

Probability and Statistics: Lecture-40

Monsoon-2020

by Dr. Pawan Kumar (IIIT, Hyderabad)

on November 16, 2020

National exit poll

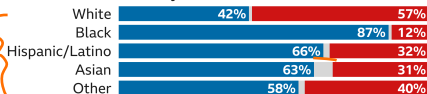
Support by gender, ethnicity, age group and education

■ Biden ■ Other ■ Trump

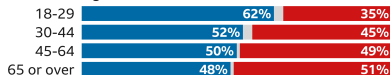
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Age



Education level



Sample size: 15,318 respondents

All figures have a margin of error which is wider for smaller sub-groups

Source: Edison Research/NEP via Reuters, 4 Nov, 17.00 EST (22.00 GMT)

BBC

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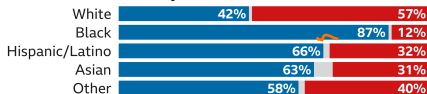
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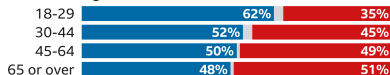
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- * On the left, US exit poll results
- * Poll on Trump Vs Biden
- * Sample size of 15,318
- * Error margin shown in grey
- * Draw conclusions from the sample data
- * Will inference fail? How much it can fail?
- * How confident we are of this?

	NDA	MAHAGATHBANDHAN	LJP	OTHERS
JAN KI BAAT	104	128	6	5
C-VOTER	116	120	1	6
MY AXIS	80	150	4	9
TV9 BHARATVARSH	115	120	4	4
POLL OF POLLS	104	129	4	6

Handwritten orange annotations on the table:

- Arrows pointing to the first column (poll names).
- Arrows pointing to the NDA column.
- Arrows pointing to the MAHAGATHBANDHAN column.
- Arrows pointing to the LJP column.
- Arrows pointing to the OTHERS column.
- A bracket grouping the MAHAGATHBANDHAN values (128, 120, 150, 120, 129).
- An arrow pointing to the value 80 in the MY AXIS row, NDA column.

POLL OF ALL POLLS				
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BIHAR ASSEMBLY ELECTIONS RESULTS 2020

TOTAL SEATS 243

NDA 125 **MGB 110** **OTH 8**

BJP	74	RJD	75	LJP	1
JD(U)	43	CONG	19	AIMIM	5
HAM	4	CPI-ML	11	BSP	1
VIP	4	CPM	3	OTHERS	1
		CPI	2		

- * On the left, poll of polls showing clear majority for MAHAGATHBANDHAN
- * After election, NDA has full majority
- * How do we estimate such errors?

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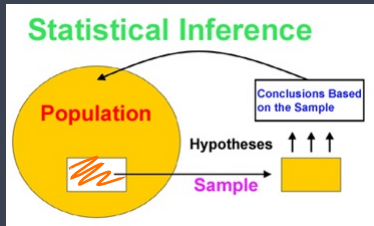
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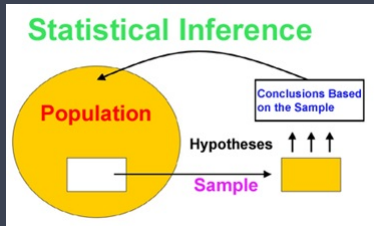


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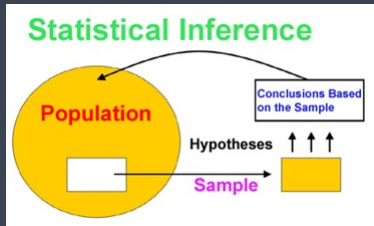
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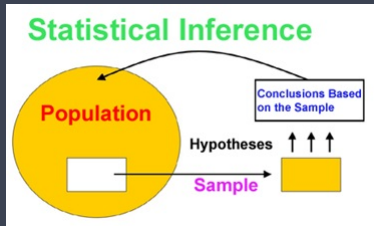
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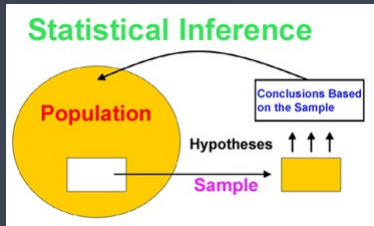
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Statistical Inference Problem

To determine an unknown quantity, get some data, and then estimate the required quantity using this data.

