Label Shift Quantification

with Robustness Guarantees via Distribution Feature Matching

Bastien Dussap[†], Gilles Blanchard[†], Badr-Eddine Chérief-Abdellatif[§]

†Université Paris-Saclay, CNRS, Inria, Laboratoire de mathématiques d'Orsay §CNRS, LPSM, Sorbonne Université, Université Paris Cité

September 19, 2023







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Introduction

Model

- ullet \mathcal{X} : the data space.
- \mathcal{Y} : the label space, $\{1, \dots, c\}$.
- $\mathbb{P}_i = p(X|Y=i)$, conditional distribution.
- $\mathbb{P}_1, \dots, \mathbb{P}_c$: A list of *c* distributions, one for each class.

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Label Shift

A "source" distribution.

$$\mathbb{P} = \sum_{i=1}^{c} \beta_i \mathbb{P}_i$$

Training set.

$$\{(x_j, y_j)\}_{j \in [n]} \in (\mathcal{X} \times \mathcal{Y})^n$$

$$\hat{\mathbb{P}}_i := \frac{1}{n_i} \sum_{j \in [n]: y_i = i} \delta_{x_j}(\cdot)$$

A "target" distribution.

$$\mathbb{Q} = \sum_{i=1}^{c} \frac{\alpha_{i}^{*}}{i} \mathbb{P}_{i}$$

Testing set.

$$\{x_{n+j}\}_{j\in[m]}\in\mathcal{X}^m$$

$$\hat{\mathbb{Q}} := \frac{1}{m} \sum_{j=1}^{m} \delta_{\mathsf{x}_{n+j}}(\cdot)$$

 \longrightarrow The distributions differ only on the marginal Y.

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Contaminated Label Shift

A "source" distribution.

$$\mathbb{P} = \sum_{i=1}^{c} \beta_i \mathbb{P}_i$$

Training set.

$$\{(x_j, y_j)\}_{j \in [n]} \in (\mathcal{X} \times \mathcal{Y})^n$$

$$\hat{\mathbb{P}}_i := \frac{1}{n_i} \sum_{j \in [n]: y_i = i} \delta_{x_j}(\cdot)$$

A "target" distribution.

$$\mathbb{Q} = \sum_{i=1}^{c} \alpha_{i}^{*} \mathbb{P}_{i} + \alpha_{0}^{*} \mathbb{Q}_{0}$$

Testing set.

$$\{x_{n+j}\}_{j\in[m]}\in\mathcal{X}^m$$

$$\hat{\mathbb{Q}} := \frac{1}{m} \sum_{i=1}^{m} \delta_{\mathsf{x}_{n+j}}(\cdot)$$

 $\longrightarrow \mathbb{Q}_0$ is unknown.

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Learning to Quantify

Goal: Quantification

Using $\{(x_j, y_j)\}_{j \in [n]} \in (\mathcal{X} \times \mathcal{Y})^n$ and $\{x_{n+j}\}_{j \in [m]} \in \mathcal{X}^m$, estimate α^* .

- González, Castaño, Chawla, and Coz "A review on quantification learning". In ACM Computing Surveys. 2017.
- Esuli, Fabris, Moreo and Sebastiani "Learning to Quantify". In Springer Nature, 2023

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Methods

Classify and Count (CC)

Use a classifier \hat{f} .

 \rightarrow Count the number of times your classifier outputs each class.

$$\hat{lpha}_{cc} = \left(\frac{1}{m}\sum_{j=1}^{m}1_{\hat{f}(x_{n+j})=i}\right)_{i} = \hat{\mathbb{Q}}(\hat{f}(x)=i)$$

Problem

$$\alpha_{cc} = \mathbb{Q}(\hat{f}(x) = i) \neq \mathbb{Q}(y = i) = \alpha^*$$

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Adjusted Classify and Count

Confusion matrix

$$\alpha_{cc} = M_{\hat{f}} \times \alpha^*$$

where M_f is the confusion matrix of \hat{f} .

