1. Etercises 12. A: |

1. Use the data for the N=2,954 white females and black females in example 12.2 to compute a 95%. Confidence interval for the population proportion with at least a backelois degree P(B/WF SCHL>21).

Complete P(B/WF SCHL>21) Str = NI 1 + N2 12.

= 0.469 (verify)

then $\hat{V}(\hat{P}(B|WFS(HL \ge 21)Str) = (\frac{N_1}{N})^2 (\frac{N_1 - n_1}{N_1}) \frac{\hat{P}_1(1-\hat{P}_1)}{n_1 - 1} + (\frac{N_2}{N})^2 (\frac{N_2 - n_2}{N_2}) \frac{\hat{P}_2(1-\hat{P}_2)}{n_2 - 1} \approx 0.00185 \text{ (ver:fy)},$

and finally (verify)

(P(B/WF S(HLZZI) Str = 1.96 ()(P(B/WF S(HLZZI) Str))

= (0.38, 0.55)

As is reported in Example 12.2, we see $\hat{P}_1 = \frac{37}{50} = 0.74$ and $\hat{P}_2 = \frac{12}{50} = 0.24$. $N_1 = 1350$, $N_2 = 1604$.

So $\hat{P}_{13}/WF_{51}/WF_{5$

Next, we see $\left(\frac{1350}{2954}\right)^2 \left(\frac{1350-50}{1350}\right) \left(\frac{6.74(1-0.74)}{50-1}\right)$ + $\left(\frac{1604}{2954}\right)^2 \left(\frac{1604-50}{1604}\right) \left(\frac{0.24(1-0.24)}{50-1}\right)$

= 0.000787. + 0.001063...

= 0.00185

Finally, we see the C. I is

(0.469 - (1.96)(0.00185), 0.469 + (1.96)(0.00185))= (0.38, 0.55).

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1. Show that (12.1), (12.7), (12.4), and (12.5) are true.

(12.1)
$$\overline{\gamma}_{M} = \frac{N_{1}}{N} \overline{\gamma}_{M} + \cdots + \frac{N_{H}}{N} \overline{\gamma}_{HM}$$

(12.2) $\overline{\gamma}_{SH} = \frac{N_{1}}{N} \overline{\gamma}_{1} + \cdots + \frac{N_{H}}{N} \overline{\gamma}_{H}$

(12.3) $E(\overline{\gamma}_{SH}) = \overline{\gamma}_{M}$

(12.4) $V(\overline{\gamma}_{SH}) = \sum_{h=1}^{N} \left(\frac{N_{h}}{N}\right)^{2} V(\overline{\gamma}_{h}) = \sum_{h=1}^{H} \left(\frac{N_{h}}{N}\right)^{2} \left(\frac{N_{h} - n_{h}}{N_{h}}\right) \frac{S_{h}^{2}}{n_{h}}$

First we consider (12.1).

We know $\overline{\gamma}_{M} = \sum_{i=1}^{N} \overline{\gamma}_{i}$ and $N_{i} + \cdots + N_{H} = N$.

For each subpopulation, $\overline{\gamma}_{hM} = \sum_{i=1}^{N} \overline{\gamma}_{i} \Leftrightarrow N_{h} \overline{\gamma}_{h} = \sum_{i=1}^{N} \gamma_{i}$.

We see $\sum_{i=1}^{N} \gamma_{i} = \sum_{j=1}^{N} \gamma_{j} + \cdots + \sum_{k=1}^{N} \gamma_{k} = N_{i} \overline{\gamma}_{k} + \cdots + N_{H} \overline{\gamma}_{H}$.

Next we consider ystr in (12.2).

The proof is almost exactly the same as above, but instead of pieces of the population we take weighted average of means of pieces of the sample.

Next we consider (12.3). Prove E(yshr) = yu. E(Ystr) = E(NI Y, + ... + NH YH) from provious section. = E(N/ 7,) + ... + (NH7H) = N/ E(7)+...+ NME(YM) since the sample mean is an unbiased estimator of the population mean, we know E (Th) = Thu. so we have Nyn+ ... + NN ynu = Yu, from previous section of homework. Next we consider (12.4).

Prove $V(\overline{y}_{SHr}) = \sum_{h=1}^{H} \left(\frac{N_h}{N}\right)^2 V(\overline{y}_h) = \sum_{h=1}^{H} \left(\frac{N_h}{N}\right)^2 \frac{N_h - n_h s_h^2}{N_h n_h}$ $V(\bar{\gamma}_{S+r}) = V(\frac{N_1}{N}\bar{\gamma}_1 + \dots + \frac{N_H}{N}\bar{\gamma}_H) = V(\frac{N_1}{N}\bar{\gamma}_1) + \dots + V(\frac{N_H}{N}\bar{\gamma}_h)$ due to independence of the Strater. $= \left(\frac{N_{i}}{N}\right)^{2} \vee \left(\tilde{\gamma_{i}}\right)^{2} + \cdots + \left(\frac{N_{H}}{N}\right)^{2} \vee \left(\tilde{\gamma}_{H}\right)^{2}.$ For (12.5). We want $\hat{V}(\hat{y}_{str})$, an estimator for the above. we simply to replate V with V in all places above. This means we replace to population 52 with astimator sampling 52.