Final Project

Benji Gold and Sam Alkalay

Introduction

Baseball, America's pastime, has a long and storied tradition that dates back well over 100 years. Since the 1850's, some form of statistics measuring how good a player is has been tracked. This began through the use of the box score, which tracked basic statistics, such as hits, runs, and errors, from which a player's batting average can be constructed. Over one hundred years later, a pioneering statistician by the name of Bill James introduced new statistical concepts, such as on-base percentage and runs created, in his annual Baseball Abstract (Lee 2018). As technology has improved, the statistics being tracked became more and more sophisticated. Then, in 2015 analytics in baseball took a giant leap. With the introduction of Statcast, teams were able to track novel metrics, such as a batter's exit velocity (the speed of the baseball as it comes off the bat, immediately after a batter makes contact) and barrel percentage (the percentage of baseballs hit off of the player's barrel) ("Statcast Search"). Around the league, teams adopted these new statistics to try and gain a competitive advantage, through which they would be able to better predict a player's potential. However, is this actually the case? While these new statistics are widely used, it is unclear whether they actually provide any useful information for predicting a player's potential. This research project intends to explore that idea through the use of a logistic regression model to predict whether a player is an all-star. The research question of interest is:

Do old or new wave statistics do a better job at predicting whether a player is selected as an all-star?

The response variables of interest are: All.Star: Whether a player is selected as an all-star.

For our analysis, we have selected two datasets. The first is from Baseball Reference, which consists of standard statistics that offer a broad view of a player's performance in a particular season. The second is from Statcast, which consists of each player's primary position. The final data file we have was compiled from baseballsavant.com with a mix of more traditional stats and statcast stats. This complete file can be found in the stats.csv file. Once that was done, we entered a players position, salary, and team from baseball prospectus. we used wikipedia to find rosters for the 2019 all-star game and created a categorical variable column with that information. ## Methodology Our original research question involved predicitng whether or not a player was an all-star in 2019 based on their full season statistics for that season. Using our domain knowledge we realized that whether or not a palyer is designated an All-star is not simpily a measure of skill (which typical baseball statistics are designed to measure), but of popularity. This is true because of how all-stars are determined. The starters for the all-star team in each league are determined by fan vote so a bad player who is popular with fans might still be selected an all-star, regardless of their overall statistical skill. Because of this, our first decision in our methodology was to create two new categorical variables, one for if a player hit more than 40 Home Runs in 2019 and one for if a player had a batting average higher than 300. These variables are meant to measure fame based off a semi-arbitrary criteria based on how fans view baseball. Our domain knowledge tells us that a player is generally considered elite by fans if they hit more than 40 home runs or have an average higher than 300. Home runs and batting average are sort of hueristic statistics that fans use to measure player quality and thus these effects should be included in our model. The addition of these variables can be found in the packages and data section on code line ________________________

Part of our research question was comparing models using advanced statistics to those that use more typical baseball statistics. We divided the stats in our data set into three categories: General stats, typical stats, and advanced stats. The general stats are basic player background statitsics like age, and position. The typical stats were statistics we could find on the basic page of a players baseballreference page. The advanced stats were things we could only find on baseballsavant. All analyses discussed were performed twice, once on the general variablegs combined with the typical statistics and once with the general variables combined with the advanced stats.

Our response variable is a categorical variable, with a 1 if a player was an all-star and a 0 if the player-was not an all-star. Thus, we decided to perform a logistic regression on our data. Prior to running our logistic regression we performed a lasso on both subsets of the data. We choose to use a lasso because we wanted to optimize the predictive power of the model. We also worried about correlation among our predictors, which would make step-wise selection methods work poorly.

We ran the lasso on both sets of models and then ran logistic regressions, measuring our success by the sensitivity of the model with a 32% probability threshold for categorization. The typical statistics model had a sensitivity of 55.5% and the advanced stats model had a sensitivity of about 44.5%. Neither of these results were very encouraging. Because of this, we went back and thought about our model. We realized that many of our predictors were highly correlated, which would make the lasso work porley for variable selection (which is obvious if you look at the lasso in our appendix and see that variables that our domain knowledge said were likely to be relevant such as batting average were selected out of the model). We also think that we had a problem with scarce data. Only 54 players in our data set were all-stars in 2019, which makes it difficult for our model to differentiate between all-stars and non-all stars, even with our low probability threshold for categorization as an all-star because it didn't have a large sample of all-stars to train on. The results for this can be found in the appendix for a view of our results, but because of the low sensitivity and methodological problems of these models we decided to try something new instead.

Because of these difficulties, we slightly reevaluated our research question. We decided to instead attempt to predict slugging percentage and on-base percentage from the advanced statistics (we couldn't predict using the typical statistics because these response variables are highly correlated, and in fact directly calculated from the typical statistics we measured). This is a valuable project because it can help us compare the outcomes we get from measuring performance through advanced metrics to the ones we get using typical statistics. We also used this to evaluate tendancies of where a player should be placed in a batting order, comparing the popular knowledge about batting order determiniations, which are based on the typical box-score statistics, to what advanced statcast knowledge says you should do.

