Additional Hypothesis Test Applications Solutions

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Exercises

1. Golf balls

Repeat the analysis of the golf ball problem from earlier this semester.

a. Load the data and tally the data into a table. The data is in golf_balls.csv.

```
golf_balls <- read_csv("data/golf_balls.csv")</pre>
```

```
head(golf_balls)
```

```
tally(~number,data=golf_balls)
```

```
## number
## 1 2 3 4
## 137 138 107 104
```

b. Using the function chisq.test conduct a hypothesis test of equally likely distribution of balls. You may have to read the help menu.

```
chisq.test(tally(~number,data=golf_balls),p=c(.25,.25,.25,.25))
```

```
##
## Chi-squared test for given probabilities
##
## data: tally(~number, data = golf_balls)
## X-squared = 8.4691, df = 3, p-value = 0.03725
```

c. Repeat part b. but assume balls with the numbers 1 and 2 occur 30% of the time and balls with 3 and 4 occur 20%.

```
chisq.test(tally(~number,data=golf_balls),p=c(.3,.3,.2,.2))
```

```
##
## Chi-squared test for given probabilities
##
## data: tally(~number, data = golf_balls)
## X-squared = 2.4122, df = 3, p-value = 0.4914
```

2. Bootstrap hypothesis testing

Repeat the analysis of the MLB data from the less on but this time generate a bootstrap distribution of the F statistic.

Read in the data.

```
mlb_obp <- read_csv("data/mlb_obp.csv")</pre>
```

```
mlb_obp <- mlb_obp %>%
mutate(position=as.factor(position))
```

```
favstats(obp~position,data=mlb_obp)
```

```
Q1 median
                                                                         n missing
##
    position
                                        QЗ
                min
                                             max
                                                       mean
           C 0.219 0.30000 0.3180 0.35700 0.405 0.3226154 0.04513175
## 1
                                                                        39
## 2
           DH 0.287 0.31625 0.3525 0.36950 0.412 0.3477857 0.03603669
                                                                                 0
## 3
           IF 0.174 0.30800 0.3270 0.35275 0.437 0.3315260 0.03709504 154
                                                                                 0
           OF 0.265 0.31475 0.3345 0.35300 0.411 0.3342500 0.02944394 120
## 4
                                                                                 0
```

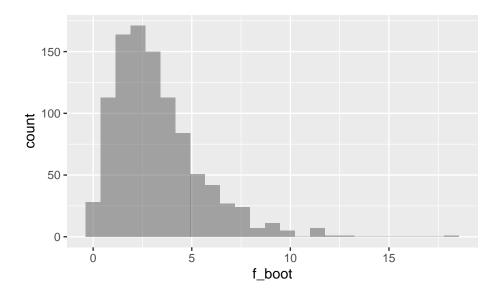
We need a function to resample the data, we will use the resample() from the mosaic package.

library(broom)

```
f_boot <- function(x){
  aov(obp~position,data=resample(x)) %>%
  tidy() %>%
  summarize(stat=meansq[1]/meansq[2]) %>%
  pull()
}
```

```
set.seed(541)
results<-do(1000)*f_boot(mlb_obp)</pre>
```

```
results %>%
gf_histogram(~f_boot)
```



Now the confidence interval for the F-statistic is:

```
cdata(~f_boot,data=results)
```

```
## lower upper central.p
## 2.5% 0.3546682 8.724895 0.95
```

We are 95% confident that the F statistic is in the interval (0.35, 8.72) which includes 1 so we fail to reject.

3. Test of variance

We have not performed a test of variance so we will create our own.

a. Using the MLB from the lesson, subset on IF and OF.

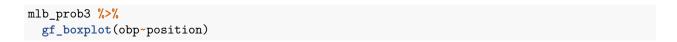
```
mlb_prob3 <- mlb_obp %>%
  filter(position=="IF" | position=="OF") %>%
  droplevels()
```

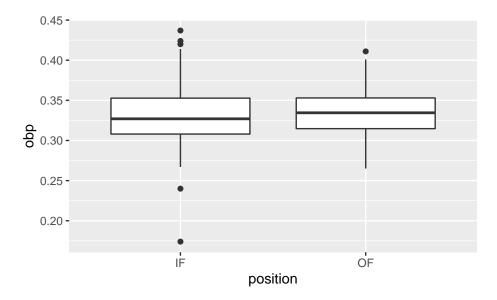
summary(mlb_prob3)

```
##
    position
                   obp
##
    IF:154
              Min.
                     :0.1740
    OF:120
              1st Qu.:0.3100
##
##
              Median :0.3310
##
              Mean
                      :0.3327
              3rd Qu.:0.3530
##
              Max.
                      :0.4370
```

The function droplevels() gets rid of C and DH in the factor levels.

b. Create a side-by-side boxplot.





The hypotheses are: H_0 : $\sigma_{IF}^2 = \sigma^2 OF$. There is no difference in the variance of on base percentage for infielders and outfielders. H_A : $\sigma_{IF}^2 \neq \sigma^2 OF$. There is a difference in variances.

c. Use the differences in sample standard deviations as your test statistic. Using a permutation test, find the p-value and discuss your decision.

```
mlb_prob3 %>%
  group_by(position) %>%
  summarize(stat=sd(obp))
## 'summarise()' ungrouping output (override with '.groups' argument)
## # A tibble: 2 x 2
     position
                stat
     <fct>
               <dbl>
## 1 IF
              0.0371
## 2 OF
              0.0294
obs <- mlb_prob3 %>%
  summarize(stat=sd(obp[position=="IF"])-sd(obp[position=="OF"])) %>%
  pull()
obs
```

[1] 0.007651101

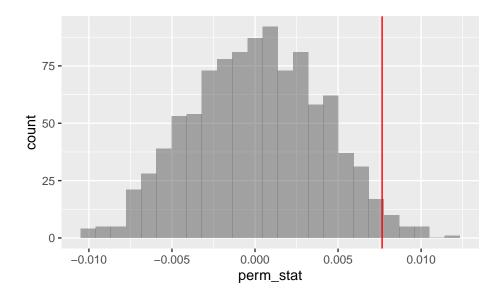
Let's write a function to shuffle the position.

```
perm_stat <- function(x){
   x %>%
   mutate(position=shuffle(position)) %>%
   summarize(stat=sd(obp[position=="IF"])-sd(obp[position=="OF"])) %>%
   pull()
}
```

```
set.seed(443)
results<-do(1000)*perm_stat(mlb_prob3)</pre>
```

```
results %>%
  gf_histogram(~perm_stat) %>%
  gf_vline(xintercept=obs,color="red")
```

Warning: geom_vline(): Ignoring 'mapping' because 'xintercept' was provided.



The p-value is

```
2*prop1(~(perm_stat>=obs),data=results)
```

```
## prop_TRUE
## 0.04395604
```

This is a two sided test since we did not know in advance which variance would be larger. We reject the hypothesis of equal variance but the p-value is too close. The conclusion is suspect. We need more data.

d. Create a bootstrap distribution of the differences in sample standard deviations, and report a 95% confidence interval. Compare with part c.

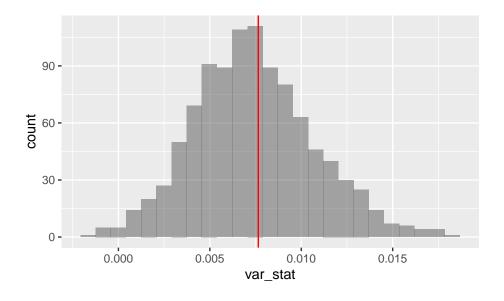
Let's write a function.

```
var_stat <- function(x){
  resample(x) %>%
  summarize(stat=sd(obp[position=="IF"])-sd(obp[position=="OF"])) %>%
  pull()
}
```

```
set.seed(827)
results<-do(1000)*var_stat(mlb_prob3)</pre>
```

```
results %>%
  gf_histogram(~var_stat) %>%
  gf_vline(xintercept=obs,color="red")
```

Warning: geom_vline(): Ignoring 'mapping' because 'xintercept' was provided.



File Creation Information

 \bullet File creation date: 2020-07-28

• Windows version: Windows 10 x64 (build 18362)

R version 3.6.3 (2020-02-29)
mosaic package version: 1.7.0
tidyverse package version: 1.3.0

• openintro package version: 2.0.0

• broom package version: 0.7.0