

## 1 MLE vs. MAP

Let  $D$  denote the observed data and  $\theta$  the parameter. While MLE only maximizes a likelihood distribution  $p(D|\theta)$ , MAP takes a more Bayesian approach. MAP assumes that the parameter  $\theta$  *is also a random variable and has its own distribution*. Recall that using Bayes' rule, the posterior distribution can be seen as the product of likelihood and prior:

$$p(\theta|D) = \frac{p(D|\theta)p(\theta)}{p(D)} \propto \underbrace{p(D|\theta)}_{\text{likelihood}} \underbrace{p(\theta)}_{\text{prior}}$$

Suppose that the data consists of  $n$  i.i.d. observations  $D = \{x_1, \dots, x_n\}$ . MAP tries to infer the parameter by maximizing the posterior distribution:

$$\begin{aligned}\hat{\theta}_{\text{MAP}} &= \arg \max_{\theta} p(\theta|D) \\ &= \arg \max_{\theta} p(D|\theta)p(\theta) \\ &= \arg \max_{\theta} \left[ \prod_{i=1}^n p(x_i|\theta) \right] p(\theta) \\ &= \arg \max_{\theta} \left( \sum_{i=1}^n \log p(x_i|\theta) \right) + \log p(\theta)\end{aligned}$$

Note that since both of these methods are point estimates (they yield a value rather than a distribution), neither of them are completely Bayesian. A faithful Bayesian would use a model that yields a posterior distribution over all possible values of  $\theta$ , but this is often intractable or very computationally expensive.

Now suppose we have a coin with unknown bias  $\theta$ . We are trying to find the bias of the coin by maximizing the underlying distribution. You tossed the coin  $n = 10$  times and 3 of the tosses came as heads.

(a) **What is the MLE of the bias of the coin  $\hat{\theta}_{\text{MLE}}$ ?**

(b) Suppose we know that the bias of the coin is distributed according to  $\theta \sim N(0.8, 0.09)$ , i.e., we are rather sure that the bias should be around 0.8.<sup>1</sup>

**What is the MAP estimate of the coin bias  $\hat{\theta}_{\text{MAP}}$ ?** You can leave your result as a polynomial equation on  $\theta$ .

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<sup>1</sup>This is a somewhat strange choice of prior, since we know that  $0 \leq \theta \leq 1$ . However, we will stick with this example for illustrative purposes.

(c) What if our prior is  $\theta \sim N(0.5, 0.09)$  or  $N(0.8, 1)$  instead?

**How does the difference between the new MAP estimates and MLE estimate change and why?**

(d) **What if our prior is that  $\theta$  is uniformly distributed in the range  $(0, 1)$ ?**

## 2 Probabilistic Interpretation of Lasso

Let's start with the probabilistic interpretation of least squares. We're given labels  $y \in \mathbb{R}$ , data  $\mathbf{x} \in \mathbb{R}^d$ , and Gaussian noise  $z \sim \mathcal{N}(0, \sigma^2)$ , where  $y = \mathbf{w}^T \mathbf{x} + z$ . Recall from lecture and the previous discussion that this results in a probabilistic linear model given by:

$$p(y|\mathbf{x}; \mathbf{w}) \sim \mathcal{N}(\mathbf{w}^T \mathbf{x}, \sigma^2)$$

However, maximum likelihood estimates (MLE) can overfit to the training data (analogous to how fitting a very high dimensional polynomial to data leads to large coefficients and extreme behavior at unseen points). To ameliorate this issue, we can assume a zero-mean Laplace prior on each component of the parameter  $w_j \sim \text{Laplace}(0, t)$ :

$$p(w_j) = \frac{1}{2t} \exp\left\{-\frac{1}{t}|w_j|\right\}$$
$$p(\mathbf{w}) = \prod_{j=1}^d p(w_j) = \left(\frac{1}{2t}\right)^d \cdot \exp\left\{-\frac{1}{t} \sum_{j=1}^d |w_j|\right\}$$

Assume that  $t$  is a known constant. Here, we will see that this modification results in a new objective called Lasso regression.

