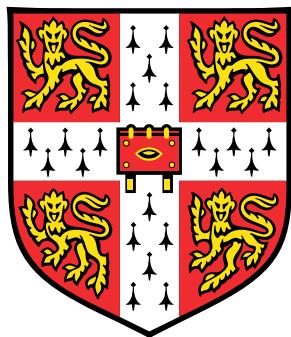


# **Optimal Importance Sampling in Quantum Monte Carlo for Lattice Models**



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This dissertation is submitted for the degree of  
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St. John's College

July 2021



I would like to dedicate this thesis to ...



## **Declaration**

I hereby declare that except where specific reference is made to the work of others, the contents of this dissertation are original and have not been submitted in whole or in part for consideration for any other degree or qualification in this, or any other university. This dissertation is my own work and contains nothing which is the outcome of work done in collaboration with others, except as specified in the text and Acknowledgements. This dissertation contains fewer than 15,000 words including appendices, figure legends, and tables.

Blaž Stojanovič  
July 2021



## **Acknowledgements**

- Zala
- Parents/Family
- Nikos/Tutor/Sally
- Prosen/Kosec+Jure
- Austen



## **Abstract**

This is where you write your abstract ...



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# Nomenclature

## Other Symbols

$\mathbb{E}$       Expectation

$\sigma\text{-field}$  ( $\sigma$ -algebra)

$\mathcal{F}$       Filtration

$\mathbb{P}$       Measure

$\mathcal{N}$       The Gaussian distribution

$\{X\}$       Random variable

$\Gamma$       Transition rate matrix

$\mathbb{R}$       The set of real numbers

$S$       State space of a Markov process

$\{X_t\}$       Stochastic process

Cov      Covariance

Var      Variance

$W_t$       Wiener process, mathematical Brownian motion

## Acronyms / Abbreviations

cdf      Cumulative density function

CNN      Convolutional Neural Network

DTMC      Discrete time Markov Chain

DL Deep Learning

DMC Diffusion Quantum Monte Carlo

DTMC Discrete time Markov Chain

e.g. Exempli gratia ("for the sake of an example")

FP Fokker-Planck

GAN General Adversarial Network

i.e. Id est ("it is")

i.i.d Independent and identically distributed

ML Machine Learning

NN Neural Network

p.b.c Periodic boundary condition

pdf Probability density function

QMC Quantum Monte Carlo

RN Radon-Nikodym

SDE Stochastic Differential Equations

s.p. Stochastic process

s.t. Such that

VAE Variational Autoencoder

VMC Variational Quantum Monte Carlo

w.r.t With respect to

# **Chapter 1**

## **Introduction**

### **1.1 Thesis Contributions**

### **1.2 Thesis Structure**



# Chapter 2

## On the quantum many-body problem

### 2.1 Lattice models

#### 2.1.1 Historical introduction

- I think starting from the magnetism point of view might be the best way to go, slowly lead into the field of condensed matter theory and lattice models.

#### 2.1.2 The Schrödinger equation and the Feynman path integral

The wavefunction

$$\Psi(r_1, \dots, r_N) \quad (2.1)$$

the Schrödinger equation

$$i \frac{\partial \psi(\mathbf{r}, t)}{\partial t} = \hat{H} \psi(\mathbf{r}, t) \quad (2.2)$$

for a single particle in an external potential  $\hat{V}(\mathbf{r})$  the Hamiltonian is

$$\hat{H} \phi(\mathbf{r}) = -\frac{1}{2} \nabla^2 \phi(\mathbf{r}) + \hat{V}(\mathbf{r}) \phi(\mathbf{r}). \quad (2.3)$$

Alternatively to the Schrödinger equation one can use an integral Green's function representation to express the wavefunction  $\psi$  at some future time  $t_2$  given initial condition  $\psi(\mathbf{r}, t_1)$  as

$$\psi(\mathbf{r}_2, t_2) = \int \mathcal{K}(\mathbf{r}_2, t_2; \mathbf{r}_1, t_1) \psi(\mathbf{r}_1, t_1) d\mathbf{r}_1. \quad (2.4)$$

The solution to equation

$$\left( i \frac{\partial}{\partial t_2} - H_{\mathbf{r}_2} \right) \mathcal{K}(\mathbf{r}_2, t_2; \mathbf{r}_1, t_1) = i\delta(\mathbf{r}_1 - \mathbf{r}_2)\delta(t_1 - t_2) \quad (2.5)$$

and the *propagator*  $\mathcal{K}(\mathbf{r}_2, t_2; \mathbf{r}_1, t_1)$  is expressed using the Feynman path integral

$$\mathcal{K}(\mathbf{r}_2, t_2; \mathbf{r}_1, t_1) = \int_{\substack{\mathbf{r}(t_1)=\mathbf{r}_1 \\ \mathbf{r}(t_2)=\mathbf{r}_2}} \mathcal{D}\mathbf{r}(t) \exp \left( i \int_{t_1}^{t_2} \mathcal{L}(\mathbf{r}, \dot{\mathbf{r}}) dt \right), \quad (2.6)$$

where  $\mathcal{L}$  is the classical Lagrangian function of the system

$$\mathcal{L}(\mathbf{r}, \dot{\mathbf{r}}) = \frac{1}{2} \dot{\mathbf{r}}^2 - \hat{V}(\mathbf{r}), \quad (2.7)$$

and the integral is over all paths that satisfy the endpoint conditions.

### 2.1.3 Examples of lattice models

$$\hat{\sigma}_i^x = \begin{pmatrix} 0 & 1 \\ 1 & 0 \end{pmatrix}_i \quad \hat{\sigma}_i^y = \begin{pmatrix} 0 & -i \\ i & 0 \end{pmatrix}_i \quad \hat{\sigma}_i^z = \begin{pmatrix} 1 & 0 \\ 0 & -1 \end{pmatrix}_i \quad (2.8)$$

#### Transverse-field Field Ising model

$$\hat{H}_{\text{Ising}} = -J \sum_{\langle i,j \rangle} \hat{\sigma}_i^z \hat{\sigma}_j^z - h \sum_i \sigma_i^x \quad (2.9)$$

#### Heisenberg model

$$\hat{H}_{\text{Heisenberg}} = -\frac{1}{2} \sum_{j=1}^N [J_x \hat{\sigma}_j^x \hat{\sigma}_{j+1}^x + J_y \hat{\sigma}_j^y \hat{\sigma}_{j+1}^y + J_z \hat{\sigma}_j^z \hat{\sigma}_{j+1}^z + h \hat{\sigma}_j^z] \quad (2.10)$$

#### Bose-Hubbard model

$$\hat{H}_{\text{BH}} = -t \sum_{\langle i,j \rangle} \hat{b}_i^\dagger \hat{b}_j + \frac{U}{2} \sum_i \hat{n}_i (\hat{n}_i - 1) - \mu \sum_i \hat{n}_i \quad (2.11)$$

## 2.2 Approaches to the quantum many-body problem

The quantum many-body problem, which amounts to solving the  $3N$ -dimensional Schrödinger equation, underpins a large part of quantum chemistry, condensed matter physics and materials science. The problem is notoriously hard to solve and very few systems with analytical solutions exist, most of them constrained in some artificial way such that they lend themselves to mathematical analysis. Examples include the Hookium atom, an analogue of helium where two electrons interact with the nucleus via a Hookean potential, Spherium, a system of two electrons confined to the surface of a sphere, and the Luttinger liquid of fermions in a one-dimensional conductor. Great efforts have been made in the nearly 100 years since the conception of the Schrödinger equation, in developing both analytical and numerical techniques to produce insights into quantum systems. Perhaps the most impactful was the development of various approximate methods that solve the many-body problem with our limited computational resources. While there is ongoing work on quantum simulators and computers that could greatly speed-up solving quantum problems [14, 10], we here discuss methods one can use with a classical computer. The commonality of all mentioned methods is that they try to tame the exponential growth of the underlying Hilbert space w.r.t the system size, they differ in how they achieve this. The three most common assumptions/simplifications to the many-body problem employed in condensed matter and quantum chemistry are the Born-Oppenheimer approximation, that electronic motion is instantaneous compared to the nuclear motion, the use of chemical basis sets, which transforms the PDE into an algebraic problem, and usually neglecting relativistic effects.

### Hartree-Fock

One of the most common approaches to the many-body problem is to replace the original interacting many-body problem with a set of independent-particle problems with effective potential. **Hartree-Fock** (HF) approaches solve an auxiliary system of independent electrons in a self-consistent field and assume that the wave function (for fermions) can be represented as a single Slater determinant. The HF method does not include electron correlation, which makes it a good approximation only in systems where correlation contributions are small.

### **Post-Hartree-Fock methods**

Post-HF methods, such as Coupled Cluster, Configuration interaction and Møller-Plesset theory include correlation by considering a linear combination of Slater determinants. They can be extremely accurate but come at a high computational cost.

### **Density Functional Theory**

Alternatively **Density Functional Theory** (DFT) reformulates the many-body electron problem in terms of the 3-dimensional electron density  $n(\mathbf{r})$ , which is found by minimising the total energy functional  $E[n(\mathbf{r})]$  [20]. In practice this is done by solving the Kohn-Sham auxiliary system. DFT is in theory exact, however only if the true energy functional  $E[n(\mathbf{r})]$  is known. As this is not the case, much research has been done in constructing different energy functionals with varying degrees of accuracy, starting with local functionals e.g. LSDA and continuing towards more heavily parameterised, non-local formulations. DFT provides a good trade-off between accuracy and computation time, it is used extensively for simulating large systems as linear scaling variants of DFT exist [41].

### **Dynamical Mean Field Theory**

DMFT [19] is a framework that is specialised in solving strongly correlated systems. It is intuitively similar to Weiss Mean Field Theory in classical statistical physics. The main idea is to map an intractable lattice problem into an impurity model in an effective medium, a many-body local problem which can be solved with any standard approach (QMC, DFT, exact diagonalisation, etc.). This mapping between lattice and impurity model is exact, the approximation comes in neglecting spatial fluctuations of the lattice self-energy  $\Sigma$ , the contribution to energy due to particle interaction with medium. DMFT assumes that  $\Sigma$  is a function of frequency and not momentum  $\Sigma(k, \omega) = \Sigma(\omega)$ , which only holds in the infinite coordination case. Time fluctuations are taken into account, i.e. the effective medium is not static in DMFT, which is an advantage over other static mean field theories.

### **Density Matrix Renormalization group**

DMRG [46] is considered the state of the art method for solving one-dimensional lattice problems, it has been widely adopted in condensed matter physics, first used to solve the system of a spin-0 particle in a box. It is an iterative method based on the renormalization group [47], and uses matrix product states as the variational ansatz. The method has also been extended for time evolution of systems [13], and higher dimensions [44].

