

Stat453_Assignment03

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Questions

Question 1. The yield of a chemical process is being studied. The two most important variables are thought to be the pressure and the temperature. Three levels of each factor are selected, and a factorial experiment with two replicates is performed. The yield data follow:

Temperature	Pressure		
	200	215	230
150	90.4	90.7	90.2
	90.2	90.6	90.4
160	90.1	90.5	89.9
	90.3	90.6	90.1
170	90.5	90.8	90.4
	90.7	90.9	90.1

```
rm(list = ls())
Temperature=rep(c(rep(150,2),rep(160,2),rep(170,2)),3)
Pressure<-c(rep(200,6), rep(215,6),rep(230,6))
Yield = c(90.4, 90.2, 90.1, 90.3, 90.5, 90.7,
          90.7, 90.6, 90.5, 90.6, 90.8, 90.9,
          90.2, 90.4, 89.9, 90.1, 90.4, 90.1)
q1_data <- data.frame(Yield, Temperature, Pressure)

#Using ANOVA

res.aov <-aov(Yield~ factor(Temperature)*factor(Pressure),data=q1_data)
summary(res.aov)

##              Df Sum Sq Mean Sq F value    Pr(>F)
## factor(Temperature)      2  0.3011   0.1506     8.469 0.008539 **
## factor(Pressure)         2  0.7678   0.3839    21.594 0.000367 ***
## factor(Temperature):factor(Pressure)  4  0.0689   0.0172     0.969 0.470006
## Residuals                9  0.1600   0.0178
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
res.aov$coefficients
```

```
##               (Intercept)
##               9.030000e+01
##      factor(Temperature)160
##               -1.000000e-01
##      factor(Temperature)170
##               3.000000e-01
##      factor(Pressure)215
##               3.500000e-01
##      factor(Pressure)230
##               -8.888948e-15
## factor(Temperature)160:factor(Pressure)215
##               1.573213e-14
## factor(Temperature)170:factor(Pressure)215
##               -1.000000e-01
## factor(Temperature)160:factor(Pressure)230
##               -2.000000e-01
## factor(Temperature)170:factor(Pressure)230
##               -3.500000e-01
```

#Multiple comparisons

```
TUKEY <- TukeyHSD(x=res.aov, conf.level=0.95)
TUKEY
```

```
##      Tukey multiple comparisons of means
##      95% family-wise confidence level
##
## Fit: aov(formula = Yield ~ factor(Temperature) * factor(Pressure), data = q1_data)
##
## $'factor(Temperature)'
```

	diff	lwr	upr	p adj
160-150	-0.1666667	-0.38159536	0.04826203	0.1313152
170-150	0.1500000	-0.06492869	0.36492869	0.1809078
170-160	0.3166667	0.10173797	0.53159536	0.0066518

```
##
## $'factor(Pressure)'
```

	diff	lwr	upr	p adj
215-200	0.3166667	0.1017380	0.53159536	0.0066518
230-200	-0.1833333	-0.3982620	0.03159536	0.0944905
230-215	-0.5000000	-0.7149287	-0.28507131	0.0002951

```
##
## $'factor(Temperature):factor(Pressure)'
```

	diff	lwr	upr	p adj
160:200-150:200	-1.000000e-01	-0.62747453	0.42747453	0.9959112
170:200-150:200	3.000000e-01	-0.22747453	0.82747453	0.4489114
150:215-150:200	3.500000e-01	-0.17747453	0.87747453	0.2916707
160:215-150:200	2.500000e-01	-0.27747453	0.77747453	0.6429825
170:215-150:200	5.500000e-01	0.02252547	1.07747453	0.0397898
150:230-150:200	-1.421085e-14	-0.52747453	0.52747453	1.0000000
160:230-150:200	-3.000000e-01	-0.82747453	0.22747453	0.4489114
170:230-150:200	-5.000000e-02	-0.57747453	0.47747453	0.9999713
170:200-160:200	4.000000e-01	-0.12747453	0.92747453	0.1812587
150:215-160:200	4.500000e-01	-0.07747453	0.97747453	0.1099996
160:215-160:200	3.500000e-01	-0.17747453	0.87747453	0.2916707

```
## 170:215-160:200 6.500000e-01 0.12252547 1.17747453 0.0147551
## 150:230-160:200 1.000000e-01 -0.42747453 0.62747453 0.9959112
## 160:230-160:200 -2.000000e-01 -0.72747453 0.32747453 0.8316762
## 170:230-160:200 5.000000e-02 -0.47747453 0.57747453 0.9999713
## 150:215-170:200 5.000000e-02 -0.47747453 0.57747453 0.9999713
## 160:215-170:200 -5.000000e-02 -0.57747453 0.47747453 0.9999713
## 170:215-170:200 2.500000e-01 -0.27747453 0.77747453 0.6429825
## 150:230-170:200 -3.000000e-01 -0.82747453 0.22747453 0.4489114
## 160:230-170:200 -6.000000e-01 -1.12747453 -0.07252547 0.0241003
## 170:230-170:200 -3.500000e-01 -0.87747453 0.17747453 0.2916707
## 160:215-150:215 -1.000000e-01 -0.62747453 0.42747453 0.9959112
## 170:215-150:215 2.000000e-01 -0.32747453 0.72747453 0.8316762
## 150:230-150:215 -3.500000e-01 -0.87747453 0.17747453 0.2916707
## 160:230-150:215 -6.500000e-01 -1.17747453 -0.12252547 0.0147551
## 170:230-150:215 -4.000000e-01 -0.92747453 0.12747453 0.1812587
## 170:215-160:215 3.000000e-01 -0.22747453 0.82747453 0.4489114
## 150:230-160:215 -2.500000e-01 -0.77747453 0.27747453 0.6429825
## 160:230-160:215 -5.500000e-01 -1.07747453 -0.02252547 0.0397898
## 170:230-160:215 -3.000000e-01 -0.82747453 0.22747453 0.4489114
## 150:230-170:215 -5.500000e-01 -1.07747453 -0.02252547 0.0397898
## 160:230-170:215 -8.500000e-01 -1.37747453 -0.32252547 0.0023855
## 170:230-170:215 -6.000000e-01 -1.12747453 -0.07252547 0.0241003
## 160:230-150:230 -3.000000e-01 -0.82747453 0.22747453 0.4489114
## 170:230-150:230 -5.000000e-02 -0.57747453 0.47747453 0.9999713
## 170:230-160:230 2.500000e-01 -0.27747453 0.77747453 0.6429825
```

(a) Analyze the data and draw conclusions. Use $\alpha = 0.05$.

