Model validation

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Outline

- Normal distribution
- Linear models
- Assumption violation

Recall:

$$y_i = \alpha + \mathbf{x}_i \boldsymbol{\beta} + \epsilon_i \sim \mathcal{N}(0, \sigma^2)$$

- the systematic component: the regression line
- the random component: left-over stuff

Our models make assumptions. Sometimes, these are realistic.

All models are wrong, but some are useful

Is is important to:

- Know which assumptions your model makes
- Consider if they are realistic
- Determine if they are violated
- Assess the consequences

Tests make assumptions too, but more on that later

Linear regression assumptions

$$y_i = \alpha + \sum_{k=1}^p x_{ik} \beta_k + \epsilon_i \sim \mathcal{N}(0, \sigma^2)$$

- Linearity
- Normality: normally distributed errors
- Homoscedastcity: same variance for all errors
- Independence of errors
- Lack of perfect multicollinearity $(\mathbf{x}_1 \neq \mathbf{x}_2)$
- No outliers
- No error in the explanatory variables

Violating assumptions

- Linearity: Eek
- Normality: cannot trust tests and confidence intervals
- Homoscedasticity: cannot trust tests and confidence intervals
- Independence of errors: Eek
- Lack of perfect multicollinearity
- No outliers: biased parameter estimates



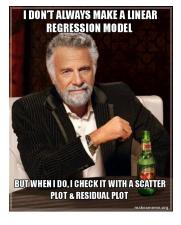
Assumption checking: methods

Tests:

- Score test for non-constant error variance
- car::ncvTest
- Durbin-Watson test for Autocorrelated errors car::durbinWatsonTest
- Bonferroni outlier test
 - car::outlierTest
- Shapiro-Wilk normality test shapiro.est

Rather: **Residual plots**

- plot(model)
- Various packages (e.g., car or DHARMa)



Errors vs. Residuals

The true model:

$$y_i = \alpha + \sum_{k=1}^p x_{ik} \beta_k + \epsilon_i \sim \mathcal{N}(0, \sigma^2)$$

The fitted model:

$$y_i = \hat{\alpha} + \sum_{k=1}^{p} x_{ik} \hat{\beta}_k + \hat{\epsilon}_i$$

$$\begin{aligned} & \blacktriangleright \ \epsilon_i = y_i - \alpha - \sum_{i=1}^n x_{ik} \beta \\ & \blacktriangleright \ \hat{\epsilon}_i = y_i - \hat{\alpha} - \sum_{i=1}^n x_{ik} \hat{\beta} \end{aligned}$$

$$\hat{\epsilon}_i = y_i - \hat{\alpha} - \sum_{i=1}^n x_{ik} \hat{\beta}_i$$

Residuals are the **observed** deviation from the regression line.

In R

```
model <- lm(y~x)
residuals(model) # estimated random part
fitted(model) # estimated systematic part</pre>
```

Residual plots

Residual plots should have no discernible patterns.

- No curvature
- No shapes whatsoever
- No outliers
- No change in variance

This can be difficult with few observations.

Standardised residuals

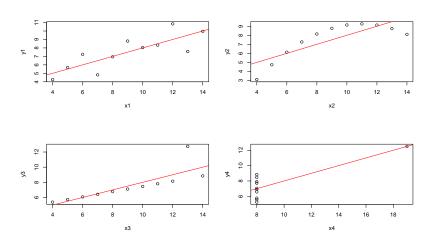
$$\frac{\hat{\epsilon_i}}{s\sqrt{1-h_i}}, \qquad \text{where } h_i \text{ is "influence"} \tag{1}$$

Studentized residuals

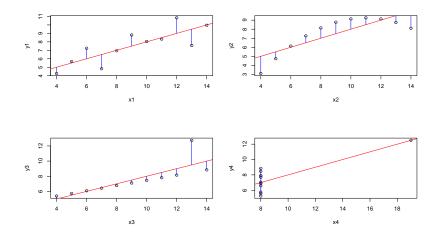
- omit observation i
- re-compute residual variance
- calculate standardised residual (with hat matrix)

rstand or rstudent

Introducing: Anscombe's quartet



All of these regressions have the same parameter estimates



Our toolbox

- Residuals versus fitted: $\hat{\epsilon}_i \ vs. \ \hat{\mu}_i$
- Observed versus theoretical quantiles:

$$Pr(X \le y_i) \ vs. \ Pr(X \le x)$$

- \triangleright Cook's distance plot: $\hat{\epsilon}_i \ vs. \ \hat{\mu}_i$
- Standardized residuals vs. fitted (Scale-location)
- Leverage and influence

