

Probability and Statistics: Lecture-33

Monsoon-2020

by Dr. Pawan Kumar (IIIT, Hyderabad)
on October 28, 2020

» Conditional PDF, Conditional Probability, and Conditional CDF...

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$$P(X \in A | Y = y) = \int_A \underbrace{f_{X|Y}(x | y)}_{\text{conditional PDF}} dx$$

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3. The **conditional CDF** of X given $Y = y$ is

$$\underbrace{F_{X|Y}(x | y)} = P(X \leq x | Y = y) = \int_{-\infty}^x f_{X|Y}(x | y) dx$$

» Solved Example...

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Example (Solved Example)

Let X, Y be two jointly continuous RVs with joint PDF

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Let X, Y be two jointly continuous RVs with joint PDF

$$f_{XY} = \begin{cases} \frac{x^2}{4} + \frac{y^2}{5} + \frac{xy}{6} & 0 \leq x \leq 1, 0 \leq y \leq 2 \\ 0 & \text{otherwise} \end{cases}$$

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For $0 \leq y \leq 2$, find the following

1. The conditional PDF of X given $Y = y$

2. $P\left(X < \frac{1}{2} \mid Y = y\right)$

given

$$f_{X|Y}(x|y) = \frac{f_{XY}(x,y)}{f_Y(y)}$$

» Answer to previous problem...

We find the marginal PDF of y .

$$f_Y(y) = \int_0^1 \left(\frac{x^2}{4} + \frac{y^2}{4} + \frac{xy}{6} \right) dx$$
$$= \left[\frac{x^3}{3 \cdot 4} + \frac{y^2 x}{4} + \frac{x^2 y}{2 \cdot 6} \right]_0^1$$

$$= \frac{1}{12} + \frac{y^2}{4} + \frac{y}{12}$$

$$= \frac{3y^2 + y + 1}{12}$$

$$\text{for } 0 \leq y \leq 2$$

$$f_{X|Y}(x|y) = \frac{f_{XY}(x,y)}{f_Y(y)}$$

$$= \frac{\frac{x^2}{4} + \frac{y^2}{4} + \frac{xy}{6}}{\frac{3y^2 + y + 1}{12}}$$

$$= \frac{30x^2 + 24y^2 + 20xy}{24 \cdot 5 \cdot 8}$$

$$\frac{3y^2 + y + 1}{12}$$

$$= \frac{30x^2 + 24y^2 + 20xy}{12(3y^2 + y + 1)}$$
$$=$$

» Answer to previous problem...

$$\textcircled{b} \quad P(x < \frac{1}{2} | y = y)$$
$$= \int_0^{\frac{1}{2}} \underbrace{f_{x|y}(x|y)}_{\substack{\uparrow \\ \text{found}}} dx$$

↑
output is a fn of y

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1. **Expected value** of X given $Y = y$ is

$$E[X \mid Y = y] = \int_{-\infty}^{\infty} x f_{X|Y}(x \mid y) dx$$

2. **Conditional expectation** of function of RV

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$$E[\underbrace{g(X)} \mid Y = y] = \int_{-\infty}^{\infty} \underbrace{g(x)} \underbrace{f_{X|Y}(x \mid y)} dx$$

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$$E[g(X) \mid Y = y] = \int_{-\infty}^{\infty} g(x) f_{X|Y}(x|y) dx$$

3. **Conditional Variance** of X given $Y = y$ is

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$$\text{Var}(X \mid Y = y) = \underline{E[X^2 \mid Y = y]} - (E[X \mid Y = y])^2$$

» Solved Example on Conditional Expectation and Conditional Variance

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For $0 \leq y \leq 2$, find the following

1. Find $E[X | Y = 1]$ and $\text{Var}(X | Y = 1)$

» Answer to previous problem...

$$\underline{E[X|Y=1]} = \int_{-\infty}^{\infty} \underbrace{(x f_{X|Y}(x|y))}_{\uparrow \text{ found!}} \bigg|_{y=1} dx$$

$$= \int_0^1 \underbrace{(x f_{X|Y}(x|y))}_{\text{found!}} \bigg|_{y=1} dx$$

$$= \int_0^1 x \frac{\cancel{2x^2}^2}{\cancel{2x^2}^2} dx$$

$$\begin{aligned} &\hookrightarrow \text{Var}[X|Y=1] \\ &= \underline{E[X^2|Y=1]} - \left(\underline{E[X|Y=1]} \right)^2 \end{aligned}$$

$$E[X^2|Y=1]$$

$$= \int_{-\infty}^{\infty} x^2 f_{X|Y}(x|y) dx$$

$$= \int_0^1 x^2 f_{X|Y}(x|y) dx$$

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If X and Y are **independent**, we have

$$\begin{aligned} E[XY] &= E[X] E[Y] \\ E[g(X)h(Y)] &= E[g(X)] E[h(Y)] \end{aligned}$$

