# Bertinelli Gabriele Rlab06

June 3, 2024

## 1 RLab06 - Gabriele Bertinelli (2103359)

```
[]: library(tidyverse)
    library(gridExtra)
    library(latex2exp)
    library(emdbook)
    library(bayestestR)
    library(coda)
    library(magrittr)
    library(rstan)

set.seed(2103359)
```

```
[2]: options(mc.cores = parallel::detectCores())
    rstan_options(auto_write = TRUE)
```

## 1) 1.1

```
p1_like <- poisson(n1, lambda_list)
p2_like <- poisson(n2, lambda_list)</pre>
```

```
[4]: # posterior with uniform prior
     p.post_star_unif1 <- p1_like * 1</pre>
     p.post_unif1 <- p.post_star_unif1 / (sum(p.post_star_unif1)*delta.t)</pre>
     my.median = function(mu, post, dt=0.001){
         for (i in seq(1, length(mu))){
             int = sum(post[1:i]*dt)
              if (int>0.4999){med=mu[i]
                           break
              }
         }
         return (med)
     }
     mean_unif1 <- sum(p.post_unif1 * lambda_list)*delta.t</pre>
     median_unif1 <- my.median(lambda_list, p.post_unif1)</pre>
     variance_unif1 <- sum(((mean_unif1-lambda_list)^2)*p.post_unif1)*delta.t</pre>
     lower_unif1 <- ncredint(lambda_list, p.post_unif1, level = 0.95)[['lower']]</pre>
     upper_unif1 <- ncredint(lambda_list, p.post_unif1, level = 0.95)[['upper']]</pre>
     p.post_star_unif2 <- p2_like * 1</pre>
     p.post_unif2 <- p.post_star_unif2 / (sum(p.post_star_unif2)*delta.t)</pre>
     mean_unif2 <- sum(p.post_unif2 * lambda_list)*delta.t</pre>
     median_unif2 <- my.median(lambda_list, p.post_unif2)</pre>
     variance unif2 <- sum(((mean unif2-lambda list)^2)*p.post unif2)*delta.t</pre>
     lower_unif2 <- ncredint(lambda_list, p.post_unif2, level = 0.95)[['lower']]</pre>
     upper_unif2 <- ncredint(lambda_list, p.post_unif2, level = 0.95)[['upper']]</pre>
     result_unif <- data.frame(</pre>
                      mean = c(mean_unif1, mean_unif2),
                      median = c(median_unif1, median_unif2),
                      variance = c(variance_unif1, variance_unif2),
                      1.ci = c(lower_unif1, lower_unif2),
                      u.ci = c(upper_unif1, upper_unif2),
                      row.names = c('1st set', '2nd set')
```

[5]:

```
print(sprintf('Mean value Set 1: %.3f - Mean value Set 2: %.3f', __
     →result_unif$mean[1], result_unif$mean[2]))
     print(sprintf('Median value Set 1: %.3f - Median value Set 2: %.3f',,,
     →result_unif$median[1], result_unif$median[2]))
     cat('\n')
     print(sprintf('SD Set 1: %.4f - SD Set 2: %.4f', result_unif$variance[1],__
     →result_unif$variance[2]))
     cat('\n')
     print(sprintf('0.95 CI Set 1 [%.3f - %.3f]', result_unif$l.ci[1], result_unif$u.

ci[1]))
     print(sprintf('0.95 CI Set 2 [%.3f - %.3f]', result_unif$l.ci[2], result_unif$u.

    ci[2]))
    [1] "Mean value Set 1: 0.615 - Mean value Set 2: 0.704"
    [1] "Median value Set 1: 0.613 - Median value Set 2: 0.702"
    [1] "SD Set 1: 0.0031 - SD Set 2: 0.0025"
    [1] "0.95 CI Set 1 [0.508 - 0.724]"
    [1] "0.95 CI Set 2 [0.607 - 0.802]"
[6]: options(repr.plot.width = 17, repr.plot.height = 8)
     # plot
     unif.post.plot <- ggplot() +</pre>
             geom_line(aes(lambda_list, p.post_unif1, color='Set 1'), lwd=1.3) +
             geom_line(aes(lambda_list, p.post_unif2, color='Set 2'), lwd=1.3) +
             labs(x=(TeX(r'(\lambda)')), y ='Posterior', title='Posterior - Uniform

     ⇔Prior', color='') +
             theme_minimal(base_size=18) +
             scale_color_manual(values = c('Set 1' = 'firebrick3', 'Set 2' = 
     # zoom in plot
     unif.post.plot.zoom <- ggplot() +</pre>
             geom_line(aes(lambda_list, p.post_unif1, color='Set 1'), lwd=1.7) +
             geom_area(aes(x = lambda_list[lambda_list > lower_unif1 & lambda_list < _ u
     →upper_unif1],
                           y = p.post_unif1[lambda_list > lower_unif1 & lambda_list_
     → < upper_unif1],
                           fill='CI Set 1'), alpha=0.3) +
             geom_vline(aes(xintercept=mean_unif1, color='Mean Set 1'), __
      →linetype='dashed', lwd=1.7) +
```

```
geom_line(aes(lambda list, p.post_unif2, color='Set 2'), lwd=1.7) +
        geom_area(aes(x = lambda_list[lambda_list > lower_unif2 & lambda_list <__
 →upper_unif2],
                      y = p.post_unif2[lambda_list > lower_unif2 & lambda_list_
 \rightarrow upper unif2],
                      fill='CI Set 2'), alpha=0.3) +
        geom_vline(aes(xintercept=mean_unif2, color='Mean Set 2'),__
 →linetype='dashed', lwd=1.7) +
        labs(x=(TeX(r'(\lambda)')), y ='Posterior', title='Posterior Zoom -⊔
 →Uniform Prior', color='', fill='') +
        scale_color_manual(values = c('Set 1' = 'firebrick3', 'Mean Set 1' = u
 'Set 2' = 'dodgerblue4', 'Mean Set 2' =

    dodgerblue4')) +

        scale_fill_manual(values = c('CI Set 1' = 'firebrick3', 'CI Set 2' = 

    dodgerblue4')) +

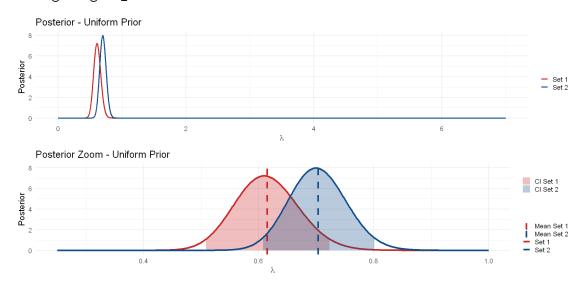
        theme_minimal(base_size = 18) +
       ylim(0, 8) + xlim(0.25, 1)
grid.arrange(grobs=list(unif.post.plot, unif.post.plot.zoom), nrow=2, ncol=1)
```

