#### INFERENCE : AND STATISTICAL SAMPLING

Sampling is a technique of selecting a subset of a population to make statistical inferences from them and estimate characteristics of the whole population. Eg. (1) Vaccine trials (11) Poll prediction

Some important terms 4 définitions.

- I Population: It is the set of all aggregate objects under study (usually denoted by N)
- 2 Sample It is a subset of the population that is considered to study the behaviour of the population. (The sample size is usually denoted by n)

If n = 30, it is considered to be a large sample n < 30, - small sample

The statistical constants of the population such as Mean, Variance etc are called <u>Parameters</u> (H, J2) The statistical constants of the sample which is used to estimate the parameters are called Statistics (x & s2)

Sampling can be done in different ways but one which is most common is Random Sampling. This means every member of the population has an equal chance of being selected. It can be done in 2 ways

- (1) With Replacement: If N is the size of the population & n is the size of sample, then there are N possible samples.
- (ii) Without Replacement: " Con possibilities

Sampling Distribution: A sampling distribution is the frequency distribution of a statistic over many random samples from a single propulation.

let us say we have different samples of sizen dearn from a population of size N. For each of these samples, we can compute mean, variance etc and they soll not be the same If we group these according to their frequencies, then the frequency distribution so generated is called the Sampling Distribution. In general, we can have a sampling distribution of mean, variance etc.

The standard deviation of the sampling distribution is called the Standard Error.

Egr A population consists of 3 members [1,23]. Form a sampling distribution of the mean for random samples of size 2 with replacement

Note that: N=3,  $\mu = \frac{1+2+3}{3} = 2$ ,  $\sigma^2 = \frac{1}{N} \sum (x-\mu)^2 = \frac{2}{3}$ 

The random samples of size 2 with replacement one:  $\{(1,1), (1,2), (1,3), (2,1), (2,2), (2,3), (3,1), (3,2), (3,3)\} \ \{=N^n=3^2=9\}$  Means of these samples are: 1,1.5,2,1.5,2,2.5,2,2.5,3 Sampling distribution of mean is:

$$\frac{\pi i}{1} \quad \frac{f_{1}}{1} \quad \frac{f_{2}}{1} \quad \frac{$$

Central limit Theorem: If we have a population N with mean  $\mu$  and  $s,d,\sigma$  is  $N(\mu,\sigma)$  and we take sufficiently large samples from the population with replacement, then the distribution of the sample means will be approximately normally distributed.

Hypothesis: The assumption that we make regarding the parameter of the population is called Hypothesis. Fest of Significance or Fest-of Hypothesis:

An important aspect of sampling theory is to study or assert that the parameter of the population is the same as the statistic obtained from the random sample.

The process or test that decides whether to accept the hyprothesis or not is called the test of significance or test of hyprothesis.

In order to agree at a decision, we want to make certain assumptions known as hypotheses which may or may not be time.

The hypothesis is accepted or rejected based on some statestical tests. We decide whether the sample statistic is significantly different from the parameter at a desired level of significance. Hence these tests are called tests of significance.

There are 2 types of hypotheses.

(1) Null Hypotheses; (Hypotheses of No deference)
The hypothesis which assumes that there is no

significant difference between sample statistic and the population parameter is called the Niell Hypothesis & it is devoted by Ho.

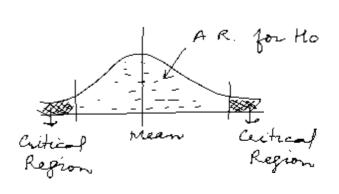
(ii) Alternate Hypothesis: Under this hypothesis, we assume that there is a significant defference between sample 4 population aspects

Type I & Type II	Errors	Deci	ision
Type I Evor if a true	Hypotheses	Accepted	Rejected
hypothesis is rejected	Frue		Type I aron
Type II Erron: if a false		Type II	
hubothesis is selected.			

level of Significance: The probability level below which we reject the hypothesis is called the level of significance (l.o.s). It can be 11/2, 21/2, 51/2, 101/2 etc. In general, 11/4 51/2 hos are considered for this, we refer to the table of values under the columns 0.01 or 0.05 resp

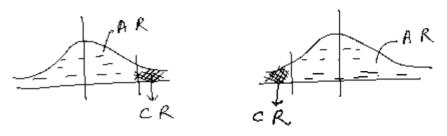
Cutical Region and Acceptance Region:

C.R. -> Region that corresponds to the rejection of hypothesis
AR. -> " acceptance " "



One tailed and two tailed Fests:

A Test on statistical hypothesis where the alternate hypothesis is one sided is called one tailed test there Ho:  $\mu = \mu_0$ ;  $\mu_1: \mu > \mu_0$  or  $\mu < \mu_0$ 



A test on statistical hypothesis where the alternate hypothesis is two sided is called two tailed test.

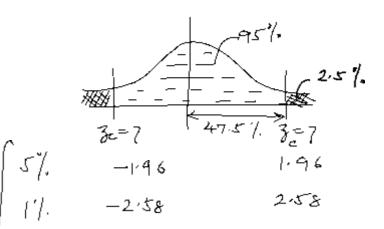
Here Ho!  $\mu = ho$ ; H,  $\mu \neq \mu_0$ 

Normally, we use 2 tailed test. If we are using a one tailed test, then for the loss of L, we have to refer to 2d value in the 2 tailed table.

