

MATH 503: Mathematical Statistics

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Homework 9 Solutions

1. Let $X_{1j}, X_{2j}, \dots, X_{a_j j}$ represent independent random samples of sizes a_j from a normal distribution with means μ_j and variances σ^2 , $j = 1, 2, \dots, b$. Show that

$$\sum_{j=1}^b \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{..})^2 = \sum_{j=1}^b \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{.j})^2 + \sum_{j=1}^b a_j (\bar{X}_{.j} - \bar{X}_{..})^2,$$

or $Q' = Q'_3 + Q'_4$. Here,

$$\bar{X}_{..} = \frac{\sum_{j=1}^b \sum_{i=1}^{a_j} X_{ij}}{\sum_{j=1}^b a_j} \text{ and } \bar{X}_{.j} = \frac{\sum_{i=1}^{a_j} X_{ij}}{a_j}.$$

If $\mu_1 = \mu_2 = \dots = \mu_b$, show that Q'/σ^2 and Q'_3/σ^2 have chi-square distributions. Note that Q'_3 and Q'_4 are independent, and hence Q'_4/σ^2 also has a chi-square distribution.

Solution:

$$\begin{aligned} \sum_{j=1}^b \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{..})^2 &= \sum_{j=1}^b \sum_{i=1}^{a_j} [(X_{ij} - \bar{X}_{.j}) + (\bar{X}_{.j} - \bar{X}_{..})]^2 \\ &= \sum_{j=1}^b \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{.j})^2 + \sum_{j=1}^b \sum_{i=1}^{a_j} (\bar{X}_{.j} - \bar{X}_{..})^2 + 2 \sum_{j=1}^b \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{.j})(\bar{X}_{.j} - \bar{X}_{..}), \end{aligned}$$

where

$$\sum_{j=1}^b \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{.j})(\bar{X}_{.j} - \bar{X}_{..}) = \sum_{j=1}^b (\bar{X}_{.j} - \bar{X}_{..}) \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{.j}) = \sum_{j=1}^b (\bar{X}_{.j} - \bar{X}_{..}) \underbrace{\left(\sum_{i=1}^{a_j} X_{ij} - a_j \bar{X}_{.j} \right)}_{=0}$$

Thus,

$$\begin{aligned} Q' = \sum_{j=1}^b \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{..})^2 &= \sum_{j=1}^b \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{.j})^2 + \sum_{j=1}^b \sum_{i=1}^{a_j} (\bar{X}_{.j} - \bar{X}_{..})^2 \\ &= \sum_{j=1}^b \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{.j})^2 + \sum_{j=1}^b (\bar{X}_{.j} - \bar{X}_{..})^2 \sum_{i=1}^{a_j} 1 \\ &= \sum_{j=1}^b \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{.j})^2 + \sum_{j=1}^b a_j (\bar{X}_{.j} - \bar{X}_{..})^2 = Q'_3 + Q'_4. \end{aligned}$$

Note that

$$\frac{Q'}{\sigma^2} = \frac{\sum_{j=1}^b \sum_{i=1}^{a_j} (X_{ij} - \bar{X}_{..})^2}{\sigma^2} = \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \bar{X}_{..}}{\sigma} \right)^2,$$

where

$$\begin{aligned} \chi_{\sum_{j=1}^b a_j}^2 &= \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \mu}{\sigma} \right)^2 = \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{(X_{ij} - \bar{X}_{..}) + (\bar{X}_{..} - \mu)}{\sigma} \right)^2 \\ &= \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \bar{X}_{..}}{\sigma} \right)^2 + \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{\bar{X}_{..} - \mu}{\sigma} \right)^2 + 2 \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \bar{X}_{..}}{\sigma} \right) \left(\frac{\bar{X}_{..} - \mu}{\sigma} \right), \end{aligned}$$

where

$$\begin{aligned} \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \bar{X}_{..}}{\sigma} \right) \left(\frac{\bar{X}_{..} - \mu}{\sigma} \right) &= \left(\frac{\bar{X}_{..} - \mu}{\sigma} \right) \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \bar{X}_{..}}{\sigma} \right) \\ &= \left(\frac{\bar{X}_{..} - \mu}{\sigma} \right) \left[\frac{1}{\sigma} \left(\underbrace{\sum_{j=1}^b \sum_{i=1}^{a_j} X_{ij} - \bar{X}_{..} \sum_{j=1}^b a_j}_{=0} \right) \right], \end{aligned}$$