(a) Recall that the MLE objective finds the parameters that maximize the likelihood of the data,

$$\begin{aligned} \mathbf{w}^* &= \arg \max_{\mathbf{w}} L(\mathbf{w}) \\ &= \arg \max_{\mathbf{w}} p(Y_1, \dots, Y_n, |\mathbf{w}, \mathbf{X}_1, \dots, \mathbf{X}_n, \sigma^2) \\ &= \arg \max_{\mathbf{w}} \prod_{i=1}^n p(Y_i | \mathbf{X}_i, \mathbf{w}, \sigma^2). \end{aligned}$$

When working in a Bayesian framework, we instead focus on the posterior probability of the parameters (the unknown quantity) conditioned the data (the evidence):

$$\text{Posterior} = p(\text{unknowns} | \text{evidence}) = p(\mathbf{w} | Y_1, \dots, Y_n, \mathbf{X}_1, \dots, \mathbf{X}_n, \sigma^2)$$

**Derive the MAP objective as a function of the log-likelihood  $\ell(\mathbf{w})$  and the prior  $p(\mathbf{w})$ .**

- (b) **Fill in the terms of the MAP objective you derived, assuming Gaussian noise and a Laplace prior on the parameter.**

- (c) **Using your answer from the previous part, show that maximizing the MAP objective is equivalent to minimizing the following:**

$$J(\mathbf{w}) = \sum_{i=1}^n (Y_i - \mathbf{w}^T \mathbf{X}_i)^2 + \lambda \|\mathbf{w}\|_1$$

**What is the constant  $\lambda$  in terms of given quantities?**

### 3 Independence and Multivariate Gaussians

To review, a covariance matrix  $\Sigma \in \mathbb{R}^{N \times N}$  for a random variable  $X \in \mathbb{R}^N$  with the following values, where  $\text{cov}(X_i, X_j) = \mathbb{E}[(X_i - \mu_i)(X_j - \mu_j)]$  is the covariance between the  $i$ -th and  $j$ -th elements of the random vector  $X$ :

$$\Sigma = \begin{bmatrix} \text{cov}(X_1, X_1) & \dots & \text{cov}(X_1, X_n) \\ \dots & & \dots \\ \text{cov}(X_n, X_1) & \dots & \text{cov}(X_n, X_n) \end{bmatrix} = \mathbb{E}[(X - \mu)(X - \mu)^\top]. \quad (1)$$

Recall that the density of an  $N$  dimensional Multivariate Gaussian Distribution  $\mathcal{N}(\mu, \Sigma)$  is defined as follows when  $\Sigma$  is positive definite:

$$f(x) = \frac{1}{\sqrt{(2\pi)^N |\Sigma|}} \exp\left\{-\frac{1}{2}(x - \mu)^\top \Sigma^{-1}(x - \mu)\right\}. \quad (2)$$

Here,  $|\Sigma|$  denotes the determinant of the matrix  $\Sigma$ .

(a) For  $X = [X_1, \dots, X_n]^\top \sim \mathcal{N}(\mu, \Sigma)$ , **verify that if  $X_i, X_j$  are independent (for all  $i \neq j$ ), then  $\Sigma$  must be diagonal, that is,  $X_i, X_j$  are uncorrelated.**

(b) Let  $N = 2$ ,  $\mu = \begin{pmatrix} 0 \\ 0 \end{pmatrix}$ , and  $\Sigma = \begin{pmatrix} \alpha & \beta \\ \beta & \gamma \end{pmatrix}$ . Suppose  $X = \begin{pmatrix} X_1 \\ X_2 \end{pmatrix} \sim \mathcal{N}(\mu, \Sigma)$ . **Show that  $X_1, X_2$  are independent if  $\beta = 0$ .** Recall that two continuous random variables  $W, Y$  with joint density  $f_{W,Y}$  and marginal densities  $f_W, f_Y$  are independent if  $f_{W,Y}(w, y) = f_W(w)f_Y(y)$ .

- (c) Consider a data point  $x$  drawn from a  $N$ -dimensional zero mean Multivariate Gaussian distribution  $\mathcal{N}(0, \Sigma)$ , as shown above. Assume that  $\Sigma^{-1}$  exists. **Prove that there exists matrix  $A \in \mathbb{R}^{N,N}$  such that  $x^\top \Sigma^{-1} x = \|Ax\|_2^2$  for all vectors  $x$ . What is the matrix  $A$ ?**
- (d) Let's constrain  $x$  to be on the unit sphere. In other words, the  $\ell_2$  norm (or magnitude) of vector  $x$  is 1 ( $\|x\|_2 = 1$ ). In this case, **what are the maximum and minimum values of  $\|Ax\|_2^2$ ? In other words,  $\max_{x:\|x\|_2=1} \|Ax\|_2^2$  and  $\min_{x:\|x\|_2=1} \|Ax\|_2^2$ ?**
- (e) If we had  $X_i \perp\!\!\!\perp X_j \forall i, j$  ( $\perp\!\!\!\perp$  denotes independence), **what is the intuitive meaning for the maximum and minimum values of  $\|Ax\|_2^2$ ?** Suppose you wanted to choose an  $x$  on the unit sphere to maximize the density function  $f(x)$  in Eq (2); **what  $x$  should you choose?**