Our methods for this were very similar to our methods for the all-star regressions, except we started by removing players who were on the Washington Nationals, so we could use them as a test data set for our model. We then used a lasso for the same reasons described above (fears of correlation among the predictors as well as our ability to optimize for preditive performance). We then ran a linear regression to predict slugging percentage and on-base percentage seperatly. we ran lassos to seperatly select variables for these two different regressions. Batting average and slugging percentage are continous variables thus making linear regression work quiet well. We chose not to include any interactions because each of the statistics are calculated to represent different aspects of player performance and are thus not meant to be taken together.

Following the linear regression we ran diagnostics such as making a QQ-Plot and Residual plot (which can be seen in the code chunks labelled statcast-obp, and slg-predict in the results section). Neither regression shows a clear pattern in the residual plots, though their do seem to be certian areas (over 0.3 for the OBP regression) where there is more data than lower on-base percentage values. There is also some worry about constant variance in the On-base percentage regression around the fitted value of 0.2, but in general, the assumptions of linearity and constant variance seem to be valid for both plots. The QQ-plots are meant to test normality. The on-base percentage QQ-Plot devaites from normality, but this deviation is largely in the tails, thus we think this assumption is fairly satisfied. The QQ-plot doesnt' deviate much for the slugging percentage regression and thus we think normality is fine for the slugging percentage regression. Independence is also satisfied because we are looking at the player level. If we were looking at typical statistics their migiht be some violation of independence based on the team the player played for but advanced statistics are designed such that they are independent. This is why exit-velocity is used rather than say, hit distance. Exit velocity cannot be altered by environmental conditions (and thus what team a player plays for), but distance a ball travels is altered by the elevation of a stadium. These linear regressions satisfy our assumptions and can thus be used for analysis.

Packages and Data

```
library(tidyverse)
library(tidymodels)
library(glmnet)
library(caret)
library(MASS)
library(lme4)
stats <- read.csv("data/stats.csv")


stats <- replace(stats, stats =="", NA)
stats <- stats %>%
  drop_na() %>%
  mutate(AVG300 = case_when(batting_avg >= .3 ~ "Greater than 300", TRUE ~ "Less than 300"
         HR40 = case_when(b_home_run >= 40 ~ "Greater than 40", TRUE ~ "Less than 40"), pi
view(stats)
```

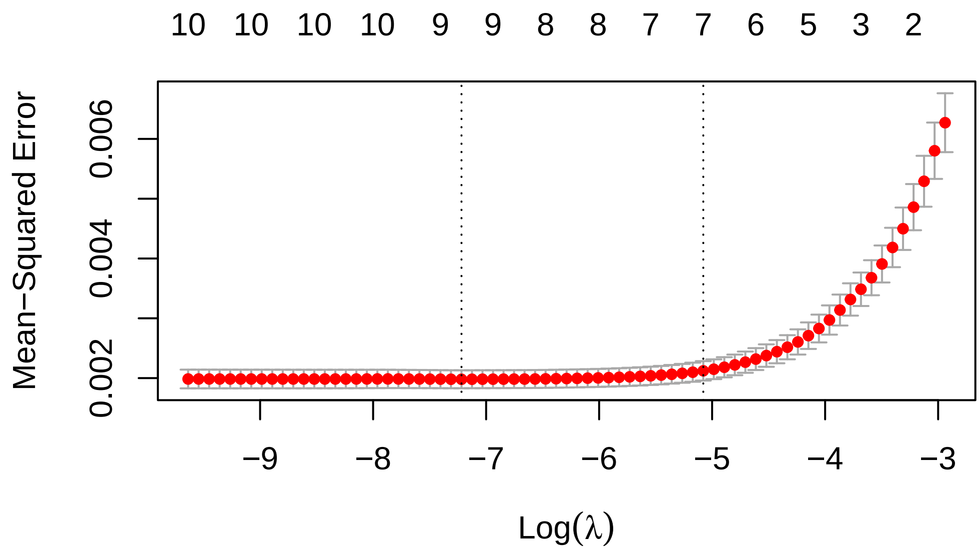
Results

```
# Remove Nationals from Data
rol_stats <- stats |>
  filter(Team != "WAS")
```

```
# obp percentage lasso for rol
set.seed(0)
y <- rol_stats$on_base_percent
x <- model.matrix(on_base_percent ~ launch_angle_avg + sweet_spot_percent +
                    barrel + solidcontact_percent + flareburner_percent +
                    hard_hit_percent + avg_hyper_speed + z_swing_percent +
                    oz_swing_percent + meatball_swing_percent, data = rol_stats)
m_lasso_cv <- cv.glmnet(x, y, alpha = 1)
best_lambda <- m_lasso_cv$lambda.min
best_lambda
```

```
[1] 0.0007332828
```

```
plot(m_lasso_cv)
```




```r
m_best <- glmnet(x, y, alpha = 1, lambda = best_lambda)
m_best$beta
```

```
11 x 1 sparse Matrix of class "dgCMatrix"
                                 s0
(Intercept)                   .
launch_angle_avg          -0.0001492831
sweet_spot_percent         0.0021040413
barrel                     0.0013140348
solidcontact_percent       0.0019525793
flareburner_percent        0.0036463362
hard_hit_percent           0.0009244833
avg_hyper_speed               .
z_swing_percent            0.0004967093
oz_swing_percent          -0.0028515915
meatball_swing_percent     0.0007020832
```