} easily
verified

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$$E[XY] = E[X] E[Y]$$

$$E[g(X)h(Y)] = E[g(X)] E[h(Y)]$$

* If we are given joint PDF of X and Y , $f_{XY}(x, y)$. If we can write

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$$f_{XY}(x, y) = f_1(x) f_2(y),$$

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If X and Y are **independent**, we have

$$E[XY] = E[X] E[Y]$$

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* If we are given joint PDF of X and Y , $f_{XY}(x, y)$. If we can write

$$f_{XY}(x, y) = \underbrace{f_1(x)} \underbrace{f_2(y)},$$

then X and Y are **independent**.

» Solved Example...

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$$f_{X,Y}(x,y) = f_X(x) f_Y(y)$$

Example (Solved Example)

Are the following RVs denoted by X and Y with the following PDF independent?

a) $f_{X,Y}(x,y) = \begin{cases} 2e^{-x-2y} & x, y > 0 \\ 0 & \text{otherwise} \end{cases}$ *yes, Independent*

b) $f_{X,Y}(x,y) = \begin{cases} 8xy & 0 < x < y < 1 \\ 0 & \text{otherwise} \end{cases}$

$$(8x)(y)$$

» Answer to previous problem...

Compute the marginals $f_X(x)$,
 $f_Y(y)$

$$\begin{aligned} f_X(x) &= \int_0^{\infty} 2e^{-x-2y} dy \\ &= 2e^{-x} \int_0^{\infty} e^{-2y} dy = 2e^{-x} \left[\frac{e^{-2y}}{-2} \right]_0^{\infty} \\ &= -e^{-x} [-1] = \boxed{e^{-x}} \end{aligned}$$

$$\begin{aligned} f_Y(y) &= \int_0^{\infty} 2e^{-x-2y} dx \\ &= 2e^{-2y} \int_0^{\infty} e^{-x} dx \\ &= 2e^{-2y} \left[\frac{e^{-x}}{-1} \right]_0^{\infty} \\ &= -2e^{-2y} [e^{-x}]_0^{\infty} \\ &= \boxed{2e^{-2y}} \\ \Rightarrow f_X(x) f_Y(y) &= f_{XY}(x, y) \end{aligned}$$

» Answer to previous problem...

$$\textcircled{b} \quad f_x(x) = \int_x^1 8xy \, dy$$

$$= 8x \left[\frac{y^2}{2} \right]_{\textcircled{x}}^1 = 4x \left[y^2 \right]_x^1 \\ = 4x(1-x^2)$$

$$f_Y(y) = \int_0^y 8xy \, dx$$

$$= 8y \int_0^{\textcircled{y}} x \, dx$$

$$= 8y \left[\frac{x^2}{2} \right]_0^y$$

$$= 4y^3$$

Here $f_x(x) \cdot f_Y(y)$

$$\neq f_{XY}(x,y)$$

$\Rightarrow x, y$ are not ind.

» Solved Example...

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Example (Solved Example)

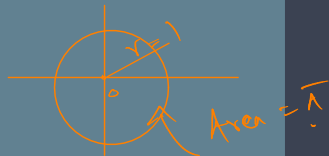
Consider a point (X, Y) chosen uniformly at random from the following disc

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$$D = \{(x, y) \mid x^2 + y^2 \leq 1\}$$



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The joint PDF of X and Y is given by

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$$D = \{(x, y) \mid x^2 + y^2 \leq 1\}$$

The joint PDF of X and Y is given by

$$f_{XY}(x, y) = \begin{cases} c & (x, y) \in D \\ 0 & \text{otherwise} \end{cases}$$

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1. Find the constant c

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2. Find the marginal PDFs $f_X(x)$ and $f_Y(y)$

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2. Find the marginal PDFs $f_X(x)$ and $f_Y(y)$
3. Find the conditional PDF of X given $Y = y$, where $-1 \leq y \leq 1$

» Solved Example...

