

Example Class 6

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- The Jeffreys Prior is proportional to the square root of the determinant of the Fisher information matrix

$$f(\theta) \propto |I(\theta)|^{\frac{1}{2}}$$

$$I(\theta) = -E \left\{ \frac{\partial^2 \log L(\theta|y)}{\partial \theta^2} \right\}$$

- It is invariant to re-parametrization.
- That is, the relative probability assigned to a volume of a probability space using a Jeffreys prior will be the same regardless of the parameterization used to define the Jeffreys prior.

Jeffreys Prior (Parameterization)

- For a monotone variable transformation $\eta = h(\theta), \theta = h^{-1}(\eta)$
- Note that

$$I(\eta) = -\mathbb{E} \left(\frac{\partial^2 \log f(x|\eta)}{\partial \eta^2} \right) = -\mathbb{E} \left[\frac{\frac{\partial^2}{\partial \eta^2} f(x|\eta)}{f(x|\eta)} - \left(\frac{\frac{\partial}{\partial \eta} f(x|\eta)}{f(x|\eta)} \right)^2 \right] = \mathbb{E} \left(\frac{\partial \log f(x|\eta)}{\partial \eta} \right)^2$$

where

$$\mathbb{E} \left(\frac{\frac{\partial^2}{\partial \eta^2} f(x|\eta)}{f(x|\eta)} \right) = \int \frac{\frac{\partial^2}{\partial \eta^2} f(x|\eta)}{f(x|\eta)} f(x|\eta) dx = \frac{\partial^2}{\partial \eta^2} \int f(x|\eta) dx = 0.$$

- So we have

$$\begin{aligned} |I(\eta)| &= \left| \mathbb{E} \left(\frac{\partial^2 \log f(x|\eta)}{\partial \eta^2} \right) \right| \\ &= \mathbb{E} \left(\frac{\partial \log f(x|h(\theta))}{\partial \eta} \right)^2 = \mathbb{E} \left(\frac{\partial \log f(x|h(\theta))}{\partial \theta} \frac{\partial \theta}{\partial \eta} \right)^2 \\ &= |I(\theta)| \left| \frac{\partial \theta}{\partial \eta} \right|^2. \end{aligned}$$

- We thus have $|I(\eta)|^{1/2} = |I(\theta)|^{1/2} \left| \frac{\partial \theta}{\partial \eta} \right|$.

How to Borrow Information from Historical Data

- Power Prior (γ is the power parameter):

$$\pi(\theta|D_0, \gamma) \propto \underbrace{L(\theta|D_0)^\gamma}_{\text{discounted historical information}} \underbrace{\pi(\theta)}_{\text{prior information unrelated to historical data}}$$

- Posterior:

$$f(\theta|D_0, D, \gamma) \propto L(\theta|D)\pi(\theta|D_0, \gamma) \propto L(\theta|D)L(\theta|D_0)^\gamma\pi(\theta)$$

- γ weights the historical data D_0 relative to the current data D
- With a flexible γ ,

$$\pi(\theta, \gamma|D_0) \propto \frac{L(\theta|D_0)^\gamma\pi(\theta)}{\int L(\theta|D_0)^\gamma\pi(\theta)d\theta}\pi(\gamma)$$

Example (Normal Case)

- Historical data $\mathbf{x}_0 = (x_{01}, \dots, x_{0n_0})'$
- Current data $\mathbf{x} = (x_1, \dots, x_n)'$, $x_i \stackrel{i.i.d}{\sim} N(\mu, \sigma^2)$
- For simplicity, we assume that the variance σ^2 is known.
- Power prior with fixed γ