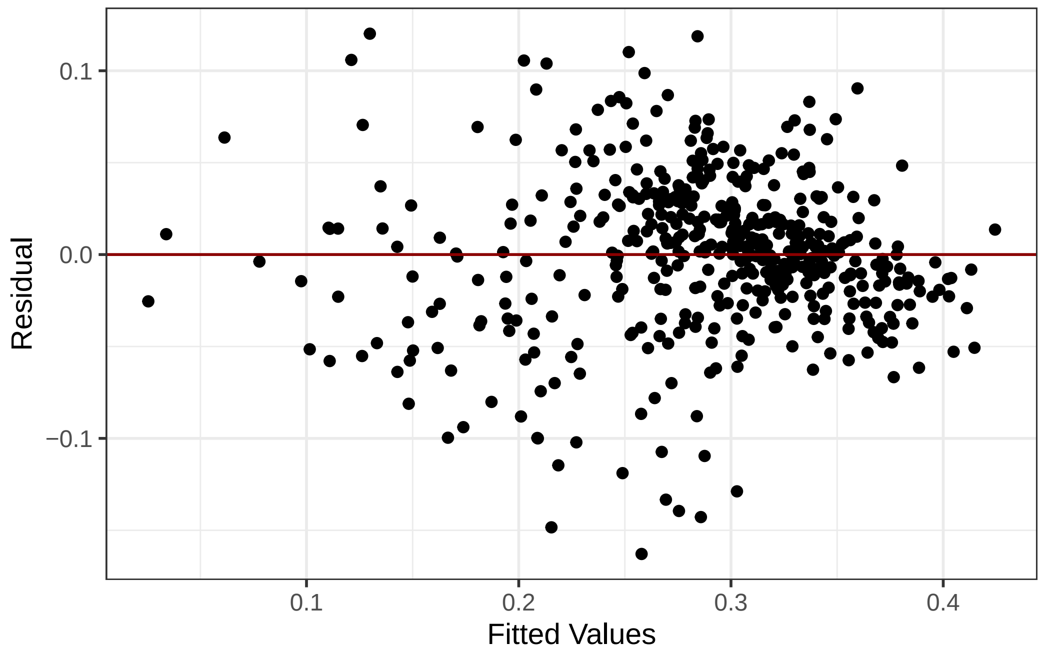
```r
# obp percentage prediction
m3 <- lm(on_base_percent ~ launch_angle_avg + sweet_spot_percent +
                    barrel + solidcontact_percent + flareburner_percent +
                    hard_hit_percent + z_swing_percent +
```

```r
                      oz_swing_percent + meatball_swing_percent,
   data = rol_stats)

summary(m3)
```

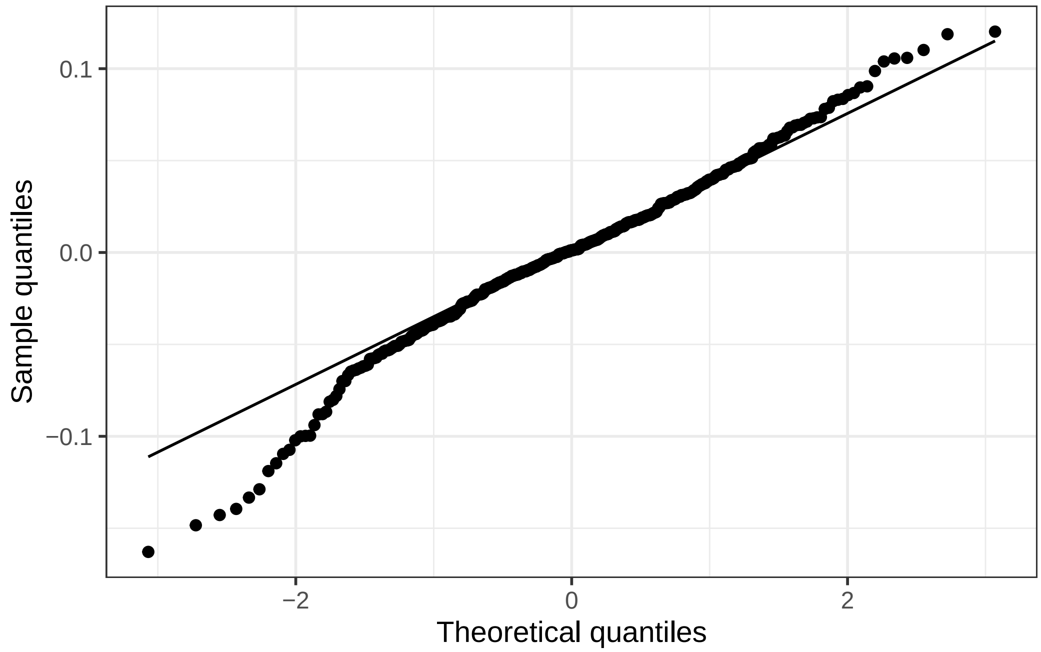
```

Call:
lm(formula = on_base_percent ~ launch_angle_avg + sweet_spot_percent +
    barrel + solidcontact_percent + flareburner_percent + hard_hit_percent +
    z_swing_percent + oz_swing_percent + meatball_swing_percent,
    data = rol_stats)

Residuals:
      Min        1Q    Median        3Q       Max
-0.162876 -0.022911  0.001195  0.026808  0.120152

Coefficients:
                         Estimate Std. Error t value Pr(>|t|)
(Intercept)             0.0660827  0.0198729   3.325 0.000955 ***
launch_angle_avg       -0.0004484  0.0003492  -1.284 0.199777
sweet_spot_percent      0.0021709  0.0004241   5.119 4.53e-07 ***
barrel                  0.0013401  0.0001692   7.923 1.80e-14 ***
solidcontact_percent    0.0022328  0.0008563   2.607 0.009421 **
flareburner_percent     0.0037954  0.0004981   7.619 1.49e-13 ***
hard_hit_percent        0.0008752  0.0003014   2.903 0.003871 **
z_swing_percent         0.0006523  0.0004437   1.470 0.142190
oz_swing_percent       -0.0030238  0.0003236  -9.343  < 2e-16 ***
meatball_swing_percent  0.0007344  0.0002548   2.883 0.004131 **
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.04353 on 456 degrees of freedom
Multiple R-squared:  0.7052,    Adjusted R-squared:  0.6994
F-statistic: 121.2 on 9 and 456 DF,  p-value: < 2.2e-16
```

```r
m3_aug <- augment(m3)
m3_aug |>
ggplot(aes(x = .fitted, y = .resid)) +
  geom_point() +
  geom_hline(yintercept = 0, color = "darkred") +
  labs(x = "Fitted Values",
```

```
       y = "Residual") +
  theme_bw()
```



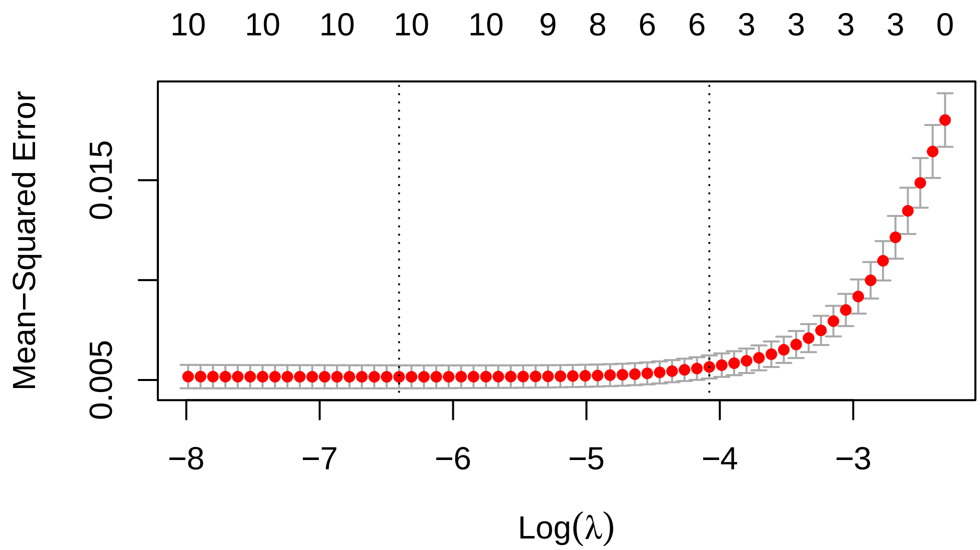

```
ggplot(m3_aug, aes(sample = .resid)) +
stat_qq() +
stat_qq_line() +
theme_bw() +
labs(x = "Theoretical quantiles",
y = "Sample quantiles")
```




```
# slugging percentage lasso
set.seed(0)
y <- rol_stats$slg_percent
x <- model.matrix(slg_percent ~ launch_angle_avg + sweet_spot_percent +
                    barrel + solidcontact_percent + flareburner_percent +
                    hard_hit_percent + avg_hyper_speed + z_swing_percent +
                    oz_swing_percent + meatball_swing_percent, data = rol_stats)
m_lasso_cv <- cv.glmnet(x, y, alpha = 1)
best_lambda <- m_lasso_cv$lambda.min
best_lambda
```

```
[1] 0.001653873
```

```
plot(m_lasso_cv)
```




```r
m_best <- glmnet(x, y, alpha = 1, lambda = best_lambda)
m_best$beta
```

```
11 x 1 sparse Matrix of class "dgCMatrix"
                                 s0
(Intercept)                .
launch_angle_avg           0.0004048935
sweet_spot_percent         0.0039725436
barrel                     0.0035324802
solidcontact_percent       0.0016992743
flareburner_percent        0.0027488668
hard_hit_percent           0.0014540632
avg_hyper_speed            0.0046269135
z_swing_percent            0.0010559425
oz_swing_percent          -0.0005609647
meatball_swing_percent     0.0006883043
```

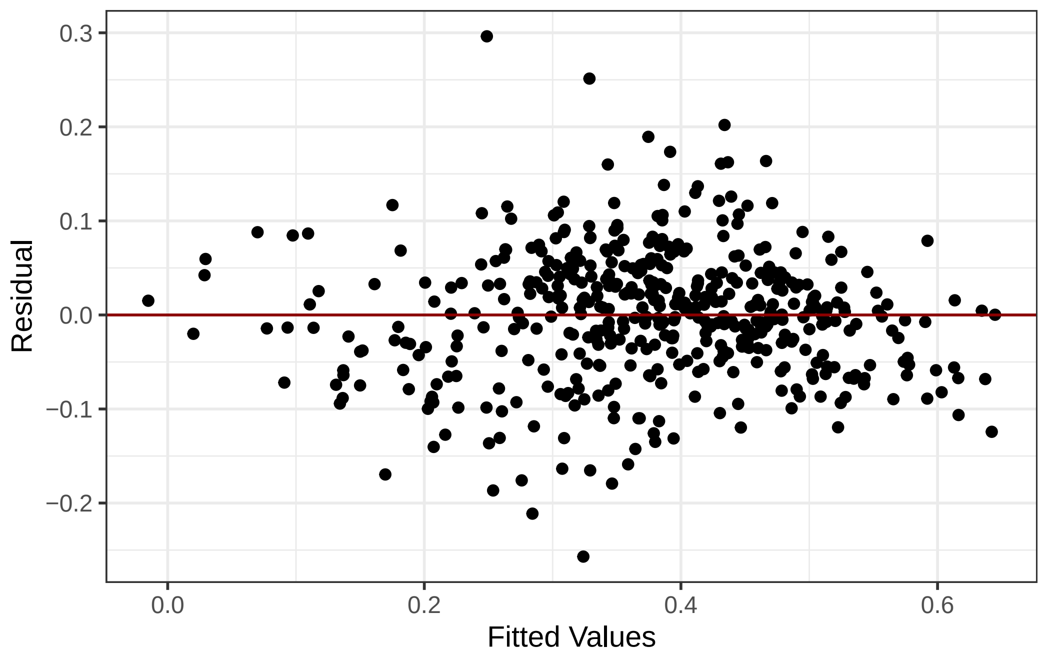
```r
# slugging percentage prediction
m4 <- lm(slg_percent ~ launch_angle_avg + sweet_spot_percent +
                    barrel + solidcontact_percent + flareburner_percent +
                    hard_hit_percent + avg_hyper_speed + z_swing_percent +
```

```r
                      oz_swing_percent + meatball_swing_percent,
   data = rol_stats)

summary(m4)
```

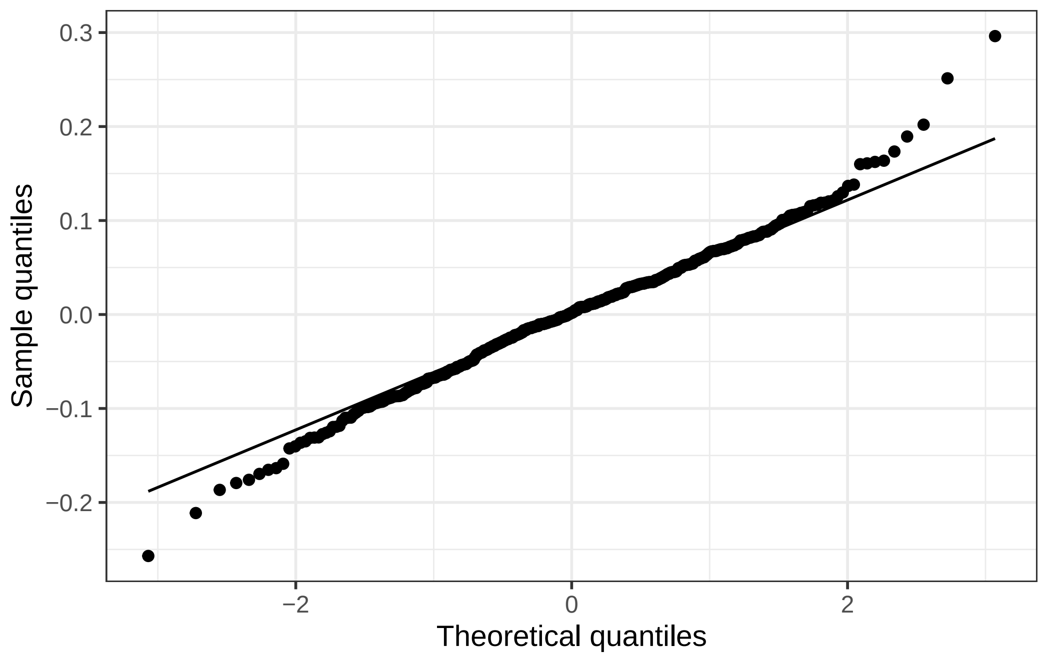
```

Call:
lm(formula = slg_percent ~ launch_angle_avg + sweet_spot_percent +
    barrel + solidcontact_percent + flareburner_percent + hard_hit_percent +
    avg_hyper_speed + z_swing_percent + oz_swing_percent + meatball_swing_percent,
    data = rol_stats)

Residuals:
      Min        1Q    Median        3Q       Max
-0.256940 -0.041724  0.001751  0.040773  0.296249

Coefficients:
                         Estimate Std. Error t value Pr(>|t|)
(Intercept)            -0.0790676  0.0327432  -2.415  0.01614 *
launch_angle_avg        0.0004505  0.0005610   0.803  0.42241
sweet_spot_percent      0.0038602  0.0006829   5.653 2.79e-08 ***
barrel                  0.0035814  0.0002773  12.916  < 2e-16 ***
solidcontact_percent    0.0021829  0.0013866   1.574  0.11612
flareburner_percent     0.0031794  0.0008195   3.880  0.00012 ***
hard_hit_percent        0.0010022  0.0010543   0.951  0.34233
avg_hyper_speed         0.0067277  0.0060541   1.111  0.26704
z_swing_percent         0.0013267  0.0007133   1.860  0.06351 .
oz_swing_percent       -0.0008821  0.0005243  -1.683  0.09314 .
meatball_swing_percent  0.0007281  0.0004096   1.777  0.07618 .
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 0.06994 on 455 degrees of freedom
Multiple R-squared:  0.7345,    Adjusted R-squared:  0.7287
F-statistic: 125.9 on 10 and 455 DF,  p-value: < 2.2e-16
```

```r
m4_aug <- augment(m4)
m4_aug |>
ggplot(aes(x = .fitted, y = .resid)) +
  geom_point() +
  geom_hline(yintercept = 0, color = "darkred") +
```

```
  labs(x = "Fitted Values",
       y = "Residual") +
  theme_bw()
```




```
ggplot(m4_aug, aes(sample = .resid)) +
stat_qq() +
stat_qq_line() +
theme_bw() +
labs(x = "Theoretical quantiles",
y = "Sample quantiles")
```



