




## Variable Selection

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# Overview

- Linear regression
  - Variable selection
  - Penalties
  - Grouped penalties
  - Tuning parameters
  - Caveats
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## Multiple linear regression model

- Observe data  $(\underline{y_i}, \overset{\downarrow}{x_{i1}}, \dots, \overset{\uparrow}{x_{ip}})$  for subjects  $1, \dots, \underline{n}$ . Want to estimate  $\beta_0, \beta_1, \dots, \beta_p$  in the model

$$\underline{y_i} = \beta_0 + \beta_1 \overset{\uparrow}{x_{i1}} + \dots + \beta_p \overset{\uparrow}{x_{ip}} + \epsilon_i; \epsilon_i \overset{iid}{\sim} (0, \sigma^2)$$

- Assumptions: residuals have mean zero, constant variance, and are independent
- Estimate parameters using OLS

(This covers a lot of ground – general goodness of linear models, interpretation, inference, unbiasedness, diagnostics, ...)

# Multiple linear regression

- Let

$$\mathbf{y} = \begin{bmatrix} \underline{y_1} \\ \vdots \\ \underline{y_n} \end{bmatrix}, \quad \mathbf{X} = \begin{bmatrix} 1 & \underline{x_{11}} & \cdots & \underline{x_{1p}} \\ \vdots & \downarrow & & \vdots \\ 1 & x_{n1} & \cdots & x_{np} \end{bmatrix}, \quad \boldsymbol{\beta} = \begin{bmatrix} \beta_0 \\ \vdots \\ \beta_p \end{bmatrix}, \quad \boldsymbol{\epsilon} = \begin{bmatrix} \epsilon_1 \\ \vdots \\ \epsilon_n \end{bmatrix}$$

- Then we can write the model in a more compact form:

$$\underline{\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \boldsymbol{\epsilon}}$$

# Matrix notation

- Matrix notation provides a compact way to discuss regression models and related areas
- $X$  is called the *design matrix*
- In settings where there are many predictors, specifying the design matrix is typically easier than a “formula interface”

# OLS Estimation

- OLS estimate found by minimizing the RSS:

$$\begin{aligned}
 \hat{\beta}_{OLS} &= \arg \min_{\beta} [\text{RSS}(\beta)] \\
 &= \arg \min_{\beta} \left[ \sum_{i=1}^n \left( \underbrace{y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j}}_{\text{residual}} \right)^2 \right] \quad \checkmark \\
 &= \arg \min_{\beta} \left[ (\mathbf{y} - \mathbf{X}\beta)^T (\mathbf{y} - \mathbf{X}\beta) \right] \\
 &= \underline{(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}}
 \end{aligned}$$

# Variable selection

- Often we need to do *variable selection*
  - ▶ Sometimes  $p > n$
  - ▶ Sometimes we have many variables and want to identify the “important” ones
  - ▶ Sometimes we have many variables and want to remove “unimportant” ones

# Methods for variable selection

- Subset selection
  - ▶ More traditional methods
- Penalization / shrinkage
  - ▶ More recent methods



# Subset selection

- Forward / backward selection
  - Best subset selection
- 
- Won't speak for everyone, but I don't like these ...
    - ▶ Often not feasible
    - ▶ Can be unstable
    - ▶ Overall uncertainty is hard to assess, so inference is suspect

# Penalized regression

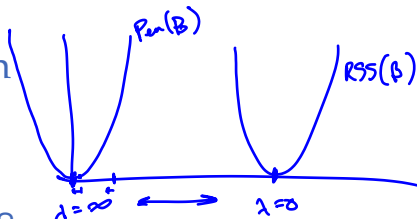
- Trade some bias for lower variance and overall MSE
  - ▶ Can outperform OLS in some important ways
- Rather than a subset selection approach, all parameters stay in the model but we restrict their effect
- Penalize the size of the coefficients – “unimportant” variables will have their coefficients shrunk towards to zero