#### Warning message:

"Removed 6250 rows containing missing values or values outside the scale range (`geom\_line()`)."

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### 1.2)

```
[7]: #posterior with Jeffrey prior
     lambda_list <- seq(from=0.0001, to=10, by=delta.t)</pre>
     p1_like <- poisson(n1, lambda_list)</pre>
     p2_like <- poisson(n2, lambda_list)</pre>
     p.post_star_Jeffrey1 <- p1_like * (1/sqrt(lambda_list))</pre>
     p.post_Jeffrey1 <- p.post_star_Jeffrey1 / (sum(p.post_star_Jeffrey1)*delta.t)</pre>
     mean_Jeffrey1 <- sum(p.post_Jeffrey1 * lambda_list)*delta.t</pre>
     median Jeffrey1 <- my.median(lambda list, p.post Jeffrey1)</pre>
     variance_Jeffrey1 <- sum(((mean_Jeffrey1-lambda_list)^2)*p.post_Jeffrey1)*delta.</pre>
      ⇔t
     lower_Jeffrey1 <- ncredint(lambda_list, p.post_Jeffrey1, level = 0.</pre>
      →95)[['lower']]
     upper_Jeffrey1 <- ncredint(lambda_list, p.post_Jeffrey1, level = 0.</pre>
      p.post_star_Jeffrey2 <- p2_like * (1/sqrt(lambda_list))</pre>
     p.post_Jeffrey2 <- p.post_star_Jeffrey2 / (sum(p.post_star_Jeffrey2)*delta.t)</pre>
     mean_Jeffrey2 <- sum(p.post_Jeffrey2 * lambda_list)*delta.t</pre>
     median_Jeffrey2 <- my.median(lambda_list, p.post_Jeffrey2)</pre>
     variance_Jeffrey2 <- sum(((mean_Jeffrey2-lambda_list)^2)*p.post_Jeffrey2)*delta.</pre>
      ⇔t
     lower_Jeffrey2 <- ncredint(lambda_list, p.post_Jeffrey2, level = 0.</pre>
      →95)[['lower']]
     upper_Jeffrey2 <- ncredint(lambda_list, p.post_Jeffrey2, level = 0.</pre>
      →95)[['upper']]
     result_jeff <- data.frame(</pre>
                      mean = c(mean_Jeffrey1, mean_Jeffrey2),
                      median = c(median_Jeffrey1, median_Jeffrey2),
                      variance = c(variance_Jeffrey1, variance_Jeffrey2),
                      1.ci = c(lower_Jeffrey1, lower_Jeffrey2),
                      u.ci = c(upper_Jeffrey1, upper_Jeffrey2),
                      row.names = c('1st set', '2nd set')
```

```
[8]: print(sprintf('Mean value Set 1: %.3f - Mean value Set 2: %.3f', __
     →result_jeff$mean[1], result_jeff$mean[2]))
     print(sprintf('Median value Set 1: %.3f - Median value Set 2: %.3f',,,
     →result jeff$median[1], result jeff$median[2]))
     cat('\n')
     print(sprintf('SD Set 1: %.4f - SD Set 2: %.4f', result_jeff$variance[1], __
     →result_jeff$variance[2]))
     cat('\n')
     print(sprintf('0.95 CI Set 1 [%.3f - %.3f]', result_jeff$l.ci[1], result_jeff$u.

ci[1]))
     print(sprintf('0.95 CI Set 2 [%.3f - %.3f]', result_jeff$1.ci[2], result_jeff$u.
      →ci[2]))
    [1] "Mean value Set 1: 0.612 - Mean value Set 2: 0.702"
    [1] "Median value Set 1: 0.611 - Median value Set 2: 0.700"
    [1] "SD Set 1: 0.0031 - SD Set 2: 0.0025"
    [1] "0.95 CI Set 1 [0.506 - 0.721]"
    [1] "0.95 CI Set 2 [0.605 - 0.800]"
[9]: options(repr.plot.width = 17, repr.plot.height = 8)
     # plot
     jeff.post.plot <- ggplot() +</pre>
             geom_line(aes(lambda_list, p.post_Jeffrey1, color='Set 1'), lwd=1.3) +
             geom_line(aes(lambda_list, p.post_Jeffrey2, color='Set 2'), lwd=1.3) +
             labs(x=(TeX(r'(\lambda)')), y ='Posterior', title='Posterior - Jeffrey□
     →Prior', color='') +
             theme_minimal(base_size=18) +
             scale_color_manual(values = c('Set 1' = 'firebrick3', 'Set 2' = 
     # zoom in plot
     jeff.post.plot.zoom <- ggplot() +</pre>
             geom_line(aes(lambda_list, p.post_Jeffrey1, color='Set 1'), lwd=1.7) +
             geom_area(aes(x = lambda_list[lambda_list > lower_Jeffrey1 &__
      →lambda_list < upper_Jeffrey1],</pre>
                           y = p.post_unif1[lambda_list > lower_Jeffrey1 &_
      →lambda_list < upper_Jeffrey1],</pre>
                           fill='CI Set 1'), alpha=0.3) +
             geom_vline(aes(xintercept=mean_Jeffrey1, color='Mean Set 1'),__
      →linetype='dashed', lwd=1.7) +
             geom_line(aes(lambda_list, p.post_Jeffrey2, color='Set 2'), lwd=1.7) +
             geom area(aes(x = lambda list[lambda list > lower Jeffrey2 &___
      →lambda_list < upper_Jeffrey2],</pre>
```

```
y = p.post_unif2[lambda_list > lower_Jeffrey2 &_
 →lambda_list < upper_Jeffrey2],</pre>
                      fill='CI Set 2'), alpha=0.3) +
        geom_vline(aes(xintercept=mean_Jeffrey2, color='Mean Set 2'),__
 →linetype='dashed', lwd=1.7) +
        labs(x=(TeX(r'(\lambda)')), y ='Posterior', title='Posterior Zoom -□
 →Jeffrey Prior', color='', fill='') +
        scale_color_manual(values = c('Set 1' = 'firebrick3', 'Mean Set 1' = 
 'Set 2' = 'dodgerblue4', 'Mean Set 2' = __

    dodgerblue4')) +

        scale_fill_manual(values = c('CI Set 1' = 'firebrick3', 'CI Set 2' = 

    dodgerblue4')) +

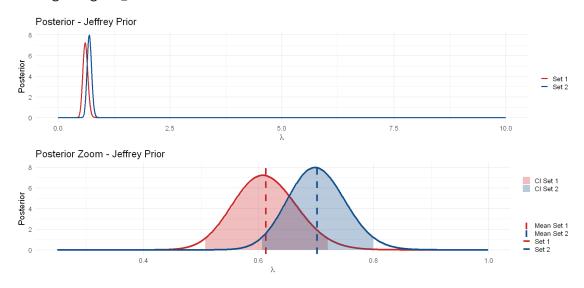
        theme_minimal(base_size = 18) +
       ylim(0, 8) + xlim(0.25, 1)
grid.arrange(grobs=list(jeff.post.plot, jeff.post.plot.zoom), nrow=2, ncol=)
```