### 2.2.1 Stochastic methods - Quantum Monte Carlo

Quantum Monte Carlo is a class of methods that uses statistical sampling to directly deal with high-dimensional integration that arises from working with the many-body wave function. QMC methods are among the most accurate achieving chemical accuracy for smaller systems [15], and can in principle achieve any degree of statistical precision sought. A large ecosystem of QMC methods exists, and they have been adapted to study almost any quantum system imaginable, from discrete to continuous state space, fermionic and bosonic systems, as well as both finite and zero temperatures. Even though QMC methods are not computationally the cheapest, they have reasonable storage requirements as the wave function does not need to be stored directly. Moreover, the high computational cost of QMC methods can be aided by parallelisation and use of hardware acceleration, as the core calculation is repetitive and usually involves generating (pseudo)-random numbers, performing a simple calculation and in the end averaging over the results.

#### Variational quantum Monte Carlo (VMC)

The most straightforward QMC approach is based on the variational principle, which provides a clear path towards a solution to the ground state problem. Simply use a *trial wave function*  $\Psi_T$  to parameterise the ground state and optimise the parameters of  $\Psi_T$  to reach the lowest-energy state. This lowest variational state should capture the behaviour of the ground state if the ansatz is expressive enough. The flexibility of easily defining the trial wave function for a variety of different problems and the ease of its evaluation is a clear advantage over methods described in the previous section. Moreover, given that the variational wave function should encapsulate the main aspects of the system studied it provides intuition into the system itself. Development of trial functions has played a key role in the applicability of VMC, famous examples of trial wave functions include the Slater-Jastrow and Backflow wave functions. The drawback of VMC is that the variational wave function might contain a bias that cannot be avoided through optimisation of the parameters alone, see Fig. 2.1. VMC necessarily contains two steps, first is the estimation of the variational energy and second is the optimisation of the parameters. Any expectation of an operator  $\hat{O}$  can be expressed in terms of the trial wave function as

$$\langle \hat{O} \rangle = \frac{\langle \Psi_T | \hat{O} | \Psi_T \rangle}{\langle \Psi_T | \Psi_T \rangle} = \frac{\sum_x \langle \Psi_T | x \rangle \langle x | \hat{O} | \Psi_T \rangle}{\sum_x \langle \Psi_T | x \rangle \langle x | \Psi_T \rangle}, \quad (2.12)$$

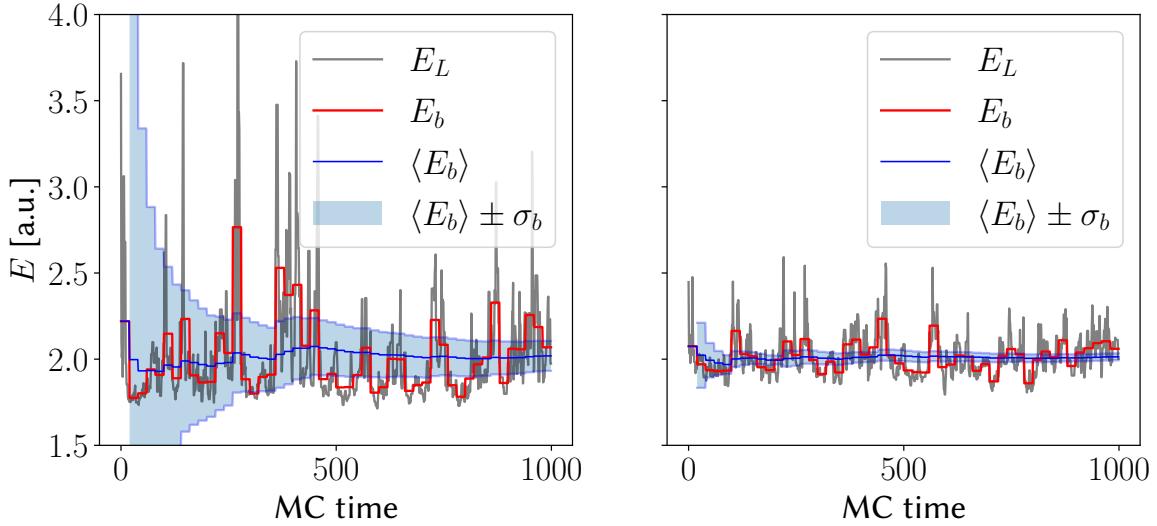


Fig. 2.1 **Ansatz quality in VMC.** Appropriateness of the variational wave function limits the quality of VMC, a poor choice of ansatz results in typical spikes in local energy and biased result (**left**), as well as slower convergence than a good trial wave function (**right**). Figures show the local energy  $E_L$ , reblocked average energy  $\langle E_b \rangle$  and variance  $\sigma_b$  of a VMC simulation of Hookium.

where  $|x\rangle$  are orthogonal and normal states of the Hilbert space. If we rewrite the above expression as

$$\langle \hat{O} \rangle = \frac{\sum_x |\Psi_T(x)|^2 \hat{O}_L(x)}{\sum_x |\Psi_T(x)|^2}, \quad (2.13)$$

with  $\hat{O}_L$  being the *local operator*

$$\hat{O}_L(x) = \frac{\langle x | \hat{O} | \Psi_T \rangle}{\langle x | \Psi_T \rangle}, \quad (2.14)$$

we can interpret  $|\Psi(x)|^2 / \sum_x |\Psi(x)|^2$  as a probability. Meaning that eq. (2.13) can be estimated as an average of the local operator  $\hat{O}_L$

$$\langle \hat{O} \rangle \approx \frac{1}{M} \sum_{m=1}^M \hat{O}_L(x_m), \quad (2.15)$$

sampled from this probability distribution. The sampling can be done using Markov Chain Monte Carlo (MCMC). The second step of the procedure is variational optimisation of the trial wave function, where the optimal parameters of the approximation are found by minimising the *cost function*. The straightforward choice of the variational energy  $E_V$  as a cost function turns out to be inferior to minimizing the *variance* of the energy  $\sigma_E$  [15]. This

is because  $\sigma_E$  obeys the *zero-variance* property, meaning that if  $\Psi_T$  is an exact eigenvalue of the Hamiltonian

$$\hat{H}|\Psi_T\rangle = E_V|\Psi_T\rangle, \quad (2.16)$$

then the local energy  $E_L$  is constant and equal to  $E_V$

$$E_L(x) = \Psi_T(x)^{-1}\hat{H}\Psi_T(x) = \Psi_T(x)^{-1}E_V\Psi_T(x) = E_V, \quad (2.17)$$

irrespective of the sampled configuration  $x$  and hence has zero variance. The zero-variance property has important consequences for numerical stability of optimisation, it means that energy variance minima are robust to finite sampling. Minimizing the variance of energy drives the trial wave function towards eigenstates of the Hamiltonian. Moreover, the statistical error of any expectation value  $\langle \hat{O} \rangle$  is proportional to the variance of  $\hat{O}$ , making low variance doubly desirable. There are several approaches to updating the parameters each iteration, gradient descent, stochastic reconfiguration [42], and the linear method [32] are just a few examples. It is crucial that the methods are robust to statistical noise and converge quickly as the MC step can be expensive to perform. Moreover they are only as good as the estimates of the energy (variance) gradients w.r.t the parameters.

The first application of VMC was to the ground state  ${}^4\text{He}$  [29] and it was later extended for studying many-body fermionic systems [8]. Time-dependant variants of this method exist [5] and VMC has been used to study non-equilibrium properties of bosonic [7, 6], and fermionic [22] systems.

### Projector QMC (PMC) techniques

PMC is a class of QMC methods which are in essence nothing more than stochastic implementations of the power method to obtain the dominant eigenvector of a matrix or a kernel function [16]. Their distinct advantage over VMC is that they are not constrained by our parametrisation of the trial wave function, as they can describe arbitrary probability distributions. PMC methods are based on the imaginary Schrödinger equation

$$\partial_t |\Psi_t\rangle = -\hat{H}|\Psi_t\rangle. \quad (2.18)$$

Its formal solution, the time propagation of an initial wave function  $|\Psi_0\rangle$  at  $t = 0$ , is written as

$$|\Psi_t\rangle = e^{-\hat{H}t}|\Psi_0\rangle. \quad (2.19)$$

From the spectral decomposition of the operator  $e^{-\hat{H}t}$  in terms of eigenstates  $|\Phi_n\rangle$  and eigen-energies  $E_n$  of the Hamiltonian  $\hat{H}$

$$e^{-\hat{H}t} = \sum_n e^{-E_n t} |\Phi_n\rangle\langle\Phi_n|, \quad (2.20)$$

it follows that the term corresponding to the ground state of the system  $|\Phi_0\rangle$  decays the slowest. Thus starting in some initial state and propagating for a long imaginary time  $it$  leads into the ground state with the decay rate giving the ground state energy  $E_0$  as

$$\lim_{t \rightarrow \infty} |\Psi_t\rangle \propto e^{-E_0 t} |\Phi_0\rangle, \quad (2.21)$$

where  $|\Phi_0\rangle$  is the corresponding state of  $E_0$ . This of course holds if the eigenstates of  $\hat{H}$  are all positive, which can be achieved by shifting the potential by a constant energy  $E_c$ , which doesn't change the ground state wave function. The basic step of a PMC simulation is the projection step, where an existing ensemble of configurations is projected into a new one, this projection  $\hat{P}$  is done in such a way that eq. (2.21) is satisfied

$$|\Phi_0\rangle = \lim_{n \rightarrow \infty} \hat{P}^n |\Psi_0\rangle. \quad (2.22)$$

