## Comments/Solutions, MTH 9831 Fall 2015 HW #11 Baruch College MFE Program

Please note that these are *possible* solutions, and that there are often many ways to approach and solve a problem or prove a theorem.

We start with  $N(t) \sim \text{Poisson}(\lambda)$ ,  $M(t) = N(t) - \lambda t$  a compensated Poisson process, Q(t) a compound Poisson process with jump distribution  $P(Y_1 = y_m) = p_m$ , m = 1, 2, ..., M, and  $M_Q(t) = Q(t) - \beta \lambda t$  (where  $\beta = E(Y_1)$ ) is the compensated process for Q.

## Problem 1 (quadratic variation of compound Poissons)

Letting  $\Delta Q(s) = Q(s) - Q(s-)$ , we have

$$\begin{split} [Q,Q](t) &= \sum_{0 < s \le t} (\Delta Q(s))^2 = \sum_{j=1}^{N(t)} Y_j^2 \\ E([Q,Q](t)) &= E\left(\sum_{j=1}^{N(t)} Y_j^2\right) = E(N(t))E(Y_1^2) \quad \text{(by Wald's identity)} \\ &= (\sigma^2 + \beta^2)\lambda t \quad \text{(setting } \sigma^2 = Var(Y_1) = E(Y_1^2) - E(Y_1)^2) \\ [M_Q, M_Q](t) &= [M_Q^c, M_Q^c](t) + [J, J](t) \quad (M_Q^c(t) = -\beta\lambda t, J(t) = Q(t)) \\ &= [Q, Q](t). \end{split}$$

### Problem 2 (Itô isometry with jumps, compensated Poisson)

The first part, showing that for  $\phi$  a left-continuous square-integrable M(t)-adapted process,

$$E\left[\left(\int_0^t \phi(s)dM(s)\right)^2\right] = \lambda E\left[\int_0^t \phi^2(s)ds\right],$$

is a special case of Problem 3 with Q(t) = N(t). Hence, we know that  $\int_0^t \phi(s) dM(s)$  is a martingale starting at 0 since M(t) is a martingale, and so

$$\begin{split} Var\left(\int_0^t 2^{M(s-)}dM(s)\right) &= E\left[\left(\int_0^t 2^{M(s-)}dM(s)\right)^2\right] \\ &= \lambda E\left[\int_0^t 4^{M(s-)}ds\right] \\ &= \lambda E\left[\int_0^t e^{\ln(4)M(s-)}ds\right] = \lambda E\left[\int_0^t e^{\ln(4)(N(s-)-\lambda s)}ds\right]. \end{split}$$

At this point, we argue that the Lebesgue integrals  $\int_0^t e^{\ln(4)(N(s-)-\lambda s)} ds = \int_0^t e^{\ln(4)(N(s)-\lambda s)} ds$  since  $e^{\ln(4)N(s-)}$  and  $e^{\ln(4)N(s)}$  only differ at finitely many points a.s., so their ds integrals match. We swap this out and use the MGF of N(s),  $E(e^{uN(s)}) = e^{\lambda s(e^u-1)}$ , to get

$$Var\left(\int_{0}^{t} 2^{M(s-)} dM(s)\right) = \lambda \int_{0}^{t} e^{\ln(4)(-\lambda s)} E(e^{\ln(4)N(s)}) ds$$
$$= \lambda \int_{0}^{t} e^{\ln(4)(-\lambda s)} e^{3\lambda s} ds$$
$$= \lambda \int_{0}^{t} e^{(3-\ln(4))(\lambda s)} ds = \frac{e^{(3-\ln(4))\lambda t} - 1}{3 - \ln(4)}.$$

#### Problem 3 (Itô isometry with jumps, compound compensated Poisson)

We need to show that

$$E\left[\left(\int_0^t \phi(s)dM_Q(s)\right)^2\right] = \lambda E(Y_1^2) E\left[\int_0^t \phi^2(s)ds\right].$$

We'll start with the differential

$$X(t) = \int_0^t \phi(s)dM_Q(s), \ dX(t) = \phi(t)dM_Q(t), \ X(0) = 0.$$

Then differentiate  $X^2(t)$ :

$$\begin{split} d(X^2(t)) &= 2X(t)dX(t) + d[X,X](t) \\ \Longrightarrow X^2(t) &= 2\int_0^t X(s)dX(s) + \int_0^t d[X,X](s) &= 2\int_0^t X(s-)dX(s) + \sum_{0 < s \le t} \Delta X(s)^2 \\ &= 2\int_0^t X(s-)dX(s) + \sum_{j=1}^{N(t)} Y_j^2 \int_{S_{j-1}}^{S_j \wedge t} \phi^2(s)dN(s), \end{split}$$

where  $S_j$  is the time of the jth jump of N,  $S_0 = 0$  (I'm being a little flaky on the notation here, but  $S_j - S_{j-1} \sim \exp(\lambda)$  for each j, IID). Since the  $Y_j$  are IID, we'll be able to combine all these integrals into one integral in the expectation.

 $M_Q(t)$  is a martingale, so X(t) is a martingale, so  $\int_0^t X(s-)dX(s)$  is a martingale since X(s-) is left-continuous. Hence, since the  $Y_j$  are IID and independent of N(t), and  $N(t) = M(t) + \lambda t$  where M(t) is a martingale,

$$\begin{split} E\left[X^2(t)\right] &= E\left[\left(\int_0^t \phi(s)dM_Q(s)\right)^2\right] = E\left[\sum_{j=1}^{N(t)} Y_j^2 \int_{S_{j-1}}^{S_j \wedge t} \phi^2(s)dN(s)\right] \\ &= E\left[Y_1^2\right] E\left[\int_0^t \phi^2(s)dN(s)\right] = \lambda E\left[Y_1^2\right] E\left[\int_0^t \phi^2(s)ds\right]. \end{split}$$

# Problem 4 Shreve Vol. 2, Ch. 11 #11.6, p. 526 (BM and compound Poisson are independent)

We'll do the same thing as in #11.5 from the previous homework set: let

$$X(t) = u_1 W(t) - \frac{1}{2} u_1^2 t + u_2 Q(t) - \lambda t (\phi(u_2) - 1)$$

$$= X^c(t) + J(t);$$

$$X^c(t) = u_1 W(t) - \frac{1}{2} u_1^2 t - \lambda t (\phi(u_2) - 1);$$

$$J(t) = u_2 Q(t);$$

where  $\phi$  is the MGF of Y(t), so that  $Y(t) = e^{X(t)}$ . We again show that Y(t) is a martingale. First, the continuous terms:

$$dX^{c}(t) = u_{1}dW_{t} + \left[ -\frac{1}{2}u_{1}^{2} - \lambda(\phi(u_{2}) - 1) \right]dt; \qquad d[X^{c}, X^{c}](t) = u_{1}^{2}dt.$$

Then the jump terms:

$$\Delta X(t) = \Delta J(t) = u_2 \Delta Q(t) = u_2 Y_{N(t)} \Delta N(t)$$
  
$$\Delta e^{X(t)} = e^{X(t)} - e^{X(t-)} = e^{X(t-)} (e^{\Delta X(t)} - 1) = e^{X(t-)} (e^{u_2 Y_{N(t)}} - 1) \Delta N(t).$$

We'll consider the compound Poisson process  $H(t) = \sum_{j=1}^{N(t)} (e^{u_2 Y_j} - 1)$ , which looks just like the above, and so we combine via Itô to give us, for any  $s \ge 0$ ,

$$\begin{split} \Delta H(u) &= dH(u) = (e^{u_2 Y_{N(u)}} - 1) \Delta N(u) \\ &\quad H(t) - \lambda(\phi(u_2) - 1)t \text{ is a martingale} \\ Y(t) &= e^{X(t)} = e^{X(s)} + \int_s^t e^{X(u-)} dX^c(u) + \frac{1}{2} \int_s^t e^{X(u-)} d[X^c, X^c](u) \\ &\quad + \sum_{s < u \le t} e^{X(u-)} (e^{u_2 Y_{N(u)}} - 1) \Delta N(t) \\ &= e^{X(s)} + \int_s^t e^{X(u-)} u_1 dW_u - \int_s^t e^{X(u-)} \left[ \frac{1}{2} u_1^2 + \lambda(\phi(u_2) - 1) \right] du \\ &\quad + \int_s^t e^{X(u-)} \frac{1}{2} u_1^2 du + \int_s^t e^{X(u-)} dH(u) \\ &= e^{X(s)} + \int_s^t e^{X(u-)} u_1 dW_u + \int_s^t e^{X(u-)} d\left[ H(u) - \lambda(\phi(u_2) - 1) u \right]. \end{split}$$

Thus, Y(t) is a martingale, so by the optional stopping theorem,

$$E[e^{u_1N_1(t)+u_2N_2(t)-\lambda_1t(e^{u_1}-1)-\lambda_2t(e^{u_2}-1)}] = E[Y(0)] = 1$$

$$\implies E[e^{u_1N_1(t)+u_2N_2(t)}] = e^{\lambda_1t(e^{u_1}-1)}e^{\lambda_2t(e^{u_2}-1)} = E[e^{u_1N_1(t)}]E[e^{u_2N_2(t)}].$$

which implies that  $N_1(t)$  and  $N_2(t)$  are independent.

#### Problem 5 (change of measure for stock price, with jumps)

We are given the SDE

$$dS(t) = \mu S(t)dt + \sigma S(t)dB(t) + S(t-)dQ(t).$$

In the no-jumps case (where  $y_m = 0$  for all m, for instance, so that Q(t) = 0 and  $\beta = 0$ ), we merely change the measure via the market price of risk process  $\theta = \theta(t) = \frac{\mu - r}{\sigma}$ . Girsanov's theorem then says that, for fixed T > 0,

$$\tilde{B}(t) = B(t) + \int_0^t \theta(s)ds = B(t) + \left(\frac{\mu - r}{\sigma}\right)t$$

$$Z(t) = \exp\left[-\int_0^t \theta(s)dB(s) - \frac{1}{2}\int_0^t \theta(s)^2ds\right] = \exp\left[-\theta B(t) - \frac{1}{2}\theta^2t\right], \ 0 \le t \le T$$

yields the  $\tilde{P}$ -martingale  $(\tilde{P}(A) = \int_A Z(T) dP$  for  $A \in \mathcal{F}(T)$ )

$$\begin{split} d(e^{-rt}S(t)) &= e^{-rt} \left( -rS(t)dt + dS(t) \right) \\ &= e^{-rt} \left( (\mu - r)S(t)dt + \sigma S(t)dB(t) \right) \\ &= e^{-rt} \left( (\mu - r)S(t)dt + \sigma S(t)d\tilde{B}(t) - \sigma \theta(t)dt \right) = e^{-rt}S(t)d\tilde{B}(t). \end{split}$$

Rewriting the SDE in terms of the martingale  $M_Q(t)$  instead of Q(t) yields (where, on the continuous measure dt, we can swap S(t) with S(t-1)

$$dS(t) = \mu S(t)dt + \sigma S(t)dB(t) + S(t-)d[Q(t) - \beta \lambda t + \beta \lambda t]$$
  
=  $(\mu + \beta \lambda)S(t)dt + \sigma S(t)dB(t) + S(t-)dM_O(t).$ 

With jumps in play, we still need the dt term to be zero to make  $e^{-rt}S(t)$  a  $\tilde{P}$ -martingale. To change measure, we can change the Poisson intensity from  $\lambda$  to  $\tilde{\lambda}$ , and change the probabilities of  $Y_1$  to  $\tilde{p}_m = \tilde{P}(Y_1 = y_m)$  so that  $\tilde{\beta} = \tilde{E}(Y_1) = \sum_{m=1}^M \tilde{p}_m y_m$ . However, we cannot change the  $y_m$ . We have

$$\begin{split} \tilde{B}(t) &= B(t) + \int_0^t \theta(s) ds = B(t) + \theta t \\ \tilde{M}_Q(t) &= Q(t) - \tilde{\beta} \tilde{\lambda} t \end{split}$$

with the constant  $\theta$  to be determined, yielding the  $\tilde{P}$ -martingale

$$\begin{split} d(e^{-rt}S(t)) &= e^{-rt} \left( -rS(t)dt + dS(t) \right) \\ &= e^{-rt} \left( (\mu - r)S(t)dt + \sigma S(t)dB(t) + S(t-)dQ(t) \right) \\ &= e^{-rt} \left( (\mu - r - \sigma\theta + \tilde{\beta}\tilde{\lambda})S(t)dt + \sigma S(t)d\tilde{B}(t) + S(t-)d\tilde{M}_Q(t) \right), \end{split}$$

which is only a martingale if the dt term is zero, with

$$\theta = \frac{\mu - r + \tilde{\beta}\tilde{\lambda}}{\sigma}.$$

Now, defining  $\tilde{\lambda} = \sum_{m=1}^{M} \tilde{\lambda}_m$ , where  $\tilde{p}_m = \frac{\tilde{\lambda}_m}{\tilde{\lambda}}$  is the change of measure on  $Y_1$  (from  $p_m = \frac{\lambda_m}{\lambda}$ ). Thus, our change of measure process changes B(t) to  $\tilde{B}(t)$  and  $M_Q(t)$  to  $\tilde{M}_Q(t)$ :

$$Z(t) = \exp\left[-\theta B(t) - \frac{1}{2}\theta(s)^2 t + (\lambda - \tilde{\lambda})t\right] \prod_{j=1}^{N(t)} \frac{\tilde{\lambda}\tilde{p}(Y_j)}{\lambda p(Y_j)}$$

so that, for fixed  $T>0,\, \tilde{P}(A)=\int_A Z(T)dP,\, A\in\mathcal{F}(T)$  makes  $e^{-rt}S(t)$  a  $\tilde{P}$ -martingale.

#### Problem 6 (price a BSM-Euro call with jumps)

We found the risk-neutral measure in #5 above, so we are granted the risk-neutral measure here: we are given the stock price model  $S(t) = S^*(t)e^{Q(t)}$  with

$$S^*(t) = S(0)e^{\sigma B(t) + \mu t}, \ \mu = r - \frac{\sigma^2}{2} - \lambda(E(e^{Y_1}) - 1); \ Q(t) = \sum_{j=1}^{N(t)} Y_j, \ Y_j \sim N(\mu_0, \sigma_0^2).$$

Thus, the time-zero cost of a European call option with strike price K and expiration t is simply the risk-neutral discounted expected value of the payoff function, as usual:

$$C(0) = e^{-rt}E[(S(t) - K)^{+}].$$

To calculate this expected value, we can condition on the number of jumps up to time t, i.e. on the event  $\{N(t) = n\}$ . If N(t) = n, then Q(t) is a normal random variable, independent of B(t), and so we can add all the normals together and scale to a standard normal Z to find when S(t) > K

probabilistically.

$$N(t) = n, S(t) > K \implies \mu t + \sigma B(t) + Q(t) > \ln\left(\frac{K}{S(0)}\right)$$

$$\implies \mu t + \left[\sigma B(t) + \sum_{j=1}^{N(t)=n} Y_j\right] > \ln\left(\frac{K}{S(0)}\right) \qquad \text{(sum of n+1 indep. normals)}$$

$$\implies \mu t + \mu_0 n + Z\sqrt{\sigma_0^2 n + \sigma^2 t} > \ln\left(\frac{K}{S(0)}\right) \qquad (Z \sim N(0, 1))$$

$$\implies Z > \frac{\ln\left(\frac{K}{S(0)}\right) - (\mu t + \mu_0 n)}{\sqrt{\sigma_0^2 n + \sigma^2 t}}$$

$$\implies Z' \stackrel{\text{(d)}}{=} -Z < \frac{\ln\left(\frac{S(0)}{K}\right) + (\mu t + \mu_0 n)}{\sqrt{\sigma_0^2 n + \sigma^2 t}} =: d_-(n) \quad (Z' \sim N(0, 1)).$$

Thus, the price of the call option is

$$\begin{split} C(0) &= e^{-rt} E[(S(t) - K)^{+}] \\ &= e^{-rt} \sum_{n=0}^{\infty} E[(S(t) - K)^{+} \mid N(t) = n] P(N(t) = n) \\ &= \frac{e^{-rt - \lambda}}{\sqrt{2\pi}} \sum_{n=0}^{\infty} \frac{\lambda^{n}}{n!} \int_{-\infty}^{d_{-}(n)} \left( S(0) e^{\mu t + \mu_{0} n - x \sqrt{\sigma_{0}^{2} n + \sigma^{2} t}} - K \right) e^{-\frac{x^{2}}{2}} dx \\ &= e^{-\lambda} \sum_{n=0}^{\infty} \frac{\lambda^{n}}{n!} \left( S(0) N(d_{+}(n)) - K e^{-rt} N(d_{-}(n)) \right), \end{split}$$

Note that setting  $\lambda = 0$  kills off all the jumps and reduces this to the n = 0 case (using  $0^0 = 1$ ), which is the usual BSM-Euro call price.