## Stat453\_Assignment01

Belina Jang

2023-01-22

## Questions

1) Suppose that we are testing H0:  $\mu 1 = \mu 2$  versus H1:  $\mu 1 > \mu 2$  with a sample size of n1 = n2 = 10. Both sample variances are unknown but assumed equal. Using R, find p-values for the following observed values of the test statistics:

```
\# (a) t0 = 2.45
t0 = 2.45
df = 18 \# n1+n2-2
pt(t0, df, lower.tail = FALSE)
## [1] 0.01237173
# (b) t0 = -3.60
t0 = -3.60
pt(t0, df, lower.tail = FALSE)
## [1] 0.9989764
\# (c) t0 = 1.96
t0 = 1.96
pt(t0, df, lower.tail = FALSE)
## [1] 0.03283198
# (d) t0 = -2.19
t0 = -2.19
pt(t0, df, lower.tail = FALSE)
## [1] 0.9790323
```

(1-e) Repeat (a)-(d) for the case when the alternative hypothesis is two-sided.

```
\# (a) t0 = 2.45
t0 = 2.45
pt(abs(t0), df, lower.tail = FALSE)*2
## [1] 0.02474346
# (b) t0 = -3.60
t0 = -3.60
pt(abs(t0), df, lower.tail = FALSE)*2
## [1] 0.002047199
\# (c) t0 = 1.96
t0 = 1.96
pt(abs(t0), df, lower.tail = FALSE)*2
## [1] 0.06566396
# (d) t0 = -2.19
t0 = -2.19
pt(abs(t0), df, lower.tail = FALSE)*2
## [1] 0.04193532
```

2) The time to repair an electronic instrument is a normally distributed random variable measured in hours. The repair time for 16 such instruments chosen at random are as follows:

```
159 280 101 212
224 379 179 264
222 362 168 250
149 260 485 170

hours = c(159, 280, 101, 212, 224, 379, 179, 264, 222, 362, 168, 250, 149, 260, 485, 170)
```

(2-a) Is the normality assumption appropriated? Justify.

Hours

```
# normality_check
shapiro.test(hours)

##

## Shapiro-Wilk normality test
##

## data: hours
## W = 0.91997, p-value = 0.1685
```

 $\therefore$  The p-value is 0.1685, which is greater than 0.05, so we fail to reject the null hypothesis that the data is normally distributed. So the normality assumption is appropriate.

(2-b) You wish to know if the mean repair time exceeds 225 hours. Set up appropriate hypotheses for investigating this issue.

H0:  $\mu$  <= 225: null hypothesis mean is less than or equal to 225 hours H1:  $\mu$  > 225: alternative hypothesis mean exceeds 225 hours

(2-c) Using R, test the hypotheses you formulated in part (b). What are your conclusions? Use a=0.01.

```
a = 0.01
sample_mean = mean(hours) # 241.5
sample_sd = sd(hours) # 98.72588
t0 = (sample_mean - 225)/(sample_sd/sqrt(16)) # 0.6685177
df = 16-1 # 15

print("p-value")

## [1] "p-value"
pt(t0, df, lower.tail = FALSE) # 0.2569801

## [1] 0.2569801
t.test(hours, mu=225, alternative = "greater")$p.value < a</pre>
```

 $\therefore$  The p-value is 0.2569801, which is greater than 0.01, so we fail to reject the null hypothesis that the mean repair time is less than or equal to 225 hours.

Therefore, we cannot conclude that the mean repair time exceeds 225 hours.

(2-d) By hand, construct a 95 percent confidence interval on mean repair time to test your hypothesis in (b). Show your work.