# Ridge regression

Start with RSS and add a penalty:

$$\begin{aligned}
 \hat{\beta}_R &= \arg \min_{\beta} [RSS(\beta) + \lambda Pen(\beta)] \\
 &= \arg \min_{\beta} \left[ \underbrace{\sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2}_{\text{RSS}} + \lambda Pen(\beta) \right] \\
 &= \arg \min_{\beta} \left[ \underbrace{\sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2}_{\text{RSS}} + \lambda \underbrace{\sum_{j=1}^p \beta_j^2}_{\text{Penalty}} \right]
 \end{aligned}$$

# Notes on ridge regression



- Not unbiased
- Lower variance than OLS
- Avoid subset selection; add tuning parameter selection

- ▶ For “small” values of  $\lambda$ ,  $\hat{\beta}_R \approx \hat{\beta}_{OLS}$
- ▶ For “large” values of  $\lambda$ ,  $\hat{\beta}_R \approx 0$

- Results in coefficients that are small but non-zero
- Doesn't have to penalize everything

✓ • (Has a closed-form solution)

## Standardizing predictors

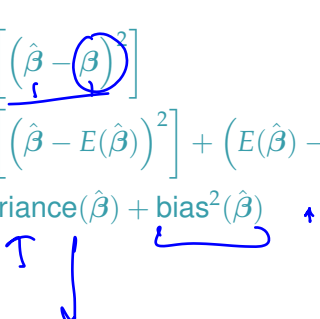
- OLS estimates are scale equivariant: multiplying a predictor by a constant rescales the coefficient but leaves the model (and the fitted values) unchanged
- In contrast, ridge regression coefficient estimates can change substantially when multiplying a given predictor by a constant
  - ▶ Rescaling a predictor changes the impact in of the coefficient in the penalty
- Therefore, it is best to apply ridge regression (and other penalized methods) after standardizing the predictors, using the formula

# Tuning parameters

Before considering other penalization methods, a brief digression ...

- Goal is to trade some increase in the bias of  $\hat{\beta}$  for a decrease in the variance of  $\hat{\beta}$
- Mean squared error formalizes this:

$$\begin{aligned}
 \underline{MSE(\hat{\beta})} &= E \left[ \underbrace{(\hat{\beta} - \beta)^2}_{\text{variance}} \right] \\
 &= E \left[ (\hat{\beta} - E(\hat{\beta}))^2 \right] + (E(\hat{\beta}) - \beta)^2 \\
 &= \text{variance}(\hat{\beta}) + \text{bias}^2(\hat{\beta}) \quad \uparrow
 \end{aligned}$$



# MSE for predictions

MSE for  $\beta$  isn't feasible in practice

- MSE for predictions is easier, and incorporates dependence on  $\hat{\beta}$  through the fitted values

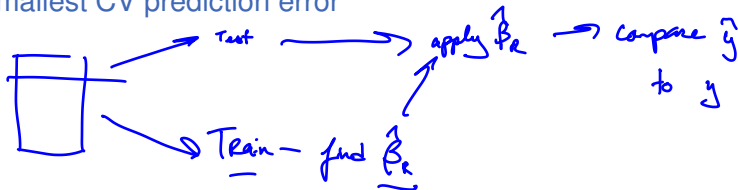
$$MSE(\hat{y}) = E \left[ (\hat{y} - y)^2 \right]$$

$\uparrow$                        $\uparrow$

- Can evaluate this using cross-validation

# Cross validation

- Set aside some subset of the data as a “test” data set
- Use remaining data to “train” the model (i.e. estimate parameters  $\hat{\beta}_R$ )
- Compute  $E(\hat{y}_{i,\lambda} - y_i)^2$  across all  $i$  in the “test” set
- Evaluate many values of  $\lambda$ , and choose the one with the smallest CV prediction error