## Warning message:

"Removed 9250 rows containing missing values or values outside the scale range (`geom\_line()`)."

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"Removed 9250 rows containing missing values or values outside the scale range (`geom\_line()`)."



## 2) Uniform prior

```
[10]: | stan_unif <- stan_model(file = './model/sm_ex6.2_unif.stan')</pre>
      y1 <- n1
      N1 <- length(y1)
      v2 <- n2
      N2 <- length(y2)
      dataList1 \leftarrow list(y = y1, N = N1)
      dataList2 \leftarrow list(y = y2, N = N2)
[11]: stanFit1 <- sampling(object = stan_unif, data = dataList1, chains = 3,
                            iter = 10000, warmup = 1000, thin = 1)
      stanFit2 <- sampling(object = stan_unif, data = dataList2, chains = 3,</pre>
                            iter = 10000, warmup = 1000, thin = 1)
[12]: print(sprintf('STAN fit with dataset 1'))
      cat('\n')
      print(stanFit1)
     [1] "STAN fit with dataset 1"
     Inference for Stan model: anon_model.
     3 chains, each with iter=10000; warmup=1000; thin=1;
     post-warmup draws per chain=9000, total post-warmup draws=27000.
               mean se_mean
                               sd
                                     2.5%
                                               25%
                                                       50%
                                                                75%
                                                                      97.5% n_eff Rhat
     lambda
                                                                       0.73 9871
               0.61
                        0.00 0.06
                                     0.51
                                              0.58
                                                      0.61
                                                              0.65
            -183.31
                        0.01 0.73 -185.39 -183.46 -183.03 -182.85 -182.79 11869
     lp__
                                                                                     1
     Samples were drawn using NUTS(diag_e) at Mon Jun 3 15:05:21 2024.
     For each parameter, n_eff is a crude measure of effective sample size,
     and Rhat is the potential scale reduction factor on split chains (at
     convergence, Rhat=1).
[13]: print(sprintf('STAN fit with dataset 2'))
      cat('\n')
      print(stanFit2)
     [1] "STAN fit with dataset 2"
     Inference for Stan model: anon_model.
     3 chains, each with iter=10000; warmup=1000; thin=1;
     post-warmup draws per chain=9000, total post-warmup draws=27000.
               mean se_mean
                               sd
                                     2.5%
                                               25%
                                                       50%
                                                                75%
                                                                      97.5% n_eff Rhat
     lambda
               0.70
                        0.00 0.05
                                     0.61
                                              0.67
                                                      0.70
                                                               0.74
                                                                       0.81 8469
            -266.76
                        0.01 0.70 -268.78 -266.91 -266.48 -266.31 -266.26 12499
                                                                                     1
     lp__
```