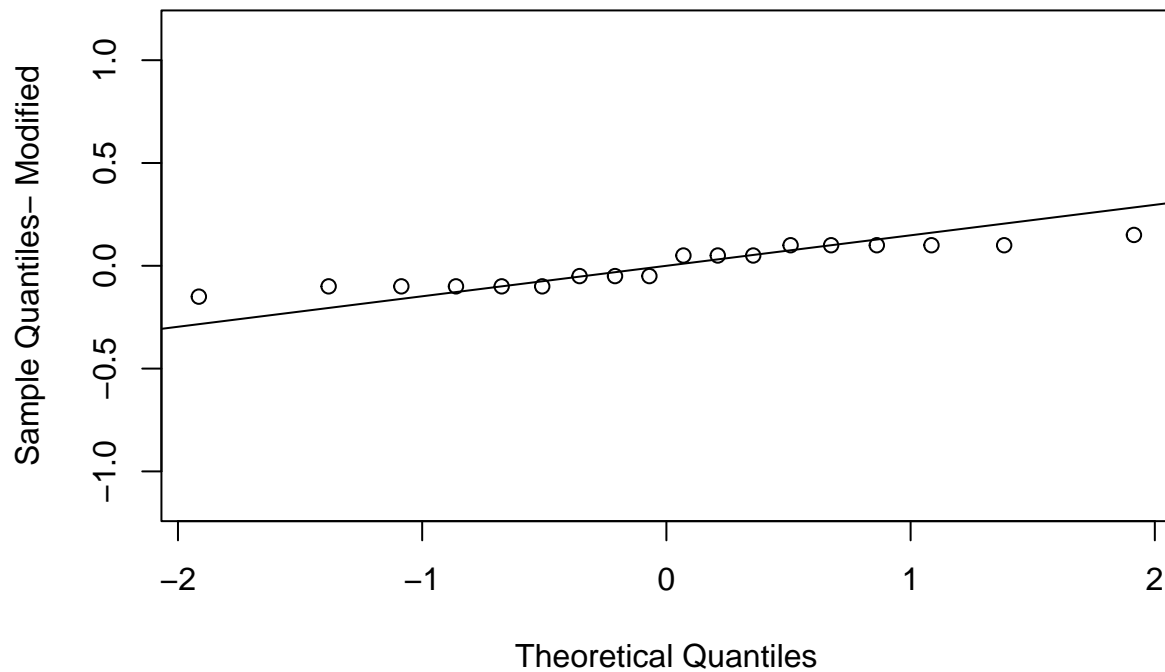
\therefore The p-value (0.470006) for the interaction term is bigger than 0.05, therefore the effect of interaction is not significant. So we can just look at the Tukey test result for the main effects. Temperature: p-value (0.0066518) for the difference between levels 170 and 160 $< \alpha = 0.05$, therefore the effect of temperature is significant. Pressure: p-values for the difference between levels are smaller than $\alpha = 0.05$ except for the difference between 230 and 200 (p-value: 0.0944905), therefore the effect of pressure is significant. So both temperature and pressure are significant factors, but interaction is not significant.

(b) Prepare appropriate residual plots and comment on the model's adequacy.

```
#Normality
residuals=res.aov$residuals
qqnorm(residuals, ylim=c(min(residuals)-1,max(residuals)+1), main = "Normal Q-Q Plot for Residuals",
       xlab = "Theoretical Quantiles", ylab = "Sample Quantiles- Modified",
       plot.it = TRUE, datax = FALSE)

qqline(residuals, datax = FALSE, distribution = qqnorm)
```

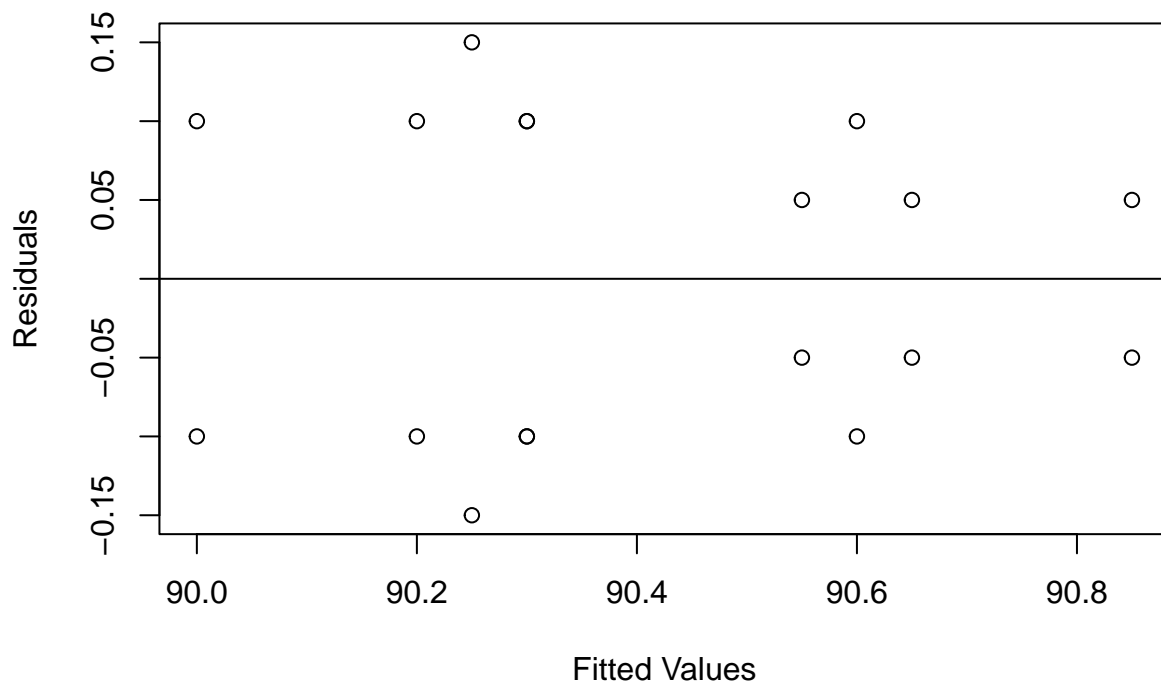
Normal Q-Q Plot for Residuals



```
#Test normality using Shapiro Wilks  
shapiro.test(residuals)
```

```
##  
##  Shapiro-Wilk normality test  
##  
## data:  residuals  
## W = 0.87366, p-value = 0.02046
```

```
#Check Variance  
Fitted_values=res.aov$fitted.values  
plot(Fitted_values,residuals,ylab="Residuals",xlab="Fitted Values")  
abline(h=0)
```



∴ Normality plot does not seem to violate normality assumption but the p-value of the shapiro test is smaller than $\alpha=0.05$, so it might not be normally distributed. Therefore there's an indication of violation of the normality assumption. Since the variance plots seem spread out, there's no indication of violation of the constant variance assumption.

