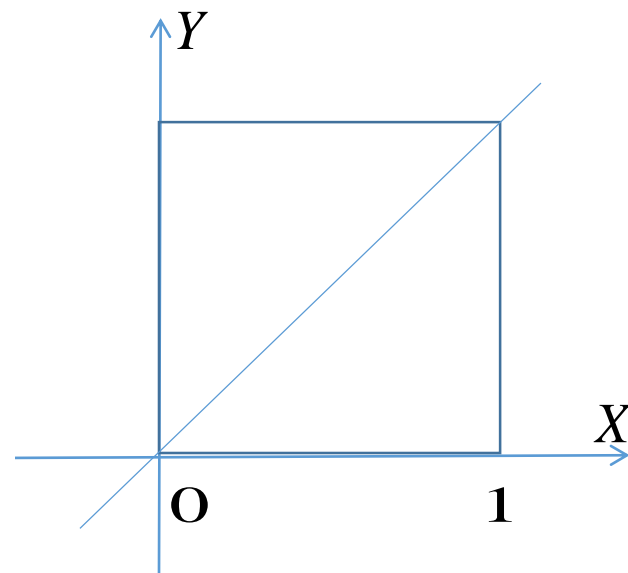


第四章 习题

1. ($P_{118} - 1$) 设随机变量 X 与 Y 独立, 并且都在区间 $(0,1)$ 上服从均匀分布, 试求随机变量 $Z = |X - Y|$ 的数学期望。

$$\begin{aligned}\text{解: } E(Z) &= E|X - Y| = \int_0^1 \int_0^1 |x - y| f(x, y) dx dy \\ &= \int_0^1 \int_0^y y - x dx dy + \int_0^1 \int_y^1 x - y dx dy \\ &= \int_0^1 \left(xy - \frac{1}{2} x^2 \right)_0^y dy + \int_0^1 \left(\frac{1}{2} x^2 - xy \right)_y^1 dy \\ &= \int_0^1 \frac{1}{2} y^2 dy + \int_0^1 \frac{1}{2} - \frac{1}{2} y^2 - y dy = \frac{1}{3}\end{aligned}$$



- 2. ($P_{118} - 2$) 设随机变量 X 与 Y 独立, 且都服从均数为0, 方差为1/2的正态分布, 求随机变量 $|X - Y|$ 的方差。

- 解: 令 $Z = X - Y$, 则 $Z \sim N(0,1)$.

$$D|X - Y| = D|Z| = E|Z|^2 - (E|Z|)^2 = EZ^2 - (E|Z|)^2 = 1 - (E|Z|)^2$$

$$E|Z| = \int_{-\infty}^{+\infty} |z| \varphi(z) dz = \int_{-\infty}^{+\infty} |z| \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz = 2 \int_0^{+\infty} z \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz$$

$$= \int_0^{+\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} dz^2 = -\frac{2}{\sqrt{2\pi}} e^{-\frac{z^2}{2}} \Big|_0^{+\infty} = \sqrt{\frac{2}{\pi}}$$

$$D|X - Y| = 1 - (E|Z|)^2 = 1 - \frac{2}{\pi}$$

• 3. ($P_{118} - 3$) 将一枚硬币重复掷 n 次, 以 X 和 Y 分别表示正面向上和反面向上的次数, 试求 X 和 Y 的相关系数 ρ_{XY} .

• 解: $X \sim B\left(n, \frac{1}{2}\right); \quad Y \sim B\left(n, \frac{1}{2}\right); \quad X + Y = n$

$$\begin{aligned} \text{Cov}(X, Y) &= \text{Cov}(X, n - X) = \text{Cov}(X, n) - \text{Cov}(X, X) \\ &= -D(X) = -\frac{n}{4} \end{aligned}$$

$$\rho_{XY} = \frac{\text{Cov}(X, Y)}{\sqrt{D(X)}\sqrt{D(Y)}} = \frac{-\frac{n}{4}}{\sqrt{\frac{n}{4}}\sqrt{\frac{n}{4}}} = -1$$

