

Data Analytics for Data Scientists

Design of Experiments (DoE)

Lecture 09: Factorial designs

2025

Prof. Dr. Jürg Schwarz

Program: 16:15 until 17:55

| | |
|--------------|---|
| 16:15 | Begin of the lesson |
| | Lecture: Jürg Schwarz <ul style="list-style-type: none">◦ Design of experiments – A look back◦ Full factorial designs with control◦ Fractional factorial designs◦ Preview of Lecture 10 |
| | Tutorial: Students / Jürg Schwarz / Assistants <ul style="list-style-type: none">◦ Working on the exercise<ul style="list-style-type: none">◦ Support by Jürg Schwarz / Assistants |
| 17:55 | End of the lesson |

Designs of experiments – A look back

Factorial designs

Review of Lecture 03

Slide 3

Why conduct trials / experiments?

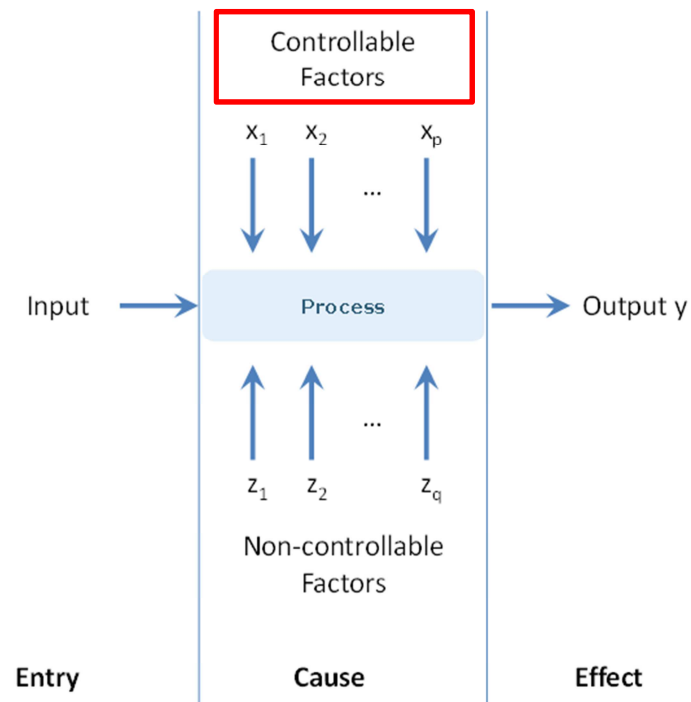
Cause and effect (causality)

A **trial / experiment** is carried out to discover a **cause-and-effect relationship** in a process.

The lady tasting tea

A lady declares that by tasting a cup of tea made with milk she can discriminate whether the milk or the tea infusion was first added to the cup. We will consider the problem of designating an experiment by means of which this assertion can be tested.

Fisher in *The Design of Experiments* (1935)



Review of Lecture 04

Slide 12

Summary

General DoE

Trial and error

- Changing many factors per run

DoE according to the principles of One-factor-at-a-time (OFAT)

One-factor-at-a-time (OFAT)

- Changing one factor per run

Principles of DoE

Design according to the principles of DoE

Full factorial designs – e.g. 2^k factorial designs

- The factor levels are determined before the experiment.
- All possible factor combinations are varied.
- Variants of full factorial designs → [Dean et al. \(2017\)](#)

Fractional factorial designs – e.g. 2^{k-1} factorial designs

- The factor levels are determined before the experiment.
- Only a (balanced) part of the possible combinations of factors are varied.
- Variants, for example «Latin Square» → [Dean et al. \(2017\)](#) / [Lecture 09](#)

Example of a **full factorial** design

Average dwell time on a website

How do factors IV1 and IV2 affect the dwell time DV, measured in seconds?