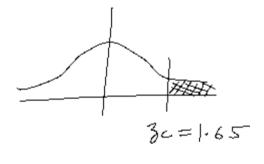
 $x = \mu - \sigma$   $3 = x - \mu = \mu - \sigma - \mu = -1$   $x = \mu + \sigma$ ,  $3 = \frac{\mu + \sigma - \mu}{\sigma} = -1$ 

Mean Mean

P(μ-σ < n < μ+σ) = P(-1<3<1) = 2 P(0<3<1)=2(0.3413) =0.6826



One tailed test ;



Fest of Significance for Single Mean:

(to test whether the difference between sample mean & population mean is significant or not)

Under null hypothesis,  $z = \frac{\overline{x} - \mu}{\overline{t_n}}$ 

where  $\bar{n} = \text{Sample mean}$   $\mu = \text{population mean}$   $\sigma = s.d.$  of the population n = sample size

y o u not known, we use 
$$z = \frac{\overline{x} - \mu}{\frac{\Delta}{5\pi}}$$

d = sample s.d.

If the los is I and 3x is the critical value, then

$$-3d < 3 = \frac{\overline{x} - \mu}{\overline{5}\overline{n}} < 3d$$

Also, the limits of the population mean  $\mu$  are given by

Ex A normal population has a mean of 6.8 & sd. of 1.5

A sample of 400 members gave a mean of 6.75. Is the difference significant?

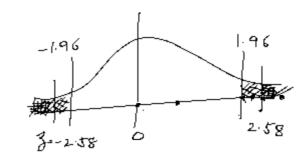
Given 
$$\mu = 6.8$$
,  $\sigma = 1.5$ ,  $n = 400$ ,  $\bar{x} = 6.75$   
Ho:  $\bar{x} = \mu$  is the sample mean 4 propulation mean are same  $|\bar{x}| = |\bar{x} - \mu| = |\bar{x} - \mu| = |6.75 - 6.8| = |-0.667| = 0.667$   
 $|\bar{x}| = |\bar{x} - \mu| = |6.75 - 6.8| = |-0.667| = 0.667$ 

=> Ho is accepted at 51/, Los

Ex The mean weight obtained from a random rample of size 100 is 64 gm, the s.d. of the weight dist. in the population is 3 gm. Can we say that the mean weight of the population is 66 gm at 5% l.O.A?

Also set up 99% Confidence interval for the mean weight of the population.

$$\frac{H_0: \mu = \pi; \quad 3 = \frac{\pi - \mu}{\frac{\pi}{5n}} = \frac{64 - 66}{\frac{3}{\sqrt{100}}} = -6.667$$



99% confidence limits are

$$= \overline{2} \pm 2.58 \frac{5}{5n}$$

$$= 64 \pm 2.58 \frac{3}{5100} = 64 \pm 0.774$$

$$= (63.226, 64.774)$$

HW

Ex. 1 Sugar is backed in bags by an automatic machine with mean contents of a bag as 1.0 kg. A random sample of 36 bags is selected 4 mean is found to be 0.997 kg. If a s d. of 0.01 kg is acceptable, determine it the machine needs adjustment

2) The average zenc concentration recovered from a sample of zenc measurements in 40 locations is found to be 2.54 gm per mellilitie. Find the 95%.

C.I. for the mean zinc concentration in the river assuming the 11 d is 0.32 gm.

Note that Ig 131 < 3c, to is accepted 131 > 3c, to is rejected

# Z-test for Difference between Fire Means:

let \$1, \$2 be the means of 2 samples of sizes my 4 ns from propulations with means pu, pls and s.d. 10 01, 02.

Then 
$$3 = (\frac{\chi_1 - \chi_2}{-\chi_1 - \chi_2}) - (\mu_1 - \mu_2)$$
 is a standard normal variable

Under the mill hypothesis  $\mu_1 = \mu_2$ ,

Remarks (1) If of 2 4 52 are unknown, then

$$\sqrt{\frac{A_1^2}{n_1} + \frac{A_2^2}{n_2}} \text{ is used as an estimate of S.E. of } \frac{\pi_1 - \pi_2}{\pi_1}$$

$$4 \text{ then } \otimes \rightarrow 3 = \frac{\pi_1 - \pi_2}{\sqrt{\frac{A_1^2}{n_1} + \frac{A_2^2}{n_2}}}$$

(11) If the two populations have the same variance  $\sigma^{\perp}$ , then  $3 = \frac{\overline{x_1} - \overline{x_2}}{\sigma \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$ 

$$3 = \frac{\overline{x_1 - \overline{x_2}}}{\sigma \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}}$$

Ex Two random samples of 100 students each of the sahools A & B were drawn. The CGPA of students from school A had mean 2.82 with s.d. 0.63 and that of school B had mean 2.43 with s.d. 0.65.

Does the data indicate any difference in the mean CGPA from two schools?

Given 
$$n_1 = n_2 = 100$$
,  $\overline{\chi}_1 = 2.82$ ,  $S_1 = 0.63$   
 $\overline{\chi}_2 = 2.43$ ,  $S_2 = 0.65$   
Ho:  $\mu_1 = \mu_2$   
 $3 = \frac{\overline{\chi}_1 - \overline{\chi}_2}{\sqrt{n_1} + \frac{S^2}{n_2}} = \frac{2.82 - 2.43}{\sqrt{(0.63)^2 + (0.65)^2}} = 4.308$ 

As 3c (5%, los.)=1.96

3c (1%, los.)=2.58

Since 3 > 3c => Ho is rejected

There is a difference in the mean CGPA of
2 schools

Ex. The mean heights in two large samples of 1000 and 2000 men are 67.5 inches 4 680 inches resp. Can the samples be regarded as drawn from the same normal population with s.d 2.5 inches.

Ho: 
$$\mu = \mu_1$$
  $3 = \frac{\pi_1 - \pi_2}{\sigma + \frac{1}{1 + n_2}} = \frac{67.5 - 68.0}{2.5 \sqrt{\frac{1}{1000} + \frac{1}{2000}}} = -5.16$ 

131=5.16 > 3c(for 11, 85% los) = to rejected

Ex. A random sample of 50 electric light tubes of type A gave mean of 1400 hrs 4 sd. 200 hrs and type B 4 4 1200 hrs & sd. 100 hrs.

Is there any significant difference between the two types of tubes?

the: M= M= m= m= 50

$$3 = \frac{\overline{x_1 - x_2}}{\sqrt{\frac{x_1^2}{n_1} + \frac{x_2^2}{n_2}}} = \frac{1400 - 1200}{\sqrt{\frac{(200)^2}{50} + \frac{(100)^2}{50}}} = 6.32$$

As 3 > 3c (for 17. 45%, l.o.s), the is rejected

Ex. The mean marks of 2 sections are as given.

Section A: 32 students have average 72 with s.d. 8 section B. 36 " " 1, 75 with s.d. 6

Can we say that section B is better than section A?

Ho:  $\mu_1 = \mu_2$  is there is no significant diff between the 2 sections

$$3 = \frac{\overline{x_1 - x_2}}{\sqrt{\frac{A_1^2}{m_1} + \frac{A_2^2}{n_2}}} = \frac{72 - 75}{\sqrt{\frac{(8)^2}{32} + \frac{(6)^2}{36}}} = -1.73$$

$$|3| = 1.73$$
  $3c(5\%, 1.0 s, one tailed) = 1.65$   $3c(1\%, 1.0 s, one tailed) = 2.33$ 

=> to is accepted at 1% los, rejected at 5% hos.

#### Fest of Significance for Single Proportion:

This test is used to find a significant difference between proportion of the sample and that of the population.

$$3 = \frac{\beta - P}{\sqrt{\frac{PQ}{n}}} \qquad ----- \varnothing$$

where p = observed (sample) proportion of success P = population proportion of success & = 1-P = " " " failure n = sample size

marks: I The probable limits for the observed proportion of successes are: Pt 32/PD / 32 = significant value of 3at los. L 2. If Pis not known, the limits for the proportion in the population are : \$\pm 3d \frac{pq}{n}

Ex. A coin was tossed 400 times & the heads lurned up 220 times Fest the hyprothesis that the coin is unbiased Ho: Coin is unbiased in P=0.5, 8=1-0.5=05  $p = \frac{270}{400} = \frac{11}{20}$ ; Under the  $3 = \frac{\frac{11}{20} - \frac{1}{2}}{\sqrt{\frac{(\frac{1}{2})(\frac{1}{2})}{400}}} = 2$ 3c = { 1.96 at 5/. LOS - Accept to at 1% los. Reject Ho at 5% LOS

Ex In a sample of 1000 people in a state, 540 are rice eaters & the rest are wheat eaters. Can we assume at 11. 45% los that both rice 4 wheat are equally popular in the state?

Ho: Rice & wheat are equally propular is  $P = \frac{1}{2}$  $p = \frac{570}{1000} = 0.54$ , q = 1-0.54 = 0.46, n = 1000

 $3 = \frac{0.54 - 0.5}{\sqrt{(0.5)(0.5)}} = 2.5248$   $3z = \begin{cases} 1.96 & \text{at } 5\%, (0.5) \\ 2.58 & \text{at } 1\%, (0.5) \end{cases}$ 

>> Accept the at 1% Los (3 < 3c)
Reject to at 5% " (3 > 3c)

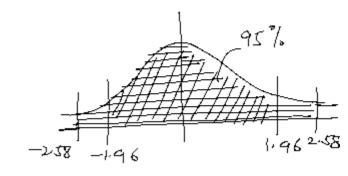
Ex During testing in a sample of 300 chips, 10 are found to be defective can the manufacturer's claim that 2% of the chips are defective be accepted?

Ho: p = P is p = 0.02 m = 300,  $p = \frac{10}{300} = \frac{1}{30}$ , P = 0.02, A = 1-0.02= 0.98

 $3 = \frac{\frac{1}{30} - 0.02}{\sqrt{\frac{(0.02)(0.98)}{300}}} = 1.6495 < 3e(=1.96 at 5'/.1.01)$ =) Ho is accepted.

Ex. (1) A manufacturer claims that only 4%, of his products HW, are defective. Arandom sample of 600 products contained 36 defectives. Fest the manufacturer's claim

(2) A die is thrown 9000 times and a throw of 3 or 4 is observed 3240 times. Can the die be regarded as unbiased wints 95% confidence interval?



3 ∈ (-2.58 to 2.58) come 99%

$$3 = \frac{p_1 - p_2}{\sqrt{p_R(\frac{1}{n_1} + \frac{1}{n_2})}}$$
 where  $P = \frac{n_1 p_1 + n_2 p_2}{n_1 + n_2}$ ,  $Q = 1 - P$ 

En A machine produced 16 defectives in a batch of 500. After overhauling, it produced 3 defectives in a batch of 100. Has the machine improved ?

Green 
$$n_1 = 500$$
,  $p_1 = \frac{16}{500}$ ,  $n_2 = 100$ ,  $p_2 = \frac{3}{100}$ 

 $\underline{H_0}$ :  $p_1 = p_2$   $\underline{H_1}$ :  $p_2 < p_1$  (one tailed test)

$$P = \frac{n_1 p_1 + n_2 p_2}{n_4 + n_2} = \frac{(500)(\frac{16}{500}) + (100)(\frac{3}{100})}{500 + 100} = \frac{19}{600}$$

$$\Rightarrow 8 = 1 - P = 1 - \frac{19}{600} = \frac{581}{600}$$

$$3 = \frac{\frac{16}{500} - \frac{3}{100}}{\sqrt{\frac{19}{600}(\frac{581}{600})(\frac{1}{500} + \frac{1}{100})}} = 0.10426$$

3c= 1.65 5%. Los.

Inference ?? => Ho is accepted

of the machine hasn't improved with overhauling

## Fests of Significance of Small Samples:

#### Degrees of Freedom; D.F. are the number of

independent values that a statistical analysis can estimate.

I.e. it is the number of values that are fee to vary as we estimate parameters.

Fypically, the df. is equal to the sample size minus the number of parameters we need to calculate.

Fests for Small Samples. n < 30

## Student's t-test:

(Fest of significance of the mean of a small sample) (Given by GOSSET)

let  $x_i (i=1,2,...,n)$  be a random sample of size n from a normal population with mean  $\mu$  4 variance  $\sigma^2$ .

to: There is no significant difference between the sample mean (x) & the population mean (4)

Then 
$$t = \frac{\overline{x} - \mu}{\frac{A}{5\pi}}$$
 where  $A^2 = \frac{1}{n} \sum_{i=1}^{n} (\pi_i - \overline{\pi})^2$   $\longrightarrow \mathfrak{S}$ 

st is an unbiased estimate of the propulation variance of 4 it follows t-distribution with (m) d.f

# Confidence limits for $\mu$ : (by t-test) C. limits for $\mu$ for $\lambda$ los = $\pi \pm \frac{\Delta}{5\pi} t_{\lambda}$

Ex. A random sample of size 16 has mean 53. The sum of oquares of the deviation from mean is 135 can this sample be regarded as taken from a population having 56 as mean ? Also obtain 95%. 499%. confidence limits for the mean of the population

Ho: sample mean = population mean H1 14 \$56

Given: n = 16,  $\bar{x} = 53$ ,  $\mu = 56$ ,  $\sum (x - \bar{x})^{\frac{1}{2}} = 135$   $s^{2} = \frac{1}{n-1} \sum (x - \bar{x})^{\frac{1}{2}} = \frac{1}{16-1} (135) = 9$ ,  $\Rightarrow s = 3$ 

 $t_{0.05}(15d.f.) = 2.131$ 

As t>to => two is rejected

951, C.I. = x + (2.131)

 $=53\pm\frac{3}{\sqrt{16}}(2\cdot131)=(51.4018,54.5982)$ 

 $\frac{99'/.CI}{\sqrt{n}} = \overline{x} \pm \frac{3}{\sqrt{n}} (2.947) = 53 \pm \frac{3}{\sqrt{16}} (2.947)$  = (50.7898, 55.2102)

Ex A sample of 20 items has mean 42 units and s of 5 units. Fest the hypothesis that it is a landom sample from a normal population with mean 45 units

Green 
$$n=20$$
,  $x=42$ ,  $S=5$ ,  $(\mu=45)$ ,  $d\cdot f=20-16$   
(Sample s.d.)

Green 
$$n=20$$
,  $\bar{x}=42$ ,  $S=5$ ,  $(\mu=45)$ ,  $d\cdot f\cdot = 20-1$   
 $= 19$   
(Sample s.d.)  
 $\frac{H_0: \mu=45}{s^2=\frac{n}{n-1}s^2} \Rightarrow s^2=\frac{20}{20-1}(5)^2=26.3159$   
 $s=5.1298$ 

$$t = \frac{\pi - \mu}{\sqrt{\pi}} = \frac{42 - 45}{\frac{5 \cdot 1298}{\sqrt{20}}} = -2.6153$$

$$\Rightarrow |k| = 2.6153, \quad t_{0.05} (19 d.f.) = 2.093$$
As  $k > t_{0.05} (19 d.f.) \Rightarrow Ho is rejected$ 

(HW) check at 1%. LOS

En A machine is expected to produce nails of (F.W) length 3 cm. A random sample of 25 nails gave an average length of 3.1 cm with s.d 0.3 cm. can it be said that the machine is producing nails as per specification at 5% l.as?

A random sample of 10 students have I.R.15
70, 120, 110, 101, 88,83,95,98, 107, 100.

Does this data support the assumption of a population mean I.R. of 100?

Ho 
$$\mu = \bar{\chi}$$

Given  $n = 10$ ,  $\bar{\chi} = \frac{1}{10} \left[ 70 + 120 + 110 + - - - + 100 \right] = 97.2$ 
 $\vec{\lambda} = \frac{1}{n-1} \sum_{i} \left[ (\pi_i - \bar{\chi})^2 = \frac{1}{10-1} \left[ (70 - 97.2)^2 + (120 - 97.2)^2 + - - - - + (100 - 97.2)^2 \right]$ 
 $= \frac{1}{4} \left[ (833.6) \right] = 203.7333$ 

⇒ 14.2735

$$t = \frac{\pi - \mu}{\frac{3}{\sqrt{n}}} = \frac{97.2 - 100}{\frac{14.2735}{\sqrt{10}}} = -0.6203$$

 $\Rightarrow |t| = 0.6203$   $t_{0.05}(9 \text{ d.f.}) = 2.262$ As  $|t| < t_{0.05}$ ,  $\Rightarrow H_0$  is accepted

Ex the life time of electric bulbs for a random sample of 10 from a large consignment gave the following table.

Item: 1 2 3 4 5 6 7 8 9 10 lefe in: 1.2 4.6 3.9 4.1 5.2 3.8 3.9 4.3 4.4 5.6

Can we accept the hypothesis that the average life time of bulbs is 4500 hrs?

to: x=11 is an life of bulbs is 4500 his

$$\overline{x} = \frac{1}{10} \left[ 1.2 + 4.6 + - - 5.6 \right] \times 1000 = 4100 \text{ fms.}$$

$$S^{2} = \frac{1}{10-1} \left[ \left( 1.2 - 4.1 \right) + \left( 4.6 - 4.1 \right)^{2} + - - \cdot + \left( 5.6 - 4.1 \right)^{2} \right] \left( 1000 \right)^{2}$$

$$= 1.3844 \times 10^{6}$$

$$\Rightarrow \Delta = 1.1789 \times 10^{3} = 1178.9$$

$$t = \frac{\overline{x} - M}{\sqrt{n}} = \frac{4100 - 4500}{178.9} = -1.0729$$

$$t_{0.05} (9 \text{ d.f.}) = 2.262$$

As It < to.05 (9 df.) => Ho is accepted

Ex. The height of 12 men from a city is

70,67,62,68,61,68,70,64,65,69,71,68 inches

Is it reasonable to believe that the average height of men in the city is 65 inches?

Does the result suggest a mean survival time of 12.5 days? Construct a 99% C. I.

For Sample t-test:

$$t = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{s_x^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}} ; \quad x_x^2 = \frac{(n_{r1}) x_1^2 + (n_2 - 1) x_2^2}{n_1 + n_2 - 2}$$

where  $S_1^2$  &  $S_2^2$  are the estimates of the variances of the I4 II propulations

X1, x2 - means of the 2 samples

The t-statistic given above follows (M+N2-2) degrees of freedom:

Ex 4no samples showed the following results.

Sample A: 44 44 56 46 47 47 58 53 49 55

Sample B: 35 38 37 32 40 39 36 41

Fest if the averages of the two propulations is the same.

Ho: 
$$\mu_1 = \mu_2$$
 (H1:  $\mu \neq \mu_2$ )

$$\overline{x}_1 = 49.9$$
 ,  $\overline{x}_2 = 37.25$ 

$$\begin{aligned} s_{1}^{+} &= \int_{m_{1}-1}^{1} \sum_{i} \left( x_{i}_{1} - \overline{x}_{1}_{1} \right)^{2} \\ &= \int_{q}^{1} \left[ \left( 44 - 49.9 \right)^{2} + \left( 44 - 49.9 \right)^{2} + \left( 55 - 49.9 \right)^{2} \right] \\ &= - - + \left( 55 - 49.9 \right)^{2} \end{aligned}$$

$$4^{2} = \frac{1}{7} \left[ (35 - 37.25)^{2} + - - - + (41 - 37.25)^{2} \right] = 8.5$$

$$A_{x}^{2} = \frac{9(26.7666) + 7(8.5)}{10 + 8 - 2} = 18.775$$

$$t = \frac{49.9 - 37.25}{\sqrt{(18.775)(\frac{1}{10} + \frac{1}{8})}} = 6.155$$

As tous (16 d.f.) = 2.120

→ Ho is rejected

Ex: A group of 9 boys were fed on diet A 4 another H.W. group of 8 boys were fed on diet B for a period of 6 months, the following increase in weights were then recorded. Fest whether diets ARB are significantly different in terms of increase in weight Diet A - 5 6 8 1 10 4 3 9 6

Diet 8: 2 3 6 8 9 1 2 7

#### F - Distribution;

becomes relevant when we want to compare the variances of two samples

let  $X = \{x_1, x_2, ..., x_m\}$ ,  $Y = \{y_1, y_2, ..., y_m\}$  be the values of two independent random samples drawn from the same normal population, with variance  $\sigma^2$ 

Then
$$F = \frac{s_1^2}{s_2^2} \text{ where } s_1^2 = \frac{1}{m-1} \sum_{i=1}^{m} (\pi_i - \bar{\chi})^2$$

$$4 \quad s_2^2 = \frac{1}{m-1} \sum_{j=1}^{m} (y_j - \bar{y})^2$$

\$\frac{1}{\times 4 \times \frac{1}{\times 4 \times \frac{1}{\times 4 \times \frac{1}{\times 4 \times \frac{1}{\times 1} \times \frac{1}{\times 4 \times \frac{1}{\times 1} \ti

Remark . We always consider the larger of the two values of A1 4 A2 in the numeration

Ex. This samples of sizes 8 & 9 give a sum of squares of devibtions from their respective means equal to 91 inches 4 160 inches. Can these be regarded as drawn from the normal populations with the same variance?

Green 
$$n = 8$$
,  $m = 9$   
 $\sum (x - \overline{x})^2 = 91$ ,  $\sum (y - \overline{y})^2 = 160$   
 $\Delta_1^2 = \frac{1}{8 - 1}(91) = 13$ ;  $\Delta_2^2 = \frac{1}{9 - 1}(160) = 20$ 

to:  $\sigma_1^2 = \sigma_2^2$  12. the two samples come from the same

$$F = \frac{A_{2}^{2}}{\delta_{1}^{2}} = \frac{20}{13} = 1.5385 \text{ with } (9-1, 8-1) \text{ d.f.}$$

$$i_{5}. (8,7) \text{ d.f.}$$

$$F_{0,0,r}(8,7d.f) = 3.73$$

$$\begin{cases} \partial_1 = 8 \\ \partial_2 = 7 \end{cases}$$

Ex Ino independent samples of sizes 746 have the following values:

Examene if samples have been taken from normal populations having the same variance.

Means of the samples: \(\overline{\pi} = 31.286, \overline{\pi} = 28.167

$$\Delta_1^2 = \frac{1}{7-1} \ge (\pi (-3).286)^2 = 5.238$$

$$A^{2} = \frac{1}{6-1} \sum_{j} (y_{j}' - 28.167)^{2} = 5.367$$

$$f = \frac{5.367}{5.238} = 1.0245 \qquad (5,6 d.f.) \qquad \left( \frac{3}{3} = 5 \right)$$

Ex. Can we conclude that the two population HIW variances are equal for the following data of students who passed out from 'state' or 'private' universities

State: 8350 8260 8130 8340 8070

hivate: 7890 8140 7900 7950 7840 7920

Ex Two random samples of size 9 4 13 have s.d. is 2.1 4 1.8 respectively can these be considered to have been drawn from the same normal population?

Green n=9,  $S_1=2.1$ ; m=13,  $S_2=1.8$ 

 $\Delta_{1}^{2} = \frac{m}{m-1} S_{1}^{2} \qquad ; \quad \Delta_{2}^{2} = \frac{m}{m-1} S_{2}^{2}$   $= \frac{9}{8} (2 \cdot 1)^{2} \qquad = \frac{13}{12} (1 \cdot 8)^{2}$   $= 4 \cdot 961 \qquad = 3 \cdot 51$ 

 $F = \frac{4.961}{3.57} = 1.41$  (8,12 d.f.)

Fo.05 (8,12d.f.) = 2.85

Ho:  $\sigma_1^2 = \sigma_2^2$ 

As Fealenlated < Fo.05 (8, 120 +)

= Ho is accepted

the same subject Section A has 16 students 4B has 25 students, In an exam, though there was

no significant difference in mean grades, class A has a s.d of 9 whereas B has 12.

Can we conclude that variability in B is higher than #?

Green 
$$n = 16$$
,  $S_1 = 9$ ;  $m = 25$ ,  $S_2 = 12$   
to  $\sigma_1^2 = \sigma_2^2$ 

$$\Delta_{1}^{2} = \frac{16}{16-1}(9)^{2}$$

$$= 86.4$$

$$A_{2}^{2} = \frac{25}{25-1}(12)^{2}$$

$$= 150$$

$$F = \frac{\Delta_2^2}{\Delta_1^2} = \frac{150}{(86.4)^2} = 1.736 \quad (24, 15 d.f.)$$

H.W. ities is given below.

can we say that there is more variability in city B for daily wages?

#### Chi Square Fest:

(Non-parametric test of significance) It is used for testing;

- (1) Independence of Attributes
- (ii) Goodness of Fit

$$\chi_{i}^{2} = \frac{\left(Q_{i}^{2} - E_{i}\right)^{2}}{E_{i}}, \qquad \chi_{i}^{2} = \sum_{i} \chi_{i}^{2}$$

where Oi - observed frequency

Ei - expected frequency

I follows the square dist into (2-1)(c-1) d.f. Ex In a health survey, 400 individuals were asked if they were vaccinated against flu and whether they were subsequently attacked by flu. The following data was obtained test whether

vaccinali	on prever	ts f	lu.
	Vaccinoted	Not	nated Folal
Affacked	60	85	145
by flee Didn't get flu	190	65	255
flu Folat	250	120	400

Ho: The attributes are independent is vaccination does not prevent flue.

If the attributes are independent,

$$\begin{cases}
\text{then} & P(A \cap B) = P(A) P(B) \\
P(A \cap B) = \frac{60}{400} = 0.15
\end{cases}$$

$$P(A) P(B) = \left(\frac{250}{400}\right) \left(\frac{145}{400}\right) = 0.2265$$

How to calculate the (enperted values):

vaccinated, attacked by flu = (250) (145/400) × 400

= 90.625

To calculate X

			+
Cell	Observed freq. (Oi)	Expected frequency (Ei)	$\chi_i^2 = \frac{(0i - Ei)^2}{Ei}$
	60	$\frac{250 \times 145}{400} = 90.625$	10.349
2_	85	$\frac{150 \times 145}{400} = 54.375$	17. 248
3	190	255 X 250 = 159.375	5.885
4	-2.6	$\frac{255 \times 150}{400} = 95.625$	9.808
			$\sum_{i} \chi_{i}^{\dagger} = 43.29$

 $\chi^2 = 43.29$ 

$$2=2, C=2$$

2=2, c=2 (2-1)(c-1) df, = 1 df $\chi^2_{0.45}(|d.f.) = 3.84$ 

As X calculated > xo.os (1 d.f)

Ho is rejected

a catching flu is not indept of vaccination

En Perform Chi-square Fest of independence of attributes for the following data:

Eye	Mark_	colour	Total
Black	40	20	60
Blue	2 <i>0</i>	30	50
Brown	60	30	90
	120	80 [	200

Ho; The attributes are independent

Class	0 i	Ei	$\chi_{i}^{2} = \frac{\left(O_{i} - \vec{E}_{i}\right)^{2}}{E_{i}}$
1	4-0	$\frac{60 \times 120}{200} = 36$	0.4444
2	20	60 × 80 = 24	0.6667
3	20	50 × 120 = 30	3.33 33
4	30	50 × 80 = 20	5.0
5_ '	60	$\frac{90 \times 120}{200} = 57$	0.6667
6	30	$\frac{90 \times 80}{200} = 36$	· 0
			x= 11.1111

$$d.f. = (2-1)(c-1)$$

$$= (3-1)(2-1) = 2 d.f$$

$$V_{0,45}^{2}(2d.f.) = 5.99$$

As the > the > 2 (2d.f)

=> Ho is rejected

3) Eye & havi colours are dependent

# Chi-Square Fest for Goodness of Fit:

Ex A die is thrown 276 times and the results are given in the table

No. appeared on the die : 2 3 4 5 6

Frequency: 44 52 42 50 49 39

Fest if the die is unbiased.

Ho: The die is unbidsed

The expected frequency for each number under the mell hypotheses is equal to  $\frac{276}{6} (= 46)$ 

ND. On the die	Observed frep (Or)	Expected Freq (Ei)	$\chi_{i}^{2} = \frac{\left(O_{i} - E_{i}\right)^{2}}{E_{i}}$
<u> </u>	44	4-6	0.0869
2	52	46	0.7826
3	42	46	0.3478
4	50	46	6.3478
5	49	46	0.1956
6	39	46	1.0652
		_	$\chi^2 = 2.8260$

of f = 5 $\chi_{0.05}$  (5 d.f.) = 11.1

As Newholed < XO.O. (5d.f.) => Hoss accepted is die is unsubseed.

The demand for a particular space part in a shop was found to vary day to day. In a sample study, the following information was obtained Days: Mon Fre Wed Thu Fri Sat No. of parts: 124 125 110 120 126 115 Fest if the demand depends on the day of the week.  $E_i = \frac{124 + 125 + - - 115}{6} = \frac{120}{6}$ d.f. 5

Ex A sample analysis of examination results of 500 students was done It was found that 220 students had failed, 170 had secured third class, 90 had secured second class and 20 had secured first class. Do these figures support that the general examination result is in the ratio 4:3:2:1 for the various categories?

failed =  $\frac{4}{10} \times 500 = 200$ Expected Freq. III class =  $\frac{3}{10} \times 500 = 150$ II class =  $\frac{2}{10} \times 500 = 100$ 2 I class = 10×500 = 50 Category  $\varepsilon_{\iota}$  $\chi^{2} = 23.667$ to: Results are in the ratio 4:3:2:1 Fael 220 2 200  $III \subset$ 170 120 2,667 II C. 90 100 d.f. = 3 / Xo.os (3d.f.)=7.81 ユく 20 50

4. is rejected

18

Ex A survey of 800 families with 4 children each revealed the following distribution:

No. of boys: 0 1 2 3 4

No of girls: 4 3 2 1 0

No. of families: 32 178 290 236 64

Is the result consistent with the hypothesis that male 4 female buths are equally probable?

Ho: Male 4 Jemale births are equally probable  $\beta = 9 = \frac{1}{2}$ 

Expected free  $p(0boy, 4guls) = {}^{4}C_{0} p^{0} q^{4-0} = (\frac{1}{2})^{4}$ Exp. free for  $(0B, 4G) = 800 \times (\frac{1}{2})^{4} = 50$ 

 $P(18,34) = {}^{4}G + {}^{1}G + {}^{1}G = {}^{4}(\frac{1}{2})^{4} = \frac{1}{4}$ 

Exp for (18,39) = 800 x = 200

" "  $(28, 29) = 800 \times {}^{4}G_{2}(\frac{1}{2})^{4} = 300$ 

 $(38,19) = 800 \times \frac{1}{4} = 200$ 

" "  $(48,09) = 800 \times \frac{1}{16} = 50$ 

No. of boys in the family	OŁ	Eι	$\gamma'_{ij} = \frac{(O_{i} - O_{i})^{2}}{E}$	-Ei) -
	32	50	6,48	
1	178	200	2,42	
2	290	300	0-3333	X=19.6
3	236	200	6.48	,
4	64	5	3.9.7_	

d.f. = 4
$$\chi_{0.05}^{2} (4 d.f.) = 9.49$$

$$\chi_{0.05}^{2} (4 d.f.) = 7.49$$
As  $\chi_{0.05}^{2} (4 d.f.) = 7.49$ 

$$\chi_{0.05}^{2} (4 d.f.) = 7.49$$

2.6.21

Chi Square Fest for Goodness of Fit for Poisson Distribution:

Remark: If a test statistic is being calculated by making use of the given data (e.g. µ in Poisson dist.), then degrees of freedom associated with a sample of size n is n-2 (& not n-1) since we lose I degree of freedom for the calculation of a parameter.

Ex Fit a Poisson distribution for the given data and test the goodness of fit.

, ≥f= 1000

Ho: The given data follows Poisson distribution.

As 
$$\mu = \frac{\sum fx}{\sum f} = \frac{(0)419 + (1)352 + (2)(154) + 3(56) + 4(19)}{419 + 352 + 154 + 56 + 19}$$

As 
$$P(x) = \frac{e^{-\mu}\mu^{x}}{x!} = \frac{e^{-0.904}(0.904)^{x}}{x!}$$

$$P(x=0) = \frac{e^{-0.904}(1)}{(1)} = 0.4049$$

Expected free for n=0 is equal to (0.4049)(1000) = 404.9

	OL	Εί	$\chi_{i}^{*} = \frac{(0i - Ei)^{2}}{Ei}$
-	419	404.9466	0.4877
1	352	366.0717	0.5408
2	154	165.4644	0.79.43
3	56	49.8599	0.7564
4	19	11.2683	5.3056
1	<u> </u>		$\chi^2 = 7.8846$

$$d.f = 5 - 2 = 3$$

$$\chi^2_{0.05}(3d.f.) = 7.81$$

$$\chi^2_{0,01}(3d.f) = 11.3$$

Reject the at 5%, Accept at 1% los

Ex The number of accidents per day as recorded in a city over a period is given below. Fest if the data follows Poisson distribution

$$\mathcal{H} = \frac{\sum fx}{\sum f} = \frac{0(173) + 1(168) + \dots + 5(1)}{173 + 168 + 37 + 18 + 3 + 1} = 0.7825$$

$$A_{3} P m = \frac{e^{-\mu} \mu^{x}}{x!} = \frac{e^{-0.7815}}{x!}, \quad \Sigma f = 400$$

$$P(x=1) = 43.12$$

Ho: The data follows Porsson distribution

\ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \ \	Oi	EL	7L 2
0	173	182.9046	
1	168	14311228	
2	37	55.9968	
3	18	4.6058	
4	3	2.8573	
5	<u> </u>	0.4472	}

$$a.f. = 6 - 1 - 1 = 4$$

$$\int \chi^2 = 12.7811$$

		<del></del> _	
7	Οί	Ei	ti_
0		}	{
1	}		1
2			
1/3		7	
74	4-	[1-1] ×400	
<u></u>			-7

d.f. = 4 - 1 = 3  $\chi^{2}_{0.05}(3df) = ...$ 

1=-

Ho.

Ex The table below gives the number of students passed & failed by three examiners A,B,C in an exam. Fest the hypothesis that the proportion of students failed by the 3 examiners are equal

<del> </del>					Fotal
+	Passed	   50	47	56	153
	Failed	5	14	8	27
1	Column	55	61	64	G.T=180
	total	<u> </u>			

Ho: The proportion of students failed by examiners A, B 4 C are equal

$$d.f. = (3-1)(2-1) = 2$$

$\sqrt{i}$	نره	EL	1/2
1	50	$\frac{153 \times 55}{180} = 46.75$	
2	47	51.85	
3	57	54,4	
4	5	8.25	
2	14	9.15	' 
6	8	9.6	

$$\chi^2 = 4.8441$$

$$\chi^2_{0.05}$$
 (2 d.f.) = 5.99

Ex The table below gives the number of good & MD bad parts produced by each of the three shifts in a factory. Yest if the production of bad parts is independent of the shift in which they were produced

Sheft	Good parts	Berd parts
Day	960	4-0
Evening	940	50
Night	950	45

(Similar to the previous one)

Ex A set of 5 identical coins is tossed 320 times and the results are shown in the table

No. of heads: 0 1 2 3 4 5

Obs. Frequency(Oi): 16 27 72 112 71 22

Expected (EL):

Fest the goodness of fit for Binomial dist.

Ho: Benomial dist is a good fit for the data

	OL	ΕĹ	12=(01-E1)2
0	16	10	3.6
	27	50	10.58
2_	72	100	7,84
3	112	100	1.44
4	71	50	8.82
5	22	10	14,4
<u> </u>	<u> </u>		•

$$\textcircled{3}$$
 for  $z=0$ ,  $5\zeta_0(\frac{1}{2})^5 = \frac{1}{2^5} = \frac{1}{5}$  Frep $(x=0) = \frac{320}{2^5}$ 

x=1  $2C^{1}(\frac{r}{r})_{2}=20$ x=2 52(1)5 =

χ<sub>0.05</sub> (5 d.f.) = 11.1

=> Ho is rejected

Ex: In 90 throws of a die, face I turned up 9 times, face 200 3 turned 27 times, face 400 5 turned 3 6 times & 6 turned up 18 times. Fest at 5 % 1.0.1. & 11.1.0.1. if the die is fair

Ho: Die is fair

face turned

1

9  $4 \times 90 = 15$ 2013

27  $\frac{2}{6} \times 90 = 30$ 4015

36  $\frac{3}{4} \times 90 = 15$   $\frac{3}{4} \times 90 = 15$