so

$$\chi_{\sum_{j=1}^b a_j}^2 = \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \bar{X}_{..}}{\sigma} \right)^2 + \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{\bar{X}_{..} - \mu}{\sigma} \right)^2,$$

where

$$\sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{\bar{X}_{..} - \mu}{\sigma} \right)^2 = \left(\frac{\bar{X}_{..} - \mu}{\sigma} \right)^2 \sum_{j=1}^b a_j = \left(\frac{\bar{X}_{..} - \mu}{\sigma / \sqrt{\sum_{j=1}^b a_j}} \right)^2 \sim \chi_1^2,$$

thus we see that $\frac{Q'}{\sigma^2} = \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \bar{X}_{..}}{\sigma} \right)^2 \sim \chi_{(\sum_{j=1}^b a_j) - 1}^2$. Similarly,

$$\begin{aligned} \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \mu}{\sigma} \right)^2 &= \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \bar{X}_{.j} + \bar{X}_{.j} - \mu}{\sigma} \right)^2 \\ &= \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \bar{X}_{.j}}{\sigma} \right)^2 + \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{\bar{X}_{.j} - \mu}{\sigma} \right)^2 \\ &= \sum_{j=1}^b \sum_{i=1}^{a_j} \left(\frac{X_{ij} - \bar{X}_{.j}}{\sigma} \right)^2 + \sum_{j=1}^b a_j \left(\frac{\bar{X}_{.j} - \mu}{\sigma} \right)^2 = \frac{Q'_3}{\sigma} + \chi_b^2, \end{aligned}$$

thus $\frac{Q'_3}{\sigma} \sim \chi_{(\sum_{j=1}^b a_j) - b}^2$.

Finally, since $\frac{Q'}{\sigma} = \frac{Q'_3}{\sigma} + \frac{Q'_4}{\sigma}$ where $\frac{Q'_4}{\sigma} \sim \chi_{(\sum_{j=1}^b a_j) - 1}^2$ and $\frac{Q'_3}{\sigma} \sim \chi_{(\sum_{j=1}^b a_j) - b}^2$, this implies that $\frac{Q'_4}{\sigma} \sim \chi_{(\sum_{j=1}^b a_j) - 1 - [(\sum_{j=1}^b a_j) - b]}^2 = \chi_{b-1}^2$.

2. Solve the following using R or SAS: The following are observations associated with independent random samples from three normal distributions having equal variances and respective means μ_1, μ_2, μ_3 .

I	II	III
0.5	2.1	3.0
1.3	3.3	5.1
-1.0	0.0	1.9
1.8	2.3	2.4
	2.5	4.2
		4.1

Compute the F-statistic that is used to test $H_0 : \mu_1 = \mu_2 = \mu_3$.

As shown in the below output, $F=6.388$, whether determined via R or SAS.

R code and output:

```
> data <- read.table("C:/courses/MATH 503/data/Exercise9.2.4data.txt",header=TRUE)
> summary(aov(obs ~ factor(group),data=data))
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
factor(group)	2	19.10	9.548	6.388	0.0129 *
Residuals	12	17.94	1.495		

```
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

SAS code shown below; corresponding output provided in Figure 1.

```
filename exer924 'C:\courses\MATH 503\data\Exercise9.2.4data.txt';
data exercise924;
infile exer924 delimiter='09'x firstobs=2;
input obs group;
run;

proc anova data=exercise924;
class group;
model obs = group;
;

run;
```

The SAS System

The ANOVA Procedure

Class Level Information

Class	Levels	Values
group	3	1 2 3

Number of Observations Read 15

Number of Observations Used 15

The SAS System

The ANOVA Procedure

Dependent Variable: obs

Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	2	19.09633333	9.54816667	6.39	0.0129
Error	12	17.93700000	1.49475000		
Corrected Total	14	37.03333333			

R-Square	Coeff Var	Root MSE	obs Mean
0.515653	54.74327	1.222600	2.233333

Source	DF	Anova SS	Mean Square	F Value	Pr > F
group	2	19.09633333	9.54816667	6.39	0.0129

Figure 1: Problem 2 SAS output

3. Solve the following by hand, and by using either R or SAS: Let μ_1, μ_2, μ_3 be, respectively, the means of three normal distributions with a common but unknown variance σ^2 . In order to test, at the $\alpha = 0.05$ significance level, the hypothesis $H_0 : \mu_1 = \mu_2 = \mu_3$ against all possible alternative hypotheses, we take an independent random sample of size 4 from each of these distributions. Determine whether we reject or fail to reject H_0 if the observed values from these three distributions are, respectively:

X1	5	9	6	8
X2	11	13	10	12
X3	10	6	9	9

Solved in any of the three manners, we reject H_0 at the 5% significance level ($F = 7.875$; p-val = 0.0105).