Black-Box Shift Estimation (BBSE)

$$\hat{\alpha} = \hat{M}_{\hat{\mathbf{f}}}^{-1} \hat{\alpha}_{cc}$$



Lipton, Wang, and Smola. "Detecting and correcting for label shift with black box predictor". In ICML, 2018.

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Methods

Label Shift (\mathcal{LS})

$$\mathbb{Q} = \sum_{i=1}^{c} \frac{\alpha_i^* \mathbb{P}_i}{\alpha_i^*}$$

Distribution Matching

Using a Pseudo-Distance D:

$$\hat{\alpha} = \operatorname*{arg\,min}_{\alpha \in \Delta^c} \, D\left(\sum_{i=1}^c \alpha_i \hat{\mathbb{P}}_i, \hat{\mathbb{Q}}\right)$$

Where $\Delta^c := \{x \in \mathbb{R}^c_+ : \sum x_i = 1\}.$

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Embedding

 \to Let $\phi: \mathcal{X} \to \mathcal{F}$ be a fixed feature mapping from \mathcal{X} into a Hilbert space \mathcal{F} (possibly $\mathcal{F} = \mathbb{R}^D$).

Embedding

$$\phi(\mathbb{P}) := \mathbb{E}_{X \sim \mathbb{P}}[\phi(X)] \in \mathcal{F}$$

 $\longrightarrow \phi$ is chosen so that $\phi(\mathbb{P})$ characterizes the distribution \mathbb{P} .

Pseudo-Distance

$$D_{\phi}(\mathbb{P}, \mathbb{Q}) = \|\phi(\mathbb{P}) - \phi(\mathbb{Q})\|_{\mathcal{F}}$$

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Distribution Feature Matching

Label Shift (\mathcal{LS})

$$\mathbb{Q} = \sum_{i=1}^{c} \alpha_{i}^{*} \mathbb{P}_{i}$$

Distribution Feature Matching (DFM)

$$\hat{\alpha} = \operatorname*{arg\,min}_{\alpha \in \Delta^c} \left\| \sum_{i=1}^c \alpha_i \phi(\hat{\mathbb{P}}_i) - \phi(\hat{\mathbb{Q}}) \right\|_{\mathcal{F}} \tag{\mathcal{P}}$$

where $\Delta^c := \{x \in \mathbb{R}^c_+ : \sum x_i = 1\}$

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Examples

Examples

- $\phi(x) = (1\{\hat{f}(x) = i\})_{i=1,\dots,c} \in \mathbb{R}^c$. For \hat{f} a hard-classifer.
- $\phi(x) = \hat{f}(x) \in \mathbb{R}^c$. For \hat{f} a soft-classifer.
- \bullet $\phi(x)$ of a neural network : $\hat{f}(x) = w^T \phi(x) + b \in \mathbb{R}^c$
- $\phi(x) = (y \mapsto k(x,y)) \in \mathcal{H}_k$
- Lipton, Wang, and Smola. "Detecting and correcting for label shift with black box predictor". In ICML, 2018.
- lyer, Nath, and Sarawagi. "Maximum mean discrepancy for class ratio estimation: Convergence bounds and kernel selection". In ICML, 2014.
- Kawakubo, Christoffel du Plessis and Sugiyama. "Computationally efficient class-prior estimation under class balance change using energy distance". In IEICE Transactions on Information and Systems, 2016.

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Maximum Mean Discrepancy

Kernel embedding of distributions

$$k(x,y) = \langle \phi(x), \phi(y) \rangle_{\mathcal{H}_k},$$

with
$$\phi(x) = (y \mapsto k(x, y))$$
.

$$\phi \colon \mathcal{M}_1^+(\mathcal{X}) \to \mathcal{H}_k$$

$$\mathbb{P}\mapsto \mathbb{E}_{X\sim \mathbb{P}}[\phi(X)]=\phi(\mathbb{P})$$

Maximum Mean Discrepancy

$$D_{\phi}(\mathbb{P}, \mathbb{Q}) = \|\phi(\mathbb{P}) - \phi(\mathbb{Q})\|_{\mathcal{H}_{k}}$$

$$= \mathbb{E}_{\mathbb{P}, \mathbb{P}}[k(X, X)] + \mathbb{E}_{\mathbb{Q}, \mathbb{Q}}[k(Y, Y)] - 2\mathbb{E}_{\mathbb{P}, \mathbb{Q}}[k(X, Y)]$$