(c) Under what conditions would you operate this process (i.e., which conditions maximize the yield)?

```
q1_data["Fitted"] = res.aov$fitted.values
q1_data[res.aov$fitted.values == max(res.aov$fitted.values),]
```

```
##      Yield Temperature Pressure Fitted
## 11  90.8           170       215  90.85
## 12  90.9           170       215  90.85
```

∴ Since the interaction term is not significant, we don't need to worry about the effect of interaction, so temperature at 170 (since the diff of 170-160 is positive and significant) and pressure at 215 (since the diff of 215-200 is positive and significant) would maximize the yield.

Question 2. The C. F. Eye Care company manufactures lenses for transplantation into the eye following cataract surgery. An engineering group has conducted an experiment involving two factors to determine their effect on the lens polishing process. The results of this experiment are summarized in the following ANOVA display:

Source	<i>DF</i>	<i>SS</i>	<i>MS</i>	<i>F</i>	<i>P</i> - value
Factor A	?	?	0.0833	0.05	0.952
Factor B	?	96.333	96.333	57.80	<0.001
Interaction	2	12.167	6.0833	3.65	?
Error	6	10.000	?		
Total	11	118.667			

(a) Complete the table

```
dfB = floor(96.333/96.333) #1
dfA = 2/dfB #2
SSA = 0.0833*dfA #0.1666
MSE = 10/6 #1.66667
PInteraction = pf(3.65, 2, 6, lower.tail = FALSE) #0.09181187
```

Source	<i>DF</i>	<i>SS</i>	<i>MS</i>	<i>F</i>	<i>P</i> - value
Factor A	2	0.1666	0.0833	0.05	0.952
Factor B	1	96.333	96.333	57.80	<0.001
Interaction	2	12.167	6.0833	3.65	0.0918
Error	6	10.000	1.6667		
Total	11	118.667			

(b) How many replicates are in this experiment?

\therefore Since $a-1 = 2$ and $b-1 = 1$, $a = 3$ and $b = 2$. Since $df(\text{error}) = 6 = ab(n-1)$ and $df(\text{Total}) = abn-1 = 11$, $n = 2$. Therefore there are 2 replicates in this experiment.

(c) Does the effect of factor B depend on the level of factor A? Justify.

\therefore Since the p-value of the interaction term is bigger than $\alpha=0.05$, the effect of the interaction term is not significant which shows that the effect of factor B does not depend on the level of factor A.

Question 3. An experiment was performed to improve the yield of a chemical process. Four factors were selected, and two replicates of a completely randomized experiment were run. The results are shown in the following table:

Treatment Combination	Replicate		Treatment Combination	Replicate	
	I	II		I	II
(1)	90	93	<i>d</i>	98	95
<i>a</i>	74	78	<i>ad</i>	72	76
<i>b</i>	81	85	<i>bd</i>	87	83
<i>ab</i>	83	80	<i>abd</i>	85	86
<i>c</i>	77	78	<i>cd</i>	99	90
<i>ac</i>	81	80	<i>acd</i>	79	75
<i>bc</i>	88	82	<i>bcd</i>	87	84
<i>abc</i>	73	70	<i>abcd</i>	80	80

```
A=rep(c(-1,1),8)
B=rep(c(rep(-1,2),rep(1,2)),4)
C=rep(c(rep(-1,4),rep(1,4)),2)
D=c(rep(-1,8),rep(1,8))
rep1 = c(90, 74, 81, 83, 77, 81, 88, 73,
          98, 72, 87, 85, 99, 79, 87, 80)
rep2 = c(93, 78, 85, 80, 78, 80, 82, 70,
          95, 76, 83, 86, 90, 75, 84, 80)
A = rep(A, rep(2, 16))
B = rep(B, rep(2, 16))
C = rep(C, rep(2, 16))
D = rep(D, rep(2, 16))
yield = cbind(rep1,rep2)
yield.vec = c(t(yield))
yield_data = data.frame(A, B, C, D, yield.vec)
```

(a) Estimate the factor effects.

```
#Using ANOVA
res.aov <-aov(yield.vec~A*B*C*D,data=yield_data)
res.aov$coefficients
```

```
## (Intercept)          A          B          C          D          A:B
## 82.78125    -4.53125   -0.65625   -1.34375   1.96875   2.03125
##          A:C          B:C          A:D          B:D          C:D          A:B:C
## 0.34375    -0.28125   -1.09375   -0.09375   0.84375   -2.59375
##          A:B:D          A:C:D          B:C:D          A:B:C:D
## 2.34375    -0.46875   -0.46875   1.21875
```

```
res.aov$coefficients[-1]*2 #factor effects
```

```
##          A          B          C          D          A:B          A:C          B:C          A:D          B:D          C:D
```

```
## -9.0625 -1.3125 -2.6875  3.9375  4.0625  0.6875 -0.5625 -2.1875 -0.1875  1.6875
##   A:B:C   A:B:D   A:C:D   B:C:D A:B:C:D
## -5.1875  4.6875 -0.9375 -0.9375  2.4375
```

Factor effects are as follows: A = -9.0625, B = -1.3125, C = -2.6875, D = 3.9375, AB = 4.0625, AC = 0.6875, BC = -0.5625, AD = -2.1875, BD = -0.1875, CD = 1.6875, ABC = -5.1875, ABD = 4.6875, ACD = -0.9375, BCD = -0.9375, ABCD = 2.4375.