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 - * Here $\hat{\theta}$ is random variable, because it depends on random sample

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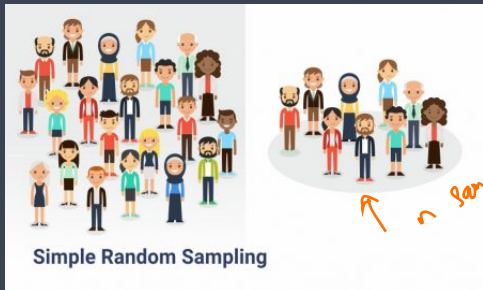
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- * We use the **prior** knowledge that $\Theta \sim \text{Bernoulli}(p)$

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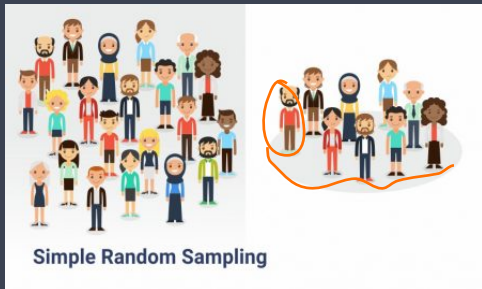


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 - * that is, working with **independently and identically distributed** makes analysis simpler

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$$\underbrace{F_{X_1}(x)} = \underbrace{F_{X_2}(x)} = \dots = \underbrace{F_{X_n}(x)}, \quad \text{for all } \underline{x \in \mathbb{R}}$$

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$$\bar{X} = \frac{X_1 + X_2 + \dots + X_n}{n}$$

» Recall: Properties of Sample Mean...

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converges in distribution to the standard normal random variable

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$$\lim_{n \rightarrow \infty} P(Z_n \leq x) = \Phi(x), \quad \text{for all } x \in \mathbb{R},$$

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Properties of sample mean, \bar{X}

1. $E[\bar{X}] = \mu, \quad \text{Var}(\bar{X}) = \frac{\sigma^2}{n}$

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$$f_{X_{(1)}, \dots, X_{(n)}}(x_1, \dots, x_n) = \begin{cases} n! f_X(x_1) f_X(x_2) \dots f_X(x_n) & \text{for } x_1 \leq x_2 \leq \dots \leq x_n \\ 0 & \text{otherwise} \end{cases}$$

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Let X_1, X_2, \dots, X_4 be a random variable from the Uniform(0,1) distribution, and let $X_{(1)}, X_{(2)}, X_{(3)}, X_{(4)}$ be the **order statistics** of X_1, X_2, \dots, X_4 .

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$$f_{X_{(2)}}(x) = \frac{4!}{(2-1)!(4-2)!} f_X(x) [F_X(x)]^{2-1} [1-F_X(x)]^{4-2}$$

» Answer to previous problem...

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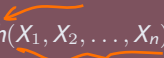
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» Unbiased Estimator is not Necessarily a Good Estimator...

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Fact

Show that unbiased estimator is **not** necessarily a good estimator.

Ex: x_1, x_2, \dots, x_n Random sample, $\theta = E[x_i] = E[x]$

$$\hat{\theta} = \bar{x} = \frac{x_1 + x_2 + \dots + x_n}{n}$$

If we choose $\hat{\theta}_1 = x_1$, then $\hat{\theta}_1$ is also an unbiased estimator of θ .
 $B(\hat{\theta}_1) = E[\hat{\theta}_1] - \theta = E[x_1] - \theta = \theta - \theta = 0$

Observe: $\hat{\theta}_1$ is probably not as good as sample mean \bar{x} .
need other measures to ensure that estimator is
"good" estimator.
Better: $E[(\hat{\theta} - \theta)^2]$

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» Answer to previous problem...

Solⁿ we have

$$\begin{aligned} \text{MSE}[\hat{\theta}_1] &= E[(\hat{\theta}_1 - \theta)^2] \\ &= E[(x_1 - E[x_1])^2] = \text{Var}(x_1) = \underline{\underline{\sigma^2}} \end{aligned}$$

$$\begin{aligned} \text{To find } \text{MSE}(\hat{\theta}_2) &= E[(\hat{\theta}_2 - \theta)^2] \\ &= E[(\bar{x} - \theta)^2] = \end{aligned}$$

$$\text{Var}(\bar{x} - \theta) + (E[\bar{x} - \theta])^2$$

$$\begin{aligned} &= \text{Var}(\bar{x}) + 0 \\ \Rightarrow \text{MSE}[\hat{\theta}_2] &= \text{Var}(\bar{x}) = \underline{\underline{\frac{\sigma^2}{n}}} \end{aligned}$$

As $n > 1$

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» Answer to previous problem...