• 4. ($P_{118} - 5$) 对于事件 A 和 B , 满足 $0 < P(A) < 1$, $0 < P(B) < 1$, 定义

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$$X = \begin{cases} 0 & A \text{ 不发生} \\ 1 & A \text{ 发生} \end{cases} \quad Y = \begin{cases} 0 & B \text{ 不发生} \\ 1 & B \text{ 发生} \end{cases} \quad \text{则 } X \sim \begin{pmatrix} 0 & 1 \\ 1 - P(A) & P(A) \end{pmatrix}$$

• 证明随机变量 X 与 Y 独立的充分必要条件是 X 与 Y 不相关。

• 证明: $X \sim B\{1, P(A)\}$, $Y \sim B\{1, P(B)\}$, $XY \sim B\{1, P(AB)\}$

• $EX = P(X = 1) = P(A)$, $EY = P(Y = 1) = P(B)$

$$EXY = P(X = 1, Y = 1) = P(AB)$$

- X 与 Y 独立，必有 X 与 Y 不相关。 反之，若 X 与 Y 不相关：

- 则 $Cov(X, Y) = EXY - EXEY = P(AB) - P(A)P(B) = 0$

- $P(AB) = P(A)P(B)$ A 与 B 独立

$$P(\overline{A}B) = P(\overline{A})P(B); \quad P(A\overline{B}) = P(A)P(\overline{B}); \quad P(\overline{A}\overline{B}) = P(\overline{A})P(\overline{B})$$

$X Y$	0	1	
0	$P(\overline{A}\overline{B})$	$P(\overline{A}B)$	$P(\overline{A})$
1	$P(A\overline{B})$	$P(AB)$	$P(A)$
	$P(\overline{B})$	$P(B)$	



$X Y$	0	1	
0	$P(\overline{A})P(\overline{B})$	$P(\overline{A})P(B)$	$P(\overline{A})$
1	$P(A)P(\overline{B})$	$P(A)P(B)$	$P(A)$
	$P(\overline{B})$	$P(B)$	

• 5. ($P_{118} - 6$) 设 X 为只取非负整数值的离散型随机变量,

• 试证明: $EX = \sum_{n=1}^{+\infty} P(X \geq n)$

• 证明: $E(X) = \sum_{n=0}^{+\infty} nP(X = n)$ (设 $P(X = n) = p_n, n = 0, 1, 2, \dots$)

$$\begin{aligned} &= 0 \times P(X = 0) + 1 \times P(X = 1) + \dots + nP(X = n) + \dots \\ &= P(X = 1) + \{P(X = 2) + P(X = 2)\} \\ &\quad + \{P(X = n) + P(X = n) + \dots + P(X = n)\} + \dots \\ &= \{P(X = 1) + P(X = 2) + \dots\} + \{P(X = 2) + P(X = 3) + \dots\} \\ &\quad + \{P(X = n) + P(X = n + 1) + \dots\} + \dots = \sum_{n=1}^{+\infty} P(X \geq n) \end{aligned}$$

- 6. ($P_{118} - 7$) 对于两个随机变量 X 与 Y ，若 EX^2, EY^2 都存在，证明

$$(EXY)^2 \leq EX^2 EY^2$$

- 证明：对任意常数 c ，考虑

$$E(cX + Y)^2 = c^2 EX^2 + 2cEXY + EY^2 \geq 0$$

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$$\text{有 } (2EXY)^2 - 4 \times EX^2 EY^2 \leq 0$$

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$$\text{则 } (EXY)^2 \leq EX^2 EY^2$$

- 7. ($P_{118} - 8$) 设随机变量 X_1, X_2, \dots, X_n 相互独立, 且 $EX_i = \mu$, $DX_i = \sigma^2$,
- $i = 1, 2, \dots, n$. 令 $\bar{X} = \frac{1}{n} \sum_{i=1}^n X_i$, $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2$
- 及 $Y_i = X_i - \bar{X}$, $i = 1, 2, \dots, n$. (1) 求 $E\bar{X}, D\bar{X}$.