DV = dwell time [s] → output with metric scaling

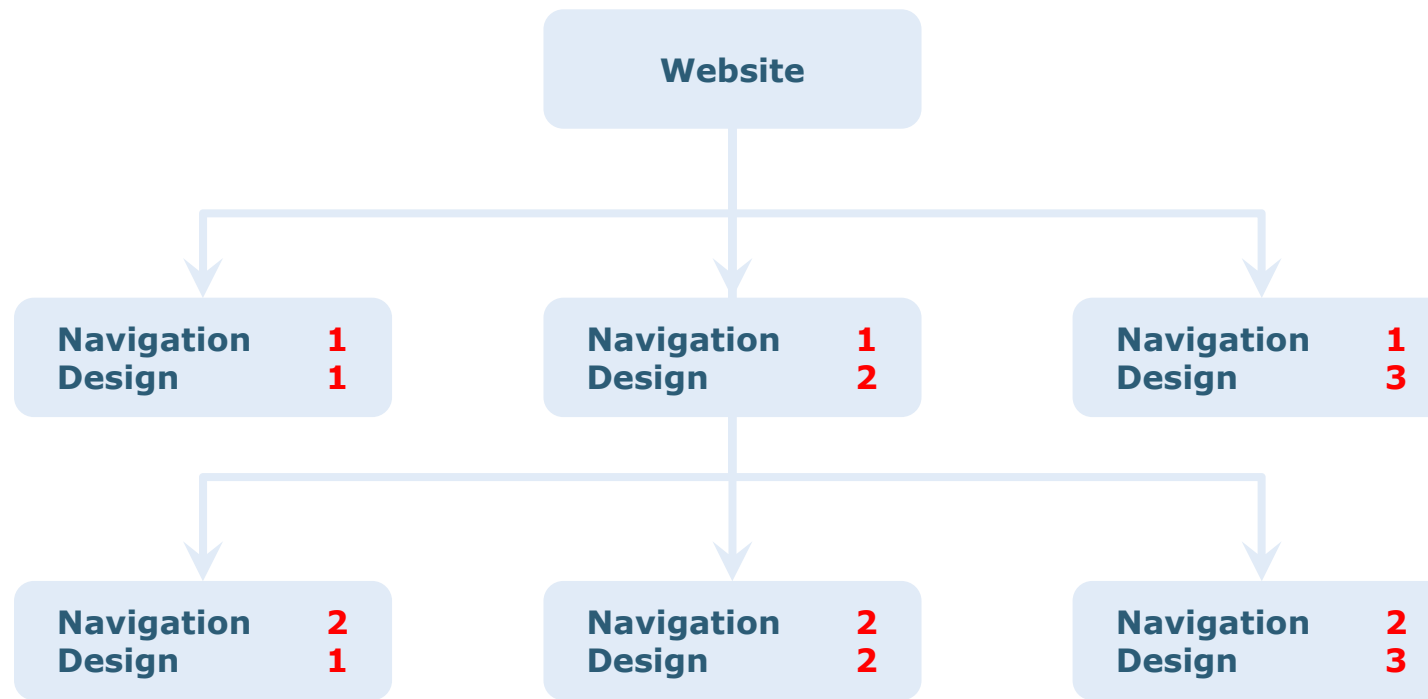
IV1 = design variants [1,2,3] → factor with 3 levels

IV2 = navigation variants [1,2] → factor with 2 levels

\bar{A}_{ij} = mean of the dwell time in group ij

| | | IV1 | | |
|-----|---|----------------|----------------|----------------|
| | | 1 | 2 | 3 |
| IV2 | 1 | \bar{A}_{11} | \bar{A}_{12} | \bar{A}_{13} |
| | 2 | \bar{A}_{21} | \bar{A}_{22} | \bar{A}_{23} |

Note: See also "Lecture 07: Addendum: Introduction to Analysis of Variance (ANOVA)"



2 Factors
6 Groups

$n_g = 384$
 $n = 2'304$

\bar{A}_{ij} = mean of the dwell time in group ij

| | | IV1 | | |
|-----|---|------|------|------|
| | | 1 | 2 | 3 |
| IV2 | 1 | 35.5 | 37.2 | 39.2 |
| | 2 | 36.9 | 40.3 | 45.0 |

ANOVA with data set *dwelltime* and R file *average_time*

Only main effects, no interaction

```
library(readxl)
dwelltime <- read_excel("<YOUR PATH>/dwelltime.xlsx")

fit <- aov(DV ~ factor(IV1) + factor(IV2), data = dwelltime)
summary(fit)
```

| | Df | Sum Sq | Mean Sq | F value | Pr(>F) | |
|-------------|------|--------|---------|---------|--------|-----|
| factor(IV1) | 2 | 13256 | 6628 | 80.71 | <2e-16 | *** |
| factor(IV2) | 1 | 6829 | 6829 | 83.16 | <2e-16 | *** |
| Residuals | 2300 | 188879 | 82 | | | |

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

There is a main effect of IV1 (levels 1, 2, 3) on DV, $F(2, 2300) = 80.71$, $p = .000$.

There is a main effect of IV2 (levels 1, 2) on DV, $F(1, 2300) = 83.16$, $p = .000$.

Design variants (1, 2, 3) and the navigation variants (1, 2) have a significant effect on dwell time.

Average dwell time **with interaction**

How do factors IV1 and IV2 affect the dwell time DV, measured in seconds?

DV = dwell time [s] → output with metric scaling

IV1 = design variants [1,2,3] → factor with 3 levels

IV2 = navigation variants [1,2] → factor with 2 levels

The two factors IV1 and IV2 could influence each other in their effect.

→ **Interaction** between IV1 and IV2

The combination of factor levels creates an autonomous effect.