Calculation by hand:

## [1] FALSE

```
a = 0.05
t_a = qt(a, df, lower.tail = FALSE) # 1.75305
lower_limit = sample_mean - t_a * sample_sd/sqrt(16) #[1] 198.2321
CI_95 = c(lower_limit, "Inf")
CI_95
## [1] "198.232138865275" "Inf"
print("95% CI: (198.2321, Inf)")
## [1] "95% CI: (198.2321, Inf)"
```

3) An article in the journal of Neurology (1998, Vol. 50, pp.1246-1252) observed that the monozygotic twins share numerous physical, psychological and pathological traits. The investigators measured an intelligence score of 10 pairs of twins. The data are obtained as follows:

```
Twin pair
Birth Order: 1 Birth Order: 2
1 5.73 6.08
25.806.22
38.427.99
46.847.44
5 6.43 6.48
6 8.76 7.99
7\ 6.32\ 6.32
8 7.62 7.60
96.596.03
10 7.67 7.52
twin_pair = c(1:10)
birth_order_1 = c(5.73, 5.80, 8.42, 6.84, 6.43, 8.76, 6.32, 7.62, 6.59, 7.67)
birth_order_2 = c(6.08, 6.22, 7.99, 7.44, 6.48, 7.99, 6.32, 7.60, 6.03, 7.52)
d = birth_order_1-birth_order_2
df2 = data.frame(twin_pair, birth_order_1, birth_order_2,d)
```

(3-a) Is the assumption that the difference in score is normally distributed reasonable?

```
# normality_check
shapiro.test(birth_order_1 - birth_order_2)

##

## Shapiro-Wilk normality test
##

## data: birth_order_1 - birth_order_2
## W = 0.96727, p-value = 0.8645
```

 $\therefore$  The p-value is 0.8645, which is greater than 0.05, so we fail to reject the null hypothesis that the data is normally distributed.

Therefore, the assumption is reasonable.

(3-b) Using R, find a 95% confidence interval on the difference in the mean score. Is there any evidence that mean score depends on birth order?

```
# 95% confidence interval
var.test(birth_order_1, birth_order_2) # 95% CI contains 1, so we assume variances are equal
##
## F test to compare two variances
##
## data: birth_order_1 and birth_order_2
## F = 1.6907, num df = 9, denom df = 9, p-value = 0.4461
## alternative hypothesis: true ratio of variances is not equal to 1
## 95 percent confidence interval:
## 0.4199363 6.8065913
## sample estimates:
## ratio of variances
##
             1.690661
a = 0.05
# paired t-test
d_{bar} = sum(d)/10
s_d = sqrt(sum((d-d_bar)^2)/(10-1))
t0 = d_bar/s_d/sqrt(10)
t_a2 = qt(a/2,10-1,lower.tail = FALSE)
lower_limit = d_bar - t_a2 * s_d/sqrt(10)
upper_limit = d_bar + t_a2 * s_d/sqrt(10)
CI_95 = c(lower_limit, upper_limit)
CI 95
```

- **##** [1] -0.2644148 0.3664148
- $\therefore$  95% confidence interval on the difference in the mean score is (-0.2644148, 0.3664148). And there's no evidence that the mean score depends on birth order, since the 95% CI contains 0 (indicating the mean values are equal).
- (3-c) Test an appropriate set of hypotheses indicating that the mean score does not depend on birth order.

```
# HO: μ1 = μ2: supports that the mean score does not depend on birth order
# H1: μ1 != μ2: supports that the mean score does depend on birth order
# Assume the datas are paired, observe p-value
t.test(birth_order_1, birth_order_2, alternative="two.sided",paired = TRUE, var.equal=TRUE)
##
## Paired t-test
##
## data: birth_order_1 and birth_order_2
```

 $\therefore$  The p-value is 0.723, which is greater than 0.05, so we fail to reject the null hypothesis that the mean scores are equal.

Therefore, since the mean scores are likely equal, there is no evidence that the mean score depends on birth order.