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Kernel

Classical kernel

- $k(x,y) = x^T y$, linear.
- $k(x, y) = (\gamma x^T y + c_0)^d$, polynomial.
- $k(x, y) = \tanh(\gamma x^T y + c_0)$, sigmoid.
- $k(x, y) = \exp(-\gamma ||x y||^2)$, gaussian.
- $k(x, y) = \exp(-\gamma ||x y||_1)$, laplacian.
- k(x,y) = ||x|| + ||y|| ||x y||, energy.
- $k(x,y) = \left(1 + \frac{\|x-y\|^2}{\sigma^2}\right)^{-1}$, cauchy.

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Optimisation Problem

Optimisation Problem (DFM)

$$\hat{\alpha} = \underset{\alpha \in \Delta^c}{\operatorname{arg\,min}} \left\| \sum_{i=1}^c \alpha_i \phi(\hat{\mathbb{P}}_i) - \phi(\hat{\mathbb{Q}}) \right\|_{\mathcal{F}}^2 \tag{\mathcal{P}}$$

where $\Delta^c := \{x \in \mathbb{R}^c_+ \colon \sum x_i = 1\}$

Solving (\mathcal{P}) amounts to solving a Quadratic Programming (QP) in dimension c. Indeed, we can rewrite the problem as:

minimise
$$\frac{1}{2}\alpha^T \hat{\mathbf{G}}\alpha + q^T \alpha$$
 (QP) subject to $\alpha \succeq 0_c$ and $\mathbf{1}_c^T \alpha = 1$,

with $q = \left(\langle \phi(\hat{\mathbb{P}}_i), \phi(\hat{\mathbb{Q}}) \rangle\right)_{i=1}^c$. This is a c-dimensional QP problem, which can be solved efficiently.

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Theoretical Analysis

Assumptions

$$\sum_{i=1}^{c} \beta_{i} \phi(\mathbb{P}_{i}) = 0 \iff \beta = 0 \tag{A_{1}}$$

and

$$\exists C > 0: \ \|\phi(x)\|_{\mathcal{F}} \le C \text{ for all } x. \tag{\mathcal{A}_2}$$

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Main Theorem under LS

Theorem

For any $\delta \in (0,1)$, with probability greater than $1-\delta$:

$$\|\hat{\alpha} - \alpha^*\|_2 \lesssim \frac{C\sqrt{\log(c/\delta)}}{\sqrt{\Delta_{\min}}} \left(\frac{\|w\|_2}{\sqrt{n}} + \frac{1}{\sqrt{m}}\right) \tag{1}$$

$$\lesssim \frac{C\sqrt{\log(c/\delta)}}{\sqrt{\Delta_{\min}}} \left(\frac{1}{\sqrt{\min_i n_i}} + \frac{1}{\sqrt{m}} \right), \tag{2}$$

where $w_i = \frac{\alpha_i^*}{\tilde{\beta}_i}$, $\tilde{\beta}_i$ the empirical proportion of class i in the Source and n_i is the number of points of class i in the Source.

The same result holds when replacing α^* by the (unobserved) vector of empirical proportions $\tilde{\alpha}$ in the target sample, both on the left-hand side and in the definition of w.

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Properties of Δ_{\min}

Definition

$$\hat{\mathbf{G}}_{ij} = \langle \phi(\hat{\mathbb{P}}_i), \phi(\hat{\mathbb{P}}_j) \rangle
\hat{\mathbf{M}}_{ij} = \langle \phi(\hat{\mathbb{P}}_i) - \overline{\phi}, \phi(\hat{\mathbb{P}}_j) - \overline{\phi} \rangle$$

with $\overline{\phi} = c^{-1} \sum_{k=1}^{c} \phi(\hat{\mathbb{P}}_k)$.