The final models for predicting on base and slugging percentage are both linear regressions. Their variables were chosen using LASSO for the advanced statistics, as defined by the data dictionary. These models were trained on data from the 2019 season for each team, aside from the Washington Nationals, which were used as a test set. We decided it was most appropriate to use a whole team as a test set because it provided a variety of positions and was most useful in terms of comparing expected lineups (lineups that would be expected to generate the most runs) to actual lineups. Now, the LASSO for the on base percentage model had a best lambda of approximately 0.000733 using 10-fold cross validation, and the variables included were launch_angle_avg, sweet_spot_percent, barrel, solidcontact_percent, flareburner_percent, hard_hit_percent, z_swing_percent, oz_swing_percent, and meatball_swing_percent. After using these variables in the linear regression, we were able to conclude that the assumptions for a linear regression were satisfied because in the residual plot we see mostly symmetrically distributed and evenly spaced observations, satisfying the criteria for linearity and constant variance. Furthermore, there is a slight correlation between the predictors and no clear deviation patterns, aside from the ends, in the Q-Q plot, satisfying the assumption for independence and normality. Moreover, 69.94% of the variation in OBP was explained by the predictors in the model (adjusted R-squared).

Among the significant variables in this linear regression, one of the most notable positive predictors was flareburner_percent. Holding all other predictors constant, for each 1% increase in flareburner_percent, expected on base percentage increases by approximately 0.380%. On the other hand, one of the most notable negative predictors was oz_swing_percent. Holding all other predictors constant, for each 1% increase in flareburner_percent, expected on base percentage decreases by approximately 0.302%. These results both make sense in the context of the game because being able to hit a variety of pitches would make a batter better, while swinging at bad pitches would make a batter worse. Moving on to slugging percentage, we observed a best lambda of approximately 0.00165, and the variables included were launch_angle_avg, sweet_spot_percent, barrel, solidcontact_percent, flareburner_percent, hard_hit_percent, avg_hyper_speed, z_swing_percent, oz_swing_percent, and meatball_swing_percent. Like the other model, all assumptions for a linear regression were met, and 72.87% of the variation was explained by the predictors in the model.

Among the significant variables in this model, the most notable predictors were sweet_spot_percent and barrel. Holding all other predictors constant, for each 1% increase in each of these predictors, expected slugging percentage increases by approximately 0.386% and 0.358% respectively. oz_swing_percent still had a negative correlation in this model, but its impact was far less. These results make sense as slugging percentage is based on how good a player's hit is, and both of these predictors are strongly associated with doubles, triples, and home runs, while OBP is only based on whether a player reaches a base.