» Relationship of MSE, Variance, and Bias...

Property

If $\hat{\theta}$ is a point estimator for θ ,

$$\text{MSE}(\hat{\theta}) = \text{Var}(\hat{\theta}) + B(\hat{\theta})^2$$

Pf.

$$\begin{aligned}\text{MSE}(\hat{\theta}) &= E[(\hat{\theta} - \theta)^2] \\ &= \text{Var}(\hat{\theta} - \theta) + \left(E[\hat{\theta} - \theta]\right)^2 \\ &= \text{Var}(\hat{\theta}) + B(\hat{\theta})^2\end{aligned}$$

Definition of Consistent Estimator

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Prf.

$$\lim_{n \rightarrow \infty} P(|\hat{\theta}_n - \theta| > \epsilon) = P(|\hat{\theta}_n - \theta|^2 > \epsilon^2) \leq \frac{E[(\hat{\theta}_n - \theta)^2]}{\epsilon^2} \quad [\text{Markov}]$$

$\rightarrow \lim_{n \rightarrow \infty} \frac{\text{MSE}[\hat{\theta}_n]}{\epsilon^2} = 0$ as $n \rightarrow \infty$

» Answer to previous problem...

» Definition of Sample Variance and Sample Standard Deviation...

Sample Variance and Sample Standard Deviation

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$$s^2 = \frac{1}{n-1} \sum_{k=1}^n (\underbrace{X_k - \bar{X}}_{\text{check}})^2 = \frac{1}{n-1} \left(\underbrace{\sum_{k=1}^n X_k^2}_{\text{check}} - \underbrace{n\bar{X}^2}_{\text{check}} \right)$$

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$$B = E[S^2 - \sigma^2] = 0$$

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$$S^2 = \frac{1}{n-1} \sum_{k=1}^n (X_k - \bar{X})^2 = \frac{1}{n-1} \left(\sum_{k=1}^n X_k^2 - n\bar{X}^2 \right)$$

We can check that sample variance is an unbiased estimator of σ^2 . The sample standard deviation is defined as

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We can check that sample variance is an unbiased estimator of σ^2 . The **sample standard deviation** is defined as

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and it is usually used as an estimator for σ . Also, S is an unbiased estimator of σ

Example (Sample Mean, Sample Variance, Sample Standard Deviation)

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Find the values of the sample mean, the sample variance, and the sample standard deviation for the observed sample.

Sample mean

$$\bar{T} = \frac{T_1 + \dots + T_6}{6} = \frac{18 + 21 + \dots + 20}{6}$$

Sample Variance

$$S^2 = \frac{1}{6-1} \sum_{k=1}^6 (T_k - \bar{T})^2 = \dots$$
$$S = \sqrt{S^2}$$

» Answer to previous problem...

Example

I have a bag that contains 3 balls.

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I have a bag that contains 3 balls. Each ball is either red or blue,

3: red
6: blue

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2. Find the value of θ that maximizes the probability of the observed sample

» Answer to previous problem...

Solⁿ Since $X_i \sim \text{Bernoulli}(\frac{\theta}{3})$.

$$P_{X_i}(x) = \begin{cases} \theta/3 & \text{for } x=1 \text{ (blue)} \\ 1-\theta/3 & \text{for } x=0 \text{ (red)} \end{cases}$$

Since X_i 's are ind., the joint PMF of X_1, X_2, X_3 and X_4 is

$$P_{X_1, X_2, X_3, X_4}(x_1, x_2, x_3, x_4)$$

$$= P_{X_1}(x_1) P_{X_2}(x_2) P_{X_3}(x_3) P_{X_4}(x_4)$$

This is called max. likelihood estimate of θ

$p = \frac{\theta}{3}$: prob. of picking blue.

$$\Rightarrow P_{X_1, X_2, X_3, X_4}(1, 0, 1, 1) = \frac{\theta}{3} \left(1 - \frac{\theta}{3}\right) \frac{\theta}{3} \frac{\theta}{3} \\ = \left(\frac{\theta}{3}\right)^3 \left(1 - \frac{\theta}{3}\right).$$

Joint PMF depends on θ .

Table θ	$P_{X_1, X_2, X_3, X_4}(1, 0, 1, 1; \theta)$
0	0
1	0.02 ←
→ 2	0.09 ←
3	0

$\theta = \underline{\underline{2}}$ max the probab.