For this purpose, a virtual third factor is «automatically» introduced in the model

→ **Interaction term** [in the R output → factor(IV1):factor(IV2)]

The number of factors and thus the number of groups remain the same →

2 Factors
6 Groups

ANOVA with data set *dwelltime* and R file *average_time***With interaction**

factor(IV1) + factor(IV2) → **without** interaction
 factor(IV1) * factor(IV2) → **with** interaction

```
library(readxl)
dwelltime <- read_excel("<YOUR PATH>/dwelltime.xlsx")

fit <- aov(DV ~ factor(IV1) * factor(IV2), data = dwelltime)
summary(fit)
```

| | Df | Sum Sq | Mean Sq | F value | Pr(>F) | |
|---|------|--------|---------|---------|----------|-----|
| factor(IV1) | 2 | 13256 | 6628 | 81.48 | < 2e-16 | *** |
| factor(IV2) | 1 | 6829 | 6829 | 83.95 | < 2e-16 | *** |
| factor(IV1):factor(IV2) | 2 | 1946 | 973 | 11.96 | 6.79e-06 | *** |
| Residuals | 2298 | 186933 | 81 | | | |
| --- | | | | | | |
| Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 | | | | | | |

There is a main effect of IV1 (levels 1, 2, 3) on DV, $F(2, 2298) = 81.48$, $p = .000$.

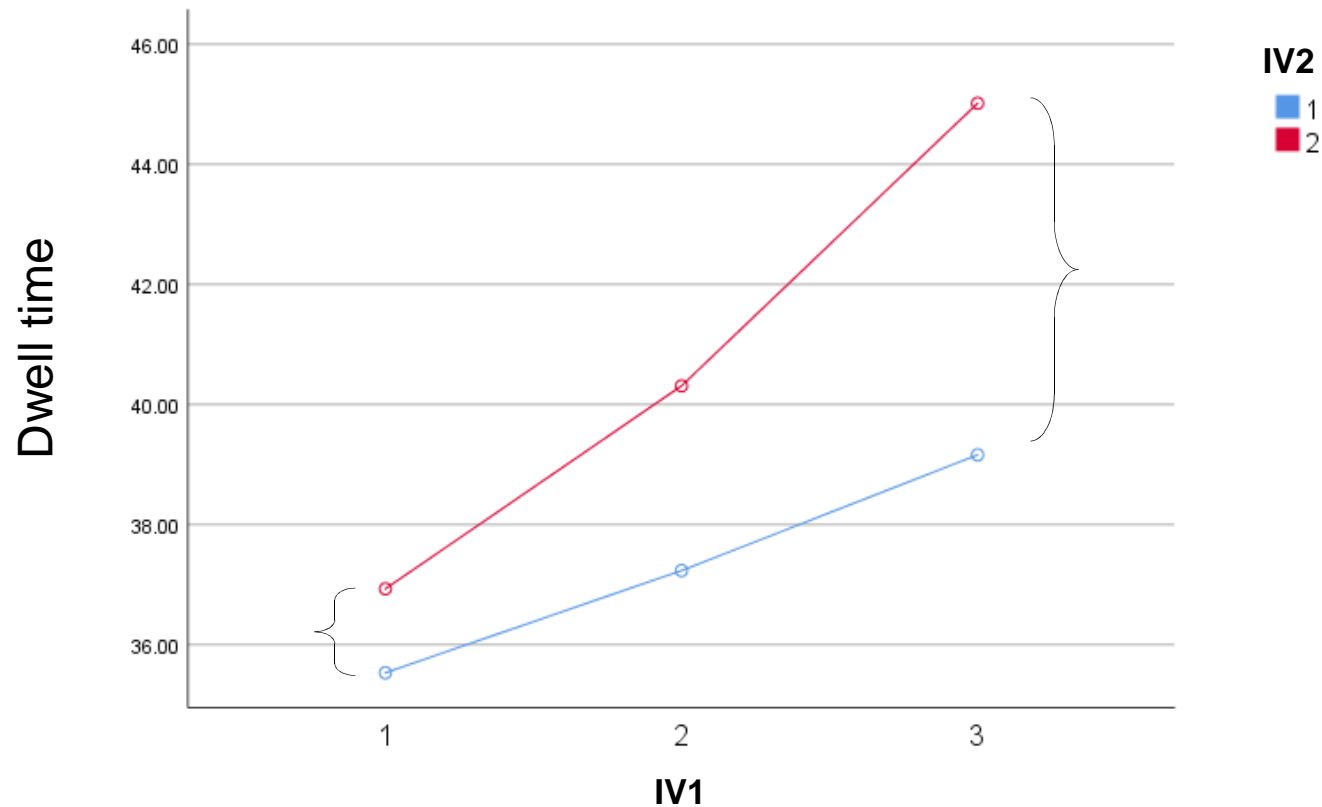
There is a main effect of IV2 (levels 1, 2) on DV, $F(1, 2298) = 83.95$, $p = .000$.