(b) Prepare an analysis of variance table, and determine which factors are important in explaining yield.

```
res.aov = aov(yield.vec~A*B*C*D, data=yield_data)
summary(res.aov)
```

```
##           Df Sum Sq Mean Sq F value    Pr(>F)
## A           1   657.0    657.0   85.816 7.87e-08 ***
## B           1    13.8     13.8    1.800 0.198445
## C           1    57.8     57.8    7.547 0.014317 *
## D           1   124.0    124.0   16.200 0.000979 ***
## A:B         1   132.0    132.0   17.245 0.000749 ***
## A:C         1     3.8      3.8    0.494 0.492302
## B:C         1     2.5      2.5    0.331 0.573296
## A:D         1    38.3     38.3    5.000 0.039945 *
## B:D         1     0.3      0.3    0.037 0.850417
## C:D         1    22.8     22.8    2.976 0.103793
## A:B:C       1   215.3    215.3   28.118 7.15e-05 ***
## A:B:D       1   175.8    175.8   22.959 0.000200 ***
## A:C:D       1     7.0      7.0    0.918 0.352162
## B:C:D       1     7.0      7.0    0.918 0.352162
## A:B:C:D     1    47.5     47.5    6.208 0.024077 *
## Residuals   16   122.5      7.7
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

∴ Factors and interactions A, C, D, AB, AD, ABC, ABD, ABCD are significant.

(c) Write down a regression model for predicting yield, assuming that all four factors were varied over the range from -1 to +1 (in coded units).

```
res.aov$coefficients
```

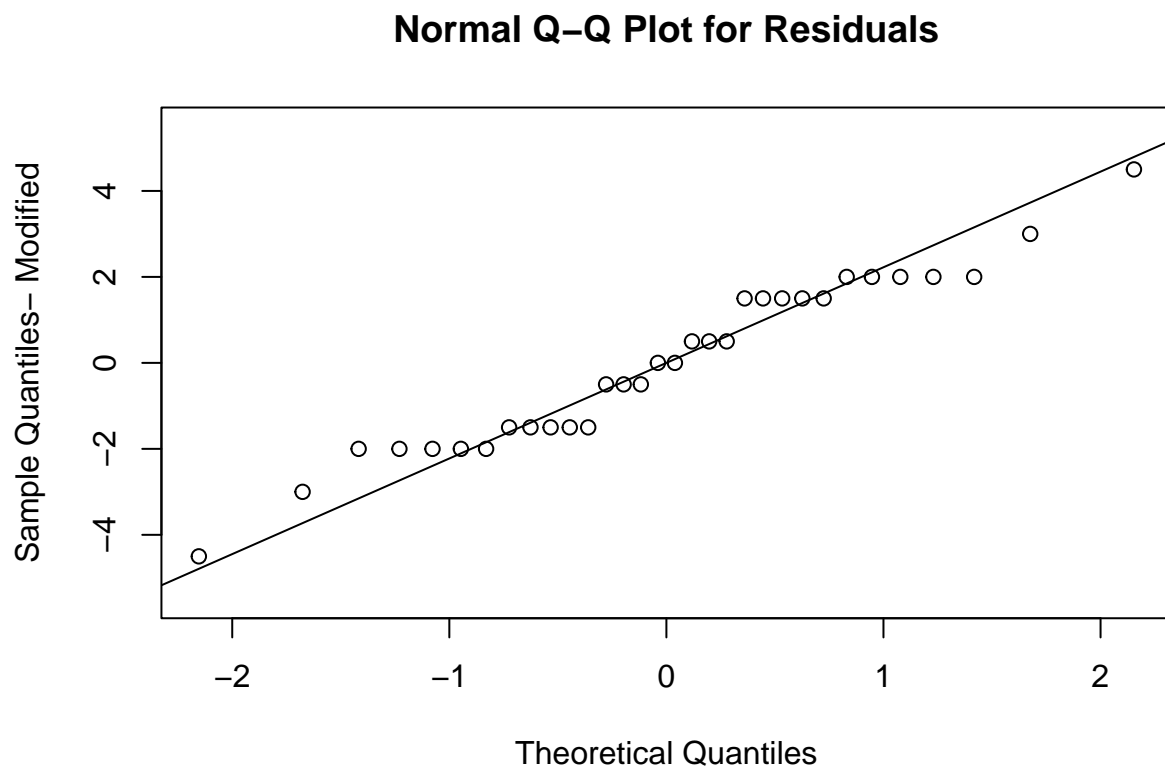
```
## (Intercept)          A          B          C          D          A:B
##   82.78125   -4.53125   -0.65625   -1.34375    1.96875    2.03125
##           A:C          B:C          A:D          B:D          C:D          A:B:C
##    0.34375   -0.28125   -1.09375   -0.09375    0.84375   -2.59375
##           A:B:D          A:C:D          B:C:D          A:B:C:D
##    2.34375   -0.46875   -0.46875    1.21875
```

yield = 82.78125 - 4.53125A - 1.34375C + 1.96875D + 2.03125AB - 1.09375AD - 2.59375ABC + 2.34375ABD + 1.21875ABCD

(d) Does the residual analysis appear satisfactory?

```
#residual analysis
#Normality
res.aov = aov(yield.vec~A*B*C*D, data=yield_data)
residuals=res.aov$residuals
qqnorm(residuals, ylim=c(min(residuals)-1,max(residuals)+1), main = "Normal Q-Q Plot for Residuals",
        xlab = "Theoretical Quantiles", ylab = "Sample Quantiles- Modified",
        plot.it = TRUE, datax = FALSE)