4) The deflection temperature under load for two different formulations of ABS plastic pipe is being studied. Two samples of 12 observations each are prepared using each formulation, and the deflection temperatures (in °F) are reported below:

```
Formulation 1 Formulation 2 206 193 192 177 176 198 188 207 210 197 185 188 205 185 194 206 200 189 187 189 178 201 197 203
```

```
formulation_1 = c(206, 193, 192, 188, 207, 210, 205, 185, 194, 187, 189, 178)
formulation_2 = c(177, 176, 198, 197, 185, 188, 206, 200, 189, 201, 197, 203)
df4 = data.frame(formulation_1,formulation_2)
```

Do the data support the claim that the mean deflection temperature under load for formulation 2 exceeds that of formulation 1?

(4-a) Use a=0.05 to perform a complete analysis in R, including normality check and the appropriate test. Use the rejection region method to test your hypothesis.

```
H0: \mu1 >= \mu2 H1: \mu1 < \mu2 # alternative hypothesis: formulation 2 has a higher mean deflection temperature
```

```
# normality check
shapiro.test(df4$formulation_1-df4$formulation_2)

##
## Shapiro-Wilk normality test
##
## data: df4$formulation_1 - df4$formulation_2
## W = 0.93569, p-value = 0.4443

# Variance
var.test(formulation_1, formulation_2)
```

```
##
## F test to compare two variances
##
## data: formulation_1 and formulation_2
## F = 1.046, num df = 11, denom df = 11, p-value = 0.9419
## alternative hypothesis: true ratio of variances is not equal to 1
## 95 percent confidence interval:
## 0.3011181 3.6334674
## sample estimates:
## ratio of variances
##
             1.045994
... The p-values of the normality test is 0.4443, which is greater than 0.05, so we fail to reject the null
hypothesis that the data is normally distributed.
95% CI of F-test contains 1, so we assume the variances are equal.
# Assuming the variances are equal
a = 0.05
s1 = sd(df4$formulation_1) # 10.17573
s2 = sd(df4\$formulation_2) # 9.949494
sp = sqrt(((12-1)*s1^2+(12-1)*s2^2)/(12+12-2)) # 10.06325
t0 = (mean(formulation_1) - mean(formulation_2))/(sp*(sqrt(1/12+1/12))) # 0.3448301
# using rejection region
qt(a, 12+12-2, lower.tail = FALSE) # ta = 1.717144
## [1] 1.717144
t0 < -qt(a, 12+12-2, lower.tail = FALSE) # False, we can't reject HO
## [1] FALSE
# using p-value, t-test
pt(t0, 22, lower.tail = TRUE) # p-value 0.6332515 -> can't reject HO
## [1] 0.6332515
t.test(formulation_1, formulation_2, var.equal = TRUE, alternative = "less") # double checking p-value
##
##
   Two Sample t-test
##
## data: formulation_1 and formulation_2
## t = 0.34483, df = 22, p-value = 0.6333
```

## alternative hypothesis: true difference in means is less than 0

## 95 percent confidence interval:

-Inf 8.471217

## sample estimates:
## mean of x mean of y
## 194.5000 193.0833

##

 $\therefore$  Since p-value of t-test is 0.6333, which is greater than 0.05, we fail to reject the null hypothesis H0. Also we can see that t0 is not in the rejection region(t0<-ta), so we cannot reject H0.

(4-b) Does the confidence interval support your conclusion on part a? Justify.

```
upper_limit = mean(formulation_1)-mean(formulation_2) + qt(a, 12+12-2, lower.tail = FALSE) * sp*(sqrt(1
CI = c("-Inf", upper_limit)
print("95% Confidence Interval")
## [1] "95% Confidence Interval"
CI # (-Inf, 8.471217)
## [1] "-Inf"
                           "8.47121668372174"
t.test(formulation_1, formulation_2, alternative = 'less', var.equal = TRUE) # double check CI
##
##
    Two Sample t-test
##
## data: formulation_1 and formulation_2
## t = 0.34483, df = 22, p-value = 0.6333
## alternative hypothesis: true difference in means is less than 0
## 95 percent confidence interval:
        -Inf 8.471217
## sample estimates:
## mean of x mean of y
   194.5000 193.0833
∴ 95% confidence interval is (-Inf, 8.471217)
The confidence interval does contain 0, so it supports H0 that \mu 1 >= \mu 2.
So, it supports my conclusion on part (a) not to reject H0.
```