There is a significant interaction, $F(2, 2298) = 11.96$, $p = .000$.

Interpretation of the **interaction**

Do the levels of the design variants (IV1) and the navigation variants (IV2) influence each other?
Is the influence of one variable different depending on the characteristics of the other variable?

Yes, if IV1 takes the values 2 or 3, the influence of IV2 increases.



Simplified: Lines are not parallel

Interpretation: Navigation variant 2 has more influence in design variants 2 and 3.

Full factorial designs

Single factorial (One)

One independent variable **IV** (factor) acts on the dependent variable **DV**

Two groups are compared

1 IV → **one IV with two levels** (dichotomous)

1 DV → metric scaled

Examples: A/B test (→ [Lecture 08](#)), RCT with treatment and control or two treatments

Analysis: One-way ANOVA (also: t-test as a special case of univariate ANOVA)

One independent variable **IV** (factor) acts on the dependent variable **DV**

Several groups are compared

1 IV → **one IV with several levels**

1 DV → metric scaled

Examples: RCT with combinations of more than two treatments and control

Analysis: One-way ANOVA

Multifactorial (Several)

Several factors act on the dependent variable **DV**

Several groups are compared

X IV → **several IV with two or more levels each**

1 DV → metric scaled

Examples: Dwell time with two IV (→ [Slide 5](#))

Analysis: Multi-factorial ANOVA (two-way ANOVA, three-way ANOVA, ...)

Fractional factorial designs → [Slide 21ff](#)

In principle multifactorial (Several)

Several factors act on the dependent variable **DV**

Several groups are compared

X IV → **several IV with two or more levels each**

1 DV → metric scaled

BUT **Not all possible combinations are considered**

→ **Latin squares**, Greek-Latin squares, ...

→ Nesting of one IV under another IV → Hierarchical designs

IV
Independent variables

DV
Dependent variable

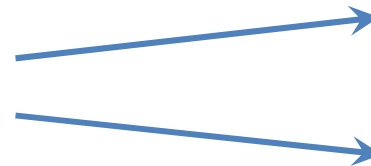
Single factorial

One IV

One DV

Two levels

1
2



A

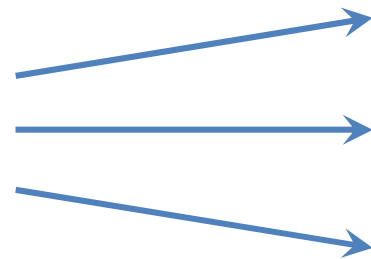
B

Divided into
two
groups

OR

Several levels

1
2
3
:



A

B

C

:

Several
groups

IV
Independent variables

DV
Dependent variable

Multifactorial

Several IV

One DV

Two levels

1
2

| |
|---|
| A |
| B |

| |
|---|
| A |
| B |

Divided into
two
groups

ADDITIONAL

Two levels

1
2

| |
|---|
| C |
| D |

| | | |
|---|----|----|
| | A | B |
| C | AC | BC |
| D | AD | BD |

Four
groups

ADDITIONAL

Several levels

1
2
3
:

| |
|---|
| E |
| F |
| G |

| | | |
|---|----|----|
| | A | B |
| C | AC | BC |
| D | AD | BD |

E F G

Several
groups

etc.

Full factorial designs with control

Review of Lecture 03

Slide 16

Control of the secondary variance

- Elimination / Keeping constant
- Repetition / Randomization / Blocking
- Nuisance variable as covariate

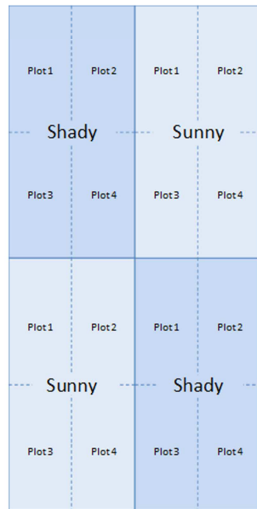
Methods for controlling secondary Variance (Individually or combined) ...

- Keeping the experimental setup constant ↔ Particularly possible in laboratory experiments
- Repetition
Several measurements are repeated on the same probands / trial objects.
- Randomization
Trial objects are assigned randomly to *Treatment* and *Control* groups to eliminate systematic bias.
- Blocking
Trial objects are grouped into homogeneous blocks based on one or more influential variables to reduce variability.
- Covariate adjustment
Nuisance variables are included as covariates in the statistical model to account for their effects.

Example of blocking: Experiment to study the effect of fertilizer on wheat yield.

Effect of two fertilizers (+/-) on the yield of two types of wheat (+/-).