Samples were drawn using NUTS(diag\_e) at Mon Jun 3 15:05:31 2024. For each parameter, n\_eff is a crude measure of effective sample size, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat=1).

[14]: chain1 <- rstan::As.mcmc.list(stanFit1, pars = c('lambda'))
summary(chain1)</pre>

Iterations = 1001:10000
Thinning interval = 1
Number of chains = 3
Sample size per chain = 9000

1. Empirical mean and standard deviation for each variable, plus standard error of the mean:

Mean SD Naive SE Time-series SE 0.6149246 0.0561459 0.0003417 0.0005661

2. Quantiles for each variable:

2.5% 25% 50% 75% 97.5% 0.5106 0.5761 0.6131 0.6520 0.7299

[15]: chain2 <- rstan::As.mcmc.list(stanFit2, pars = c('lambda'))
summary(chain2)</pre>

Iterations = 1001:10000
Thinning interval = 1
Number of chains = 3
Sample size per chain = 9000

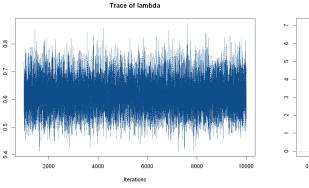
1. Empirical mean and standard deviation for each variable, plus standard error of the mean:

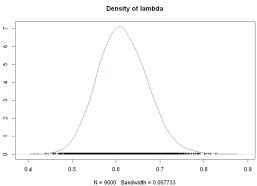
Mean SD Naive SE Time-series SE 0.7036984 0.0498203 0.0003032 0.0005383

2. Quantiles for each variable:

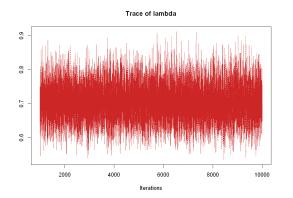
2.5% 25% 50% 75% 97.5% 0.6097 0.6697 0.7023 0.7364 0.8052

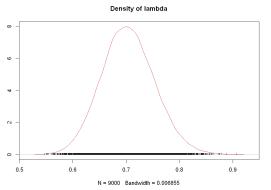
```
[16]: 1.mean.1 <- summary(chain1)$statistics[1]
      1.mean.2 <- summary(chain2)$statistics[1]</pre>
      1.med.1 <- summary(chain1)$quantiles[3]</pre>
      1.med.2 <- summary(chain2)$quantiles[3]</pre>
      print(sprintf('Mean value Set 1: %.3f - Mean value Set 2: %.3f', 1.mean.1, 1.
       \rightarrowmean.2))
      print(sprintf('Median value Set 1: %.3f - Median value Set 2: %.3f', l.med.1, 
       \rightarrow1.med.2))
      cat('\n')
      1.lo.ci.1 <- summary(chain1)$quantiles[1]</pre>
      1.up.ci.1 <- summary(chain1)$quantiles[5]</pre>
      1.lo.ci.2 <- summary(chain2)$quantiles[1]</pre>
      1.up.ci.2 <- summary(chain2)$quantiles[5]</pre>
      print(sprintf('0.95 CI Set 1 [%.3f - %.3f]', 1.lo.ci.1, 1.up.ci.1))
      print(sprintf('0.95 CI Set 2 [%.3f - %.3f]', 1.lo.ci.2, 1.up.ci.2))
      [1] "Mean value Set 1: 0.615 - Mean value Set 2: 0.704"
      [1] "Median value Set 1: 0.613 - Median value Set 2: 0.702"
      [1] "0.95 CI Set 1 [0.511 - 0.730]"
      [1] "0.95 CI Set 2 [0.610 - 0.805]"
[17]: options(repr.plot.width = 17, repr.plot.height = 6)
      plot(chain1, col='dodgerblue4')
```





```
[18]: options(repr.plot.width = 17, repr.plot.height = 6)
plot(chain2, col='firebrick3')
```





## Jeffrey's prior

```
[19]: stan_jeff <- stan_model(file = './model/sm_ex6.2_jeffrey.stan')
y1 <- n1
N1 <- length(y1)

y2 <- n2
N2 <- length(y2)

dataList1 <- list(y = y1, N = N1)
dataList2 <- list(y = y2, N = N2)</pre>
```

```
[21]: print(sprintf('STAN fit with dataset 1'))
    cat('\n')
    print(stanJeff1)
```

#### [1] "STAN fit with dataset 1"

Inference for Stan model: anon\_model.
3 chains, each with iter=10000; warmup=1000; thin=1;
post-warmup draws per chain=9000, total post-warmup draws=27000.

```
mean se_mean
                         sd
                                2.5%
                                         25%
                                                 50%
                                                          75%
                                                                97.5% n_eff Rhat
                                0.51
                                        0.57
                                                0.61
                                                         0.65
lambda
          0.61
                  0.00 0.06
                                                                 0.72 9202
lp__
       -182.81
                  0.01 0.72 -184.85 -182.98 -182.53 -182.36 -182.30 9982
                                                                                1
```

