ME 233 Advance Control II

Lecture 5 Least Squares Estimation

(ME233 Class Notes pp. LS1-LS5)

Marginal Expectation (review)

Let X and Y be continuous random vectors with joint PDF $p_{XY}(x,y)$

<u>Marginal Expectation</u> (mean) of X

$$m_X = E\{X\}$$

$$= \int_{R_x} \underbrace{\int_{R_y} x \, p_{XY}(x, y) \, dy \, dx}_{x p_X(x)}$$

Notation

Let X and Y be continuous random vectors with joint PDF $p_{XY}(x,y)$

Let x and y be respectively outcomes of X and Y and

$$x \in R_x \subseteq R^{n_x} \quad y \in R_y \subseteq R^{n_y}$$

$$p_{XY}: R_x \times R_y \to R_+$$

Marginal Expectation (review)

Let X and Y be continuous random variables with joint PDF $p_{XY}(x,y)$

Marginal Expectation (mean) of X

$$m_X = E\{X\} = \int_{R_x} x \, p_X(x) dx$$

new notation (following the ME233 class notes)

Marginal Expectation \hat{x}

- \widehat{x} is the minimum least squares marginal estimator of X, i.e.
- For any deterministic vector Z

$$E\{\|X - \hat{x}\|^2\} \le E\{\|X - z\|^2\}$$
Euclidean norm

Conditional Expectation (review)

Let X and Y be continuous random vectors with joint PDF $p_{XY}(x,y)$

Conditional Expectation (conditional mean)

of X given and outcome Y = y

$$m_{X|y} = E\{X|Y = y\}$$
$$= \int_{R_x} x \, p_{X|y}(x) dx$$

Marginal Expectation \hat{x}

$$E\{\|X - \hat{x}\|^2\} \le E\{\|X - z\|^2\}$$

Proof:

$$E\{\|X - z\|^2\} = E\{\|(X - \hat{x}) - (z - \hat{x})\|^2\}$$

$$= E\{\|X - \hat{x}\|^2 + \|z - \hat{x}\|^2 - 2(z - \hat{x})^T (X - \hat{x})\}$$

$$= E\{\|X - \hat{x}\|^2\} + \|z - \hat{x}\|^2 - 2(z - \hat{x})^T E\{X - \hat{x}\}$$

$$\geq E\{\|X - \hat{x}\|^2\}$$

Conditional Expectation (review)

Conditional Expectation (conditional mean)

of X given and outcome Y = y

$$\begin{split} m_{X|y} &= \int_{R_x} x \, p_{X|y}(x) dx \\ &= \int_{R_x} x \, \left(\frac{p_{XY}(x,y)}{p_Y(y)} \right) dx \\ &= \hat{x}|_{y} \quad \text{\tiny new notation}_{\text{\tiny (following the ME233 class notes)}} \end{split}$$

Conditional Expectation $\hat{x}|_y$

Notice that the conditional expectation $\hat{x}|_{y}$

$$\widehat{x}|_{y} = \int_{R_{x}} x \frac{p_{XY}(x,y)}{p_{Y}(y)} dx$$

can be interpreted as a function of the random variable Y.

$$\widehat{X}|_{Y} = \int_{R_{x}} x \frac{p_{XY}(x,Y)}{p_{Y}(Y)} dx$$

Conditional Expectation $\widehat{X}|_{Y}$

Lemma:

For any function $f(\cdot)$ of the random vector Y, with the appropriate dimensions

$$E\{f(Y)X\} = E\{f(Y)\hat{X}|_{Y}\}$$

we can replace X by its conditional expectation $\widehat{X}|_{Y}$

Marginal Expectation \hat{x}

$$E\{f(Y)X\} = E\{f(Y)\hat{X}|_{Y}\}$$

Proof:

First examine the left-hand side:

$$E\{f(Y)X\} = \int_{R_y} \int_{R_x} f(y)x \underbrace{p_{XY}(x,y)} dx dy$$

$$= \int_{R_y} \int_{R_x} f(y)x \underbrace{p_{X|y}(x)p_Y(y)} dx dy$$

$$= \int_{R_y} f(y) \left[\int_{R_x} x p_{X|y}(x) dx \right] p_Y(y) dy$$

Marginal Expectation \hat{x}

$$E\{f(Y)X\} = E\{f(Y)\hat{X}|_{Y}\}$$

Proof:

First examine the left-hand side:

$$E\{f(Y)X\} = \int_{R_y} f(y) \underbrace{\left[\int_{R_x} x \, p_{X|y}(x) dx\right]}_{\hat{x}|_{\mathcal{Y}}} p_Y(y) dy$$

$$E\{f(Y)X\} = \int_{R_y} f(y)\hat{x}|_y p_Y(y)dy$$

Marginal Expectation \hat{x}

$$E\{f(Y)X\} = E\{f(Y)\hat{X}|_{Y}\}$$

Proof:

Now examine the right-hand side:

$$E\{f(Y)\hat{X}|_Y\} = \int_{R_y} \int_{R_x} \underbrace{f(y)\hat{x}|_y}_y p_{XY}(x,y) dx \, dy$$
 Not a function of x

$$E\{f(Y)\hat{X}|_{Y}\} = \int_{R_{y}} f(y)\hat{x}|_{y} \underbrace{\left[\int_{R_{x}} p_{XY}(x,y)dx\right]}_{p_{Y}(y)} dy$$

Marginal Expectation \hat{x}

$$E\{f(Y)X\} = E\{f(Y)\hat{X}|_{Y}\}$$

Proof:

Therefore,

$$E\{f(Y)X\} = \int_{R_y} f(y)\hat{x}|_y p_Y(y)dy$$
$$= E\{f(Y)\hat{X}_Y\}$$

Conditional Expectation $\widehat{X}|_{Y}$

Theorem:

 $\widehat{X}|_{Y}$ is the least squares minimum estimator of ${\it X}$ given ${\it Y}$, i.e.

$$E\{||X - \hat{X}|_Y||^2\} \le E\{||X - f(Y)||^2\}$$

for all functions $f(\cdot)$ of Y of appropriate dimensions

$$||X||^2 = X^T X$$

Marginal Expectation \hat{x}

$$E\{||X - \hat{X}|_Y||^2\} \le E\{||X - f(Y)||^2\}$$

Proof:

$$E\{\|X - f(Y)\|^2\} = E\{\|(X - \hat{X}|_Y) - (f(Y) - \hat{X}|_Y)\|^2\}$$

$$= E\{\|X - \hat{X}|_Y\|^2 + \|f(Y) - \hat{X}|_Y\|^2$$

$$-2(f(Y) - \hat{X}|_Y)^T(X - \hat{X}|_Y)\}$$

$$= E\{\|X - \hat{X}|_Y\|^2\} + E\{\|f(Y) - \hat{X}|_Y\|^2\}$$

$$-2E\{(f(Y) - \hat{X}|_Y)^TX\} + 2E\{(f(Y) - \hat{X}|_Y)^T\hat{X}|_Y\}$$

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Marginal Expectation \hat{x}

$$E\{||X - \hat{X}|_Y||^2\} \le E\{||X - f(Y)||^2\}$$

Proof:

Define
$$g(Y) := (f(Y) - \hat{X}|_Y)^T$$

$$E\{\|X - f(Y)\|^{2}\} = E\{\|X - \hat{X}|_{Y}\|^{2}\} + E\{\|f(Y) - \hat{X}|_{Y}\|^{2}\}$$
$$-2E\{g(Y)X\} + 2E\{g(Y)\hat{X}|_{Y}\}$$

Since $||f(Y) - \hat{X}|_Y||^2 \ge 0$ for all outcomes,

$$E\{\|f(Y) - \hat{X}|_{Y}\|^{2}\} \ge 0$$

$$\Rightarrow E\{\|X - f(Y)\|^2\} \ge E\{\|X - \hat{X}|_Y\|^2\}$$

Conditional Expectation for Gaussians (review)

$$\text{When } \begin{bmatrix} X \\ Y \end{bmatrix} \sim N \left(\begin{bmatrix} m_X \\ m_Y \end{bmatrix}, \ \begin{bmatrix} \Lambda_{XX} & \Lambda_{XY} \\ \Lambda_{YX} & \Lambda_{YY} \end{bmatrix} \right)$$

$$X|_{y} \sim N(\widehat{x}_{y}, \ \land_{X|yX|y})$$

where

$$\widehat{x}|_{y} = \widehat{x} + \bigwedge_{XY} \bigwedge_{YY}^{-1} (y - \widehat{y})$$
$$\bigwedge_{X|yX|y} = \bigwedge_{XX} - \bigwedge_{XY} \bigwedge_{YY}^{-1} \bigwedge_{YX}$$

Conditional Mean for Gaussians

$$\text{When } \begin{bmatrix} X \\ Y \end{bmatrix} \sim N \left(\begin{bmatrix} m_X \\ m_Y \end{bmatrix}, \ \begin{bmatrix} \Lambda_{XX} & \Lambda_{XY} \\ \Lambda_{YX} & \Lambda_{YY} \end{bmatrix} \right)$$

$$\widehat{X}|_{Y} = \widehat{x} + \Lambda_{XY} \Lambda_{YY}^{-1} (Y - \widehat{y})$$

$$E\{\hat{X}|_{Y}\} = \hat{x} + \bigwedge_{XY} \bigwedge_{YY}^{-1} E\{Y - \hat{y}\}^{0}$$
$$= \hat{x}$$

Conditional Mean for Gaussians

$$\text{When } \begin{bmatrix} X \\ Y \end{bmatrix} \sim N \left(\begin{bmatrix} m_X \\ m_Y \end{bmatrix}, \ \begin{bmatrix} \Lambda_{XX} & \Lambda_{XY} \\ \Lambda_{YX} & \Lambda_{YY} \end{bmatrix} \right)$$

$$\tilde{X}|_{y} = X - (\hat{x} + \bigwedge_{XY} \bigwedge_{YY}^{-1} (y - \hat{y}))
= \tilde{X} - \bigwedge_{XY} \bigwedge_{YY}^{-1} (y - \hat{y})$$



$$\begin{split} \tilde{X}|_{Y} &= \tilde{X} - \Lambda_{XY} \Lambda_{YY}^{-1} (Y - \hat{y}) \\ &= \tilde{X} - \Lambda_{XY} \Lambda_{YY}^{-1} \tilde{Y} \end{split}$$

Conditional Mean for Gaussians

$$\text{When } \begin{bmatrix} X \\ Y \end{bmatrix} \sim N \left(\begin{bmatrix} m_X \\ m_Y \end{bmatrix}, \ \begin{bmatrix} \Lambda_{XX} & \Lambda_{XY} \\ \Lambda_{YX} & \Lambda_{YY} \end{bmatrix} \right)$$

$$\tilde{X}|_{Y} = \tilde{X} - \Lambda_{XY}\Lambda_{YY}^{-1}\tilde{Y}$$

$$E\{\tilde{X}|_{Y}\} = E\{\tilde{X}\} - \bigwedge_{XY} \bigwedge_{YY}^{-1} E\{\tilde{Y}\}^{0}$$
$$= 0$$

Least Squares Estimation: Property 1

- The conditional estimation error $\ \tilde{X}_{|_{Y}} \ \$ and $\ Y \ \$ are $\ \textit{uncorrelated}$

$$E\{\tilde{X}_{|_{Y}}\tilde{Y}^{T}\} = 0$$

• $ilde{X}_{|_Y}$ and $\hat{X}_{|_Y}$ are **orthogonal**

$$E\{\tilde{X}_{|Y}\hat{X}_{|Y}^T\} = \mathbf{0} \qquad \text{and} \qquad E\{\tilde{X}_{|Y}^T\hat{X}_{|Y}\} = \mathbf{0}$$

Least Squares Estimation: Property 1

$$E\{\tilde{X}_{|_{Y}}\tilde{Y}^{T}\}=0$$

Proof

$$\begin{split} E\{\tilde{X}_{|Y}\tilde{Y}^T\} &= E\{(\tilde{X} - \Lambda_{XY}\Lambda_{YY}^{-1}\tilde{Y})\tilde{Y}^T\} \\ &= E\{\tilde{X}\tilde{Y}^T\} - \Lambda_{XY}\Lambda_{YY}^{-1}E\{\tilde{Y}\tilde{Y}^T\} \\ &= \Lambda_{XY} - \Lambda_{XY}\Lambda_{YY}^{-1}\Lambda_{YY} \\ &= 0 \end{split}$$

Least Squares Estimation: Property 1

$$E\{\tilde{X}_{|_{Y}}\hat{X}_{|_{Y}}^{T}\}=0$$

Proof

$$\begin{split} E\{\tilde{X}_{|Y}\hat{X}_{|Y}^T\} &= E\{\tilde{X}_{|Y}(\hat{x} + \Lambda_{XY}\Lambda_{YY}^{-1}\tilde{Y})^T\} \\ &= E\{\tilde{X}_{|Y}\}\hat{x}^T + E\{\tilde{X}_{|Y}Y^T\}\Lambda_{YY}^{-1}\Lambda_{XY}^T \\ &= 0 \end{split}$$

Least Squares Estimation: Property 1

$$E\{\tilde{X}_{|Y}^T\hat{X}_{|Y}\} = 0$$

Proof

$$\Rightarrow \quad E\{\tilde{X}_{|Y}^T\hat{X}_{|Y}\} = E\{\operatorname{trace}(\tilde{X}_{|Y}\hat{X}_{|Y}^T)\}$$

Why does trace commute with expectation?

$$= trace(E\{\tilde{X}_{|Y}\hat{X}_{|Y}^T\})$$
$$= trace(0) = 0$$

Deterministic interpretation of Property 1 X $\tilde{X}_{|Y}$ $\hat{X}_{|Y}$

Recursive LS Estimation

Let X, Y and Z be jointly Gaussian R.V.s

Recursive LS Estimation

1. Assume that we already know of outcome Y=y and we have obtained $\hat{x}_{|y}=E\{X|Y=y\}$

$$\hat{x}_{|y} = \hat{x} + \Lambda_{XY} \underbrace{\Lambda_{YY}^{-1}}_{y} (y - \hat{y}) \underbrace{}_{y}$$

inverse of an M × M matrix

Recursive LS Estimation

- 1. Assume that we already know of outcome Y=y and we have obtained $\hat{x}_{|y}=\hat{x}+\Lambda_{XY}\Lambda_{YY}^{-1}\left(y-\hat{y}\right)$
- 2. Now we also know the outcome Z = z

How do we efficiently compute

$$\hat{x}_{|yz} = E\{X|Y = y, Z = z\}$$
 ?