Problem: The areas of the field varies in terms of solar radiation → nuisance factor



→ **Block factor**
Shady = 1
Sunny = 2

| Block | Plot | Fertilizer | Wheat |
|-------|------|------------|-------|
| 1 | 1 | + | + |
| 1 | 2 | - | - |
| 1 | 3 | + | - |
| 1 | 4 | - | + |
| 2 | 1 | - | - |
| 2 | 2 | + | + |
| 2 | 3 | - | - |
| 2 | 4 | + | + |

In each
with all
independ

Control by means of ...

◦ Nuisance variable as covariate

Conversion of a nuisance factor into an IV using the example of body growth in children

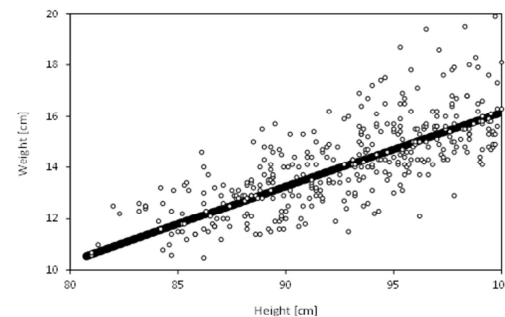
Primary variance is modeled using regression analysis.

The nuisance factor "sex" is included as an additional independent variable (control variable).

Model without sex as IV
Both sexes together

$$y = \beta_0 + \beta_1 \cdot x$$

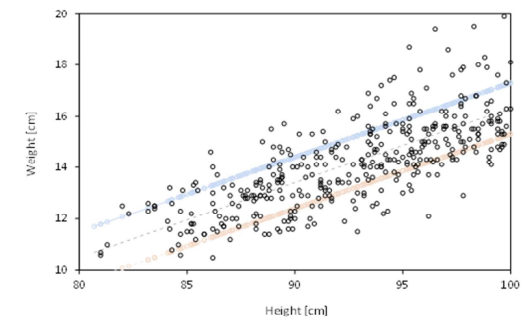
Model accuracy $R^2 = 0.557$



Model with sex as additional IV
Sexes separated

$$y = \beta_0 + \beta_1 \cdot x + \beta_2 \cdot \text{sex}$$

Model accuracy $R^2 = 0.601$



The relationship between height and body weight is estimated more precisely

↔ Secondary variance is controlled through the inclusion of sex as an additional IV.

Example of average dwell time – **with covariate**

How do factors IV1 and IV2 affect the dwell time DV, measured in seconds?

DV = dwell time [s] → output with metric scaling

IV1 = design variants [1,2,3] → factor with 3 levels

IV2 = navigation variants [1,2] → factor with 2 levels

The impact of an additional variable is taken into account:

IV3 = text amount [number of words] → **metric variable**

How can a metric variable be included in an ANOVA?

As a covariate

→ **ANCOVA** (analysis of **c**ovariance)

The number of factors and thus the number of groups remain the same →

2 Factors
6 Groups

ANOVA with data set *dwelltime* and R file *average_time***With covariate IV3**

Instead of entering the variable as a factor, enter it directly.

```
library(readxl)
dwelltime <- read_excel("<YOUR PATH>/dwelltime.xlsx")

options(contrasts = c("contr.sum", "contr.sum"))
fit <- aov(DV ~ factor(IV1) * factor(IV2) + IV3, data = dwelltime)
summary(fit)
```

| | Df | Sum Sq | Mean Sq | F value | Pr(>F) | |
|---|------|--------|---------|---------|----------|-----|
| factor(IV1) | 2 | 13256 | 6628 | 81.644 | < 2e-16 | *** |
| factor(IV2) | 1 | 6829 | 6829 | 84.119 | < 2e-16 | *** |
| IV3 | 1 | 1177 | 1177 | 14.496 | 0.000144 | *** |
| factor(IV1):factor(IV2) | 2 | 1228 | 614 | 7.562 | 0.000533 | *** |
| Residuals | 2297 | 186475 | 81 | | | |
| --- | | | | | | |
| Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 | | | | | | |

There is a main effect of IV1 (levels 1, 2, 3) on DV, $F(2, 2297) = 81.644$, $p = .000$.

There is a main effect of IV2 (levels 1, 2) on DV, $F(1, 2297) = 84.119$, $p = .000$.

There is a significant interaction, $F(2, 2297) = 7.562$, $p = .000$.

The **covariate IV3** has a significant impact, $F(1, 2297) = 14.496$, $p = .000$.

Example of average dwell time – **with block factor**

How do factors IV1 and IV2 affect the dwell time DV, measured in seconds?