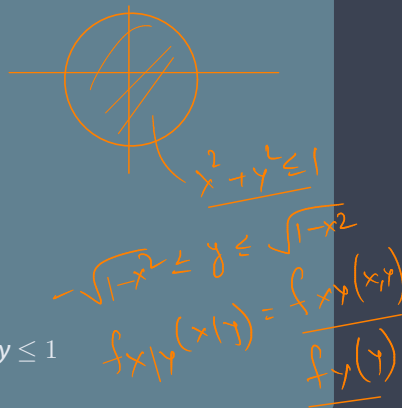
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Consider a point (X, Y) chosen uniformly at random from the following disc

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1. Find the constant c
2. Find the marginal PDFs $f_X(x)$ and $f_Y(y)$
3. Find the conditional PDF of X given $Y = y$, where $-1 \leq y \leq 1$
4. Are X and Y independent?

» Answer to previous problem...

① We have

$$\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{XY}(x, y) dx dy = 1$$

$$\Rightarrow \iint_D c dx dy = 1$$

$$\Rightarrow c \pi = 1$$

$$\Rightarrow c = 1/\pi //$$

$$\begin{aligned} \textcircled{2} \quad f_X(x) &= \int_{-\infty}^{\infty} f_{XY}(x, y) dy \\ &= \int_{-\sqrt{1-x^2}}^{\sqrt{1-x^2}} \frac{1}{\pi} dy = \frac{2\sqrt{1-x^2}}{\pi} \end{aligned}$$

Similarly, by symmetry

$$f_Y(y) = \frac{2\sqrt{1-y^2}}{\pi}$$

» Answer to previous problem...

$$\begin{aligned} \textcircled{3} \quad f_{X|Y}(x|y) &= \frac{f_{XY}(x,y)}{f_Y(y)} \\ &= \frac{1/\pi}{2\sqrt{1-y^2}} = \frac{1}{2\sqrt{1-y^2}} \end{aligned}$$

$$f_{X|Y}(x|y) = \begin{cases} \frac{1}{2\sqrt{1-y^2}} & \sqrt{1-y^2} \leq x \leq \sqrt{1-y^2} \\ 0 & \text{otherwise} \end{cases}$$

Note that given $y=y$,
 x is uniformly distributed
on $[-\sqrt{1-y^2}, \sqrt{1-y^2}]$
 $X|Y=y \sim \text{Unif}(-\sqrt{1-y^2}, \sqrt{1-y^2})$

④ No.

» Law of Total Probability...

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B_i 's
portion of sample

$$P(A) = \sum_{i=1}^K P(A|B_i) P(B_i)$$

\nwarrow \nwarrow \nwarrow PDF

Law of total probability

1. Law of Total Probability

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$$P(A) = \int_{-\infty}^{\infty} \underbrace{P(A \mid X = x)}_{\text{orange bracket}} f_X(x) dx$$


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$$P(A) = \int_{-\infty}^{\infty} P(A \mid X = x) f_X(x) dx$$

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$$P(A) = \int_{-\infty}^{\infty} P(A | X = x) f_X(x) dx$$

2. Law of Total Expectation

$$\underline{E[Y]} = \int_{-\infty}^{\infty} \underbrace{E[Y | X = x]}_{f_{X|Y}} \underbrace{f_X(x)}_{f_X(x)} dx = E[\underbrace{E[Y | X]}_{X}]$$

$Y|X \sim \text{Exponential}(\lambda)$
 $E[Y|X] = \lambda$

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Law of total probability

1. Law of Total Probability

$$P(A) = \int_{-\infty}^{\infty} P(A \mid X = x) f_X(x) dx$$

2. Law of Total Expectation

$$E[Y] = \int_{-\infty}^{\infty} E[Y \mid X = x] f_X(x) dx = E[E[Y \mid X]]$$

3. Law of Total Variance

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Law of total probability

1. Law of Total Probability

$$P(A) = \int_{-\infty}^{\infty} P(A \mid X = x) f_X(x) dx$$

2. Law of Total Expectation

$$E[Y] = \int_{-\infty}^{\infty} E[Y \mid X = x] f_X(x) dx = E[E[Y \mid X]]$$

3. Law of Total Variance

$$\text{Var}(Y) = \underline{E[\text{Var}(Y \mid X)]} + \underline{\text{Var}(E[Y \mid X])}$$

» Solved Example...