Flavours of PMC differ in the choice of  $\hat{P}$ , the most popular Diffusion Monte Carlo (DMC) [15, 35] works with the time-dependent Green's function  $G(x', t'; x, t)$  of eq. (2.18)

$$\Psi(x, t) = \int G(x, t; x', t') \Psi(x', t') dx', \quad (2.23)$$

while Green's function MC (GFMC) [24, 25] uses the time integrated version of the Green's function

$$\Psi^{(n+1)}(x) = E \int G(x, x') \Psi^{(n)}(x') dx'. \quad (2.24)$$

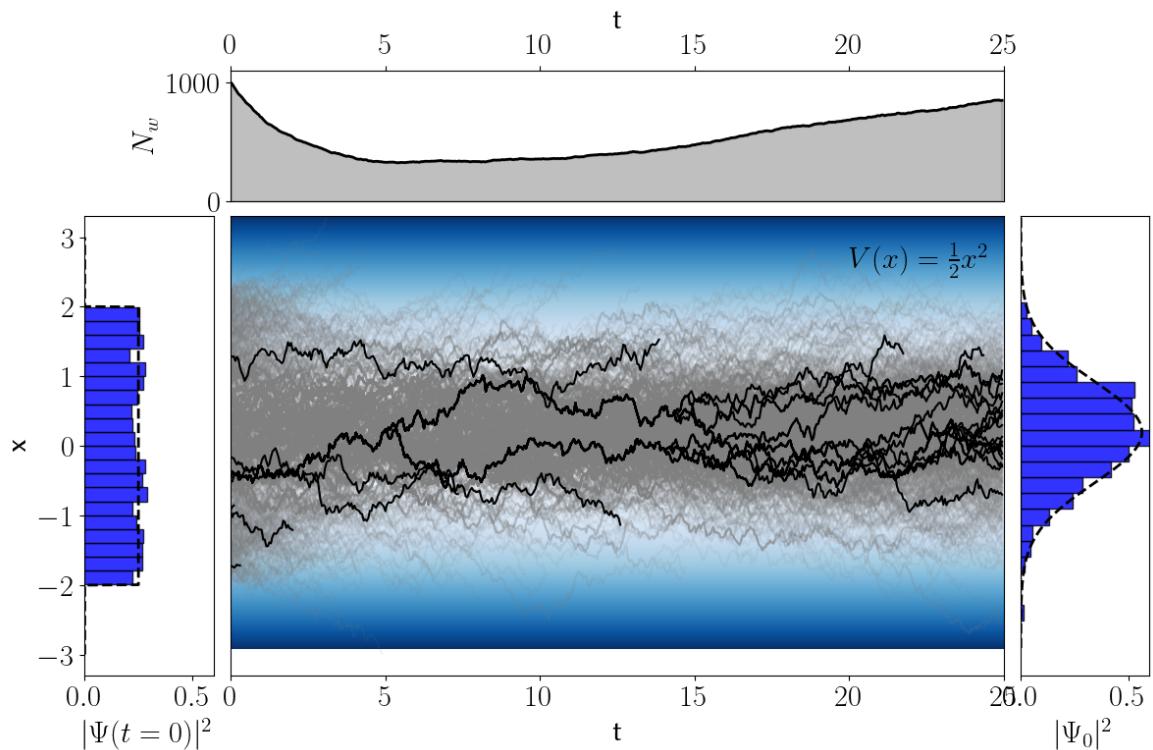
Both formulations are exact, but need some additional approximations to be made practical for use, as Green's functions are not known for a general system. In DMC the Green's function

$$G(x', t'; x, t) = \langle x | e^{-(t-t')[\hat{T} + \hat{V} - E_c]} | x' \rangle, \quad (2.25)$$

is approximated for short times  $\tau = t - t'$  using Trotter-Suzuki formula

$$G(x' \rightarrow x; \tau) = \underbrace{(2\pi\tau)^{-3N/2} e^{-\frac{(x-x')^2}{2\tau}}}_{\text{ordinary diffusion}} \cdot \underbrace{e^{-\tau[V(\mathbf{R}) + V(\mathbf{R}') - 2E_c]/2}}_{\text{reweighting} = \text{birth/death}} + \mathcal{O}(\tau^3), \quad (2.26)$$

where the kinetic term is recognised to be ordinary diffusion. In practice eq. (2.26) is implemented as a simulation of a diffusion process, but instead of weighting the paths of the walkers, the potential contribution to  $G$  is interpreted as a probability of a walker to either branch or die, which is numerically more stable. This stochastic process converges to the ground state for sufficiently long times, see Fig. 2.2. Reptation quantum Monte Carlo [35] (RMC) is an alternative formulation which only uses a single walker, and instead of branching and dying the MC moves mutate the path of that single walker. Using a



**Fig. 2.2 Diffusion Monte Carlo simulation of harmonic oscillator**, starting with  $N_w = 1000$  walkers,  $\tau = 0.05$ ,  $E_c = 0.25$  and uniformly sampling their initial positions from  $(-2, 2)$  (**left**). The number of walkers at each step decreases rapidly before slowly increasing (**top**) the number of walkers is controlled by adjusting  $E_c$ . Walker paths, with a few highlighted in black to emphasise death/birth process (**middle**), diffuse into the approximate ground state of the HO  $u_0(x) = \frac{1}{\pi^{\frac{1}{4}}} e^{-\frac{1}{2}x^2}$  (**right**).

trial wave function  $\Psi_T$  as a guiding function for importance sampling is an important improvement over vanilla DMC. This introduces a *drift* into the diffusion process, which leads the walkers into regions of large values of  $\Psi_T$  and greatly improves the statistical efficiency of the method. The guiding wave function is usually obtained by means of VMC. So far we have conveniently assumed that the wave function is positive everywhere in the

domain, this is not generally true, e.g. in fermionic systems, and poses a problem for PMC methods.

### The sign problem

Projector Monte Carlo methods can only operate with positive distributions, and as such they fall apart when applied to fermionic or frustrated systems [16]. A straightforward modification to the sampling scheme allows us to sample from a mixed-sign distribution. We sample from the distribution normally when it is positive, but sample from its absolute value and change the sign of the observable, when it is negative. The issue with this approach is that the population of configurations is split between positive and negative regions, the averages over both are comparable in size and cancel out, leading to a large statistical error compared to the observable. We refer to the accompanying exponential decrease [16] in sampling efficiency with system size and temperature, **the sign problem**. Its general solution was shown to be NP-hard [43], and as it is believed that  $P \neq NP$ , this implies that no *general* polynomial-time solutions exist. However, this does not mean that the problem cannot be avoided in special cases, the search for solutions is still an area of active research [21, 4, 2]. In practice the sign problem is remedied by the *fixed-node* [3] or *constrained-path* [48] approximation. Fixed-node imposes a boundary condition into the projection such that the projected state shares the nodal surface with the trial wave function. The projected state is now only exact when the nodal surface is exact.

#### 2.2.2 Machine Learning and the quantum many-body problem

# Chapter 3

## Feynman-Kac: connecting Quantum Mechanics and Stochastic Processes

In this chapter we will provide a bridge between the quantum many-body problem discussed in the previous chapter and stochastic processes. This will entail introducing the Feynman-Kac formula and relating it to the Fokker-Planck equation and optimal control approaches to QM. Moreover, a probabilistic view of the cost function will lead us to proposals for loss functions that can be used to learn optimal transition rates and consequently sample the ground state.

The field of stochastic processes is a vast body of work, approached from different angles by mathematicians, physicists and engineers. A necessary consequence of this is that the literature ranges from extremely thorough and rigorous [36, 37] to more applied and intuitive [39]. For this reason, the mentioned discussion will be preceded by an overview of the mathematical notation, lemmas and results from stochastic processes and measure theory that underpin some core ideas of this thesis. To avoid including a whole textbook of material on measure and stochastic processes some concepts will not be rigorously defined, the text will point to relevant literature where this is the case.

### 3.1 Stochastic processes

#### 3.1.1 Fundamentals

This brief, more formal, discussion of stochastic processes is based mostly upon classic texts [12, 36, 37] and borrows some intuitions from [39]. The most basic quantity that we will need is the **probability space**.

**Definition 3.1.1 (Probability space)** *The probability space is a tuple  $(\Omega, \mathcal{F}, \mathbb{P})$ , where  $\Omega$  is the sample space,  $\mathcal{F}$  is a  $\sigma$ -field, and  $\mathbb{P}$  is the measure.*

The sample space is simply the set of all possible outcomes. A canonical example would be the roll of a 6-sided dice,  $\Omega = \{1, 2, 3, 4, 5, 6\}$ . Without measure  $\mathbb{P}$ , the tuple  $(\Omega, \mathcal{F})$  is termed a **measurable space**.

**Definition 3.1.2 ( $\sigma$ -field)** *A  $\sigma$ -field  $\mathcal{F}$  on a set  $\Omega$ , is a nonempty collection of subsets of  $\Omega$  that includes  $\Omega$  itself, is closed under complement, i.e. if  $A \in \mathcal{F}$  then  $A^c \in \mathcal{F}$ , and is closed under countable unions,  $\cup_i A_i \in \mathcal{F}$  if  $A_i \in \mathcal{F}$  is a countable union of sets.*

The main utility of the  $\sigma$ -field to us is its use in defining measures. We want to be able to assign a non-negative real number to all subsets of  $\Omega$ , as well as the size of the union of the disjoint sets to be the sum of their individual sizes. This is not always possible, a counterexample for the real line being Vitali sets. The collection  $\mathcal{F}$ , must thus only include *measurable* sets, which are precisely the ones that satisfy the constraints imposed by the  $\sigma$ -field.

**Definition 3.1.3 (Measure)** *A non-negative countably additive set function  $\mu : \mathcal{F} \rightarrow \mathbb{R}$  that satisfies*

- i)  $\mu(A) \geq \mu(\emptyset) = 0$  for all  $A \in \mathcal{F}$
- ii) if  $A_i \in \mathcal{F}$  is a countable sequence of disjoint sets, then  $\mu(\cup_i A_i) = \sum_i \mu(A_i)$

is a **measure**.

If  $\mu(\Omega) = 1$ , then  $\mu$  is a **probability measure** and will be denoted by  $\mathbb{P}$ . With this notion we are now able to define a random variable (r.v) and a stochastic process (s.p.).

**Definition 3.1.4 (Random variable)** *A random variable  $X$  defined on  $\Omega$  is a real-valued measurable function  $X(\omega)$ ,  $X : \Omega \rightarrow \mathbb{R}^d$ .*

For a function to be measurable, we require that its preimage  $X^{-1}$  is in the  $\sigma$ -field  $\mathcal{F}$

$$X^{-1}(B) = \{\omega : X(\omega) \in B\} \in \mathcal{F}, \quad (3.1)$$

and that this holds for every Borel set  $B$  in the Borel  $\sigma$ -field<sup>1</sup> of  $\mathbb{R}^d$ , which is simply the smallest  $\sigma$ -field that contains all measurable sets in  $\mathbb{R}^d$ .

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<sup>1</sup>For a proper definition of the Borell set see ch. 3 of [38].

A random variable  $X$  induces a probability measure  $\mu$  on  $\mathbb{R}^d$  called its **distribution**, this is done by setting  $\mu(A) = P(X \in A)$  for Borel sets  $A$ . Moreover, the distribution is usually given in terms of a **distribution function**  $F(x)$

$$F(x) = \mathbb{P}(\{\omega \in \Omega : X(\omega) \leq x\}) = \mathbb{P}(X \leq x), \quad (3.2)$$

and  $X$  is said to have a **density function**  $f(x)$  if  $F(x)$  can be written as

$$F(x) = \int_{-\infty}^x f(y)dy. \quad (3.3)$$

In essence, the random variable provides a connection between the less familiar probability measure  $\mathbb{P}$  and the cumulative distribution function (CDF).

### 3.1.2 Stochastic process

**Definition 3.1.5 (Stochastic process)** *Given a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  and a measurable (state) space  $(E, \mathcal{E})$ , we define the collection  $\{X_t : t \in T\}$  of set  $T$  indexed and  $(E, \mathcal{E})$  valued random variables a **stochastic process**.*

By far the most common case for the index set  $T$ , is time  $T = \mathbb{R}^+$ . Such s.p's are called *temporal*, examples include the model of velocity of a Brownian particle under influence of friction, in Fig. 3.1, or the Black-Scholes model. Nevertheless, the index set is not limited to time, as is often the case with Gaussian Process regression [34]. In this thesis we will mostly deal with temporal s.p's of the kind that do not "see into the future". This notion is formalized using **filtrations**. A filtration  $\mathbb{F} = (\mathcal{F}_t)_{t \in T}$  on a probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  is just an increasing sequence or order of  $\sigma$ -fields

$$\mathcal{F}_s \subset \mathcal{F}_t \text{ if } 0 \leq s \leq t < \infty \quad (3.4)$$

The filtration associated to a process that records its "past behaviour" at each time is called the **natural filtration**.

