Lecture 7 Linear Regression and Extensions

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Extending Linear Regressions (LR)

General setup

$$y_i^* = x_i'\beta + u_i, \quad u_i|x_i \sim_{i.i.d.} t_{\nu}(0, \sigma^2)$$

$$\mathbb{E}(y_i) = G(x_i'\beta), \quad i = 1, \dots, n$$

- ▶ Choice of link function $G(\cdot)$
 - standard LR

$$G(x_i'\beta) = x_i'\beta \quad \Rightarrow \quad y_i = y_i^*$$

▶ tobit censored LR ($\nu = \infty$; $\mathbb{N}(0,1)$ p.d.f. ϕ , c.d.f. Φ)

$$G(x_i'\beta) = x_i'\beta + \frac{\phi(-x_i'\beta/\sigma)}{1 - \Phi(-x_i'\beta/\sigma)}\sigma \quad \Rightarrow \quad y_i = \max\{y_i^*, 0\}$$

▶ binary probit LR ($\nu = \infty$; $\sigma^2 = 1$)

$$G(x_i'\beta) = \Phi(x_i'\beta) \quad \Rightarrow \quad y_i = 1\{y_i^* > 0\}$$

The Road Ahead...

- ► LR with Gaussian errors
- ► LR with Student-*t* errors
- Marginal likelihood
- ► Tobit censored LR
- ▶ Binary probit LR

LR with Gaussian Errors

Conditionally conjugate prior

$$\beta \sim \mathbb{N}(\beta_0, B_0), \qquad \sigma^2 \sim \mathbb{IG}\text{-}2(\alpha_0/2, \delta_0/2)$$

- Gibbs algorithm
 - ▶ step 1: choose $\beta = \beta^{(0)}$, $\sigma^2 = \sigma^{2(0)}$, set g = 0
 - step 2: sample recursively

$$\beta^{(g)} \sim \mathbb{N}(\beta_1^{(g+1)}, B_1^{(g+1)}), \qquad \sigma^{2(g+1)} \sim \mathbb{IG}\text{-}2(\alpha_1/2, \delta_1^{(g+1)}/2)$$

where

$$\begin{array}{lcl} B_1^{(g+1)} & = & (\sigma^{-2(g)}X'X + B_0^{-1})^{-1} \\ \beta_1^{(g+1)} & = & B_1^{(g+1)}(\sigma^{-2(g)}X'y + B_0^{-1}\beta_0) \\ \alpha_1 & = & \alpha_0 + n \\ \delta_1^{(g+1)} & = & \delta_0 + (y - X\beta^{(g+1)})'(y - X\beta^{(g+1)}) \end{array}$$

▶ step 3: set g = g + 1 and go to step 2

LR with Student-t Errors

Conditional likelihood

$$f(y_i|\beta,\sigma^2,\lambda_i)=\mathbb{N}(x_i'\beta,\lambda_i^{-1}\sigma^2), \quad \lambda_i\sim \mathbb{G}(\nu/2,\nu/2)$$
 (latent)

▶ Gibbs sampler for $\pi(\beta, \sigma^2, \lambda|y)$

$$\beta|y,\lambda,\sigma^{2} \sim \mathbb{N}(\beta_{1},B_{1})$$

$$\sigma^{2}|y,\beta,\lambda \sim \mathbb{IG}\text{-}2(\alpha_{1}/2,\delta_{1}/2)$$

$$\lambda_{i}|y,\beta,\sigma^{2} \sim \mathbb{G}(\nu_{1}/2,\nu_{2i}/2), \qquad i=1,\ldots,n$$

where $\Lambda = \operatorname{diag}(\lambda_i)$ and

$$B_{1} = (\sigma^{-2}X'\Lambda X + B_{0}^{-1})^{-1}$$

$$\beta_{1} = B_{1}(\sigma^{-2}X'\Lambda y + B_{0}^{-1}\beta_{0})$$

$$\alpha_{1} = \alpha_{0} + n$$

$$\delta_{1} = \delta_{0} + (y - X\beta)'\Lambda(y - X\beta)$$

$$\nu_{1} = \nu + 1$$

$$\nu_{2i} = \nu + \sigma^{-2}(y_{i} - x'_{i}\beta)^{2}$$

Python Pseudo Code

To be added...