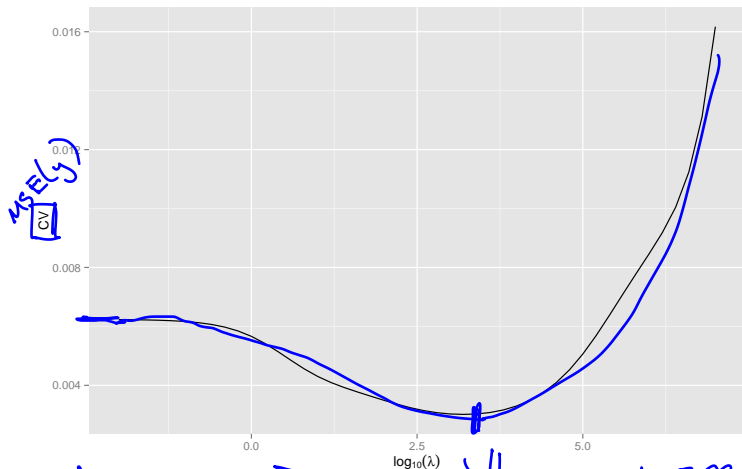


# Over and under fitting

Over- and under-fitting are common problems

- Over-fitting means you're fitting the current (training) data too well, and will make bad predictions for future (test) data
- Under-fitting means you're not fitting the current data well enough, and will make bad predictions for future (test) data
- Over-fitting is a problem of high variance; under-fitting is a problem of high bias

# Cross validation



$\lambda \approx 0$   
low bias  
high var



$\lambda!!$

$\lambda \rightarrow \infty$   
high bias  
low var

# Lasso penalization



- Lasso (least absolute shrinkage and selection operator) is a more recent penalized regression estimator
- Basic form is similar to that of ridge regression, but penalty function is different:

$$\begin{aligned}
 \hat{\beta}_L &= \arg \min_{\beta} [\underbrace{RSS(\beta)} + \underbrace{\lambda ||\beta||_1}] \\
 &= \arg \min_{\beta} \left[ \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 + \lambda \sum_{j=1}^p \underbrace{|\beta_j|} \right]
 \end{aligned}$$

- Quite popular – broadly used, many adaptations

# Lasso penalization

## Some properties of Lasso penalties

- No closed form solution (although there are some computationally useful tricks)
- The different penalty form means Lasso has a tendency to shrink coefficients *all the way* to zero
- Can be useful as an <sup>^</sup>“automated” variable selection approach
- Still have to choose  $\lambda$ ; cross validation is a popular tool for this

## Lasso vs ridge – “the picture”

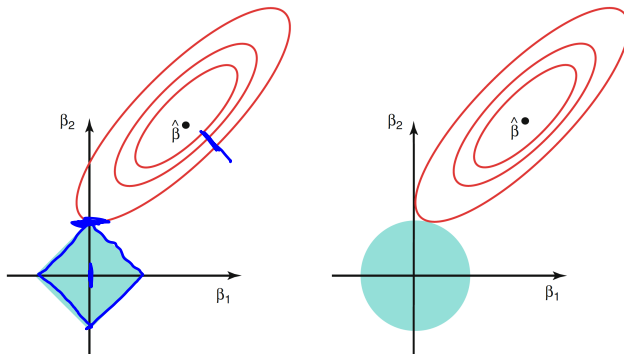


Figure: ISL 6.7, Page 222

The lasso performs  $\ell_1$  shrinkage, and there are “corners” in the constraint. If the RSS “hits” one of these corners, the coefficient corresponding to the axis is shrunk to zero.

# Lasso vs ridge – “the other picture”

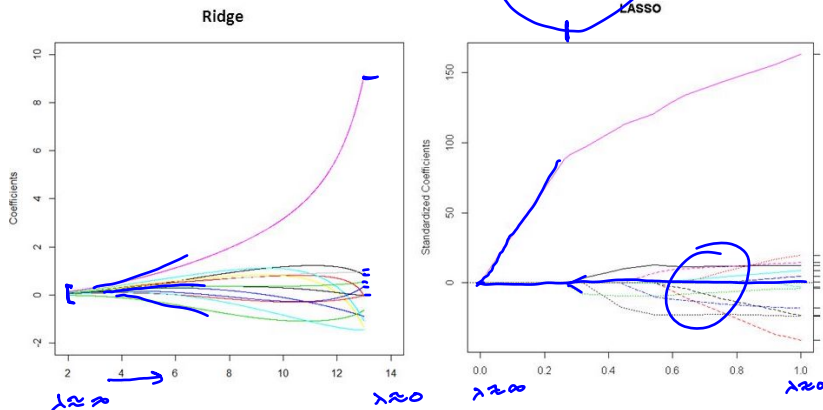


Figure: ISL 6.7, Page 222

## Comments on lasso

... especially with respect to MLR:

- Emphasis is on prediction, not inference
- Coefficients for selected variables is *not the same* as an MLR including only that subset
- When predictors are correlated, lasso tends to select one element of a group

# Bias-reducing penalties

Penalized regression framework is pretty broad:

$$\arg \min_{\beta} \left[ \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 + \lambda \underline{Pen}(\beta) \right]$$

In addition to lasso and ridge regression, other popular alternatives include

- Adaptive lasso
- MCP: minimax concave penalty
- SCAD: smoothly clipped absolute deviation

These all try to give unbiased coefficients for covariates with large effects, but operate under the same general framework



# Adaptive lasso

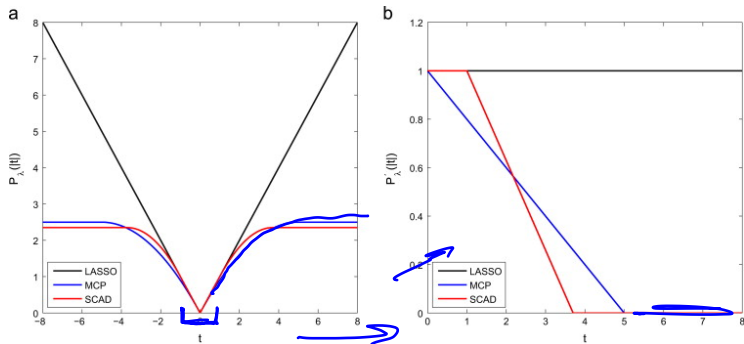
Adaptive lasso adds weights:

$$\arg \min_{\beta} \left[ \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 + \lambda \sum_{j=1}^p \overset{\downarrow}{w_j} |\beta_j| \right]$$

Weights (obtained via e.g. a first-round “standard” lasso) increase or decrease the effect of penalization on individual coefficients.

# MCP, SCAD

SCAD and MCP use a different penalty structure:



# Elastic net

- Unbiasedness is one problem with lasso; selection of a single covariate among from a correlated group is another
- The elastic net is one solution to this:

$$\arg \min_{\beta} \left[ \sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2 + \lambda \left( \underbrace{(1 - \alpha)}_{\text{ridge}} \sum_{j=1}^p \beta_j^2 + \underbrace{\alpha}_{\text{lasso}} \sum_{j=1}^p |\beta_j| \right) \right]$$

- Combines regularization via the ridge-type penalty with variable selection via the lasso penalty
- More effective to deal with groups of correlated predictors

# Group variable selection

Elastic net often “works”, but you can make grouping explicit

- Partition covariates into known groups
- Apply a relevant penalty to groups, not individual coefficients

$$\arg \min_{\beta} \left[ \underbrace{\sum_{i=1}^n \left( y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{i,j} \right)^2}_{\text{RSS}} + \lambda \sum_{g=1}^G \underbrace{Pen(\beta_g)}_{\text{penalty}} \right]$$

*Note: In the original image, the summation index 'G' is circled in blue, and there are blue arrows pointing from the 'G' to the summation and from the 'Pen' term to the underlined 'How you group matters' bullet point below.*

- Group lasso, SCAD, MCP, etc exist
- How you group matters

## Closing thoughts

- Inference (and, at least to some extent, interpretation) is challenging in variable selection methods
  - ▶ Emphasis is on prediction accuracy
  - ▶ Some work of post-selection inference exists, but hasn't yet been widely adopted
- Variable selection can be included in other models
  - ▶ Sparse PCA, for example
- Variable selection methods can have a similar goal as principal components regression, but use a very different approach

## Other directions

- ~~Variable selection can be included in other models~~
- Other approaches to variable selection (e.g. Bayesian methods ...)
- Consistency across approaches

# Sessions's big ideas

- Penalization; tuning parameters; cross validation; group penalties

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- ISLR 6