Least Squares Estimation: Property 2

Assume that $\ \, \Lambda_{ZY} = E\{\tilde{Z}\tilde{Y}^T\} = 0$ Then,

$$\hat{X}_{|YZ} = \hat{X}_{|Y} + \left(\tilde{X}_{|Y}\right)_{|Z}$$

$$\Lambda_{\tilde{X}_{|YZ}\tilde{X}_{|YZ}} = \Lambda_{\tilde{X}_{|Y}\tilde{X}_{|Y}} - \Lambda_{XZ}\Lambda_{ZZ}^{-1}\Lambda_{ZX}$$

where

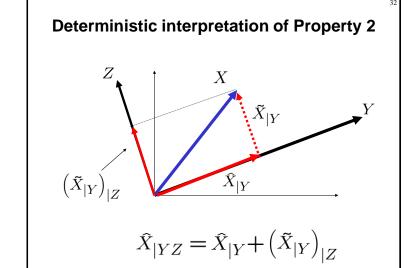
$$\begin{split} \widehat{X}_{|Y} &= \widehat{x} + \bigwedge_{XY} \bigwedge_{YY}^{-1} (Y - \widehat{y}) \\ \left(\widetilde{X}_{|Y} \right)_{|Z} &= \bigwedge_{XZ} \bigwedge_{ZZ}^{-1} (Z - \widehat{z}) \\ \bigwedge_{\widetilde{X}_{|Y} \widetilde{X}_{|Y}} &= \bigwedge_{XX} - \bigwedge_{XY} \bigwedge_{YY}^{-1} \bigwedge_{YX} \end{split}$$

Non-Recursive LS Estimation

- 1) Define the vector $W = \left[egin{array}{c} Z \\ Y \end{array} \right] \quad \hat{w} = \left[egin{array}{c} \hat{z} \\ \hat{y} \end{array} \right]$
- 2) Compute $\hat{x}_{|w} = E\{X|Y=y,\,Z=z\}$

$$\widehat{x}_{|w} = \widehat{x} + \bigwedge_{XW} \bigwedge_{WW}^{-1} (w - \widehat{w})$$

$$\lim_{n \text{ inverse of an } (p+M) \times (p+M) \text{ matrix}} p + M$$



Least Squares Estimation: Property 2

$$\left(\tilde{X}_{|Y}\right)_{|Z} = \Lambda_{XZ}\Lambda_{ZZ}^{-1}\left(Z - \hat{z}\right)$$

Proof:

$$(\tilde{X}_{|Y})_{|Z} = E\{\tilde{X}_{|Y}\} + \Lambda_{\tilde{X}_{|Y}Z}\Lambda_{ZZ}^{-1}(Z - \hat{z})$$

$$\boldsymbol{\Lambda}_{\tilde{\boldsymbol{X}}_{|Y}\boldsymbol{Z}} = E\{\tilde{\boldsymbol{X}}_{|Y}\tilde{\boldsymbol{Z}}^T\} = E\left\{\left[\tilde{\boldsymbol{X}} - \boldsymbol{\Lambda}_{\boldsymbol{X}\boldsymbol{Y}} \, \boldsymbol{\Lambda}_{\boldsymbol{Y}\boldsymbol{Y}}^{-1}\tilde{\boldsymbol{Y}}\right]\tilde{\boldsymbol{Z}}^T\right\}$$

$$=\underbrace{E\{\tilde{X}\tilde{Z}^T\}}_{\text{because }Z \text{ and }Y} - \bigwedge_{YZ} \bigwedge_{YZ} - \bigwedge_{YZ} \underbrace{A_{YY}^{-1}E\{\tilde{X}\tilde{Z}^T\}}_{\text{are uncorrelated}}$$

Least Squares Estimation: Property 2

$$\hat{X}_{|YZ} = \hat{X}_{|Y} + \left(\tilde{X}_{|Y}\right)_{|Z}$$

Proof:

$$\widehat{X}_{|YZ} = \widehat{x} + \underbrace{\Lambda_{XW}}_{XW} \underbrace{\Lambda_{WW}^{-1}}_{WW} \underbrace{(W - \widehat{w})}_{\left[\begin{array}{ccc} \Lambda_{XZ} & \Lambda_{XY} \end{array} \right] \left[\begin{array}{ccc} \tilde{Z} \\ 0 & \Lambda_{YY} \end{array} \right]}_{\left[\begin{array}{ccc} \tilde{Z} \\ \tilde{Y} \end{array} \right]}$$

$$\hat{X}_{|YZ} = \underbrace{\hat{x} + \bigwedge_{XY} \bigwedge_{YY}^{-1} \tilde{Y}}_{\hat{X}_{|Y}} + \underbrace{\bigwedge_{XZ} \bigwedge_{ZZ}^{-1} \tilde{Z}}_{|Z}$$