Samples were drawn using NUTS(diag\_e) at Mon Jun 3 15:05:46 2024. For each parameter, n\_eff is a crude measure of effective sample size,

and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat=1).

```
[22]: print(sprintf('STAN fit with dataset 2'))
    cat('\n')
    print(stanJeff2)
```

[1] "STAN fit with dataset 2"

Inference for Stan model: anon\_model.
3 chains, each with iter=10000; warmup=1000; thin=1;
post-warmup draws per chain=9000, total post-warmup draws=27000.

```
25%
                                                  50%
                                                                97.5% n_eff Rhat
          mean se_mean
                          sd
                                2.5%
                                                          75%
lambda
          0.70
                  0.00 0.05
                                0.60
                                        0.67
                                                 0.70
                                                         0.73
                                                                  0.80 9470
                  0.01 0.74 -268.55 -266.56 -266.14 -265.96 -265.91 10720
lp__
       -266.42
                                                                                1
```

Samples were drawn using NUTS(diag\_e) at Mon Jun 3 15:05:56 2024. For each parameter, n\_eff is a crude measure of effective sample size, and Rhat is the potential scale reduction factor on split chains (at convergence, Rhat=1).

```
[23]: chain1 <- rstan::As.mcmc.list(stanJeff1, pars = c('lambda'))
summary(chain1)</pre>
```

```
Iterations = 1001:10000
Thinning interval = 1
Number of chains = 3
Sample size per chain = 9000
```

1. Empirical mean and standard deviation for each variable, plus standard error of the mean:

```
Mean SD Naive SE Time-series SE 0.6097067 0.0555149 0.0003379 0.0005698
```

2. Quantiles for each variable:

```
2.5% 25% 50% 75% 97.5% 0.5060 0.5717 0.6078 0.6463 0.7230
```

```
[24]: chain2 <- rstan::As.mcmc.list(stanJeff2, pars = c('lambda'))
summary(chain2)</pre>
```

Iterations = 1001:10000

```
Thinning interval = 1
Number of chains = 3
Sample size per chain = 9000
```

1. Empirical mean and standard deviation for each variable, plus standard error of the mean:

```
Mean
                               Naive SE Time-series SE
0.6998283
               0.0504295
                              0.0003069
                                              0.0005061
```

2. Quantiles for each variable:

```
2.5%
                        75% 97.5%
          25%
                 50%
0.6041 0.6654 0.6987 0.7330 0.8022
```

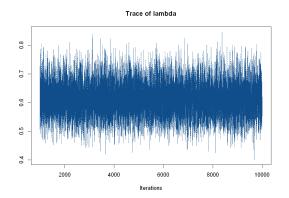
```
[25]: 1.mean.1 <- summary(chain1)$statistics[1]
      1.mean.2 <- summary(chain2)$statistics[1]</pre>
      1.med.1 <- summary(chain1)$quantiles[3]</pre>
      1.med.2 <- summary(chain2)$quantiles[3]</pre>
      print(sprintf('Mean value Set 1: %.3f - Mean value Set 2: %.3f', 1.mean.1, 1.
       \rightarrowmean.2))
      print(sprintf('Median value Set 1: %.3f - Median value Set 2: %.3f', 1.med.1, 
       \rightarrow1.med.2))
      cat('\n')
      1.lo.ci.1 <- summary(chain1)$quantiles[1]</pre>
      1.up.ci.1 <- summary(chain1)$quantiles[5]</pre>
      1.lo.ci.2 <- summary(chain2)$quantiles[1]</pre>
      1.up.ci.2 <- summary(chain2)$quantiles[5]</pre>
      print(sprintf('0.95 CI Set 1 [%.3f - %.3f]', 1.lo.ci.1, 1.up.ci.1))
      print(sprintf('0.95 CI Set 2 [%.3f - %.3f]', 1.lo.ci.2, 1.up.ci.2))
      [1] "Mean value Set 1: 0.610 - Mean value Set 2: 0.700"
```

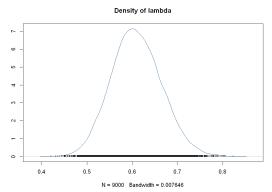
```
[1] "0.95 CI Set 1 [0.506 - 0.723]"
```

```
[26]: options(repr.plot.width = 17, repr.plot.height = 6)
     plot(chain1, col='dodgerblue4')
```

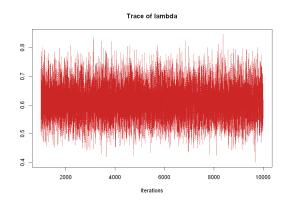
<sup>[1] &</sup>quot;Median value Set 1: 0.608 - Median value Set 2: 0.699"

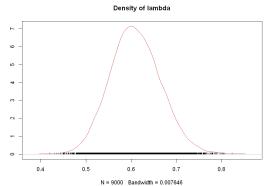
<sup>[1] &</sup>quot;0.95 CI Set 2 [0.604 - 0.802]"





[27]: options(repr.plot.width = 17, repr.plot.height = 6)
 plot(chain1, col='firebrick3')





```
3)
```