R code and output:

```
> data <- read.table("C:/courses/MATH 503/data/Exercise9.2.6data.txt",header=TRUE)
> summary(aov(obs ~ factor(group),data=data))
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
factor(group)	2	42	21.000	7.875	0.0105 *
Residuals	9	24	2.667		

```
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

SAS code shown below; corresponding output provided in Figure 2.

```
filename exer926 'C:\courses\MATH 503\data\Exercise9.2.6data.txt';
data exercise926;
infile exer926 delimiter='09'x firstobs=2;
input obs group;
run;

proc anova data=exercise926;
class group;
model obs = group;
;

run;
```

The SAS System						
The ANOVA Procedure						
Class Level Information						
Class	Levels	Values				
group	3	1	2	3		
Number of Observations Read 12						
Number of Observations Used 12						
The SAS System						
The ANOVA Procedure						
Dependent Variable: obs						
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F	
Model	2	42.00000000	21.00000000	7.88	0.0105	
Error	9	24.00000000	2.66666667			
Corrected Total	11	66.00000000				
R-Square Coeff Var Root MSE obs Mean						
0.636364 18.14437 1.632993 9.000000						
Source	DF	Anova SS	Mean Square	F Value	Pr > F	
group	2	42.00000000	21.00000000	7.88	0.0105	

Figure 2: Problem 3 SAS output

4. Let $X_{ij} = \mu + \alpha_i + \beta_j + \epsilon_{ij}$ denote independently normally distributed random variables, where $\epsilon_{ij} \sim N(0, \sigma^2)$; $i = 1, 2, \dots, a$ and $j = 1, 2, \dots, b$. Show that the maximum likelihood estimator of α_i , β_j , and μ are $\hat{\alpha}_i = \bar{X}_{i.} - \bar{X}_{..}$, $\hat{\beta}_j = \bar{X}_{.j} - \bar{X}_{..}$, and $\hat{\mu} = \bar{X}_{..}$, respectively.

We have that $X_{ij} = \mu + \alpha_i + \beta_j + \epsilon_{ij}$ where $\epsilon_{ij} \sim N(0, \sigma^2)$, i.e. $X_{ij} \sim N(\mu + \alpha_i + \beta_j, \sigma^2)$, therefore $f(x_{ij}; \mu, \alpha_i, \beta_j) = \frac{1}{\sqrt{2\pi}\sigma} e^{\frac{-1}{2\sigma^2}(x_{ij} - \mu - \alpha_i - \beta_j)^2}$.

$$\begin{aligned} L(\mu, \alpha, \beta; \mathbf{x}) &= (2\pi\sigma^2)^{-ab/2} \exp \left[\frac{-1}{2\sigma^2} \sum_{j=1}^b \sum_{i=1}^a (x_{ij} - \mu - \alpha_i - \beta_j)^2 \right] \\ \ln L(\mu, \alpha, \beta; \mathbf{x}) &= -\frac{ab}{2} \ln(2\pi) - \frac{ab}{2} \ln(\sigma^2) - \frac{1}{2\sigma^2} \sum_{j=1}^b \sum_{i=1}^a (x_{ij} - \mu - \alpha_i - \beta_j)^2. \end{aligned}$$

Minimizing $\ln L(\mu, \alpha, \beta; \mathbf{x})$ with respect to μ, α, β respectively reduces to minimizing

$$\begin{aligned} &\sum_{j=1}^b \sum_{i=1}^a (x_{ij} - \mu - \alpha_i - \beta_j)^2 \\ &= \sum_{j=1}^b \sum_{i=1}^a [(x_{ij} - \bar{x}_{i.} - \bar{x}_{.j} + \bar{x}_{..}) + (\bar{x}_{..} - \mu) + (\bar{x}_{i.} - \bar{x}_{..} - \alpha_i) + (\bar{x}_{.j} - \bar{x}_{..} - \beta_j)]^2 \\ &= \sum_{j=1}^b \sum_{i=1}^a (x_{ij} - \bar{x}_{i.} - \bar{x}_{.j} + \bar{x}_{..})^2 + \sum_{j=1}^b \sum_{i=1}^a (\bar{x}_{..} - \mu)^2 \\ &\quad + \sum_{j=1}^b \sum_{i=1}^a (\bar{x}_{i.} - \bar{x}_{..} - \alpha_i)^2 + \sum_{j=1}^b \sum_{i=1}^a (\bar{x}_{.j} - \bar{x}_{..} - \beta_j)^2, \end{aligned} \tag{1}$$