Discussion

In this study, we aimed to develop predictive models for on base percentage and slugging percentage using advanced statistics in the MLB. Our results showed that the best models for this were linear regressions, and the selected variables were significant predictors of on base and slugging percentage in the 2019 MLB season. Our findings suggest that advanced statistics can be used to predict player performance and improve team strategies. For example, our models highlighted the importance of flareburner_percent and oz_swing_percent in predicting on base percentage. These findings suggest that players who can hit a variety of pitches are more likely to perform better, while those who swing at bad pitches are less likely to reach base. These insights can be useful for teams when selecting players or developing strategies to optimize their lineups. Moreover, our models revealed that sweet_spot_percent and barrel were the most notable predictors of slugging percentage. These findings suggest that players who can consistently hit the ball at certain angles and on a specific part of their bat are more likely to generate more extra base hits. Therefore, teams can use this information to identify players who are more likely to produce high slugging percentages, which is a key factor in offensive success in baseball. Despite the promising results of our study, there are some limitations to our analysis. First, our data only covers the 2019 season, and it is possible that our models may not be generalizable to other seasons or contexts. Furthermore, this is a generalized model that does not take into account unique player and stadium advantages. For example, certain players hit far better against right handed pitchers than left handed pitchers, and a team's optimal lineup should change based on the type of pitcher they are facing. Also, stadiums at higher altitudes, like the one in Colorado, favor more powerful hitters because its altitude is higher, which makes the baseball go approximately 10% further (CITE). This would also warrant a game specific change in the lineup. However, the biggest limitation of our model is that it only predicts on base and slugging percentage for players who are already in the major leagues. While this is helpful for decisions for future seasons, the model is not able to shine any light on how a player may perform before they actually play in the major leagues. Fortunately, major league baseball franchises have a well developed farm system where players spend years in the minor leagues. This is the highest level of baseball, aside from the major league, in the United States, and they track all of the same stats that the MLB does. Therefore, a natural progression for future work would be to try and predict a player's on base and slugging percentage based on their advanced statistics in the minor leagues. This would give teams a better gauge on how players would perform in the major leagues, rather than how players would perform in future seasons after they have already made it to the majors. With this type of model, teams would have a better idea of how to strategically bring up players and put the team that gives them the best chance of winning on the field.

```r
# Subset for nationals
nationals_stats <- stats |>
  filter(Team == "WAS")

# Predict
pred_obp <- predict(m3, nationals_stats)
pred_slg <- predict(m4, nationals_stats)

# Add to DF
nationals_stats <- nationals_stats |>
  mutate(Predicted_OBP = pred_obp,
         Predicted_SLG = pred_slg)

# Display
nationals_pred_stats <- nationals_stats[ , c("last_name", "first_name", "Position",
                                             "Predicted_OBP", "Predicted_SLG")]
print(nationals_pred_stats)
```

```
      last_name  first_name Position Predicted_OBP Predicted_SLG
1       Sanchez      Anibal       SP    0.07961863     0.1275467
2  Kendrick III       Howie       1B    0.35010631     0.5018412
3        Suzuki        Kurt        C    0.33371079     0.4386308
4      Scherzer         Max       SP    0.20971532     0.2277935
5     Zimmerman        Ryan       1B    0.28173250     0.3623809
6    Hellickson      Jeremy       SP    0.18646363     0.1018744
7         Gomes         Yan        C    0.27653808     0.3802595
8        Rendon     Anthony       3B    0.41590403     0.6003806
9     Strasburg     Stephen       SP    0.24879639     0.2691283
10        Adams        Matt       1B    0.31008350     0.4743130
11       Corbin     Patrick       SP    0.14354136     0.1522589
```

```
12       Taylor  Michael A.      CF   0.30846931    0.3732054
13       Dozier       Brian      2B   0.33754471    0.4167834
14         Difo      Wilmer      SS   0.29585135    0.3388092
15        Eaton        Adam      RF   0.31271661    0.4048301
16       Turner        Trea      SS   0.32119304    0.4405392
17       Robles      Victor      CF   0.28957201    0.3891161
18    Stevenson      Andrew      PH   0.26422635    0.3259004
19         Soto        Juan      LF   0.40647774    0.5769956
20         Noll        Jake      PH   0.15891038    0.1942342
```