DV = dwell time [s] → output with metric scaling

IV1 = design variants [1,2,3] → factor with 3 levels

IV2 = navigation variants [1,2] → factor with 2 levels

The impact of an additional variable is taken into account:

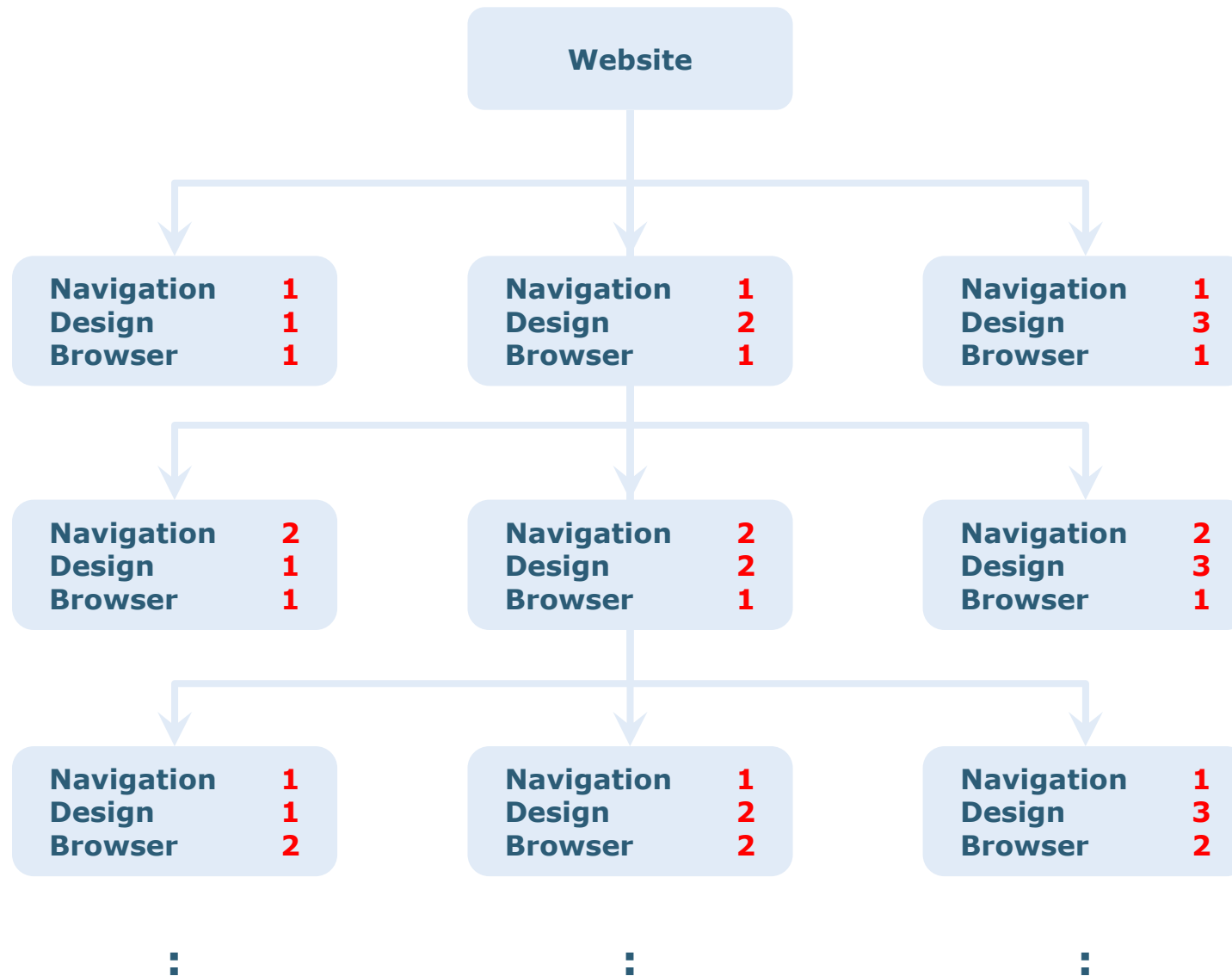
IV4 = Browser type [1,2,3] → **factor with 3 levels**

The impact of the additional variable IV4 is not part of the research question.

Therefore, although IV4 is introduced as a factor – as a **block factor** – it is not interpreted.

Note: ANOVA does not distinguish between factor and block factor.

3 Factors
18 Groups



Fractional factorial designs

Reduction of the number of groups

The more factors and the more characteristics, the larger the number of groups.

Question: How can the number of groups be reduced?

Example: A design for three factors IV1, IV2, IV3 with three levels each [1, 2, 3]

| IV1 | IV2 | IV3 |
|-----|-----|-----|
| 1 | 1 | 1 |
| 1 | 1 | 2 |
| 1 | 1 | 3 |
| 1 | 2 | 1 |
| 1 | 2 | 2 |
| 1 | 2 | 3 |
| 1 | 3 | 1 |
| 1 | 3 | 2 |
| 1 | 3 | 3 |
| 2 | 1 | 1 |
| 2 | 1 | 2 |
| 2 | 1 | 3 |

⋮

$3 \times 3 \times 3 = 27$ groups must be included

Provided that there are **no interactions**, the experiment can also be conducted successfully with a reduced number of groups.