**Definition 3.1.6 (Adapted process)** *A process  $\{X_t\}$  is said to be **adapted to the filtration**  $(\mathcal{F}_t)_{t \in T}$  if the random variable  $X_t : \Omega \rightarrow E$  is  $\mathcal{F}_t$ -measurable function for each  $t \in T$ .*

A process that is *non-anticipating*, i.e. depends only on the past and present, is adapted to the filtration  $(\mathcal{F}_t)_{t \in T}$ .

**Definition 3.1.7 (Brownian motion)** *Brownian motion or a non-anticipating Wiener process is a stochastic process  $W_t$ , with the following properties:*

- i)  $W_0 = 0$
- ii)  $W_t$  is almost surely continuous in  $t$
- iii)  $W_t$  has independent increments
- iv)  $W_t - W_s \sim \mathcal{N}(0, t - s)$  for  $0 \leq s \leq t$

A realisation of Brownian motion can be found in Fig. 3.1.

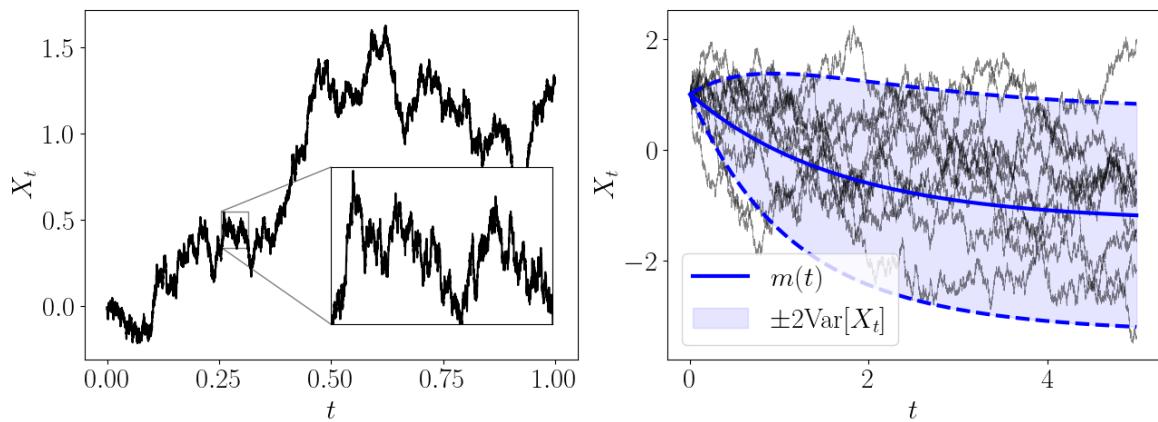


Fig. 3.1 **Brownian motion and Ornstein–Uhlenbeck process.** **left:** A single realisation of the Brownian process. **right:** Mean, variance and 10 samples of the Ornstein–Uhlenbeck process with  $\theta = 0.6$ ,  $\sigma = 1.1$ ,  $X_0 = 1.0$ ,  $\mu = -1.3$ , integrated using Euler–Maruyama method.

### 3.1.3 Integrals

In order to proceed and define stochastic differential equations (SDE's) and the Radon–Nikodyn derivative, we must spend some time discussing various integrals we will use. In particular, alongside the usual Riemann integral, we will need three more types of integrals, which we will briefly describe without mathematical derivation. The simplest kind of integral we will introduce is the integral of a stochastic process

$$I = \int_0^t X_t dt. \quad (3.5)$$

The simple appearance of the integral is deceiving as the integrand is a realisation of a  $\mathcal{F}_t$ -adapted stochastic process  $\{X_t\} : \Omega \times T \rightarrow \mathbb{R}^d$ , meaning that  $I$  itself is a random variable. However, since each realisation of  $X_t$  is almost surely continuous,  $I$  can be expanded as a Riemann sum, which converges under mean-squared norm to  $I$ , so long as the mean

$\mathbf{m}(t) = \mathbb{E}[X_t]$  and covariance  $\mathbf{k}(t, s) = \text{Cov}(X_t, X_s)$  are continuous. In practice, computing the mean and covariance of  $I$  is usually enough to understand the resulting stochastic process. Importantly, integrals of continuous functions of s.p.'s  $h(X_t)$ ,  $h : \mathbb{R} \rightarrow \mathbb{R}$  can be computed in a similar manner.

The second type of integrals we need to consider, are integrals with respect to a s.p., known as **Itô integrals**

$$Y_t = \int_0^t H_s dX_s, \quad (3.6)$$

where both  $H_s$  and  $X_s$  are stochastic processes. The result integral  $Y_t$  is itself a stochastic process which resides in the probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ , filtered by  $(\mathcal{F}_t)_{t \in T}$ . The integral can be formalised by putting slight constraints on what sort of stochastic processes  $X_s$  and  $H_t$  can be, expanding  $Y_t$  as a Riemann sum and proving convergence. Details of this procedure can be found in [37].

Finally we must define the **Lebesgue-Stieltjes integral** [17], which we need to properly define expectations of stochastic processes.

**Definition 3.1.8 (Lebesgue-Stieltjes Integral)** *Given probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  and measurable function  $f : \Omega \rightarrow \mathbb{R}$ , the Lebesgue-Stieltjes integral*

$$I = \int_A f(x) d\mathbb{P}(x), \quad (3.7)$$

*is the Lebesgue integral<sup>2</sup> with respect to measure  $\mathbb{P}$ ,  $A \in \mathcal{F}$ .*

With it we can define expectations in the probability space  $(\Omega, \mathcal{F}, \mathbb{P})$  as

$$\mathbb{E}_{\mathbb{P}}[f(x)] = \int_{\Omega} f(x) d\mathbb{P}(x). \quad (3.8)$$

For a newcomer to stochastic processes this formulation may seem redundant, can we not just calculate expectations using a Riemann integral and the PDF? We can, and when the distribution  $\mathbb{P}$  can be expressed in terms of the PDF (3.3), the Lebesgue integral can be interpreted in this way. However, stochastic processes need not admit a PDF, that is when the Lebesgue-Stieltjes integral is necessary.

### 3.1.4 Stochastic Differential Equations

In this thesis we will refer to a SDE as an informal notation of an Itô integral equation or **Itô process**.

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<sup>2</sup>For proper definition of the Lebesgue integral see ch. 1 of [38].

**Definition 3.1.9 (Itô process)** Given deterministic functions  $v : \mathbb{R}^d \times \mathbb{R}^+ \rightarrow \mathbb{R}^d$  and  $\sigma : \mathbb{R}^d \times \mathbb{R}^+ \rightarrow \mathbb{R}^{d \times d}$ , we define the **Itô process**  $X_t$  as the sum of Itô and Lebesgue integrals

$$X_{t+s} - X_t = \int_t^{t+s} \sigma(X_u, u) dW_u + \int_t^{t+s} v(X_u, u) du, \quad (3.9)$$

where  $W_t$  is a Brownian motion.

In simplified notation we can write (3.9) as

$$dX_t = \sigma(X_t, t) dW_t + v(X_t, t) dt, \quad (3.10)$$

this is what we refer to as an SDE, an example can be found in Fig. 3.1. The functions  $v$  and  $\sigma$ , we will refer to as the **drift** and **volatility** of the Itô process respectively. The most intuitive interpretation of a SDE is in terms of the time evolution of the PDF of the process  $X_t$ . It is described by the **Fokker-Planck equation**<sup>3</sup>

$$\frac{\partial p(\mathbf{x}, t)}{\partial t} = - \sum_{i=1}^N \frac{\partial}{\partial x_i} [\mu_i(\mathbf{x}, t) p(\mathbf{x}, t)] + \sum_{i=1}^N \sum_{j=1}^N \frac{\partial^2}{\partial x_i \partial x_j} [D_{ij}(\mathbf{x}, t) p(\mathbf{x}, t)], \quad (3.11)$$

where  $p(\mathbf{x}, t)$  is the PDF of the solution to the SDE and  $D = \frac{1}{2} \sigma \sigma^\top$  is the diffusion tensor. Finally we state without proof a consequence of Itô calculus, most commonly named **Itô's rule or lemma**, it is the stochastic calculus equivalent of the chain rule

**Lemma 3.1.1 (Itô's lemma)** Given an Itô process  $X_t$  as given by (3.9) and a twice differentiable scalar function  $f(X_t, t)$ , then the Itô process for  $f$  is

$$df = \frac{\partial f}{\partial t} dt + \sum_i \frac{\partial f}{\partial x_i} dx_i + \frac{1}{2} \sum_{ij} \frac{\partial^2 f}{\partial x_i \partial x_j} dx_i dx_j, \quad (3.12)$$

when compared to ordinary calculus we notice an additional quadratic term.

### 3.1.5 Radon-Nikodym Derivative and Girsanov theorem

To perform importance sampling we perform a change of measure in an integral

$$\int_A f(x) d\mathbb{P}(x) = \int_A f(x) \frac{d\mathbb{P}}{d\mathbb{Q}}(x) d\mathbb{Q}(x). \quad (3.13)$$

---

<sup>3</sup>Derivation in [39].

The function that measures the rate of change of density of one measure w.r.t another is the **Radon-Nikodym derivative**  $\frac{d\mathbb{P}}{d\mathbb{Q}}(x)$ .

**Theorem 3.1.2 (Radon-Nikodym theorem)** *Let  $\mathbb{P}$  and  $\mathbb{Q}$  be probability measures on the measurable space  $(\Omega, \mathcal{F})$ , then the measurable function **Radon-Nikodym derivative**  $\frac{d\mathbb{P}}{d\mathbb{Q}}(x) : \Omega \rightarrow [0, \infty)$  exists and*

$$\mathbb{P}(A) = \int_A \frac{d\mathbb{P}}{d\mathbb{Q}}(x)d\mathbb{Q}(x), \quad (3.14)$$

for set  $A \subseteq \mathcal{F}$ .

The RN derivative will also be useful in defining the KL divergence between two **path measures**. Properly defining the path measure would bring a lot of notational overhead, it is enough to think of it as a measure on the **path space**, i.e all possible paths of a SDE, for rigour see [28]. Finally, we state the **Girsanov theorem** that is often used for transforming or removing drift functions of SDE, it is the RN derivative between an Itô process and one with  $v = 0$  and  $\sigma = 1$ , i.e. Brownian motion.