[28]: n <- 116 #n° of sample y <- 11 #samples w/ hl bacteria

3.1)

[29]: p\_freq <- y/n
cat('The frequenties estimator is:', round(p\_freq\*100, 2), '%')</pre>

The frequenties estimator is: 9.48 %

3.2)

[30]: delta.t <- 0.0001 p <- seq(0, 1, by = delta.t)

#### 3.3)

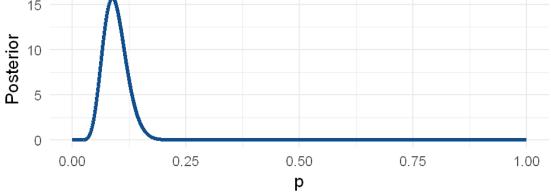
```
[31]: # Bayesian estimator for p
      beta_mean <- (sum(p.beta_post * p)) * delta.t</pre>
      beta_var <- sum(((beta_mean - p)^2) *p.beta_post) * delta.t</pre>
      beta_low <- ncredint(p, p.beta_post, level=0.95)[['lower']]</pre>
      beta_up <- ncredint(p, p.beta_post, level=0.95)[['upper']]</pre>
      beta_res <- tibble(</pre>
                       mean = beta_mean,
                       variance = beta_var,
                       1.ci = beta_low,
                       u.ci = beta_up
      beta_res
      # zoomed plot
      beta.plot.bayes <- beta_post_plot +</pre>
               geom_area(aes(x = p[p>beta_low & p<beta_up],</pre>
                              y = p.beta_post[p>beta_low & p<beta_up],</pre>
                             fill='CI'), alpha=0.3) +
               geom_vline(aes(xintercept = beta_mean, color='Mean'),__
       →linetype="dashed", linewidth=1.5) +
               scale_color_manual(values = c('Mean' = 'firebrick3')) +
               scale_fill_manual(values = c('CI' = 'dodgerblue4')) +
               xlim(0, 0.2) +
               labs(title='Zoom', color='', fill='')
      grid.arrange(grobs=list(beta_post_plot, beta.plot.bayes), nrow=2)
```

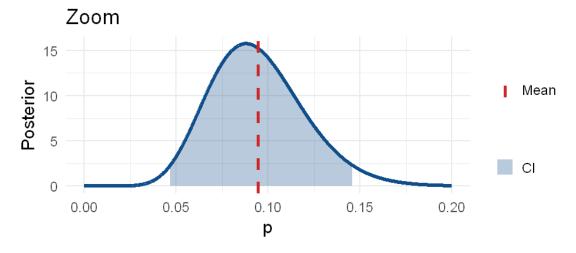
	mean	variance	l.ci	u.ci
A tibble: $1 \times 4$	<dbl $>$	<dbl></dbl>	<dbl $>$	<dbl $>$
·	0.09448819	0.0006684388	0.0467	0.1459

## Warning message:

"Removed 8000 rows containing missing values or values outside the scale range (`geom\_line()`)."







## 3.4)

## Frequentist approach

```
[32]: #Frequentist Approach
      alpha = 0.05
      p = seq(0, n, 1)
      p.freq = pbinom(p, n, prob=0.1, lower.tail = TRUE)
      lower.limit = max(p[p.freq<=alpha/2])</pre>
```

```
upper.limit = min(p[p.freq>=1-alpha/2])
alpha.star = p.freq[lower.limit] + 1 - p.freq[upper.limit]
cat('The closest we can get to alpha=0.05 is: ', alpha.star, '\n')
cat('The lower and upper limit to the 95% interval are: (', lower.limit, ', ', \upper.limit, ')\n')
cat('Since the number of successful events (',y,') fall whithin this range, we \upper.limit \upper
```

The closest we can get to alpha=0.05 is: 0.04761008 The lower and upper limit to the 95% interval are: (5, 18) Since the number of successful events (11) fall whithin this range, we CANNOT reject the null hypotesis.

Warning message:

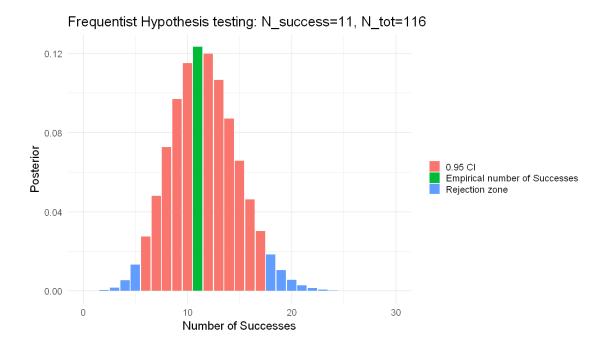
"Removed 88 rows containing missing values or values outside the scale range (`geom\_col()`)."

Warning message:

"Removed 1 row containing missing values or values outside the scale range (`geom\_col()`)."

Warning message:

"Removed 87 rows containing missing values or values outside the scale range (`geom\_col()`)."



## Bayesian approach

```
[34]: cat('Since p=0.1 fall whithin the 95% CI (',beta_low,',',beta_up,') we CANNOT

→reject the null hypothesis.')
```

Since p=0.1 fall whithin the 95% CI ( 0.0467 , 0.1459 ) we CANNOT reject the null hypothesis.

#### 3.5

```
[35]: n <- 165
y <- 9

p_freq <- y/n
cat('The frequenties estimator is:', round(p_freq*100, 2), '%')</pre>
```

The frequenties estimator is: 5.45 %

## 3.6 - 3.7

```
[36]: delta.t <- 0.0001
  p <- seq(0, 1, by=delta.t)

#likelihood
  p.like <- dbinom(y, n, prob=p)

#the two prior
  p.beta_prior <- dbeta(p, 1, 10)</pre>
```

```
p.old_prior <- p.beta_post</pre>
#the two possible posterior
p.beta_post <- (p.like * p.beta_prior) / (delta.t * sum(p.like * p.beta_prior))</pre>
p.old_post <- (p.like * p.old_prior) / (delta.t * sum(p.like * p.old_prior))</pre>
#find the two baiesian estimator for the two posterior
beta_mean <- sum(p * p.beta_post) * delta.t</pre>
beta_var <- sum(((p-beta_mean)^2) * p.beta_post) * delta.t</pre>
beta_low <- ncredint(p, p.beta_post, level = 0.95)[['lower']]</pre>
beta_up <- ncredint(p, p.beta_post, level = 0.95)[['upper']]</pre>
beta_cred_int <- c(beta_low, beta_up)</pre>
old_mean <- sum(p * p.old_post) * delta.t</pre>
old_var <- sum(((p-old_mean)^2) * p.old_post) * delta.t</pre>
old_low <- ncredint(p, p.old_post, level=0.95)[['lower']]</pre>
old_up <- ncredint(p, p.old_post, level=0.95)[['upper']]</pre>
old_cred_int <- c(old_low, old_up)</pre>
#print the result in a dataframe
result <- data.frame(</pre>
          Mean = c(beta mean, old mean),
          Variance = c(beta_var, old_var),
          1.ci = c(beta_low, old_low),
          u.ci = c(beta_up, beta_low),
          row.names = c('Prior: Beta', 'Prior: Old posterior')
result
options(repr.plot.width = 17, repr.plot.height = 8)
#plot the two posterior
prior_plot <- ggplot()+</pre>
        geom_line(aes(p, p.beta_post, color='Beta'), lwd=1.7) +
        geom_line(aes(p, p.old_post, color='Old post'), lwd=1.7) +
        scale_color_manual(values = c('Beta' = 'dodgerblue4', 'Old post' = 
theme_minimal(base_size = 18) +
        labs(x='p', y='Posterior', color='')
prior_plot_zoom <- prior_plot +</pre>
        geom vline(aes(xintercept=beta_mean, color = 'Mean - Beta'), __
 →linetype='dashed', lwd=1.5) +
        geom_area(aes(x = p[p>beta_low & p<beta_up],</pre>
                       y = p.beta_post[p>beta_low & p<beta_up],</pre>
                       fill='CI - Beta'), alpha=0.3) +
```

		Mean	Variance	l.ci	u.ci
A data.frame: $2 \times 4$		<dbl></dbl>	<dbl></dbl>	<dbl $>$	<dbl $>$
	Prior: Beta	0.05681818	0.0003027677	0.0251	0.0914
	Prior: Old posterior	0.07191781	0.0002278008	0.0436	0.0251

Scale for colour is already present.

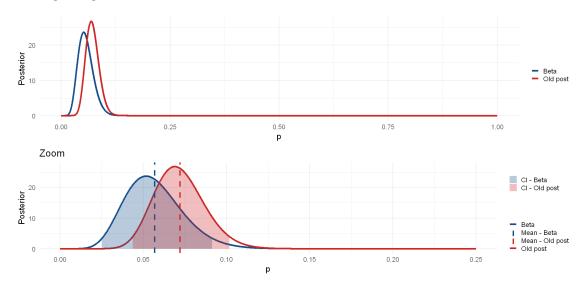
Adding another scale for colour, which will replace the existing scale.

Warning message:

"Removed 7500 rows containing missing values or values outside the scale range (`geom\_line()`)."

Warning message:

"Removed 7500 rows containing missing values or values outside the scale range ( $`geom\_line()`)$ ."