 λ_{\min} is the smallest eigenvalue of \hat{G} .

 Δ_{\min} is the second smallest eivenvalue of \hat{M} .

In particular, it holds:

 $\Delta_{\min} \geq \lambda_{\min}$.

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Properties of Δ_{min}

Let (b_1, \dots, b_c) be a c-uple of vectors of \mathbb{R}^D assumed to be linearly independant. We denote G the Gram matrix of those vectors, i.e. $G_{ij} = \langle b_i, b_i \rangle$.

Theorem

For any number of classes c, Δ_{min} is equal to $\min_{\substack{\|u\|=1\\\mathbf{1}^Tu=0}} u^T Gu$.

 $ightarrow \Delta_{\min}(b_1,\cdots,b_c)$ is always greater than the smallest eigenvalue of the Gram matrix.

Theorem

In particular for two classes, $\Delta_{min}(b_1, b_2) = \frac{1}{2} ||b_1 - b_2||^2$.

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Properties of Δ_{min}

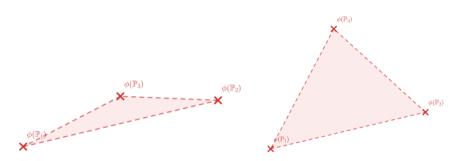


Figure: On the left Δ_{min} is small. On the right Δ_{min} is large.

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soft-DFM for CLS

A "target" distribution.

$$\mathbb{Q} = \sum_{i=1}^{c} \alpha_{i}^{*} \mathbb{P}_{i} + \alpha_{0}^{*} \mathbb{Q}_{0}$$
 (CLS)

Soft-DFM

$$\hat{\alpha}_{\text{soft}} = \underset{\alpha \in \text{int}(\Delta^c)}{\text{arg min}} \left\| \sum_{i=1}^c \alpha_i \phi(\hat{\mathbb{P}}_i) - \phi(\hat{\mathbb{Q}}) \right\|_{\mathcal{F}}^2, \tag{\mathcal{P}_2}$$

where $\operatorname{int}(\Delta^c) := \{x \in \mathbb{R}^c_+ : \sum x_i \leq 1\}$

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Main Theorems under CLS

Definition

Introduce $\bar{V} := \operatorname{Span}\{\phi(\mathbb{P}_i), i \in [c]\}$ and let $\Pi_{\bar{V}}$ be the orthogonal projection on \bar{V} .

Theorem

With probability greater than $1 - \delta$:

$$\|\hat{\alpha}_{\text{soft}} - \alpha^*\|_2 \lesssim \frac{1}{\sqrt{\lambda_{\min}}} \Big(3\epsilon_n + \varepsilon_m + \sqrt{2\alpha_0^* \epsilon_n \|\phi(\mathbb{Q}_0)\|} + \|\Pi_{\tilde{V}}(\phi(\mathbb{Q}_0))\|_{\mathcal{F}} \Big),$$

with:

$$\epsilon_n = \sqrt{\frac{\log(c/\delta)}{\min_i n_i}}; \qquad \varepsilon_m = \sqrt{\frac{\log(c/\delta)}{m}};$$

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Robustness to contamination

$$\|\hat{\alpha}_{\text{soft}} - \alpha^*\|_2 \lesssim \frac{1}{\sqrt{\lambda_{\min}}} \Big(3\epsilon_n + \varepsilon_m + \sqrt{2\alpha_0^* \epsilon_n \|\phi(\mathbb{Q}_0)\|} + \|\Pi_{\bar{V}}(\phi(\mathbb{Q}_0))\|_{\mathcal{F}} \Big)$$

Observe that the bound shows the robustness of a soft-DFM procedure against contaminations \mathbb{Q}_0 that are orthogonal to $\bar{V} := \operatorname{Span}\{\phi(\mathbb{P}_i), i \in [c]\}$.