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Definition of Likelihood and log likelihood Function

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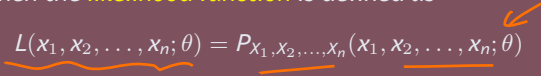
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$$\ln L(x_1, x_2, \dots, x_n; \theta)$$

» Example

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Find the likelihood function for the following random sample

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1. $X_i \sim \text{Binomial}(3, \theta)$ and we have observed $(x_1, x_2, x_3, x_4) = (1, 3, 2, 2)$
2. $X_i \sim \text{Exponential}(\theta)$ and we have observed $(x_1, x_2, x_3, x_4) = (1.23, 3.32, 1.98, 2.12)$

Solution. Recall that Random sample: x_1, x_2, x_3, x_4 are i.i.d.
 \Rightarrow the joint PMF & (PDF)
 $=$ product of marginal PMFs
 (4 PDFs respect.)

① If $x_i \sim \text{Binomial}(3, \theta)$,
 then $P_{x_i}(x_i; \theta) = \binom{3}{x_i} \theta^{x_i} (1-\theta)^{3-x_i}$

$$\begin{aligned} L(x_1, x_2, x_3, x_4; \theta) &= P_{x_1} P_{x_2} P_{x_3} P_{x_4}(x_1, x_2, x_3, x_4; \theta) \\ &= \frac{P_{x_1}}{\binom{3}{x_1}} \frac{P_{x_2}}{\binom{3}{x_2}} \frac{P_{x_3}}{\binom{3}{x_3}} \frac{P_{x_4}}{\binom{3}{x_4}} \theta^{x_1+x_2+x_3+x_4} \\ &= \frac{1}{\binom{3}{x_1} \binom{3}{x_2} \binom{3}{x_3} \binom{3}{x_4}} \theta^{12-(x_1+x_2+x_3+x_4)} (1-\theta)^{x_1+x_2+x_3+x_4} \end{aligned}$$

Since $(x_1, x_2, x_3, x_4) = (1, 3, 2, 2)$

$$\begin{aligned} \Rightarrow L(1, 3, 2, 2; \theta) &= \frac{1}{\binom{3}{1} \binom{3}{3} \binom{3}{2} \binom{3}{2}} \theta^8 (1-\theta)^4 \\ &= 27 \theta^8 (1-\theta)^4. \end{aligned}$$

» Answer to previous problem...

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Let $X_1, X_2, X_3, \dots, X_n$ be a random sample from a distribution with a parameter θ .

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» Maximum Likelihood Estimator...

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$$L(\underbrace{x_1, x_2, \dots, x_n}_{\text{data}}; \theta) = f(\theta)$$

A **maximum likelihood estimator (MLE)** of the parameter θ , shown by $\hat{\theta}_{ML}$ is a random variable $\hat{\theta}_{ML} = \hat{\theta}_{ML}(\underbrace{X_1, X_2, \dots, X_n}_{\text{data}})$ whose value when $X_1 = x_1, X_2 = x_2, \dots, X_n = x_n$ is given by $\hat{\theta}_{ML}$.

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For the following examples, find the **maximum likelihood estimator (MLE)** of θ :

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2. $X_i \sim \text{Exponential}(\theta)$ and we have observed X_1, X_2, \dots, X_n (1.23, 3.32, ...)

$$L(\underline{1, 3, 2, 2}; \theta) = 27 \theta^8 (1-\theta)^4$$
$$\frac{dL}{d\theta} = 27 \left[8\theta^7 (1-\theta)^4 - 4\theta^8 (1-\theta)^3 \right] = 0$$
$$\Rightarrow \hat{\theta}_{ML} = \frac{2}{3}$$

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Example (Example of maximum likelihood estimator)

Suppose that we have observed the random sample $X_1, X_2, X_3, \dots, X_n$, where $X_i \sim N(\theta_1, \theta_2)$ so

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Find the maximum likelihood estimators for θ_1 and θ_2 .

Better to use log likelihood.

- ① Find $L(\cdot)$
- ② Take $\ln L(\cdot)$
- ③ maximize $\ln L(\cdot)$

» Answer to previous problem...

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(CLT)

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2. Let $\hat{\Theta}_1$ be an estimator for θ such that $E[\hat{\Theta}_1] = \underline{a\theta + b}$, where $a \neq 0$. Show that

$$\hat{\Theta}_2 = \frac{\hat{\Theta}_1 - b}{a}$$

is an unbiased estimator for θ

ET possible