

$$\Rightarrow P[\underline{X} \in \underbrace{(\underline{x}, \underline{x} + d\underline{x})}_{(x_1, x_1 + dx_1) \times (x_2, x_2 + dx_2) \times \dots \times (x_k, x_k + dx_k)}] = f_{\underline{X}}(\underline{x}) d\underline{x},$$

i.e.,

$$P[x_1 < X_1 < x_1 + dx_1, x_2 < X_2 < x_2 + dx_2, \dots, x_k < X_k < x_k + dx_k] \\ = f_{X_1, X_2, \dots, X_k}(x_1, x_2, \dots, x_k) dx_1 dx_2 \dots dx_k$$

Any joint pdf  $f_{\underline{X}}$  satisfies

$$\textcircled{\text{I}} \quad f_{\underline{X}}(\underline{x}) \geq 0 \quad \forall \quad \underline{x} \in \mathbb{R}^k, \text{ and}$$

$$\textcircled{\text{II}} \quad \int \dots \int_{\mathbb{R}^k} f_{\underline{X}}(\underline{x}) d\underline{x} = \int_{-\infty}^{\infty} \dots \int_{-\infty}^{\infty} f_{X_1, \dots, X_k}(x_1, \dots, x_k) dx_1 \dots dx_k = 1.$$

Again from the joint cdf  $F_{\underline{X}}(\underline{x})$ , a joint pdf can be guessed using the following recipe:

$$f_{\underline{X}}(\underline{x}) = \begin{cases} \frac{\partial}{\partial x_1} \frac{\partial}{\partial x_2} \dots \frac{\partial}{\partial x_k} F_{\underline{X}}(\underline{x}) & \text{whenever the} \\ & \text{partial derivatives exist,} \\ 0 & \text{otherwise.} \end{cases}$$

If  $\underline{X}$  is indeed a cont random vector, then the above recipe will always work and the order

of the partial derivatives ~~does~~<sup>will</sup> not matter.

## Independence of $k$ random variables

Suppose  $\underline{X} = (X_1, X_2, \dots, X_k)$  is any (not necessarily discrete or continuous) random vector.

Def<sup>n</sup>: We say that the r.v.s  $X_1, X_2, \dots, X_k$  are independent if  $\forall \underline{u} = (u_1, u_2, \dots, u_k) \in \mathbb{R}^k$ ,

$$P(X_1 \leq u_1, X_2 \leq u_2, \dots, X_k \leq u_k) = \prod_{i=1}^k P(X_i \leq u_i),$$

$$\text{i.e., } F_{\underline{X}}(\underline{u}) = F_{X_1}(u_1) F_{X_2}(u_2) \dots F_{X_k}(u_k).$$

Roughly speaking, this means that the r.v.s  $X_1, X_2, \dots, X_k$  do not influence each other.

In parallel to the bivariate case, the following theorems can be proved.

Thm<sup>d</sup>: Suppose  $\underline{X}$  is a discrete random vector. Then  $X_1, X_2, \dots, X_k$  are independent if and only if  $p_{\underline{X}}(\underline{z}) = p_{X_1}(z_1) p_{X_2}(z_2) \dots p_{X_k}(z_k) \quad \forall \underline{z} \in \mathbb{R}^k$ .

Thm C: Suppose  $X_1, X_2, \dots, X_k$  are cont r.v.s with pdfs  $f_{X_1}, f_{X_2}, \dots, f_{X_k}$ , respectively.

Then  $X_1, X_2, \dots, X_k$  are independent if and only if  $\underline{X}$  is a cont random vector with a joint pdf

$$h(\underline{x}) = f_{X_1}(x_1) f_{X_2}(x_2) \dots f_{X_k}(x_k), \quad \underline{x} \in \mathbb{R}^k.$$

Exc: Prove Thm C and Thm D.

Remarks: Suppose  $\{X_1, X_2, X_3, X_4\}$  are ind r.v.s. Then the following facts can be shown to hold:

(i) <sup>The r.v.s in any subcollection</sup> ~~Any~~ <sup>subset</sup> of  $\{X_1, X_2, X_3, X_4\}$  are ind.

In particular  $X_i \perp\!\!\!\perp X_j \quad \forall \quad 1 \leq i < j \leq 4.$

(ii)  $X_1^2, e^{X_2}, \log(1+|X_3|), \sin X_4$  are ind.

(iii)  $X_1 + X_2 + X_3 \perp\!\!\!\perp X_4.$

(iv)  $X_1^2 + X_2^2 \perp\!\!\!\perp X_3^7 + e^{X_4}.$

Exc: Prove (i). Also verify (ii) and (iii) in the discrete case.



(104)

Exc: Verify (iii) ~~in the~~ when  $X_1, X_2, X_3$  and  $X_4$  are jointly continuous. [Hint: Use (★)]

The phrase independent and identically distributed (or i.i.d. or iid) has the same meaning. It means that a bunch of r.v.s are independent and they all have the same distribution with same parameter values.

