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- the instances that constitute a sample are iid
- the advantage that the likelihood of the sample is the product of likelihoods of the individual instances
- this assumption not valid where successive instances are dependent
- applications: letters in a word, speech recognition only certain sequences of phonemes are allowed
- a sequence can be characterized as being generated by a parametric random process



- Discrete Markov Processes
- a system at any time is in one of a set of N distinct states S<sub>1</sub>,...,S<sub>N</sub>
- the state at time t is q<sub>t</sub> = S<sub>i</sub> means that at time t the system is in state S<sub>i</sub>
- the system moves to a state with a given probability, depending on the values of the previous states

$$P(q_{t+1} = S_j | q_t = S_i, q_{t-1} = S_k, ...)$$

First order Markov model,

$$P(q_{t+1} = S_i | q_t = S_i, q_{t-1} = S_k, ...) = P(q_{t+1} = S_i | q_t = S_i) \equiv a_{ij}$$

simplifying - the transition probabilities are independent of time



- ullet  $a_{ij} \geq 0$  and  $\sum_{j=1}^N a_{ij} = 1$   $oldsymbol{\mathsf{A}} = [a_{ij}]$  is a N imes N matrix whose rows sum to 1
- it can be seen as stochastic automaton
- initial probabilities  $\pi_i$  is the probability that the first state in the sequence is  $S_i$   $\pi_i \equiv P(q_1 = S_i)$  satisfying  $\sum_{i=1}^N \pi_i = 1$
- $\Pi = [\pi_i]$  is a vector of N elements of that sum to 1
- Observable Markov model the states are observable
- the system moves from one state to another, results into an observation sequence and that is a sequence of states
- the output of the process is the set of states at each instant of time where each state corresponds to a physical observable event



an observance sequence O that is the state sequence
 O = Q = {q<sub>1</sub>q<sub>2</sub>····q<sub>T</sub>}

$$P(O = Q|\mathbf{A}, \Pi) = P(q_1) \prod_{t=2}^{T} P(q_t|q_{t-1}) = \pi_{q_1} a_{q_1 q_2} \cdots a_{q_{T-1} q_T}$$

- $\bullet$   $\pi_{q_1}$  is the probability that the first state  $q_1$ ,  $a_{q_1q_2}$  is the probability of going from  $q_1$  to  $q_2$
- N urns, each urn contains balls of only one color; there is an urn of red balls, another of blue balls and so forth
- draws balls from urns one by one and their color are shown
- three states: S<sub>1</sub>: red, S<sub>2</sub>: blue, S<sub>3</sub>: green and q<sub>t</sub> denote the color of the ball drawn at time t
- initial probabilities  $\Pi = [0.5, 0.2, 0.3]^T$



- a<sub>ij</sub> is the probability of drawing from urn j (a ball of color j) after drawing a ball of color i from urn i
- the transition matrix

$$\mathbf{A} = \left[ \begin{array}{cccc} 0.4 & 0.3 & 0.3 \\ 0.2 & 0.6 & 0.2 \\ 0.1 & 0.1 & 0.8 \end{array} \right]$$

- given Π and A it is easy to generate K random sequences each of length T
- calculate the probability of a sequence "red, red, green, green", i.e. the observation sequence  $O = \{S_1, S_1, S_3, S_3\}$

$$P(O|\mathbf{A}, \Pi) = P(S_1) \cdot P(S_1|S_1) \cdot P(S_3|S_1) \cdot P(S_3|S_3)$$
  
=  $\pi_1 \cdot a_{11} \cdot a_{13} \cdot a_{33}$   
=  $0.5 \cdot 0.4 \cdot 0.3 \cdot 0.8 = 0.048$ 



- how to learn the parameters  $\Pi$ . A given K sequences of length Twhere  $q_t^k$  is the state at time t of sequence k
- the initial probability estimate is the number of sequences starting with  $S_i$  divided by the number of sequences

$$\hat{\pi_i} = \frac{\{\text{number of sequences starting with} S_i\}}{\{\text{number of sequences}\}} = \frac{\sum_k \mathbb{1}(q_1^k = S_i)}{K}$$

