

## SUPPLEMENTAL MATERIAL

### A MATHEMATICAL ANALYSIS

In this section, we analyze the expected AAE for frequency estimation of BM. Then we define the “loading rate” to represent the load status of the data structure at a certain time. Additionally, we define “error rate” as the probability of identifying the frequency of an item that has never appeared as non-zero.

#### A.1 Average absolute error (AAE)

The calculation formula of AAE is defined as  $\frac{1}{|\Psi|} \sum_{(e_i \in \Psi)} |f_i - \tilde{f}_i|$ , where  $f_i$  is the real frequency of item  $e_i$ ,  $\tilde{f}_i$  is the estimated frequency, and the  $\Psi$  is the query set.

Suppose that BM records a data stream  $S$  with  $N$  items, and  $E$  is the item set ( $|E| = n$ ). Let  $e_i$  be the  $i^{th}$  item in  $E$ , whose actual frequency is  $f_i$ . Assume that there are  $B$  entries in each bucket of BM on average, and use  $fp()$  to refer to the  $fingerprnt()$  function. The final frequency estimation of item  $e_i$  is

$$\hat{f}_i = f_i - X_i + Y_i \quad (1)$$

where  $X_i$  is the decrement brought by the replacement strategy (when an item cannot find a vacancy in the bucket, it will decrement the smallest entry by 1, which will only lead to underestimation). And  $Y_i$  is the increment due to fingerprint collisions (which only leads to overestimation). The two error bounds of BitMatcher—the lower bound and the upper bound—result from  $X_i$  and  $Y_i$  respectively. We will calculate them separately.

**A.1.1 For  $E(X_i)$ .** We assume that the replacement strategy is “direct decrease by 1”, and no fingerprint collision occurs at this time, which means  $Y_i = 0$ . Since there are a total of  $N$  items in the data stream, and only  $2w$  buckets are used to accommodate them, on average, each bucket will have  $\frac{N}{2w}$  items inserted. And since there are  $B$  entries in each bucket on average to accommodate  $B$  types of items, and the frequency of these items increase sequentially in the bucket ( $entry_{1.cnt} \leq \dots \leq entry_{B.cnt}$ ), so the maximum number of decays for item  $e_i$  occurs in this case:  $entry_2 \sim entry_B$  are already occupied by other items,  $e_i$  is always in  $entry_1$ , and all the items coming in later are used to decrease  $e_i$ . So it follows that:  $E(X_i) \leq \max \# \text{ of decays} = \min(f_i, \frac{N}{2w} \times \frac{1}{B+1})$ .

**A.1.2 For  $E(Y_i)$ .** Assuming that  $e_i$  is already in  $\mathcal{A}_1[\tilde{h}_1(e_i)]$  (same as  $\mathcal{A}_2[\tilde{h}_2(e_i)]$ ), we introduce indicator variables  $I_{i,j}$ , ( $1 \leq i, j \leq n$ ) as:

$$I_{i,j} = \begin{cases} 1, & (i \neq j) \wedge (fp(e_i) = fp(e_j)) \wedge \\ & [\tilde{h}_1(e_i) = \tilde{h}_1(e_j) \vee \\ & \tilde{h}_2(e_i) = \tilde{h}_2(e_j)] \\ 0, & \text{else} \end{cases} \quad (2)$$

The last two conditions in the brackets ensure that  $e_i$  and  $e_j$  are mapped to at least one same bucket, and its probability can be written as

$$\begin{aligned} & Pr[\tilde{h}_1(e_i) = \tilde{h}_1(e_j) \vee \tilde{h}_2(e_i) = \tilde{h}_2(e_j)] \\ &= Pr[\tilde{h}_1(e_i) = \tilde{h}_1(e_j)] + Pr[\tilde{h}_2(e_i) = \tilde{h}_2(e_j)] \\ &\quad - Pr[\tilde{h}_1(e_i) = \tilde{h}_1(e_j) \wedge \tilde{h}_2(e_i) = \tilde{h}_2(e_j)] \\ &\leq \frac{1}{w} + \frac{1}{w} - \frac{1}{w^2} < \frac{2}{w} \end{aligned} \quad (3)$$

Here we assume  $hash(e_i)$  is much larger than  $w$ . Then we have:

$$\begin{aligned} E(I_{i,j}) &< Pr[fp(e_i) = fp(e_j)] \times \frac{2}{w} \\ &\leq \frac{1}{range(fingerprint)} \times \frac{2}{w} = \frac{1}{w2^{\mathcal{F}-1}} \end{aligned} \quad (4)$$