which, because $\sum_{i=1}^a \alpha_i = 0$ and $\sum_{j=1}^b \beta_j = 0$, Equation (1) is minimized iff $\hat{\mu} = \bar{x}_{..}$, $\hat{\alpha}_i = \bar{x}_{i.} - \bar{x}_{..}$, and $\hat{\beta}_j = \bar{x}_{.j} - \bar{x}_{..}$.

Further, note that

$$\begin{aligned} E(\hat{\mu}) &= E(\bar{X}_{..}) = \frac{1}{ab} \sum_{j=1}^b \sum_{i=1}^a E(X_{ij}) = \frac{1}{ab} \sum_{j=1}^b \sum_{i=1}^a (\mu + \alpha_i + \beta_j) = \frac{ab\mu}{ab} = \mu \\ E(\hat{\alpha}_i) &= E(\bar{X}_{i.} - \bar{X}_{..}) = E(\bar{X}_{i.}) - \mu = \frac{\sum_{j=1}^b E(X_{ij})}{b} - \mu = \frac{\sum_{j=1}^b (\mu + \alpha_i + \beta_j)}{b} - \mu \\ &= \frac{b\mu}{b} + \frac{b\alpha_i}{b} - \mu = \alpha_i \\ E(\hat{\beta}_j) &= E(\bar{X}_{.j} - \bar{X}_{..}) = \frac{\sum_{i=1}^a E(X_{ij})}{a} - \mu = \frac{\sum_{i=1}^a (\mu + \alpha_i + \beta_j)}{a} - \mu = \frac{a\mu}{a} + \frac{a\beta_j}{a} - \mu = \beta_j \end{aligned}$$

5. Solve the following using either R or SAS: Given the following observations in a two-way classification with $a = 3$, $b = 4$, and $c = 2$, compute the F-statistics used to test that all interactions are equal to zero ($\gamma_{ij} = 0$), all column means are equal ($\beta_j = 0$), and all row means are equal ($\alpha_i = 0$), respectively.

		Column			
		1	2	3	4
Row	1	3.1	4.2	2.7	4.9
		2.9	4.9	3.2	4.5
	2	2.7	2.9	1.8	3.0
		2.9	2.3	2.4	3.7
	3	4.0	4.6	3.0	3.9
		4.4	5.0	2.5	4.2

Whether solved via R or SAS, we produce the same ANOVA table; see below for code and associated output.

R code and output:

```
> data <- read.table("C:/courses/MATH 503/data/Exercise9.5.7data.txt",header=TRUE)
> summary(aov(obs~factor(row)+factor(col) + factor(row)*factor(col), data=data))
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
factor(row)	2	7.298	3.649	30.726	1.90e-05 ***
factor(col)	3	8.131	2.710	22.825	3.01e-05 ***
factor(row):factor(col)	6	3.412	0.569	4.789	0.0102 *
Residuals	12	1.425	0.119		

```
---
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

SAS code shown below; corresponding output provided in Figure 3.

```
filename exer957 'C:\courses\MATH 503\data\Exercise9.5.7data.txt';
data exercise957;
infile exer957 delimiter='09'x firstobs=2;
input obs row col;
run;

proc anova data=exercise957;
class row col;
model obs = row col row*col;
;

run;
```


The SAS System					
The ANOVA Procedure					
Class Level Information					
Class	Levels	Values			
row	3	1	2	3	
col	4	1	2	3	4
Number of Observations Read 24					
Number of Observations Used 24					

The SAS System					
The ANOVA Procedure					
Dependent Variable: obs					
Source	DF	Sum of Squares	Mean Square	F Value	Pr > F
Model	11	18.84125000	1.71284091	14.42	<.0001
Error	12	1.42500000	0.11875000		
Corrected Total	23	20.26625000			
R-Square Coeff Var Root MSE obs Mean					
0.929686 9.881039 0.344601 3.487500					
Source	DF	Anova SS	Mean Square	F Value	Pr > F
row	2	7.29750000	3.64875000	30.73	<.0001
col	3	8.13125000	2.71041667	22.82	<.0001
row*col	6	3.41250000	0.56875000	4.79	0.0102

Figure 3: Problem 5 SAS output