Marginal Likelihood

Chib method

$$m(y) = \frac{\prod_{i=1}^{n} t_{\nu}(x_{i}'\beta^{*}, \sigma^{2*})\pi(\beta^{*})\pi(\sigma^{2*})}{\pi(\beta^{*}, \sigma^{2*}|y)}, \quad \forall \theta^{*} = (\beta^{*}, \sigma^{2*}) \in \Theta$$

▶ Compute $\pi(\beta^*, \sigma^{2*}|y)$ (not involving λ) at high-density point θ^* from Gibbs output

$$\pi(\beta^*, \sigma^{2*}|y) = \pi(\beta^*|y)\pi(\sigma^{2*}|\beta^*, y)$$

- full run: $\hat{\pi}(\beta^*|y) = \frac{1}{G} \sum_{g=1}^G \pi(\beta^*|\sigma^{2(g)}, \lambda^{(g)}, y)$, where $(\theta^{(g)}, \lambda^{(g)}) \sim \pi(\theta, \lambda|y)$
- ▶ reduced run: $\hat{\pi}(\sigma^{2*}|\beta^*,y) = \frac{1}{G}\sum_{g=1}^{G}\pi(\sigma^{2*}|\beta^*,\lambda^{(g)},y)$, where $(\sigma^{2(g)},\lambda^{(g)}) \sim \pi(\sigma^2,\lambda|\beta^*,y)$

Tobit Censored LR

Model

$$y_i^* = x_i'\beta + u_i, u_i|x_i \sim_{i.i.d.} \mathbb{N}(0, \sigma^2)$$

 $y_i = \max\{y_i^*, 0\}, i = 1, ..., n$

- ► Chib (1992) introduces latent variables z for censored observations and Gibbs sampler for $\pi(\beta, \sigma^2, z|y)$
 - conditionally conjugate prior for (β, σ^2) as before
 - ▶ sample $\beta | y_z, \sigma^2 \sim \mathbb{N}(\beta_1, B_1)$, where y_z replaces $y_i = 0$ by $z_i < 0$
 - sample $\sigma^2 | y_z, \beta \sim \mathbb{IG}\text{-}2(\alpha_1/2, \delta_1/2)$
 - ▶ sample $z_i|y, \beta, \sigma^2 \sim \mathbb{TN}_{(-\infty,0)}(x_i'\beta, \sigma^2)$ (truncated normal)
 - exercise: Student-t version
- Data augmentation technique [Tanner & Wong (1987)]

Python Pseudo Code

To be added...

Binary Probit LR

Model

$$y_i^* = x_i'\beta + u_i, u_i|x_i \sim_{i.i.d.} \mathbb{N}(0,1)$$

 $y_i = 1\{y_i^* > 0\}, i = 1, \dots, n$

- ▶ Albert and Chib (1993) introduces latent variables $z = y^*$ and Gibbs sampler for $\pi(\beta, z|y)$
 - \triangleright $\beta \sim \mathbb{N}(\beta_0, B_0)$ as before; $\sigma^2 = 1$ for identification
 - ▶ sample $\beta | z \sim \mathbb{N}(\beta_1, B_1)$ (B_1 not updated)
 - ▶ sample $z_i|y,\beta \sim \mathbb{TN}_{(-\infty,0]}(x_i'\beta,1)$ if $y_i=0$ or $\mathbb{TN}_{(0,\infty)}(x_i'\beta,1)$ if $y_i=1$
 - exercise: Student-t version
- ▶ Binary logit LR: $u_i|x_i \sim_{i.i.d.} \mathbb{L}(0,1)$ (logistic distribution)

Python Pseudo Code

To be added...

Readings

- Albert & Chib (1993), "Bayesian Analysis of Binary and Polychotomous Response Data," *Journal of the American* Statistical Association
- Chib (1992), "Bayes Inference in the Tobit Censored Regression Model," *Journal of Econometrics*
- Tanner & Wong (1987), "The Calculation of Posterior Distributions by Data Augmentation (with Discussion)," Journal of the American Statistical Association