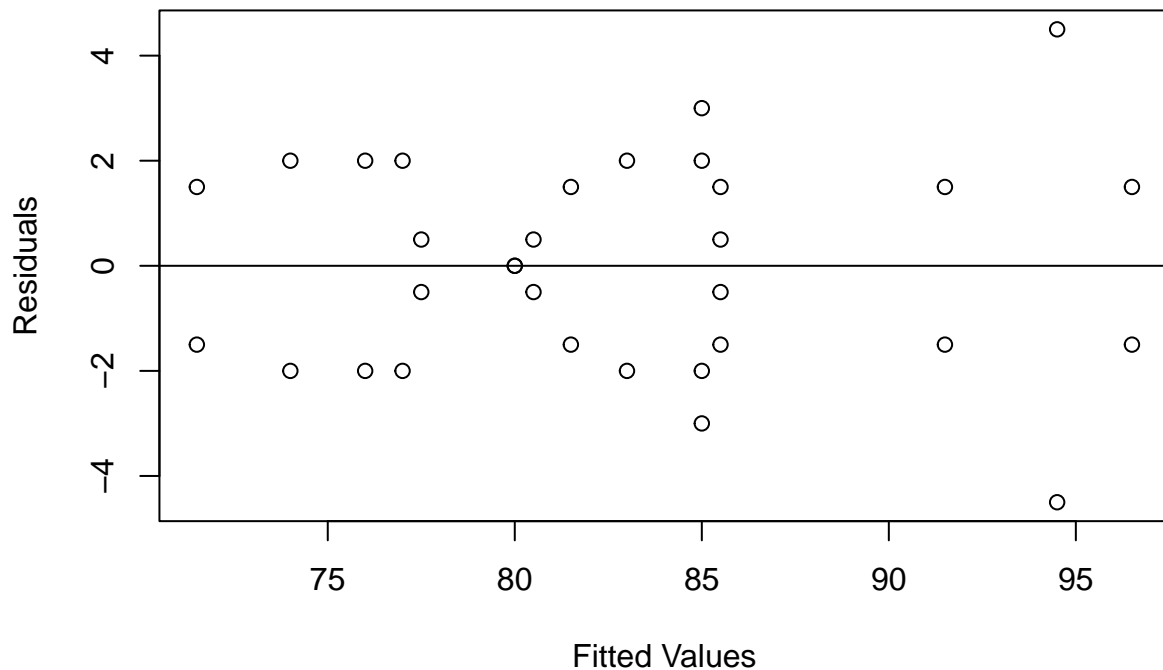
qqline(residuals, datax = FALSE, distribution = qnorm)
```



```
#Test normality using Shapiro Wilks
shapiro.test(residuals)
```

```
##
##  Shapiro-Wilk normality test
##
## data:  residuals
## W = 0.96135, p-value = 0.2989
```

```
#Check Variance
Fitted_values=res.aov$fitted.values
plot(Fitted_values,residuals,ylab="Residuals",xlab="Fitted Values")
abline(h=0)
```



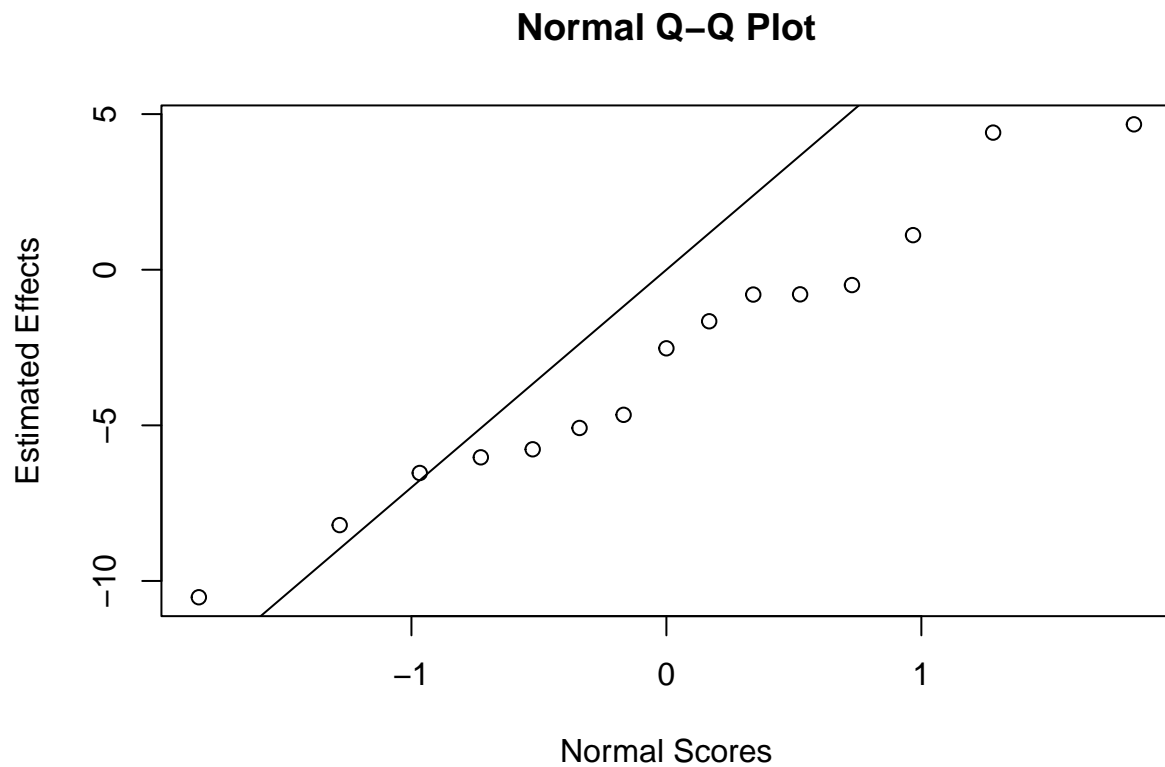
\therefore Since the p-value from Shapiro Wilks test is $0.2989 > 0.05$, we fail to reject the null hypothesis that the residuals are normally distributed. Therefore it is normally distributed. There's an indication of violation of the constant variance assumption since there seems to be a pattern in Fitted values vs. residuals plot.

Question 4. The effect estimates from a 2^4 factorial experiment are listed here.

$ABCD = -2.5251$	$AD = -1.6564$
$BCD = 4.4054$	$AC = 1.1109$
$ACD = -0.4932$	$AB = -10.5229$
$ABD = -5.0842$	$D = -6.0275$
$ABC = -5.7696$	$C = -8.2045$
$CD = 4.6707$	$B = -6.5304$
$BD = -4.6620$	$A = -0.7914$
$BC = -0.7982$	

(a) Are any of the effects significant?

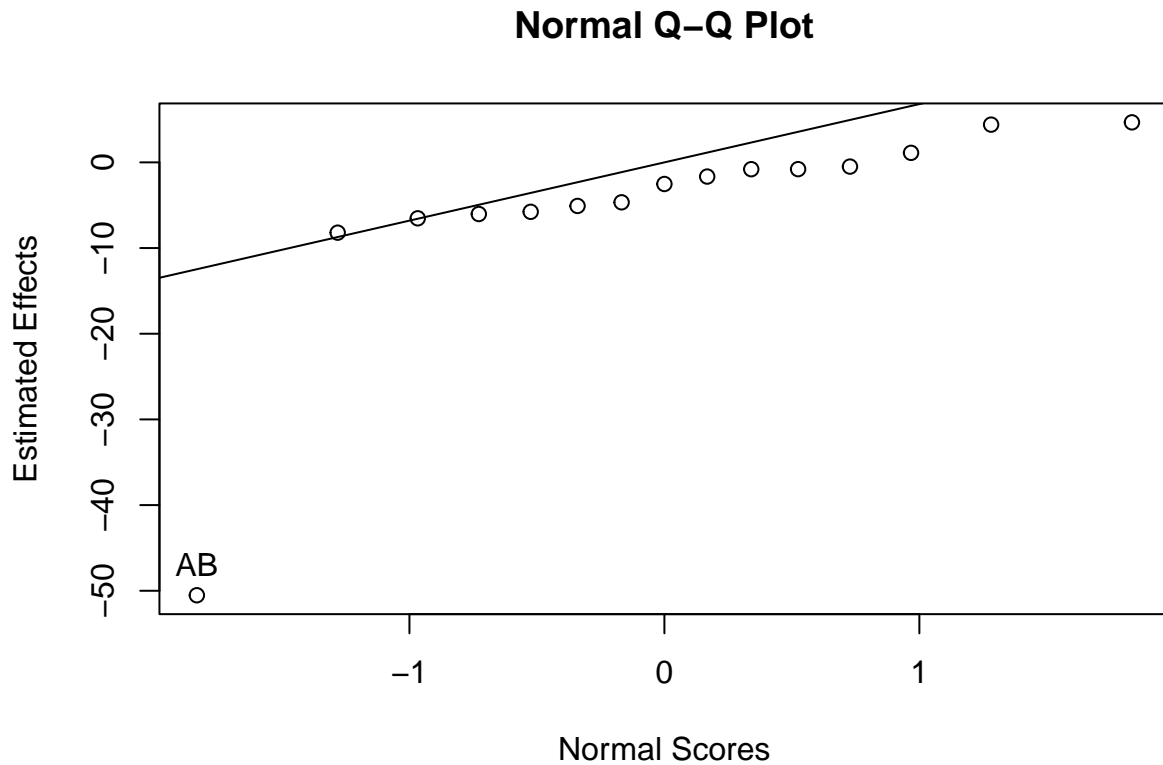
```
effects = c(-2.5251,-1.6564,4.4054,1.1109,-0.4932,-10.5229,-5.0842,-6.0275,-5.7696,-8.2045,
            4.6707,-6.5304,-4.6620,-0.7914,-0.7982)
names(effects) = c("ABCD","AD","BCD","AC","ACD","AB","ABD","D","ABC","C","CD","B","BD","A","BC")
fullnormal(effects,names(effects),alpha=0.025)
```



∴ No, all the effects doesn't seem to be significant.

(b) What happens if the effect of the interaction AB was -50.5229 instead of -10.5229?

```
effects[6] = -50.5229  
fullnormal(effects,names(labels),alpha=.025)
```



∴ The effect of the interaction AB is significant when the effect of AB is -50.5229.

Question 5. An article in *Quality and Reliability Engineering International* (2010, Vol. 26, pp. 223-233) presents a 2^5 factorial design. The experiment is shown in the following table:

A	B	C	D	E	y
-1	-1	-1	-1	-1	8.11
1	-1	-1	-1	-1	5.56
-1	1	-1	-1	-1	5.77
1	1	-1	-1	-1	5.82
-1	-1	1	-1	-1	9.17
1	-1	1	-1	-1	7.8
-1	1	1	-1	-1	3.23
1	1	1	-1	-1	5.69
-1	-1	-1	1	-1	8.82
1	-1	-1	1	-1	14.23
-1	1	-1	1	-1	9.2
1	1	-1	1	-1	8.94
-1	-1	1	1	-1	8.68
1	-1	1	1	-1	11.49
-1	1	1	1	-1	6.25
1	1	1	1	-1	9.12
-1	-1	-1	-1	1	7.93
1	-1	-1	-1	1	5
-1	1	-1	-1	1	7.47
1	1	-1	-1	1	12
-1	-1	1	-1	1	9.86
1	-1	1	-1	1	3.65
-1	1	1	-1	1	6.4
1	1	1	-1	1	11.61
-1	-1	-1	1	1	12.43
1	-1	-1	1	1	17.55
-1	1	-1	1	1	8.87
1	1	-1	1	1	25.38
-1	-1	1	1	1	13.06
1	-1	1	1	1	18.85
-1	1	1	1	1	11.78
1	1	1	1	1	26.05

```
A <- rep(c(-1,1), 16)
B <- rep(c(-1,-1,1,1), 8)
```

```

C <- rep(c(rep(-1,4),rep(1,4)),4)
D <- rep(c(rep(-1,8),rep(1,8)),2)
E <- c(rep(-1,16),rep(1,16))
y <- c(8.11,5.56,5.77,5.82,9.17,7.8,3.23,5.69,8.82,14.23,9.2,8.94,8.68,11.49,6.25,9.12,7.93,5,7.47,12,9)
q5_data <- data.frame(A,B,C,D,E,y)

```

(a) Analyze the data from this experiment. Identify the significant factors and interactions.