Least Squares Estimation: Property 2

$$\Lambda_{\tilde{X}_{|YZ}\tilde{X}_{|YZ}} = \Lambda_{\tilde{X}_{|Y}\tilde{X}_{|Y}} - \Lambda_{XZ}\Lambda_{ZZ}^{-1}\Lambda_{ZX}$$

Proof:

$$\Lambda_{\tilde{X}|YZ}^{\tilde{X}|YZ} = \Lambda_{XX} - \Lambda_{XW} \Lambda_{WW}^{-1} \Lambda_{WX}$$

$$\left[\Lambda_{XZ} \Lambda_{XY} \right] \left[\Lambda_{ZZ}^{-1} \quad 0 \atop 0 \quad \Lambda_{YY}^{-1} \right] \left[\Lambda_{ZX}^{XX} \right]$$

$$\boldsymbol{\Lambda}_{\tilde{X}_{|YZ}\tilde{X}_{|YZ}} = \underbrace{\boldsymbol{\Lambda}_{XX} - \boldsymbol{\Lambda}_{XY} \boldsymbol{\Lambda}_{YY}^{-1} \boldsymbol{\Lambda}_{YX}}_{\tilde{X}_{|Y}\tilde{X}_{|Y}} - \boldsymbol{\Lambda}_{XZ} \boldsymbol{\Lambda}_{ZZ}^{-1} \boldsymbol{\Lambda}_{ZX}$$

Least Squares Estimation : Property 3

What happens when Z and Y are correlated?

$$\Lambda_{ZY} = E\{\tilde{Z}\tilde{Y}^T\} \neq 0$$

Then,

$$\hat{X}_{|YZ} = \hat{X}_{|Y} + \underbrace{\left(\tilde{X}_{|Y}\right)_{|(\tilde{Z}_{|Y})}}_{}$$

 ${\it This warrants further explanation} \dots$

Recursive LS Estimation

Using Y, we can estimate X and Z by their conditional means:

The conditional mean of X The conditional mean of Z $\widehat{X}_{|_{V}} = \widehat{x} + \Lambda_{XY} \Lambda_{VY}^{-1} (Y - \widehat{y}) \qquad \widehat{Z}_{|_{V}} = \widehat{z} + \Lambda_{ZY} \Lambda_{VY}^{-1} (Y - \widehat{y})$

The corresponding conditional estimation errors are:

$$\tilde{X}_{|_{Y}} = X - \hat{X}_{|_{Y}} \qquad \tilde{Z}_{|_{Y}} = Z - \hat{Z}_{|_{Y}}$$

 $\underline{\text{Uncorrelated}}$ with Y (by Least Squares Property 1)

Recursive LS Estimation

We have:

The conditional mean of X The conditional mean of Z $\hat{X}_{|_Y} = \hat{x} + \Lambda_{XY} \Lambda_{YY}^{-1} (Y - \hat{y}) \qquad \hat{Z}_{|_Y} = \hat{z} + \Lambda_{ZY} \Lambda_{YY}^{-1} (Y - \hat{y})$

If we get the outcomes Y=y and Z=zThe corresponding conditional estimation errors become:

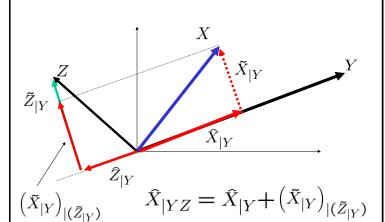
 $\tilde{X}_{|_{y}} = X - \hat{x}_{|_{y}}$

 $\tilde{z}_{|y} = z - \hat{z}_{|y}$

This is still random

This is now an outcome

Deterministic interpretation of Property 3



Least Squares Estimation : Property 3

a) Recursive estimate

$$\widehat{X}_{|YZ} = \widehat{X}_{|Y} + \left(\widetilde{X}_{|Y}\right)_{|(\widetilde{Z}_{|Y})}$$

Least Squares Estimation : Property 3

b) Recursive estimation error

$$\Lambda_{\tilde{X}_{|YZ}\tilde{X}_{|YZ}} = \Lambda_{\tilde{X}_{|Y}\tilde{X}_{|Y}} - \Lambda_{\tilde{X}_{|Y}\tilde{Z}_{|Y}} \Lambda_{\tilde{Z}_{|Y}\tilde{Z}_{|Y}}^{-1} \Lambda_{\tilde{Z}_{|Y}\tilde{X}_{|Y}}$$

where:

$$\Lambda_{\tilde{X}_{|Y}\tilde{X}_{|Y}} = \Lambda_{XX} - \Lambda_{XY}\Lambda_{YY}^{-1}\Lambda_{YX}$$

$$\Lambda_{\tilde{X}_{|Y}\tilde{Z}_{|Y}} = \Lambda_{XZ} - \Lambda_{XY}\Lambda_{YY}^{-1}\Lambda_{YZ}$$

$$\Lambda_{\tilde{Z}_{|Y}\tilde{Z}_{|Y}} = \Lambda_{ZZ} - \Lambda_{ZY}\Lambda_{YY}^{-1}\Lambda_{YZ}$$