By using **Latin squares** or related designs, the number of groups required can be significantly reduced compared to full factorial designs.

Latin square

Introduction

A Latin squares design makes it possible to study the **main effects of factors** without having to observe all combinations of treatment levels.

Compared to the full factorial design, the Latin squares design requires **smaller sample sizes**.

A Latin squares design can only be used if it follows from theory or empirical evidence that the joint effect of the factors **does not produce interactions**.

That is, **interactions** between factors **cannot be verified** by a Latin squares design.

Accordingly, if interactions are actually present, the main effects are not clearly interpretable.

Structure of a Latin square – Examples with three factors

Three factors, which must have the **same number of factor levels**.

Example 1: Factors A, B, C, each with **two factor levels** (a_1, a_2), (b_1, b_2), (c_1, c_2) $\rightarrow p = 2$

| | a_1 | a_2 |
|-------|-------|-------|
| b_1 | c_1 | c_2 |
| b_2 | c_2 | c_1 |

The factor level combination a_1b_1 is combined with c_1 , a_2b_1 with c_2 , a_1b_2 with c_2 and a_2b_2 with c_1 .

In an experiment, each of the four combinations is assigned a random sample of size n .

Example 2: Factors A, B, C, each with **three factor levels** (a_1, a_2, a_3), (b_1, b_2, b_3), (c_1, c_2, c_3)

→ **$p = 3$**

| | a_1 | a_2 | a_3 |
|-------|-------|-------|-------|
| b_1 | c_1 | c_2 | c_3 |
| b_2 | c_2 | c_3 | c_1 |
| b_3 | c_3 | c_1 | c_2 |

The factor level combination a_1b_1 is combined with c_1 , a_2b_1 with c_2 , a_3b_1 with c_3 , a_1b_2 with c_2 etc.

In an experiment, each of the nine combinations is assigned a random sample of size n .

Construction rules

In each row and in each column, each C-level appears **only once**.

For **$p = 3$** , there are **twelve** arrangements that satisfy this condition.

Two further examples

| | a_1 | a_2 | a_3 |
|-------|-------|-------|-------|
| b_1 | c_3 | c_1 | c_2 |
| b_2 | c_2 | c_3 | c_1 |
| b_3 | c_1 | c_2 | c_3 |

| | a_1 | a_2 | a_3 |
|-------|-------|-------|-------|
| b_1 | c_2 | c_1 | c_3 |
| b_2 | c_1 | c_3 | c_2 |
| b_3 | c_3 | c_2 | c_1 |

However, in the case of $p = 3$, there is only one arrangement in which the C-levels in the first row and the first column appear in a natural sequence (c_1, c_2, c_3) → **Standard form** (see above)

Balancing and decomposition

Balancing = Balance in the combination of factor levels

Decomposition = Breaking down a complete design into smaller, balanced subsets

Full* Each level of C is combined **only once** with each level of A and with each level of B.
Factorial The design is **completely balanced** in terms of the **main effects**.

Fractional** Each level of C occurs only partially with other combinations of A and B.
Factorial The Latin square is **only partially balanced with respect to first-order interaction**.

Relationship between a complete design **C** and a Latin square design **L** (balanced form)

| a ₁ | | | | | | | | | a ₂ | | | | | | | | | a ₃ | | | | | | | | | |
|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|----------------|---|---|--|
| b ₁ | | | b ₂ | | | b ₃ | | | b ₁ | | | b ₂ | | | b ₃ | | | b ₁ | | | b ₂ | | | b ₃ | | | |
| C | c ₁ | c ₂ | c ₃ | c ₁ | c ₂ | c ₃ | c ₁ | c ₂ | c ₃ | c ₁ | c ₂ | c ₃ | c ₁ | c ₂ | c ₃ | c ₁ | c ₂ | c ₃ | c ₁ | c ₂ | c ₃ | c ₁ | c ₂ | c ₃ | | | |
| | ↑ | | | | ↑ | | | | ↑ | | ↑ | | | ↑ | ↑ | | | | | ↑ | ↑ | | | | ↑ | | |
| L | 1 | | | | 2 | | | | 3 | | 4 | | | | 5 | 6 | | | | | 7 | 8 | | | | 9 | |

Arrows point to the factor level combinations in the Latin square → [Slide 23](#)

| | | | | |
|-------------------------|-------------------|---|--------------------------------------|-----------------------------------|
| C Full factorial | → 27 combinations | } | Ratio = $\frac{9}{27} = \frac{1}{3}$ | General***: Ratio = $\frac{1}{p}$ |
| L Latin square | → 9 combinations | | | |

Shifting the arrows one unit to the right (*modulo*) results in further factor level combinations, which in turn produce other Latin squares.

