

This discussion was released **Friday, November 6**.

This discussion serves as an introduction to the EM algorithm. We will first start with this [Jupyter notebook](#). We will then come back to this worksheet for a more theoretical understanding of EM algorithm.

As a reminder, if you have questions, we will answer them via the queue at oh.eecs189.org. Once you complete the Jupyter notebook, please return to this worksheet.

1 Jupyter Notebook

2 One Dimensional Mixture of Two Gaussians

Suppose we have a mixtures of two Gaussians in \mathbb{R} that can be described by a pair of random variables (X, Z) where X takes values in \mathbb{R} and Z takes value in the set $1, 2$. The joint-distribution of the pair (X, Z) is given to us as follows:

$$\begin{aligned} Z &\sim \text{Bernoulli}(0.5), \\ (X|Z = k) &\sim \mathcal{N}(\mu_k, \sigma_k), \quad k \in 1, 2, \end{aligned}$$

We use θ to denote the set of all parameters $\mu_1, \sigma_1, \mu_2, \sigma_2$.

- (a) Write down the expression for the joint likelihood $p_\theta(X = x_i, Z_i = 1)$ and $p_\theta(X = x_i, Z_i = 2)$. What is the marginal likelihood $p_\theta(X = x_i)$?
- (b) It turns out that we can compute the log-likelihood easily.

$$\begin{aligned} \ell_\theta(\mathbf{x}) &= \ln(p_\theta(X = x_1, \dots, X = x_n)) \\ &= \sum_{i=1}^n \ln(p_\theta(X = x_i)) \\ &= \sum_{i=1}^n \ln \left[\frac{1}{2} \mathcal{N}(x_i | \mu_1, \sigma_1^2) + \frac{1}{2} \mathcal{N}(x_i | \mu_2, \sigma_2^2) \right] \end{aligned}$$

Here, we use $\mathcal{N}(x|\mu, \sigma^2)$ as shorthand for the Gaussian density evaluated at x for a Normal random variable with mean μ and variance σ^2 .

This log-likelihood can be optimized, but not analytically. Taking the derivative with respect to μ_1 , for example, would give:

$$\frac{\partial \ell_{\theta}(\mathbf{x})}{\partial \mu_1} = \sum_{i=1}^n \frac{\mathcal{N}(x_i|\mu_1, \sigma_1^2)}{\mathcal{N}(x_i|\mu_1, \sigma_1^2) + \mathcal{N}(x_i|\mu_2, \sigma_2^2)} \left(\frac{x_i - \mu_1}{\sigma_1^2} \right)$$

This ratio of exponentials and linear terms makes it difficult to analytically solve for the maximum likelihood estimate.

Anyway, we still want to optimize the log likelihood: $\ell_{\theta}(x)$. However, we just saw this can be hard to solve for an MLE closed form solution. **Show that a lower bound for a single term in the log likelihood is $\ell_{\theta}(x_i) \geq \mathbb{E}_q \left[\ln \left(\frac{p_{\theta}(X=x_i, Z_i=k)}{q_{\theta}(Z_i=k|X=x_i)} \right) \right]$.** Here, this bound should hold for any distribution $q_{\theta}(Z_i = k|X = x_i)$. The expectation in the expression above to the right is using q to treat k as a random variable.

Here, you should start with:

$$\ell_{\theta}(x_i) = \ln \left(\sum_k p_{\theta}(X = x_i, Z_i = k) \right)$$

and go from there. You don't have to worry about the details of Gaussians for this problem.

(Hint: At a high level, look at what you are trying to prove. There are three things you clearly need to do just by looking at the patterns: (1) Somehow introduce the distribution q into the problem; (2) Somehow turn the sum over k into an expectation; (3) Somehow get that expectation/sum outside the logarithm.

We will stop here due to time limit and continue the theoretical setup in your homework.

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