

# HW 3

Ben Lirio

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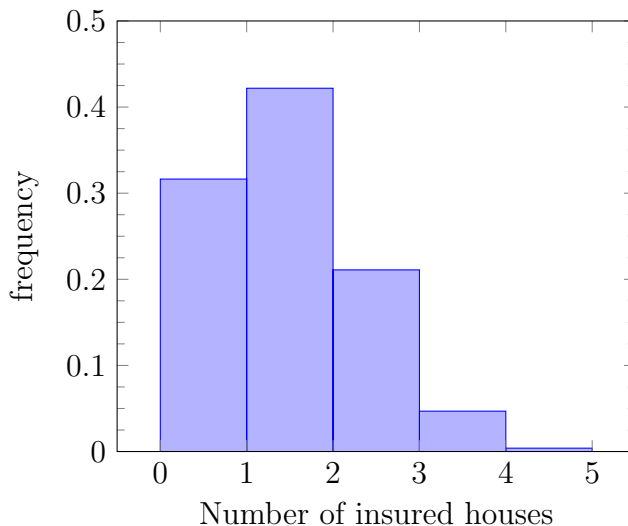
## 1. Flight

- (a) In order for a flight with  $s = 50$  seats to accomodate  $k$  passengers who show up,  $k \leq s$ . Let  $X$  represent the random variable of the number of passengers who show up, then  $P(X \leq s)$  is probability all  $k$  passengers will have a seat. This can be calculated as follows,  $P(X \leq 50) = \rho(45) + \rho(46) + \rho(47) + \rho(48) + \rho(49) + \rho(50) = .05 + .1 + .12 + .14 + .25 + .17 = .83$ . There is an **83%** chance all ticketed passengers who show up will have a seat.
- (b) The probability at least one of  $k$  passengers who show up will not be seated happens when  $k > s$ . Which can be shown as  $P(X < 50)$  which is the complement to answer 'a'. so  $P(X < 50) = 1 - P(X \leq 50) = 1 - .83 = .17$ . The probability that not all of the  $k$  will receive a seat is **17%**.
- (c) If you are the first person on standby, then if there must be one free seat on the plane for you to have a seat. This happens when  $X \leq 49$ . So answer derived from a  $P(X < 50) = P(X \leq 50) - P(X = 50) = .83 - .17 = .66$ . If you are the first person on standby there is a **66%** you will still be able to fly. Now, if you are the 3rd person on standby  $X \leq 47$  must be true.  $P(X \leq 47) = P(45) + P(46) + P(47) = .05 + .10 + .12 = .27$ . If you are the third person on standby there is a **27%** chance you will still be able to fly.

## 2. Earthquake

- (a) The probability distrabution of  $X$  is

$$\rho(k) = \begin{cases} \binom{4}{k} .25^k .75^{N-k} & 0 \leq k \leq 4 \\ 0 & otherwise \end{cases}$$



(b)

- (c) Initially I thought of calculating the expected value,  $E(X) = pn = 1.0$ , but then I interpreted the question as the mode, which value individually is most likely to be chosen, which also is **1**.
- (d) The probability at least two of the four selected houses have earthquake insurance can be represented as follows,  $P(X \geq 2) = 1 - P(X < 2) = 1 - P(X = 0) - P(X = 1) = 1 - .316406 - .421875 = 0.261719$ . The probability that at least two of the selected houses have earthquake insurance is **26.1719%**.

### 3. Allergies

- (a)  $P(X \leq 3) = \sum_{k=0}^3 \binom{25}{k} .05^k .95^{25-k} = 0.965909$ . There is a **96.5909%** chance 3 or less children in the sample will have allergies.  $P(X < 3) = P(X \leq 3) - P(X = 3) = 0.965909 - \binom{25}{3} .05^3 .95^{25-3} = 0.965909 - 0.872894 = 0.093015$ . There is a **9.3015%** chance less than 3 children will have allergies.
- (b)  $P(X \geq 4) = 1 - P(X \leq 3) = 1 - 0.965909 = 0.034091$ . There is a **3.4091%** chance 4 or more children in the sample will have allergies.
- (c)  $P(1 \leq X \leq 3) = P(X \leq 3) - P(X = 0) = .965909 - 0.27739 = 0.688519$ . There is a **68.8519%** chance between 1 and 3 children in the sample will have allergies.
- (d) The given distribution is binomial, therefore the expected value can be calculated as follows,  $E(X) = pn = (.05)(25) = 1.25$ . The expected value of the number of children in the sample size that will have allergies is **1.25**.
- (e) Let  $N = 50$  be the sample size maintaining  $p = 0.05$ . Let  $x = 0$  be the number of children of the sample size that has allergies. Calculate the probability  $P(X = x) = P(X = 0) = \binom{50}{0} .05^0 .95^{50} = 0.076945$ . There is a **7.6945%** chance none of the 50 sampled children will have allergies.

### 4. Fine Crystal

- (a) Assuming the sample population of fine crystal goblets is significantly greater than 6 I will consider this a binomial distribution. Let  $X$  be the r.v. associated with the number of "seconds". If  $N$  is the number of goblets sampled then the distribution of  $X \sim \text{Bin}(N, p)$ . In this case  $N = 6, p = .10$ . Therefore the probability of selecting only one "second" is  $P(X = 1) = \binom{6}{1} (.1)^1 (.9)^5 = \mathbf{0.354294}$ .
- (b) Similar to the previous problem but now there are 2 "seconds".  $P(X = 2) = \binom{6}{2} (.1)^2 (.9)^4 = \mathbf{0.098415}$ .
- (c) Selecting goblets one by one and counting the number of failures changes the distribution from binomial to negative binomial. Classifying a failure as "not seconds" our ending condition is when we have 4 failures. Now  $X$  will be the number of successes (counter intuitive since these are bad goblets) and I am calculating what the probability  $X + 4 \leq 5$ . The probability then is  $P(X \leq 1) = P(X = 0) + P(X = 1) = .9^4 + \binom{4}{1} (.9)^4 (.1)^1 = 0.59049 + 0.26244 = 0.85293$ . There is a **85.293%** chance that at most 5 goblets must be selected to find four that are not seconds.
5.  $Y = 3$  for the outcome  $\{SSS\}$ .  $Y = 4$  for the outcomes  $\{FSSS\}$ .  $Y = 5$  for the outcomes  $\{SFSSS\}, \{FFSSS\}, \{SFFSS\}$ .
6. By looking at a single round at a time, the following events are mutually exclusive:  $X = 0$ , a tie occurs,  $X = 1$  the first player wins,  $X = 2$  the second player wins. Because each round is independent, and the only possible outcomes are prefixed with ties, ( $\{001\}, \{00002\}, \{002\}$ ),

it is valid to condition on the fact that the last event was not a tie. Now, let  $c_i$  represent the number of ways the  $i^{th}$  player can win. To calculate these, I need to know the r.v. that correspond to the number of heads each player gets. Let  $X, Y$  represent the number of head the first and second player get respectively (since heads and tails is equally likely and the number of coins tossed is fixed I am counting possible outcomes as opposed to probabilities).  $c_1 = X > Y = \binom{3}{3}(2^2) + \binom{3}{2}(2^2 - \binom{2}{2}) + \binom{3}{1}\binom{2}{0} = 16$ . And  $c_2 = Y > X = \binom{2}{2}(\binom{3}{1} + \binom{3}{0}) + \binom{2}{1}\binom{3}{0} = 6$ . I will divide each of those probabilities by the sum of the total and I get: There is a **72.7273%** chance the first player will win and a **27.2727%** chance the second player will win.