→ A complete 3×3×3 design can be divided into **three** balanced Latin squares.

Other fractional factorial designs / other methods

Greco-Latin square

Extension of the Latin square to **four factors**.

Counterbalancing in case of repeated measures

Problem → influence of a treatment on the result in one of the following treatments.

For example learning and fatigue effects also called «carry-over» effect.

Solution «counterbalancing» → conducting as many test sequences as needed to ensure that each stage of the treatment occurs equally frequently at each stage of the study.

More → [Dean et al. \(2017\)](#)

More → [Oehlert \(2010\)](#)

For example

| | | |
|----------|--|------------|
| C | Experimental Design Plans | 607 |
| C.1 | Latin Squares | 607 |
| C.1.1 | Standard Latin Squares | 607 |
| C.1.2 | Orthogonal Latin Squares | 608 |
| C.2 | Balanced Incomplete Block Designs | 609 |
| C.3 | Efficient Cyclic Designs | 615 |
| C.4 | Alpha Designs | 616 |
| C.5 | Two-Series Confounding and Fractioning Plans | 617 |

Preview of Lecture 10

What has happened so far

We looked at full factorial designs of experiments in more detail by using an example.

We also looked at examples of full factorial designs with control.

For full factorial designs, the complexity and sample size can increase very quickly, resulting in a need for fractional factorial designs.

A classic is the implementation with the Latin square.

What follows in Lecture 10

First, the question occurs «What are large data quantities?»

Specific properties of large data quantities become apparent, which have to be considered more closely in the context of design of experiments and statistics.

Possible solutions and procedures are discussed to deal with the problems associated with large data quantities.

Appendix

Details of a research project

Research question

How do factors A, B and C affect the dependent variable (DV) dwell time, measured in seconds?

| | | |
|----|---|------------------------------|
| DV | = dwell time [s] | → output with metric scaling |
| A | = design variants [a_1, a_2, a_3] | → factor with 3 levels |
| B | = navigation variants [b_1, b_2, b_3] | → factor with 3 levels |
| C | = browser type [c_1, c_2, c_3] | → factor with 3 levels |

Sample size

Based on power analysis, 1,200 persons are collected for each combination of factors.

| | | |
|-----------------|-------------------|------------------------------------|
| Complete design | → 27 combinations | $27 \times 1,200 = 32,400$ persons |
| Latin square | → 9 combinations | $9 \times 1,200 = 10,800$ persons |

Experimental design

Creating a Latin square with R script: ##### EXERCISE 09: Factorial Designs

```
library(agricolae)
```

```
my_design_1sd <- design.1sd(trt = c("c1","c2","c3"))
my_design_1sd$sketch
```

trt = short for "Treatments"

| | [,1] | [,2] | [,3] |
|------|--------------------------|------|------|
| [1,] | Output | | |
| [2,] | This is part of the task | | |
| [3,] | in the exercise | | |

Interpretation

This is part of the task in the exercise

Sampling

A sample with $n = 1,200$ is drawn for each of the combinations This is part of the task in the exercise

| | 1 | 2 | 3 | 4 |
|---|------|---|---|---|
| 1 | AV | A | B | C |
| 2 | 43.7 | 1 | 1 | 1 |
| 3 | 28.7 | 1 | 1 | 1 |
| 4 | 40.0 | 1 | 1 | 1 |
| 5 | 41.4 | 1 | 1 | 1 |
| 6 | 29.7 | 1 | 1 | 1 |
| 7 | 48.5 | 1 | 1 | 1 |
| 8 | 34.7 | 1 | 1 | 1 |

dwelltime_Latin.xlsx

Statistical data analysis

ANOVA with data set *dweltime_Latin* and R script: ##### EXERCISE 09: Factorial Designs

```
library(readxl)
dweltime_Latin <- read_excel("<YOUR PATH>/dweltime_Latin.xlsx")
View(dweltime_Latin)

fit <- aov(DV ~ factor(A) + factor(B) + factor(C), data = dweltime_Latin)
summary(fit)
```

| | Df | Sum Sq | Mean Sq | F value | Pr(>F) | |
|---|-------|--------|---------|---------|--------|-----|
| factor(A) | 2 | 76057 | 38028 | 465.761 | <2e-16 | *** |
| factor(B) | 2 | 27159 | 13580 | 166.320 | <2e-16 | *** |
| factor(C) | 2 | 552 | 276 | 3.383 | 0.034 | * |
| Residuals | 10338 | 844075 | 82 | | | |
| --- | | | | | | |
| Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1 | | | | | | |

Interpretation

This is part of the task in the exercise

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