解: (1) $E\bar{X} = E\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n} \sum_{i=1}^n EX_i = \frac{1}{n} \times n\mu = \mu.$

$$D\bar{X} = D\left(\frac{1}{n} \sum_{i=1}^n X_i\right) = \frac{1}{n^2} \sum_{i=1}^n D(X_i) = \frac{1}{n^2} \times n\sigma^2 = \frac{\sigma^2}{n}$$

$$E\bar{X} = EX_i = \mu; \quad D\bar{X} = \frac{\sigma^2}{n}$$

• (2) 证明 $S^2 = \frac{1}{n-1} \left(\sum_{i=1}^n X_i^2 - n(\bar{X})^2 \right)$, 并求 ES^2

证明: $S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{n}{n-1} \left\{ \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2 \right\} = \frac{n}{n-1} B_2$

令 $B_2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X}_n)^2 = \frac{1}{n} \sum_{i=1}^n (X_i^2 - 2X_i \bar{X}_n + \bar{X}_n^2)$

$$= \frac{1}{n} \sum_{i=1}^n X_i^2 - 2\bar{X}_n \frac{1}{n} \sum_{i=1}^n X_i + \bar{X}_n^2 = \frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}_n^2$$

$$S^2 = \frac{n}{n-1} B_2 = \frac{n}{n-1} \left(\frac{1}{n} \sum_{i=1}^n X_i^2 - \bar{X}_n^2 \right) = \frac{1}{n-1} \left(\sum_{i=1}^n X_i^2 - n\bar{X}_n^2 \right)$$

$$\begin{aligned}
 \bullet \bullet EB_2 &= \frac{1}{n} \sum_{i=1}^n EX_i^2 - E(\bar{X}_n)^2 = \frac{1}{n} \sum_{i=1}^n (\sigma^2 + \mu^2) - \left(\frac{\sigma^2}{n} + \mu^2 \right) \\
 &= \sigma^2 + \mu^2 - \left(\frac{\sigma^2}{n} + \mu^2 \right) = \frac{n-1}{n} \sigma^2
 \end{aligned}$$

$$EX_i^2 = DX_i + E\bar{X} = \sigma^2 + \mu^2$$

$$E(\bar{X}_n)^2 = D\bar{X}_n + (E\bar{X}_n)^2 = \frac{\sigma^2}{n} + \mu^2$$

$$\text{即: } EB_2 = \frac{n-1}{n} \sigma^2$$

$$\text{而 } S^2 = \frac{1}{n-1} \sum_{i=1}^n (X_i - \bar{X})^2 = \frac{n}{n-1} B_2, \quad \text{则 } ES^2 = E\left(\frac{n}{n-1} B_2\right) = \sigma^2$$

• (3) 解:
$$\begin{aligned} \text{Cov}(X_i, \bar{X}) &= \text{Cov}\left(X_i, \frac{1}{n} \sum_{j=1}^n X_j\right) \quad i = 1, 2, \dots, n \\ &= \text{Cov}\left(X_i, \frac{1}{n} X_i\right) + \text{Cov}\left(X_i, \frac{1}{n} \sum_{\substack{j=1 \\ j \neq i}}^n X_j\right) = \frac{1}{n} \text{Cov}(X_i, X_i) \\ &= \frac{1}{n} DX_i = \frac{\sigma^2}{n} \end{aligned}$$

• (4) 解:
$$\begin{aligned} D(Y_i) &= D(X_i - \bar{X}) = D\left(\left(1 - \frac{1}{n}\right)X_i - \frac{1}{n} \sum_{\substack{j=1 \\ j \neq i}}^n X_j\right) \\ &= \left(1 - \frac{1}{n}\right)^2 DX_i + \frac{n-1}{n^2} DX_j = \frac{n-1}{n} \sigma^2 \end{aligned}$$

• (5) 解: $Cov(Y_i, Y_j) = Cov(X_i - \bar{X}_n, X_j - \bar{X}_n) \quad i \neq j \quad Y_i = X_i - \bar{X}_n$

$$= Cov(X_i, X_j) + Cov(\bar{X}_n, \bar{X}_n) - Cov(X_i, \bar{X}_n) - Cov(X_j, \bar{X}_n)$$

$$= 0 + D(\bar{X}_n) - \frac{\sigma^2}{n} - \frac{\sigma^2}{n} = -\frac{\sigma^2}{n}$$

• (6) 解: $D(Y_i - Y_j) = D\{(X_i - \bar{X}_n) - (X_j - \bar{X}_n)\} \quad i \neq j$

$$= D(X_i - X_j)$$

$$= D(X_i) + D(X_j) = 2\sigma^2$$