- an approach to the iterative computation of maximum likelihood (ML) estimates
- useful in a variety of incomplete-data problems, where algorithms such as the Newton-Raphson method may turn out to be more complicated
- it is applied to incomplete-data problems, where ML estimation is made difficult by the absence of some part of data in a more familiar and simple data structure
- on each iteration of the EM algorithm, there are two steps, called, the expectation step (E-step) and the maximization step (M-step)
- because of this, the algorithm is named as EM by Dempster



- it is closely related to the ad hoc approach to estimation with missing data,
- where the parameters are estimated after filling in initial values for the missing data
- the latter are then updated by their predicted values using these initial parameter estimates
- the parameters are then reestimated, and so on, proceeding iteratively until convergence



- the basic idea of the EM algorithm is to associate with the given incomplete-data problem, a complete-data problem for which ML estimation is computationally more tractable
- the methodology of the EM algorithm then consists in reformulating the problem in terms of this more easily solvable complete-data problem
- establishing a relationship between the likelihoods of these two problems, and
- exploiting the simpler MLE computation of the complete-data problem in the M-step of the iterative computing algorithm



Expectation Maximization Algorithm

- the E-step consists in manufacturing data for the complete-data problem, using the observed data set of the incomplete-data problem and the current value of the parameters, so that
- the simpler M-step computation can be applied to this "completed" data set
- more precisely, it is the log likelihood of the complete-data problem that is "manufactured" in the E-step
- as it is based partly on unobservable data, it is replaced by its conditional expectation given the observed data, where this E-step is effected using the current fit for the unknown parameters



- let, Y be the random vector corresponding to the observed data y, having p.d.f. postulated as g(y; Ψ), where
- $\boldsymbol{\Psi} = (\Psi_1, \dots, \Psi_d)^T$ is a vector of unknown parameters with parameter space $\boldsymbol{\Omega}$.
- ullet the vector $oldsymbol{\psi}$ can be estimated by maximum likelihood approach
- ullet the likelihood function for $oldsymbol{arphi}$ formed from the observed data $oldsymbol{y}$ is given by

$$L(\boldsymbol{\Psi}) = g(\boldsymbol{y}; \boldsymbol{\Psi})$$

ullet an estimate $\hat{oldsymbol{\psi}}$ of $oldsymbol{\psi}$ can be obtained as a solution of the likelihood equation

$$\partial L(\boldsymbol{\Psi})/\partial \boldsymbol{\Psi} = 0$$



or equivalently,

$$\partial \log L(\boldsymbol{\Psi})/\partial \boldsymbol{\Psi} = 0$$

- the notion of incomplete data comes where the complete data may contain some variables that are never observable in a data sense
- within this framework, x denotes the vector containing the augmented or so-called complete data
- let z be the vector containing the additional data, referred to as the unobservable or missing data
- the p.d.f. of the random vector \boldsymbol{X} corresponding to the complete-data vector \boldsymbol{x} is denoted by $g_c(\boldsymbol{x}; \boldsymbol{\Psi})$



$$\log L_c(\boldsymbol{\Psi}) = \log g_c(\boldsymbol{x}; \boldsymbol{\Psi})$$

- ullet formally, two sample spaces ${\mathcal X}$ and ${\mathcal Y}$ and a many-to-one mapping from ${\mathcal X}$ to ${\mathcal Y}$
- instead of observing the complete-data vector \mathbf{x} in \mathcal{X} , we observe the incomplete-data vector $\mathbf{y} = \mathbf{y}(\mathbf{x})$ in \mathcal{Y}
- it follows that

$$g(\mathbf{y}; \boldsymbol{\Psi}) = \int_{\mathcal{X}(\mathbf{y})} g_c(\mathbf{x}; \boldsymbol{\Psi}) d\mathbf{x}$$

where $\mathcal{X}(\mathbf{y})$ is the subset of \mathcal{X} determined by the equation $\mathbf{y} = \mathbf{y}(\mathbf{x})$



- the EM algorithm approaches the problem of solving the incomplete-data likelihood equation indirectly by proceeding iteratively in terms of the complete-data log likelihood function, $\log L_c(\Psi)$
- as it is unobservable, it is replaced by its conditional expectation given y, using the current fit for
- ullet let, $oldsymbol{\psi}^{(0)}$ be some initial value for $oldsymbol{\psi}$
- then on the first iteration, the E-step requires the calculation of

$$Q(\boldsymbol{\Psi}; \boldsymbol{\Psi}^{(0)}) = E_{\boldsymbol{\Psi}^{(0)}} \left\{ \log L_{c}(\boldsymbol{\Psi}) | \boldsymbol{y} \right\}$$

• the M-step requires the maximization of $Q(\Psi;\Psi^{(0)})$ with respect to Ψ over the parameter space Ω









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ullet choose $oldsymbol{\psi}^{(1)}$ such that

$$Q(\Psi^{(1)}; \Psi^{(0)}) \geq Q(\Psi; \Psi^{(0)})$$

for all $\Psi \in \Omega$

- the E and M steps are then carried out again, but this time with $\pmb{\psi}^{(0)}$ replaced by the current fit $\pmb{\psi}^{(1)}$
- \bullet on the (k+1)th iteration, the E and M steps are defined as follows:
 - **E-step** Calculate $Q(\Psi; \Psi^{(k)})$, where $Q(\Psi; \Psi^{(k)}) = E_{\Psi^{(k)}} \{ \log L_c(\Psi) | \mathbf{y} \}$
 - ▶ M-step Choose $\Psi^{(k+1)}$ to be any value of $\Psi \in \Omega$ that maximizes $Q(\Psi; \Psi^{(k)})$; that is

$$Q(\boldsymbol{\Psi}^{(k+1)}; \boldsymbol{\Psi}^{(k)}) \geq Q(\boldsymbol{\Psi}; \boldsymbol{\Psi}^{(k)})$$

for all $\Psi \in \Omega$



the E and M steps are alternated repeatedly until the difference

$$L(\boldsymbol{\Psi}^{(k+1)}) - L(\boldsymbol{\Psi}^{(k)})$$

changes by an arbitrarily small amount in the case of convergence of the sequence of likelihood values $\{L(\boldsymbol{\psi}^{(k+1)})\}$

 it is shown that the (incomplete-data) likelihood function L(♥) is not decreased after an EM iteration; that is,

$$L(\boldsymbol{\Psi}^{(k+1)}) \geq L(\boldsymbol{\Psi}^{(k)})$$

for
$$k = 0, 1, 2, ...$$

 hence convergence must be obtained with a sequence of likelihood values that are bounded above

ullet another way of expressing is that $oldsymbol{\psi}^{(k+1)}$ belongs to

$$\mathcal{M}(\boldsymbol{\Psi}^{(k)}) = \underset{\boldsymbol{\Psi}}{\operatorname{arg max}} Q(\boldsymbol{\Psi}; \boldsymbol{\Psi}^{(k)})$$

which is the set of points that maximize $Q(\Psi; \Psi^{(k)})$

- ullet it is not necessary to specify the exact mapping from ${\mathcal X}$ to ${\mathcal Y}$, nor
- the corresponding representation of the incomplete-data density g in terms of the complete-data density g_c
- all that is necessary is the specification of the complete-data vector x
 and the conditional density of X given the observed data vector y
- specification of this conditional density is needed in order to carry out the E-step



- the EM algorithm is modified to produce the maximum a posteriori (MAP) estimate in incomplete-data problems
- the computation of the MAP estimate in a Bayesian framework via the EM algorithm, corresponding to some prior density
- p(♥) for ♥ is described as follows
- let, the incomplete and complete data posterior densities for ♥ be given by
- $p(\Psi|y)$ and $p(\Psi|x)$, respectively
- ullet then the MAP estimate of $oldsymbol{\psi}$ is the value of $oldsymbol{\psi}$ that maximizes the log (incomplete-data) posterior density

$$\log p(\boldsymbol{\Psi}|\boldsymbol{y}) = \log L(\boldsymbol{\Psi}) + \log p(\boldsymbol{\Psi})$$

here $p(\cdot)$ is being used as a generic symbol for a p.d.f.



- the EM algorithm is implemented as follows to compute the MAP estimate
 - **E-step** on the (k+1)th iteration,
 - calculate the conditional expectation of the log complete-data posterior density given the observed data vector y, using the current MAP estimate \(\psi^{(k)}\) of \(\psi\)
 - that is, calculate

$$E_{\boldsymbol{\Psi}^{(k)}}\left\{\log p\left(\boldsymbol{\Psi}|\boldsymbol{x}\right)|\boldsymbol{y}\right\} = Q(\boldsymbol{\Psi};\boldsymbol{\Psi}^{(k)}) + \log p(\boldsymbol{\Psi})$$
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- ▶ M-step choose $\Psi^{(k+1)}$ to maximize over $\Psi^{(k)} \in \Omega$
- the E-step is effectively the same as seen earlier
- the M-step differs in that the objective function for the maximization process is equal to Q(Ψ; Ψ^(k)) augmented by the log prior density, log p(Ψ)
- the presence of this latter term as the result of the imposition of a Bayesian prior for

- the EM algorithm has several properties relative to other iterative algorithms such as Newton-Raphson and Fisher's scoring method for finding MLEs
- some of its advantages compared to its competitors are as follows:
 - the EM algorithm is numerically stable with each EM iteration increasing the likelihood (except at a fixed point of the algorithm)
 - under favourable conditions, the EM algorithm has reliable global convergence, i.e.
 - ▶ starting from an arbitrary point \(\psi^{(0)}\) in the parameter space, convergence is nearly always to a local maximizer
 - the EM algorithm is easily implemented, because
 - the E-step of each iteration only involves taking expectations over complete-data conditional distributions and
 - the M-step of each iteration only requires complete-data ML estimation, which is often in simple closed form



- the EM algorithm is easy to program, since no evaluation of the likelihood nor its derivatives is involved
- the cost per iteration is generally low, which can offset the larger number of iterations needed for the EM algorithm compared to other competing procedures
- by watching the monotone increase in likelihood over iterations, it is easy to monitor convergence
- the EM algorithm can be used to provide estimated values of the 'missing' data



- Some of the criticisms (disadvantages) of the EM algorithm are as follows:
 - the EM algorithm may converge slowly even in some seemingly inocuous problems and in problems where there is too much 'incomplete information'
 - in some problems, the E-step may be analytically intractable
 - the EM algorithm like the Newton-type methods does not guarantee convergence to the global maximum when there are multiple maxima
 - in this case, the estimate obtained depends upon the initial value



- in general, no optimization algorithm is guaranteed to converge to a global or local maximum, and
- the EM algorithm is not magical to this regard





Thank You