Derivation of Recursive LS Estimation

- 1) Define the vector $W = \left[egin{array}{c} Z \\ Y \end{array}
 ight] \quad \hat{w} = \left[egin{array}{c} \hat{z} \\ \hat{y} \end{array}
 ight]$
- 2) Compute $\hat{x}_{|yz} = E\{X|Y=y, Z=z\}$

$$\hat{x}_{|yz} = \hat{x} + \bigwedge_{XW} \bigwedge_{WW}^{-1} (w - \hat{w})$$
inversion of an $(p+M) \times (p+M)$ matrix

Solution: use Schur complement

Given

$$\Lambda_{WW} = \begin{bmatrix} \Lambda_{ZZ} & \Lambda_{ZY} \\ \Lambda_{YZ} & \Lambda_{YY} \end{bmatrix} \quad \text{and} \quad \Lambda_{YY}^{-1}$$

• Compute the Schur complement of Λ_{VV}

$$\Delta = \Lambda_{ZZ} - \Lambda_{ZY} \Lambda_{YY}^{-1} \Lambda_{YZ}$$
$$= \Lambda_{\tilde{Z}|Y} \tilde{Z}|Y := \Lambda_{Z|Y}$$

which is the conditional covariance

Solution: use Schur complement of Λ_{YY}

• Given

$$\Lambda_{WW} = \begin{bmatrix} \Lambda_{ZZ} & \Lambda_{ZY} \\ \Lambda_{YZ} & \Lambda_{YY} \end{bmatrix} \quad \Lambda_{Z|Y} = \Lambda_{ZZ} - \Lambda_{ZY} \Lambda_{YY}^{-1} \Lambda_{YZ}$$

Then

$$\Lambda_{WW}^{-1} = \begin{bmatrix} \Lambda_{Z|Y}^{-1} & -\Lambda_{Z|Y}^{-1} F \\ -F^T \Lambda_{Z|Y}^{-1} & \Lambda_{YY}^{-1} + F^T \Lambda_{Z|Y}^{-1} F \end{bmatrix}$$
$$F = \Lambda_{ZY} \Lambda_{YY}^{-1}$$

Non-Recursive LS Estimation

$$\hat{x}_{|yz} = \hat{x} + \bigwedge_{XW} \bigwedge_{WW}^{-1} (w - \hat{w})$$

$$W = \begin{bmatrix} Z \\ Y \end{bmatrix}$$

$$\begin{bmatrix} \bigwedge_{ZZ} & \bigwedge_{ZY} \\ \bigwedge_{YZ} & \bigwedge_{YY} \end{bmatrix}^{-1}$$

$$\begin{bmatrix} \bigwedge_{XZ} & \bigwedge_{XY} \end{bmatrix}$$

Use Schur complement

$$\hat{x}_{|yz} = \hat{x}$$

$$+ \left[\bigwedge_{XZ} \bigwedge_{XY} \right] \underbrace{\left[\bigwedge_{YZ} \bigwedge_{YY} \bigwedge_{YY} \right]^{-1}}_{\downarrow} \left[\begin{bmatrix} \tilde{z} \\ \tilde{y} \end{bmatrix} \right]$$

$$\begin{bmatrix} \bigwedge_{Z|Y}^{-1} & -\bigwedge_{Z|Y}^{-1} F \\ -F^{T} \bigwedge_{Z|Y}^{-1} & \bigwedge_{YY}^{-1} + F^{T} \bigwedge_{Z|Y}^{-1} F \end{bmatrix}$$

$$\bigwedge_{Z|Y} = \bigwedge_{ZZ} - \bigwedge_{ZY} \bigwedge_{YY}^{-1} \bigwedge_{YZ} \qquad F = \bigwedge_{ZY} \bigwedge_{YY}^{-1}$$

Use Schur complement

$$\begin{split} \hat{x}_{|yz} &= \hat{x} + \Lambda_{XY} \Lambda_{YY}^{-1} \tilde{y} \\ &+ (\Lambda_{XZ} - \Lambda_{XY} \Lambda_{YY}^{-1} \Lambda_{YZ}) \Lambda_{Z|Y}^{-1} (\tilde{z} - \Lambda_{ZY} \Lambda_{YY}^{-1} \tilde{y}) \end{split}$$

$$\boldsymbol{\Lambda}_{\boldsymbol{Z}|\boldsymbol{Y}} = \boldsymbol{\Lambda}_{\boldsymbol{Z}\boldsymbol{Z}} - \boldsymbol{\Lambda}_{\boldsymbol{Z}\boldsymbol{Y}}\boldsymbol{\Lambda}_{\boldsymbol{Y}\boldsymbol{Y}}^{-1}\boldsymbol{\Lambda}_{\boldsymbol{Y}\boldsymbol{Z}}$$