```

#Using ANOVA
res.lm<-lm(y~A*B*C*D*E, data=q5_data)
summary(res.lm)

```

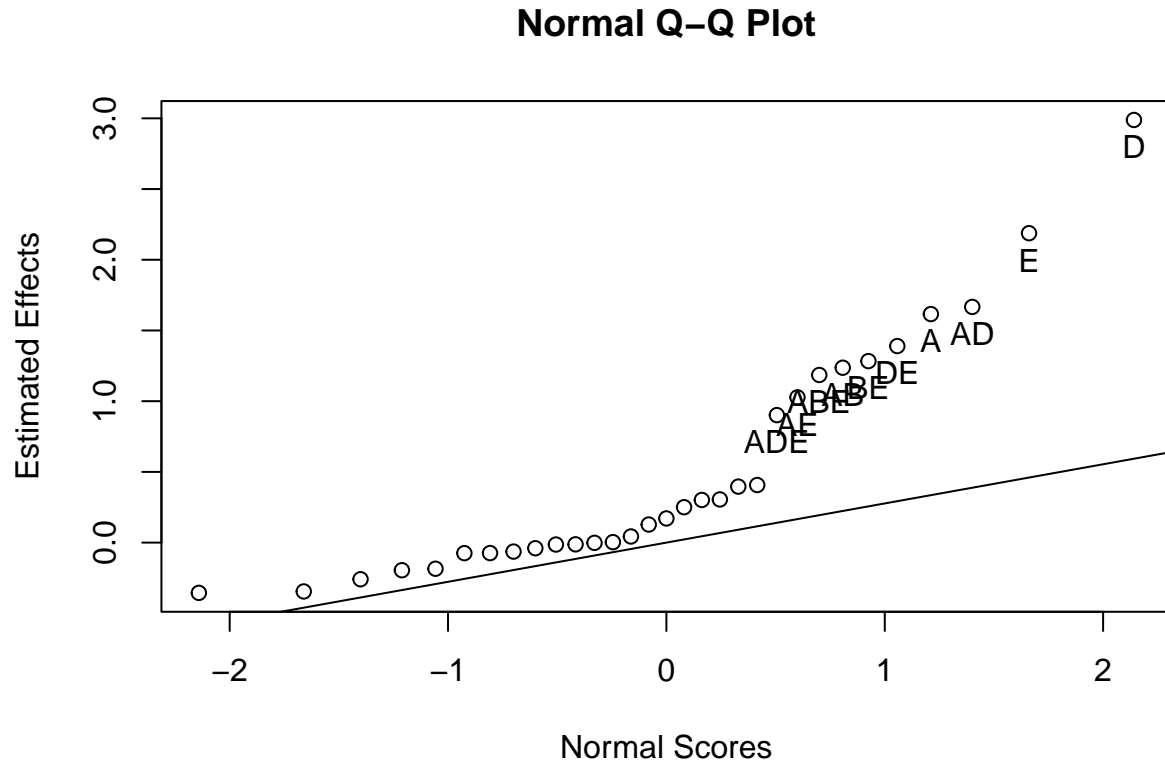
```

##
## Call:
## lm(formula = y ~ A * B * C * D * E, data = q5_data)
##
## Residuals:
## ALL 32 residuals are 0: no residual degrees of freedom!
##
## Coefficients:
##              Estimate Std. Error t value Pr(>|t|)
## (Intercept) 10.180312      NaN      NaN      NaN
## A              1.615938      NaN      NaN      NaN
## B              0.043438      NaN      NaN      NaN
## C             -0.012187      NaN      NaN      NaN
## D              2.988438      NaN      NaN      NaN
## E              2.187813      NaN      NaN      NaN
## A:B            1.236562      NaN      NaN      NaN
## A:C           -0.001563      NaN      NaN      NaN
## B:C           -0.195313      NaN      NaN      NaN
## A:D            1.666562      NaN      NaN      NaN
## B:D           -0.013438      NaN      NaN      NaN
## C:D            0.003437      NaN      NaN      NaN
## A:E            1.027187      NaN      NaN      NaN
## B:E            1.283437      NaN      NaN      NaN
## C:E            0.301562      NaN      NaN      NaN
## D:E            1.389687      NaN      NaN      NaN
## A:B:C          0.250313      NaN      NaN      NaN
## A:B:D         -0.345312      NaN      NaN      NaN
## A:C:D         -0.063437      NaN      NaN      NaN
## B:C:D          0.305313      NaN      NaN      NaN
## A:B:E          1.185313      NaN      NaN      NaN
## A:C:E         -0.259062      NaN      NaN      NaN
## B:C:E          0.170938      NaN      NaN      NaN
## A:D:E          0.901563      NaN      NaN      NaN
## B:D:E         -0.039687      NaN      NaN      NaN
## C:D:E          0.395938      NaN      NaN      NaN
## A:B:C:D       -0.074063      NaN      NaN      NaN
## A:B:C:E       -0.184688      NaN      NaN      NaN
## A:B:D:E        0.407187      NaN      NaN      NaN
## A:C:D:E        0.127812      NaN      NaN      NaN

```

```
## B:C:D:E      -0.074688      NaN      NaN      NaN
## A:B:C:D:E    -0.355312      NaN      NaN      NaN
##
## Residual standard error: NaN on 0 degrees of freedom
## Multiple R-squared:      1, Adjusted R-squared:      NaN
## F-statistic:   NaN on 31 and 0 DF,  p-value: NA
```

```
fullnormal(coef(res.lm)[-1],alpha=.05)
```



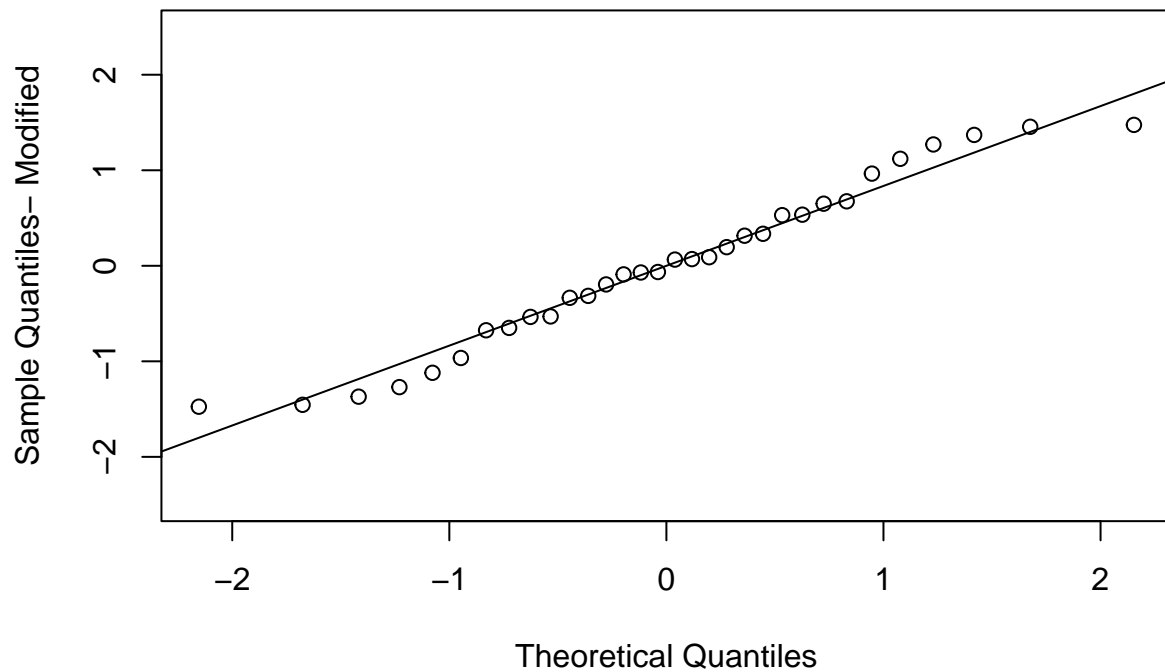
Factors and Interactions A, D, E, AB, AD, AE, BE, DE, ABE and ADE are significant.

(b) Analyze the residuals from this experiment. Are there any indications of model inadequacy or violations of the assumptions?