~~Thm~~: Back to additivity of gamma dist<sup>n</sup>.

Thm: Suppose  $X_1, X_2, \dots, X_k$  are ind r.v.s such that  $X_i \sim \text{Gamma}(\alpha_i, \lambda)$  for each  $i \in \{1, 2, \dots, k\}$ . Then

$$X_1 + X_2 + \dots + X_k \sim \text{Gamma}(\alpha_1 + \alpha_2 + \dots + \alpha_k, \lambda).$$

Proof: Exc [Hint: Use the proposition stated at the end of Pg (93) + induction on  $k$  + Remark (iii) of Pg (103).]

Cor1: If  $X_1, X_2, \dots, X_k \stackrel{iid}{\sim} \text{Exp}(\lambda)$ , then we have

$$X_1 + X_2 + \dots + X_k \sim \text{Gamma}(k, \lambda).$$

Proof: Apply the theorem stated in Pg (104) ~~in~~ in the special case  $\alpha_1 = \alpha_2 = \dots = \alpha_k = 1$ .

Cor2: If  $Z_1, Z_2, \dots, Z_k \stackrel{iid}{\sim} N(0,1)$ , then

$$Z_1^2 + Z_2^2 + \dots + Z_k^2 \sim \text{Gamma}\left(\frac{k}{2}, \frac{1}{2}\right).$$

Proof: Exc [Hint: See Pg (94) for the  $k=2$  case.]

Def<sup>n</sup>:  $\text{Gamma}\left(\frac{k}{2}, \frac{1}{2}\right)$  distribution is also called chi-squared distribution (also chi-square ~~or~~ dist<sup>n</sup> or  $\chi^2$ -dist<sup>n</sup>) with  $k$  degrees of freedom.

Remarks: ① The phrase "k degrees of freedom" is used because there are  $k$  independent (and hence "free") standard normal r.v.s in the background (i.e.,  $Z_1, Z_2, \dots, Z_k \stackrel{iid}{\sim} N(0,1)$  as in Cor 2 above) whose sum of squares follow this dist<sup>n</sup>.

② If  $X \sim \text{Gamma}(\frac{k}{2}, \frac{1}{2})$ , then we also use the notation  $X \sim \chi^2_k$ .

③ In light of Remark ② above, Cor-2 can be restated as

$$Z_1, Z_2, \dots, Z_k \stackrel{\text{iid}}{\sim} N(0,1) \Rightarrow \sum_{i=1}^k Z_i^2 \sim \chi^2_k.$$

④ Chi-squared dist<sup>n</sup> plays a very important role in statistics and related disciplines.

Till now, we were discussing the distr<sup>n</sup> of sum of two (or more) ~~joint~~ jointly cont r.v.s. We shall now deal with another scalar valued function of a <sup>cont</sup> random vector - namely, the ratio.

The distribution of ratio of two jointly cont r.v.s

Suppose  $(X, Y)$  is a cont random vector with a joint pdf  $f_{X,Y}$ . In particular,  $Y$  is a cont r.v. and hence it satisfies  $P[Y=0]=0$ .



(107)

Therefore  $Z := \frac{X}{Y}$  is a well-defined r.v.

Goal: To find the dist<sup>n</sup> of  $Z$ .

Thm: If  $(X, Y)$  is a cont random vector with a joint pdf  $f_{X,Y}(x, y)$ , then  $Z := \frac{X}{Y}$  is ~~also~~ a cont r.v. with a pdf

$$f_Z(z) = \int_{-\infty}^{\infty} |y| f_{X,Y}(yz, y) dy, \quad z \in \mathbb{R}.$$

~~If  $X \perp Y$~~

Cor: If  $X, Y$  are ind cont r.v.s with pdfs  $f_X, f_Y$  respectively, then  $Z := \frac{X}{Y}$  is also a cont r.v. with a pdf

$$f_Z(z) = \int_{-\infty}^{\infty} |y| f_X(yz) f_Y(y) dy, \quad z \in \mathbb{R}.$$

Cor: In the setup of the above thm, if  $\text{Range}(Y) \subseteq (0, \infty)$ , and ~~then~~  $\text{Range}(X) \subseteq (0, \infty)$ ,  $\mathbb{R}^+$   
then  $f_Z(z) = \int_0^{\infty} y f_{X,Y}(yz, y) dy, \quad z \in \mathbb{R}^+.$