- 1(b) is 1 if b is true and 0 otherwise
- the transition probabilities, the estimate for aii is the number of transitions from  $S_i$  to  $S_i$  divided by the total number of transitions from  $S_i$  over all sequences

$$\hat{a}_{ij} = \frac{\{\text{number of transition from } S_i \text{to } S_j\}}{\{\text{number of transitions from } S_i\}}$$

$$= \frac{\sum_k \sum_{t=1}^T 1(q_t^k = S_i \text{and } q_{t+1}^k = S_j)}{\sum_k \sum_{t=1}^T 1(q_t^k = S_i)}$$



- a<sub>12</sub> is the number of times a blue ball follows a red ball divided by the total number of red ball draws over all sequences
- in HMM, the states are not observable, but when a state is visited, an observation is recorded that is a probabilistic function of the state
- a discrete observation in each state from the set {v<sub>1</sub>, v<sub>2</sub>,..., v<sub>M</sub>}

$$b_j(m) \equiv P(O_t = v_m | q_t = S_j)$$

•  $b_j(m)$  is the observation, or emission probability that we observe  $v_m$  (m = 1, ..., M) in state  $S_j$ 



- the state sequence Q is not observed, (so called hidden), but it should be inferred from the observation sequence O
- many different state sequences Q that could have generated the same observation sequence but with different probabilities, just as,
- given an iid sample from a normal distribution, there are an infinite number of  $(\mu, \sigma)$  value pair possible and interested in the one having the highest likelihood of generating the sample
- there are two sources of randomness: randomly moving from one state to another and the observation in a state
- the hidden case corresponds to the urn-and-ball example where each urn contains balls of different colors
- let b<sub>j</sub>(m) denote the probability of drawing a ball of color m from urn

- again a sequence of ball colors observed but without knowing the sequence of urns from which the balls were drawn
- as if the urns are placed behind a curtain and somebody picks a ball at random from one of the urns and shows us only the ball, without showing us the urn from which it is picked
- the ball is returned to the urn to keep the probabilities the same
- the number of ball colors may be different from the number of urns
- example: three urns and the observation sequence is
   O = {red, red, green, blue, yellow}
- earlier, knowing the observation (ball color), the state (urn) is known because there are separate urns for separate colors and each urn contained balls of only one color



- the observable model is a special case of the hidden model where
   M = N and b<sub>j</sub>(m) is 1 if j = m and 0 otherwise
- in case of a hidden model, a ball could have been picked from any urn
- for the same observation sequence O, there may be many possible state sequences Q that could have generated O
- HMM: N states in model  $S = \{S_1, \dots, S_N\}_{(i)}$
- M distinct observation symbols in the alphabet  $V = \{v_1, \dots, v_M\}$
- transition probabilities A
- observation probabilities  $B = [b_j(m)]$
- initial state probabilities  $\Pi = [\pi_i]$



- given λ, the model can be used to generate an arbitrary number of observation sequences of arbitrary length
- interested in estimating the parameters of the model given a training set of sequences
- three basic problems of HMM: given a number of sequences of observations
  - given a model λ evaluate the probability of any given observation sequence O = {O<sub>1</sub>O<sub>2</sub>····O<sub>T</sub>}, namely, P(O|λ)
  - @ given a  $\lambda$  and O find out the  $Q = \{q_1 \cdots q_T\}$  which has the highest probability of generating O, find  $Q^*$  that maximizes  $P(Q|O,\lambda)$
  - given a training set of observation sequences, X = {O<sup>k</sup>}<sub>k</sub>, learn the model that maximizes the probability of generating X, namely, find λ\* that maximize P(X|λ)



evaluation problem

$$P(O|Q,\lambda) = \prod_{t=1}^{T} P(O_{t}|q_{t},\lambda) = b_{q_{1}}(O_{1}) \cdot b_{q_{2}}(O_{2}) \cdots b_{q_{T}}(O_{T})$$

