Package 'gre'

Version 1.0.0

Title High-dimensional general relative error

Description This package is mainly used to estimate the parameters in the high-dimensional linear multiplicative (or accelerated failure time - AFT) model. It utilizes the general relative error (GRE) criteria to build the objective function. The GRE criteria has two exact forms in this package: the least absolute relative errors (LARE) and the least product relative errors (LPRE). Both the two forms have satisfactory performance.

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❖ auccalc

Description

Function auccalc aims to calculate the area under the curve (AUC).

Usage

```
auc = auccalc(FPR,TPR)
```

Input

FPR: false positive rate set

TPR: true positive rate set

Output

auc: the area under the curve (AUC)

Examples

FPRcal

Description

Function FPRcal aims to calculate the false positive rate of the estimated beta

Usage

```
value = FPRcal(betahat,betaori)
```

Input

betahat: estimated beta (p dimensional vector)

betaori: original beta (p dimensional vector)

Output

value: the false positive rate

Examples

```
% This Demo shows how to use functions FPRcal to compute the false positive rate. betaori=[1;2;3;0;0;0;0;0;0]; betahat=[0.9;2.1;3.2;0.6;0.2;0.1;0;0;0]; result=FPRcal(betahat,betaori);
```

❖ KM

Description

Function KM creates the Kaplan-Meier weight for right censoring data.

Usage

```
weight=KM(y,Delta)
```

Input

y: n dimensional right censoring observation.

Delta: n dimensional event indicator, which has the same length with y. The element of Delta equals 1 if the event was observed, or equals 0 if the responce was censored

Output

weight: n dimensional Kaplan-Meier vector.

```
% This Demo shows how to use function KM to estimate the
weight when the data is right censoring.
clear
n = 200;
p1 = 500;
p2 = 5;
w = randn(n,p1);
x1 = w;
x2 = randn(n,p2);
x3 = zeros(n,p1*p2);
for i = 1:n
    x3(i,:) = reshape(x2(i,:)'*x1(i,:),p1*p2,1);
end
x = [x1, x2, x3];
[\sim,p] = size(x);
beta01 = zeros(p1,1);
beta02 = zeros(p2,1);
beta03 = zeros(p1*p2,1);
beta01(1:10) = 0.4+0.8*rand(10,1);
beta02(1:5) = 4+0.8*rand(5,1);
beta03(1:20) = 4+0.8*rand(20,1);
beta = [beta01;beta02;beta03];
YY = \exp(x*beta).*\exp(randn(n,1));
C = 100 * rand(n, 1);
y = zeros(n, 1);
Delta = zeros(n,1);
for i = 1:n
    y(i) = min(C(i), YY(i));
    if YY(i) \ll C(i)
        Delta(i) = 1;
    end
end
weight = KM(y, Delta);
```

❖ lare

Description

Function lare estimates the coefficients of the high dimensional AFT model with the least add relative errors (LARE) method. We use the MCP penalty in this function.

Usage

betahat = lare(x,y,lam,rr,weight)

Input

```
x: n x p covariates
y: n dimensional response vector
lam, rr: tuning parameters
weight: the weight allocated to the samples (n dimensional vector).
Default as 1/n. n is the number of samples.
```

Output

betahat: estimated coefficient (p dimensional vector)

```
% This Demo shows how to use functions lare to estimate the
coefficients of the high dimensional AFT model. We also use
the 5-fold cross validation method to select the best
lambda.
tic,
clear
n = 100;
p = 200;
maxit = 20;
maxit1 = 20;
toler = 1E-4;
nLambda = 20;
rr = 6;
lambdaRatio = 1E-5;
w = randn(n,p);
wtest = randn(n,p);
x = w;
beta = zeros(p, 1);
beta (1:10) = 0.4+0.8* rand (10,1);
y = \exp(x*beta).*\exp(randn(n,1));
beta0 = zeros(p,1);
for i = 1:100
    lam = i;
    betahat = lare(x, y, lam, 6);
    if sum(abs(betahat) <= 0.001) == p
        break;
    end
end
lambdaMax = i;
lambdaMin = lambdaMax * lambdaRatio;
```

```
loghi = log(lambdaMax);
loglo = log(lambdaMin);
logrange = loghi - loglo;
interval = -logrange/(nLambda-1);
lambda = exp(loghi:interval:loglo)';
betaset = zeros(p,nLambda);
bicset = zeros(nLambda,1);
for i = 1:nLambda
    for ttt = 1:5
        xtest = x(floor((ttt-1)*n/5)+1:floor(ttt*n/5),:);
        ytest = y(floor((ttt-1)*n/5)+1:floor(ttt*n/5),:);
        xtrain = x(setdiff(1:n,floor((ttt-
1) *n/5) +1:floor(ttt*n/5)),:);
        ytrain = y(setdiff(1:n,floor((ttt-
1) *n/5) +1: floor (ttt*n/5)),:);
     beta0 = lare(xtrain, ytrain, lambda(i), 6);
        bicset(i,1) = bicset(i,1) + sum((log(ytest) -
xtest*beta0).^2);
    end
end
bicbesti = find(bicset == min(bicset(bicset>0)));
for i = bicbesti
    betahat = lare(x,y,lambda(i),6);
    betaset(:,i) = betahat;
end
finalbeta = betaset(:,bicbesti);
plot(bicset)
toc
```

❖ Ipre

Description

Function lpre estimates the coefficients of the high dimensional AFT model with the least product relative errors (LPRE) method. We use the MCP penalty in this function.

Usage

```
betahat = lpre(x,y,lam,rr,weight)
```

Input

```
x: n x p covariates
```

y: n dimensional response vector

lam, rr: tuning parameters

weight: the weight allocated to the samples (n dimensional vector).

Default as 1/n. n is the number of samples.

Output

betahat: estimated coefficient (p dimensional vector)

```
% This Demo shows how to use functions lpre to select the
important genes, environment factors, and gene-environment
interactions under the AFT model. Note that in the Demo the
response is also right censoring, so we weight samples with
the Kaplan-Meier weight. The best lambda is chosen as
follows: we generate a same structure data without
censoring as the test set, and select the tuning parameter
which makes the sum of squares errors of the test set
smallest. We also can use the 5-fold cross validation to
select the best tuning parameter.
tic,
clear
n = 200;
p1 = 50;
p2 = 5;
maxit = 20;
maxit1 = 20;
toler = 1E-4;
nLambda = 20;
rr = 6;
lambdaRatio = 1E-5;
w = randn(n, p1);
wtest = randn(n,p1);
x1 = w;
x2 = randn(n, p2);
x3 = zeros(n,p1*p2);
x1test = wtest;
x2test = randn(n,p2);
x3test = zeros(n,p1*p2);
for i = 1:n
    x3(i,:) = reshape(x2(i,:)'*x1(i,:),p1*p2,1);
end
for i = 1:n
    x3test(i,:) = reshape(x2(i,:)'*x1(i,:),p1*p2,1);
```

```
end
x = [x1, x2, x3];
xtest = [x1test, x2test, x3test];
[\sim,p] = size(x);
beta01 = zeros(p1,1);
beta02 = zeros(p2,1);
beta03 = zeros(p1*p2,1);
beta01(1:10) = 0.4+0.8*rand(10,1);
beta02(1:5) = 0.4+0.8*rand(5,1);
beta03(1:20) = 0.4+0.8*rand(20,1);
beta = [beta01;beta02;beta03];
YY = \exp(x*beta).*\exp(randn(n,1));
YYtest = \exp(\text{xtest*beta}) \cdot \exp(\text{randn}(n, 1));
% YY = \exp(x*beta).*\exp(4*rand(n,1)-2);
% epsilon = zeros(n,1);
% for i = 1:n
      cc = rand();
9
      if cc<0.9
응
           epsilon(i) = randn();
응
      else
           epsilon(i) = trnd(1);
응
      end
% end
% YY = \exp(x*beta).*\exp(epsilon);
C = 100.*rand(n,1);
y = zeros(n, 1);
Delta = zeros(n,1);
for i = 1:n
    y(i) = min(C(i), YY(i));
    if YY(i) \le C(i)
        Delta(i) = 1;
    end
end
yold = y;
[y, index] = sort(y);
x = x(index,:);
Delta = Delta(index,:);
weight = KM(y, Delta);
r = 6;
beta0 = zeros(p,1);
for i = 1:100
    lam = i;
    betahat = lpre(x, y, lam, 6);
    if sum(abs(betahat) \le 0.001) == p
        break;
    end
end
```

```
lambdaMax = i;
lambdaMin = lambdaMax * lambdaRatio;
loghi = log(lambdaMax);
loglo = log(lambdaMin);
logrange = loghi - loglo;
interval = -logrange/(nLambda-1);
lambda = exp(loghi:interval:loglo)';
betaset = zeros(p,nLambda);
tprset = zeros(nLambda,1);
fprset = zeros(nLambda, 1);
ebicset = zeros(nLambda,1);
for i = 1:nLambda
    beta0 = lpre(x, y, lambda(i), 6);
   betaset(:,i) = beta0;
    tprset(i,1) = TPRcal(beta0,beta);
    fprset(i,1) = FPRcal(beta0,beta);
    ebicset(i,1) = 2*log(sum((log(YYtest)-
xtest*beta0).^2));
end
auc = auccalc(fprset,tprset);
ebicbesti = find(ebicset == min(ebicset(ebicset>0)));
finalbeta = betaset(:,ebicbesti);
plot(ebicset)
toc
```

❖ TPRcal

Description

Function TPRcal aims to calculate the true positive rate of the estimated beta

Usage

```
value = TPRcal(betahat,betaori)
```

Input

betahat: estimated beta (p dimensional vector)

betaori: original beta (p dimensional vector)

Output

value: the true positive rate

```
% This Demo shows how to use functions TPRcal to compute
the true positive rate.
betaori=[1;2;3;0;0;0;0;0];
betahat=[0.9;2.1;3.2;0.6;0.2;0.1;0;0;0];
result=TPRcal(betahat,betaori);
```