Use Schur complement

$$\widehat{x}_{|yz} = \underbrace{\widehat{x} + \bigwedge_{XY} \bigwedge_{YY}^{-1} \widetilde{y}}_{\widehat{x}_{|y} \leftarrow \text{expected value of } X \text{given outcome } y} \\ + (\bigwedge_{XZ} - \bigwedge_{XY} \bigwedge_{YY}^{-1} \bigwedge_{YZ}) \bigwedge_{Z|Y}^{-1} (\widetilde{z} - \bigwedge_{ZY} \bigwedge_{YY}^{-1} \widetilde{y})$$

Use Schur complement

We will now show that

$$\begin{split} \widehat{x}_{|yz} &= \widehat{x}_{|y} \\ &+ \underbrace{(\Lambda_{XZ} - \Lambda_{XY} \Lambda_{YY}^{-1} \Lambda_{YZ}) \Lambda_{Z|Y}^{-1} (\widetilde{z} - \Lambda_{ZY} \Lambda_{YY}^{-1} \widetilde{y})}_{E\{\widetilde{X}_{|Y} | \widetilde{z}_{|y}\}} \end{split}$$

The expected value of $~\tilde{X}_{|y}~$ given the outcome $~\tilde{z}_{|y}$

Computation of $\tilde{z}_{|y}$

The conditional mean of Z given Y = y:

$$\hat{z}_{|y} = \hat{z} + \Lambda_{ZY} \Lambda_{YY}^{-1} \tilde{y}$$

$$\tilde{z}_{|y} = z - \hat{z}_{|y}$$

$$\tilde{z}_{|y} = \underbrace{z - \hat{z}}_{\tilde{z}} - \Lambda_{ZY} \Lambda_{YY}^{-1} \tilde{y}$$

Computation of $E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\}$

Therefore, $\tilde{z}_{|_{\mathcal{Y}}} = \tilde{z} + \Lambda_{_{ZY}} \Lambda_{_{YY}}^{-1} \, \tilde{y}$

We will now compute $\,E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\}\, \text{using the LS result:}$

$$E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\} = \ E\{\tilde{X}_{|Y}\} + \ E\{\tilde{X}_{|Y}\tilde{Z}_{|Y}^T\}E\{\tilde{Z}_{|Y}\tilde{Z}_{|Y}^T\}^{-1}\,\tilde{z}_{|y}$$

to verify that

$$E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\} = (\Lambda_{XZ} - \Lambda_{XY}\Lambda_{YY}^{-1}\Lambda_{YZ})\Lambda_{Z|Y}^{-1}\underbrace{(\tilde{z} - \Lambda_{ZY}\Lambda_{YY}^{-1}\tilde{y})}_{\tilde{z}_{|y}}$$

Computation of $E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\}$

Using Gaussian least squares results:

$$\begin{split} E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\} &= \underbrace{E\{\tilde{X}_{|Y}\}}^{\pmb{0}} \\ &+ E\{\tilde{X}_{|Y}\tilde{Z}_{|Y}^T\}E\{\tilde{Z}_{|Y}\tilde{Z}_{|Y}^T\}^{-1}\tilde{z}_{|y} \end{split}$$

Estimation errors always have zero means

Computation of $E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\}$

Using Gaussian least squares results:

$$\begin{split} E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\} &= E\{\tilde{X}_{|Y}\tilde{Z}_{|Y}^T\}\underbrace{E\{\tilde{Z}_{|Y}\tilde{Z}_{|Y}^T\}}^{-1}\tilde{z}_{|y} \\ E\{\tilde{Z}_{|Y}\tilde{Z}_{|Y}^T\} &= \Lambda_{\tilde{Z}_{|Y}}\tilde{Z}_{|Y} = \Lambda_{Z|Y} \\ &= \Lambda_{ZZ} - \Lambda_{ZY}\Lambda_{YY}^{-1}\Lambda_{YZ} \end{split}$$

the conditional covariance

Computation of $E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\}$

Using Gaussian least squares results:

$$E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\} = E\{\tilde{X}_{|Y}\tilde{Z}_{|Y}^T\} \wedge_{Z|Y}^{-1} \tilde{z}_{|y}$$

Notice that, from the Schur complements result,

$$E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\} \; = (\bigwedge_{XZ} - \bigwedge_{XY} \bigwedge_{YY}^{-1} \bigwedge_{YZ}) \bigwedge_{Z|Y}^{-1} \tilde{z}_{|y}$$

Computation of $E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\}$

Using Gaussian least squares results:

$$\begin{split} E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\} &= \underbrace{E\{\tilde{X}_{|Y}\tilde{Z}_{|Y}^T\}}_{K|Y} \Lambda_{Z|Y}^{-1} \tilde{z}_{|y} \\ &= \underbrace{E\{(\tilde{X} - \Lambda_{XY}\Lambda_{YY}^{-1}\tilde{Y})\tilde{Z}_{|Y}^T\}}_{E\{\tilde{X}\tilde{Z}_{|Y}^T\}} + \underbrace{\Lambda_{XY}\Lambda_{YY}^{-1}E\{\tilde{Y}\tilde{Z}_{|Y}^T\}}_{E\{\tilde{Y}\tilde{Z}_{|Y}^T\}} \end{split}$$