$$\pi^{PP}(\mu|D_0, \gamma) \propto L(\mu|D_0)^\gamma \pi_0(\mu)$$

- With a flat prior $\pi_0(\mu) \propto 1$,

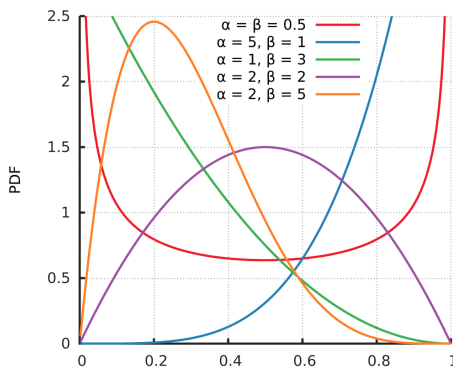
$$\pi^{PP}(\mu|D_0, \gamma) \propto e^{-\frac{\gamma \sum_{i=1}^{n_0} (x_{0i} - \mu)^2}{2\sigma^2}} \sim N\left(\bar{x}_0, \frac{\sigma^2}{n_0\gamma}\right)$$

- The posterior distribution has the form,

$$p^{PP}(\mu|D, D_0, \gamma) \propto e^{-\frac{\gamma \sum_{i=1}^{n_0} (x_{0i} - \mu)^2}{2\sigma^2}} e^{-\frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2}} \sim N\left(\frac{\gamma n_0 \bar{x}_0 + n \bar{x}}{\gamma n_0 + n}, \frac{\sigma^2}{n_0\gamma + n}\right)$$

How to Borrow Information from Historical Data

- If we specify $\pi(\gamma)$ as a Beta(a, b) distribution for fixed positive hyperparameters a and b
- ($a = 5, b = 1$) would strongly encourage borrowing
- ($a = 1, b = 5$) would strongly discourage it
- ($a = b = 1$) would be agnostic on the subject, essentially letting the data determine the degree of borrowing.



Example (Normal Case)

- Modified power prior (flexible γ) Page 4

$$\pi^{MPP}(\mu, \gamma | D_0) \propto \frac{L(\mu | D_0)^\gamma \pi(\mu)}{\int L(\mu | D_0)^\gamma \pi(\mu) d\mu} \pi(\gamma)$$

- With a Beta(a,b) prior on γ ,

$$\begin{aligned} \pi^{MPP}(\mu, \gamma | D_0) &\propto \frac{e^{-\frac{(\mu - \bar{x}_0)^2}{2\sigma^2/(n_0\gamma)}}}{\int e^{-\frac{(\mu - \bar{x}_0)^2}{2\sigma^2/(n_0\gamma)}} d\mu} \pi(\gamma) \propto \sqrt{\gamma} e^{-\frac{n_0\gamma(\mu - \bar{x}_0)^2}{2\sigma^2}} \gamma^{a-1} (1-\gamma)^{b-1} \\ &\propto N\left(\mu | \bar{x}_0, \frac{\sigma^2}{n_0\gamma}\right) \text{Beta}(\gamma | a, b) \end{aligned}$$

- Posterior

$$p^{MPP}(\mu, \gamma | D_0, D) \propto \pi^{MPP}(\mu, \gamma | D_0) L(\mu | D)$$

Example (Normal Case)

Posterior

$$p^{MPP}(\mu, \gamma | D_0, D) \propto \pi^{MPP}(\mu, \gamma | D_0) L(\mu | D)$$