```
#residual analysis
#Normality: drop C (the reduced model)
res.aov = aov(y~A*B*D*E, data=q5_data)
residuals=res.aov$residuals
qqnorm(residuals, ylim=c(min(residuals)-1,max(residuals)+1), main = "Normal Q-Q Plot for Residuals",
       xlab = "Theoretical Quantiles", ylab = "Sample Quantiles- Modified",
       plot.it = TRUE, datax = FALSE)

qqline(residuals, datax = FALSE, distribution = qnorm)
```

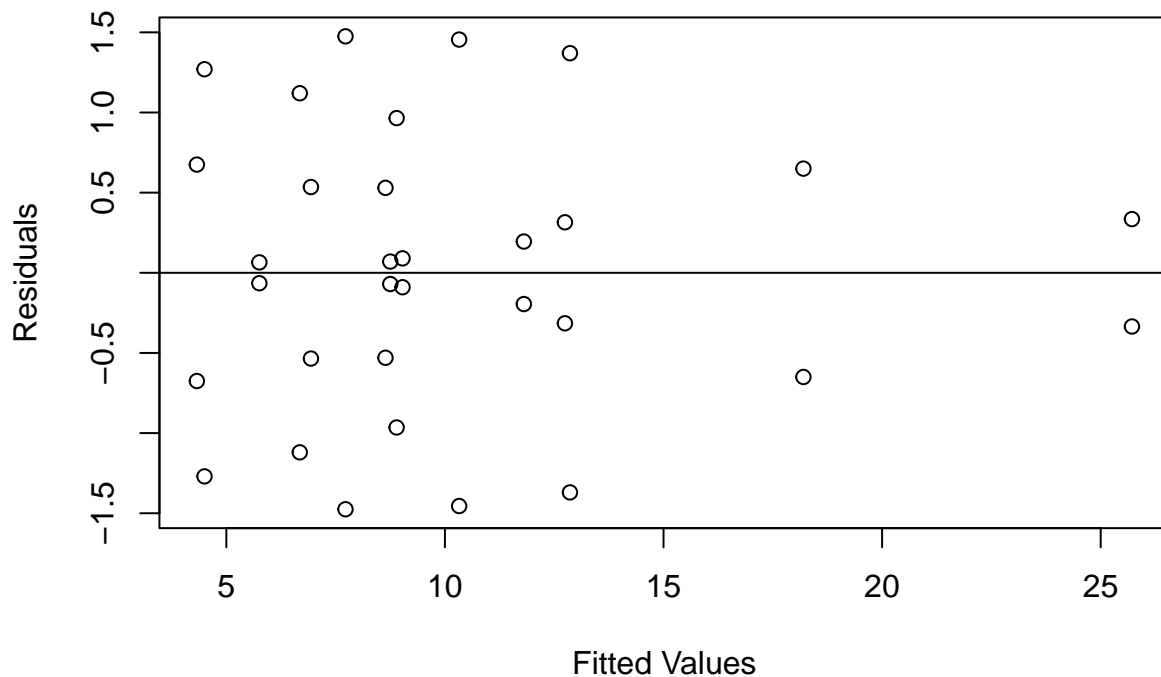
Normal Q-Q Plot for Residuals



```
#Test normality using Shapiro Wilks  
shapiro.test(residuals)
```

```
##  
##  Shapiro-Wilk normality test  
##  
## data:  residuals  
## W = 0.96505, p-value = 0.3751
```

```
#Check Variance  
Fitted_values=res.aov$fitted.values  
plot(Fitted_values,residuals,ylab="Residuals",xlab="Fitted Values")  
abline(h=0)
```



\therefore Since the p-value from Shapiro Wilks test is $0.3751 > 0.05$, we fail to reject the null hypothesis that the residuals are normally distributed. So there's no indication of violation of the normality assumption. There might be a violation of the constant variance assumption since higher fitted values seem to have smaller variance.

(c) One of the factors from this experiment does not seem to be important. If you drop this factor, what type of design remains? Analyze the data using the full factorial model for only the four active factors. Compare your results with those obtained in part (a).

```
# drop factor C
res.aov<-aov(y~A*B*D*E, data=q5_data)
summary(res.aov)
```

##	Df	Sum Sq	Mean Sq	F value	Pr(>F)
## A	1	83.56	83.56	57.233	1.14e-06 ***
## B	1	0.06	0.06	0.041	0.841418
## D	1	285.78	285.78	195.742	2.16e-10 ***
## E	1	153.17	153.17	104.910	1.97e-08 ***
## A:B	1	48.93	48.93	33.514	2.77e-05 ***
## A:D	1	88.88	88.88	60.875	7.66e-07 ***
## B:D	1	0.01	0.01	0.004	0.950618
## A:E	1	33.76	33.76	23.126	0.000193 ***
## B:E	1	52.71	52.71	36.103	1.82e-05 ***
## D:E	1	61.80	61.80	42.328	7.24e-06 ***
## A:B:D	1	3.82	3.82	2.613	0.125501


```
## A:B:E      1  44.96   44.96  30.794 4.40e-05 ***
## A:D:E      1  26.01   26.01  17.815 0.000650 ***
## B:D:E      1   0.05    0.05   0.035 0.854935
## A:B:D:E    1   5.31    5.31   3.634 0.074735 .
## Residuals 16  23.36    1.46
## ---
## Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

∴ Factor C is not important. The remaining design is a 2^4 factorial design with 2 replicates. And the same factors and interactions (A, D, E, AB, AD, AE, BE, DE, ABE and ADE) are significant as part (a).

(d) Find the settings of the active factors that maximize the predicted response.

```
# Add fitted as a column to the data frame
q5_data$Fitted = res.aov$fitted.values
q5_data[res.aov$fitted.values==max(res.aov$fitted.values),] #also can use "which"
```

```
##      A B  C D E      y Fitted
## 28 1 1 -1 1 1 25.38 25.715
## 32 1 1  1 1 1 26.05 25.715
```

The maximum fitted value is 25.715 and the settings of the active factors that maximize the predicted response are (A=1, B=1, C=-1, D=1, E=1) and (A=1, B=1, C=1, D=1, E=1).