Here we assume the length of fingerprint is  $\mathcal{F}$ . And since given the two buckets mapped by  $e_i$  and  $e_j$  (and they have a shared bucket  $\mathcal{A}$ ),  $e_i$  has a probability of  $\frac{1}{2}$  being mapped into  $\mathcal{A}$ . So

$$E(Y_i) = E\left(\sum_{j=1}^n \frac{1}{2} I_{i,j} f_j\right) \approx \frac{1}{2} \sum_{j=1}^n f_j E(I_{i,j}) = \frac{N}{w2^{\mathcal{F}}} \quad (5)$$

**A.1.3 Experimental Results.** Next, we compare the theoretical and empirical AAE. We note that according to the definition of AAE and Eq. (1), we have  $AAE = |\hat{f}_i - f_i| = |Y_i - X_i|$ . For cold item, its  $E(X_i) \leq \min(f_i, \frac{N}{2w} \times \frac{1}{B+1}) = f_i$  is a very small value. For hot item, the upper bound of  $E(X_i)$  is actually too loose, because when a hot item enters a bucket, it will quickly grow into larger entry. It won't be affected by too many replacement strategies in the process. So  $E(X_i)$  of hot item is much smaller than  $\min(f_i, \frac{N}{2w} \times \frac{1}{B+1})$ . So to sum up,  $E(X_i)$  is always a negligibly small value, and we can only consider the effect of  $Y_i$  in the experiment, that is,  $AAE \approx |Y_i|$ .

We use the CAIDA and IMC dataset. For detailed information, please refer to Section 5.1.2. The experimental results are shown in Fig. 19. The abscissa is memory, the range is 0.1 ~ 1 MB (interval is 0.1) and 1 ~ 10 MB (interval is 1). The black line is the empirical AAE, and the red line is the theoretical expectation of AAE. The fingerprint length  $\mathcal{F}$  is 8 bits. We can find that the theoretical values fit fairly well.

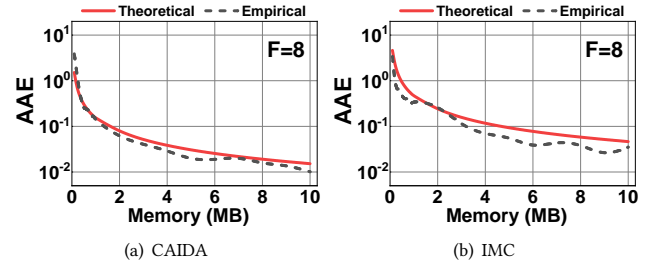


Figure 19: Empirical and theoretical AAE.

#### A.2 Loading rate

First we introduce the definition of loading rate.

**DEFINITION 2.** If the BM contains  $2w$  buckets, of which  $x$  buckets are full (no empty entries), then the **loading rate** at this time is  $\frac{x}{2w}$ .

Then we have the following theorem:

**THEOREM 3.** Assuming that the BM has  $2w$  buckets, each bucket has an average of  $B$  entries, Each array has  $k$  candidate buckets ( $k = 1$

in this paper), and  $n$  types of items arrive at this time. Then for the loading rate  $LR$  at this time, we have:

$$LR \geq 1 - \sum_{l=0}^n \left\{ \binom{n}{l} \left( \frac{1}{2w} \right)^l \left( 1 - \frac{k}{w} \right)^{n-l} \left[ \sum_{i=0}^{\min(l, B-1)} \binom{l}{i} (2k-1)^{l-i} \right] \right\}$$

PROOF. We connect BM's  $array_1$  after  $array_2$ . There are  $2k$  hash functions in total at this time (the range becomes  $0 \sim (2w-1)$ , and the insert operation is the same as before). We define the loading rate in this case as  $LR'$ . We first prove that  $LR \geq LR'$ .

We consider the following: when BM processes a data stream, if one hash method is easier to find empty buckets (with at least 1 empty entry in it) than another (assuming the total number of buckets is the same), then the loading rate will increase faster at this time. Because if the hash function maps to a full bucket, it will only perform the replacement strategy without any contribution to the loading rate. So we only need to compare the probability of finding empty buckets between these two hashing methods to know which one has a higher loading rate.

We define the probability that the original BM and the connected BM find an empty bucket at a certain moment as  $p_{ori}^{LR}$  and  $p_{new}^{LR}$ . Assume that there are  $x_1$  full buckets in  $array_1$  and  $x_2$  full buckets in  $array_2$  at this time. We have:

$$\begin{aligned} p_{ori}^{LR} &= 1 - P(\text{all } 2k \text{ mapped buckets are full}) \\ &= 1 - \left( \frac{x_1}{w} \right)^k \times \left( \frac{x_2}{w} \right)^k \\ &= 1 - \frac{(x_1 x_2)^k}{w^{2k}} \end{aligned} \quad (6)$$

$$\begin{aligned} p_{new}^{LR} &= 1 - P(\text{all } 2k \text{ mapped buckets are full}) \\ &= 1 - \left( \frac{x_1 + x_2}{2w} \right)^{2k} \\ &= 1 - \frac{\left( \frac{x_1 + x_2}{2} \right)^{2k}}{w^{2k}} \end{aligned} \quad (7)$$

Since  $\frac{x_1 + x_2}{2} \geq (x_1 x_2)^{\frac{1}{2}}$ , we have  $p_{ori}^{LR} \geq p_{new}^{LR}$ , so  $LR \geq LR'$ .

