

Appendix: Identifiability Under Latent Cell-State Confounding

A1. Derivation details

A1.1 Observational slope decomposition

Given $X = \alpha Z + \epsilon_x$ - $Y = \beta X + \gamma Z + \epsilon_y$ with zero-mean independent noises and $\text{Var}(Z)=1$,

$$\text{Var}(X) = \alpha^2 + \sigma_x^2$$

$$\begin{aligned} \text{Cov}(X, Y) &= \text{Cov}(X, \beta X + \gamma Z + \epsilon_y) &= \beta \\ \text{Var}(X) + \gamma \text{Cov}(X, Z) &= \beta \text{Var}(X) + \gamma \alpha \end{aligned}$$

Hence observational slope:

$$\text{Cov}(X, Y) / \text{Var}(X) = \beta + (\gamma \alpha) / \text{Var}(X)$$

which shows confounding-induced shift away from β .

A1.2 Anchor-adjusted bias formula

Let anchor $A = Z + u$, $\text{Var}(u)=\sigma_u^2$, and regress Y on (X, A) . By Frisch-Waugh-Lovell, coefficient on X equals slope of Y on residualized $X_r = X - \text{proj}_A(X)$.

$\text{proj}_A(X)$ coefficient is $\text{Cov}(X, A) / \text{Var}(A) = \alpha / (1 + \sigma_u^2)$.

$$\text{So } X_r = X - [\alpha / (1 + \sigma_u^2)] A$$

$$\text{Cov}(Z, X_r) = \alpha - \alpha / (1 + \sigma_u^2) = \alpha \sigma_u^2 / (1 + \sigma_u^2)$$

$$\begin{aligned} \text{Var}(X_r) &= \text{Var}(X) - \text{Cov}(X, A)^2 / \text{Var}(A) &= (\alpha^2 + \sigma_x^2) - \alpha^2 / (1 + \sigma_u^2) \\ &= \sigma_x^2 + \alpha^2 \sigma_u^2 / (1 + \sigma_u^2) \end{aligned}$$

Bias term from residual confounding is

$$\begin{aligned} \text{Bias_anchor} &= \gamma * \text{Cov}(Z, X_r) / \text{Var}(X_r) &= \gamma * \\ &\alpha * \sigma_u^2 / [\sigma_x^2 (1 + \sigma_u^2) + \alpha^2 \sigma_u^2]. \end{aligned}$$

This matches observed degradation with larger anchor noise in simulations.

A2. Output inventory

Synthetic suite outputs: - /Users/ihorkendiukhov/biodyn-work/subproject_10_identifiability_latent_
- /Users/ihorkendiukhov/biodyn-work/subproject_10_identifiability_latent_cell_state_confoundi
- /Users/ihorkendiukhov/biodyn-work/subproject_10_identifiability_latent_cell_state_confoundi
- /Users/ihorkendiukhov/biodyn-work/subproject_10_identifiability_latent_cell_state_confoundi
- /Users/ihorkendiukhov/biodyn-work/subproject_10_identifiability_latent_cell_state_confoundi
- /Users/ihorkendiukhov/biodyn-work/subproject_10_identifiability_latent_cell_state_confoundi

Tissue diagnostics outputs: - /Users/ihorkendiukhov/biodyn-work/subproject_10_identifiability_latent_cell_state_confounding
- /Users/ihorkendiukhov/biodyn-work/subproject_10_identifiability_latent_cell_state_confounding
- /Users/ihorkendiukhov/biodyn-work/subproject_10_identifiability_latent_cell_state_confounding

A3. Additional result slices

A3.1 Best and worst environment-stress cells

From `synthetic_environment_stress.csv`:

Best gain cell: - `alpha_shift=0.0`, variance ratio 1.8 - pooled RMSE 0.6280 - invariance RMSE 0.1133 - gain +0.5147

Worst gain cell: - `alpha_shift=0.0`, variance ratio 1.0 - pooled RMSE 0.7484 - invariance RMSE 4.4592 - gain -3.7109

Interpretation: invariance estimator can be numerically explosive when environments are not sufficiently separated.

A3.2 Tissue pass/fail summary

From `tissue_edge_invariance_fits.csv`: - Pass edges: 22/76. - Median R^2 pass set: 0.738. - Median R^2 fail set: 0.130.

Top pass edges by fit: 1. RELA -> CXCL8 ($R^2=0.989$, RMSE=0.0037) 2. GATA2 -> NR3C2 ($R^2=0.981$, RMSE=0.0622) 3. NFKB2 -> CXCL8 ($R^2=0.952$, RMSE=0.0229)

Worst-fit examples: 1. HIF1A -> XXYLT1 ($R^2=0.0006$) 2. STAT3 -> TYK2 ($R^2=0.0009$) 3. GATA3 -> RBMS1 ($R^2=0.0022$)

A4. Practical implementation note

For production use, the environment-based estimator should include a hard precondition check: - reject estimation when $|1/\text{VarX1} - 1/\text{VarX2}| < \text{epsilon}$.

This prevents unstable divisions and aligns estimator behavior with identifiability requirements.