**Theorem 3.1.3 (Girsanov Theorem)** *Given Itô process*

$$dX_t = dW_t + v(X_t, t)dt \quad \text{and} \quad X_0 = 0 \quad (3.15)$$

and Brownian motion  $dY_t = dW_t$ , the RN derivative of their respective path measures  $\mathbb{P}$  and  $\mathbb{P}_0$  is

$$\frac{d\mathbb{P}}{d\mathbb{P}_0} = \exp\left(-\frac{1}{2} \int_0^t |v(X_s, s)|^2 ds + \int_0^t v(X_s, s)^\top dW_s\right) \quad (3.16)$$

This *change in dynamics* as we will call it later is true in the sense, that expectations for an arbitrary functional  $h(\cdot)$  of the path from 0 to  $t$  are

$$\mathbb{E}_{\mathbb{P}_0}[h(X_t)] = \mathbb{E}_{\mathbb{P}_0}\left[\frac{d\mathbb{P}}{d\mathbb{P}_0} h(Y_t)\right]. \quad (3.17)$$

For a more general case and proof see [39].

### 3.1.6 Markov processes

We now shift our view to a special kind of s.p's, ones that satisfy the **Markov property** called **Markov processes** or **Markovian**. The property is sometimes referred to as *memorlessness*, as the future of a Markov process depends only on the present state. We can classify the processes based on the system's **state-space**  $S$ , which can be either discrete (countable) or continuous, and the **time indexing** of the system, either discrete-time

$\{X_n\}_{n \geq 0}$  or continuous-time  $\{X_t\}_{t \geq 0}$ . A taxonomy is given in Table 3.1. We will not specifically discuss Markov processes in continuous state-space, but it is important to note that any Itô process with time-homogenous drift  $v = v(X_t)$  and volatility  $\sigma = \sigma(X_t)$  is Markovian.

From now on we refer to Markov processes in countable state-space as **Markov chains**. We base our discussion on [36] and [33].

Table 3.1 **Taxonomy of Markov processes**

	Countable state-space	Continuous state-space
Discrete time	<b>index:</b> $\{X_n\}_{n \geq 0}, n \in \mathbb{Z}^+$ <b>state-space:</b> countable set $I$ <b>define:</b> stochastic $\{P\}_{ij}$ <b>example:</b> DTMC	<b>index:</b> $\{X_n\}_{n \geq 0}, n \in \mathbb{Z}^+$ <b>state-space:</b> general state-space $\Omega$ <b>define:</b> stochastic kernel $K$ <b>example:</b> Harris Chain
Continuous time	<b>index:</b> $\{X_t\}_{t \geq 0}, t \in \mathbb{R}^+ = [0, \infty)$ <b>state-space:</b> countable set $I$ <b>define:</b> rate $\{\Gamma\}_{ij}$ equiv. to jump chain $\{J_n\}_{n \geq 0}$ and hold times $\{S_n\}_{n \geq 1}$ . <b>example:</b> CTMC	<b>index:</b> $\{X_t\}_{t \geq 0}, t \in \mathbb{R}^+ = [0, \infty)$ <b>state-space:</b> general state-space $\Omega$ <b>define:</b> stochastic kernel $K$ <b>example:</b> Diffusion process

### Discrete-time Markov Chains

The simplest and most common Markov process is a Markov chain in discrete time, an example of it can be found in Fig. 3.2. Its state-space is a countable set  $I$  and we call each  $i \in I$  a **state**. We define a distribution  $\lambda$  in a familiar way

$$\lambda = \{\lambda_i : i \in I\} \quad \text{where} \quad \forall i : 0 \leq \lambda_i < \infty \quad \text{and} \quad \sum_{i \in I} \lambda_i = 1. \quad (3.18)$$

We can now set  $\lambda$  as a distribution of some random variable  $X : \Omega \rightarrow I$  as

$$\lambda_i = \mathbb{P}(X = i) = \mathbb{P}(\{\omega : X(\omega) = i\}), \quad (3.19)$$

where we are still working in the probability space  $(\Omega, \mathcal{F}, \mathbb{P})$ . A discrete-time Markov chain is defined in terms of its **transition matrix**  $P = \{p_{ij} : i, j \in I\}$ , which is a **stochastic matrix** meaning all of its rows  $\{p_{ij} : j \in I\}$  are distributions.

**Definition 3.1.10 (Discrete-time Markov chain)** A discrete time stochastic process  $\{X_n\}_{n \geq 0}$  is a **discrete-time Markov chain** with initial distribution  $\lambda$  and transition matrix  $P$  if for  $i_1, \dots, i_n + 1 \in I$  and  $n \geq 0$

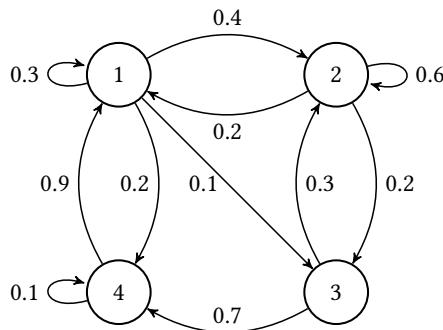
- i)  $\mathbb{P}(X_0 = i_1) = \lambda_{i_1}$
- ii)  $\mathbb{P}(X_{n+1} = i_{n+1} | X_0 = i_0, \dots, X_n = i_n) = p_{i_n i_{n+1}}$

Rewriting the second condition above, it is clear that the Markov chain is without memory

$$\mathbb{P}(X_{n+1} = i_{n+1} | X_0 = i_1, \dots, X_n = i_n) = \mathbb{P}(X_{n+1} = i_{n+1} | X_n = i_n). \quad (3.20)$$

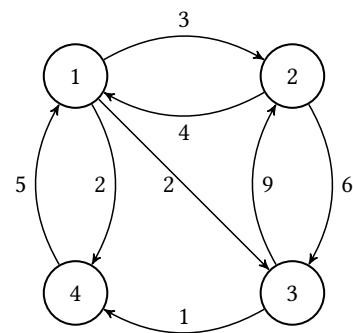
Intuitively we understand the discrete-time Markov chain as a system changing its state at discrete time intervals, each time choosing the next state according to the row of the Transition matrix corresponding to the current state.

a) Discrete-time Markov Chain



$$P = \begin{pmatrix} 0.3 & 0.4 & 0.1 & 0.2 \\ 0.2 & 0.6 & 0.2 & 0 \\ 0 & 0.3 & 0 & 0.7 \\ 0.1 & 0 & 0 & 0.1 \end{pmatrix}$$

b) Continuous-time Markov Chain



$$\Gamma_{ii} = \sum_{j \neq i} -\Gamma_{ij} \rightarrow \Gamma = \begin{pmatrix} -7 & 3 & 2 & 2 \\ 4 & -10 & 6 & 0 \\ 0 & 9 & 10 & 1 \\ 5 & 0 & 0 & -5 \end{pmatrix}$$

Fig. 3.2 Discrete and continuous time Markov Chains. **left:** Discrete-time Markov Chain defined by  $P$ . **right:** Continuous-time Markov Chain defined by  $\Gamma$ .

### Continuous-time Markov Chains

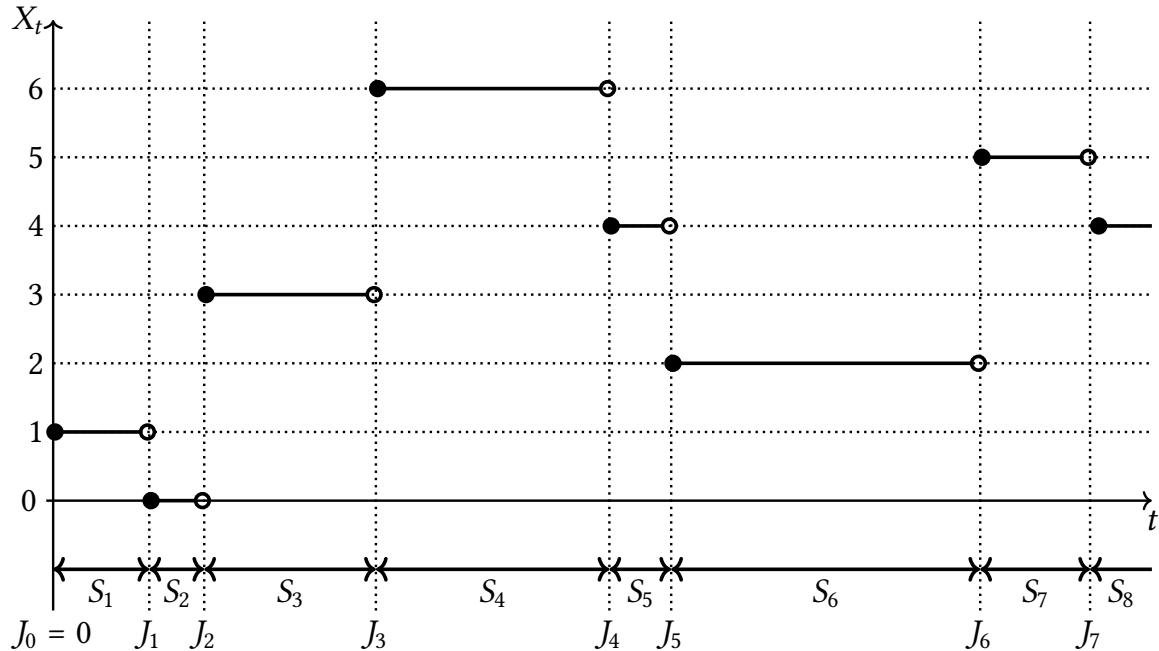
Defining a Markov chain in continuous time is slightly trickier as describing the system with a stochastic matrix does no longer suffice because transition probabilities become zero when considering an infinitesimal time. Instead a continuous-time Markov Chain (CTMC) is characterised by a **rate matrix** or **infinitesimal generator matrix**  $\Gamma$  defined on the set  $I$ . A rate matrix has the following three properties

$$\text{i)} \quad 0 \leq \Gamma_{ii} < \infty, \quad \forall i$$

$$\text{ii)} \quad \Gamma_{ij} \geq 0, \quad \forall i \neq j$$

$$\text{iii)} \quad \sum_{j \in I} \Gamma_{ij} = 0, \quad \forall i$$

While the CTMC can be interpreted in a number of ways, we shall use the so called **jump chain** and **holding times** representation, see also Fig. 3.3. We can think of a CTMC as a



**Fig. 3.3 Jump chain and Holding times.** A discrete space Markov process  $\{X_t\}_{t \geq 0}$  in continuous time. The holding times  $S_n$  are independent exponential random variables and the transition probabilities at jump times  $J_n$  are given with the jump matrix  $\Pi$ . Inspired by [33].

series of discrete jumps, where the system remains in each state for a certain holding time. This suggests that we can construct the CTMC from a discrete-time chain with stochastic matrix  $\Pi$ , which we will call the **jump matrix**, and a set of independent random variables  $\{S_n\}$  which determine the holding times. We construct matrix  $\Pi$  by rescaling rows of  $\Gamma$  so they add up to one, putting a 0 on the diagonal

$$\begin{aligned} \Pi_{ij} &= \begin{cases} \Gamma_{ij}/\Gamma_{ii} & \text{if } j \neq i \text{ and } \Gamma_{ii} \neq 0 \\ 0 & \text{if } j \neq i \text{ and } \Gamma_{ii} = 0 \end{cases} \\ \Pi_{ii} &= \begin{cases} 0 & \text{if } \Gamma_{ii} \neq 0 \\ 1 & \text{if } \Gamma_{ii} = 0. \end{cases} \end{aligned} \tag{3.21}$$

In order for the process to possess the Markov property, the distribution of holding times  $\{S_n\}$  must be exponential [33],

$$S_{n+1} \sim \text{Exp}(-\Gamma_{ii}(X_n)), \quad (3.22)$$

with exponential parameters being  $-\Gamma_{ii}$  where  $i$  is the current state. Processes with different holding time distributions are called **semi-Markov**. The jump times  $\{J_n\}$  are simply

$$J_n = S_1 + \dots + S_n. \quad (3.23)$$

**Definition 3.1.11 (Continuous-time Markov chain)** A stochastic process  $\{X_t\}_{t \geq 0}$  on set  $I$  is a **continuous-time Markov chain** if its jump chain  $\{Y_n\}_{n \geq 0}$  is a discrete-time Markov chain and its holding times  $\{S_n\}_{n \geq 1}$  are independent exponential random variables  $S_n \sim \text{Exp}(-\Gamma_{ii}(X_n))$ .

