If u is an eigenvector of $S_{W}JS_{B}$, i.e. $S_{W}JS_{B}U=\lambda U$ then $u = \frac{1}{\lambda} S_w^{-1} S_B u$ $=\frac{1}{2} \frac{1}{2} \frac{1$ $= \frac{(u^{(i)} - u^{(i)})^{T} u}{\lambda} \quad Sw \left(u^{(i)} - u^{(i)}\right)$ Thus we can take $u = \frac{S_w^{-1}(u^{(1)} - u^{(2)})}{||S_w^{-1}(u^{(1)} - u^{(2)})||}$

Remarka: Companison between PCA and LDA

(1) Both are dimension reduction techniques

projected
(2) PCA maximizes data variability,

while LDA maximizes projected data separability.

(3) PCA is unsupervised, while LDA is supervised.

4.3.3 Algorithm (LDA for binary classification)

Imput: Sample pts with labels (x^0, Y_1) , (x^0, Y_n) where $Y_i = \pm 1$

Output: 1D direction u with maximal projected data separability.

Step 1: Compute the means and sample covariance matrices of the two classes

$$\mu^{(i)} = \frac{1}{n_1} \sum_{\{i: y_i = 1\}} x^{(i)}$$

$$\mu^{(2)} = \frac{1}{n_2} \sum_{\{i: y_i = -1\}} x^{(i)}$$

$$C^{(i)} = \frac{1}{N_i} \sum_{\{j_i, j_{i=1}\}} (\chi^{(i)} - \mu^{(i)}) (\chi^{(i)} - \mu^{(i)})^T$$

$$C^{(z)} = \frac{1}{n_z} \sum_{\{j_i, j_i=1\}} (\chi^{(j)} - u^{(z)}) (\chi^{(j)} - u^{(z)})^T$$

where $n_1 = \#$ of sample pts with $x_i = 1$ $h_2 = - - - - - x_i = -1$

Step 2: Compute the between-class and within-class scatter mothix SB and Sw.

$$S_{B} = (\mathcal{M}^{(1)} - \mathcal{M}^{(1)}) (\mathcal{M}^{(1)} - \mathcal{M}^{(1)})^{\mathsf{T}}$$

$$S_{W} = n_{1} c^{(1)} + n_{2} c^{(2)}$$

Step 3: u is an eigenvector associated to the largest eigenvalue of $S_w^{-1}S_B$.

Alternatively, u can be taken as

$$U = \frac{S_{w}^{-1}(u^{(1)} - u^{(2)})}{\|S_{w}^{-1}(u^{(1)} - u^{(2)})\|}$$

Reference: "The Elements of Statistical Data Mining Inference, and Probably Hastie - Tibshirani - Eriedman, Section 4.3.	Leaving: