

# Named Continuous Distribution Application Solutions

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## Exercises

For problems 1-3 below, **1)** define a random variable that will help you answer the question, **2)** state the distribution and parameters of that random variable; **3)** determine the expected value and variance of that random variable, and **4)** use that random variable to answer the question.

1. On a given Saturday, suppose vehicles arrive at the USAFA North Gate according to a Poisson process at a rate of 40 arrivals per hour.

a) Find the probability no vehicles arrive in 10 minutes.

$X$ : number of vehicles that arrive in 10 minutes

$X \sim \text{Pois}(\lambda = 40/6 = 6.67)$  and  $E(X) = \text{Var}(X) = 6.67$ .

$P(\text{no arrivals in 10 minutes}) = P(X = 0) = \frac{6.67^0 e^{-6.67}}{0!} = e^{-6.67}$

```
exp(-40/6)
```

```
## [1] 0.001272634
```

```
##or  
dpois(0,40/6)
```

```
## [1] 0.001272634
```

or, using the exponential distribution:

$Y$ : time in minutes until the next arrival

$Y \sim \text{Expon}(\lambda = 40/60 = 0.667)$  and  $E(Y) = 1.5$  and  $\text{Var}(Y) = 2.25$ .

$$P(\text{at least 10 minutes until the next arrival}) = P(Y \geq 10) = \int_{10}^{\infty} \frac{2}{3} e^{-\frac{2}{3}y} dy$$

```
1-pexp(10,2/3)
```

```
## [1] 0.001272634
```

or using simulation:

```
set.seed(616)
mean(rpois(100000,40/6) == 0)
```

```
## [1] 0.00126
```

```
mean(rexp(100000,2/3) >=10)
```

```
## [1] 0.00127
```

b) Find the probability that at least 5 minutes will pass before the next arrival.

$Y$ : same as in part a

$$P(\text{at least 5 minutes until next arrival}) = P(Y \geq 5) = \int_5^{\infty} \frac{2}{3} e^{-\frac{2}{3}y} dy$$

```
1-pexp(5,2/3)
```

```
## [1] 0.03567399
```

c) Find the probability that the next vehicle will arrive between 2 and 10 minutes from now.

Same  $Y$  as defined above.

```
pexp(10,2/3)-pexp(2,2/3)
```

```
## [1] 0.2623245
```

d) Find the probability that at least 7 minutes will pass before the next arrival, given that 2 minute have already passed. Compare this answer to part (b). This is an example of the memoryless property of the exponential distribution.

$$P(Y \geq 7 | Y \geq 2) = \frac{P(Y \geq 7, Y \geq 2)}{P(Y \geq 2)} = \frac{P(Y \geq 7)}{P(Y \geq 2)}$$

```
(1-pexp(7,2/3))/(1-pexp(2,2/3))
```

```
## [1] 0.03567399
```

This is the same answer and a result of the memoryless property.

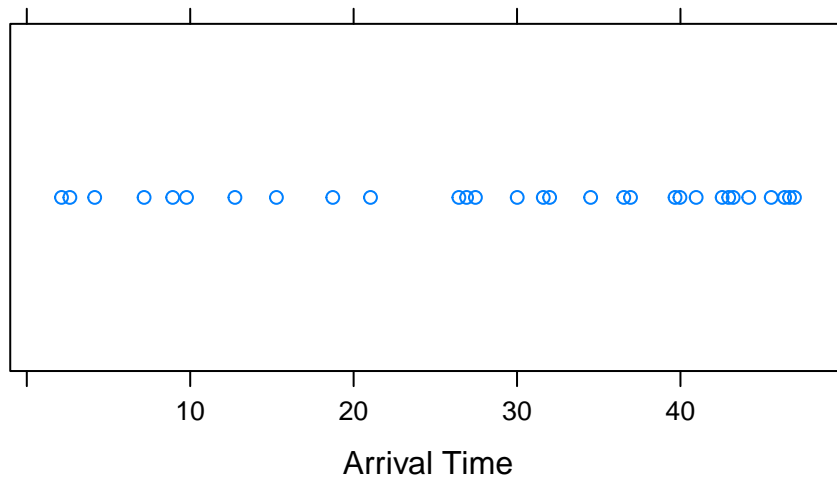
e) Fill in the blank. There is a probability of 90% that the next vehicle will arrive within \_\_\_\_ minutes. This value is known as the 90% percentile of the random variable.

```
qexp(0.9,2/3)
```

```
## [1] 3.453878
```