Computation of $E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\}$

Using Gaussian least squares results:

$$E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\} = \underbrace{E\{\tilde{X}\tilde{Z}_{|Y}^T\}}_{Z|Y} \Lambda_{Z|Y}^{-1} \tilde{z}_{|y}$$

$$E\{\tilde{X}\tilde{Z}_{|Y}^T\} = E\{\tilde{X}(\tilde{Z} - \Lambda_{ZY}\Lambda_{YY}^{-1}\tilde{Y})^T\}$$

$$= E\{\tilde{X}\tilde{Z}^T\} - E\{\tilde{X}\tilde{Y}^T\}\Lambda_{YY}^{-1}\Lambda_{YZ}$$

$$= \Lambda_{XZ} - \Lambda_{XY}\Lambda_{YY}^{-1}\Lambda_{YZ}$$

Computation of $E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\}$

Therefore,

$$E\{\tilde{X}_{|Y}\tilde{Z}_{|Y}^T\} = \Lambda_{XZ} - \Lambda_{XY}\Lambda_{YY}^{-1}\Lambda_{YZ}$$

and

$$E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\} = (\Lambda_{XZ} - \Lambda_{XY}\Lambda_{YY}^{-1}\Lambda_{YZ})\Lambda_{Z|Y}^{-1}\tilde{z}_{|y}$$

Non-Recursive LS Estimation Error

$$\Lambda_{\widetilde{X}|W}\widetilde{X}|W} = \Lambda_{XX} - \Lambda_{XW}\Lambda_{WW}^{-1}\Lambda_{WX}$$

$$W = \begin{bmatrix} Z \\ Y \end{bmatrix}$$

$$\begin{bmatrix} \Lambda_{ZZ} & \Lambda_{ZY} \\ \Lambda_{YZ} & \Lambda_{YY} \end{bmatrix}^{-1}$$

$$[\Lambda_{XZ} & \Lambda_{XY}]$$

Use Schur complement

$$\Lambda_{\tilde{X}|YZ}\tilde{X}|YZ} = \Lambda_{XX} \\
- \left[\Lambda_{XZ} \quad \Lambda_{XY}\right] \left[\begin{array}{ccc} \Lambda_{ZZ} & \Lambda_{ZY} \\ \Lambda_{YZ} & \Lambda_{YY} \end{array}\right]^{-1} \left[\begin{array}{ccc} \Lambda_{ZX} \\ \Lambda_{ZY} \end{array}\right] \\
\left[\begin{array}{ccc} \Lambda_{Z|Y}^{-1} & -\Lambda_{Z|Y}^{-1}F \\ -F^{T}\Lambda_{Z|Y}^{-1} & \Lambda_{YY}^{-1} + F^{T}\Lambda_{Z|Y}^{-1}F \end{array}\right]$$

Use Schur complement

$$\Lambda_{\tilde{X}|YZ}^{\tilde{X}|YZ} = \Lambda_{XX} - \Lambda_{XY}\Lambda_{YY}^{-1}\Lambda_{YX}$$
$$-(\Lambda_{XZ} - \Lambda_{XY}\Lambda_{YY}^{-1}\Lambda_{YZ})\Lambda_{Z|Y}^{-1}(\Lambda_{ZX} - \Lambda_{ZY}\Lambda_{YY}^{-1}\Lambda_{YX})$$

$$\Lambda_{Z|Y} = \Lambda_{ZZ} - \Lambda_{ZY} \Lambda_{YY}^{-1} \Lambda_{YZ}$$

Summary

The conditional mean is the least squares estimator:

$$E\{||X - \hat{X}|_Y||^2\} \le E\{||X - f(Y)||^2\}$$

For Gaussians, the conditional mean is an affine function

$$\widehat{x}|_{y} = \widehat{x} + \Lambda_{XY} \Lambda_{YY}^{-1} (y - \widehat{y})$$

Summary

The conditional mean can be computed recursively:

1. If we first know of outcome Y = y

$$\hat{x}_{|_{y}} = \hat{x} + \bigwedge_{XY} \bigwedge_{YY}^{-1} \tilde{y}$$

Summary

The conditional mean can be computed recursively:

2 If we afterwards know of outcome Z = z

$$\begin{split} \hat{z}_{|y} &= \hat{z} + \bigwedge_{ZY} \bigwedge_{YY}^{-1} \tilde{y} \\ \tilde{z}_{|y} &= z - \hat{z}_{|y} \end{split}$$

then

$$\hat{x}_{|yz} = \quad \hat{x}_{|y} \quad + \ E\{\tilde{X}_{|Y}|\tilde{z}_{|y}\}$$

Course Outline

· Unit 0: Probability

Finished

- Unit 1: State-space control, estimation
- Unit 2: Input/output control
- · Unit 3: Adaptive control