Next we calculate  $LR'$ . Note that the loading rate at a certain moment ( $n$  types of items have arrived) is actually the probability that a bucket is full at this time, which is  $1 - P(\text{it has a vacancy})$ . Assuming that there is no fingerprint collision at this time, there are  $B$  entries in each bucket on average. When we look at a bucket, we consider the probability of the following event: the candidate bucket for  $l$  of  $n$  items is this bucket, and this bucket currently has  $i$  entries occupied. Let this probability be  $P_i^l$ , with a little knowledge of combinatorics, we have:

$$P_i^l = \binom{n}{l} \left( \frac{2k}{2w} \right)^l \left( 1 - \frac{2k}{2w} \right)^{n-l} \times \binom{l}{i} \left( \frac{1}{2k} \right)^i \left( \frac{2k-1}{2k} \right)^{l-i} \quad (8)$$

Sum it over  $l$  and  $i$  to get:  $P(\text{it has a vacancy}) = \sum_{l,i} P_i^l$ . Therefore, we have  $LR \geq LR' = 1 - \sum_{l,i} P_i^l$ , and the result in Theorem 3 can be obtained by substituting Eq. 8 into it.  $\square$

**THEOREM 4.** Under the conditions of Theorem 3, and we consider the fingerprint effect. Assuming that the theoretical upper bound of the loading rate is  $LR_{opt}$ , we have:

$$LR_{opt} = \frac{n}{2wB^2} \left[ \sum_{i=0}^{B-1} \frac{2^F - i}{2^F} \right] \quad (9)$$

PROOF. The proof is omitted.  $\square$

We use the CAIDA dataset to validate our conclusions, which contains approximately 165K different kinds of items. As shown in Fig. 20, we show the theoretical optimal value and theoretical lower bound of the loading rate of BM, and the empirical value is recorded every 5K items. The experimental results validate our theory, and the empirical value is relatively close to the ideal value.

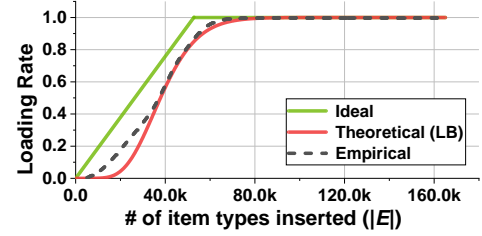


Figure 20: Empirical and theoretical LR.

### A.3 Error rate

After processing a data stream, if we query the frequency of an item that has not yet appeared, we do not expect it to return non-zero results. Of course, for a probabilistic algorithm like sketch, this is inevitable when memory is tight. But we expect this probability to be as small as possible. Since there are a number of applications such as blacklist (when some items are blocked) or path verification which would be impacted by it. We return 0 when we do not find the corresponding fingerprint and there is an empty entry in two candidate buckets. And we return the minimum value of two buckets if there is no empty entry. This raises the question: will we identify more non-occurring items as non-zero compared to normal CM-Sketch?

First we introduce the definition of error rate.

**DEFINITION 3.** After processing a data stream, the probability that an algorithm identify the frequency of a non-occurring item as non-zero is defined as **error rate**.

Our next step is to show that this error rate of BM is smaller than CM. Suppose  $d$  and  $w_{CM}$  are the depth and width of CM sketch respectively,  $w_{BM}$  is the width of BitMatcher,  $n$  represents the type of item in the data stream (i.e.,  $|E|$ ),  $M$  stands for memory usage (unit is MB).

**A.3.1 For CM Sketch.** After inserting all items in stream, we look at a counter in a row, which has probability  $P_{1\_cnt\_0}$  to be zero. Here  $P_{1\_cnt\_0} = Pr[n \text{ kinds of item are mapped to the remaining } (w_{cm} - 1) \text{ counters}] = \left( 1 - \frac{1}{w_{CM}} \right)^n$ . For a item that does not appear, we define  $p_{CM}^{ER}$  as the error rate of CM, we have:

$$\left\{ \begin{aligned} p_{CM}^{ER} &= (1 - P_{1\_cnt\_0})^d = \left( 1 - \left( 1 - \frac{1}{w_{CM}} \right)^n \right)^d \\ d \times w_{CM} \times 2 &= M \times 1024 \times 1024 \end{aligned} \right. \quad (10)$$

Note that in the paper we take the counter size of CM-Sketch as 16 bits (2 bytes).