5) Photoresist is a light-sensitive material applied to semiconductor wafers so that the circuit pattern can be imaged on to the wafer. After application, the coated wafers are baked to remove the solvent in the photoresist mixture and to harden the resist. Here are measurements of photoresist thickness (in kÅ) for eight wafers baked at two different temperatures. Assume that all of the 16 runs were made in random order. Note: a wafer cannot be baked twice.

```
95 °C 100 °C
11.176 5.623
7.089 6.748
8.097 7.461
11.739 7.015
11.291 8.133
10.759 7.418
6.467 3.772
8.315 8.963
```

```
temp95 = c(11.176, 7.089, 8.097, 11.739, 11.291, 10.759, 6.467, 8.315)
temp100 = c(5.623, 6.748, 7.461, 7.015, 8.133, 7.418, 3.772, 8.963)
df5 = data.frame(temp95, temp100)
```

(5-a) Is there evidence to support the claim that the higher baking temperature results in wafers with a lower mean photoresist thickness? Use a = 0.05 and justify your answer.

```
# H0: \mu 1 <= \mu 2
# H1: \mu 1 > \mu 2
a = 0.05
var.test(temp95, temp100) # 95% CI contains 1 so we assume the variances are equal.
##
##
   F test to compare two variances
## data: temp95 and temp100
## F = 1.7326, num df = 7, denom df = 7, p-value = 0.4855
## alternative hypothesis: true ratio of variances is not equal to 1
## 95 percent confidence interval:
## 0.3468649 8.6539742
## sample estimates:
## ratio of variances
             1.732559
# Assuming variances are equal
s1 = sd(temp95)
s2 = sd(temp100)
sp = sqrt(((8-1)*s1^2+(8-1)*s2^2)/(8+8-2))
t0 = (mean(temp95)-mean(temp100))/(sp*(sqrt(1/8+1/8)))
print("p-value")
## [1] "p-value"
pt(t0, 8+8-2, lower.tail = FALSE) # p-value 0.009423718
## [1] 0.009423718
t.test(temp95, temp100, alternative = "greater", var.equal= TRUE) # double checking p-value
##
##
   Two Sample t-test
##
## data: temp95 and temp100
## t = 2.6549, df = 14, p-value = 0.009424
\#\# alternative hypothesis: true difference in means is greater than 0
## 95 percent confidence interval:
## 0.8330468
                    Tnf
```

```
## sample estimates:
## mean of x mean of y
## 9.366625 6.891625
```

 $\therefore$  The p-value is 0.009423718, which is smaller than 0.05, therefore, there's a very strong evidence against H0, so we reject H0.

Therefore, there is an evidence to support the claim that the higher baking temperature results in wafers with a lower mean photoresist thickness.

(5-b) Find a 95% confidence interval on the difference in means. Provide a practical interpretation of this interval.

```
a = 0.05
s1 = sd(temp95)
s2 = sd(temp100)
sp = sqrt(((8-1)*s1^2+(8-1)*s2^2)/(8+8-2))
t0 = (mean(temp95)-mean(temp100))/(sp*(sqrt(1/8+1/8)))
lower_limit = mean(temp95)-mean(temp100) - qt(a, 8+8-2, lower_tail = FALSE) * sp * sqrt(1/8+1/8)

CI = c(lower_limit, "Inf")
print("95% CI")

## [1] "95% CI"

CI # (0.8330468, Inf) lower limit only since it's a one sided test
```

```
## [1] "0.833046759185779" "Inf"
```

 $\therefore$  The confidence interval, (0.8330468, Inf), shows where the absolute value of the mean difference  $|\mu 1-\mu 2|$  lies. Since the 95% confidence inteval does not contain 0, it indicates that there is statistically significant difference between two mean values.

6) The following are the burning times (in minutes) of chemical flares of two different formulations. The design engineers are interested in both the means and variance of the burning times.

```
Type 1 Type 2 65 82 64 56 81 67 71 69 57 59 83 74 66 75 59 82 82 70 65 79
```

```
type_1 = c(65, 82, 81, 67, 57, 59, 66, 75, 82, 70)
type_2 = c(64, 56, 71, 69, 83, 74, 59, 82, 65, 79)
df6 = data.frame(type_1, type_2)
```

(6-a) Test the hypotheses that the two variances are equal. Use a = 0.05.

```
# HO: sigma1^2 = sigma2^2
# H1: sigma1^2 != sigma2^2
a = 0.05
f0 = sd(type_1)^2/sd(type_2)^2
f0 # 0.9782168
## [1] 0.9782168
f_a2 = qf(a/2, 10-1, 10-1, lower.tail = FALSE)
f_a2 #4.025994
## [1] 4.025994
print("abs(f0) > f_a2")
## [1] "abs(f0) > f a2"
abs(f0) > f_a2 # FALSE so we can't reject the null hypothesis that the two variances are equal. slide 1
## [1] FALSE
var.test(type_1, type_2) # 0.9744
##
##
   F test to compare two variances
##
## data: type_1 and type_2
## F = 0.97822, num df = 9, denom df = 9, p-value = 0.9744
## alternative hypothesis: true ratio of variances is not equal to 1
## 95 percent confidence interval:
## 0.2429752 3.9382952
## sample estimates:
## ratio of variances
           0.9782168
##
```

Since,  $abs(f0) > f_a2$  is false and p-value is 0.744, we can't reject the null hypothesis that the two variances are equal. Therefore, variances are equal.

(6-b) Using the results of (a), test the hypotheses that the mean burning times are equal. Use a=0.01. What is the p-value for this test?

```
# variances are equal and unknown

# H0: \mu 1 = \mu 2

# H1: \mu 1 != \mu 2

sp = sqrt(((10-1)*sd(type_1)^2+(10-1)*sd(type_2)^2)/(10+10-2))

t0 = (mean(type_1)-mean(type_2))/(sp*(sqrt(1/10+1/10)))

pt(abs(t0), 10+10-2, lower.tail = FALSE)*2
```

```
t.test(type_1, type_2, var.equal = TRUE) # double check p-value
```

```
##
## Two Sample t-test
##
## data: type_1 and type_2
## t = 0.048008, df = 18, p-value = 0.9622
## alternative hypothesis: true difference in means is not equal to 0
## 95 percent confidence interval:
## -8.552441 8.952441
## sample estimates:
## mean of x mean of y
## 70.4 70.2
```

 $\therefore$  The p-value 0.9622 is greater than 0.01, so we fail to reject the null hypothesis that the mean burning times are equal.

Therefore, we can assume that mean burning times are equal.

(6-c) Discuss the role of the normality assumption in this problem. Check the assumption of normality for both types of flares.

```
# normality check: The test rejects the hypothesis of normality when the p-value \leq 0.05. shapiro.test(type_1) #p-value \geq 0.05, so we do not reject the hyphothesis, the data is likely normal.
```

```
##
## Shapiro-Wilk normality test
##
## data: type_1
## W = 0.91359, p-value = 0.3065
shapiro.test(type_2) #p-value > 0.05, so we do not reject the hyphothesis, the data is likely normal.
```

```
##
## Shapiro-Wilk normality test
##
## data: type_2
## W = 0.95478, p-value = 0.7251
```

... In the part (a,b), the basic assumption required is that the samples are drawn from two independent normal populations, so it's important to check for the normality. And both Type 1 and Type 2 datas are normal.