#### Biomedical Engineering Degree

## 3. Hypothesis Testing: Two-Sample Inference

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#### References

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## Outline

Introduction

- Testing for the equality of two means
  - Paired samples
  - Independent Samples

Testing for the equality of two variances

#### Introduction

We want to compare two populations

$$X_1 \sim \mathcal{N}(\mu_1, \sigma_1^2)$$
 vs  $X_2 \sim \mathcal{N}(\mu_2, \sigma_2^2)$ 

by using **two random samples** of  $n_1$  and  $n_2$  **observations** from  $X_1$  and  $X_2$ , respectively. We are interested in:

- Testing for the equality of two population means:
  - Paired or matched samples
  - Independent samples
    - Equal vs unequal variances
- Testing for the equality of two population variances
  - Independent samples



## Paired vs Independent samples

- Two samples are said to be **paired** when each data point in the first sample is matched and is related to a unique data point in the second sample.
  - ▶ Prices of books in Casa del libro vs Amazon
  - Longitudinal or follow-up studies, where the same group of people is followed over time.
- Two samples are said to be independent when the data points in one sample are unrelated to the data points in the second sample
  - Cross-sectional studies, where the participants are seen at only one point in time.

## Outline

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  - Paired samples
  - Independent Samples

Testing for the equality of two variances

#### Paired t Test

- ullet Let  $X_1$  be a population with mean  $\mu_1$ , and  $X_2$  be a population with mean  $\mu_2$
- ullet Suppose we have a random sample of n paired observations from these two populations and let

$$d_{1} = x_{1,1} - x_{1,2}$$

$$d_{2} = x_{2,1} - x_{2,2}$$

$$\vdots$$

$$d_{n} = x_{n,1} - x_{n,2}$$

represent n differences with

- Sample mean:  $\overline{d}$
- Quasi-standard deviation<sup>1</sup>: s<sub>d</sub>

Let assume that the population of differences is normal  $D \sim \mathcal{N}(\mu_d, \sigma_d)$ 

## Example

i	SBP level while not using OCs $(x_{i1})$	SBP level while using OCs $(x_{12})$	$d_i^{\star}$
1	115	128	13
2	112	115	3
3	107	106	-1
4	119	128	9
5	115	122	7
6	138	145	7
7	126	132	6
8	105	109	4
9	104	102	-2
10	115	117	2

#### Where:

$$\bullet$$
  $\overline{d} = \frac{1}{10} \sum_{i=1}^{10} d_i = 4.8$ 

• 
$$s_d = \frac{1}{n-1} \sum_{i=1}^{10} (d_i - \overline{d})^2 = 5.56$$



#### Paired t Test

To test the hypothesis

$$H_0: \mu_d = 0 \text{ vs } H_1: \mu_d \neq 0$$

with unknown  $\sigma_d$  with a significance level of  $\alpha$ 

#### Paired t-Test

Compute

$$t = \frac{\bar{d}}{s_d/\sqrt{n}} \sim t_{n-1}$$

- if  $|t| > t_{n-1,1-\alpha/2}$ , then we reject  $H_0$
- if  $|t| \leq t_{n-1,1-\alpha/2}$ , then we accept  $H_0$

## Paired t Test. Confidence interval

To test the hypothesis

$$H_0: \mu_d = 0 \text{ vs } H_1: \mu_d \neq 0$$

with unknown  $\sigma_d$  with a significance level of  $\alpha$ 

#### Paired t-Test

Compute the confidence interval

$$\bar{d} \pm t_{n-1,1-\alpha/2} \frac{s_d}{\sqrt{n}}$$

- if  $100\%(1-\alpha)$  CI does not contain 0, then we reject  $H_0$
- if  $100\%(1-\alpha)$  CI does contain 0, then we accept  $H_0$

## Example

#### Paired t-Test

Assess the statistical significance of the example data shown in the previous table. Use:

- The critical-value method
  - p-value method

## **Example solution**

Using the critical-value approach, first we compute the statistic

$$t = \frac{\bar{d}}{s_d/\sqrt{n}} = \frac{4.8}{4.56/\sqrt{10}} = 3.32$$

and then:  $t_{n-1,1-\alpha/2}={ t (df=9).ppf(0.975)}=2.262.$  Thus,  $t\geq t_{9,0.975}$ 

ullet On the other hand, using the p-value approach, we calculate

$$\begin{array}{lcl} p = 2*P(t_9 > 3.32) & = & 2\cdot(1-P(t_9 \leq 3.32)) = \\ & = & 2*(1-\texttt{t(df=9).cdf(3.32)}) = 0.0089 \end{array}$$

then, the results are statistically significant to reject  $H_0$ 

We REJECT the null hypothesis  $H_0$  at a significance level of 0.05



## t—Test for independent samples with equal variances

- ullet Let  $X_1$  be a population with mean  $\mu_1$  and variance  $\sigma_1^2$
- ullet Let  $X_2$  be a population with mean  $\mu_2$  and variance  $\sigma_2^2$

Let assume that both population are **normally distributed** with unknown but equal variances

$$\sigma^2=\sigma_1^2=\sigma_2^2$$

- Suppose we have a random sample of  $n_1$  observations from  $X_1$  and an independent random sample of  $n_2$  observations from  $X_2$
- Thus, we have access to:
  - $\bar{x}_1, s_1$ , for  $X_1$  population
  - $ightharpoonup ar x_2, s_2$ , for  $X_2$  population



#### Variance estimation

- Since both populations have equal variances, we can estimate  $\sigma$  from either  $n_1$  or  $n_2$  observations. Thus,
  - On the one hand,  $\hat{\sigma} = s_{*,1}$
  - On the other hand,  $\hat{\sigma} = s_{*,2}$
- ... or we could use both of them (weighted average):

#### Pooled estimate of the variance

$$s^{2} = \frac{(n_{1} - 1)s_{1}^{2} + (n_{2} - 1)s_{2}^{2}}{n_{1} + n_{2} - 2}$$

## t-Test for independent samples with equal variances

We want to test the hypothesis

$$H_0: \mu_1 = \mu_2 \quad \text{vs} \quad H_1: \mu_1 \neq \mu_2$$

with unknown  $\sigma$  with a significance level of  $\alpha$ 

## t-Test for independent samples with equal variances

Compute

$$t = \frac{\bar{x}_1 - \bar{x}_2}{s\sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} \sim t_{n_1 + n_2 - 2}$$

- if  $|t| > t_{n_1+n_2-2,1-\alpha/2}$ , then we reject  $H_0$
- ullet if  $|t| \leq t_{n_1+n_2-2,1-lpha/2}$ , then we accept  $H_0$

## t-Test for independent samples with equal variances

## Example

A study attempted to assess the effect of the presence of a moderator on the number of ideas generated by a group. Groups of four members, with or without moderator, were observed. For a random sample of four groups with a moderator, the mean number of ideas generated per group was 78.0, and the sample quasi-standard deviation was 24.4. For an independent sample of four groups without a moderator, the mean number of ideas generated was 63.5, and the sample quasi-standard deviation was 20.2. Assuming that the populations distributions are normal with equal variances, test the null hypothesis  $(\alpha=0.1)$  that the population means are equal against the alternative that the true mean is higher for groups with a moderator.

## Example solution

- We know that  $\bar{x}_1 = 78.0$ ,  $s_1 = 24.4$ , and  $n_1 = 4$
- We also know that  $\bar{x}_2 = 63.5$ ,  $s_1 = 20.2$ , and  $n_2 = 4$
- Then,

$$s^{2} = \frac{(n_{1} - 1)s_{1}^{2} + (n_{2} - 1)s_{2}^{2}}{n_{1} + n_{2} - 2} = \frac{(4 - 1)24.4^{2} + (4 - 1)20.2^{2}}{4 + 4 - 2} = 501.7$$

• Therefore  $s = \sqrt{501.7} = 22.4$  . Using this value, we can calculate t

$$t = \frac{\bar{x}_1 - \bar{x}_2}{s\sqrt{\frac{1}{n_1} + \frac{1}{n_2}}} = \frac{78.0 - 63.5}{22.4\sqrt{\frac{1}{4} + \frac{1}{4}}} = 0.915$$

and then  $t_{n_1+n_2-2,1-\alpha/2}=t_{4+4-2,1-0.1/2}={ t(df=6).ppf(0.95)}=1.94$ 

• Since  $t < t_{n_1 + n_2 - 2, 1 - \alpha/2}$ , then

We ACCEPT the null hypothesis  $H_0$  at a significance level of 0.1



## t-Test for independent samples with equal variances

• We want to test the hypothesis

$$H_0: \mu_1 = \mu_2 \text{ vs } H_1: \mu_1 \neq \mu_2$$

with unknown  $\sigma$  with a significance level of  $\alpha$ 

#### Compute the confidence interval

$$\bar{x}_1 - \bar{x}_2 \pm t_{n_1 + n_2 - 2, 1 - \alpha/2} \cdot s \sqrt{\frac{1}{n_1} + \frac{1}{n_2}}$$

- if  $100\%(1-\alpha)$  CI does not contain 0, then we reject  $H_0$
- if  $100\%(1-\alpha)$  CI does contain 0, then we accept  $H_0$

## Test for independent samples with unequal and known

- **Variances** Let  $X_1$  be a population with mean  $\mu_1$  and variance  $\sigma_1^2$ 
  - ullet Let  $X_2$  be a population with mean  $\mu_2$  and variance  $\sigma_2^2$

### Let assume that both population are **normally distributed with known** variances

- Suppose we have a random sample of  $n_1$  observations from  $X_1$  and an **independent** random sample of  $n_2$  observations from  $X_2$
- Thus, we have access to:
  - $\bar{x}_1, \sigma_1$ , for  $X_1$  population
  - $\bar{x}_2, \sigma_2$ , for  $X_2$  population

#### Homework!

If we want to test the hypothesis

$$H_0: \mu_1 = \mu_2 \quad \text{vs} \quad H_1: \mu_1 \neq \mu_2$$

with a significance level of  $\alpha$ , what would be test we should apply?<sup>a</sup>

<sup>a</sup>Hint: if  $\overline{X}_i \sim \mathcal{N}(\mu_i, \frac{\sigma_i}{n_i})$ , i = 1, 2, how is distributed  $\overline{X}_1 - \overline{X}_2$ ?

## Test for independent samples with **unequal and known** variances: solution

• We want to test the hypothesis

$$H_0: \mu_1 = \mu_2 \quad \text{vs} \quad H_1: \mu_1 \neq \mu_2$$

with unknown  $\sigma_1$ ,  $\sigma_2$  and a significance level of  $\alpha$ 

## $t-{\sf Test}$ for independent samples with unknown and unequal variances

#### Compute

$$z = \frac{\bar{x}_1 - \bar{x}_2}{\sqrt{\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}}} \sim \mathcal{N}(0, 1)$$

- if  $|z|>z_{1-\alpha/2}$ , then we reject  $H_0$
- if  $|z| \leq z_{1-\alpha/2}$ , then we accept  $H_0$

## Test for independent samples with **unequal and unknown** variances

- ullet Let  $X_1$  be a population with mean  $\mu_1$  and variance  $\sigma_1^2$
- ullet Let  $X_2$  be a population with mean  $\mu_2$  and variance  $\sigma_2^2$

Let assume that both population are normally distributed with unknown variances

- Suppose we have a random sample of  $n_1$  observations from  $X_1$  and an **independent** random sample of  $n_2$  observations from  $X_2$
- Thus, we have access to:
  - $\bar{x}_1, s_1$ , for  $X_1$  population
  - $ightharpoonup \bar{x}_2, s_2$ , for  $X_2$  population

# *t*—Test for independent samples with unknown and unequal variances

We want to test the hypothesis

$$H_0: \mu_1 = \mu_2 \quad \text{vs} \quad H_1: \mu_1 \neq \mu_2$$

with unknown  $\sigma_1\text{, }\sigma_2$  and a significance level of  $\alpha$ 

## $t{\operatorname{\mathsf{-Test}}}$ for independent samples with unknown and unequal variances

Compute

$$t = \frac{x_1 - x_2}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} \sim t_d \text{ (Satterthwaite approximation)}$$

- if  $|t| > t_{d,1-\alpha/2}$ , then we reject  $H_0$
- ullet if  $|t| \leq t_{d,1-lpha/2}$ , then we accept  $H_0$  where

$$d = \frac{\left(s_1^2/n_1 + s_2^2/n_2\right)^2}{\left(s_1^2/n_1\right)^2/(n_1 - 1) + \left(s_2^2/n_2\right)^2/(n_2 - 1)}$$

Also know as Welch's t-test. You can find a deep explanation here

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3 Testing for the equality of two variances

## Testing for the equality of two variances

- Let  $X_1$  be a population with mean  $\mu_1$  and variance  $\sigma_1^2$
- Let  $X_2$  be a population with mean  $\mu_2$  and variance  $\sigma_2^2$

#### Let assume that both population are normally distributed

- Suppose we have a random sample of  $n_1$  observations from  $X_1$  and an **independent** random sample of  $n_2$  observations from  $X_2$
- Thus, we have access to:
  - $\bar{x}_1, s_1$ , for  $X_1$  population
  - $\bar{x}_2, s_2$ , for  $X_2$  population

## Testing for the equality of two variances

We want to test the hypothesis

$$H_0: \sigma_1 = \sigma_2 \quad \text{vs} \quad H_1: \sigma_1 \neq \sigma_2$$

with a significance level of  $\boldsymbol{\alpha}$ 

## Testing for the equality of two variances

Compute the statistic

$$f = \frac{s_1^2}{s_2^2} \sim F_{n_1 - 1, n_2 - 1}$$
 (F distribution)

- $\bullet$  if  $f > F_{n_1-1,n_2-1,1-\alpha/2},$  or  $f < F_{n_1-1,n_2-1,\alpha/2}$  then we reject  $H_0$
- ullet Otherwise we accept  $H_0$



#### F-distribution

• if  $x_1, x_2, \ldots, x_n$  and  $y_1, y_2, \ldots, y_m$  denote r.v.'s following a  $\mathcal{N}(0, 1)$  distribution. Then, the r.v.

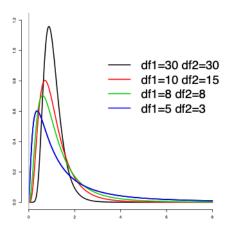
$$F = \frac{\frac{1}{n} \sum_{i=1}^{n} x_i^2}{\frac{1}{m} \sum_{i=1}^{m} y_i^2}$$

follows a  $F_{n,m}$  with n and m degrees of freedom

We can view it as a ratio of two normalized chi-square r.v.'s

$$\frac{s_1^2}{s_2^2} = \frac{\frac{1}{n_1 - 1}}{\underbrace{\frac{1}{n_1 - 1}}} \underbrace{\frac{(n_1 - 1)s_1^2}{\sigma^2}}_{\underbrace{\frac{1}{n_2 - 1}}} \sim F_{n_1 - 1, n_2 - 1}$$

## F-distribution



## Testing for the equality of two variances

### Example

For a random sample of 17 newly issued AAA-rated industrial bonds, the quasi-variance of maturities (in years squared) was 123.35. For an independent random sample of 11 issued CCC-rated industrial bonds, the quasi-variance of maturities was 8.02. If the respective population variances are denoted  $\sigma_1$  and  $\sigma_2$ , perform a two-sided test at a  $5\,\%$  level.

## **Example solution**

• Calculate the *f* statistic

$$f = \frac{s_1^2}{s_2^2} = \frac{123.35}{8.02} = 15.38 \sim F_{16,10}$$

where

- $F_{16,10,1-\alpha/2} = f(16,10).ppf(0.975) = 3.496$
- $ightharpoonup F_{16,10,lpha/2} = exttt{f(16,10).ppf(0.025)} = 0.335$

and since  $f > F_{16,10,1-\alpha/2}$  then

We REJECT the null hypothesis  $H_0$  at a significance level of 0.05