**A.3.2 For BitMatcher.** After inserting all items in stream, we also define  $P_{BM}^{ER}$  as the error rate of BM. Let's first look at a bucket in  $array_1$ , the probability of which is full is defined as  $P_{1\_occ}$ , we have: (we assume there are no fingerprint collisions)

$$\begin{aligned} P_{1\_occ} &= 1 - \sum_{i=0}^{B-1} Pr[i \text{ entries occupied}] \\ &= 1 - \sum_{i=0}^{B-1} \binom{n}{i} \left(\frac{1}{w_{BM}}\right)^i \left(1 - \frac{1}{w_{BM}}\right)^{n-i} \end{aligned}$$

Since  $n$  and  $w_{cc}$  are quite large, we have:

$$P_{1\_occ} \approx 1 - \left(1 - \frac{1}{w_{BM}}\right)^n \sum_{i=0}^{B-1} \frac{n^i}{i! w_{BM}^i} \quad (11)$$

Furthermore, we have:

$$P_{BM}^{ER} = Pr[A_1[h_1(e)] \text{ is full} \wedge A_2[h_2(e)] \text{ is full}] \quad (12)$$

We then want to compute an upper bound on  $P_{BM}^{ER}$ . We give an instance of BM that has been inserted, assuming that there are  $x_1$  buckets in  $array_1$  are full, and  $x_2$  buckets in  $array_2$  are full. At this time, according to our algorithm,  $P_{BM}^{ER} = \frac{x_1 x_2}{w_{BM}^2}$ . And, if we connect  $array_2$  after  $array_1$ , and treat it as a whole array, we choose a bucket, and the probability that it is full is defined as  $P'_{BM} = \frac{x_1 + x_2}{2w_{BM}}$ . Note that according to the definition of  $P'_{cc}$ , it is actually  $P_{1\_occ}(2w_{BM})$ . We then have:

$$\begin{aligned} P_{BM}^{ER} - P'_{BM} &= \frac{x_1 x_2}{w_{BM}^2} - \frac{x_1 + x_2}{2w_{BM}} \\ &= \frac{2x_1 x_2 - x_1 w_{BM} - x_2 w_{BM}}{2w_{BM}^2} \\ &< 0 \quad (\text{since } x_1, x_2 < w_{BM}) \end{aligned} \quad (13)$$

So  $P'_{BM} = P_{1\_occ}(2w_{BM})$  is an upper bound of  $P_{BM}^{ER}$ :

$$\begin{cases} P_{BM}^{ER} < P_{1\_occ}(2w_{BM}) \\ \approx 1 - \left(1 - \frac{1}{(2w_{BM})}\right)^n \sum_{i=0}^{B-1} \frac{n^i}{i! (2w_{BM})^i} \\ 2 \times w_{BM} \times 8 = M \times 1024 \times 1024 \end{cases}$$

Note that in the paper we take the bucket size of BitMatcher as 64 bits (8 bytes).

**A.3.3 Experimental Results.** Below we draw the images of theoretical  $P_{CM}^{ER}$  and  $P_{BM}^{ER}$ 's upper bound, we take  $d = 4$ ,  $n = 0.1\text{Million} = 100000$ , and the range of  $M$  is  $0.1 \sim 2MB$ .

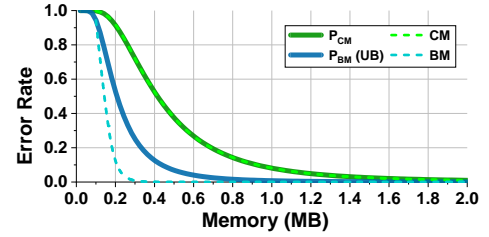


Figure 21: Empirical and theoretical ER.

Note that we also draw the results of  $P_{CM}^{ER}$  and  $P_{BM}^{ER}$  in the real experiment (Fig. 21: green-solid line represents theoretical  $P_{CM}^{ER}$ , blue-solid line represents theoretical upper bound of  $P_{BM}^{ER}$ , green-dash line represents empirical  $P_{CM}^{ER}$ , blue-dash line represents empirical  $P_{BM}^{ER}$ ). We can find that: (1). The theoretical upper bound of  $P_{BM}^{ER}$  is much lower than the theoretical results of  $P_{CM}^{ER}$ , and we have also selected different  $n$  (e.g. 1 million) to plot, and the results are similar to the above Fig. 21. (2). The real  $P_{CM}^{ER}$  is very close to the theoretical value. (3). The real  $P_{BM}^{ER}$  is slightly lower than the theoretical upper bound.