### 3.8)

#### Frequentist approach

```
[37]: alpha = 0.05
      p = seq(0, n, 1)
      p.freq = pbinom(p, n, prob=0.1, lower.tail = TRUE)
      lower.limit = max(p[p.freq<=alpha/2])</pre>
      upper.limit = min(p[p.freq>=1-alpha/2])
      alpha.star = p.freq[lower.limit] + 1 - p.freq[upper.limit]
      cat('The closest we can get to alpha=0.05 is: ', alpha.star, '\n')
      cat('The lower and upper limit to the 95% interval are: (', lower.limit, ', ', |
      →upper.limit, ')\n')
      cat('Since the number of successful events (', y, ') slightly fall whithin this⊔
       \rightarrowrange, we are not sure if we CAN or
          CANNOT reject the null hypotesis. \n')
      options(repr.plot.width=12, repr.plot.height=7)
      h0.plot <- ggplot()+
          geom_col(aes(x=p, y=dbinom(p, n, 0.1), fill='Confidence Intervall: 95%'))+
          geom_col(aes(x=p[p.freq<=alpha/2], y=dbinom(p[p.freq<=alpha/2], n, 0.1),__</pre>
       →fill='Rejection range'))+
          geom_col(aes(x=p[p.freq>=1-alpha/2], y=dbinom(p[p.freq>=1-alpha/2], n, 0.
       →1), fill='Rejection range'))+
          geom_col(aes(x=y, y=dbinom(y, n, 0.1), fill='Empirical number of_

¬Successes'))+
          xlim(0, 30) +
          labs(x='Number of Successes', y = 'Posterior', title='Frequentist∪
       →Hypothesis testing: N_success=9, N_tot=165') +
          theme_minimal(base_size = 18)
      h0.plot
```

```
The closest we can get to alpha=0.05 is: 0.04532508

The lower and upper limit to the 95% interval are: (8, 24)

Since the number of successful events (9) slightly fall whithin this range, we are not sure if we CAN or

CANNOT reject the null hypotesis.

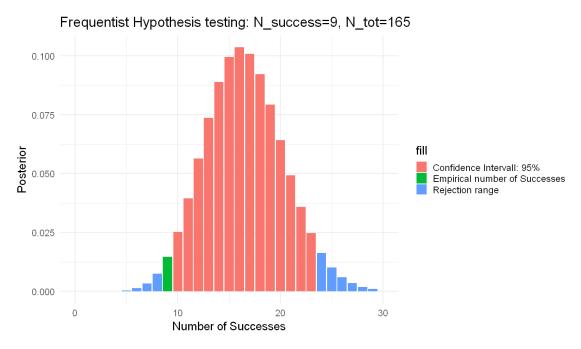
Warning message:

"Removed 137 rows containing missing values or values outside the scale range (`geom_col()`)."

Warning message:

"Removed 1 row containing missing values or values outside the scale
```

range (`geom\_col()`)."
Warning message:
"Removed 136 rows containing missing values or values outside the scale
range (`geom\_col()`)."



### Bayesian approach

```
[38]: cat('Using the old posterior as a prior:\n')
cat('Since p=0.1 slightly falls whithin the 95% CI (',old_low,',',old_up,') we

→ are not sure if we CAN or
CANNOT reject the null hypotesis.')

cat('\n')
cat('\nUsing the beta(1,10) as a prior:\n')
cat('since p=0.1 slightly falls outside the 95% CI (',beta_low,',',beta_up,')
→ we are not sure if we CAN or
CANNOT reject the null hypotesis.')

Using the old posterior as a prior:
```

```
4)
[39]: stan 4 <- stan model(file = './model/sm ex6.4.stan')
      dataList \leftarrow list(y = 11, n = 116)
[40]: stanFit <- sampling(object = stan_4, data = dataList, chains = 3,
                           iter = 10000, warmup = 1000, thin = 1)
[41]: print(sprintf('STAN fit'))
      cat('\n')
      print(stanFit)
     [1] "STAN fit"
     Inference for Stan model: anon_model.
     3 chains, each with iter=10000; warmup=1000; thin=1;
     post-warmup draws per chain=9000, total post-warmup draws=27000.
            mean se_mean
                                         25%
                                                50%
                                                       75% 97.5% n_eff Rhat
                            sd
                                 2.5%
            0.09
                                 0.05
                                        0.08
                                               0.09
                                                             0.15 9655
                    0.00 0.03
                                                      0.11
     lp__ -40.24
                    0.01 0.71 -42.25 -40.42 -39.96 -39.78 -39.73 8851
     Samples were drawn using NUTS(diag_e) at Mon Jun 3 15:06:20 2024.
     For each parameter, n_eff is a crude measure of effective sample size,
     and Rhat is the potential scale reduction factor on split chains (at
     convergence, Rhat=1).
[42]: chain <- rstan::As.mcmc.list(stanFit, pars = c('p'))
      summary(chain)
     Iterations = 1001:10000
     Thinning interval = 1
     Number of chains = 3
     Sample size per chain = 9000
     1. Empirical mean and standard deviation for each variable,
        plus standard error of the mean:
                                 SD
                                          Naive SE Time-series SE
               Mean
          0.0940470
                         0.0259066
                                         0.0001577
                                                        0.0002645
     2. Quantiles for each variable:
        2.5%
                 25%
                         50%
                                  75%
                                        97.5%
```

0.04975 0.07531 0.09211 0.11045 0.15011

```
[43]: l.mean.1 <- summary(chain)$statistics[1]
l.med.1 <- summary(chain)$quantiles[3]

print(sprintf('Mean value of p: %.3f', l.mean.1))
print(sprintf('Median value of p: %.3f', l.med.1))
cat('\n')

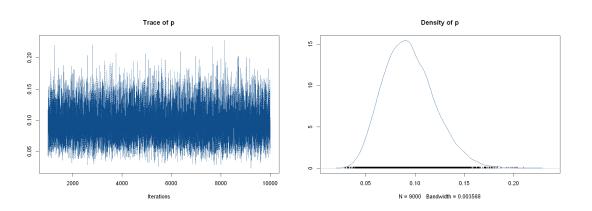
l.lo.ci.1 <- summary(chain)$quantiles[1]
l.up.ci.1 <- summary(chain)$quantiles[5]

print(sprintf('0.95 CI [%.3f - %.3f]', l.lo.ci.1, l.up.ci.1))

[1] "Mean value of p: 0.094"
[1] "Median value of p: 0.092"

[1] "0.95 CI [0.050 - 0.150]"

[44]: options(repr.plot.width = 17, repr.plot.height = 6)
plot(chain, col='dodgerblue4')</pre>
```



[]: