

Canonical Variates (Linear Discriminant) Analysis

Overview of Discriminant Analysis

Discrimination

Given an $n \times p$ data matrix, X , and a grouping of the n specimens into g groups, find the linear combination of the variables, Xa , that best discriminates between the groups.

$$\vec{y}_{\text{discrim}} = a_1 \vec{x}_1 + a_2 \vec{x}_2 + \cdots + a_p \vec{x}_p$$

Classification

Given g groups, define a function that assigns an object with unknown assignment to the 'best' group.

Fisher's Discriminant Function

- Applies to the two-group case.
- Solution: find \mathbf{a} that maximizes the ratio of the squared group mean difference to within-group variance:

$$F = \frac{(\bar{\mathbf{x}}_1 \mathbf{a} - \bar{\mathbf{x}}_2 \mathbf{a})^2}{\mathbf{a}' \mathbf{W} \mathbf{a}}$$

where

- $\bar{\mathbf{x}}_1 = \frac{1}{n_1} \sum_{i=1}^{n_1} \mathbf{x}_{1i}$ (row-vector of means of group 1)
- $\bar{\mathbf{x}}_2 = \frac{1}{n_2} \sum_{i=1}^{n_2} \mathbf{x}_{2i}$ (row-vector of means of group 2)
- $\mathbf{W} = \frac{1}{n_1 + n_2 - 2} \sum_{i=1}^2 \sum_{j=1}^{n_i} (\mathbf{x}_{ij} - \bar{\mathbf{x}}_i)(\mathbf{x}_{ij} - \bar{\mathbf{x}}_i)'$ (w/in-group pooled covariance matrix)
- n_i indicates the number of observations in the i th group and the $\mathbf{x}_{i1}, \dots, \mathbf{x}_{in_i}$ represent the specific observations (as vectors).

Geometry of the Two-Group Discriminant Function

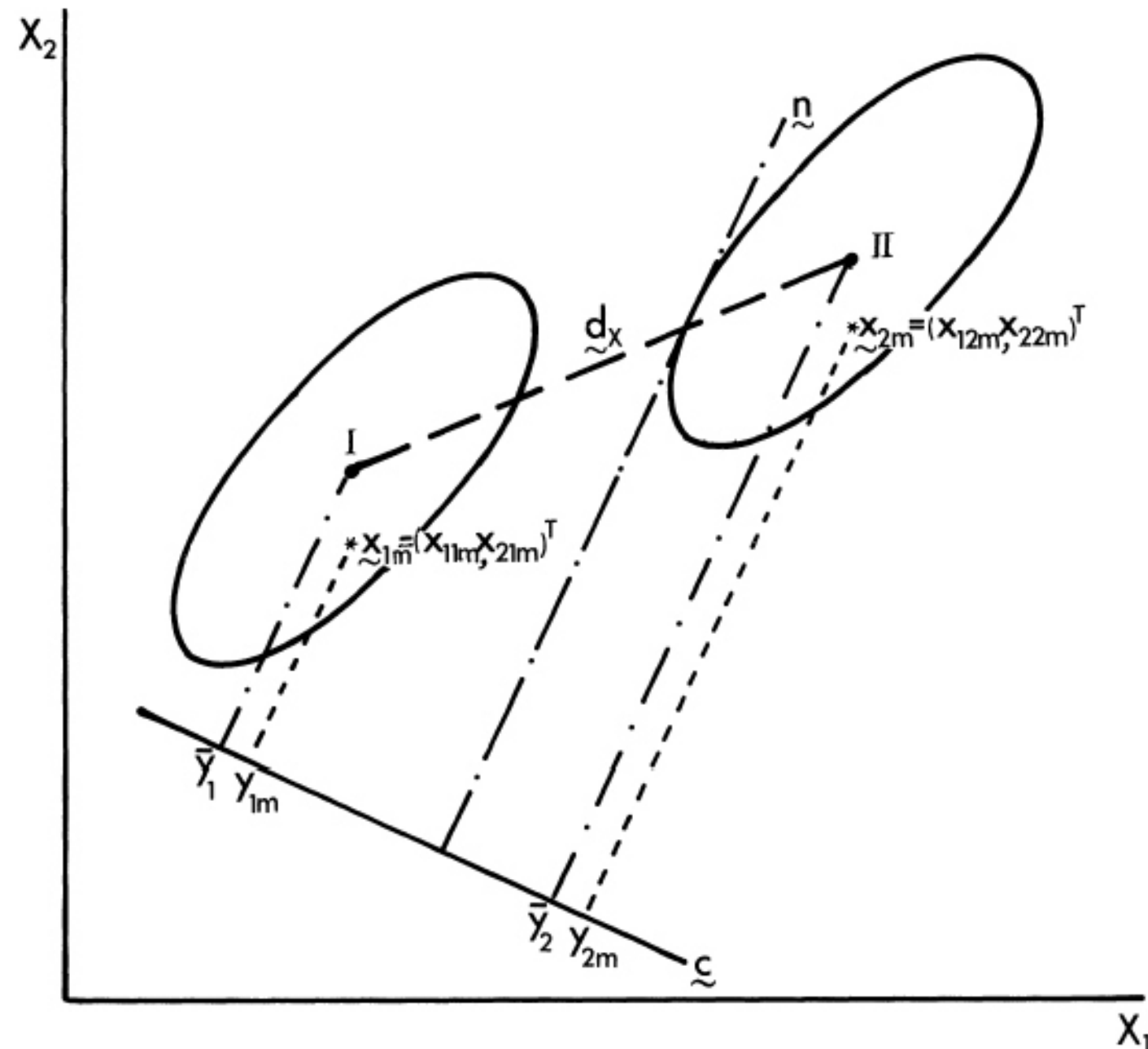


FIG. 4.—Representation of the discriminant function for two groups and two variables, showing the group means and associated 95% concentration ellipses. The vector c is the discriminant vector. The points \bar{y}_1 and \bar{y}_2 represent the discriminant means for the two groups.

The discriminant vector can be constructed by drawing the tangent n to the concentration ellipse at the point of intersection with the line d joining the group means; the discriminant vector is orthogonal to the tangent n .

Fisher's LDF

$$F = \frac{(\mathbf{a}' \bar{\mathbf{x}}_1 - \mathbf{a}' \bar{\mathbf{x}}_2)^2}{\mathbf{a}' \mathbf{W} \mathbf{a}}$$

Maximizing F gives:

$$\mathbf{a} = c \mathbf{W}^{-1} (\bar{\mathbf{x}}_1 - \bar{\mathbf{x}}_2)'$$

where c is an arbitrary constant (usually taken to be 1).

Fisher's LDF as Classification

Fisher's solution can also be setup as a classification solution using regression.

- setup a dummy variable, y that takes the values:
 - $y_1 = n_2 / (n_1 + n_2)$ for observations in group 1
 - $y_2 = -n_1 / (n_1 + n_2)$ for observations in group 2
- Solve the standard multivariate regression, $y = Xb + e$
- Allocate unknown individual to group 1 if it's predicted y is closer to y_1 than to y_2 , otherwise assign to group 2.

What if there are more than two groups?

The multi-group equivalent of Fisher's LDF is called 'Canonical Variate Analysis' (CVA).

- straight forward extension of Fisher's solution
- Find \mathbf{a} that maximizes the ratio of between-group to within-group variance:

$$F = \frac{\mathbf{a}' \mathbf{B} \mathbf{a}}{\mathbf{a}' \mathbf{W} \mathbf{a}}$$

- \mathbf{W} is within-group matrix (as defined previously)
- \mathbf{B} is the between-group covariance matrix
 - $\mathbf{B}_w = \frac{1}{g-1} \sum_{i=1}^g n_i (\bar{\mathbf{x}}_i - \bar{\mathbf{x}})(\bar{\mathbf{x}}_i - \bar{\mathbf{x}})'$ where $\bar{\mathbf{x}}$ is the "grand-mean", $\bar{\mathbf{x}}_i$ is the mean in group i , and n_i is the sample size in group i (weighted version)
 - $\mathbf{B}_u = \frac{1}{g-1} \sum_{i=1}^g (\bar{\mathbf{x}}_i - \bar{\mathbf{x}})(\bar{\mathbf{x}}_i - \bar{\mathbf{x}})'$ (unweighted version)

Geometry of CVA

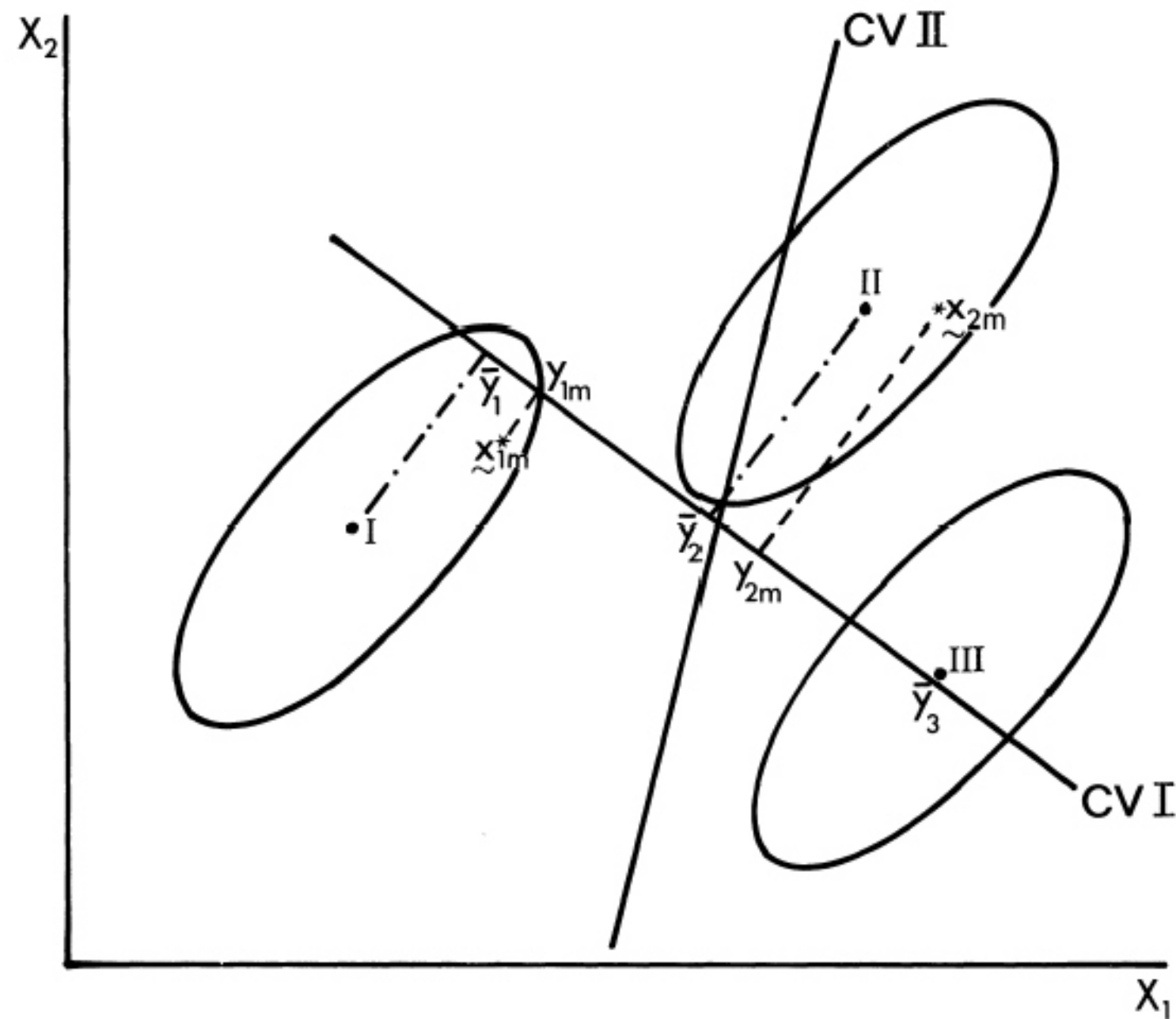


FIG. 2.—Representation of the canonical vectors for three groups and two variables. The group means (I, II and III) and 95% concentration ellipses are shown. The vectors CVI and CVII are the two canonical vectors. In the text, CVI = \mathbf{c} . The points y_{1m} and y_{2m} represent the canonical variate scores corresponding to the first canonical vector for the observations \mathbf{x}_{1m} and \mathbf{x}_{2m} .

CVA Solution

Maximizing F leads to the following:

$$(B - lW)a = 0$$

- l is an eigenvalue of $W^{-1}B$
- a is an eigenvector of $W^{-1}B$

There will be $s = \min(p, g - 1)$ non-zero eigenvalues.

Organize the eigenvectors, a_i , as columns of a $p \times s$ matrix A .

- The ***canonical variates*** are given by $Y = XA$
- The mean of the i -th group in the canonical variates space is given by $\bar{y}_i = \bar{x}_i A$, where \bar{x}_i is the mean row-vector for group i .

CVA as a two-stage rotation I

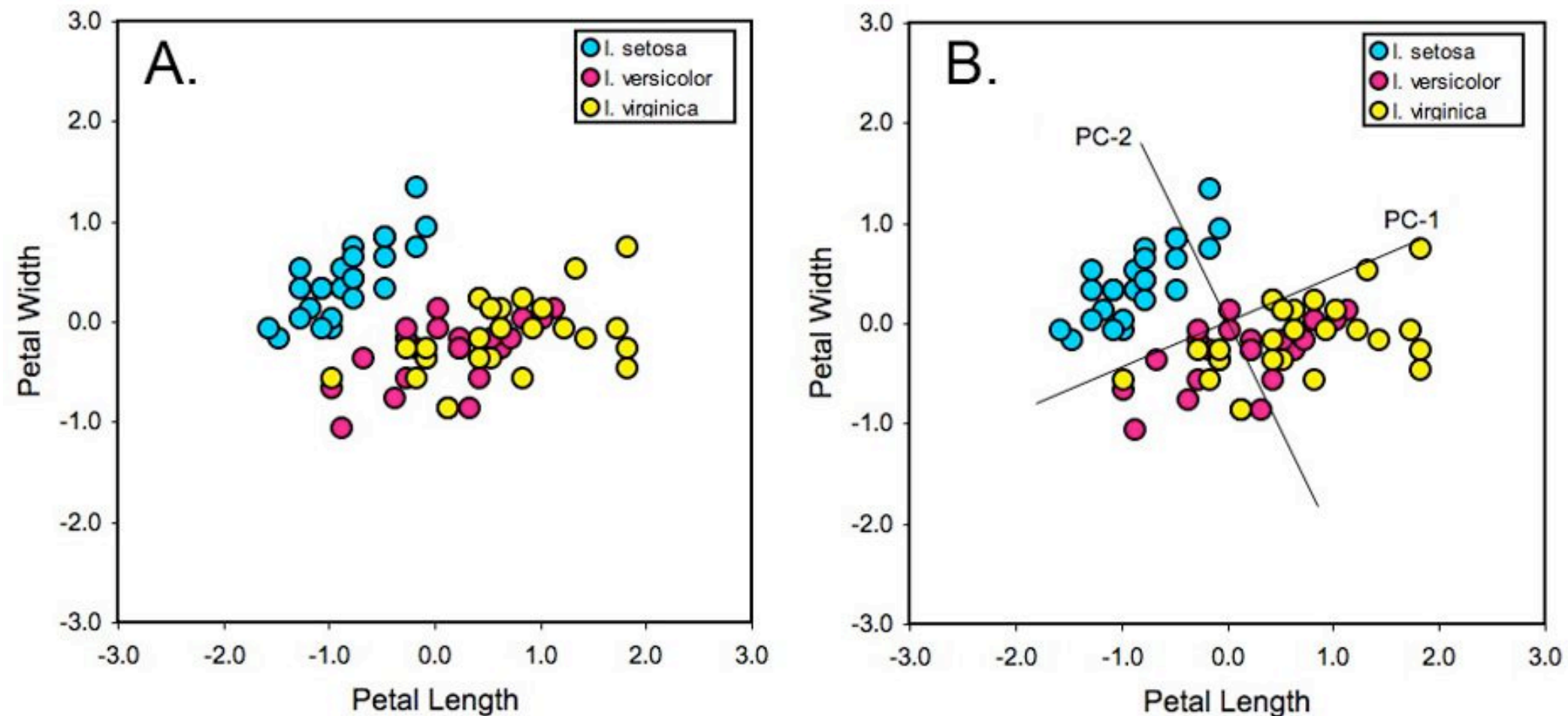


Figure 2. Stage 1 CVA implicit rotation. A. Scatterplot of first two *Iris* variables for example dataset. B. Orientation of the two pooled-sample principal components of the within-groups SSQCP matrix (W).

From MacLeod 2010, PalaeoMath newsletter

CVA as a two-stage rotation II

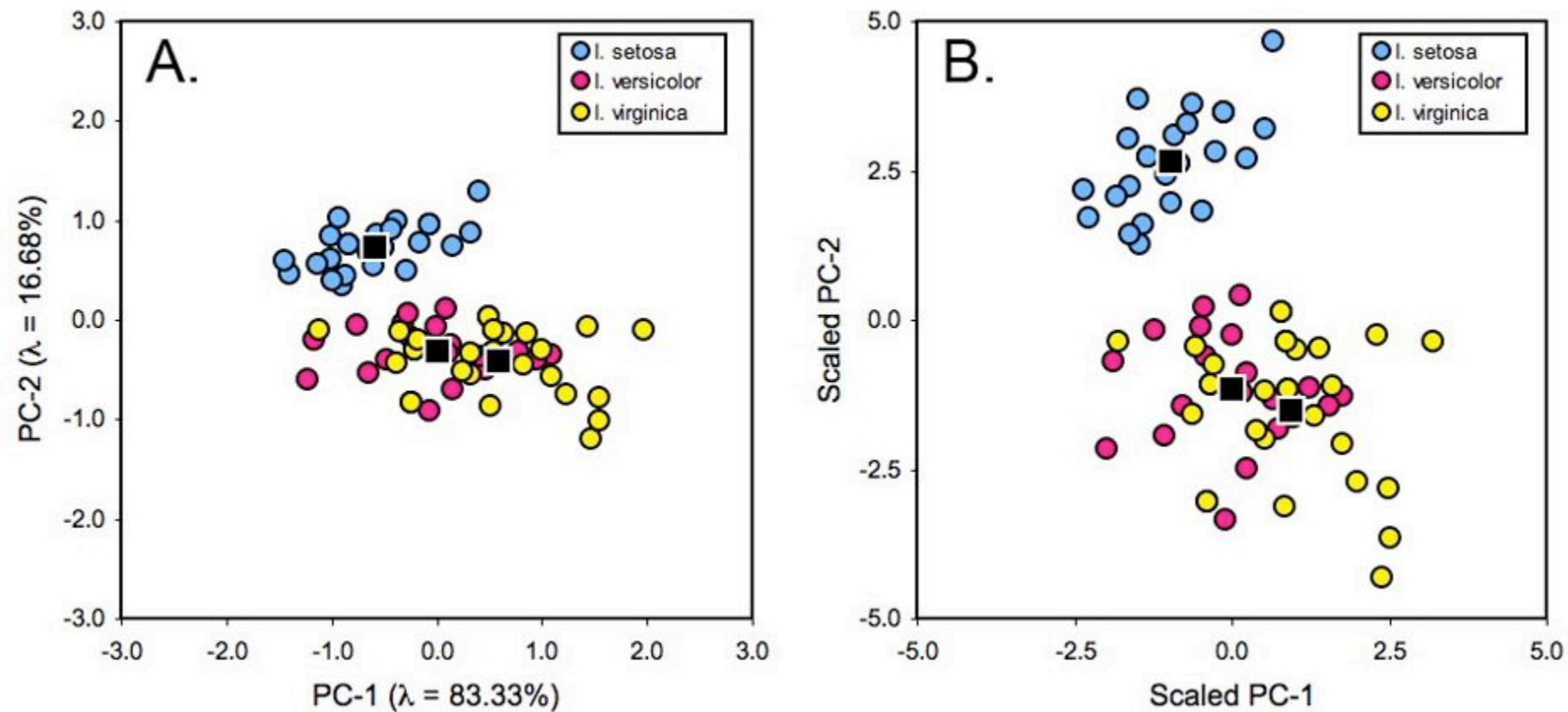


Figure 3. Intermediate scaling operation of a CVA. A. Scatterplot of *Iris* PC scores for the Stage 1 rotation (see Fig. 2). B. Result of scaling the two within-groups principal components by the square roots of their associated eigenvalues. Note difference in separation of the group centroids (black squares) after scaling.

From MacLeod 2010, PalaeoMath newsletter

CVA as a two-stage rotation III

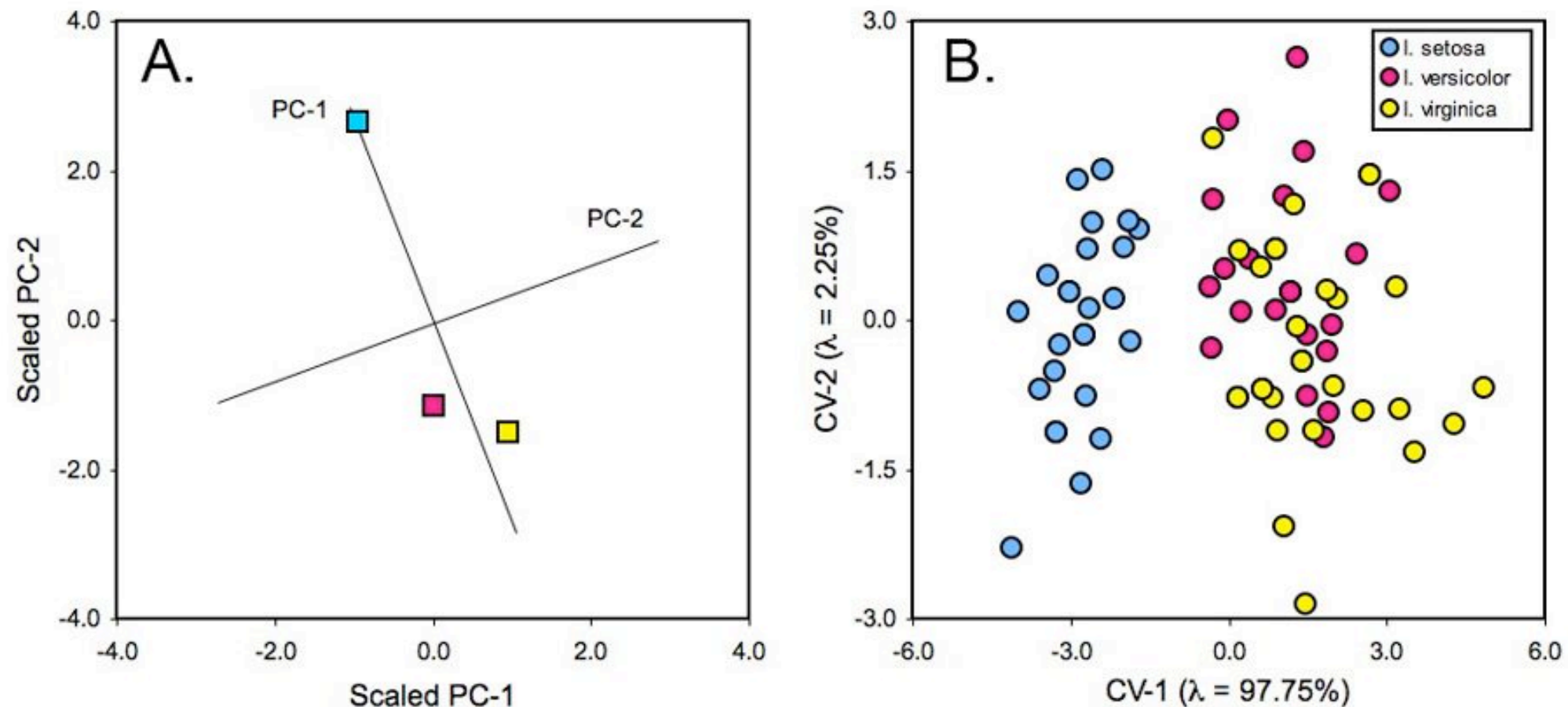


Figure 4. Stage 2 CVA implicit rotation. A. *Iris* group centroids plotted in the within-groups orthogonal-orthonormal space (see Fig. 3B) with between groups PC (= CVA) axes. B. Reduced *Iris* dataset plotted in the space defined by the CVA axes.

CVA as a two-stage rotation IV

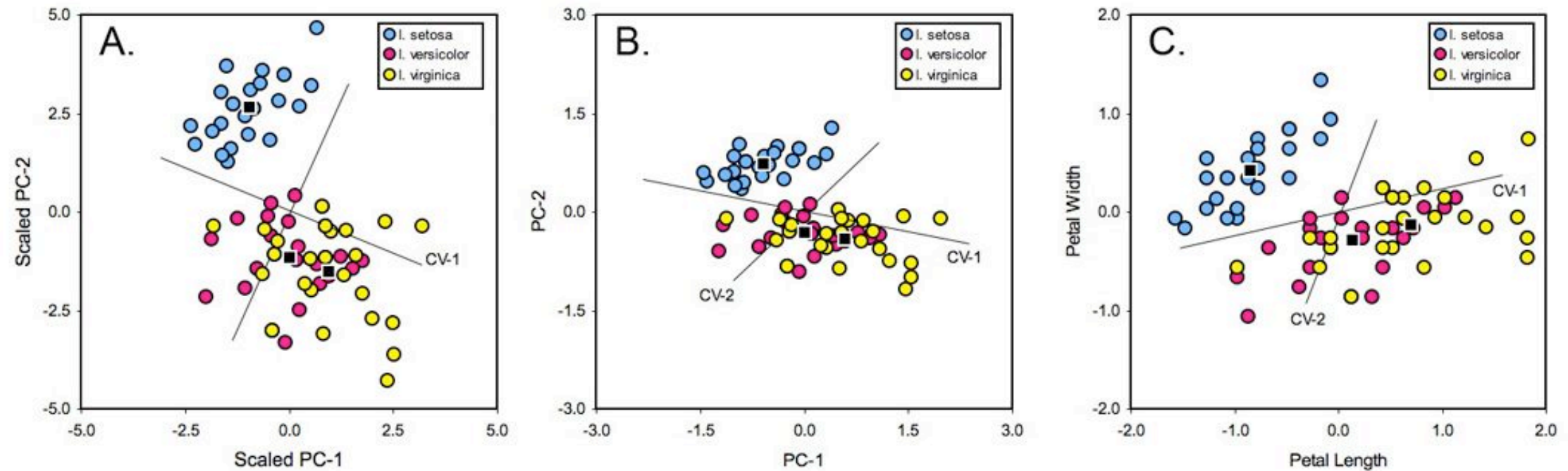


Figure 5. Back-calculation of final CVA axis orientation through the intermediate stages of the canonical rotations and scalings. A. Orientation of final CVA axes in the space of the scaled within-groups principal components (compare to Fig. 3A). B. Orientation of final CVA axes in the space of the raw within-groups principal components (compare to Fig. 3B). C. Orientation of final CVA axes in the space of the original variables (compare to Fig. 2).

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Similarities and Differences between CVA and PCA

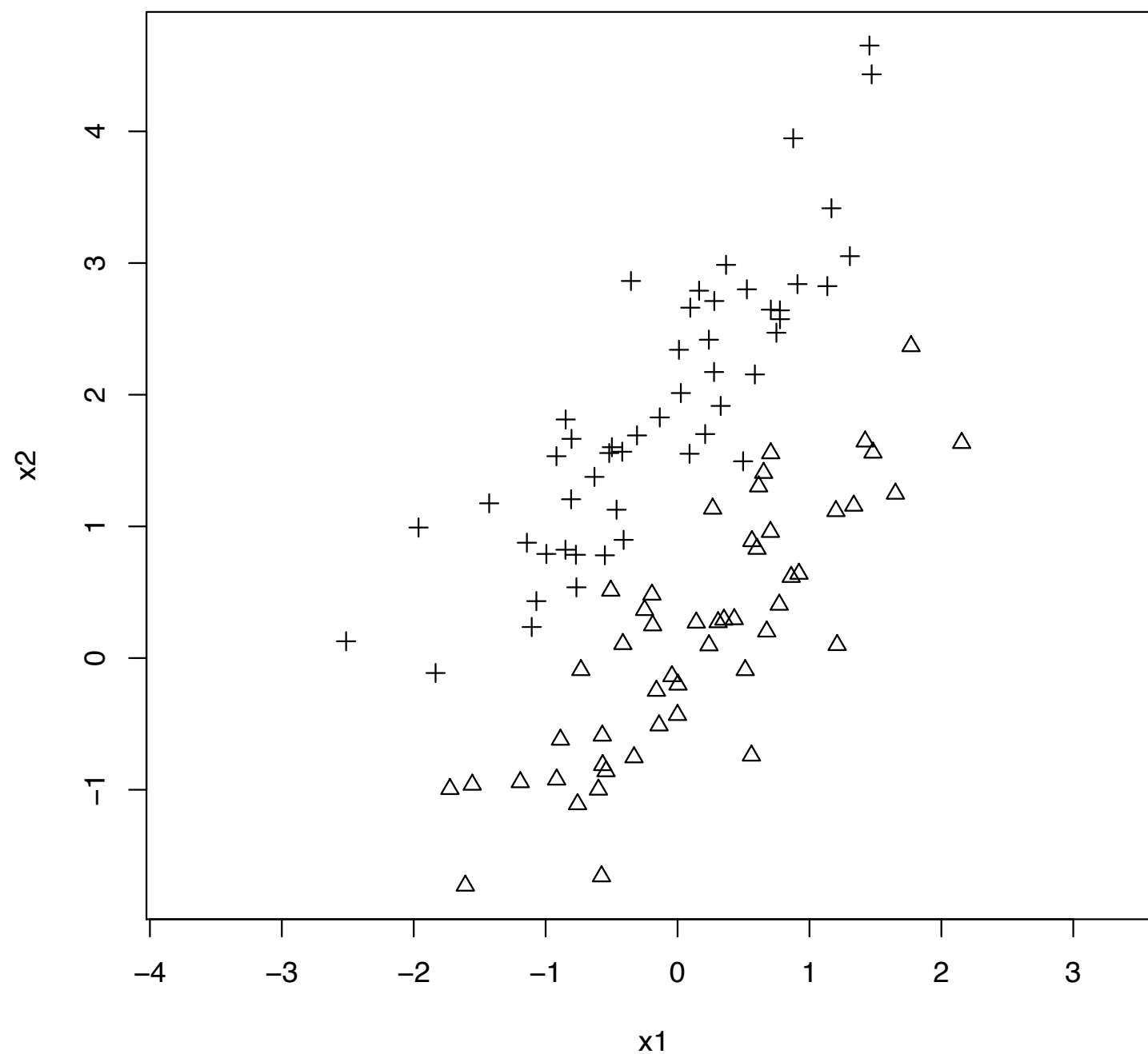
PCA:

- Uncorrelated over the whole sample
- orthogonal transformation from the original variates, x , to the new variates y . PC axes at right angles to each other in the space of the original variables.

CVA:

- Canonical variates are uncorrelated both *within* and *between* groups
- Canonical variates have equal variance *within* groups, but in decreasing order *between* groups
- non-orthogonal transformation, CV axes *not* at right angle to each other in the original frame of reference.

PCA vs CVA: A Motivating Example



What is the direction of PC1? What is the direction of CVI?

Are any of the groups significantly different in the canonical variate space?

To test:

- $H_0 : \mu_1 = \mu_2 = \dots = \mu_3$
- H_1 : at least one μ_i differs from the rest

A couple of approaches:

- Compare the largest eigenvalue, l_1 , of $W^{-1}B$ to critical values in a F-table. H_0 is rejected for large values (> 1).
- Likelihood approach:
 - Wilks' lambda, $\Lambda = |W|/|B + W| = \prod_{i=1}^p (1 + l_i)^{-1}$
 - there is an approximation that has a χ^2 distribution.

Both boil down to a consideration of eigenvalues of $W^{-1}B$.

Which groups are different? Where does an unassigned observation belong?

Within groups the canonical variates are:

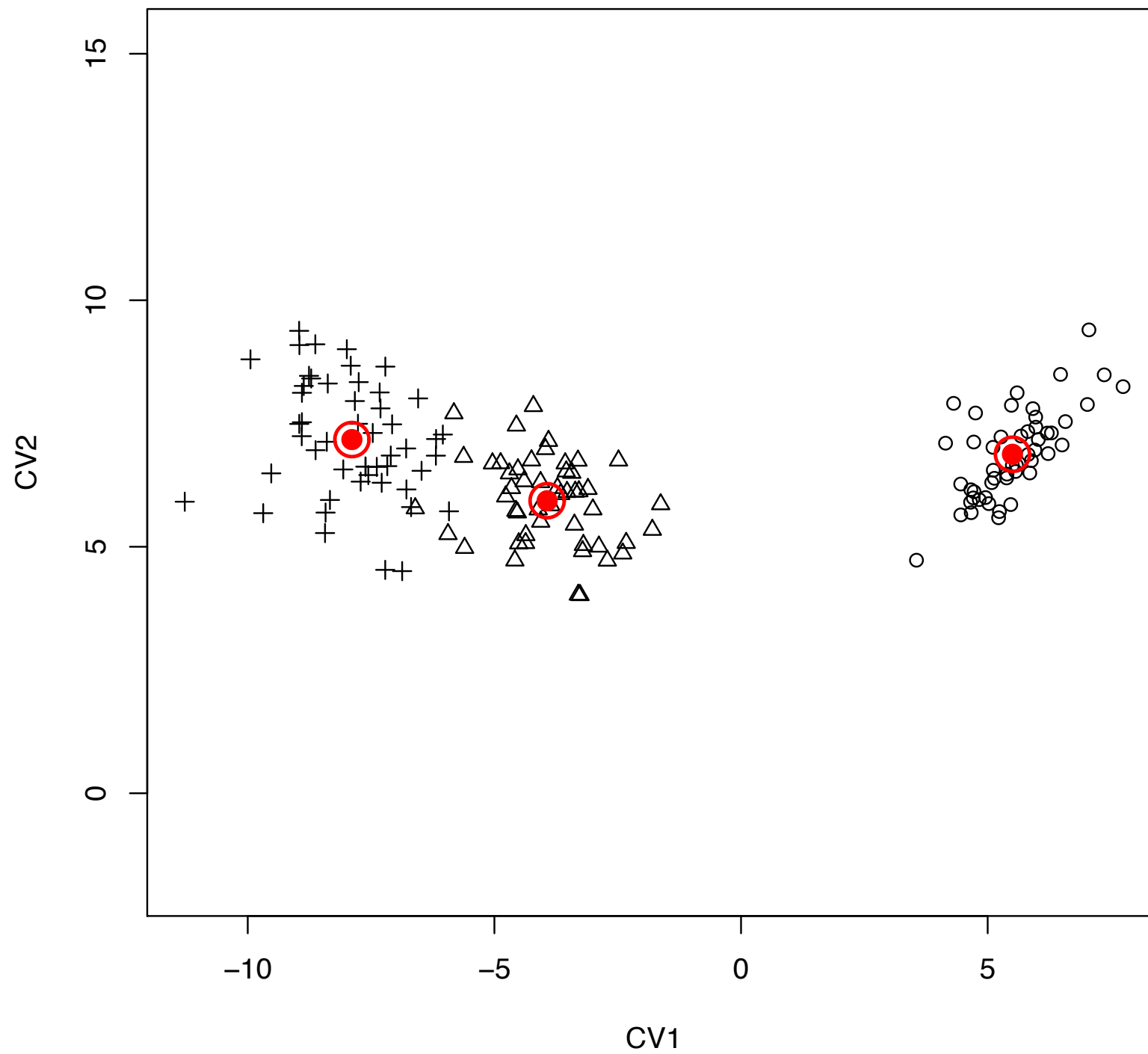
- uncorrelated
- have unit variance

If we assume multivariate normality of the data then we can exploit this to draw confidence intervals around the group means in the canonical variate space.

A $100(1-\alpha)$ percent confidence region for the true mean \mathbf{v}_i is given by:

- hypersphere centered at $\bar{\mathbf{y}}_i$
- with radius $(\chi^2_{\alpha,r}/n_i)^{1/2}$ where r is the number of canonical variate dimensions considered

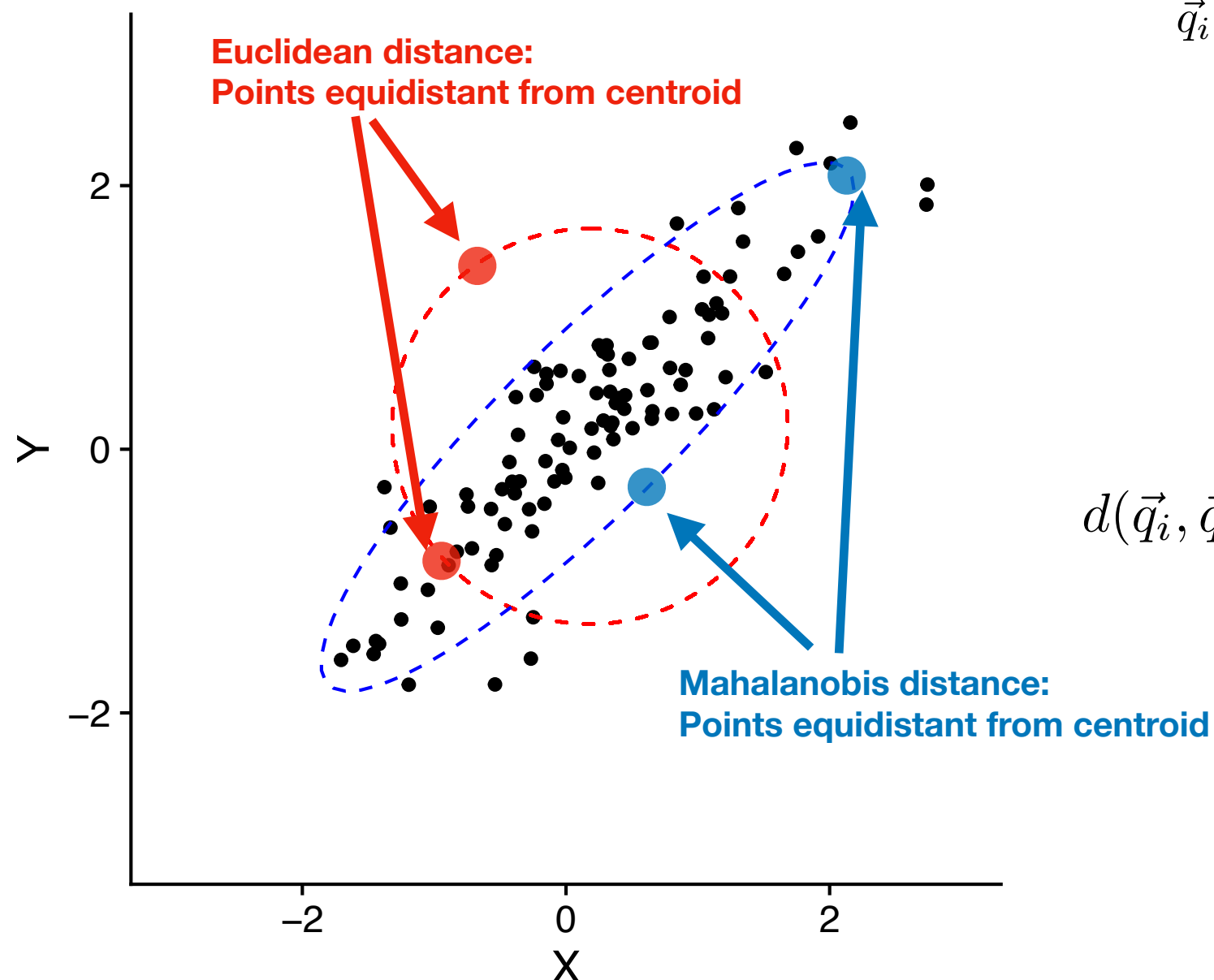
Illustration of group means and tolerance regions



Mahalanobis distance

Mahalanobis distance takes into account the covariance structure of the data.

- Deviations in the directions of greatest variance are "downweighted"



\vec{q}_i, \vec{q}_j : points i and j

\mathbf{S} : covariance matrix

$$\vec{q}_i = (\mathbf{X}_{i,1}, \mathbf{X}_{i,2}, \dots, \mathbf{X}_{i,p})$$

$$\vec{q}_j = (\mathbf{X}_{j,1}, \mathbf{X}_{j,2}, \dots, \mathbf{X}_{j,p})$$

$$d(\vec{q}_i, \vec{q}_j) = \sqrt{(\vec{q}_i - \vec{q}_j)^T \mathbf{S}^{-1} (\vec{q}_i - \vec{q}_j)}$$