$$\begin{aligned} p^{MPP}(\mu, \gamma | D_0, D) &\propto \gamma^{1/2} e^{-\frac{n_0 \gamma (u - \bar{x}_0)}{2\sigma^2}} e^{-\frac{\sum_{i=1}^n (x_i - \mu)^2}{2\sigma^2}} \gamma^{a-1} (1 - \gamma)^{b-1} \\ &\propto \gamma^{1/2} \gamma^{a-1} (1 - \gamma)^{b-1} e^{-\left[\frac{\left(u - \frac{n_0 \gamma \bar{x}_0 + n \bar{x}}{n_0 \gamma + n} \right)^2}{2\sigma^2 / (n_0 \gamma + n)} + \frac{n_0 \gamma \bar{x}_0^2 - \frac{(n_0 \gamma \bar{x}_0 + n \bar{x})^2}{n_0 \gamma + n}}{2\sigma^2} \right]} \\ &\propto \gamma^{1/2} \gamma^{a-1} (1 - \gamma)^{b-1} e^{-\frac{\left(u - \frac{n_0 \gamma \bar{x}_0 + n \bar{x}}{n_0 \gamma + n} \right)^2}{2\sigma^2 / (n_0 \gamma + n)}} e^{-\frac{nn_0 \gamma (\bar{x} - \bar{x}_0)^2}{2\sigma^2 (n + n_0 \gamma)}} \\ p^{MPP}(\mu | D_0, D, \gamma) &\sim N \left(\frac{\gamma n_0 \bar{x}_0 + n \bar{x}}{\gamma n_0 + n}, \frac{\sigma^2}{n_0 \gamma + n} \right) \\ p^{MPP}(\gamma | D_0, D) &\sim N \left(\bar{x} | \bar{x}_0, \frac{\sigma^2 (n + n_0 \gamma)}{nn_0 \gamma} \right) \text{Beta}(\gamma | a, b) \end{aligned}$$

Commensurate Power Prior

- The hierarchical power prior do not directly parametrize the commensurability of the historical and new data since the posterior of γ is free of the current data D .
- Commensurate Power Prior: assume different parameters in the historical and current data.

$$\pi(\theta, \gamma, \tau | D_0) \propto \int \frac{L(\theta_0 | D_0)^\gamma \pi(\theta_0)}{\int L(\theta_1 | D_0)^\gamma \pi(\theta_1) d\theta_1} \pi(\theta | \theta_0, \tau) d\theta_0 \cdot \pi(\gamma | \tau) \pi(\tau)$$

- The prior of θ has the mean θ_0 and precision τ and the power parameter γ depends on τ . Thus, τ parameterizes commensurability between the historical and current data.
- For example, $\theta \sim N(\theta_0, 1/\tau)$, $\gamma \sim \text{Beta}(g(\tau), 1)$

Example (Normal Case)

- Commensurate Power Prior:

$$\pi(\mu_0) \propto 1, \pi(\mu|\mu_0, \tau) \sim N(\mu_0, 1/\tau), \pi(\gamma|\tau) \sim \text{Beta}(g(\tau), 1)$$

$$\begin{aligned}\pi^{CPP}(\mu, \gamma, \tau|D_0) &\propto \int \frac{L(\mu_0|D_0)^\gamma \pi(\mu_0)}{\int L(\mu_0|D_0)^\gamma \pi(\mu_0)} \pi(\mu|\mu_0, \tau) d\mu_0 \cdot \pi(\gamma|\tau) \pi(\tau) \\ &\sim N\left(\mu|\bar{x}_0, \frac{1}{\tau} + \frac{\sigma^2}{\gamma n_0}\right) \text{Beta}(g(\tau), 1) \pi(\tau)\end{aligned}$$

$$\begin{aligned}p^{CPP}(\mu|\gamma, \tau, D_0, D) &\sim N\left(\frac{\gamma n_0 \tau \sigma^2 \bar{x}_0 + n(\gamma n_0 + \sigma^2 \tau) \bar{x}}{\gamma n_0 \tau \sigma^2 + n(\gamma n_0 + \sigma^2 \tau)}, \right. \\ &\quad \left. \frac{(\gamma n_0 + \sigma^2 \tau) \sigma^2}{\gamma n_0 \tau \sigma^2 + n(\gamma n_0 + \sigma^2 \tau)}\right)\end{aligned}$$

$$p^{CPP}(\gamma, \tau|D, D_0) \sim N\left(\bar{x}|\bar{x}_0, \frac{\sigma^2}{n} + \frac{1}{\tau} + \frac{\sigma^2}{\gamma n_0}\right) \text{Beta}(g(\tau), 1) \pi(\tau)$$

Hierarchical Commensurate Prior

- General case of commensurate priors: weighting the influence of prior information relative to its consistency with data from the concurrent study.

$$\pi(\theta|\theta_0, D_0, \tau) \propto L(\theta_0|D_0)p(\theta|\theta_0, \tau)\pi_0(\theta)$$

- Note that the full conditional posterior distribution for θ_0 would be independent of the current data. Thus, θ_0 should be integrated out of the prior when the $\int L(\theta_0|D_0)p(\theta|\theta_0, \tau)d\theta_0$ is tractable.
- When $\tau \rightarrow 0$, $\pi(\theta|\theta_0, D_0, \tau) \rightarrow \pi_0(\theta)$.
- When $\tau \rightarrow \infty$, $\pi(\theta|\theta_0, D_0, \tau) \rightarrow L(\theta|D_0)\pi_0(\theta)$

Example (Normal Case)

- Hierarchical Commensurate Prior: consider $\pi_0(\mu) \propto 1$

$$\begin{aligned}\pi^{HCP}(\mu, \tau | D_0) &= \int L(\mu_0 | D_0) \pi(\mu | \mu_0, \tau) d\mu_0 \pi_0(\mu) \pi(\tau) \\ &\propto N\left(\mu | \bar{x}_0, \frac{1}{\tau} + \frac{\sigma^2}{n_0}\right) \pi(\tau)\end{aligned}$$

Example (Binary Case)

- Historical data: y_0, n_0
- Current data: $y, n, y \sim \text{Bin}(n, \theta)$
- Power prior with a fixed γ and a $\text{Beta}(a, b)$ prior on θ :

$$\begin{aligned}\pi^{PP}(\theta|D_0, \gamma) &\propto L(\theta|D_0)^\gamma \pi_0(\theta) \propto \theta^{\gamma y_0} (1 - \theta)^{\gamma(n_0 - y_0)} \theta^{a-1} (1 - \theta)^{b-1} \\ &\sim \text{Beta}(a + \gamma y_0, b + \gamma(n_0 - y_0))\end{aligned}$$

- Posterior:

$$p^{PP}(\theta|D, D_0, \gamma) \sim \text{Beta}(a + y + \gamma y_0, b + (n - y) + \gamma(n_0 - y_0))$$

Example (Binary Case w/ Modified Power Prior)

- Modified power prior (with a $\text{Beta}(c, d)$ on γ):

$$\begin{aligned}\pi^{MPP}(\theta, \gamma | D_0) &\propto \frac{\Gamma(a + b + \gamma n_0)}{\Gamma(a + \gamma y_0) \Gamma(b + \gamma(n_0 - y_0))} \theta^{\gamma y_0 + a - 1} (1 - \theta)^{\gamma(n_0 - y_0) + b - 1} \\ &\quad \gamma^{c-1} (1 - \gamma)^{d-1} \\ &\propto \text{Beta}(\theta | a + \gamma y_0, b + \gamma(n_0 - y_0)) \text{Beta}(\gamma | c, d)\end{aligned}$$

- Posterior:

$$\begin{aligned}p^{MPP}(\theta | D, D_0, \gamma) &\sim \text{Beta}(a + \gamma y_0 + y, b + \gamma(n_0 - y_0) + n_0) \\ p^{MPP}(\gamma | D, D_0) &\propto \frac{\Gamma(a + b + \gamma n_0)}{\Gamma(a + \gamma y_0) \Gamma(b + \gamma(n_0 - y_0))} \\ &\quad \cdot \frac{\Gamma(a + \gamma y_0 + y) \Gamma(b + \gamma(n_0 - y_0) + (n - y))}{\Gamma(a + b + \gamma n_0 + n)} \\ &\quad \cdot \text{Beta}(\gamma | c, d)\end{aligned}$$