Proof of Thm: We need to show that  $\forall a \in \mathbb{R}$ ,

$$(2) \dots P(Z \leq a) = \int_{-\infty}^a h(z) dz,$$

where  $h(z) = \int_{-\infty}^{\infty} |y| f_{X,Y}(yz, y) dy, z \in \mathbb{R}.$

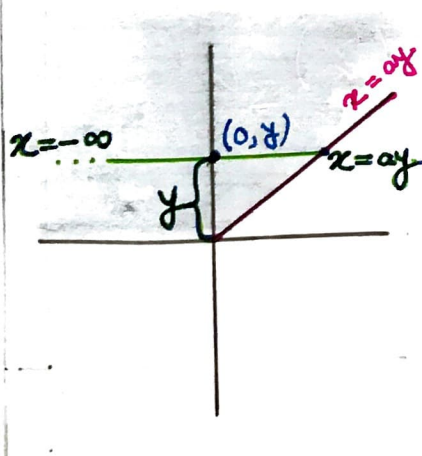
~~Fix  $z \in \mathbb{R}$ .~~ Take  $a \in \mathbb{R}.$

$$\text{LHS of (2)} = P(Z \leq a) = P\left(\frac{X}{Y} \leq a\right)$$

$$= \underbrace{P(X \leq aY, Y > 0)}_{\text{I}} + \underbrace{P(X \geq aY, Y < 0)}_{\text{II}}$$

$$[\because P(Y=0)=0]$$

The first term  $\text{I} = P(X \leq aY, Y > 0)$



$$\stackrel{(*)}{=} \iint_{\substack{y > 0 \\ x \leq ay}} f_{X,Y}(x, y) dx dy$$

$$= \int_0^{\infty} \int_{-\infty}^{ay} f_{X,Y}(x, y) dx dy$$

Note that in the inner integral,  $y$  is a constant.



(109)

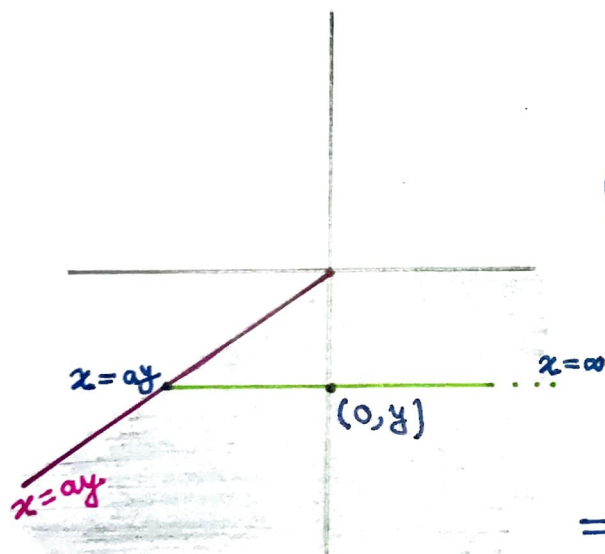
Put  $x = yz$  in the inner integral.

$$\Rightarrow dx = y dz \quad \text{and} \quad z = \frac{x}{y}$$

$$\Rightarrow \textcircled{\text{I}} = \int_0^{\infty} \int_{-\infty}^a y f_{X,Y}(yz, y) dz dy.$$

Similarly, the second term =  $\textcircled{\text{II}}$

$$= P[X \geq aY, Y < 0]$$



$$\stackrel{(*)}{=} \iint_{\substack{y < 0 \\ x \geq ay}} f_{X,Y}(x, y) dx dy$$

$$= \int_{-\infty}^0 \int_{ay}^{\infty} f_{X,Y}(x, y) dx dy$$

Again put  $x = yz$  in the inner integral.

$$\Rightarrow dx = y dz \quad \text{and} \quad z = \frac{x}{y} \quad (\text{but } y < 0).$$

$$\begin{aligned} \Rightarrow \textcircled{\text{II}} &= \int_{-\infty}^0 \int_a^{-\infty} y f_{X,Y}(yz, y) dz dy \\ &= \int_{-\infty}^0 \int_{-\infty}^a (-y) f_{X,Y}(yz, y) dz dy \end{aligned}$$

Therefore, LHS of (2)

$$= \textcircled{\text{I}} + \textcircled{\text{II}}$$

$$= \int_0^{\infty} \int_{-\infty}^a y f_{x,y}(y\bar{z}, y) d\bar{z} dy + \int_{-\infty}^0 \int_{-\infty}^a (-y) f_{x,y}(y\bar{z}, y) d\bar{z} dy$$

$$= \int_0^{\infty} \int_{-\infty}^a |y| f_{x,y}(y\bar{z}, y) d\bar{z} dy + \int_{-\infty}^0 \int_{-\infty}^a |y| f_{x,y}(y\bar{z}, y) d\bar{z} dy$$

$$= \int_{-\infty}^{\infty} \int_{-\infty}^a |y| f_{x,y}(y\bar{z}, y) d\bar{z} dy$$

$$\stackrel{\text{Fubini}}{=} \int_{-\infty}^a \int_{-\infty}^{\infty} |y| f_{x,y}(y\bar{z}, y) dy d\bar{z}$$

$$= \int_{-\infty}^a h(\bar{z}) d\bar{z}$$

$$= \text{RHS of (2)}.$$

This completes the proof.