f) Use the function `stripplot()` to visualize the arrival of 30 vehicles using a random sample from the appropriate exponential distribution.

```
set.seed(202)
stripplot(cumsum(rexp(30,2/3)),xlab="Arrival Time")
```



2. Suppose time until computer errors on the F-35 follows a Gamma distribution with mean 20 hours and variance 10.

a) Find the probability that 20 hours pass without a computer error.

$X$ : time in hours until next computer error.

$X \sim \text{Gamma}(\alpha = 40, \lambda = 2)$

We need to find  $\alpha$  and  $\lambda$  from the given moments.

$$E(X) = 20 = \frac{\alpha}{\lambda}$$

$$\text{Var}(X) = 10 = \frac{\alpha}{\lambda^2}$$

Notice that  $\frac{E(X)}{\text{Var}(X)} = \lambda = \frac{20}{10} = 2$  and then using  $E(X) = 20 = \frac{\alpha}{\lambda}$  we get  $\alpha = 40$ .

$P(X \geq 20)$ :

```
1-pgamma(20,shape=40,rate=2)
```

```
## [1] 0.4789711
```

b) Find the probability that 45 hours pass without a computer error, given that 25 hours have already passed. Does the memoryless property apply to the Gamma distribution?

$$P(X \geq 45 | X \geq 25) = \frac{P(X \geq 45, X \geq 25)}{P(X \geq 25)} = \frac{P(X \geq 45)}{P(X \geq 25)}$$

```
(1-pgamma(45,40,2))/(1-pgamma(25,40,2))
```

```
## [1] 1.77803e-08
```

No, the memoryless property does not apply to the Gamma distribution.

c) Find  $a$  and  $b$  where there is a 95% probability that the time until next computer error will be between  $a$  and  $b$ . (Note: technically, there are many answers to this question, but find  $a$  and  $b$  such that each tail has equal probability.)

```
qgamma(c(0.025,0.975),40,2)
```

```
## [1] 14.28829 26.65714
```

So in the time interval  $[14, 29, 26.66]$ .

```
qgamma(.95,40,2)
```

```
## [1] 25.46987
```

Another answer is between  $[0, 25, 47]$ .

3. Suppose PFT scores in the cadet wing follow a normal distribution with mean 330 and standard deviation 50.

a) Find the probability a randomly selected cadet has a PFT score higher than 450.

$X$ : PFT score of a randomly selected cadet

$X \sim \text{Norm}(\mu = 330, \sigma = 50)$

$E(X) = 330$  and  $\text{Var}(X) = 50^2 = 2500$ .

```
1-pnorm(450,330,50)
```

```
## [1] 0.008197536
```

b) Find the probability a randomly selected cadet has a PFT score within 2 standard deviations of the mean.

Need  $P(230 \leq X \leq 430)$ .

```
pnorm(430,330,50)-pnorm(230,330,50)
```

```
## [1] 0.9544997
```

c) Find  $a$  and  $b$  such that 90% of PFT scores will be between  $a$  and  $b$ .

Need  $a$  such that  $P(X \leq a) = 0.05$  and  $b$  such that  $P(X \geq b) = 0.05$ :

```
qnorm(0.05,330,50)
```

```
## [1] 247.7573
```

```
qnorm(0.95,330,50)
```

```
## [1] 412.2427
```

d) Find the probability a randomly selected cadet has a PFT score higher than 450 given he/she is among the top 10% of cadets.

Need  $P(X > 450 | X > x_{0.9})$  where  $x_{0.9}$  is the 90th percentile of  $X$ .

The 90th percentile is:

```
qnorm(0.9,330,50)
```

```
## [1] 394.0776
```

$$P(X > 450 | X > x_{0.9}) = \frac{P(X > 450, X > x_{0.9})}{P(X > x_{0.9})} = \frac{P(X > 450, X > 394.08)}{P(X > x_{0.9})} = \frac{P(X > 450)}{0.1}$$

This is assuming that  $x_{0.9} < 450$ . Otherwise the problem is trivial and the probability is 1.

```
(1-pnorm(450,330,50))/0.1
```

```
## [1] 0.08197536
```

4. Let  $X \sim \text{Beta}(\alpha = 1, \beta = 1)$ . Show that  $X \sim \text{Unif}(0, 1)$ . Hint: write out the beta distribution pdf where  $\alpha = 1$  and  $\beta = 1$ .

The beta pdf is:

$$f_X(x) = \frac{\Gamma(\alpha + \beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1}$$

When  $X \sim \text{Beta}(\alpha = 1, \beta = 1)$ , this becomes:

$$f_X(x) = \frac{\Gamma(2)}{\Gamma(1)\Gamma(1)} x^{1-1} (1-x)^{1-1} = 1$$

5. When using R to calculate probabilities related to the gamma distribution, we often use `pgamma`. Recall that `pgamma` is equivalent to the cdf of the gamma distribution. If  $X \sim \text{Gamma}(\alpha, \lambda)$ , then

$$P(X \leq x) = \text{pgamma}(x, \alpha, \lambda)$$

The `dgamma` function exists in R too. In plain language, explain what `dgamma` returns. I'm not looking for the definition found in R documentation. I'm looking for a simple description of what that function returns. Is the output of `dgamma` useful? If so, how?

The `dgamma` function returns the value of probability density function. While this is not a probability, it is still a useful quantity. It can be said that larger densities ( $f(x)$ ) imply that values near  $x$  are more likely to occur than values associated with smaller densities. It is also useful when computing conditional probability distributions.

6. Advanced. You may have heard of the 68-95-99.7 rule. This is a helpful rule of thumb that says if a population has a normal distribution, then 68% of the data will be within one standard deviation of the mean, 95% of the data will be within two standard deviations and 99.7% of the data will be within three standard deviations. Create a function in **R** that has two inputs (a mean and a standard deviation). It should return a vector with three elements: the probability that a randomly selected observation from the normal distribution with the inputted mean and standard deviation lies within one, two and three standard deviations. Test this function with several values of *mu* and *sd*. You should get the same answer each time.

```
rulethumb<-function(mu,sd){
  pnorm(mu+c(1,2,3)*sd,mu,sd)-pnorm(mu-c(1,2,3)*sd,mu,sd)
}

rulethumb(15,12)
```

```
## [1] 0.6826895 0.9544997 0.9973002
```

7. Derive the mean of a general uniform distribution,  $U(a, b)$ .

From the definition

$$\begin{aligned}
 E(X) &= \int_a^b xf(x)dx = \\
 &= \int_a^b \frac{x}{b-a}dx = \\
 &= \frac{1}{b-a} \int_a^b xdx = \frac{1}{b-a} \cdot \frac{x^2}{2} \Big|_a^b = \\
 &= \frac{1}{b-a} \cdot \frac{b^2 - a^2}{2} = \frac{1}{b-a} \cdot \frac{(b-a)(b+a)}{2} = \frac{(b+a)}{2}
 \end{aligned}$$

## File creation information

- File creation date: 2020-09-14
- Windows version: Windows 10 x64 (build 18362)
- R version 3.6.3 (2020-02-29)
- **mosaic** package version: 1.7.0
- **tidyverse** package version: 'r