An equivalent formulation is in terms of **competing exponentials**. Transitions  $\Gamma_{j \rightarrow k}$  from  $j$  to  $k$  are defined as independent exponential random variables  $\tau_{j \rightarrow k}$

$$\tau_{j \rightarrow k} \sim \text{Exp}(\Gamma_{jk}), \quad j \neq k \quad (3.24)$$

the next state is then chosen as

$$Y_{n+1} = \operatorname{argmin}_k \tau_{j \rightarrow k}. \quad (3.25)$$

The chain  $\{Y_n\}_{n \geq 0}$  along with times

$$S_n = \min_k \tau_{j \rightarrow k}, \quad (3.26)$$

gives the full description of the CTMC. With this formulation in mind we now interpret  $\Gamma_{ii}$  as the rate of *leaving* current state and  $\Gamma_{ij}$  as the rate of *going* from  $i$  to  $j$ .

## 3.2 The Feynman-Kac formula

The Feynman path integral formulation introduced in chapter 2 was extensively used by physicists for decades, even in the absence of a formal mathematical formulation which is hard to define because of the difficulties with defining an appropriate measure on the path space. Kac [23] provided a rigorous formulation of the *real-valued* case of the Feynman path

integral, and the resulting **Feynman-Kac** formula provides a bridge between *parabolic* partial differential equations and stochastic processes.

To illustrate the Feynman-Kac formula let us consider a single particle with Hamiltonian

$$\hat{H} = -\frac{d^2}{dx^2} + V(x) \quad (3.27)$$

and the Schrödinger equation in *imaginary time*, which is of the elliptic type,

$$\partial_t |\psi_t\rangle = -\hat{H} |\psi_t\rangle. \quad (3.28)$$

Its formal solution, the time propagation of an initial wave function  $|\phi_0\rangle$  at  $t = 0$ , is written as

$$|\psi_t\rangle = e^{-\hat{H}t} |\phi_0\rangle. \quad (3.29)$$

From the spectral decomposition of the operator  $e^{-\hat{H}t}$  in terms of eigenstates  $|\phi_n\rangle$  and eigen-energies  $E_n$  of the Hamiltonian  $\hat{H}$

$$e^{-\hat{H}t} = \sum_n e^{-E_n t} |\phi_n\rangle \langle \phi_n|, \quad (3.30)$$

it follows that the term corresponding to the ground state of the system  $|\phi_0\rangle$  decays the slowest. Thus starting in some initial state and propagating for a long imaginary time  $it$  leads into the ground state with the decay rate giving the ground state energy as

$$\lim_{t \rightarrow \infty} |\psi_t\rangle \propto e^{-E_0 t} |\phi_0\rangle, \quad (3.31)$$

where  $E_0$  is the ground state energy and  $|\phi_0\rangle$  is the corresponding state. Kac noticed that the kinetic term of the Lagrangian in (2.6) could be interpreted as a measure on Brownian walks, and a solution to the imaginary time Schrödinger equation can be written as

$$\psi(x, t) = \mathbb{E}_{X \sim \text{Brownian with } X_t=x} \left[ \exp \left( - \int_0^t V(X_\tau, \tau) d\tau \right) \psi(X_0, 0) \right], \quad (3.32)$$

where only the **endpoint** at time  $t$  of the Brownian process fixed, whereas the starting point at time  $t = 0$  is not,  $\psi(x, 0)$  encodes the initial condition into this representation. When there is no external potential  $V(x) = 0$ , the Schrödinger equation in imaginary time is the

diffusion equation and the Feynman-Kac solution is simply

$$\begin{aligned}\psi(x, t) &= \mathbb{E}_{X \sim \text{Brownian with } X_t=x} [\psi(X_0, 0)] \\ &= \frac{1}{\sqrt{2\pi t}} \int e^{-(x-x')^2/2t} \psi_0(x') dx'\end{aligned}\quad (3.33)$$

An illustration of the Feynman-Kac approach to the problem with no external potential  $V(x)$  in 1D is depicted in Fig. 3.5. The role of the potential in the Feynman-Kac formula is to weight the Brownian paths, in turn defining the Feynman-Kac *path measure*  $\mathbb{P}_{\text{FK}}$  it is related to the Brownian measure  $\mathbb{P}_0$  by the Radon-Nykodym derivative

$$\frac{d\mathbb{P}_{\text{FK}}}{d\mathbb{P}_0} = \mathcal{N} \exp \left( - \int V(X_t) dt \right), \quad (3.34)$$

where  $\mathcal{N}$  is a normalizing constant. Intuitively we can understand the measure as assigning more weight to Brownian paths that spend more time in the attractive region ( $V(x) < 0$ ) than in repulsive regions ( $V(x) > 0$ ), this is illustrated in Fig. 3.4. Moreover, this new stochastic

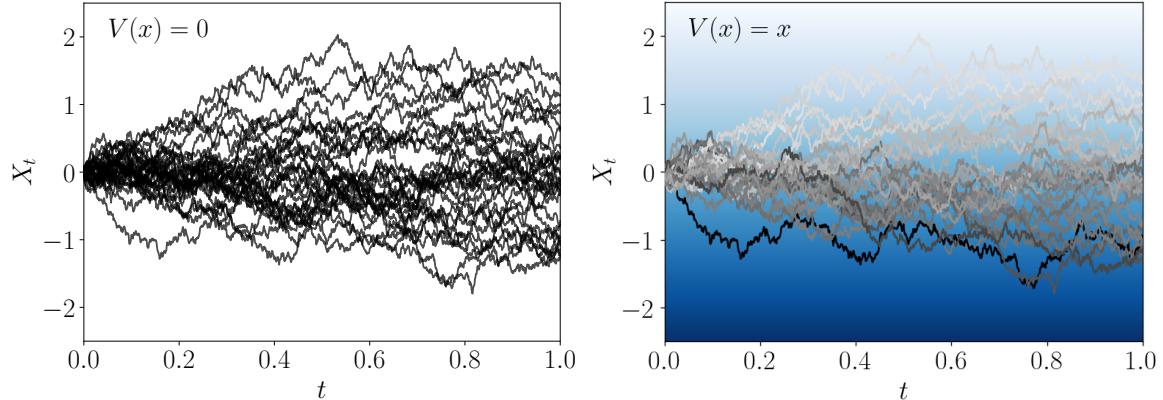


Fig. 3.4 **Feynman-Kac measure in a linear potential.** **left:**  $N = 30$  Brownian paths. **right:** The paths colored ( $P(\text{black}) = 1, P(\text{white}) = 0$ ) by their likelihood under the Feynman-Kac measure with  $V(x) = x$ .

process is Markovian, meaning that a clear connection exists between the imaginary time Schrödinger equation and a SDE of form (3.10) with time-homogeneous  $\sigma$  and  $v$ . Indeed, in the continuous case we have discussed so far, the mapping between the Fokker-Planck equation (3.11) and the Schrödinger equation exists in the form of a similarity transform.

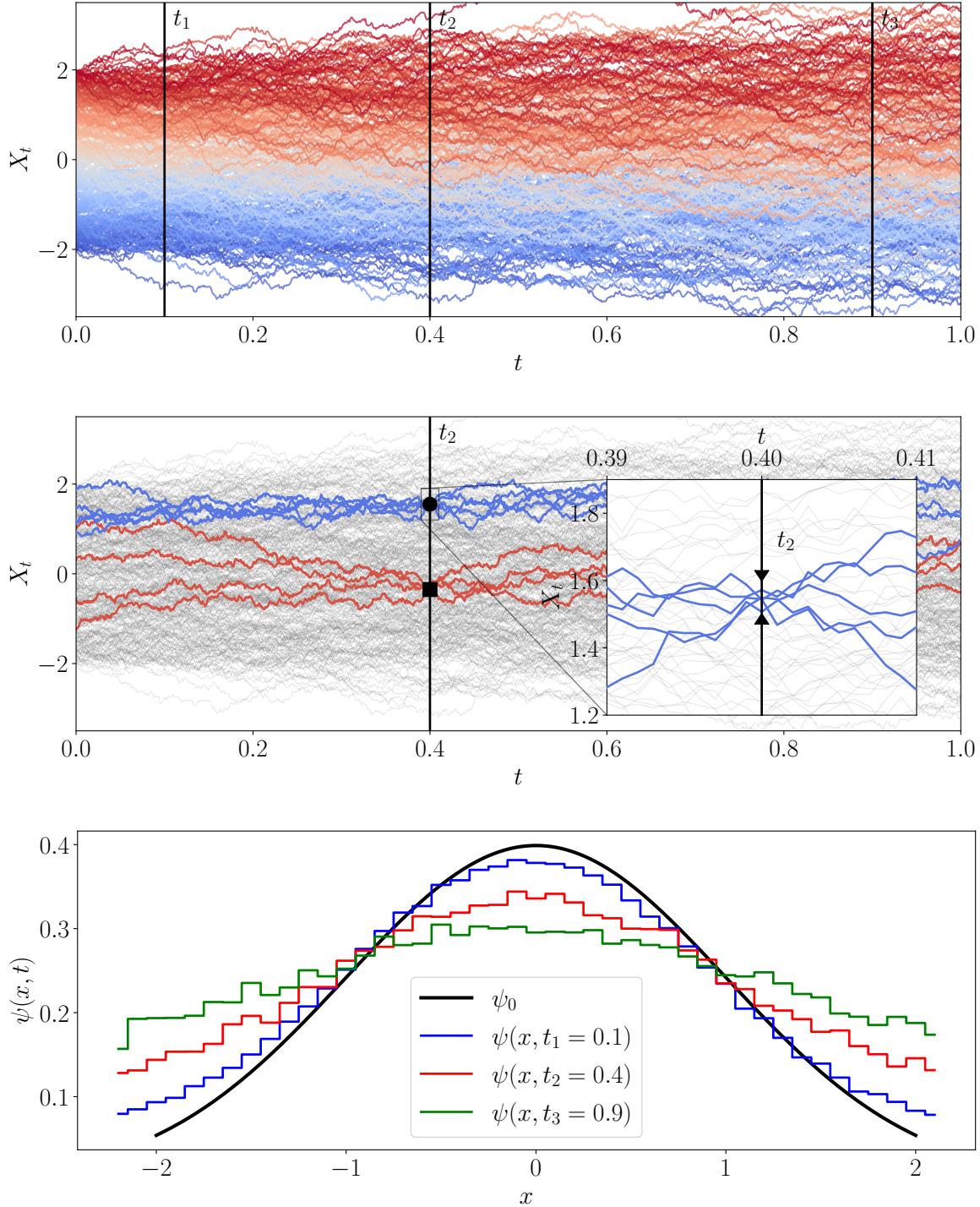


Fig. 3.5 **Feynman-Kac for a free particle in 1D.** **top:**  $N = 400$  Brownian walks starting from different  $x_0$ , the color signifies initial position. In order to evaluate  $\psi$  between  $x - \frac{\delta x}{2}$  and  $x + \frac{\delta x}{2}$  at some time  $t$  we must first find Brownian paths that end there. **middle:** The paths that pass through at  $x \in (1.5, 1.6)$  (blue) and through  $x \in (-0.4, -0.3)$  (red) are colored, others are left in grey. **bottom:** Time evolution of the initial condition  $\psi_0 = \frac{1}{\sqrt{2\pi}} e^{-\frac{1}{2}x^2}$ , by estimating  $\mathbb{E}[\psi(X_0, 0)]$  from the filtered paths at each timestep.

Starting from the FP equation of a stochastic process with constant volatility  $\sigma = 1$

$$dX_t = dW_t + v(X_t) dt, \quad (3.35)$$

and drift  $v(x) = -U'(x)$  given as a gradient of some potential function  $U(x)$ , the PDF  $\rho(t, x)$  of the process is

$$\frac{\partial \rho}{\partial t} = \frac{\partial}{\partial x} \left[ \frac{\partial \rho}{\partial x} + U'(x)\rho \right]. \quad (3.36)$$

We can define the function

$$\psi(x, t) = \frac{\rho(x, t)}{\sqrt{\rho_0(x)}}, \quad (3.37)$$

with  $\rho_0$  being the stationary distribution of the FP equation

$$\frac{\partial}{\partial x} \left[ \frac{\partial \rho}{\partial x} + U'(x)\rho \right] = 0 \quad \rightarrow \quad \rho_0(x) \propto \exp(-U(x)), \quad (3.38)$$

which satisfies the imaginary time Schrödinger equation (3.28) with the Hamiltonian

$$H = -\frac{\partial^2}{\partial x^2} \overbrace{-\frac{U''}{2} + \frac{U'^2}{4}}^{\equiv V(x)}. \quad (3.39)$$

The ground state of this Hamiltonian has zero energy and is

$$\psi_0(x) = \sqrt{\rho_0(x)}. \quad (3.40)$$

In other words, the quantum ground state probability distribution  $|\psi_0|^2$  is the same as classical stationary distribution  $\rho_0$  of the stochastic process  $X_t$  in the literature referred to as the *Nelson's ground state process* [31, 1]. This connection is one that our computational method will exploit, as the ability to efficiently sample from the stochastic process with correct drift  $v$  is equivalent to sampling from the ground state of the quantum system. Even though the connection is simple, it comes with a caveat. Starting from the Schrödinger equation one needs to find the drift  $v(x)$  and while the connection with the potential of the Hamiltonian is clear-cut in this simple example, this is not the case in many-body systems, i.e. the *inverse problem* of finding the stochastic process of a given Hamiltonian is difficult, and is one of the core problems approached in this thesis.

### 3.3 Stoquastic Hamiltonians and Lattice-model representations

A similar connection between stochastic processes and the ground state exists in the discrete space, the difference being that instead of FP we have the Master equation

For a Markov process over discrete states, we can write the master equation as

$$\frac{\partial P_j}{\partial t} = \sum_{k \neq j} [\Gamma_{k \rightarrow j} P_k - \Gamma_{j \rightarrow k} P_j]. \quad (3.41)$$

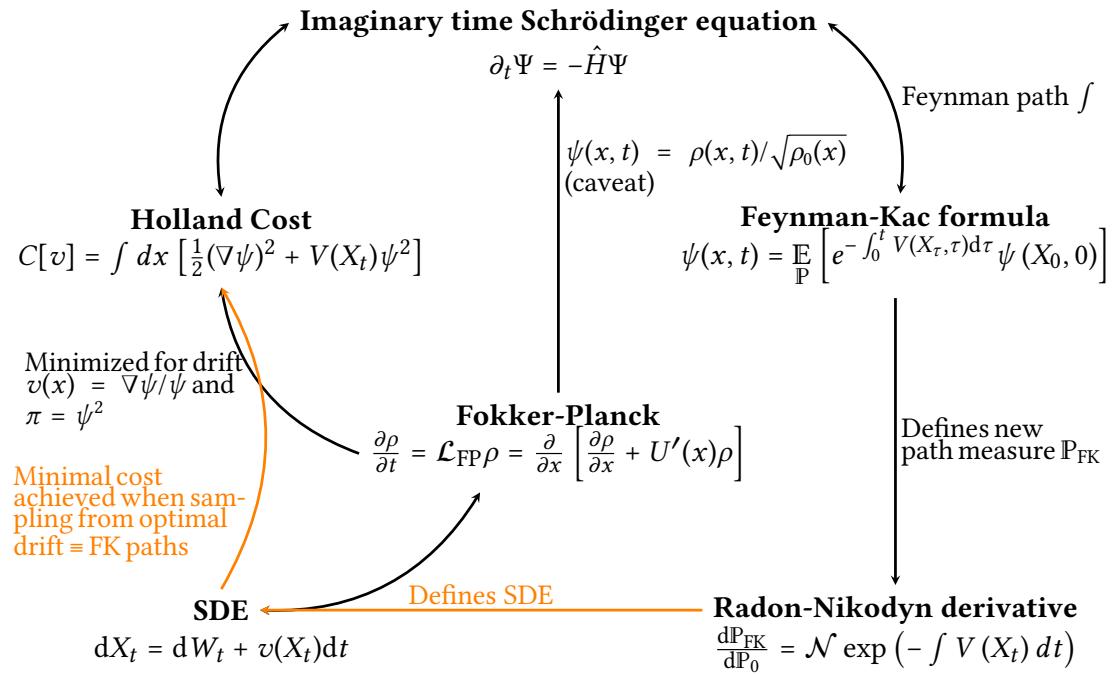


Fig. 3.6 QM, stochastic processes and optimal control

## 3.4 Quantum Mechanics, Control and loss functions

### 3.4.1 Continuous space

### 3.4.2 Discrete space



# **Chapter 4**

## **Machine Learning**

### **4.1 Overview of ML approaches to the Quantum many-body problem**

#### **4.1.1 Neural Network Ansatzes**

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### **4.2 Neural Networks**

#### **4.2.1 Convolutional Neural Networks**

### **4.3 Generative models**

#### **4.3.1 Connection between our approach and Score-Based models**



# Chapter 5

## Methodology

### 5.1 Monte Carlo Importance Sampling

The most common application of Monte Carlo methods is evaluation of integrals in high dimensional space. There MC has a distinct advantage over quadrature methods, as the statistical error decreases with the square root of samples irregardless of the dimensionality of the problem. Integrals of a function  $g(\mathbf{R})$

$$I = \int g(\mathbf{R})d\mathbf{R}, \quad (5.1)$$

where  $\mathbf{R}$  is the *configuration* of the system or simply a *walker*, can be integrated by use of an *importance function*  $P(\mathbf{R})$ , where  $\int d\mathbf{R}P(\mathbf{R}) = 1$  and  $P(\mathbf{R}) \geq 0$ . The integral can be rewritten in the form

$$\int g(\mathbf{R})d\mathbf{R} = \int \frac{g(\mathbf{R})}{P(\mathbf{R})}P(\mathbf{R})d\mathbf{R} = \int f(\mathbf{R})P(\mathbf{R})d\mathbf{R}, \quad (5.2)$$

where  $f(\mathbf{R}) = g(\mathbf{R})/P(\mathbf{R})$ . The importance function  $P(\mathbf{R})$  can be interpreted as a probability density. If we generate an infinite number of random uncorrelated configurations  $\mathbf{R}_m$  from the distribution  $P(\mathbf{R})$ , the sample average is a good estimator of the integral  $I$

$$I = \lim_{M \rightarrow \infty} \left\{ \frac{1}{M} \sum_{m=1}^M f(\mathbf{R}_m) \right\}, \quad (5.3)$$

and for an approximation with a finite number of samples

$$I \approx \frac{1}{M} \sum_{m=1}^M f(\mathbf{R}_m). \quad (5.4)$$

Under conditions where the central limit theorem holds [15], the estimator is normally distributed with variance  $\sigma_f^2/M$ , which can also be estimated from the samples as

$$\frac{\sigma_f^2}{M} \approx \frac{1}{M(M-1)} \sum_{m=1}^M \left[ f(\mathbf{R}_m) - \frac{1}{M} \sum_{n=1}^M f(\mathbf{R}_n) \right]^2. \quad (5.5)$$

## 5.2 Metropolis-Hastings Algorithm

The integration technique from the previous section relies on our ability to obtain samples from a probability distribution  $P(\mathbf{R})$ . In the case of QMC these distributions are high-dimensional and cannot be directly sampled from. Moreover their normalisations are usually not known. The Metropolis-Hastings algorithm [18], see Algorithm 1, avoids direct sampling from the distribution  $P(\mathbf{R})$  and is insensitive to its normalisation. It uses a Markov process whose stationary distribution  $\pi(\mathbf{R})$  is the same as  $P(\mathbf{R})$  to generate a sequence of configurations  $\{\mathbf{R}_n\}_P$  that are drawn from  $P(\mathbf{R})$ . A Markov process is completely defined with its transition probability  $P(\mathbf{R} \rightarrow \mathbf{R}')$ , which is the probability of transitioning from state  $\mathbf{R}$  to state  $\mathbf{R}'$ . For the process to have a unique stationary distribution two conditions must be met, the process must be *ergodic* and it must obey *detailed balance*

$$P(\mathbf{R})P(\mathbf{R} \rightarrow \mathbf{R}') = P(\mathbf{R}')P(\mathbf{R}' \rightarrow \mathbf{R}), \quad (5.6)$$

rewritten as

$$\frac{P(\mathbf{R})}{P(\mathbf{R}')} = \frac{P(\mathbf{R}' \rightarrow \mathbf{R})}{P(\mathbf{R} \rightarrow \mathbf{R}')}. \quad (5.7)$$

The right transition probability  $P(\mathbf{R} \rightarrow \mathbf{R}')$  is not known, but we can express it with a trial move transition probability  $T(\mathbf{R} \rightarrow \mathbf{R}')$  which we sample and acceptance probability  $A(\mathbf{R} \rightarrow \mathbf{R}')$  as

$$P(\mathbf{R} \rightarrow \mathbf{R}') = T(\mathbf{R} \rightarrow \mathbf{R}')A(\mathbf{R} \rightarrow \mathbf{R}'). \quad (5.8)$$

For equation (5.7) to hold, the acceptance probability must be

$$A(\mathbf{R} \rightarrow \mathbf{R}') = \min \left( 1, \frac{T(\mathbf{R}' \rightarrow \mathbf{R})P(\mathbf{R}')}{T(\mathbf{R} \rightarrow \mathbf{R}')P(\mathbf{R})} \right). \quad (5.9)$$

Thus to sample from any probability distribution we need only have the ability to calculate probabilities  $P(\mathbf{R})$  and to sample from a trial transition probability  $T(\mathbf{R} \rightarrow \mathbf{R}')$ . The efficiency of the algorithm depends on the amount of trial moves that we reject. All trial moves would

be accepted if  $T(R \rightarrow R') = P(R')$ , which would just mean sampling from  $P$  directly and is the very problem we are trying to solve with Metropolis-Hastings.

---

**Algorithm 1:** Metropolis-Hastings

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**Result:** A set of configurations  $\{R_n\}_P$  sampled from  $P$

Initialize walker at random configuration  $R$ ;

**while** no. samples less than  $N$  **do**

- Generate new configuration  $R'$  with transition probability  $T(R \rightarrow R')$ ;
- Accept the move  $(R \rightarrow R')$  with probability
- $A(R \rightarrow R') = \min\left(1, \frac{T(R' \rightarrow R)P(R)}{T(R \rightarrow R')P(R')}\right)$ ;
- Append  $R$  to the set of configuration;

**end**

---

## 5.3 Gradient based optimisation

### 5.3.1 Gradient estimation

- **TODO:** Rewrite with a more general tone, do not only talk about the ELBO. Still use Mohammeds review! It is very good.

In order to perform gradient descent on the ELBO objective, we need to be able to evaluate its gradients with respect to parameters  $\theta$  and  $\phi$ . Taking the gradient w.r.t generative parameters  $\theta$  is straightforward, because we can change the order of the expectation operator and the gradient, leaving us with

$$\begin{aligned} \nabla_{\theta} \mathcal{L}_{\theta, \phi}(x) &= \nabla_{\theta} \mathbb{E}_{q_{\phi}(z|x)} [\log p_{\theta}(x, z) - \log q_{\phi}(z|x)] \\ &\simeq \nabla_{\theta} (\log p_{\theta}(x, z) - \log q_{\phi}(z|x)) \\ &= \nabla_{\theta} (\log p_{\theta}(x, z)), \end{aligned} \tag{5.10}$$

where  $\simeq$  denotes an unbiased estimator. This reversing of the order of operations is not possible when taking gradients w.r.t variational parameters  $\phi$  because the expectation  $\mathbb{E}_{q_{\phi}(z|x)}$  is performed w.r.t the approximate posterior  $q_{\phi}(z|x)$ . The gradient could be estimated with a vanilla Monte Carlo estimator, but it has very high variance and is not practical [27].

The problem of stochastic gradient estimation of an expectation of a function is a well studied problem that transcends machine learning and has a variety of applications [11, 40]. Different estimators differ in from and their properties, variance being one of the most important. In their review [30] Mohamed et al. categorise MC gradient estimators into three categories

### Score-function estimator

*Score-function estimator:* The score function is a logarithm of a probability distribution w.r.t to distributional parameters. It can be used as a gradient estimator

$$\begin{aligned}\nabla_{\theta} \mathbb{E}_{p_{\theta}(\mathbf{x})}[f(\mathbf{x})] &= \nabla_{\theta} \int p_{\theta}(\mathbf{x}) f(\mathbf{x}) d\mathbf{x} \\ &= \mathbb{E}_{p_{\theta}(\mathbf{x})}[f(\mathbf{x}) \nabla_{\theta} \log p_{\theta}(\mathbf{x})].\end{aligned}\tag{5.11}$$

The score-function estimator is compatible with any cost function, it requires that the measure  $p_{\theta}(\mathbf{x})$  is differentiable and easy to sample. Very importantly it is applicable to both discrete and continuous distribution, but has a drawback of having high variance.

### Pathwise estimator

Continuous distributions can be sampled either directly by generating samples from the distribution  $p_{\theta}(\mathbf{x})$  or indirectly, by sampling from a simpler base distribution  $p(\epsilon)$  and transforming the variate through a deterministic path  $g_{\theta}(\epsilon)$ . Using this, it is possible to move the source of randomness in such a way that the objective is differentiable. In essence this approach pushes the parameters of the measure into the cost function which is then differentiated. The estimator is

$$\begin{aligned}\nabla_{\theta} \mathbb{E}_{p_{\theta}(\mathbf{x})}[f(\mathbf{x})] &= \nabla_{\theta} \int p_{\theta}(\mathbf{x}) f(\mathbf{x}) d\mathbf{x} \\ &= \nabla_{\theta} \int p(\epsilon) f(g_{\theta}(\epsilon)) d\epsilon \\ &= \mathbb{E}_{p(\epsilon)}[\nabla_{\theta} f(g_{\theta}(\epsilon))].\end{aligned}\tag{5.12}$$

This was the gradient estimator originally used in the VAE implementation [27] there named as the *reparametrization trick*, see also Figure 5.1. In many cases the transformation paths are so simple they can be implemented in one line of code, referred to as *one-liners*. The pathwise-estimator can only be used on differentiable cost functions, but is easy to implement and crucially has lower variance than the score-function estimator.

### Measure-valued gradient estimator

Which exploits the properties of signed-measures, is beyond the scope of this report.

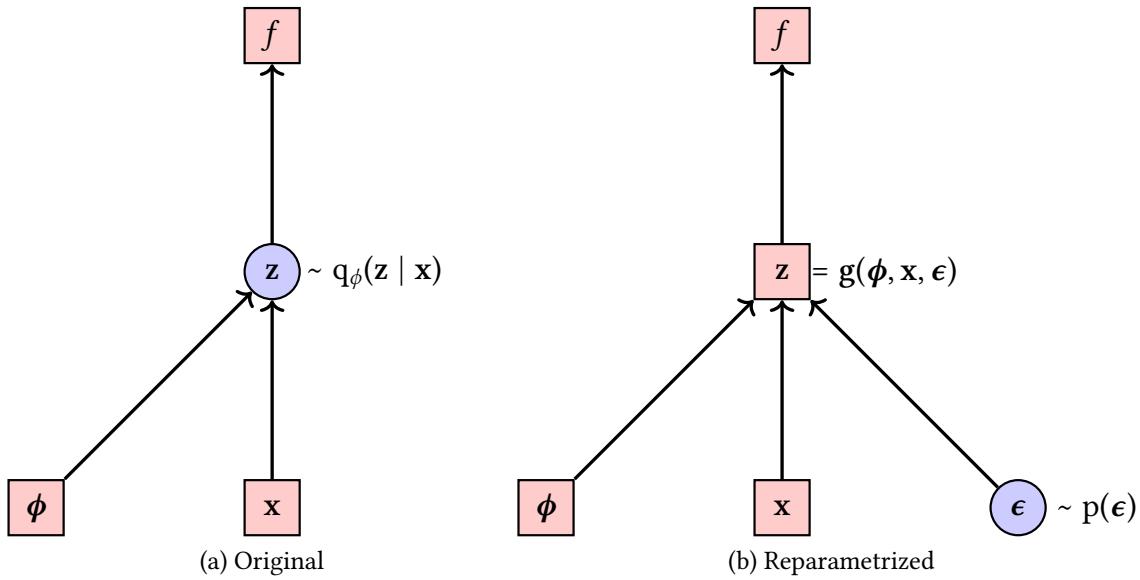


Fig. 5.1 The reparametrization trick, adapted from [26]. The stochasticity of the  $z$  node is pushed out into a separate input to the same node, resulting in deterministic gradients w.r.t  $\phi$  through the node.

### 5.3.2 Automatic differentiation

## 5.4 *Optimal sampling*: optimal sampling in lattice models



# Chapter 6

## Results

### 6.1 Single Particle on a Lattice

$$\partial_t \psi_j = \frac{1}{2} [\psi_{j+1} + \psi_{j-1} - 2\psi_j] + V_j \psi_j \quad (6.1)$$

$$\psi_j(t) = \mathbb{E}_{X \sim \text{SRW} \text{ with } X_t=j} \left[ \exp \left( - \int_0^t V(X_\tau, \tau) d\tau \right) \psi_{X_0}(0) \right] \quad (6.2)$$

### 6.2 Transverse-field Ising model

### 6.3 Heisenberg model

Heisenberg ferromagnet

$$\hat{H}_F = -\frac{1}{2} \sum_j [\hat{\sigma}_j^x \hat{\sigma}_{j+1}^x + \hat{\sigma}_j^y \hat{\sigma}_{j+1}^y + \hat{\sigma}_j^z \hat{\sigma}_{j+1}^z] \quad (6.3)$$

The XY model.

$$\begin{aligned} \hat{H}_{XY} &= - \sum_j [\hat{\sigma}_j^x \hat{\sigma}_{j+1}^x + \hat{\sigma}_j^y \hat{\sigma}_{j+1}^y] = H_F + \frac{1}{2} \sum_j \hat{\sigma}_j^z \hat{\sigma}_{j+1}^z \\ &= -\mathcal{W} + \sum_j [n_j (1 - n_{j+1}) + n_{j+1} (1 - n_j)] \end{aligned} \quad (6.4)$$

$$\psi_{s_{1:N}}(t) = \underset{\Sigma_{[0,t]} \sim \text{SEP}}{\mathbb{E}} \left[ \exp \left( - \int_0^t dt' \sum_j [n_j(1-n_{j+1}) + n_{j+1}(1-n_j)] \right) \psi_{\Sigma_0}(0) \right] \quad (6.5)$$

## 6.4 Bose-Hubbard model

# **Chapter 7**

## **Conclusions**

**7.1 Direction for further work**

**7.2 Remarks**



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# **Appendix A**

## **Hyperparameters**



## **Appendix B**

### **Fixed-Node Feynman-Kac formula**



## **Appendix C**

### **Additional results**



## **Appendix D**

### **Additional Derivations**