- it can not be calculated because the state sequence is not known
- the probability of the state sequence Q

$$P(Q|\lambda) = P(q_1) \prod_{t=2}^{T} P(q_t|q_{t-1}) = \pi_{q_1} a_{q_1 q_2} \cdots a_{q_{T-1} q_T}$$

the joint probability is

$$P(O, Q|\lambda) = P(q_1) \prod_{t=2}^{T} P(q_t|q_{t-1}) \prod_{t=1}^{T} P(O_t|q_t)$$

$$= \pi_{q_1} b_{q_1}(O_1) a_{q_1 q_2} b_{q_2}(O_2) \cdots a_{q_{T-1} q_T} b_{q_T}(O_T)$$

 P(O|λ) can be computed by marginalizing over the joint, namely, by summing up over all possible Q

$$P(O|\lambda) = \sum_{\text{all possible } Q} P(O, Q|\lambda)$$

- this is not practical since there are N<sup>T</sup> possible Q, assuming that all the probabilities are nonzero
- efficient procedure to calculate P(O|λ) called forward backward procedure
- it is based on the idea of dividing the observation sequence into two parts: the first one starting from time 1 until time t, and the second one from time t + 1 until T
- define the forward variable  $\alpha_t(i)$  as the probability of observing the partial sequence  $\{O_1 \cdots O_t\}$  until time t and being in  $S_i$  at time t, given the model  $\lambda$

$$\alpha_t(i) \equiv P(O_1 \cdots O_t, q_t = S_i | \lambda)$$

- it can be calculated recursively
- initialization

$$\alpha_1 \equiv P(O_1, q_1 = S_i | \lambda)$$

$$= P(O_1 | q_1 = S_i, \lambda) P(q_1 = S_i | \lambda) = \pi_i b_i(O_1)$$

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Recursion

$$\alpha_{t+1}(j) \equiv P(O_1 \cdots O_{t+1}, q_{t+1} = S_j | \lambda)$$
  
=  $P(O_1 \cdots O_{t+1} | q_{t+1} = S_j, \lambda) P(q_{t+1} = S_j | \lambda)$ 

$$= P(O_1 \cdots O_t | q_{t+1} = S_j, \lambda) P(O_{t+1} | q_{t+1} = S_j, \lambda) P(q_{t+1} = S_j | \lambda)$$

$$= P(O_1 \cdots O_t, q_{t+1} = S_j | \lambda) P(O_{t+1} | q_{t+1} = S_j, \lambda)$$

$$= P(O_{t+1}|q_{t+1} = S_j, \lambda) \sum_i P(O_1 \cdots O_t, q_t = S_i, q_{t+1} = S_j|\lambda)$$

$$= P(O_{t+1}|q_{t+1} = S_j, \lambda)$$

$$\sum_{i} P(O_1 \cdots O_t, q_{t+1} = S_j|q_t = S_i, \lambda) P(q_t = S_i|\lambda)$$

$$= P(O_{t+1}|q_{t+1} = S_j, \lambda)$$

$$\sum_{i} P(O_1 \cdots O_t|q_t = S_i, \lambda) P(q_{t+1} = S_j|q_t = S_i, \lambda) P(q_t = S_i|\lambda)$$

$$= P(O_{t+1}|q_{t+1} = S_j, \lambda)$$

$$\sum_{i} P(O_1 \cdots O_t, q_t = S_i, \lambda) P(q_{t+1} = S_j|q_t = S_i, \lambda)$$

$$= \left[ \sum_{i=1}^{N} \alpha_t(i) a_{ij} \right] b_j(O_{t+1})$$

 $\alpha_t(i)$  explains the first t observations and ends in state  $S_i$  Multiply  $\alpha_t(i)$  by the probability  $a_{ii}$  to move to state  $S_i$ 

- there are N possible previous states, need to sum up over all such possible previous S<sub>i</sub>, b<sub>j</sub>(O<sub>t+1</sub>) then is the probability for the (t+1)st observation while in state S<sub>j</sub> at time t+1
- it is easy to calculate the probability of the observation sequence

$$P(O|\lambda) = \sum_{i=1}^{N} P(O, q_T = S_i|\lambda) = \sum_{i=1}^{N} \alpha_T(i)$$

- α<sub>T</sub>(i) is the probability of generating the full observation sequence and ending up in state S<sub>i</sub> and need to sum up over all such possible final states
- computing  $\alpha_t(i)$  is  $\mathcal{O}(N^2T)$  and solves the evaluation problem
- similarly backward variable  $\beta_t(i)$  is the probability of being in  $S_i$  at time t and observing the partial sequence  $O_{t+1} \cdots O_T$

$$\beta_t(i) \equiv P(O_{t+1} \cdots O_T | q_t = S_i, \lambda)$$

initialization (arbitrarily to 1)  $\beta_T(i) = 1$ Recursion

$$\beta_t(i) \equiv P(O_{t+1} \cdots O_T | q_t = S_i, \lambda)$$

$$= \sum_j P(O_{t+1} \cdots O_T, q_{t+1} = S_j | q_t = S_i, \lambda)$$

$$= \sum_{j} P(O_{t+1} \cdots O_T | q_{t+1} = S_j, q_t = S_i, \lambda) P(q_{t+1} = S_j | q_t = S_i, \lambda)$$

$$= \sum_{i} P(O_{t+1}|q_{t+1} = S_{j}, q_{t} = S_{i}, \lambda)$$

$$P(O_{t+2}\cdots O_T|q_{t+1}=S_j, q_t=S_i, \lambda)P(q_{t+1}=S_j|q_t=S_i, \lambda)$$



$$= \sum_{j} P(O_{t+1}|q_{t+1} = S_j, \lambda)$$

$$P(O_{t+2} \cdots O_T|q_{t+1} = S_j, \lambda) P(q_{t+1} = S_j|q_t = S_i, \lambda)$$

$$= \sum_{j=1}^{N} a_{ij} b_j (O_{t+1}) \beta_{t+1}(j)$$

• when in state  $S_i$ , we can go to N possible next states  $S_j$ , each with probability  $a_{ij}$ 



- finding the state sequence  $Q = \{q_1 q_2 \cdots q_T\}$  having the highest probability of generating the observation sequence  $O = \{O_1 O_2 \cdots O_T\}$  given the model  $\lambda$
- $\gamma_t(i)$  the probability of being in state  $S_i$  at time t given O and  $\lambda$

$$\gamma_t(i) \equiv P(q_t = S_i | O, \lambda) = \frac{P(O|q_t = S_i, \lambda)P(q_t = S_i | \lambda)}{P(O|\lambda)}$$

$$= \frac{P(O_1 \cdots O_t | q_t = S_i, \lambda) P(O_{t+1} \cdots O_T | q_t = S_i, \lambda) P(q_t = S_i | \lambda)}{\sum_{j=1}^{N} P(O_i, q_t = S_j | \lambda)}$$

$$= \frac{P(O_1 \cdots O_t, q_t = S_i | \lambda) P(O_{t+1} \cdots O_T | q_t = S_i, \lambda)}{\sum_{j=1}^{N} P(O_i | q_t = S_j, \lambda) P(q_t = S_j | \lambda)}$$

$$= \frac{\alpha_t(i) \beta_t(i)}{\sum_{j=1}^{N} \alpha_t(j) \beta_t(j)}$$



- $\alpha_t(i)$  and  $\beta_t(i)$  split the sequence,  $\alpha_t(i)$  explains the starting part of the sequence until time t and ends in  $S_i$  and the  $\beta_t(i)$  takes it from there and explains the ending part until time T
- $\alpha_t(i)\beta_t(i)$  explains the whole sequence given that at time t the system is in state  $S_i$ . It is normalized by dividing this over all possible intermediate states that can be traversed at time t, and guarantee that  $\sum_i \gamma_t(i) = 1$
- to find the state sequence, for each time step t, choose the state that has the highest probability

$$q_t^* = \underset{i}{\operatorname{arg\,max}} \gamma_t(i)$$

- this may choose S<sub>i</sub> and S<sub>j</sub> as the most probable states at time t and t+1 even when a<sub>ij</sub> = 0.
- to find the single best state sequence, the Viterbi algorithm is used