Classifier (BBSE)

The feature space is of the same dimension as the number of sources hence the orthogonal component will always be 0 and we expect no robustness property for BBSE.

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Robustness to contamination

$$\|\hat{\alpha}_{\text{soft}} - \alpha^*\|_2 \lesssim \frac{1}{\sqrt{\lambda_{\min}}} \Big(3\epsilon_n + \varepsilon_m + \sqrt{2\alpha_0^* \epsilon_n \|\phi(\mathbb{Q}_0)\|} + \|\frac{\Pi_{\bar{V}}(\phi(\mathbb{Q}_0))}{\|\mathcal{F}\|} \Big)$$

Observe that the bound shows the robustness of soft-DFM procedure against contaminations \mathbb{Q}_0 that are orthogonal to $\bar{V} := \operatorname{Span}\{\phi(\mathbb{P}_i), i \in [c]\}$.

Gaussian kernel (KMM)

Two embeddings $\phi(\mathbb{P})$ and $\phi(\mathbb{P}')$ will be close to orthogonal if \mathbb{P} and \mathbb{P}' are well-separated.

We anticipate that KMM will be robust against contamination distributions \mathbb{Q}_0 whose main mass is far away from the source distributions.

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Experiments

The source is a list of 5 Gaussian distributions. α_0^* ranges from from 0 to 0.3. We repeated the experiments with different dimensions.

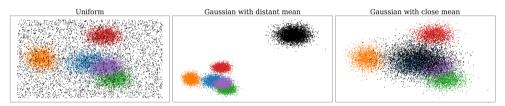


Figure: Three kinds of noise: On the left \mathbb{Q}_0 is uniformly distributed over the data range, in the middle \mathbb{Q}_0 is Gaussian with a mean distant from the other means and on the right \mathbb{Q}_0 is Gaussian with a similar mean to the others

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Experiments

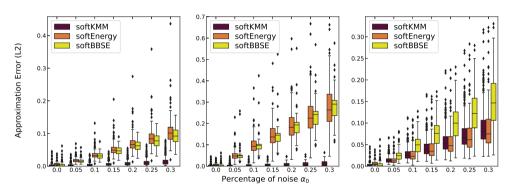


Figure: Robustness of the algorithms to three types of noise. Left: uniform noise; middle: noise is a new class far from the others; right: noise is a new class in the middle of the others.

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Experiments on Cytometry dataset

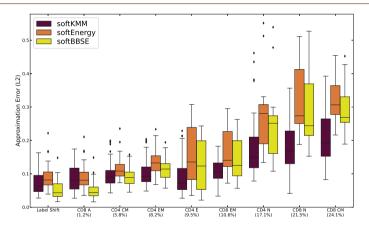


Figure: Each column represents the error — computed using the ℓ_2 norm between the true proportions and the estimated proportions — obtained when some class is absent from the source but present in the target distribution. The first column gives the results when no class is discarded. The class are sorted according to the average proportions they represent in the samples (x labels mention class held out from the source and its proportion)

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Conclusion

- We introduced a general approach for Label Shift Quantification.
- We provided a general theoretical analysis of DFM, improving over previously known bounds derived for specific instantiations only.
- We proposed the use of Random Fourier Features to speed up the computation of kernel-based approaches and obtain an explicit finite-dimensional vectorization (or "sketch") of the distributions.
- We analysed theoretically the behavior of DFM under departures from the label shift hypothesis, a situation not studied in earlier works, and put into light a robustness against certain types of perturbations, depending on the feature mapping ϕ used.

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Thank you for your attention.

Kernel Mean Matching

Kernel Mean Embedding

$$\phi(\mathbb{P}) = \mathbb{E}_{\mathbb{P}}[k(x,\cdot)]$$

Kernel Trick

$$\langle \phi(\mathbb{P}), \phi(\mathbb{Q}) \rangle = \mathbb{E}_{\mathbb{P},\mathbb{P}}[k(x,x')] + \mathbb{E}_{\mathbb{Q},\mathbb{Q}}[k(y,y')] - 2\mathbb{E}_{\mathbb{P},\mathbb{Q}}[k(x,y)]$$

→ Methods using Kernels are quadratic.

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Random Fourier Features

Random Fourier Features are based on Bochner's Theorem:

Theorem

A continuous function φ on \mathbb{R}^D is positive definite if and only if φ is the Fourier transform of a non-negative measure.

RFF

$$k(x,y) = k(x - y)$$

$$= \mathbb{E}_{\omega \sim \Lambda_k} [e^{i\omega^T (x - y)}]$$

$$= \mathbb{E}_{\omega \sim \Lambda_k} [\cos (\omega^T (x - y))]$$

Random Fourier Features

Using a sample $(\omega_i)_{i=1}^{D/2}$ i.i.d. from Λ_k :

$$\phi(x) = \sqrt{\frac{2}{D}} \left[\cos(\omega_i^T x), \sin(\omega_i^T x) \right]_{i=1}^{D/2}$$

is such that

$$k(x, y) = \mathbb{E}[\phi(x)^T \phi(y)],$$

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Random Fourier Feature Matching

Random Fourier Feature Matching (RFFM)

DFM method using:

$$\phi(x) = \sqrt{\frac{2}{D}} \left[\cos(\omega_i^T x), \sin(\omega_i^T x) \right]_{i=1}^{D/2}$$

Relying on RFF with D Fourier features induces a complexity of O(D(n+m)) since we only have to compute $\phi(\hat{\mathbb{P}}_i)$ and $\phi(\hat{\mathbb{Q}})$.

Computing $\phi(\hat{\mathbb{P}})$ reduces to a matrix multiplication, for which GPU are well suited.



Rahimi and Recht. "Random features for large-scale kernel machines". In Advances in neural information processing systems, 2007.

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Main Theorems under CLS

Theorem

If $\alpha_0^* = 0$, for any $\delta \in (0,1)$, with probability greater than $1 - \delta$:

$$\begin{split} \|\hat{\alpha}_{\text{soft}} - \alpha^*\|_2 &\lesssim \frac{C\sqrt{\log(c/\delta)}}{\sqrt{\lambda_{\min}}} \left(\frac{\|w\|_2}{\sqrt{n}} + \frac{1}{\sqrt{m}}\right) \\ &\lesssim \frac{C\sqrt{\log(c/\delta)}}{\sqrt{\lambda_{\min}}} \left(\frac{1}{\sqrt{\min_i n_i}} + \frac{1}{\sqrt{m}}\right), \end{split}$$

Gram Matrix

 λ_{\min} is the smallest eigenvalue of \hat{G} and Δ_{\min} the second smallest eivenvalue of \hat{M} . In particular, it holds:

 $\Delta_{\min} \geq \lambda_{\min}$.

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Theorem Deltamin

Definition

For any number of classes c and any vectors $\{b_1, \dots, b_c\}$, we define:

$$\Gamma(b_1,\cdots,b_c) := \min_{(I_1,I_2)\in\mathcal{P}_2([c])} d^2(C_{I_1},C_{I_2}),$$

with the following:

$$\mathcal{P}_{2}([c]) = \{I_{1}, I_{2} \subset \{1, \cdots, c\} | |I_{1} \cap I_{2}| = 0, |I_{1} \cup I_{2}| = c, |I_{1}|, |I_{2}| > 0\}$$

$$C_{I} = \left\{ \sum_{j \in I} \lambda_{j} b_{j} | \lambda \in \Delta^{|I|} \right\}$$

$$d^{2}(A, B) = \inf_{\substack{x \in A \\ y \in B}} ||x - y||_{2}^{2}$$

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Theorem Deltamin

Theorem

For any number of classes c and any vectors $\{b_1, \dots, b_c\}$:

$$\frac{c}{4} \Gamma(b_1, \dots, b_c) \ge \Delta_{min}(b_1, \dots, b_c) \ge \frac{1}{2} \Gamma(b_1, \dots, b_c), \tag{3}$$

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