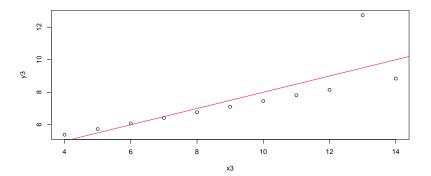
Leverage: hat matrix

$$\hat{\mu}_i = \hat{\alpha} - \sum_{i=1}^n x_{ik} \hat{\beta}$$

but we can also write:

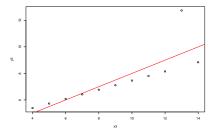
$$\hat{\mu} = \mathbf{H}\mathbf{y}$$

- ► **H** is the "hat" matrix
- $ightharpoonup \mathbf{H} = \mathbf{X}(\mathbf{X}^{\top}\mathbf{X})^{-1}\mathbf{X}^{\top}\mathbf{y}$
- Its diagonal gives "leverage" and is between 0 and 1
- Leverage: observation that is far from others on the explanatory variable



An observation that has undue impact on the regression line.

Outliers



- Outliers are observations far away from the regression line
- Can be erroneous
 - Might be real data
- Remove with caution
- Better: sensitivity analysis

Cook's distance

Cook's distance: how much the fitted values change if we remove an observation

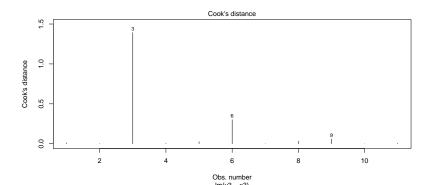
$$D_{j} = \frac{\sum_{i=1}^{n} (\hat{y}_{i} - \hat{y}_{i(-j)})^{2}}{\sigma^{2}}$$

- $\triangleright \hat{y}_i$ prediction
- $\hat{y}_{i(-j)}$ prediction for model with data point j removed
- $ightharpoonup \sigma^2$ residual variance

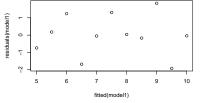
What is influential?

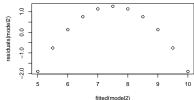
Large values of D_i mean a large influence

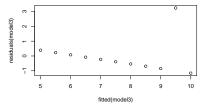
Usually $D_i > 1$, or 4/n

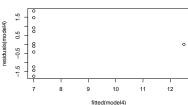




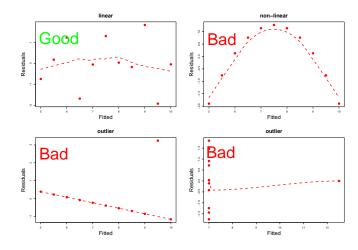




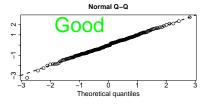


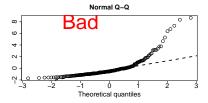


Anscombe's quartet: residuals vs. fitted



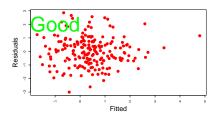
Residual diagnostics: normality with QQ-plot

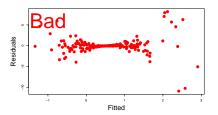




- 1. compare residuals to theoretical quantiles
- 2. points should follow the line

Residual diagnostics: constant variance





- compare residuals to fitted
- 2. points should be randomly distributed

Addressing assumption violations

- Curvature? Add quadratic term
- Outlier? Sensitivity analysis (if not error)
- Normality? GLM
- Constant variance? GLM

Alternative: data transformation

Data transformation

$$\mathbb{E}\{\mathsf{h}(y_i)|x_i\} = \alpha + \beta x_i \tag{2}$$

- $ightharpoonup h(\cdot)$ data transformation (e.g. $\sqrt(\cdot)$ or $\log(\cdot)$)
- To stabilize variance
- To normalize residuals

Data transformation

$$\mathbb{E}\{\mathsf{h}(y_i)|x_i\} = \alpha + \beta x_i \tag{2}$$

- \blacktriangleright h(·) data transformation (e.g. $\sqrt(\cdot)$ or log(·))
- To stabilize variance
- To normalize residuals

Not the same as

$$h(\mathbb{E}\{y_i|x_i\}) = \alpha + \beta x_i \tag{3}$$

Use a Generalised linear model instead

Summary

Now we have a valid model We still need to check ${f how}$ good it is (more on that tomorrow)