» Solved Example...

$$P(X \perp Y) = P(X) \quad \text{if } X, Y \text{ ind.}$$

Example (Example)

Let X, Y be two independent Uniform(0, 1) RVs. Find $P(X^3 + Y > 1)$.

$$P(X^3 + Y > 1) = \int_{-\infty}^{\infty} P(X^3 + Y > 1 \mid X=x) f_X(x) dx \quad (\text{Law of total probability})$$

$$= \int_0^1 P(X^3 + Y > 1 \mid X=x) \frac{f_X(x) dx}{\text{because } X \sim \text{Unif}(0,1)}$$

$$\begin{aligned} & \int_0^1 P(Y > 1 - x^3) \cdot 1 \cdot dx = \int_0^1 1 - P(Y \leq 1 - x^3) dx \\ & \text{because } Y \sim \text{Unif}(0,1) \quad \text{and } X, Y \text{ ind.} \\ & = \int_0^1 1 - (1 - x^3) dx = \left[\frac{x^4}{4} \right]_0^1 = \frac{1}{4} \end{aligned}$$

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Example (Solved example)

Let $X \sim \text{Uniform}(1, 2)$.

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$$Y \mid X \sim \text{Exponential}(X).$$

1. Find $E[Y]$

» Solved Example...

Example (Solved example)

Let $X \sim \text{Uniform}(1, 2)$. For a given $X = x$, Y is an exponential RV with parameter $\lambda = x$, that is

$$Y \mid X = x \sim \text{Exponential}(x).$$

That is

$$Y \mid X \sim \text{Exponential}(X).$$

1. Find $E[Y]$
2. Find $\text{Var}(Y)$

» Answer to previous problem...

$$a) E[Y] = \int_{-\infty}^{\infty} E[Y|X=x] \underbrace{f_X(x)}_{dx}$$

(by Law of Total Expectation)

$$= \int_1^2 E[Y|X=x] \cdot 1 \cdot dx$$

$$= \int_1^2 \frac{1}{x} dx = \ln \underline{\underline{2}}$$

Another way

$$E[Y] = E[E[Y|X]]$$

$$= E\left[\frac{1}{X}\right] = \int_1^2 \frac{1}{x} \cdot f_X(x) dx$$

$$= \int_1^2 \frac{1}{x} dx = \ln 2$$

$$b) \text{Var}(Y) = E[Y^2] - (E[Y])^2$$

we would prefer \leftarrow Law of total var.

$$\text{Var}(Y) = E[\text{Var}(Y|X)] + \text{Var}(E[Y|X])$$

» Answer to previous problem...

We have

$$\begin{aligned}\text{Var}(Y) &= E\left[\frac{1}{x^2}\right] + \text{Var}\left(\frac{1}{x}\right) \\&= E\left[\frac{1}{x^2}\right] + E\left[\frac{1}{x^2}\right] - \left(E\left[\frac{1}{x}\right]\right)^2 \\&= E\left[\frac{2}{x^2}\right] - (\ln 2)^2 \\&= \int_1^2 \frac{2}{x^2} \cdot f_X(x) dx \\&= \frac{1}{2}\end{aligned}$$

» Expectation of Function of Two RVs...

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Definition of Expectation of Function of RVs

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$$E[g(X, Y)] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} g(x, y) f_{XY}(x, y) dx dy$$

» Example of Expectation of Function of Two RVs...

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Example (Example)

Let X, Y be two jointly continuous RVs with joint PDF

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$$f_{XY}(x, y) = \begin{cases} x + y & 0 \leq x, y \leq 1 \\ 0 & \text{otherwise} \end{cases}$$

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Find $E[XY^2]$.

Handwritten notes:
 $= E[g(x, y)]$
 $E[XY^2] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} (xy^2) f_{XY}(x, y) dx dy = \int_0^1 \int_0^1 (xy^2)(x+y) dx dy$
 $= \int_0^1 \int_0^1 (x^2y^2 + xy^3) dx dy = \int_0^1 \left(\frac{y^2}{3} + \frac{y^3}{3} \right) dy = \dots$

» Computing CDF of Function of Two RVs...

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Example (Example)

Let X, Y be two independent $\text{Uniform}(0, 1)$ RVs, and $Z = XY$. Find the CDF and PDF of Z .