Variable Selections and Regressions we tried

```
# LASSO Variable Selection Basic Stats
y <- stats$All.Star
x <- model.matrix(All.Star ~ player_age + b_ab + b_total_pa + b_total_hits + b_home_run +
                    b_double + b_triple + b_home_run * HR40 + b_strikeout + b_walk +
                    batting_avg + slg_percent + on_base_percent + Position, data = stats)
m_lasso_cv <- cv.glmnet(x, y, alpha = 1)
best_lambda <- m_lasso_cv$lambda.min
best_lambda
```

```
[1] 0.0005971366
```

```
m_best <- glmnet(x, y, alpha = 1, lambda = best_lambda)
m_best$beta
```

```
29 x 1 sparse Matrix of class "dgCMatrix"
                                    s0
(Intercept)                .
player_age                -0.0029767462
b_ab                      -0.0016283645
b_total_pa                 .
b_total_hits               0.0043184447
b_home_run                 0.0117187722
AVG300Less than 300       -0.1242223268
batting_avg                .
b_double                   0.0025726213
b_triple                   0.0043507975
HR40Less than 40          -0.1082317461
```

```
b_strikeout                      -0.0007458214
b_walk                            0.0033086510
slg_percent                       0.0627793959
on_base_percent                  -0.3042934878
Position2B                        0.0395393326
Position3B                       -0.0346355227
PositionC                         0.0650214403
PositionCF                        0.0537743645
PositionCH                        0.0079932117
PositionDH                       -0.0385415609
PositionDNP                      -0.0066508435
PositionLF                       -0.0511356574
PositionPH                        0.0034339028
PositionRF                        0.0018072204
PositionSP                        0.1709299345
PositionSS                        0.0457111491
AVG300Less than 300:batting_avg   .
b_home_run:HR40Less than 40       0.0002396631
```

```
# LASSO Variable Selection Advanced Stats
y <- stats$All.Star
x <- model.matrix(All.Star ~ player_age + launch_angle_avg + sweet_spot_percent +
                    barrel + solidcontact_percent + flareburner_percent +
                    hard_hit_percent + avg_hyper_speed + z_swing_percent +
                    oz_swing_percent + meatball_swing_percent, data = stats)
m_lasso_cv <- cv.glmnet(x, y, alpha = 1)
best_lambda <- m_lasso_cv$lambda.min
best_lambda
```

```
[1] 0.005820234
```

```
m_best <- glmnet(x, y, alpha = 1, lambda = best_lambda)
m_best$beta
```

```
12 x 1 sparse Matrix of class "dgCMatrix"
                                s0
(Intercept)             .
player_age             -0.0018096273
launch_angle_avg       -0.0001599061
sweet_spot_percent      .
```

```
barrel                    0.0080466625
solidcontact_percent     -0.0030825651
flareburner_percent      -0.0018259472
hard_hit_percent         -0.0018130937
avg_hyper_speed           .
z_swing_percent           .
oz_swing_percent          .
meatball_swing_percent -0.0021930005
```

```
#Basic model
m1 <- glm(All.Star ~ player_age + b_ab +  b_total_hits +
                    b_double + b_triple + b_home_run + b_strikeout +
                    b_bb_percent + AVG300 + slg_percent +
                    on_base_percent + Position,
  data = stats,
  family = "binomial"
)
tidy(m1)
```

```
# A tibble: 24 x 5
   term                estimate std.error statistic p.value
   <chr>                  <dbl>     <dbl>     <dbl>   <dbl>
 1 (Intercept)         -5.96      2.78       -2.15   0.0319
 2 player_age          -0.0369    0.0570     -0.647  0.518
 3 b_ab                -0.00925   0.00925    -1.00   0.317
 4 b_total_hits         0.0450    0.0300      1.50   0.133
 5 b_double             0.0182    0.0411      0.443  0.657
 6 b_triple            -0.0469    0.116      -0.404  0.686
 7 b_home_run           0.0719    0.0524      1.37   0.170
 8 b_strikeout         -0.00188   0.00930    -0.203  0.839
 9 b_bb_percent         0.155     0.105       1.48   0.140
10 AVG300Less than 300 -0.617     0.768      -0.804  0.421
# ... with 14 more rows
```

```
m1_aug <- augment(m1) %>%
  mutate(prob = exp(.fitted)/(1 + exp(.fitted)),
         pred_leg = ifelse(prob > 0.32, "All-Star", "Not All-Star"))
table(m1_aug$pred_leg, m1_aug$All.Star)
```

```
              0   1
  All-Star      22  30
  Not All-Star 410  24
```

```
#Advanced model
m2 <- glm(All.Star ~ player_age + launch_angle_avg +
                    barrel + solidcontact_percent + flareburner_percent +
                    hard_hit_percent + meatball_swing_percent,
  data = stats,
  family = "binomial"
)
tidy(m2)
```

```
# A tibble: 8 x 5
  term                   estimate std.error statistic  p.value
  <chr>                     <dbl>     <dbl>     <dbl>    <dbl>
1 (Intercept)             1.38       1.76       0.785 4.32e- 1
2 player_age             -0.0361     0.0468    -0.772 4.40e- 1
3 launch_angle_avg       -0.00881    0.0283    -0.312 7.55e- 1
4 barrel                  0.0852     0.0136     6.27  3.56e-10
5 solidcontact_percent   -0.0805     0.0943    -0.854 3.93e- 1
6 flareburner_percent    -0.0129     0.0379    -0.341 7.33e- 1
7 hard_hit_percent       -0.0256     0.0263    -0.974 3.30e- 1
8 meatball_swing_percent -0.0358     0.0162    -2.21  2.72e- 2
```

```
m2_aug <- augment(m2) %>%
  mutate(prob = exp(.fitted)/(1 + exp(.fitted)),
         pred_leg = ifelse(prob > 0.32, "All-Star", "Not All-Star"))
table(m2_aug$pred_leg, m2_aug$All.Star)
```

```
              0   1
  All-Star      22  20
  Not All-Star 410  34
```

ARTICLE ABOUT BATTING ORDER STRATEGY: https://www.sportsbettingdime.com/guides/strategy/ba
order-sabermetrics/